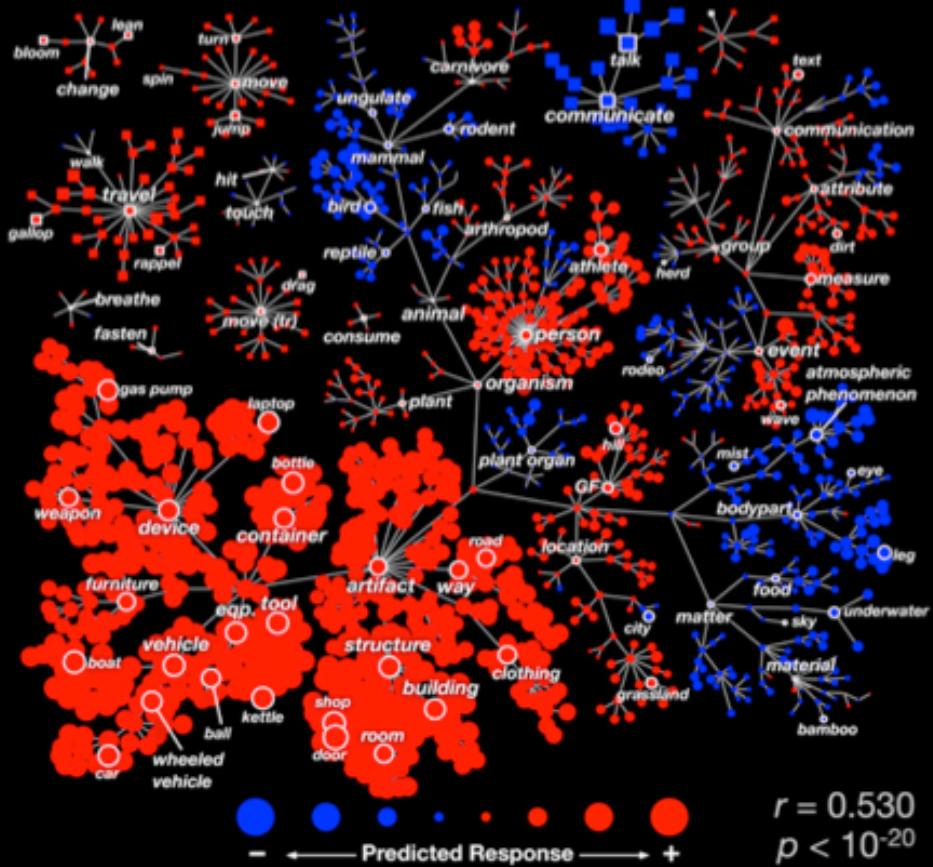
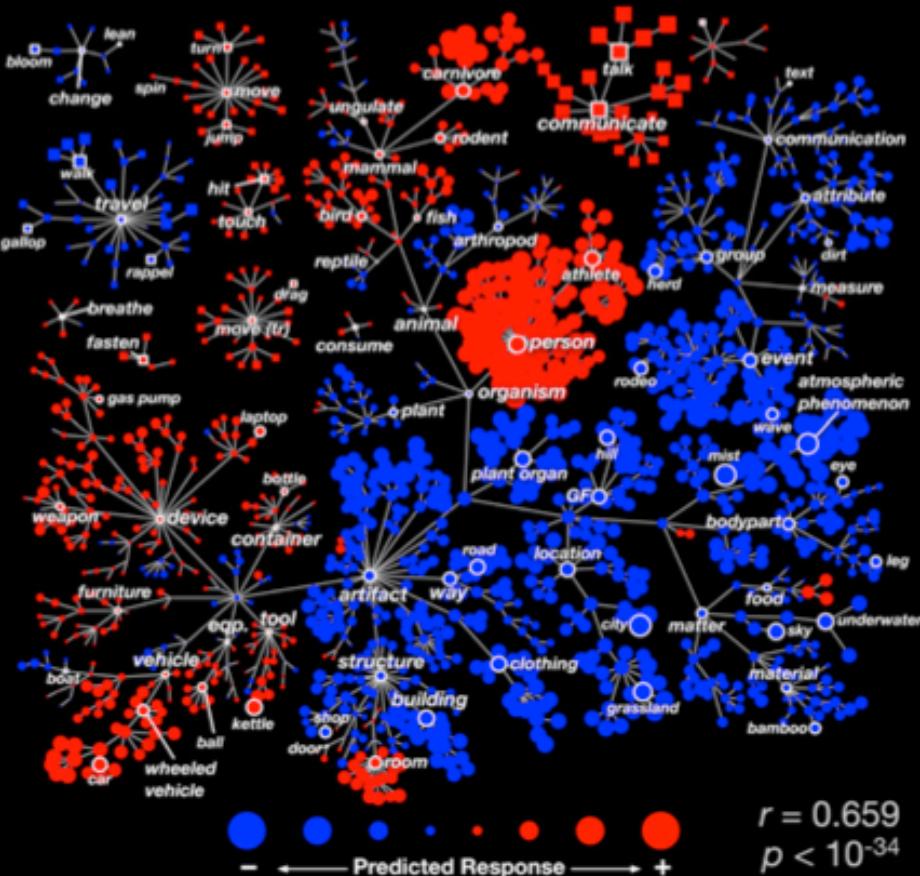
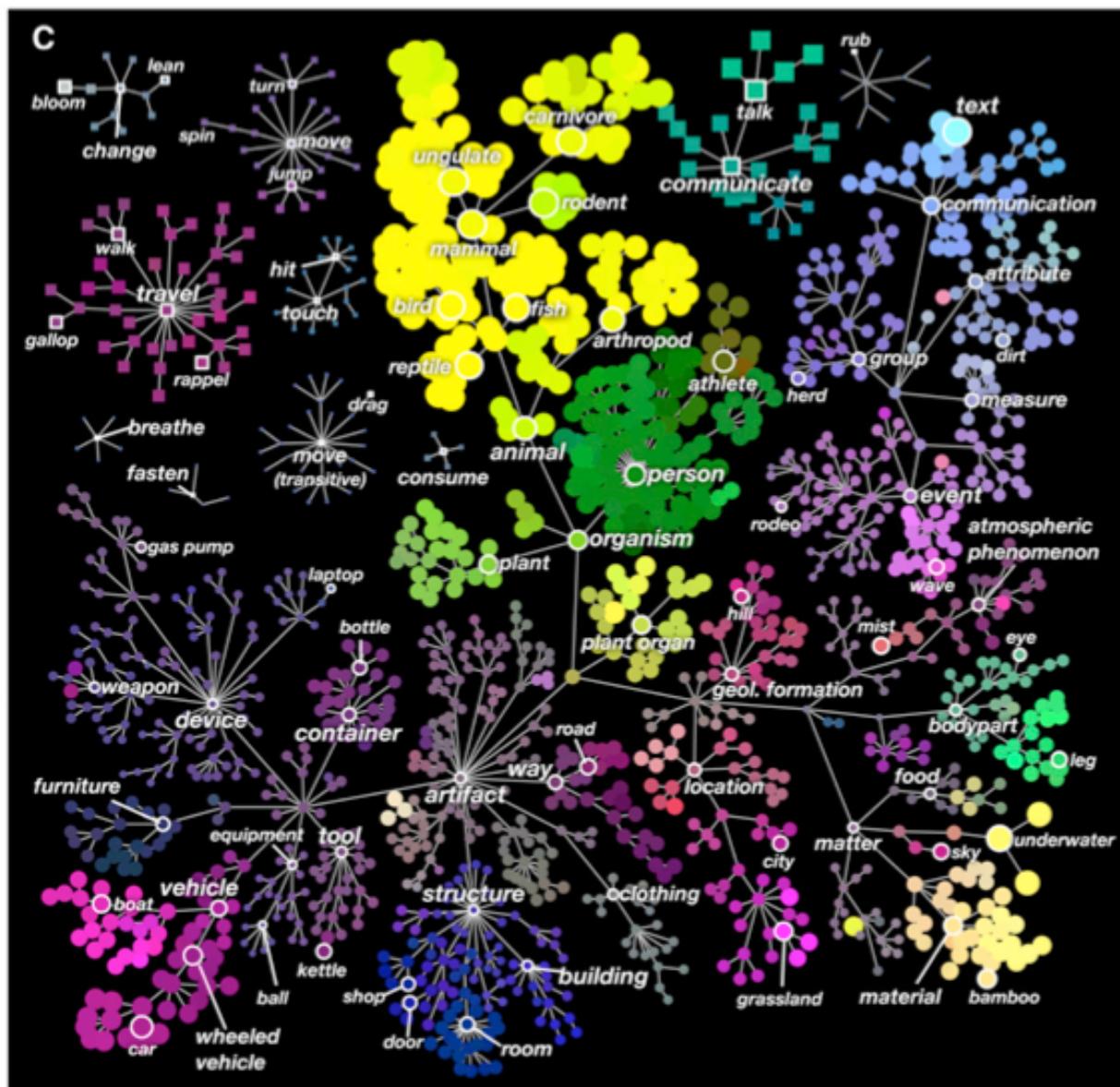
