
Hypothesis Testing & Tests

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Learning Objectives: Hypothesis Tests

- Describe the difference between a hypothesis test and a confidence interval
- Correctly use vocabulary of hypothesis testing:
 - Null Hypothesis: H_0
 - Reject or Fail to Reject the Null Hypothesis
 - Alternate Hypothesis: H_1
 - p -value
- Test hypotheses for:
 - a single proportion
 - two proportions
 - two means
- Distinguish when to use one-way and two-way hypothesis test

Descriptive vs. Inferential Statistics

- **Descriptive** or Summary Statistics
 - Goal: to **describe the features** of a collection of data in a quantitative way
 - Measures of Central Tendency:
 - mean, median, and mode
 - Measures of Variability, Dispersion, or Spread:
 - range, variance, standard deviation, quartiles
- **Inferential** or Inductive Statistics
 - Goal: to summarize a **sample of the data** to infer or draw conclusions about the **population** from which the sample is drawn
 - Hypothesis testing
 - A/B testing
 - p -value
 - t -tests, χ^2 -tests, F -tests
 - Confidence intervals

Hypothesis Testing

Hypothesis testing tell us *how extreme* the observed result is compared to what random chance might produce, and it helps us make decisions on that basis.

Lecture Outline

- Hypothesis Testing
 - Testing Procedure
 - Null Hypothesis & Alternative Hypothesis
 - p -value and Degrees of Freedom (df)
 - Type I and Type II Errors
- Exemplar Tests: Parametric & Nonparametric
 - T-Tests and Wilcoxon Tests
 - Single Sample: T-Test
 - Two-Sample: Independent Groups
 - Paired Two-Sample: Dependent Groups
 - Multiple Samples: Independent Groups

Statistical Distributions & Functions in R

Distribution	Random Number Generator	Density	Distribution	Quantile
Normal	r norm	d norm	p norm	q norm
t	rt	dt	pt	qt
F	rf	df	pf	qf
χ^2	rchisq	dchisq	pchisq	qchisq

{d p q r} *distribution_abbreviation*()

- **d** = density
- **p** = distribution function
- **q** = quantile function
- **r** = random generation

- **pnorm(a)** $\equiv P(X \leq a)$: probability that a or smaller number occurs
- **pnorm(b) - pnorm(a)** $\equiv P(a \leq X \leq b)$: probability that the variable falls between two points
- **qnorm()**: given the cumulative probability distribution, it returns the quantile

Statistical Distributions: Mean and Variance

Distribution	Degrees of freedom	Mean	Variance
Normal		μ	σ^2
t	n	0	$n/(n - 2)$
F	n_1 and n_2	$n_2/(n_2 - 2)$	a/b
χ^2	r	r	$2r$

$$a = 2n_2^2(n_1 + n_2 - 2)$$

$$b = n_1(n_2 - 2)^2 (n_2 - 4)$$

Reminder: Statistic & its Proxy

Aim	Model Statistic	Sample Statistic	Proxy Statistic	Formula for Proxy
Estimate the mean μ of a normal distribution with known variance σ^2	μ	m	Z-statistic	$Z \sim \frac{m - \mu}{\sigma / \sqrt{n}}$
Estimate the variance σ^2 of a normal distribution with known mean μ	σ^2	S^2	χ^2 -statistic	$\chi_{n-1}^2 \sim (n - 1) \frac{S^2}{\sigma^2}$
Estimate the mean μ of a normal distribution with un-known variance σ^2	μ	m	t -statistic	$T_{n-1} \sim \frac{m - \mu}{S / \sqrt{n}}$

Ex.	Proxy Statistic	Distribution	Degrees of Freedom (df)
1	Z-statistic	$N(0, 1)$	
2	χ^2 -statistic	$\chi^2(n - 1)$	$n - 1$
3	t -statistic	T_{n-1}	$n - 1$

Hypothesis Testing: Procedure

- Step 1: Define **a statistic** that obeys a certain **distribution** if the hypothesis is correct:
 - Ex-1: The mean μ from a normal distribution with known variance σ^2
 - Ex-2: The variance σ^2 from a normal distribution with known mean μ
 - Ex-3: The mean μ from a normal distribution with unknown variance σ^2
- Step 2 (optional): Transform the statistic to a **proxy statistic** with the **proxy distribution** of better understood properties/characteristics:
 - Ex-1: Z-statistic from a uniform normal distribution, $N(0,1)$
 - Ex-2: χ_{n-1} -statistic from a χ^2 distribution with n df
 - Ex-3: T_{n-1} -statistic from a t -distribution with $n - 1$ df
- Step 3: Calculate the statistic (original/proxy) from the **sample**
- Step 4: Compute the **probability** (the **p-value**) of this sample with this statistic to be drawn from this distribution (original/proxy)
 - **Reject the hypothesis** if probability is **low** (e.g., **p-value** < 0.05)
 - **Fail to reject the hypothesis** otherwise (e.g., **p-value** \geq 0.05)

Important Note

DO NOT SAY: We **ACCEPT** the Hypothesis

INSTEAD: We **FAIL TO REJECT** the Hypothesis

- Given the sample we had to calculate the statistic

Null Hypothesis vs. Alternative Hypothesis

- **Null Hypothesis** (H_0): what is considered to be true:
 - **Example:** $H_0 : \mu = \mu_0$: We want to test a hypothesis that the unknown mean μ for a sample from a normal distribution with known variance σ^2 is equal to a specific constant μ_0
- **Alternative Hypothesis** (H_1): If the null hypothesis is rejected:
 - **Example:** $H_1 : \mu \neq \mu_0$

Examples: Null and Alternative Hypotheses

- Null = “no difference between the means of group A and group B”
 - Alternative = “A is different from B” (could be bigger or smaller)
- Null = “ $A \leq B$ ”
 - Alternative = “ $B > A$ ”
- Null = “B is not $x\%$ greater than A”
 - Alternative = “B is $x\%$ greater than A”

- The Null and Alternative Hypotheses must account for all possibilities.
- The nature of the null hypothesis determines the structure of the hypothesis test.

