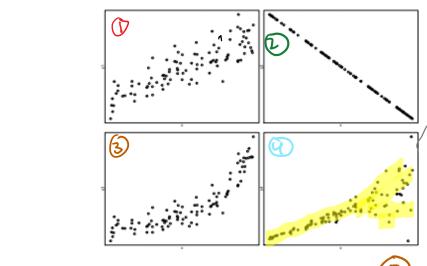
### Preliminary activity II



note: frend in difference does not matter which × you pich has different noise, while others have the same noise,

- Description: increasing/positive - (incar
- 3 Description: -increasing/positive - has noise - curved
- Description: decocasing higgside - no noise
- 1 Description: -in creasing/positive - has noise

Heteroscedasticity: the circumstance in which the variability of a variable is unequal across the range of values of a second variable that predicts it.

# Kegression analysis

- · Statisfical methodology that utilizes the relation between variables.
- · Predicts a response variable (or outcome) from the relation between the response and other variables
- · Regression analysis is used in many diciplines such as:
  - Business'.
    - i) Forecasting: predicting future demand for a product.
    - ii) Optimization: fine time manufacturing and delivery processes

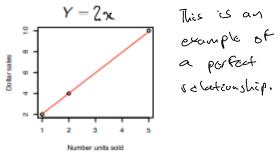
### Functional relation

▶ Relation of the form

$$Y = f(X),$$

where X, Y are variables, and f is a function.

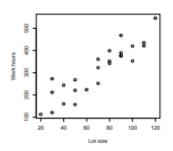
Example: Relation between dollar sales (Y) of a product sales sold \$2 per unit and number of units sold (X):



All observations fall on the line of functional relationship.

#### Statistical relation

- Not a perfect relation.
- Example: A company produces replacement parts. It produces lots of varying size. The relation between the lot size and work hours is a statistical relation.

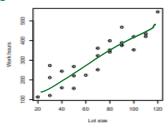


Note: There is still a trend but a greater amount of notice

#### Statistical relation

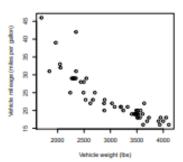
- ► Perfect relation?
  No! >>>> two is noise and data points are scattered around the trend.
- Two lots with X = 40 have different Y.
- Linear or non-linear statistical relation?

  Linear statistical relation



### Statistical relation

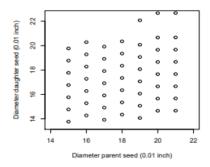
- Example: Weight and mileage for 54 cars.
- Functional or statistical relation?
- Linear statistical relation?



#### Galton's early considerations of regression

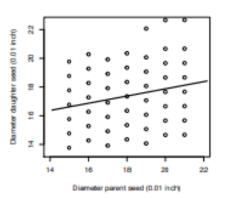
- Sir Francis Galton, English Victorian statistician, sociologist, psychologist, anthropologist, etc.
- ▶ Work on inherited characteristics of sweet peas ⇒ initial conceptualization of linear regression.
- In 1975, Galton distributed packets of sweet pea seeds to seven friends who harvested seeds from the new generations of plants and returned them to Galton.
- Galton plotted the diameter of the daughter seeds against the diameter of the mother seeds [Galton, 1894].

#### Galton's early considerations of regression



## Galton's early considerations of regression

- Mean diameter of daughter seeds from a particular diameter of mother seed approximately a straight line with positive slope Tendency of diameter of daughter seeds to vary with diameter of mother seeds
- Constant variability for diameter of daughter seeds from a particular diameter of mother seed
   Random scatter around this tendency



## Notation and general concepts

- Model: mathematical expression to describe the behavior of a random variable of interest
- Response variable or outcome Y: variable of interest
- Predictor or independent variables X: know constant variables thought to provide information on the behavior of Y
- Subscript on Y and X identifies the particular unit from which the observation was taken (X<sub>5</sub> for unit 5)
- Parameters: control behavior of the model; usually represented by Greek letters (β, σ); unknown constants to be estimated from the sample
- Note: A model is a look accurate but can be close to matching reality.

