

# Financial Mathematics 33000

## Lecture 6

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## Fundamental theorem in continuous time

### Black-Scholes

# Fundamental theorem for continuous time Itô processes

$\exists$  equivalent martingale measure  $\mathbb{P} \Rightarrow$  No arb

Proof:

- ▶ Given: under  $\mathbb{P}$ , each member of the discounted asset prices

$$\frac{1}{B_t} \mathbf{X}_t := \left( \frac{X_t^1}{B_t}, \frac{X_t^2}{B_t}, \dots, \frac{X_t^N}{B_t} \right)$$

is a martingale Itô process. Need to prove: No arbitrage

- ▶ For any self-financing trading strategy  $\Theta_t$  with value  $V_t$ , we'll show that  $V/B$  is a martingale, where  $V_t := \Theta_t \cdot \mathbf{X}_t$ .

After we show this, we're done, by the familiar argument:

If  $V_0 = 0$ , then  $V_0/B_0 = 0$ , hence  $\mathbb{E}(V_T/B_T) = 0$ .

If also  $V_T \geq 0$ , then  $V_T/B_T \geq 0$ , so  $V_T/B_T = 0$ , hence  $V_T = 0$

Conclusion:  $\Theta$  is not an arbitrage.

# Fundamental theorem

- To see that  $V/B$  is a martingale, let  $A_t := 1/B_t$ . Then

$$\begin{aligned}
 d(V_t/B_t) &= d(A_t V_t) = V_t dA_t + A_t dV_t + dA_t dV_t \\
 &= \Theta_t \cdot \mathbf{X}_t dA_t + A_t (\Theta \cdot d\mathbf{X}_t) + (dA_t)(\Theta_t \cdot d\mathbf{X}_t) \\
 &= \Theta_t \cdot (\mathbf{X}_t dA_t + A_t d\mathbf{X}_t + dA_t d\mathbf{X}_t) \\
 &= \Theta_t \cdot d(A_t \mathbf{X}_t) = \sum_{n=1}^N \theta_t^n d(A_t X_t^n)
 \end{aligned}$$

Since each  $A_t X_t^n$  is a martingale,  $V/B$  is a martingale also.

- Idea: A martingale is the cumulative PnL from betting on zero- $\mathbb{E}$  games. Varying your bet size across games and across time still produces, collectively, a zero-expectation game. Can't risklessly make something from nothing by playing zero-expectation games.

# Fundamental theorem

No arb  $\Rightarrow \exists$  equivalent martingale measure  $\mathbb{P}$ :

Intuition of proof: Same as in L2, L3.

- ▶ Define  $\mathbb{P}$  by defining the  $\mathbb{P}_t$ -probability of an event to be the time- $t$  **price**, in units of  $B$ , of an (“Arrow-Debreu”) asset that pays 1 unit of  $B$  at time  $T$  if the event occurs, else 0. (But what if the A-D asset does not exist and can’t be replicated?)
- ▶ Martingale property holds because any asset  $X$  can be replicated by portfolio of  $X_T(\omega_j)/B_T(\omega_j)$  units of the A-D asset for each  $\omega_j$ . Value portfolio by summing **quantity**  $\times$  **price**.

$$\frac{X_0}{B_0} = \sum_{j=1}^J \frac{X_T(\omega_j)}{B_T(\omega_j)} \cdot \mathbb{P}(\omega_j) = \mathbb{E}\left(\frac{X_T}{B_T}\right)$$

$p(A) =$  PRICE of contract that pays:  $\begin{cases} \$1 & \text{if } A \\ \$0 & \text{if not } A. \end{cases}$

## Fundamental theorem: Comments

- ▶ Idea: The  $\mathbb{P}$  probability of an event is simply the *price* (in units of  $B$ ) of a asset that pays 1 unit of  $B$  iff that event occurs.
- ▶ Note: In this entire proof, we never assumed that  $B$  is the bank account, and never assumed that it is riskless. It is enough to assume that  $B$  is some asset with positive price process.

In some applications, it may be easier to normalize using some such asset (some *numeraire*) that is *not* the bank account.

By default, if we say risk-neutral or martingale measure without specifying the numeraire, it is understood to be the bank account.

# Option pricing

In L5, we did this by *replication*.

In L6, let's do it by martingale methods: Option price equals the *expected discounted payoff*, under a martingale measure  $\mathbb{P}$ . Why?

- ▶ By the Fundamental theorem.

How do we calculate  $\mathbb{P}$ -expectations (denoted by  $\mathbb{E}$ )?

- ▶ In many cases, a model is already specified under risk-neutral measure. Then simply work directly under the given measure.

But what if the model is specified under physical measure?

- ▶ We know how  $S$  behaves with respect to physical measure  $\mathbb{P}$ .

How does  $S$  behave wrt  $\mathbb{P}$ ? All risk driven by  $W$ . So let's see what changing measure does to  $W$ , then find what it does to  $S$ .