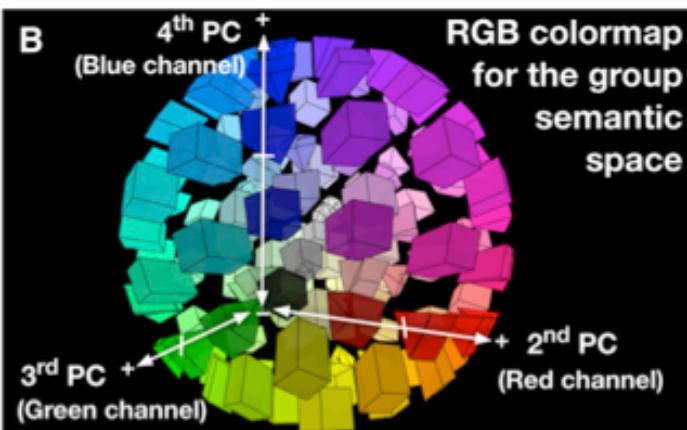
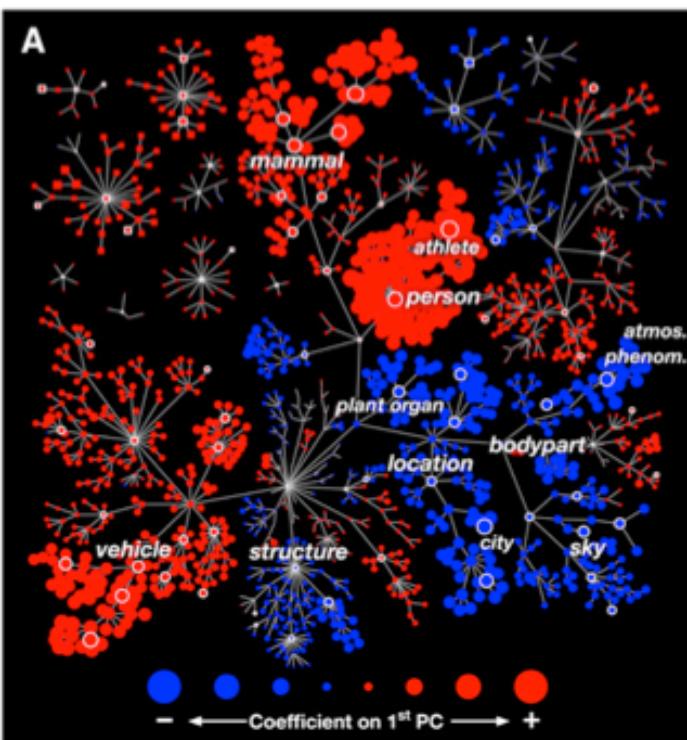
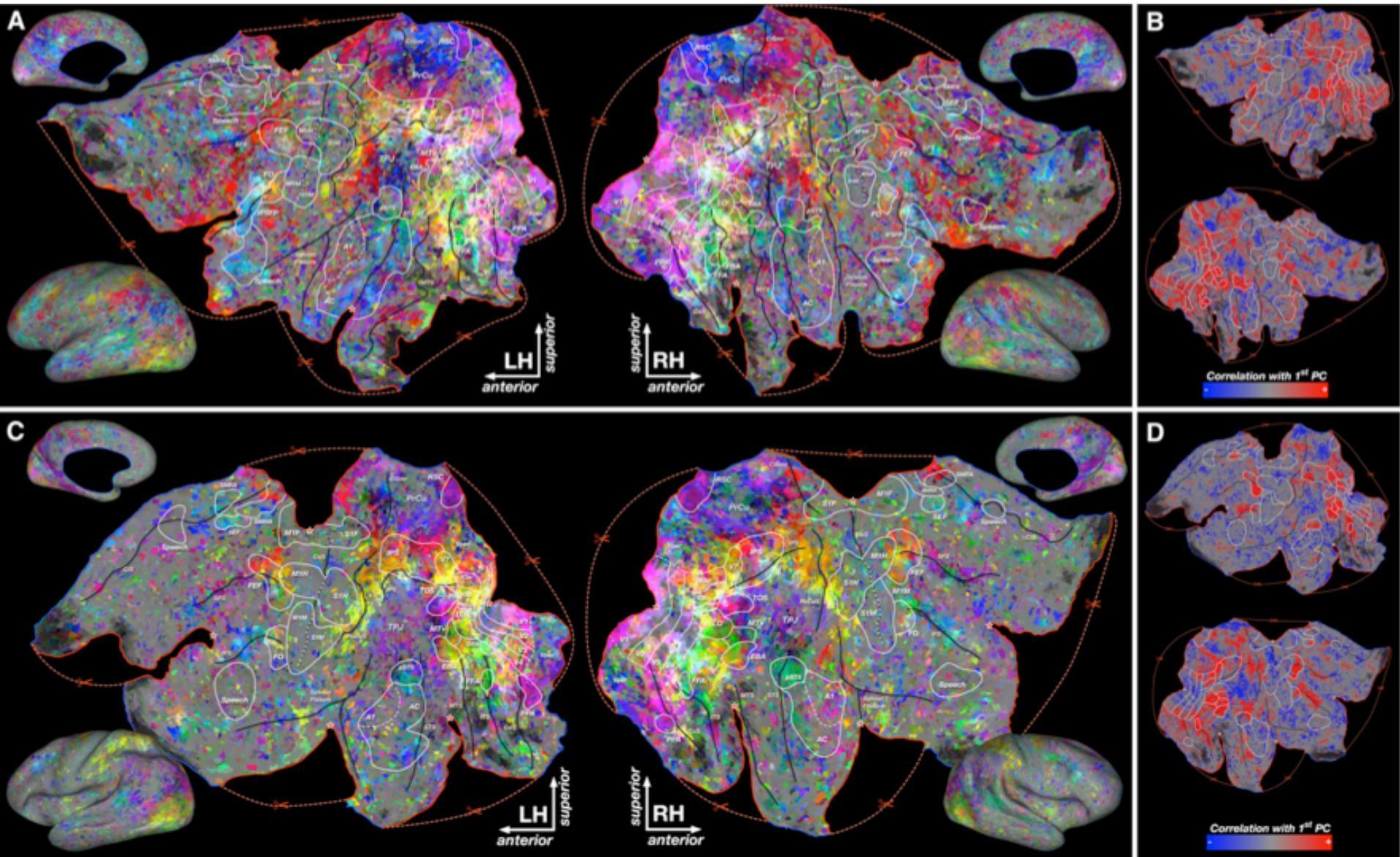


