Tutorial 3 (Abubakari Sumaila Salpawuni)

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 Q_1 – solution

1.(a) Given $X \sim U(0,1)$. Define g(x) as the area the square such that;

Area =
$$g(x) = X^2$$

$$E(g(X)) = \int_0^1 g(x) f_X(x) dx$$

$$E(X^2) = \int_0^1 x^2 dx = \frac{1}{3} x^3 \Big|_0^1$$

$$= \frac{1}{3} \text{sq. units}$$

1.(b) If $X_i \sim Poi(\lambda_i)$, then it moment generating function, m.g.f, is $M_X(t) = e^{[\lambda_i(e^t-1)]}$. Since the variables are independent, we have;

$$M_{Y} = M_{X_{1}}(t) \cdot M_{X_{2}}(t) \dots \cdot M_{X_{n}}(t)$$

$$= e^{[\lambda_{1}(e^{t}-1)]} \cdot e^{[\lambda_{2}(e^{t}-1)]} \cdot \dots \cdot e^{[\lambda_{n}(e^{t}-1)]}$$

$$= e^{[(\lambda_{1}+\lambda_{2}+\dots+\lambda_{n})(e^{t}-1)]}$$

$$= e^{[\sum_{i=1}^{n} \lambda_{i}(e^{t}-1)]}$$

$$\therefore \sum_{i=1}^n X_i \sim \operatorname{Poi}(\sum_i^n \lambda_i)$$

1.(c) Given a density function, one way to solve this is to use the method of transformation.

Let $y = g(x) = \frac{1}{x}$. The density of g(y) can be computed as;

$$g(y) = \left| \frac{dx}{dy} \right| \cdot f(W(y))$$

where W(y) is the inverse function of g(x).

$$y = \frac{1}{x} \implies x = \frac{1}{y}$$
$$\left| \frac{dx}{dy} \right| = -\frac{1}{y^2} = \frac{1}{y^2}$$

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$$g(x) = \frac{1}{y^2} \times \frac{1}{(n-1)!\beta^n} \left(\frac{1}{y}\right)^{n-1} e^{-\frac{1}{y\beta}}$$
$$= \frac{1}{(n-1)!\beta^n} \left(\frac{1}{y}\right)^{(n+2)-1} e^{-\frac{1}{y\beta}}$$
$$\therefore g\left(\frac{1}{x}\right) = \operatorname{Gamma}(n+2,\beta)$$

Q_2 – solution

This is clearly a Bayesian problem. Define D as an event that a subject has the disease; T^+ be event that a subject's test result returns positive, and the event that a subject's test result T^- returns negative. Thus far, we proceed with the following pieces of information;

$$P(D) = 0.015$$
 (prevalence)
 $P(T^+|D) = 0.97$ (sensitivity)
 $P(T^-|D') = 0.95$ (specificity)

2.(a) i. We partition T into T^+ and T^- . Thus,

$$P(T^{+}|D) = P(D)P(T^{+}|D) + P(D')P(T^{+}|D')$$

= 0.015 × 0.97 + (1 – 0.015) × (1 – 0.95)
= 0.905975

ii. We're interested in $P(D|T^+)$, the probability you have the disease given you tested positive to the disease. Using the Bayesian formulation, we proceed as follows:

$$P(D|T^{+}) = \frac{P(D)P(T^{+}|D)}{P(D)P(T^{+}|D) + P(D')P(T^{+}|D')}$$
$$= \frac{0.015 \times 0.97}{0.905975}$$
$$= 0.0161(1.61%)$$

2.(b)
$$P(A)=0.40, P(B)=0.10, P(C)=0.50$$

Let D be the event that a defective product is produced $P(D|A)=0.02$
 $P(D|B)=0.03$
 $P(D|C)=0.04$

i.

$$P(D) = P(A)P(D|A) + P(B)P(D|B) + P(C)P(D|C)$$

= 0.40(0.02) + 0.10(0.03) + 0.50(0.04)
= 0.031

ii.
$$\alpha$$
. $P(A|D) = \frac{P(A)P(D|A)}{P(D)} = \frac{0.40 \times 0.02}{0.031} = \frac{8}{31}$
 β . $P(B|D) = \frac{P(B)P(D|A)}{P(D)} = \frac{0.10 \times 0.03}{0.031} = \frac{3}{31}$
 γ . $P(C|D) = \frac{P(C)P(D|A)}{P(D)} = \frac{0.50 \times 0.04}{0.031} = \frac{20}{31}$

