# ECE 361 Probability for Engineers (Fall, 2016) Lecture 6a

### §3.2 Cumulative distribution functions

**Example.** The maximum of several discrete uniform RVs. Let  $X = \max\{X_1, \dots, X_k\}$  where  $X_i \sim \text{Uni}([n])$  for  $i \in [k]$ . Find the PMF for  $X, p_X$ . We first find the CDF  $F_X$  for  $j \in [n]$ :

$$F_X(j) = \mathbb{P}(X \le j) = \mathbb{P}(\max\{X_1, \dots, X_k\} \le j) = \mathbb{P}(X_1 \le j, \dots, X_k \le j) = \mathbb{P}(X_1 \le j) \cdots \mathbb{P}(X_k \le j) = F_{X_1}(j)^k. \tag{1}$$

Thus  $F_X(j) = (j/n)^k$  and the PMF is:

$$p_X(j) = F_X(j) - F_X(j-1) = \left(\frac{j}{n}\right)^k - \left(\frac{j-1}{n}\right)^k, \ j \in [n].$$
 (2)

Now consider the same argument for continuous uniform RVs. Let  $X = \max\{X_1, \dots, X_k\}$  where  $X_i \sim \text{Uni}([0, t])$  for  $i \in [k]$ . Find the PDF for X,  $f_X$ . We first find the CDF  $F_X$  for  $s \in [0, t]$ :

$$F_X(s) = \mathbb{P}(X \le s) = \mathbb{P}(\max\{X_1, \dots, X_k\} \le s) = \mathbb{P}(X_1 \le s, \dots, X_k \le s) = \mathbb{P}(X_1 \le s) \cdots \mathbb{P}(X_k \le s) = F_{X_1}(s)^k$$
(3)

Thus  $F_X(s) = (s/t)^k$  and the PDF is:

$$f_X(s) = \frac{\mathrm{d}}{\mathrm{d}s} F_X(s) = \frac{\mathrm{d}}{\mathrm{d}s} \left(\frac{s}{t}\right)^k = \frac{ks^{k-1}}{t^k} = \frac{k}{s} \left(\frac{s}{t}\right)^k, \ s \in [0, t]. \tag{4}$$

#### The geometric and exponential CDFs

Recall the geometric PMF  $(X \sim \text{Geo}(p))$  and exponential PDF  $(Y \sim \text{Exp}(\lambda))$ :

$$p_X(n) = (1-p)^{n-1}p, \ n \in \mathbb{N}$$
  

$$f_Y(y) = \lambda e^{-\lambda y}, \ y \in \mathbb{R}_+$$
(5)

We find their CDFs by summing and integrating, respectively:

$$F_X(k) = \sum_{l=1}^k p_X(l) = \sum_{l=1}^k (1-p)^{l-1} p = 1 - (1-p)^k, \ k \in \mathbb{N}$$

$$F_Y(y) = \int_0^y f_Y(t) dt = \int_0^y \lambda e^{-\lambda t} dt = 1 - e^{-\lambda y}, \ y \in \mathbb{R}_+$$
(6)

If we fix  $\lambda$  for Y and p for X and choose  $\delta = -\log(1-p)/\lambda$  then one can show  $F_Y(n\delta) = F_X(n)$  for each n. A plot of the CDFs nearly coincide, see Fig. 3.8.

# Example

Two towns, A and B are 50 miles apart. Car 1 leaves town A and car 2 leaves town B independently of one another. The departure time of each car is random and uniformly distributed between 12 and 1 o'clock. The speed of each car is 100 miles per hour and they travel directly to meet one another. Denote by X the distance between the meeting point and A. Find  $F_X$ , the CDF for X. Hints: i) let  $T_A, T_B$  be independent uniform RVs on [0,1], and let  $Y = T_A - T_B$  be their difference; ii) find the CDF for Y with support  $\mathcal{Y} = [-1,1]$  noting the two cases  $y \in [-1,0]$  and  $y \in [0,1]$ ; iii) find an expression for the RV X as a linear function of Y provided |Y| < 1/2, and consider X for Y > 1/2 and Y < -1/2 separately.

**Solution.** Let  $T_A, T_B$  be the random times that the two cars leave their respective times, where  $T_A \sim \text{Uni}[0,1]$  and  $T_B \sim \text{Uni}[0,1]$ . For  $|T_A - T_B| \ge 1/2$  the cars will meet at one of the towns: for  $T_A - T_B > 1/2$  they meet at A while for

 $T_B - T_A > 1/2$  they meet at B. For all other pairs  $(T_A, T_B)$  they will meet on the road. Let T be the time they meet, which is characterized by the fact that the sum of the distances they have traveled must equal 50:

$$100(T - T_A) + 100(T - T_B) = 50 \Leftrightarrow T = \frac{2(T_A + T_B) + 1}{4}.$$
 (7)