A**Voxel AV-8592 (left PPA)****B****Voxel AV-19987 (right precuneus)**





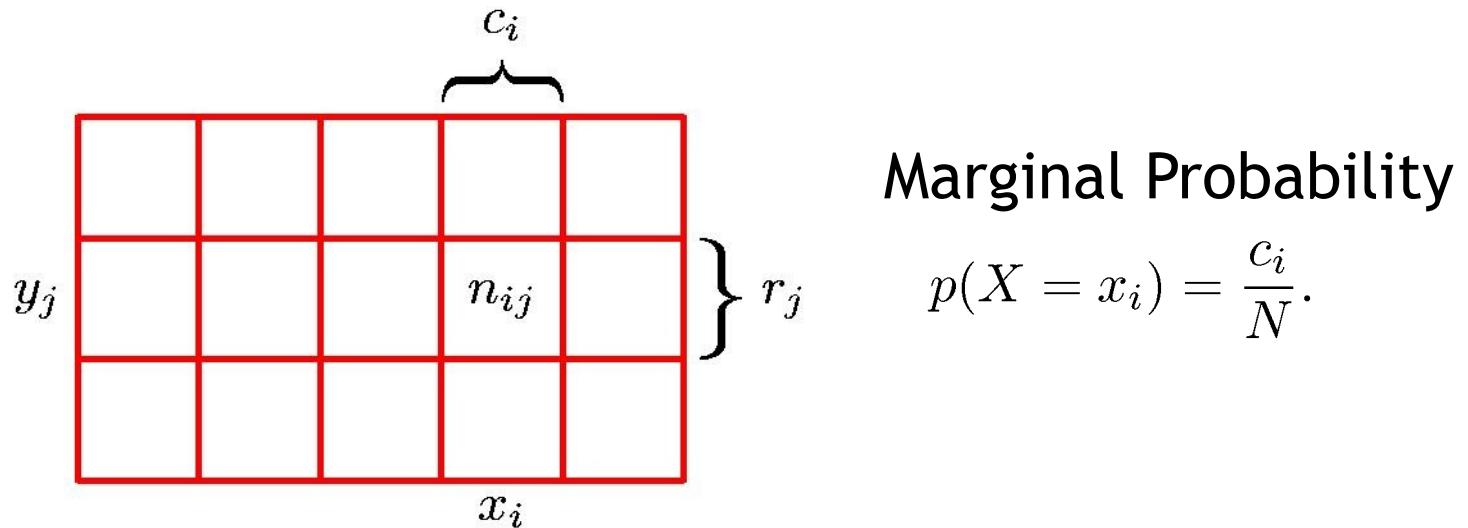
MACHINE LEARNING

PROBABILITY THEORY

LECTURE 4: BASICS, BAYES, DISTRIBUTIONS

Probability Theory

$$p(X, Y)$$



Joint Probability

$$p(X = x_i, Y = y_j) = \frac{n_{ij}}{N}$$

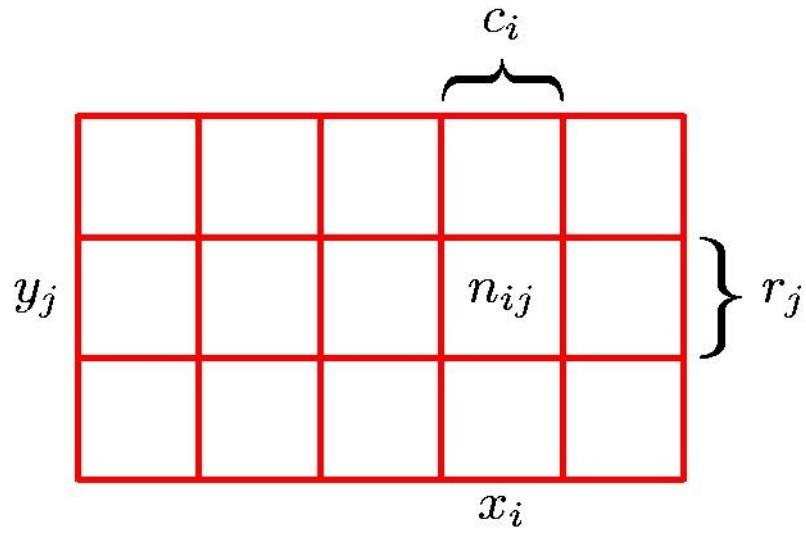
Marginal Probability

$$p(X = x_i) = \frac{c_i}{N}.$$

Conditional Probability

$$p(Y = y_j | X = x_i) = \frac{n_{ij}}{c_i}$$

Probability Theory



Sum Rule

$$p(X = x_i) = \frac{c_i}{N} = \frac{1}{N} \sum_{j=1}^L n_{ij}$$
$$= \sum_{j=1}^L p(X = x_i, Y = y_j)$$

Product Rule

$$\begin{aligned} p(X = x_i, Y = y_j) &= \frac{n_{ij}}{N} = \frac{n_{ij}}{c_i} \cdot \frac{c_i}{N} \\ &= p(Y = y_j | X = x_i) p(X = x_i) \end{aligned}$$

The Rules of Probability

Sum Rule

$$p(X) = \sum_Y p(X, Y)$$

Product Rule

$$p(X, Y) = p(Y|X)p(X)$$

Bayes' Theorem

$$p(Y|X) = \frac{p(X|Y)p(Y)}{p(X)}$$

$$p(X) = \sum_Y p(X|Y)p(Y)$$

posterior \propto likelihood \times prior

Example Problem 1

1. (Kohler) Bad news at the doctor's

After your yearly checkup the doctor has bad news. You tested positive for a serious disease and the test is 99% accurate. That is, the probability that the test result is positive given that you have the disease is 0.99. The test is very accurate: the probability that you do not have the disease given the test result is negative is also 0.99.

The only good news is that the disease is rare, striking only one in 10,000 people. What are your chances given this information?

Solution

Let's let T^+ stand for the test being positive and T^- stand for the test being negative.

Similarly, let's let D^+ stand for having the disease and D^- stand for not having the disease.

What we want to know is $P(D^+|T^+)$, the probability that we actually have the disease given that the test was positive.

Solution

Let's let T^+ stand for the test being positive and T^- stand for the test being negative.

Similarly, let's let D^+ stand for the having the disease and D^- stand for not having the disease.

What we want to know is $P(D^+|T^+)$, the probability that we actually have the disease given that the test was positive. Using Bayes rule:

$$P(D^+|T^+) = \frac{P(T^+|D^+)(P(D^+))}{P(T^+)},$$

$$= \frac{P(T^+|D^+)(P(D^+))}{P(T^+, D^+) + P(T^+, D^-)},$$

$$\frac{P(T^+|D^+)(P(D^+))}{P(T^+|D^+)P(D^+) + P(T^+|D^-)P(D^-)},$$

$$= \frac{1}{1 + \frac{P(T^+|D^-)P(D^-)}{P(T^+|D^+)P(D^+)}},$$

$$= \frac{1}{1 + \frac{(0.01)(0.9999)}{0.99} \cdot (10^{-4})},$$

Example Problem 2

2. (MacCay) Monte Hall On a game show a contestant is told the rules as follows:

There are three doors labeled 1, 2 and 3. A single prize has been hidden behind one of them. You get to select one door. Initially your chosen door will *not* be opened. Instead the gameshow host will open one of the two remaining doors and will do so in a way so as to not reveal the prize.

At this point you will be given a fresh choice of door. You can either stick to your choice or choose the remaining unopened door.

Should the contestant stick with the originally chosen door or switch? Or does it matter?

Let H_i represent the hypothesis that the prize is behind door i .

And without loss of generality let's assume the contestant chooses door 1 and the gameshow host opens door 3.