3.(c) Using Matrix approach, Let $\hat{y} = X\beta + \epsilon$ such that;

$$\beta = (X'X)^{-1}X'y$$

$$X'X = \begin{pmatrix} 1 & 1 & \dots & 1 \\ 5 & 6 & \dots & 10 \\ 1 & 1 & \dots & 6 \end{pmatrix} \begin{pmatrix} 1 & 5 & 1 \\ 1 & 6 & 1 \\ \vdots & \vdots & \vdots \\ 1 & 10 & 6 \end{pmatrix}$$

$$= \begin{pmatrix} 10 & 70 & 40 \\ 70 & 514 & 298 \\ 40 & 298 & 194 \end{pmatrix}$$

$$(X'X)^{-1} = \begin{pmatrix} 2.2179 & -0.3374 & -0.0610 \\ -0.3374 & 0.0691 & 0.0366 \\ -0.0610 & 0.0366 & 0.0488 \end{pmatrix}$$

$$X'y = \begin{pmatrix} 60 \\ 449 \\ 264 \end{pmatrix}$$

$$\hat{\beta} = \begin{pmatrix} 2.2179 & -0.3374 & -0.0610 \\ -0.3374 & 0.0691 & 0.0366 \\ -0.0610 & 0.0366 & 0.0488 \end{pmatrix} \begin{pmatrix} 60 \\ 449 \\ 264 \end{pmatrix}$$

$$= \begin{pmatrix} -2.3146 \\ 1.1195 \\ 0.1098 \end{pmatrix}$$

 $\hat{y} = -2.3146 + 1.1195X_1 + 0.1098X_2$

 Q_3 – solution

3.(a) i.

$$\int_{-\infty}^{\infty} x f(x) dx = \int_{-1}^{0} \frac{2x|x|}{5} dx + \int_{0}^{2} \frac{2x|x|}{5} dx$$

$$E(X) = -\frac{2}{5} \int_{-1}^{0} x^{2} dx + \frac{2}{5} \int_{0}^{2} x^{2} dx$$

$$= -\frac{2}{15} x^{3} \Big|_{-1}^{0} + \frac{2}{15} x^{3} \Big|_{0}^{2}$$

$$= -\frac{2}{15} + \frac{16}{15}$$

$$= \frac{14}{15}.$$

ii.

$$Var(X) = E(X^{2}) - [E(X)]^{2}$$

$$= -\int_{-1}^{0} \frac{2x^{2}|x|}{5} dx + \int_{0}^{2} \frac{2x^{2}|x|}{5} dx - \left(\frac{14}{15}\right)^{2}$$

$$= -\frac{1}{10}x^{4}\Big|_{-1}^{0} + \frac{1}{10}x^{4}\Big|_{0}^{2} - \left(\frac{14}{15}\right)^{2}$$

$$= -\frac{1}{10} + \frac{16}{10} - \left(\frac{14}{15}\right)^{2} = \frac{17}{10} - \frac{196}{225}$$

$$= \frac{95}{102}.$$

3.(b)

$$Y \sim U(-1,8) = \begin{cases} \frac{1}{9}, & -1 < x < 8 \\ 0, & \text{otherwise} \end{cases}$$

For the equation $2x^2 + 4Yx + 3Y + 2 = 0$ to have real roots, the *discriminant* must be non-negative, thus;

$$b^{2} - 4ac \ge 0 \implies (4Y)^{2} - 4(2)(3Y + 2) \ge 0$$
$$\implies 16Y^{2} - 24Y + 16 \ge 0$$
$$\implies 2Y^{2} - 8Y + 2 \ge 0$$
$$\implies Y \ge 2 \quad \text{or} \quad Y \le -\frac{1}{2}$$

Using the probability density function above, and the idea of mutually exclusive events, we have;

$$P(Y \le -\frac{1}{2} \cup Y \ge 2) = P(Y \le -\frac{1}{2}) + P(Y \ge 2)$$

$$= \int_{-1}^{\frac{1}{2}} f_Y(y) dy + \int_{2}^{8} f_Y(y) dy$$

$$= \frac{1}{9} (\frac{1}{2} + 6)$$

$$= \frac{13}{18}.$$

3.(c) Note that the moment generating function for the Exponential distribution is given as $M_X(t) = \frac{\lambda}{\lambda - t}$, $t < \lambda$. Comparing this with the above function, we may infer that;

$$f(x) = \begin{cases} \frac{1}{2}e^{-\frac{1}{2}\lambda t}, & t < \lambda \\ 0, & \text{otherwise} \end{cases}$$

i. The cumulative density function for the Exponential distribution is thus;

$$1 - F(x) = e^{-\lambda t}, \quad t < \lambda$$

$$\therefore P(X > \text{In4}) = e^{-\frac{1}{2}\text{In4}} = \frac{1}{2}.$$

ii.

$$E(X) = \frac{d}{dx} M_X(t) \Big|_{t=0}$$
$$= \frac{0.5}{(0.5 - t)^2} \Big|_{t=0}$$
$$= 2$$

$$E(X^{2}) = \frac{d^{2}}{dx^{2}} M_{X}(t) \Big|_{t=0}$$
$$= \frac{2(0.5)}{(0.5-t)^{3}} \Big|_{t=0}$$
$$= 8$$

$$Var(X) = E(X^2) - [E(X)]^2$$

= $8 - 2^2 = 4$
 $\therefore Var(X) = 4$.