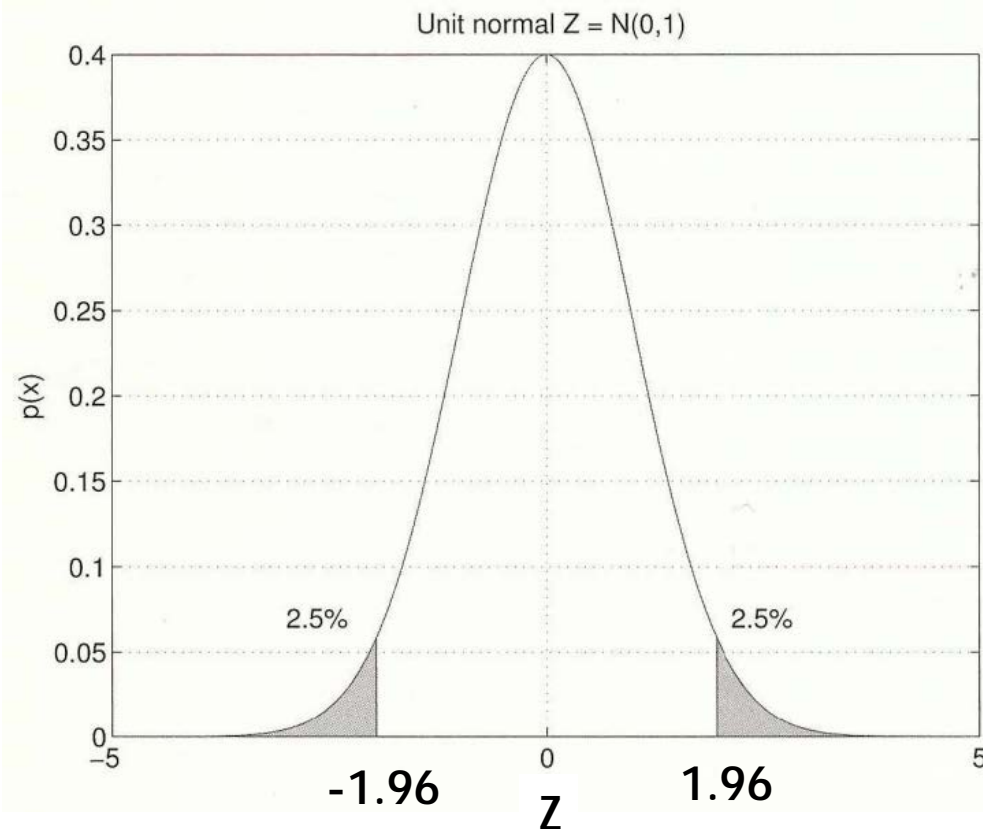
H1 Hypothesis: One-Way or Two-Way Tests?

- A **directional** alternative hypothesis:
 - B is better than A
 - Use a **one-way or one-tail** hypothesis test
 - An extreme chance results in only one direction count toward the p -value
 - A one-tail hypothesis test often fits the nature of A/B decision making
 - decision is required and one option is assigned “default” unless the other proves better
- A **bi-directional** alternative hypothesis
 - A is different from B (could be bigger or smaller)
 - Use a **two-way or two-tail** hypothesis
 - Extreme chance results in either direction count toward the the p -value
 - R software typically provides a two-tail test in default output

Summary: Key Ideas

- A **null hypothesis** embodies the notion that nothing special has happened
 - any effect you observe is due to random chance
- The **hypothesis test** assumes the null hypothesis is true, creates a “null model” (a probability model), and tests whether the effect you observe is a reasonable outcome of that model

