

MA-485
HW8

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Exercise 7.1

Ans: $X = \text{exponential}(1)$.

$$f(x) = \begin{cases} e^{-x}, & x \in [0, \infty) \\ 0 & \text{otherwise.} \end{cases}$$

The range of x is $[0, \infty]$.

(a) $y = 2x - 1$

$$\begin{aligned} \text{if } x=0, \quad y &= 2(0) - 1 \\ &= 0 - 1 \\ &= -1 \\ &\equiv \end{aligned}$$

$$\text{if } x=\infty, \quad y = 2(\infty) - 1 \\ = \infty$$

\therefore Range of y is $[-1, \infty]$

(b) $y = \frac{1}{x+1}$

$$\text{if } x=0 \Rightarrow y = \frac{1}{0+1} = \frac{1}{1} = 1$$

$$\text{if } x=\infty, \Rightarrow y = \frac{1}{\infty+1} = 0$$

Range of y is $0 \leq y \leq 1$

(2)

$$Y = \frac{1}{X+1} \text{ with range } 0 \leq Y \leq 1$$

Exercise 7.2

Ans. $Y = \frac{1}{X-1}$

$$X = U(0, 2)$$

Range of Y :-

$$\text{If } X=0, \text{ then } Y = \frac{1}{0-1} = -1$$

$$\text{If } X=2, \text{ then } Y = \frac{1}{2-1} = 1$$

\therefore Range of Y :- -1 to 1 or
 $-1 \leq Y \leq 1$

SExercise 7.3

Ans. Suppose X is a uniform random variable on the interval $(0, 1)$.

(3)

$$F(x) = \frac{1}{1-x} = 1, 0 < x < 1$$

$$\text{Here } Y = x^3$$

$$\therefore x = y^{1/3}$$

$$= g^{-1}(y)$$

$$\text{Hence } \left| \frac{dx}{dy} \right| = \left| \frac{1}{3} y^{-2/3} \right|$$

$$= \frac{1}{3 y^{2/3}}$$

Range of y if $x \geq 0, Y \geq 0$ & $x=1 \Leftrightarrow y \geq 1$.

$$\therefore 0 < x < 1 \Rightarrow 0 < y < 1$$

$$h(y) = f(x) \left| \frac{dx}{dy} \right|$$

$$\therefore h(y) = F \left[g^{-1}(y) \right] \left| \frac{dx}{dy} \right|$$

$$\therefore h(y) = 1 \times \frac{1}{3 y^{2/3}}, 0 < y < 1$$

(4)

$$\therefore f(y) = \frac{1}{3y^{2/3}}, \quad 0 < y < 1$$

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Exercise 7.4

Ans: $0 < x \leq 1 \Rightarrow 0 < 4x \leq 4 \Rightarrow 0 - 7 < 4x - 7 \leq 4 - 7$
 $\Rightarrow -7 < 4x - 7 < -3$

Range (y) = $(-7, -3)$

$$F_x(x) = P(X \leq x) = \int_{-\infty}^x f_{4x}(u) du$$

$$= \begin{cases} 0 & ; \text{ if } x < 0 \\ \int_{-\infty}^x 0 du & ; \text{ if } 0 \leq x < 1 \\ \int_{-\infty}^x 1 du & ; \text{ if } x \geq 1 \end{cases}$$

$$F_x(x) = \begin{cases} 1 & ; \text{ if } x \in (0, 1) \\ 0 & \text{otherwise} \end{cases}$$

$$F_X(x) = \begin{cases} 0 & ; \text{ if } x < 0 \\ x & ; \text{ if } 0 \leq x < 1 \\ 1 & ; \text{ if } x \geq 1 \end{cases}$$

$$F_y(y) = P(Y \leq y) = P(4x - 7 \leq y) \\ = P\left(X \leq \frac{y+7}{4}\right)$$

(5)

$$F_X\left(\frac{y+7}{4}\right) = \begin{cases} 0 & \text{if } \frac{y+7}{4} < 0 \\ \frac{y+7}{4} & \text{if } 0 \leq \frac{y+7}{4} < 1 \\ 1 & \text{if } \frac{y+7}{4} \geq 1 \end{cases}$$

$$= \begin{cases} 0 & \text{if } y < -7 \\ \frac{y+7}{4} & \text{if } -7 \leq y < -3 \\ 1 & \text{if } y \geq -3 \end{cases}$$

$$F_Y(y) = F_X\left(\frac{y}{4}\right) = \begin{cases} 0 & \text{if } y < -7 \\ \frac{y}{4} & \text{if } -7 \leq y < -3 \\ 1 & \text{if } y \geq -3 \end{cases}$$

$$= \begin{cases} 1/4 & \text{if } -7 \leq y < -3 \\ 0 & \text{otherwise} \end{cases}$$