# Girsanov's theorem

Theorem: If  $W$  is a Brownian motion under  $\mathbb{P}$ ,  
and if  $\mathbb{P}$  is a probability measure on  $\mathcal{F}_T^W$  that is equivalent to  $\mathbb{P}$ ,  
then there exists an adapted process  $\lambda$  such that for all  $t \in [0, T]$ ,

$$\tilde{W}_t := W_t + \int_0^t \lambda_s ds$$

is Brownian motion under  $\mathbb{P}$ . Therefore:

- ▶  $d\tilde{W}_t = dW_t + \lambda_t dt$ , and  $\tilde{W}$  is BM under  $\mathbb{P}$  but not under  $\mathbb{P}$
- ▶  $dW_t = d\tilde{W}_t - \lambda_t dt$ , and  $W$  is BM under  $\mathbb{P}$  but not under  $\mathbb{P}$



# Girsanov: an analogy

skip

No proof, but here is an *analogy* on a sample space  $\Omega = \{\omega_1, \dots, \omega_6\}$ .

Let  $X(\omega_1) = X(\omega_2) = X(\omega_3) = 10$ ,  $X(\omega_4) = X(\omega_5) = X(\omega_6) = 25$ .

- ▶ Let  $P(\omega) = 1/6$  for each  $\omega$ . Then  $X \sim \text{Uniform}\{10, 25\}$  under  $P$ .
- ▶ But if  $\mathbb{P}$  assigns probability  
 $1/12$  to each of  $\omega_1, \omega_2, \omega_3$ ,  
 and  $1/4$  to each of  $\omega_4, \omega_5, \omega_6$ ,  
 then  $X$  is *not*  $\text{Uniform}\{10, 25\}$  under  $\mathbb{P}$ .
- ▶ However,  $\tilde{X} := X + \lambda$  is  $\text{Uniform}\{10, 25\}$  under  $\mathbb{P}$ ,  
 where  $\lambda(\omega_4) := 15$  and  $\lambda(\omega) := 0$  for  $\omega \neq \omega_4$ .

$X$  under  $\mathbb{P}$  does not have the same distribution as  $X$  under  $P$ .

But  $X$  *plus drift* under  $\mathbb{P}$  has the same distribution as  $X$  under  $P$ .

# Girsanov: some intuition

skip

No proof, but here is some *intuition*:

- $W$  is BM under  $\mathbb{P}$ . After changing measure to  $\mathbb{P}$ , the  $W$  may not still be BM, but it is plausible that it is a martingale plus drift:

$$dW_t = \lambda_t dt + \sigma_t dB_t$$

where  $B$  is a BM under  $\mathbb{P}$ , and  $\sigma_t$  is some adapted process. So

$$(dW_t)^2 = (\lambda_t dt + \sigma_t dB_t)^2$$

hence  $dt = \sigma_t^2 dt$ , so  $\sigma_t = \pm 1$ . Define  $\tilde{W}$  by  $d\tilde{W}_t = \sigma_t dB_t$ .

- Then  $W$  can be shown to be  $\mathbb{P}$ -BM. And, as claimed,

$$dW_t = \lambda_t dt + d\tilde{W}_t.$$

Fundamental theorem in continuous time

Black-Scholes

# Black-Scholes via martingale approach

Black-Scholes dynamics

$$dB_t = rB_t dt \qquad B_0 = 1$$

$$dS_t = \mu S_t dt + \sigma S_t dW_t \qquad S_0 > 0$$

where  $W$  is BM under physical measure  $\mathbb{P}$ , and  $\sigma > 0$ .

No arb implies that  $\exists \mathbb{P}$ , equivalent to  $\mathbb{P}$ , such that  $S/B$  is a  $\mathbb{P}$ -MG.

Hence by Girsanov,  $\exists \lambda$  such that  $\tilde{W}_t := W_t + \int_0^t \lambda_s ds$  is  $\mathbb{P}$ -BM.

Substitute  $d\tilde{W}_t = dW_t + \lambda_t dt$  into the SDE of  $S$ :

$$\begin{aligned} dS_t &= \mu S_t dt + \sigma S_t (d\tilde{W}_t - \lambda_t dt) \\ &= (\mu - \lambda_t \sigma) S_t dt + \sigma S_t d\tilde{W}_t \end{aligned}$$

But can we say anything about  $\mu - \lambda_t \sigma$ ?

Under  $\mathbb{P}$ , every tradeable asset  $X$  has drift  $rX$

This page does not assume that  $X$  is a GBM.

Assume only that  $X$  is an Itô process.

- ▶ Under  $\mathbb{P}$ , the discounted price  $X/B$  is a MG, hence has zero drift.
- ▶ By Itô's rule,  $X/B$  has dynamics

$$\begin{aligned} d(X_t/B_t) &= d(e^{-rt}X_t) = e^{-rt}dX_t - re^{-rt}X_tdt + d(e^{-rt})dX_t \\ &= e^{-rt}(dX_t - rX_tdt), \end{aligned}$$

so  $dX_t - rX_tdt$  has no drift term.

- ▶ Therefore the drift term of  $dX_t$  must be  $rX_tdt$ .