A priori we must assume that

$$P(H_i) = \frac{1}{3}, i = 1, 2, 3$$

The key observation is that if the contestant chose a door with the prize, then the host can choose either of the remaining doors, but if the contestant *didn't* choose the prize then the host is restricted in choosing the door to open. Where D is the door the host is thinking of picking, these options can be expressed as;

$$P(D = 2|H_1) = \frac{1}{2},$$

$$P(D = 3|H_1) = \frac{1}{2},$$

$$P(D = 2|H_2) = 0,$$

$$P(D = 3|H_2) = 1,$$

$$P(D = 2|H_3) = 1,$$

$$P(D = 3|H_3) = 0$$

Now for Bayes rule:

$$P(H_i|D = 3) = \frac{P(D = 3|H_i)p(H_i)}{P(D = 3)}$$

So, enumerating the options,

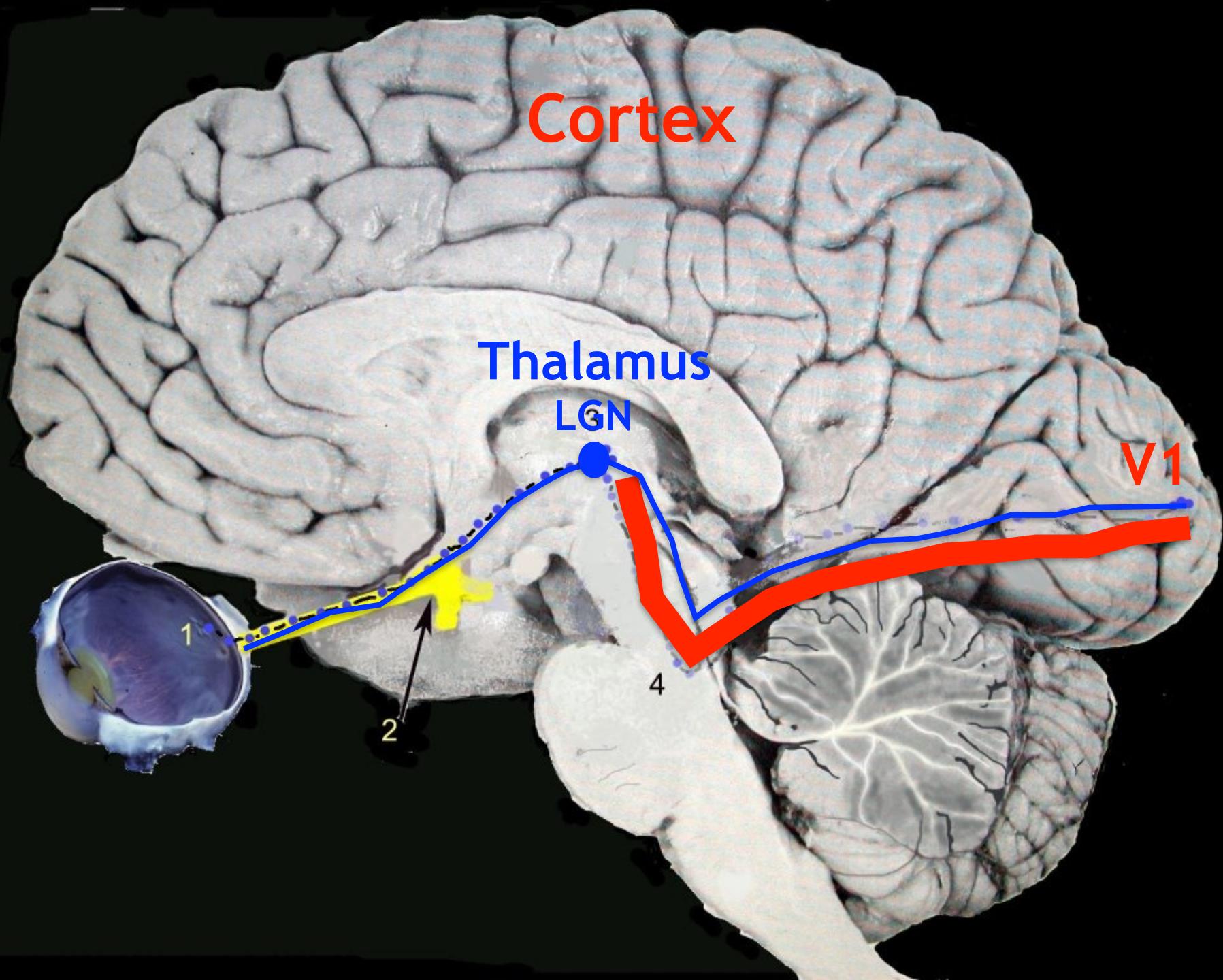
$$P(H_1|D = 3) = \frac{\frac{1}{2} \times \frac{1}{3}}{P(D = 3)}$$

$$P(H_2|D = 3) = \frac{1 \times \frac{1}{3}}{P(D = 3)},$$

$$P(H_3|D = 3) = \frac{0 \times \frac{1}{3}}{P(D = 3)}$$

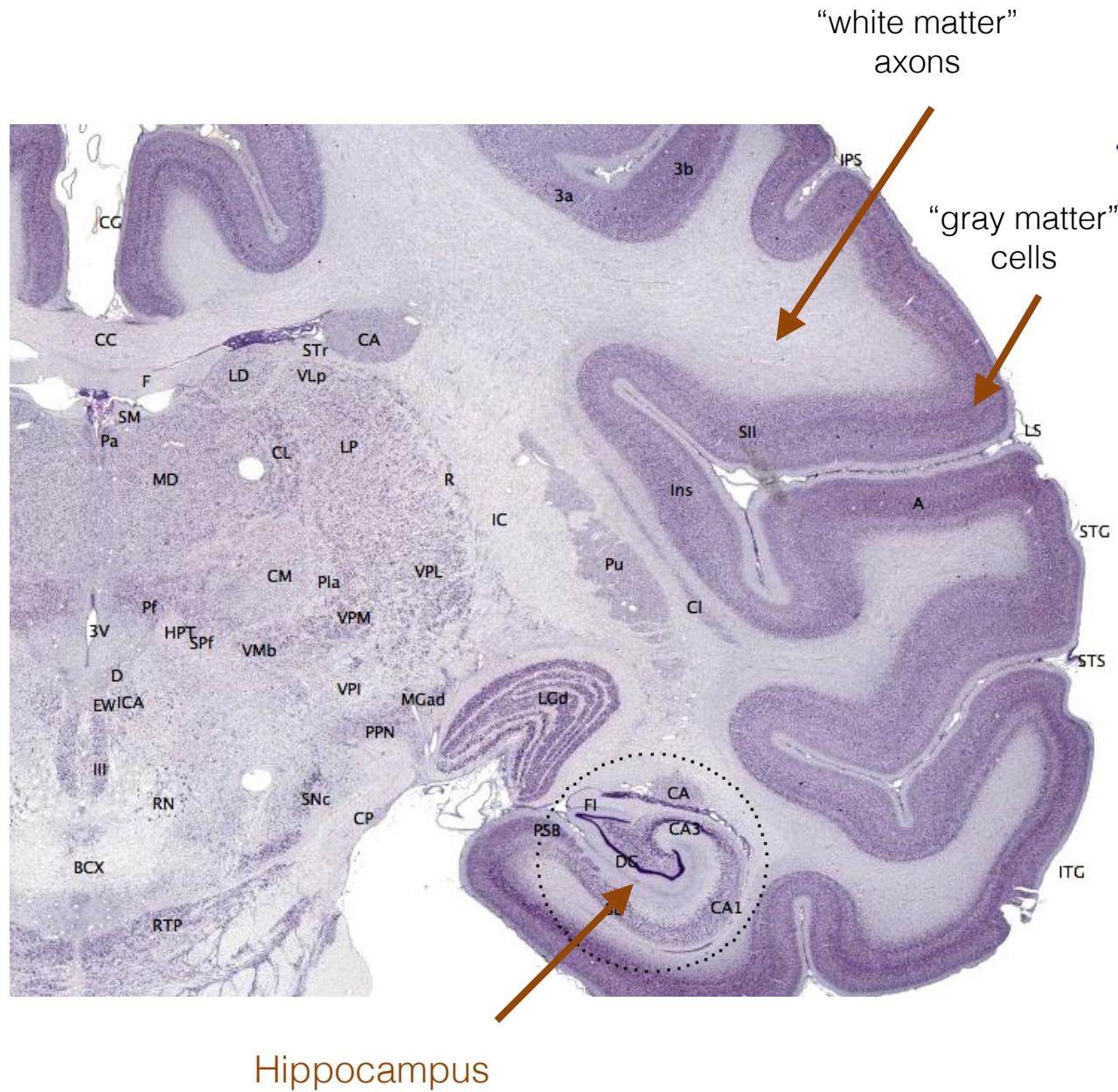
By inspection, after door 3 has been opened, H_2 is twice as likely as H_1 , so the contestant should switch.