Two-sided Confidence Interval for $Z \sim N(0, 1)$



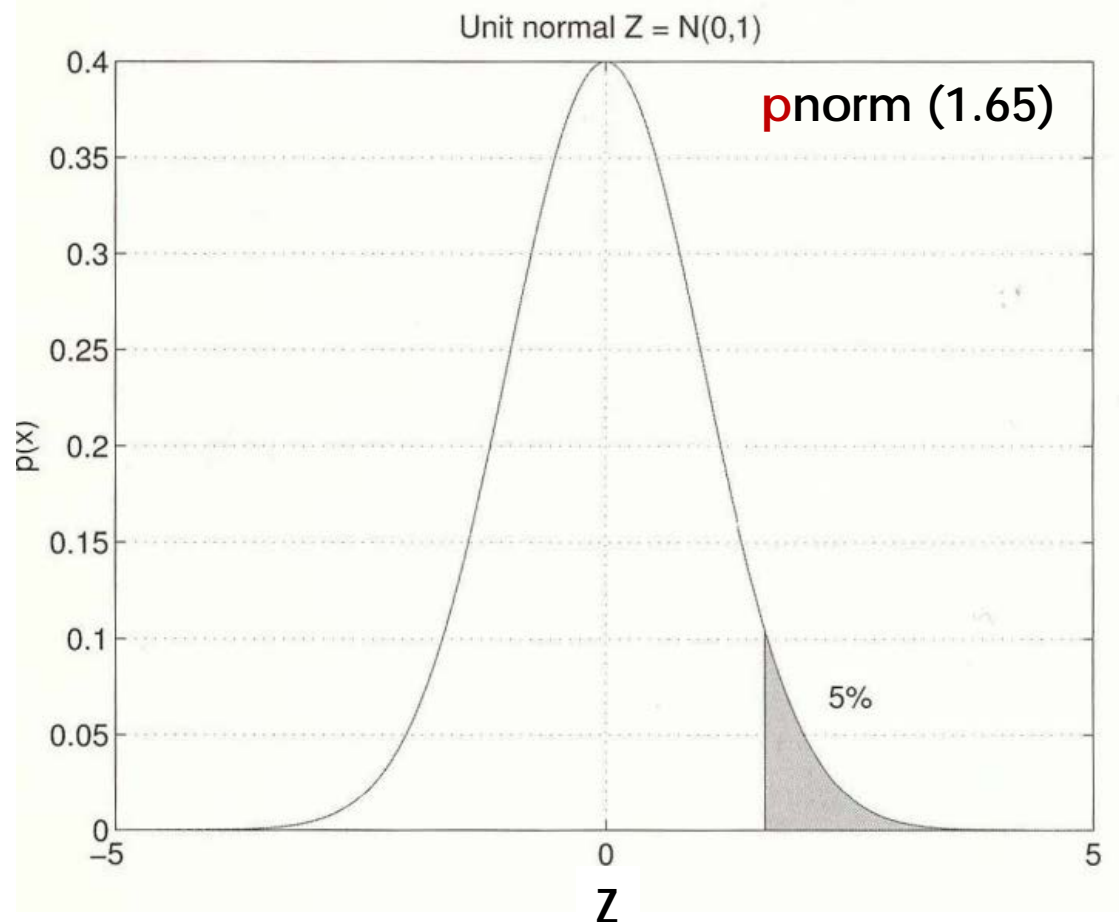
95% of the unit normal distribution lies between - 1.96 and 1.96

$$P\{ |Z - 0| < 1.96 \} = 0.95$$

$$\text{pnorm}(1.96) - \text{pnorm}(-1.96)$$

What is $(1 - \text{pnorm}(1.96))$?