### Examples

 Dollar sales of a product sales sold \$2 per unit and number of units sold

$$Y = \beta X$$

▶ Diameter of daughter seeds and diameter of mother seeds

$$Y = \beta X + \varepsilon$$

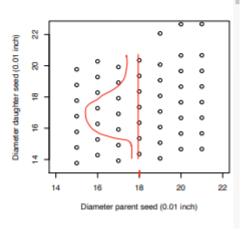
### Basic concepts

# Two characteristics of a statistical relation:

- 1. Tendency of Y to vary with X
- Random scatter around this tendency

### In a regression model:

- 1. The mean of Y vary in a systematic fashion with X
- Probability distribution of Y for any given value of X



### Data collection for regression analysis

#### Observational study

- Investigator has no control over the explanatory variables (X)
- Limitation: not adequate for cause-and-effect
   A strong association does not necessarily means a
   cause-and-effect relationship

#### Experiment

- Investigator exercises control over the explanatory variables
   (X) through random assignment
- Random assignment balances out effect of other variables that might affect Y
- Gold standard for cause-and-effect conclusions

Note: observational
studies can't
conclude cause and
corrolation.

### Example of observational study

Study the relationship between age of employees (X) and number of days of illness last year (Y)

- ► Observational data because we can't control age or # of sich days-
- An observed association between X and Y does not necessarily imply that X explains Y

- Note: There may be other factors that we have not looked at.

### Example of experiment

Study the relationship between productivity and length of training of analysts working in a bank:

- 1. 30 analysts considered
- randomly select 10 analysts that will be trained for 2 week; randomly select 10 other analysts that will be trained for 5 weeks; the 10 remaining will be trained for 8 weeks
- productivity of the 30 analysts observed for a fixed time after the training
- Experiment because investigators can manipulate the value of y

   eg. 8 uscoles

   This could have cause and effect.

### Cause-and-effect / Causation

- We observe an association between Y and X
- Does changing one of the variables imply the other to change?
- Mechanisms that can result in an observed association between Y and X:

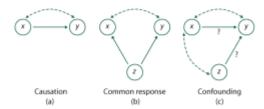


Figure 1: The dashed arrows represent association and the solid ones cause and effect link. The variable x is explanatory, y is response, and z is a lurking variable.

Regression analysis by itself provides no information about causation. Be careful in drawing causal conclusions

### Overview of the steps in regression analysis

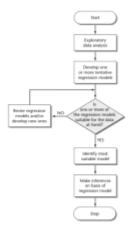


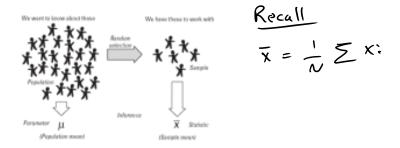
Figure 2: The steps in regression analysis [Kutner et al., 2004, p.14]

### Three main purposes of regression analysis

- Describe: describe the relation between diameter of daughter seeds and diameter of mother seeds.
- Control: control the length of training to maximize productivity constrained by costs.
- 3. Predict: predict future demand for a product.

### Parameters, estimators, and estimates

- Parameter: quantity of interest, quantity describing a population (or model).
  - A parameter is a constant (constant/random) quantity.
- Estimator: rule for calculating an estimate of parameter.
  An estimator is a \_\_\_\_\_\_(constant/random) quantity.
- Estimate: result of the estimator (for a given sample).
  An estimate is a \_\_\_\_\_\_\_ (constant/random) quantity.



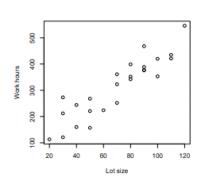
## Toluca company example<sup>1</sup>

- Toluca Company produces replacement parts for refrigeration equipment
- Produces lots of varying size
- Cost improvement: find optimal lot size
- Key input: relationship between lot size and labor hours
- ▶ Data: lot size X and work hours Y for 25 production runs

Run	Lot size	Work hour	
i	$X_i$	$Y_i$	takes to produce the items
1	80	399	takes to produce
2	30	121	
			ex/399 hours to create
24	80	342	
25	70	323	_ 80 fridges
			$=$ $\delta$

<sup>&</sup>lt;sup>1</sup>From [Kutner et al., 2004], page 19

#### Toluca company example



From the scatter plot: - looks like a linear model

### Simple linear model

Suppose we have n observed pairs  $(X_i, Y_i)$ , i = 1, ..., n. The simple linear model is:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i,$$

β<sub>0</sub> and β<sub>1</sub> are parameters,
X<sub>i</sub> is the observed value of X on unit i, and
ε<sub>i</sub> are random errors that have zero mean E(ε<sub>i</sub>) = 0, with common variance Var(ε<sub>i</sub>) = σ², and pairwise independent.