What is $P(D=3)$?



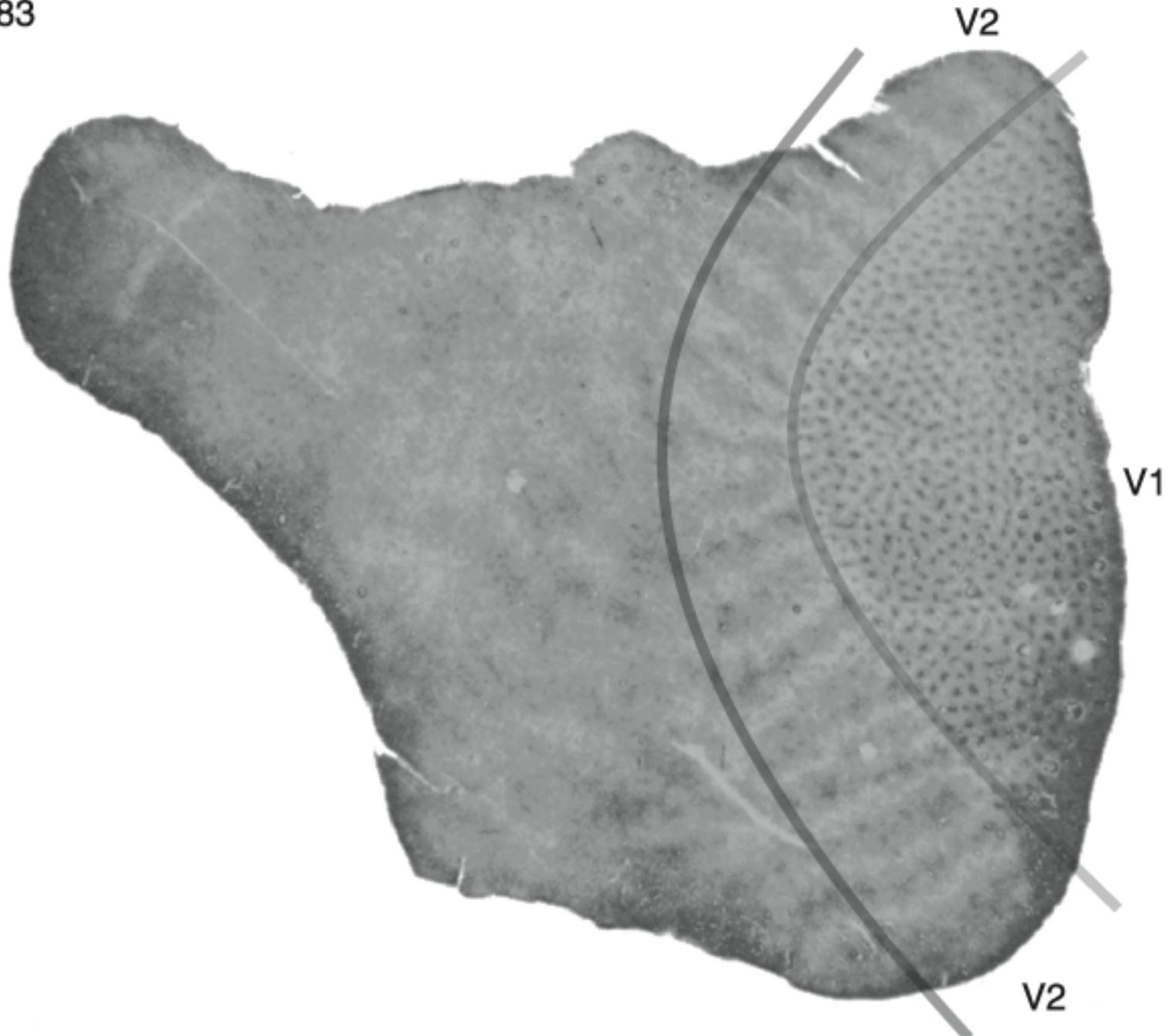
A slice through the macaque brain cortex

The cortex has been likened to a “Pizza” in that the cells are located in a thin outer layer 6 credit cards thick. This slice shows the cortical layers terminating in the Hippocampus, which has a critical role in the creation of new memories.