One-sided Confidence Interval for $Z \sim N(0, 1)$

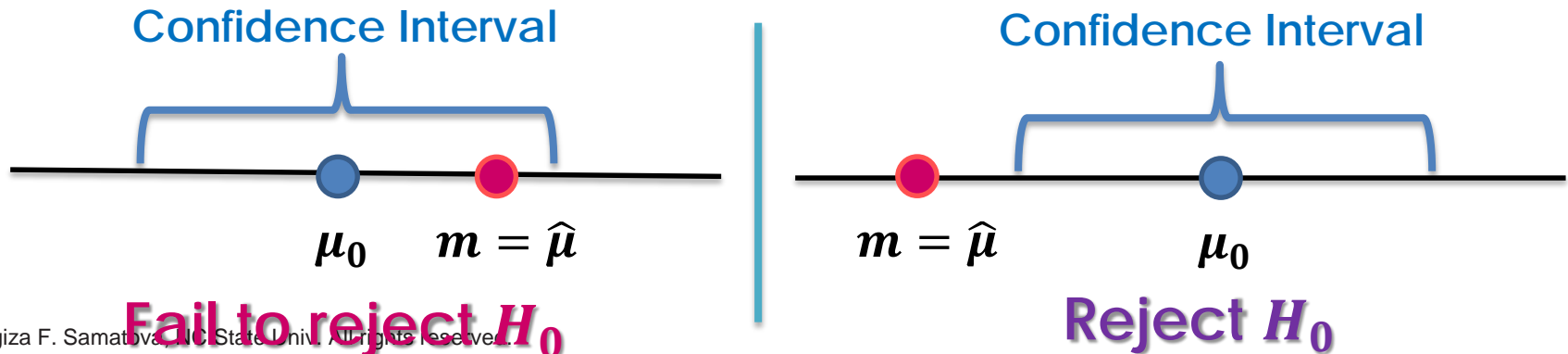


95% of the unit normal distribution lies below 1.64

$$P\{ Z < 1.64 \} = 0.95$$

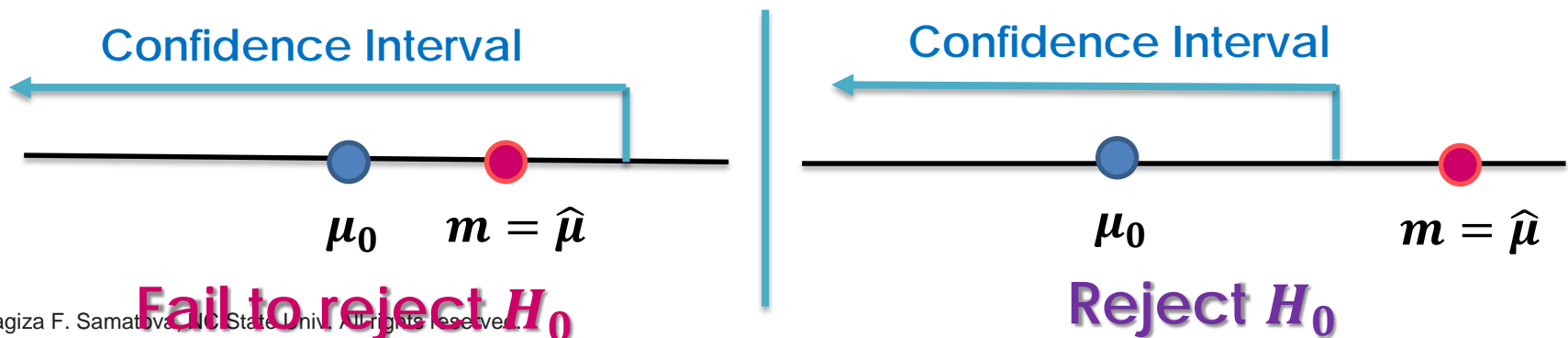
Significance Level: Two-sided Test

- Null Hypothesis: $H_0 : \mu = \mu_0$
- Alternative Hypothesis: $H_1 : \mu \neq \mu_0$
- **Significance Level** (α): We fail to reject the null hypothesis with *level of significance* α if the estimate of the sample statistic lies within the $100(1 - \alpha)$ percent *two-sided confidence interval* (CI) for the hypothesized value of the statistic:
 - m is the point estimate of μ :
 - We **fail to reject** H_0 if m is close to μ_0 , i.e., within the confidence interval, namely, if $Z \sim \frac{m - \mu_0}{\sigma / \sqrt{n}} \in (-z_{\alpha/2}, z_{\alpha/2})$
 - We **reject** H_0 if m is too far from μ_0 , i.e., outside the confidence interval, namely, if $Z \sim \frac{m - \mu_0}{\sigma / \sqrt{n}} \notin (-z_{\alpha/2}, z_{\alpha/2})$



Significance Level: One-sided Test

- Null Hypothesis: $H_0 : \mu \leq \mu_0$
- Alternative Hypothesis: $H_1 : \mu > \mu_0$
- **Significance Level** (α): We fail to reject the null hypothesis with *level of significance* α if the estimate of the sample statistic lies within the $100(1 - \alpha)$ percent *one-sided confidence interval* (CI) for the hypothesized value of the statistic:
 - m is the point estimate of μ :
 - We **fail to reject** H_0 if m is close to μ_0 , i.e., within the confidence interval, namely, if $Z \sim \frac{m - \mu_0}{\sigma / \sqrt{n}} \in (-\infty, z_\alpha)$
 - We **reject** H_0 if m is too far from μ_0 , i.e., outside the confidence interval, namely, if $Z \sim \frac{m - \mu_0}{\sigma / \sqrt{n}} \notin (-\infty, z_\alpha)$



Exercise: Test the null hypothesis

- **Null Hypothesis** (H_0): **what is considered to be true**:
 - $H_0 : \mu = \mu_0$: We want to test a hypothesis that the **unknown** mean μ for a sample from a normal distribution with **unknown** variance σ^2 is equal to a specific constant μ_0
- Hint: Use **t-statistic** rather than **Z-statistic** from the previous examples

Solution: Test the null hypothesis

- **Null Hypothesis** (H_0): what is considered to be true:
 - $H_0 : \mu = \mu_0$: We want to test a hypothesis that the **unknown** mean μ for a sample from a normal distribution with **unknown** variance σ^2 is equal to a specific constant μ_0

Use **t**-statistic: $T_{n-1} \sim \frac{m - \mu}{S / \sqrt{n}}$

Two-sided Test:

- We **fail to reject** H_0 at significance level α if

$$T_{n-1} \sim \frac{m - \mu_0}{S / \sqrt{n}} \in (-t_{\alpha/2, n-1}, t_{\alpha/2, n-1})$$

- We **reject** H_0 at significance level α if

$$T_{n-1} \sim \frac{m - \mu_0}{S / \sqrt{n}} \notin (-t_{\alpha/2, n-1}, t_{\alpha/2, n-1})$$

Example: T-Test Hypothesis Testing in R

Null Hypothesis (H_0): The average tip is equal to \$2.50

```
> data(tips, package = "reshape2")
> head (tips)
  total_bill  tip  sex smoker day  time size
1    16.99  1.01 Female    No  Sun  Dinner    2
2    10.34  1.66   Male    No  Sun  Dinner    3
3    21.01  3.50   Male    No  Sun  Dinner    3
4    23.68  3.31   Male    No  Sun  Dinner    2
5    24.59  3.61 Female    No  Sun  Dinner    4
6    25.29  4.71   Male    No  Sun  Dinner    4
> unique (tips$sex)
[1] Female Male
Levels: Female Male
> unique (tips$day)
[1] Sun  Sat  Thur Fri
Levels: Fri Sat Sun Thur
```

One-Sample T-Test (cont.)

Null Hypothesis (H_0): The average tip is equal to \$2.50

```
> t.test(tips$tip, alternative="two.sided", mu=2.5)
```

one sample t-test

data: tips\$tip

$t = 5.6253$, $df = 243$, $p\text{-value} = 5.08e-08$

alternative hypothesis: true mean is not equal to 2.5

95 percent confidence interval:

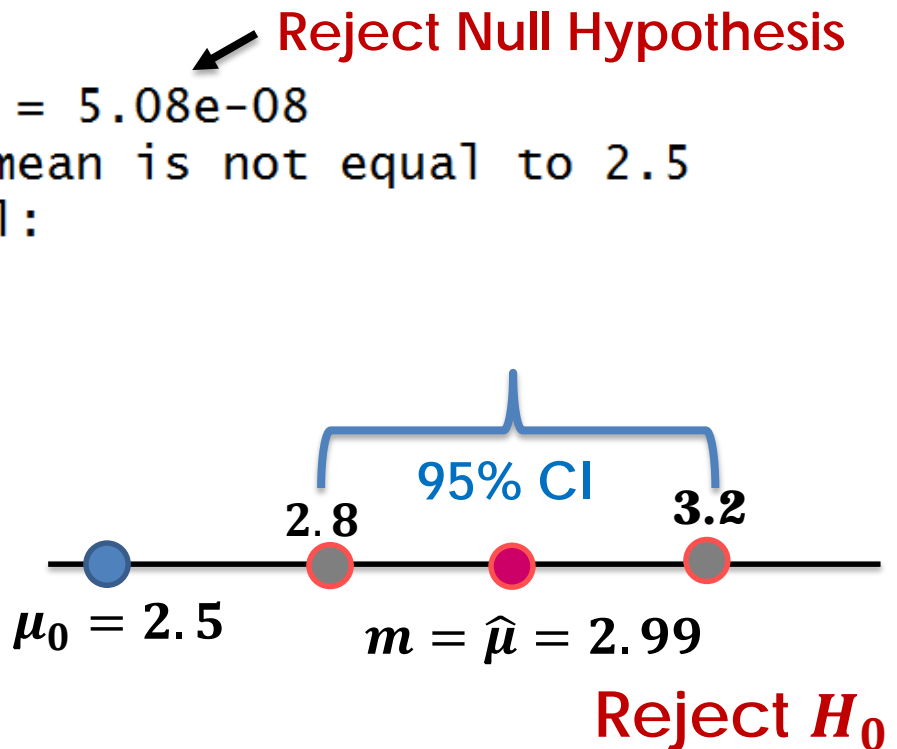
2.823799 3.172758 ←

sample estimates:

mean of x

2.998279 ←

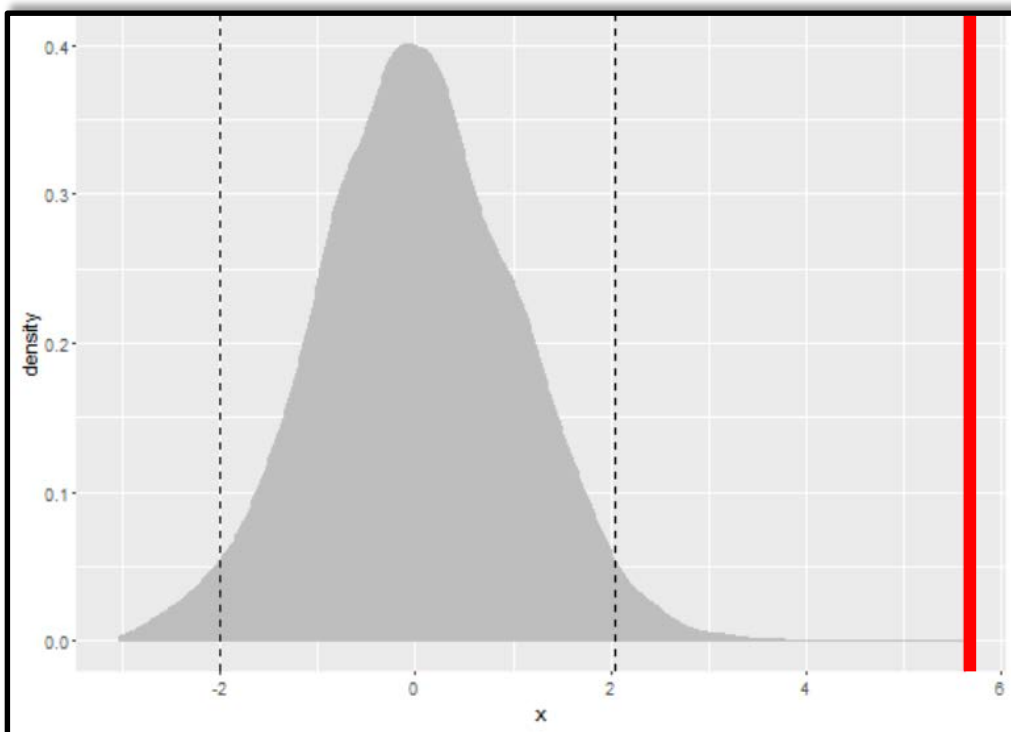
- The p -value (less than 0.05) indicates the null hypothesis should be rejected



Conclusion: The mean is not equal to \$2.50

Examine t -statistic & its probability visually

```
22 randT <- rt(3000, df=NROW(tips)-1)
23 tipTTest <- t.test(tips$tip,
24                   alternative="two.sided",
25                   mu=2.5)
26 require(ggplot2)
27 ggplot(data.frame(x=randT)) +
28   geom_density(aes(x=x), fill="grey", color="grey") +
29   geom_vline(xintercept=tipTTest$statistic, color="red") +
30   geom_vline(xintercept=mean(randT) +
31             c(-2,2)*sd(randT), linetype=2)
```



Probability of t-statistic

p-value = 5.08e-08

t-statistic = 5.62

t-distribution and t-statistic for tip data:

- dashed lines are two sd's from the mean in either direction
- thick red line (t-statistic) is far outside the distribution → reject null hypothesis → true mean is not equal to \$2.50

What about one-sided T-Test

Null Hypothesis (H_0): The average tip is less than \$2.50

```
> t.test(tips$tip, alternative="greater", mu=2.5)
```