Yes, the variable Y is uniformly distributed
 $Y \sim U(-7, -3)$

 \equiv

Giving 7.5

Ans: Given $X \sim U(-2, 2)$

$$\text{So, } f_X(x) = \begin{cases} \frac{1}{4}, & -2 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

(6)

Now ~~half~~ half of X :

$$F_Y(y) = P(Y \leq -1)$$

$$= P(|X| \leq y)$$

$$= P(-y \leq X \leq y)$$

$$\Rightarrow \int_{-y}^y \frac{1}{4} dx = \frac{2y}{4} = \frac{y}{2}, 0 < y < 2$$

Hence, $f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{1}{2}, 0 < y < 2$

Yes Y is a uniform R.V. in $(0, 2)$, i.e.,
 ~~X~~ $\sim U(0, 2)$.

Exercise 7.6

Ans: $F_Y(y) = P(Y \leq y) = P(\tan(X) \leq y)$

This happens precisely when $P(Y \leq y) = P(X \leq \arctan(y))$

But as $X \sim U\left(-\frac{\pi}{2}, \frac{\pi}{2}\right)$ we have:

(7)

$$P(X \leq \arctan(y)) = \frac{1}{2} + \frac{1}{\pi} \arctan(y), \text{ as the}$$

cumulative distribution of this variable is:-

$$F_x(t) = \frac{t + \frac{\pi}{2}}{\pi} = \frac{1}{2} + \frac{t}{\pi}$$

$\therefore F_x(y) = \frac{1}{2} + \frac{1}{\pi} \arctan(y)$ is the required cumulative distribution of this variable.

PDF on differentiating :-

$$f_x(y) = \frac{d}{dy} F_x(y) = \frac{d}{dy} \left(\frac{1}{2} + \frac{1}{\pi} \arctan(y) \right)$$

i.e. $f_x(y) = \frac{1}{\pi(1+y^2)}$ is the distribution function.

Exercise 8.1

Ans. By the normal distribution table :-

$$\begin{aligned} P(-0.9 < Z < 2.65) &= \phi(2.65) - \phi(-0.9) \\ &\approx 0.81915 \end{aligned}$$

(8)

$$(a) P(Z > -0.29) = 1 - \Phi(-0.29) = 0.614092$$

$$(b) P(-1.65 < Z < -0.18) = \Phi(-0.18) - \Phi(-1.65) \\ = 0.379105$$

$$(c) P(|Z| \leq 2.2) = \Phi(2.2) - \Phi(-2.2) \\ = 0.972193$$

Exercise 8.2

Given :- $X \sim N(3, 9)$ i.e. $\mu = 3$, $\sigma^2 = 9$

$$(a) P(2 < X < 5) = P\left[\frac{2-3}{3} < \frac{X-\mu}{\sigma} < \frac{5-3}{3}\right] \\ = P[-0.3333 < Z < 0.6667] \\ = P[Z < 0.6667] - P[Z < -0.3333] \\ = 0.7475 - 0.3694 \\ = 0.3781$$

(9)

(a)

$$P[X > 0] = 1 - P[X \leq 0]$$

$$= 1 - P\left[\frac{X-\mu}{\sigma} \leq \frac{0-3}{3}\right]$$

$$= 1 - P[Z \leq -1]$$

$$= 1 - 0.1587$$

$$= 0.8413$$

(b)

$$P[|X-3| > 6] = P[X-3 < -6] + P[X-3 > 9]$$

$$= P[X < -6+3] + P[X > 6+3]$$

$$= P[X < -3] + P[X > 9]$$

$$= P\left[\frac{X-\mu}{\sigma} < \frac{-3-3}{3}\right]$$

$$+ P\left[\frac{X-\mu}{\sigma} > \frac{9-3}{3}\right]$$

$$= P[Z < -2] + P[Z > 2]$$

$$= 0.0227 + 0.0227$$

$$= 0.0454$$

≈

(10)

Exercise 8.3

Ans:

Given:-

Mean, $\mu = 180$ Variance, $\sigma^2 = 900$ Standard deviation, $\sigma = 30$

$$P(X > 240) = 1 - P(X < 240)$$

$$= 1 - P \left[(X - \mu) / \sigma < (240 - 180) / 30 \right]$$

$$= 1 - P(z < 2)$$

$$= 1 - 0.9772$$

$$= 0.0228 \times 100$$

$$= 2.28\%$$

B

Exercise 8.4

Ans:

$$\mu = 270$$

$$\sigma^2 = 100$$

$$\sigma = \sqrt{100} = 10$$

$$P(245 < X < 295) = ?$$

$$z = \frac{X - \mu}{\sigma} = \frac{245 - 270}{10} = -2.5$$

$$z = \frac{X - \mu}{\sigma} = \frac{295 - 270}{10} = 2.5$$

(11)