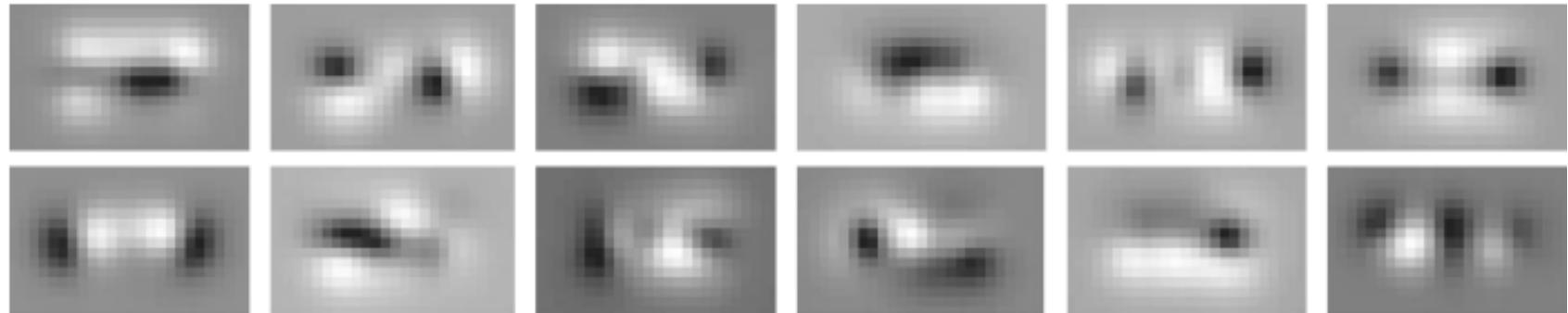


Cytochrome Oxidase in Monkey Visual Cortex

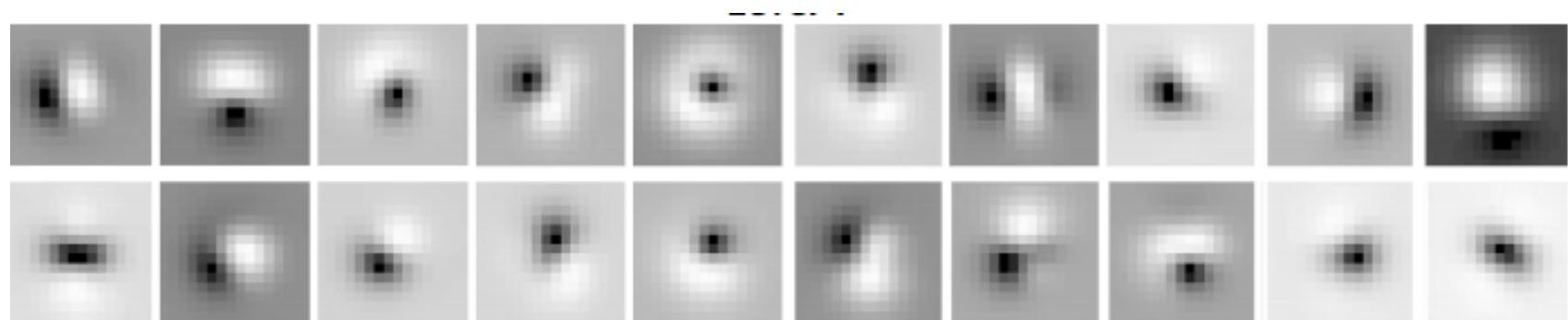
Tootell et al 1983



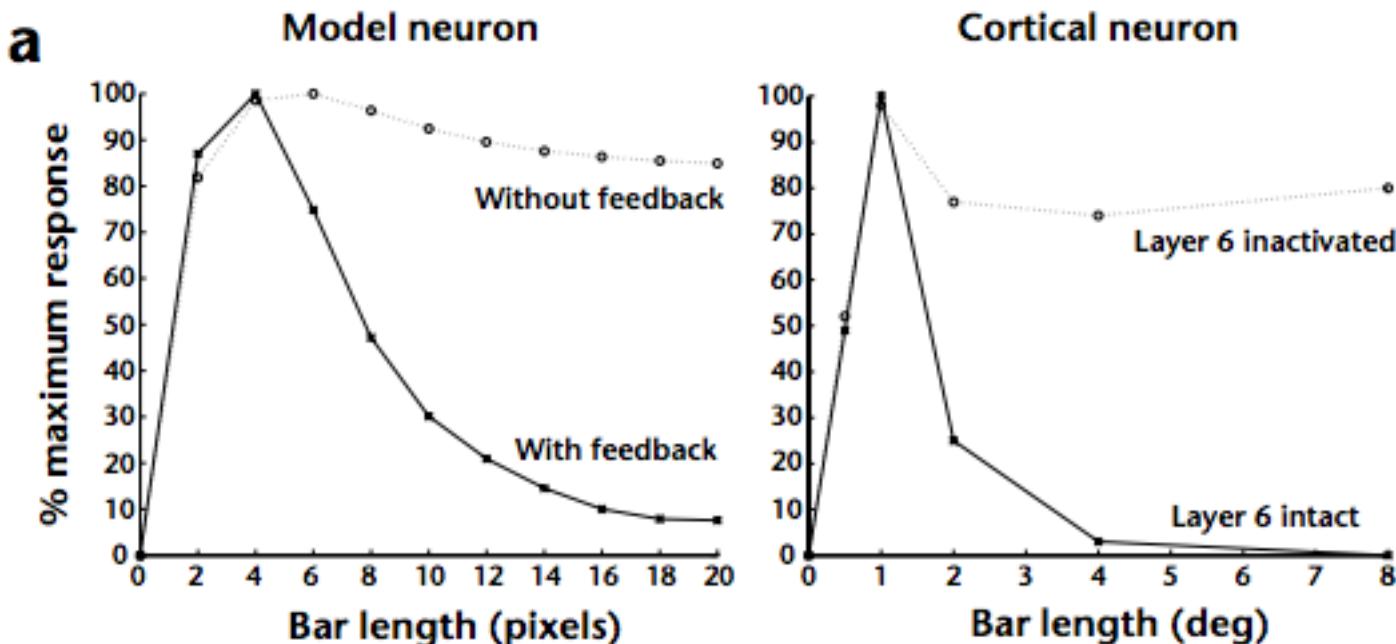
V2



V1

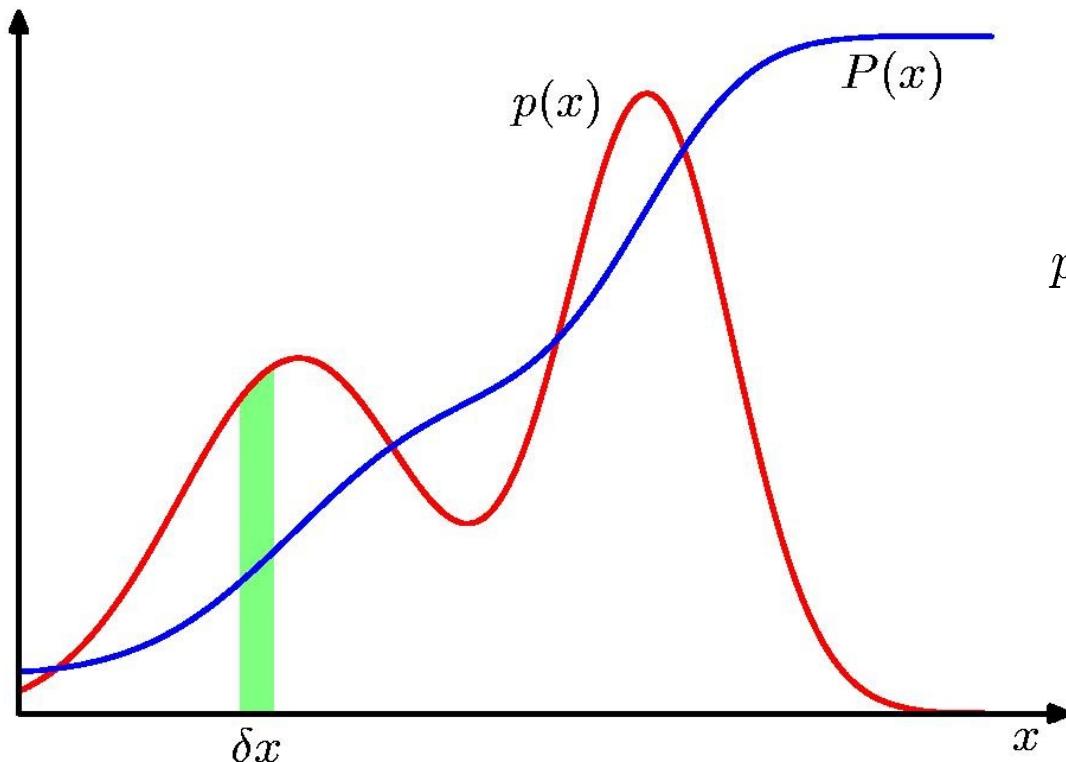


$$P(V1, V2) = P(V1|V2)P(V2)$$



Rao, R., Ballard, D. Predictive coding in the visual cortex:
a functional interpretation of some extra-classical receptive-field effects.
Nat Neurosci **2**, 79–87 (1999)

Probability Densities



$$p(x \in (a, b)) = \int_a^b p(x) dx$$

$$P(z) = \int_{-\infty}^z p(x) dx$$

$$p(x) \geq 0$$

$$\int_{-\infty}^{\infty} p(x) dx = 1$$

Expectations

$$\mathbb{E}[f] = \sum_x p(x)f(x)$$

$$\mathbb{E}[f] = \int p(x)f(x) dx$$

$$\mathbb{E}_x[f|y] = \sum_x p(x|y)f(x)$$


Conditional Expectation
(discrete)

$$\mathbb{E}[f] \simeq \frac{1}{N} \sum_{n=1}^N f(x_n)$$

Approximate Expectation
(discrete and continuous)

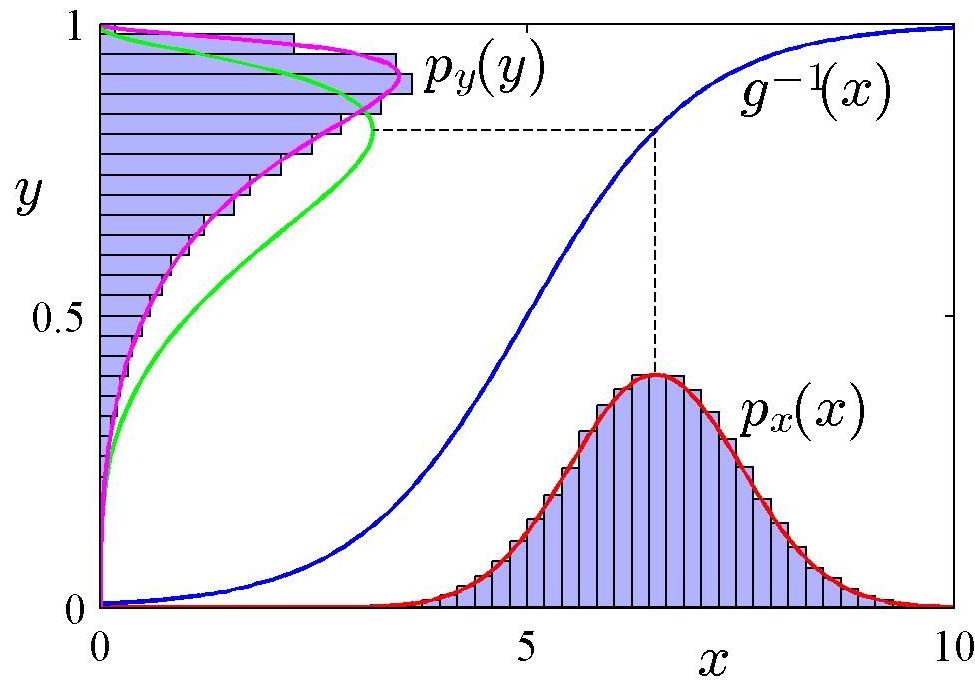
Variances and Covariances

$$\text{var}[f] = \mathbb{E} \left[(f(x) - \mathbb{E}[f(x)])^2 \right] = \mathbb{E}[f(x)^2] - \mathbb{E}[f(x)]^2$$