One sample t-test

```
data: tips$tip
```

```
t = 5.6253, df = 243, p-value = 2.54e-08
```

```
alternative hypothesis: true mean is greater than 2.5
```

```
95 percent confidence interval:
```

```
2.852023      Inf
```

```
sample estimates:
```

```
mean of x
```

```
2.998279
```

 **Reject Null Hypothesis**

- The p -value (less than 0.05) indicates the null hypothesis should be rejected

Conclusion: **The mean is greater than \$2.50**

Comments on p -value & degrees of freedom

- **p -value**: The probability, if the null hypothesis were correct, of getting as extreme, or more extreme, a result for the tested statistic (e.g., the estimated mean):
 - It is a measure of how extreme the statistic is
 - If the statistic is too extreme, we conclude that H_0 should be rejected
 - Typical p -value to reject H_0 : 0.10, 0.05 or 0.01 to be too extreme
- **Degrees of freedom (df)**: Represents the effective number of observations:
 - Usually, df is the number of observations minus the number of parameters being estimated

***P*-values: Six principles**

- ***P*-values can indicate how incompatible the data are with a specified statistical model**
- ***P*-values do *not* measure the probability that the studied hypothesis is true, or probability that the data were produced by random chance alone**
- **Scientific conclusions and business or policy decisions should not be based only on whether *p*-value passes a specific threshold**
- **Proper inference requires full reporting and transparency**
- ***P*-value, or statistical significance, does *not* measure the size of an effect or the importance of a result**
- **By itself, a *p*-value does *not* provide a good measure of evidence regarding a model of hypothesis**

Guidelines by the American Statistical Association (ASA)

Type I and Type II Errors, Power Function

	Decision	
Truth	Fail to reject H_0	Reject H_0
True	Correct	Type I Error
False	Type II Error	Correct (Power)

- **Type I Error:** Reject the null hypothesis H_0 , when H_0 is correct
 - The significance level α set before the test defines how much Type I Error we can tolerate
 - Typical values for $\alpha = 0.1, 0.05, 0.01$
- **Type II Error:** Fail to reject the null hypothesis H_0 , when H_0 is false
 - Fail to reject the null hypothesis when the true mean μ is unequal to μ_0 .
 - The probability that H_0 is not rejected when the true mean is μ is a function of μ :

$$\beta(\mu) = P_{\mu}\left\{-z_{\alpha/2} \leq \frac{\bar{m} - \mu_0}{\sigma/\sqrt{n}} \leq z_{\alpha/2}\right\}$$
- **Power function of the test ($1 - \beta(\mu)$):** The probability of rejection when μ is the true value
 - Type II error probability increases as μ and μ_0 get closer

Comparing **Two** Groups of Observations

- **Parametric vs. Nonparametric**
 - Parametric tests are more powerful if the underlying assumptions hold true → Always try parametric tests first
 - Nonparametric tests are more appropriate when the assumptions are grossly unreasonable (e.g., rank ordered data)
- **Dependent vs. Independent** Groups
 - **Paired Tests** (paired = TRUE) for *dependent* groups

Examples: Hypothesis Tests

Sample	Paired	Null Hypothesis	Assumptions	Test
One Sample		$H_0 : \mu = \mu_0$	i.i.d. $N(\mu, \sigma^2)$	t.test()
Two Samples	No	$H_0 : \sigma_1^2 = \sigma_2^2$	Normally distributed	F-test: var.test() Bartlett: bartlett.test()
Two Samples	No	$H_0 : \sigma_1^2 = \sigma_2^2$	Non-parametric	Ansari-Bradley test: ansari.test()
Two Samples	No	$H_0 : \mu_1 = \mu_2$	$\sigma_1^2 = \sigma_2^2$	t.test(var.equal=TRUE)
Two Samples	No	$H_0 : \mu_1 = \mu_2$	$\sigma_1^2 \neq \sigma_2^2$	Welch t-test t.test(var.equal=FALSE)
Two samples	No	$p_1(x) = p_2(x)$ p : probab. distr	Non-parametric	Wilcoxon rank sum wilcox.test ()
Two Samples	Yes	$H_0 : \mu_1 = \mu_2$	$\sigma_1^2 \neq \sigma_2^2$	t.test(paired=TRUE)
Two samples	Yes	$p_1(x) = p_2(x)$ p : probab. distr	Non-parametric	wilcox.test (paired=TRUE)

Non-parametric Test of Equal Variance

$$H_0 : \sigma_1^2 = \sigma_2^2$$

- Input: Two **independent** samples (i.e., two groups of observations)
- Null Hypothesis: The variances of two populations are equal
- Assumption: The data does not appear to be normally distributed
 - Hence, parametric tests can not be applied:
 - Neither F-test (var.test) nor Bartlett test can be applied
- **Ansari-Bradley Test**: `ansari.test()`
 - Non-parametric (no assumptions about population distribution)
 - Fail to reject the null hypothesis if the p -value is large, i.e.,
 - in this case, we conclude that the test indicates that the variances are equal

Ex: Ansari-Bradley Test: Equality of Variances

H_0 : The variances in tips between female and male groups are equal

```
> aggregate (tip ~ sex, data = tips, var)
```

Quick look into variances

```
      sex      tip
1 Female 1.344428
2  Male 2.217424
```

