

ECE 361 Probability for Engineers (Fall, 2016)

Lecture 5b

§3.1 Continuous RVs and their PDFs

Expectation

The expectation of a continuous RV is:

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx \quad (1)$$

This definition should be expected given the corresponding definition for discrete RVs – recall that integration is the limiting process of summation. Similarly, for any function g :

$$\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx. \quad (2)$$

The variance of X is

$$\text{var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \int_{-\infty}^{\infty} (x - \mathbb{E}[X])^2 f_X(x) dx. \quad (3)$$

As with discrete RVs:

$$0 \leq \text{var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2. \quad (4)$$

For linear functions $Y = aX + b$:

$$\mathbb{E}[Y] = a\mathbb{E}[X] + b, \quad \text{var}(Y) = a^2 \text{var}(X). \quad (5)$$

Example. Mean and variance of a uniform RV. Let $X \sim \text{Uni}[a, b]$. The mean is:

$$\mathbb{E}[X] = \int_a^b x \frac{1}{b-a} dx = \frac{1}{b-a} \frac{1}{2} x^2 \Big|_a^b = \frac{a+b}{2}. \quad (6)$$

The second moment is:

$$\mathbb{E}[X^2] = \int_a^b x^2 \frac{1}{b-a} dx = \frac{1}{b-a} \frac{1}{3} x^3 \Big|_a^b = \frac{b^3 - a^3}{3(b-a)} = \frac{a^2 + ab + b^2}{3}. \quad (7)$$

The variance is:

$$\text{var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{a^2 + ab + b^2}{3} - \left(\frac{a+b}{2}\right)^2 = \frac{(b-a)^2}{12}. \quad (8)$$

Exponential RV

An exponential RV has PDF

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{else} \end{cases} \quad (9)$$

Note this is a legitimate PDF since it integrates to one:

$$\int_0^{\infty} f_X(x) dx = \int_0^{\infty} \lambda e^{-\lambda x} dx = -e^{-\lambda x} \Big|_0^{\infty} = 1 \quad (10)$$

We find the probability that X exceeds some value $a \in \mathbb{R}_+$:

$$\mathbb{P}(X \geq a) = \int_a^{\infty} f_X(x) dx = -e^{-\lambda x} \Big|_a^{\infty} = e^{-\lambda a}. \quad (11)$$

Thus the probability that $X > a$ decays exponentially in a at rate λ , hence the name for this distribution. We can find the mean and variance:

$$\mathbb{E}[X] = \frac{1}{\lambda}, \quad \text{var}(X) = \frac{1}{\lambda^2}. \quad (12)$$

To compute $\mathbb{E}[X]$ we require integration by parts:

$$\mathbb{E}[X] = \int_0^\infty x \lambda e^{-\lambda x} dx = (-xe^{-\lambda x}) \Big|_0^\infty + \int_0^\infty e^{-\lambda x} dx = 0 - \frac{e^{-\lambda x}}{\lambda} \Big|_0^\infty = \frac{1}{\lambda}. \quad (13)$$

Similarly for $\mathbb{E}[X^2]$:

$$\mathbb{E}[X^2] = \int_0^\infty x^2 \lambda e^{-\lambda x} dx = (-x^2 e^{-\lambda x}) \Big|_0^\infty + \int_0^\infty 2x e^{-\lambda x} dx = 0 + \frac{2}{\lambda} \mathbb{E}[X] = \frac{2}{\lambda^2}. \quad (14)$$

Combining these gives the variance:

$$\text{var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}. \quad (15)$$

I will write $X \sim \text{Exp}(\lambda)$ to denote an exponential RV with rate λ .

§3.2 Cumulative distribution functions

In Chapter 2 we used the probability mass function (PMF) to characterize discrete RVs and in §3.1 we used the probability density function (PDF) to characterize continuous RVs. We now introduce a third concept, the cumulative distribution function (CDF) that applies to both types of RVs. If X is an RV we say $X \sim F_X$ to mean X is a RV with CDF F_X , where:

$$F_X(x) = \mathbb{P}(X \leq x) = \begin{cases} \sum_{k \leq x} p_X(k), & \text{discrete} \\ \int_{-\infty}^x f_X(t) dt, & \text{continuous} \end{cases} \quad (16)$$

Note the helpful illustrations comparing PMFs, PDFs, and CDFs in Fig. 3.6 and 3.7. Key properties of CDFs include:

- To obtain the CDF from the PMF for a discrete RV we compute $F_X(x) = \sum_{k \leq x} p_X(k)$ for each $x \in \mathcal{X}$. To obtain the CDF from the PDF for a continuous RV we compute $F_X(x) = \int_{-\infty}^x f_X(t) dt$ for each x .
- To obtain the PMF from the CDF for a discrete RV we compute $p_X(k) = F_X(k) - F_X(k-1)$ at each k . To obtain the PDF from the CDF for a continuous RV we compute $f_X(x) = \frac{d}{dx} F_X(x)$.
- Also note the CDF is monotone non-decreasing in x : if $x \leq y$ then $F_X(x) \leq F_X(y)$. This is a simple consequence of the observation that $(-\infty, x] \subset (-\infty, x + \delta]$ and thus $\mathbb{P}(X \in (-\infty, x]) \leq \mathbb{P}(X \in (-\infty, x + \delta])$, i.e., $F_X(x) \leq F_X(x + \delta)$.
- Finally, note the limits:

$$\lim_{x \rightarrow -\infty} F_X(x) = 0, \quad \lim_{x \rightarrow \infty} F_X(x) = 1. \quad (17)$$

References

- [1] *Introduction to Probability, 2nd Edition* by Dimitri P. Bertsekas and John N. Tsitsiklis, Athina Scientific Press, 2008.