色: 上をら, 1 # j

### Simple linear model

#### Exercise 1

Show that the random errors satisfy

$$E(\varepsilon_i \varepsilon_j) = \begin{cases} 0 & \text{if } i \neq j \\ \sigma^2 & \text{if } i = j \end{cases}$$

Recall Assumptions about random errors

- 1) E( E;) =0
- 2) Var (21) = c2
- 3) pairwise independent; thus (ou(Ei,E;)=0 i+)

For this proof there are 2 cases i=j & i + j

· i=j:

wts E(2; . 2;) = E(2;2) = 62

we know: Var (2,) = 62

No Nor 
$$(2i) = E(2i) - (E(2i))^2$$
, first assumption and  $E(2i) = 0$ 

So  $Var(2i) = E(2i) - (E(2i))^2$ , second assumption facts  $Var(2i) = 6^2$ 

$$6^2 = 6^2 - 0$$

$$6^2 = 6^2 - 0$$

e ifj: we want to show E(E(E) =0

$$0 = E(z; z_j) - E(z_i)E(z_j), first assumption$$

$$0 = E(z; z_j)$$

$$0 = E(z; z_j)$$

### Important features

Simple linear model

constant => understand constant as not random.  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$ 

Parameters are always

constant. We don't know

them but they are combat

- The response Y<sub>i</sub> is a sum of two terms:
  - A constant term
  - A random tern

The outcome Yi is random (constant/random)

2.  $E(Y_i) = \beta_0 + \beta_1 X_i$ , where  $E(Y_i)$  is a shortcut for  $E(Y_i|X_i)$ 4: 12 constant + random = rundom the mean of Y when  $X = X_i$ .  $E(y_i) = E(B_0 + B_1 \times i + E_i) = E(B_0) + E(B_1 \times i) + E(E_i)$ , linearity = Bo + Bix:

Thus, the functional relationship between the true mean of  $Y_i$ and  $X_i$  is a straight line with intercept  $\beta_0$  and slope  $\beta_i$ 

#### Important features

Simple linear model

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

3.  $Var(Y_i) = \sigma^2$ , where  $Var(Y_i)$  is a shortcut for  $Var(Y_i|X_i)$  the mean of Y when  $X = X_i$ .  $Var(y_i) = Var(\frac{y_i}{x_i}) = Var(\frac{y_i}$ 

$$Var(y_i) = Var(\frac{g_0 + R_1 X_1}{constant} + \xi_i) = Var(\xi_i) = 6^2$$

4. The outcomes  $Y_i$  are pairwise independent because the errors  $\varepsilon_i$  are pairwise independent.

Recall

Variance is not

Incar.

You could use expediction but

too hard.

Bo + By Xi is a constant

Var (constant + reV) = Var(r.V)

#### Reminder: normal distribution

A random variable X is normal if its probability density function is

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right\},\,$$

where  $\mu \in \mathbb{R}$  and  $\sigma \in \mathbb{R}^+$  are the parameters of the distribution. We say that X is normally distributed with mean  $E(X) = \mu$  and variance  $Var(X) = \sigma^2$  and we write

$$X \sim \mathcal{N}(\mu, \sigma^2)$$
.

Additional (tsymption)

Sometimes we make an additional assumption that random errors are normally distributed.

- This assumption is only used when explicitly staded

### Simple linear model with normal errors

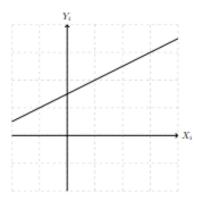
- The random errors are sometimes assumed to be normally distributed.
- Simple linear model with normal errors:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

where

- β<sub>0</sub> and β<sub>1</sub> are parameters,
- ε<sub>i</sub> are independently and identically distributed (i.i.d.) with normal distribution with mean 0 and variance σ<sup>2</sup>.
- In what follows, we suppose a simple linear model (errors not necessarily normal) unless otherwise specified.