Ex: Ansari-Bradley Test: Equality of Variances

H_0 : The variances in tips between female and male groups are equal

```
> shapiro.test(tips$tip[tips$sex == "Female"])
```

shapiro-wilk normality test

```
data: tips$tip[tips$sex == "Female"]  
W = 0.9568, p-value = 0.005448
```

```
> shapiro.test(tips$tip[tips$sex == "Male"])
```

shapiro-wilk normality test

```
data: tips$tip[tips$sex == "Male"]  
W = 0.8759, p-value = 3.708e-10
```

*Check the assumptions:
test for normality of tip distributions*

- $p\text{-value} < 0.05$: **the null hypothesis should be rejected**
- Conclusion: groups are not normally distributed

```
> ansari.test(tip ~ sex, tips)
```

Ansari-Bradley test

```
data: tip by sex  
AB = 5582.5, p-value = 0.376  
alternative hypothesis: true ratio of scales  
is not equal to 1
```

*Assumption appears to be correct:
apply a non-parametric test*

- $p\text{-value} > 0.05$: **fail to reject the null hypothesis**
- According to this test, the results were not significant;
- Conclusion: the variances are equal

Ex: T-Test: Equality of Means

H_0 : Female and male groups are, on average, tipped equally

- Based on the Ansari-Bradley test, the variances in tips between two groups are equal
- Hence, a standard two sample t-test can be used rather than the Welch test for unequal variances

*Check the assumptions:
test for equal variances*

*Assumption appears to be correct:
apply a standard two sample t-test*

```
> t.test (tip ~ sex, data = tips, var.equal=TRUE)
```

```
Two Sample t-test
```

```
data:  tip by sex
t = -1.3879, df = 242, p-value = 0.1665
alternative hypothesis: true difference in means
is not equal to 0
95 percent confidence interval:
 -0.6197558  0.1074167
sample estimates:
mean in group Female    mean in group Male
      2.833448           3.089618
```

- $p\text{-value} > 0.05$: **fail to reject the null hypothesis**
- According to this test, the results were not significant;
- Conclusion: female and male workers are tipped roughly equally

Paired Two-Sample T-Test: Dependent Groups

H_0 : Fathers and sons have equal heights, on average

```
install.packages("UsingR")  
require(UsingR)  
head(father.son)
```

Check the assumptions:

- *test for normal distribution*
- *test for equal variances*

```
> t.test(father.son$fheight, father.son$sheight, paired=TRUE)
```

Paired t-test

```
data: father.son$fheight and father.son$sheight  
t = -11.7885, df = 1077, p-value < 2.2e-16  
alternative hypothesis: true difference in means is not equal  
to 0  
95 percent confidence interval:  
 -1.1629160 -0.8310296  
sample estimates:  
mean of the differences  
 -0.9969728
```

Reject H_0

- $p\text{-value} < 0.05$: **the null hypothesis should be rejected**
- Conclusion: fathers and sons (at least for this data set) have different heights

Wilcoxon Rank Sum Test

Non-parametric comparison of two (in)dependent groups

H_0 : Both groups are sampled from the same probability distribution:

$$p_1(x) = p_2(x)$$

- Assumptions for using Wilcoxon Rank Sum Test: `wilcox.test()`:
 - Two groups are *independent*
 - If two groups are *dependent* then use parameter **paired = TRUE**
 - Unable to meet the parametric assumptions of a t-test or ANOVA
 - Outcome variables are severely *skewed* or
 - Outcome variables are *ordinal* in nature (**rank ordered data**):
 - Probability of obtaining higher scores is greater in one population than the other

Example: Wilcoxon Rank Sum Test

Non-parametric comparison of two independent groups

H_0 : Incarceration rates are the same in Southern & non-Southern states

```
library(MASS)
head(UScrime)
# So: Southern vs non-Southern state
# Prob: Probability of incarceration
# (i.e., being imprisoned if committed a crime)
with(UScrime, by(Prob, So, median))

wilcox.test(Prob ~ So, data = UScrime)
```

wilcoxon rank sum test

data: Prob by So

w = 81, p-value = 8.488e-05 ←

Reject H_0

alternative hypothesis: true location shift
is not equal to 0

- $p\text{-value} < 0.05$: the null hypothesis should be rejected
- Conclusion: incarceration rates are not the same

Example: Paired Wilcoxon Signed Rank Test

Non-parametric comparison of two dependent groups

H_0 : Unemployment rates are the same
for younger and older males in Alabama

*Check the assumptions:
null hypothesis for
normality is rejected*

```
library(MASS)
head(UScrime)
sapply(UScrime[c("U1", "U2")], median)
with (UScrime, wilcox.test(U1, U2, paired = TRUE))
```

```
wilcoxon signed rank test with continuity
correction
```

```
data:  U1 and U2
V = 1128, p-value = 2.464e-09 ← Reject  $H_0$ 
alternative hypothesis: true location shift
is not equal to 0
```

- $p\text{-value} < 0.05$: **the null hypothesis should be rejected**
- Conclusion: unemployment rates are not the same

Example: Paired T-Test

Parametric comparison of two **dependent** groups

H_0 : Unemployment rates are the same
for younger and older males in Alabama

```
library(MASS)
head(UScrime)
sapply(UScrime[c("U1", "U2")],
       function(x) c(mean=mean(x), sd=sd(x)))
with (UScrime, t.test(U1, U2, paired = TRUE))
```

Paired t-test

```
data:  U1 and U2
t = 32.4066, df = 46, p-value < 2.2e-16
alternative hypothesis: true difference in
means is not equal to 0
95 percent confidence interval:
 57.67003 65.30870
sample estimates:
mean of the differences
      61.48936
```