The random distance from A, say X, is  $X = 100(T - T_A) = 25 + 50(T_B - T_A)$  when  $|T_B - T_A| < 1/2$ . Combining:

$$X = \begin{cases} 50, & T_B - T_A > 1/2\\ 0, & T_B - T_A < -1/2\\ 25 + 50(T_B - T_A), & |T_B - T_A| < 1/2 \end{cases}$$
(8)

It remains to find the CDF of the difference of two uniform RVs. Let  $Y = T_B - T_A$  with support  $\mathcal{Y} = [-1, 1]$ . Then, using Fig. 1 the two cases are  $y \in [-1, 0]$  and  $y \in [0, 1]$ . For  $y \in [0, 1]$ :

$$F_Y(y) = 1 - \mathbb{P}(Y > y) = 1 - \mathbb{P}(T_B - T_A > y) = 1 - \int_0^{1-y} \mathbb{P}(T_B - T_A > y | T_A = t) f_{T_A}(t) dt$$

$$= 1 - \int_0^{1-y} (1 - F_{T_B}(t+y)) f_{T_A}(t) dt.$$
(9)

Using the distributions for  $T_A$  and  $T_B$  we obtain:

$$F_Y(y) = 1 - \int_0^{1-y} (1 - (t+y)) 1 dt = \frac{1}{2} + y - \frac{y^2}{2}.$$
 (10)

For  $y \in [-1, 0]$ :

$$F_Y(y) = \mathbb{P}(Y \le y) = \mathbb{P}(T_B - T_A \le y) = \int_{-y}^1 \mathbb{P}(T_B - T_A \le y | T_A = t) f_{T_A}(t) dt = \int_{-y}^1 F_{T_B}(t + y) f_{T_A}(t) dt. \tag{11}$$

Using the distributions for  $T_A$  and  $T_B$  we obtain:

$$F_Y(y) = \int_{-y}^1 (t+y) 1 dt = \frac{1}{2} + y + \frac{y^2}{2}.$$
 (12)

Combining, the CDF for Y is:

$$F_Y(y) = \begin{cases} \frac{1}{2} + y + \frac{y^2}{2}, & -1 \le y \le 0\\ \frac{1}{2} + y - \frac{y^2}{2}, & 0 \le y \le 1 \end{cases}$$
 (13)

We now express X in terms of Y:

$$X = \begin{cases} 0, & Y > 1/2\\ 50, & Y < -1/2\\ 25 + 50Y, & |Y| < 1/2 \end{cases}$$
 (14)

For the boundary cases:

$$\mathbb{P}(X=0) = \mathbb{P}(Y>1/2) = 1 - F_Y(1/2) = 1 - \left(\frac{1}{2} + \frac{1}{2} - \frac{(1/2)^2}{2}\right) = 1/8$$

$$\mathbb{P}(X=50) = \mathbb{P}(Y<-1/2) = F_Y(-1/2) = \frac{1}{2} + (-1/2) + \frac{(-1/2)^2}{2} = 1/8$$
(15)

For the remaining cases:

$$F_X(x) = \mathbb{P}(X \le x) = \mathbb{P}(25 + 50Y \le x) = \mathbb{P}(Y \le (x - 25)/50) = \begin{cases} \frac{1}{2} \left(\frac{25 + x}{50}\right)^2, & 0 < x \le 25\\ \frac{1}{2} \left(\frac{-x^2 + 150x - 625}{50^2}\right), & 25 < x < 50. \end{cases}$$
(16)

All together:

$$F_X(x) = \begin{cases} \frac{1}{8}, & x = 0\\ \frac{1}{2} \left(\frac{25+x}{50}\right)^2, & 0 < x \le 25\\ \frac{1}{2} \left(\frac{-x^2+150x-625}{50^2}\right), & 25 < x < 50\\ \frac{1}{8}, & x = 50 \end{cases}$$
(17)

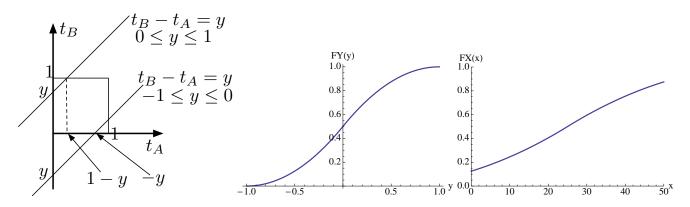


Figure 1: **Left:** the difference of two uniform RVs  $Y = T_B - T_A$  for  $T_A, T_B \sim \text{Uni}[0, 1]$  has support  $\mathcal{Y} = [-1, 1]$ , and the CDF is found by considering the two cases  $y \in [-1, 0]$  and  $y \in [0, 1]$ . **Middle:** the CDF for Y. **Right:** the CDF for X for 0 < x < 50. Note the CDF has  $F_X(0) = 1/8$  and  $F_X(50) = 1/8$ .

# References

[1] Introduction to Probability, 2nd Edition by Dimitri P. Bertsekas and John N. Tsitsiklis, Athina Scientific Press, 2008.