### Interpretation of the regression parameters



- ▶ If the scope of the model includes X = 0, the intercept  $\beta_0$  is the mean of Y when X = 0 (no meaning otherwise)
- The slope β<sub>1</sub> is the change in the mean of Y per unit increase of X

### Estimation of the parameters

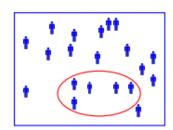
► Postulated model:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

- Observed values (X<sub>i</sub>, Y<sub>i</sub>)
- Parameters β<sub>0</sub> and β<sub>1</sub> unknown and to be estimated from the sample.



- 1. Least squares
- 2. Maximum likelihood
- $\Rightarrow$  Estimates  $\widehat{\beta}_0$  and  $\widehat{\beta}_1$



- See the rodel as describing the population
- Solect scriple out random
- observe X & g value
- estimate Po and Bi trun the sample

Me hat means estimate, not true value.

#### Method of least squares

Simple linear model:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

- ▶ Parameters  $\beta_0$  and  $\beta_1$  to be estimated from the data.
- ▶ **Goal**: find the best estimates  $\widehat{\beta}_0$  and  $\widehat{\beta}_1$  given the data.
- ▶ What does best mean?
- Least square: best by criterion

$$Q(\beta_0,\beta_1) = \sum_{i=1}^{n} (\underbrace{Y_i - \beta_0 - \beta_1 X_i})^2$$

$$\underbrace{Y_i - \underbrace{Y_i - \underbrace{Y_i}}_{0 - y_i}}_{\text{is the deviation of } Y_i \text{ from its expected value.}}$$

Least square estimators of β<sub>0</sub> and β<sub>1</sub>: β̂<sub>0</sub> and β̂<sub>1</sub> that minimize criterion Q.

the sum of the evuces to get the best me

#### Least square estimators

Find  $\widehat{\beta}_0$  and  $\widehat{\beta}_1$  that minimize criterion

$$Q(\beta_0, \beta_1) = \sum_{i=1}^{n} (Y_i - \beta_0 - \beta_1 X_i)^2$$

given the data.

- 1. Write the normal equations (derivatives of Q set to 0).
- 2. Find the critical points (solution of the normal equations).
- Determine whether the critical point is a maximum or a minimum (we will skip this step).

1) 
$$\frac{\partial Q}{\partial B_0} = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$$
 (-1)  
 $O = -2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (-xi)  
 $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (-xi)  
2)  $\frac{\partial Q}{\partial B_1} = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (-xi)  
 $O = -2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (0)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (0)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (2)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (3)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (4)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (5)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (6)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (7)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (8)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (9)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (2)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (3)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (4)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (5)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (6)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (7)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (8)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (9)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (2)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (3)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (4)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (5)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (6)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (7)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (8)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (9)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i - B_0 - B_1 x_i)$  (1)  $O = 2 \sum_{i=1}^{\infty} (y_i$ 

#### Least square estimators

Least square estimators of 
$$\beta_1$$
 and  $\beta_0$ : 
$$\widehat{\beta}_1 = \frac{\sum_{i=1}^n X_i Y_i - \frac{1}{n} \sum_{i=1}^n X_i \sum_{i=1}^n Y_i}{\sum_{i=1}^n X_i^2 - \frac{1}{n} \left(\sum_{i=1}^n X_i\right)^2},$$
 
$$\widehat{\beta}_2 = \overline{Y} - \widehat{\beta}_1 \overline{X}$$

### Least square estimators

#### Exercise 2

Show that

$$\widehat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \overline{X})(Y_i - \overline{Y})}{\sum_{i=1}^n (X_i - \overline{X})^2}.$$

#### Regression equation

Regression equation or fitted regression line

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 X$$

where  $\widehat{Y}$  is the estimated mean of the response variable at level  $\boldsymbol{X}$  of the explanatory.

#### Gauss-Markov theorem

#### Theorem 1

Consider the simple linear model

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$
.