← **Reject H_0**

- the mean difference (61.5) is large to warrant rejection of H_0 that the mean unemployment rate for older and younger males is the same
- younger males have a higher rate
- **probability of obtaining a sample difference that large if population means are equal is $2.2e^{-16}$**

Comparing More than Two Groups

- **Parametric vs. Nonparametric**
 - Parametric tests: **ANOVA** ← later as part of Experiment Design
 - Nonparametric tests: **Kruskal Wallis or Friedman**
- **Dependent vs. Independent Groups : Nonparametric Tests**
 - Independent Groups: **Kruskal Wallis** Test: `kruskal.test()`
 - Dependent Groups: **Friedman** Test: `friedman.test()`

Hypothesis Testing

STATISTICAL SIMULATION

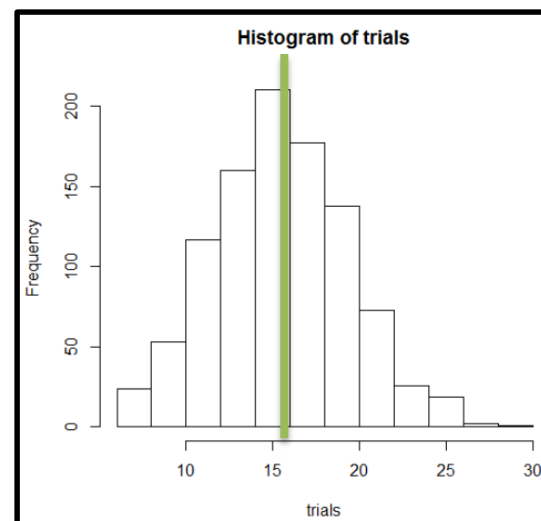
Null Hypothesis via Simulation

Null Hypothesis, H_0	Hat Model for Null Hypothesis	Quantitative Translation
Advertisements A and B are equally good (Goal: Reject H_0)	A single hat with 0s (no clicks) and 1s (clicks) where the 1s are the total clicks (A+B) and 0s are the total page-views	Ads A and B have the same click-through rate
A new targeted chemotherapy for advanced breast cancer is no more effective than the standard tamoxifen (Goal: Reject H_0)	A single hat for all patients (regardless of the group), with the number of days the person survived as values	Median survival time in a clinical trial is the same for both groups
A new & costly manufacturing process will not increase the chip-processing speed worth the investment (Goal: Reject H_0)	A single hat with all chips tested (both new and existing) in the sample, with the processing speed as values	Mean processing speed in a pilot new-process sample falls short of a 25% improvement

Example: Hypothesis Testing via Simulation

Null Hypothesis: Reduction in the return rate to 8% have occurred by chance

- An online merchant has historically experienced 10% return rate in the “kitchen gadget” category.
- To increase returns, the merchant does a pilot in which it adds additional explanatory information and pictures about products to its website
- Out of the next 200 purchases, 16 (8%) are returned.
- Is the pilot effective?



- 16 or fewer returns are not unusual
- $p\text{-value} = 0.564 \rightarrow$ Fail to Reject Null Hypothesis

```
5 # Hat: One one and nine zeros,  
6 # representing the null model of 10% returns  
7 hat <- c(rep(1,1),rep(0,9))  
8  
9 resample.statistic <- function() {  
10   draw <- sample(hat, 200, replace = TRUE)  
11   ret <- sum(draw)  
12   return(ret)  
13 }  
14  
15 n <- 1000  
16 trials <- replicate(n, res())  
17 pval <- sum(ifelse(trials<=16,1,0))  
18 cat("Estimated p-value: ",pval/n,"\n")  
19 hist(trials)
```

Basic Two-Sample Hypothesis Test: Concept

1. Establish a null model, or the **null hypothesis**
 - This represents a world in which nothing unusual is happening except by chance
 - Often, this null model is that the two samples come from the same population
2. Examine **pairs of resamples** drawn repeatedly from the null model to see how much they differ from one another
 - Alternatively, use formulas to learn about distribution of sample differences
 - If the observed difference is rarely encountered in this chance model, then we **Reject the Null Hypothesis**: random chance is not responsible

Basic Two-Sample Hypothesis Test: Details

1. Make sure you clearly understand:
 - the sizes of the two original samples
 - the statistic used to measure the difference between sample A and sample B:
 - the difference in means, proportions, ratio of proportions
 - the value of that statistic for the original two samples
2. Create a **hat** that represent the **null model**
 - e.g.: a hat with all the observed body weights in sample A and sample B
3. Draw two resamples of the same size as the original samples from the hat
 - with or without replacement: similar results unless small samples (<10)
4. Record the value of the statistic of interest
5. Repeat steps 3 and 4 many times for 1,000 trials
 - more trial can produce greater accuracy
6. Note a proportion of trials that yields a value for the statistic as large as that observed (known as **p-value**)

Alternative H_1 : Hypothesis Tests

The **alternative hypothesis** is the theory you would like to accept, assuming that your results disprove the null hypothesis.

Null Hypothesis, H_0 :

- No-fault reporting in hospitals is no more effective than the regular systems

Quantitatively:

- The no-fault system and the regular one both reduce errors to the same degree

Hat Model for Null H_0 :

- A single hat with the total number of errors for both groups

Alternative Hypothesis, H_1 :

- No-fault reporting in hospitals is **BETTER** the regular system; it reduces errors more
- **One-way**: Hypothesis is the question of whether a treatment is *better* than the control

Average error reduction

Control	1.88
No-fault	2.80
Difference	0.92