$$\begin{aligned}\text{cov}[x, y] &= \mathbb{E}_{x,y} [\{x - \mathbb{E}[x]\} \{y - \mathbb{E}[y]\}] \\ &= \mathbb{E}_{x,y}[xy] - \mathbb{E}[x]\mathbb{E}[y]\end{aligned}$$

$$\begin{aligned}\text{cov}[\mathbf{x}, \mathbf{y}] &= \mathbb{E}_{\mathbf{x},\mathbf{y}} [\{\mathbf{x} - \mathbb{E}[\mathbf{x}]\}\{\mathbf{y}^T - \mathbb{E}[\mathbf{y}^T]\}] \\ &= \mathbb{E}_{\mathbf{x},\mathbf{y}}[\mathbf{x}\mathbf{y}^T] - \mathbb{E}[\mathbf{x}]\mathbb{E}[\mathbf{y}^T]\end{aligned}$$

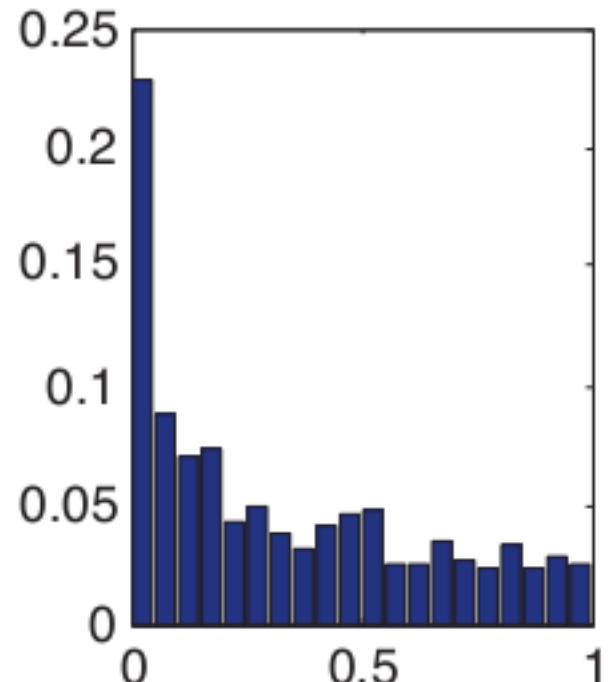
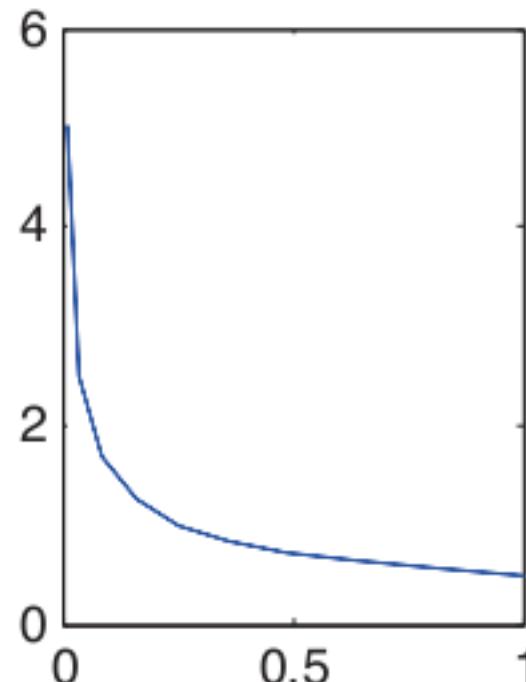
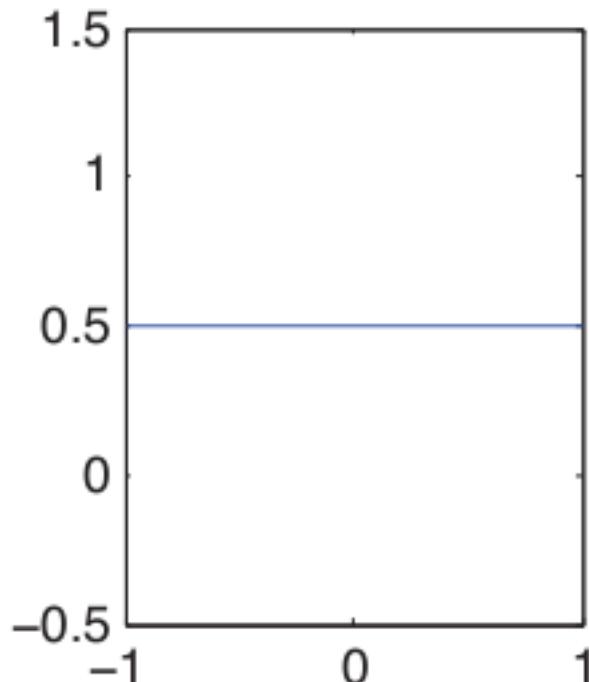
Transformed Densities



$$\begin{aligned} p_y(y) &= p_x(x) \left| \frac{dx}{dy} \right| \\ &= p_x(g(y)) |g'(y)| \end{aligned}$$