Suppose that the following assumptions concerning the random errors (called Gauss-Markov assumptions) are satisfied:

- They have mean zero: E(ε<sub>i</sub>) = 0,
- They are homoscedastic: Var(ε<sub>i</sub>) = σ<sup>2</sup> < ∞, and</p>
- ▶ There are uncorrelated  $Cov(\varepsilon_i, \varepsilon_j) = 0, \forall i \neq j$ .

Then the least squares estimators  $\widehat{\beta}_0$  and  $\widehat{\beta}_1$  are unbiased and have minimum variance among all unbiased linear estimators.

#### Proof of the Gauss-Markov theorem

Step 1 Exercise: Prove that the least squares estimators are unbiased, i.e. prove that

$$E(\widehat{\beta}_1) = \beta_1$$
 and  $E(\widehat{\beta}_0) = \beta_0$ 

Step 2 To be proven later: The least squares estimators have minimum variance among all unbiased linear estimators.

### Toluca company example

Using R, we find:

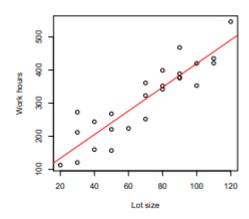
$$\sum_{i=1}^{n} X_i = 1750 \quad \sum_{i=1}^{n} Y_i = 7807 \quad \sum_{i=1}^{n} X_i Y_i = 617180$$

$$\sum_{i=1}^{n} X_i^2 = 142300 \quad n = 25$$

#### Exercise 3

- 1. Compute the least squares estimates of  $\beta_1$  and  $\beta_0$ .
- 2. What is the regression equation?
- 3. Interpret the parameters.

### Toluca company example



### Toluca company example: R output

```
Call:
lm(formula = Hours ~ Size, data = toluca)
Residuals:
            1Q Median
                            3Q
   Min
-83.876 -34.088 -5.982 38.826 103.528
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 62.366 26.177 2.382 0.0259 *
             3.570
                        0.347 10.290 4.45e-10 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 48.82 on 23 degrees of freedom
Multiple R-squared: 0.8215, Adjusted R-squared: 0.8138
F-statistic: 105.9 on 1 and 23 DF, p-value: 4.449e-10
```

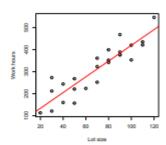
### Preliminary exercise: Toluca company example

 Regression equation (or fitted regression line)

$$\hat{Y} = 62.37 + 3.5702X$$

#### where

- X is the lot size, and
- Y is the work hours.



#### What is

- The predicted work hours for a new production run for a lot size of 60?
- The estimated population mean work hours for a lot size of 60?

### Predicted values and residuals

Fitted regression line:

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 X$$

Fitted value: value of Y computed from the regression line:

$$\widehat{Y}_i = \widehat{\beta}_0 + \widehat{\beta}_1 X_i.$$

Fitted value  $\hat{Y}_i$  used as:

- Prediction of the value of Y for particular value X<sub>i</sub> of X.
   Sometimes written Ŷ<sub>predi</sub>.
   Estimate of the population mean of Y for particular value X<sub>i</sub>
- ▶ A residual is the deviation of the observed value Y<sub>i</sub> from the fitted value  $\hat{Y}_i$

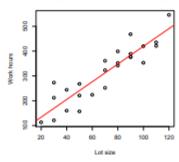
$$e_i = Y_i - \widehat{Y}_i$$

### Toluca company example

For production run 6, the lot size was 60 and 224 work hours were required.

#### Exercise 4

- 1. What is the fitted value for this observation?
- 2. What is the residual?
- 3. Where can we read the observed work hours (Y), the fitted work hours, and the residual in the scatterplot?



### **Exercises**

- ► From the textbook<sup>2</sup>
  - ▶ 1.20
  - ▶ 1.21
- ► From the slides<sup>3</sup>
  - ► Slide 26
  - ► Slide 37
  - ► Slide 40
  - ► Slide 41
  - ► Slide 47

<sup>&</sup>lt;sup>2</sup>Solutions in the student manual (CD) provided with the textbook

<sup>&</sup>lt;sup>3</sup>Partial solutions posted on Quercus