$$\text{So, } P(245 < z < 295) = P(-2.5 < z < 2.5)$$

$$= P(z < 2.5) - [1 - P(z < 2.5)] \\ = 0.9938 - 0.0062 \\ = 0.9876$$

∴ Probability that he could nevertheless be the father of the child = $1 - P(245 < z < 295)$

$$= 1 - 0.9876 \\ = 0.0124 \\ \underline{\underline{}}$$

Exercise 8.5

Ans: $I = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)/2} dx dy$

$R = r \cos \theta, y = r \sin \theta$ with $0 < r < \infty$
 $0 < \theta < 2\pi$

$$\therefore I = \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} \cdot r dr d\theta$$

$$= [0]_0^{2\pi} [-e^{-r^2/2}]_0^{\infty}$$

$$= 2\pi(0+1) \\ = 2\pi$$

(12)

$$\Rightarrow \int_{-\infty}^{\infty} e^{-x^2/2} dx \cdot \int_{-\infty}^{\infty} e^{-y^2/2} dy = 2\pi$$

$$\Rightarrow \left(\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx \right) \left(\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-y^2/2} dy \right) = 1 \times 1$$

$$\Rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx = 1$$

Hence proved

Exercise 8.6

Quesn 8.6 Let Z be a standard normal variable.

Probability density function is given by:

$$f_Z(z) = \begin{cases} \frac{1}{\sqrt{2\pi}} e^{-z^2/2}, & -\infty < z < \infty \\ 0, & \text{otherwise} \end{cases}$$

Let $Y = e^Z$

$$P(Y \leq y) = F_Y(y)$$

(B)

$$F_X(y) = P[e^X \leq y]$$

$$= P[\tau \leq \ln y]$$

$$= \Phi(\ln y)$$

We get the density by differentiating :-

$$f_Y(y) = \frac{1}{y} \frac{d}{dy} \Phi(\ln y)$$

$$= \frac{1}{y} \frac{1}{\sqrt{2\pi}} e^{-\frac{(\ln y)^2}{2}}$$

$$f_Y(y) = \frac{1}{y} \frac{1}{\sqrt{2\pi}} e^{-\frac{(\ln y)^2}{2}} ; 0 < Y < \infty$$

which is the pdf of a log normal distribution with parameters $\mu=0, \sigma^2=1$

(a) Here, X is a standard normal variable,
i.e. $X \sim N(0, 1)$

∴ The pdf of X is :-

(14)

$$f_2(x) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}} ; -\infty < x < \infty$$

Let $\Phi(x) = \text{cdf of } X = P(X \leq x)$.

Let another continuous random variable

$$\text{Y be } Y = \sqrt{|x|}$$

Now, the cdf of Y is,

$$\Rightarrow F_Y(y) = P(Y \leq y)$$

$$= P(\sqrt{|x|} \leq y)$$

$$= P(|x| \leq y^2)$$

$$= P(-y^2 \leq x \leq y^2)$$

$$= \Phi(y^2) - \Phi(-y^2)$$

$$= \Phi(y^2) - [1 - \Phi(y^2)]$$

$$= 2\Phi(y^2) - 1, \text{ from } 0 \leq y < \infty.$$

$$\left[\begin{array}{l} \because \Phi(-k) = 1 - \Phi(k) \\ (\forall k) \end{array} \right]$$

$$\therefore \text{We get } \therefore F_Y(y) = \begin{cases} 0 &; y < 0 \\ [2\Phi(y^2) - 1] &; y \geq 0 \end{cases}$$

(15)

Hence the pdf of Y is

$$f_Y(y) = \frac{d}{dy} [F_Y(y)]$$

$$\begin{aligned} \left[\frac{d}{dy} [\Phi(x)] \right] &= \frac{d}{dy} [2 \cdot \Phi(y^2) - 1] \\ \left[= f_X(x) \right] &= 2 \cdot d_x(y^2) \cdot (2y) \\ &= 4y f_X(y^2) \end{aligned}$$

$$\left(\text{Using the form of } f_X(x) \right) = 4y \cdot \left[\frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{(y^2)^2}{2}} \right]$$

$$= 2 \cdot \frac{\sqrt{2}}{\sqrt{\pi}} \cdot y \cdot e^{-\frac{y^4}{2}} ; \text{ from } y \geq 0,$$

\therefore The pdf of Y is

$$\Rightarrow f_Y(y) = \begin{cases} \frac{2\sqrt{2}}{\sqrt{\pi}} \cdot y \cdot e^{-\frac{y^4}{2}} ; \text{ from } 0 \leq y < \infty \\ 0 ; \text{ otherwise.} \end{cases}$$

Z