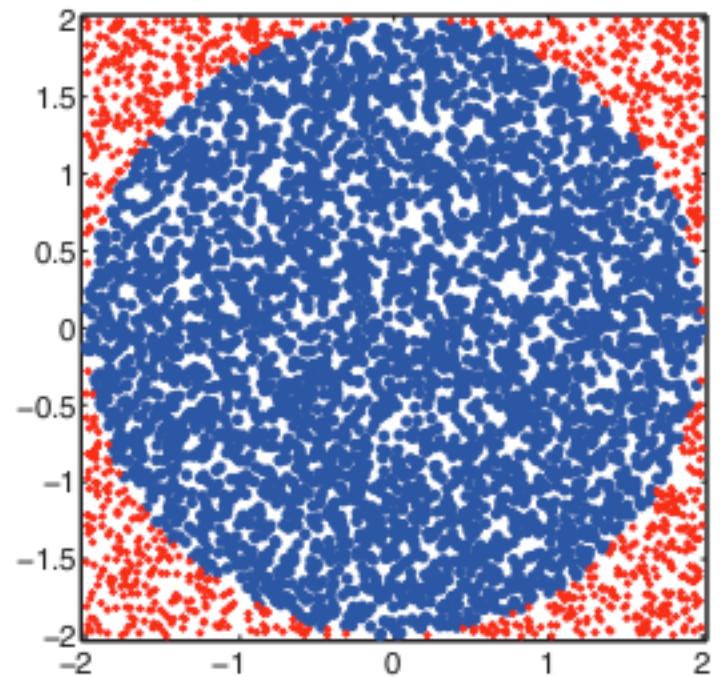
Transformed densities by sampling

e.g. $p(y)$ given $y=x^2$ and $p(x)$ is uniform

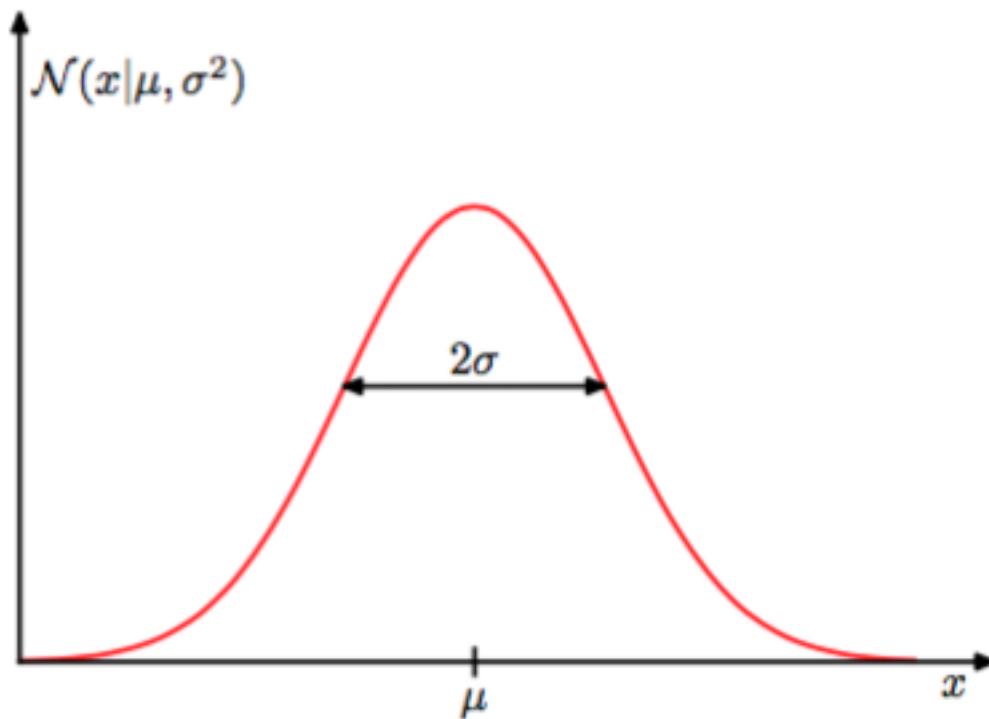


Monte Carlo sampling for π

$$\begin{aligned} I &= (2r)(2r) \int \int f(x, y)p(x)p(y)dxdy \\ &= 4r^2 \int \int f(x, y)p(x)p(y)dxdy \\ &\approx 4r^2 \frac{1}{S} \sum_{s=1}^S f(x_s, y_s) \end{aligned}$$

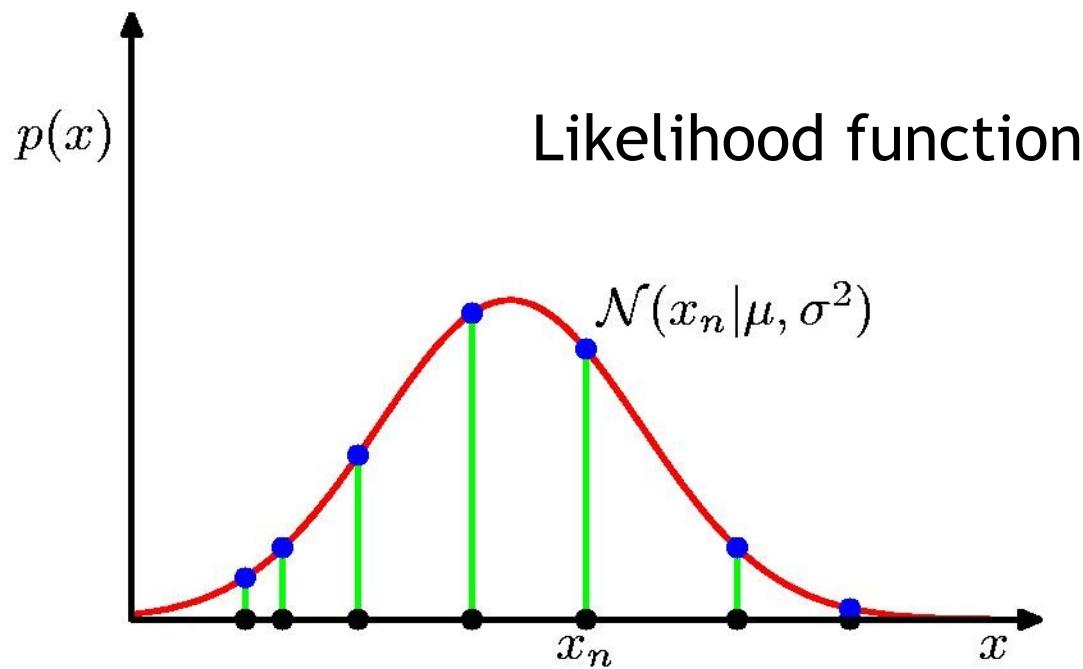


Gaussian (Normal) distribution



$$N(x|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp \left\{ -\frac{1}{2\sigma^2}(x - \mu)^2 \right\}$$

Gaussian Parameter Estimation



$$p(\mathbf{x}|\mu, \sigma^2) = \prod_{n=1}^N \mathcal{N}(x_n|\mu, \sigma^2)$$

Gaussian parameters by Maximum (Log) Likelihood

$$\ln p(\mathbf{x}|\mu, \sigma^2) = -\frac{1}{2\sigma^2} \sum_{n=1}^N (x_n - \mu)^2 - \frac{N}{2} \ln \sigma^2 - \frac{N}{2} \ln(2\pi)$$

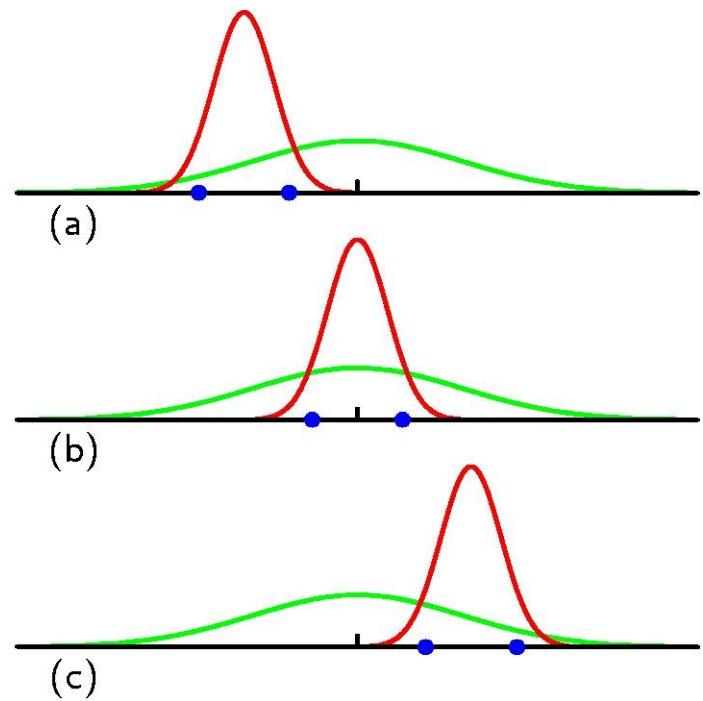
$$\mu_{\text{ML}} = \frac{1}{N} \sum_{n=1}^N x_n \quad \sigma_{\text{ML}}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \mu_{\text{ML}})^2$$

Properties of μ_{ML} and σ_{ML}^2

$$\mathbb{E}[\mu_{\text{ML}}] = \mu$$

$$\mathbb{E}[\sigma_{\text{ML}}^2] = \left(\frac{N-1}{N} \right) \sigma^2$$

$$\begin{aligned}\tilde{\sigma}^2 &= \frac{N}{N-1} \sigma_{\text{ML}}^2 \\ &= \frac{1}{N-1} \sum_{n=1}^N (x_n - \mu_{\text{ML}})^2\end{aligned}$$



End of Lec 4: Basic Probabilities