Under  $\mathbb{P}$ , the GBM  $S$  is still GBM, but with drift  $r$

- Applying this to  $S$ , we have  $(\mu - \lambda_t \sigma)S_t = rS_t$ , and

$$\boxed{dS_t = rS_t dt + \sigma S_t d\tilde{W}_t}$$

where  $\tilde{W}$  is  $\mathbb{P}$ -BM.

Rate of growth changes from  $\mu$  to  $r$ . Volatility stays the same.

- By L4, therefore, under  $\mathbb{P}$ , conditional on  $\mathcal{F}_t^W$

$$\boxed{\log S_T \sim \text{Normal}(\log S_t + (r - \sigma^2/2)(T - t), \sigma^2(T - t))}$$

Compare: under  $\mathbb{P}$ , conditional on  $\mathcal{F}_t^W$ ,

$$\log S_T \sim \text{Normal}(\log S_t + (\mu - \sigma^2/2)(T - t), \sigma^2(T - t))$$

# Lognormal distribution

Here's a more general calculation, allowing different rates for growth and discounting, on an underlying  $X$ , not necessarily a stock price.

- ▶ Let  $t < T$ . Let  $R_{grow}$  and  $r$  be constants.
- ▶ Assume that (conditional on the time- $t$  information  $\mathcal{F}_t$ ) the random variable  $X_T$  has lognormal  $\mathbb{P}$ -distribution

$$\log X_T \sim \text{Normal}(\log X_t + (R_{grow} - \sigma^2/2)(T - t), \sigma^2(T - t))$$

where  $X_t > 0$ , and  $\sigma > 0$  is a constant.

- ▶ One way that this distribution could arise is from the dynamics

$$dX_t = R_{grow}X_t dt + \sigma X_t dW_t \quad X_0 > 0$$

where  $W$  is  $\mathbb{P}$ -BM.

## Conclusion: the Black-Scholes call price formula

Then, letting  $\mathbb{E}$  denote expectation wrt  $\mathbb{P}$ ,

$$e^{-r(T-t)}\mathbb{E}_t(X_T - K)^+ = C^{BS}(X_t, t, K, T, R_{grow}, r, \sigma)$$

where the function  $C^{BS}$  is defined for  $X > 0, K > 0, \sigma > 0, t < T$  by

$$C^{BS}(X, t, K, T, R_{grow}, r, \sigma) := e^{-r(T-t)}[FN(d_1) - KN(d_2)],$$

and

$$F := Xe^{R_{grow}(T-t)} = \mathbb{E}_t X_T$$

and

$$d_{1,2} := d_{+,-} := \frac{\log(F/K)}{\sigma\sqrt{T-t}} \pm \frac{\sigma\sqrt{T-t}}{2}.$$



Proof of formula: decompose  $(X_T - K)^+$   $\mathbb{1}_A = \begin{cases} 1 & \text{if } A \\ 0 & \text{if not } A \end{cases}$

Proof: Using the fact that  $x^+ = x\mathbf{1}_{x>0}$  for all  $x$ ,

$$\mathbb{E}\mathbb{1}_A = \mathbb{P}(A)$$

$$\begin{aligned} \mathbb{E}_t(X_T - K)^+ &= \mathbb{E}_t(X_T - K)\mathbf{1}_{X_T > K} \\ &= \mathbb{E}_t(X_T\mathbf{1}_{X_T > K}) - K\mathbb{E}_t\mathbf{1}_{X_T > K} \\ &= \underbrace{F_t\mathbb{E}_t((X_T/F_t)\mathbf{1}_{X_T > K})}_{\text{will see this is } N(d_1)} - K\underbrace{\mathbb{P}_t(X_T > K)}_{\text{will see this is } N(d_2)}, \end{aligned}$$

where

$$F_t := X_t e^{R_{grow}(T-t)} = \mathbb{E}_t X_T.$$

To proceed, first note that the event  $X_T > K$  is equivalent to  $\log(X_T/F_t) > \log(K/F_t)$ , and that (conditional on time- $t$  information)

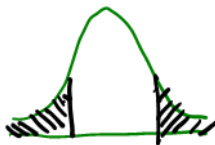
$$\log(X_T/F_t) \sim \text{Normal}(-(\sigma^2/2)(T-t), \sigma^2(T-t)).$$

## Proof of formula: Second term of decomposition

Now consider separately each term in the decomposition.

For the second term,

$$\begin{aligned}
 \mathbb{P}_t(X_T > K) &= \mathbb{P}_t\left(\log(X_T/F_t) > \log(K/F_t)\right) \\
 &= \mathbb{P}_t\left(\frac{\log(X_T/F_t) + (\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}} \right. \\
 &\quad \left. > \frac{\log(K/F_t) + (\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}\right) \\
 &= N\left(-\frac{\log(K/F_t) + (\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}\right) \\
 &= N(d_2)\Big|_{X=X_t}
 \end{aligned}$$



## Proof of formula: First term of decomposition

In the first term, we have  $\mathbb{E}_t((X_T/F_t)\mathbf{1}_{X_T>K})$

$$\begin{aligned}
 &= \mathbb{E}_t[e^{\log(X_T/F_t)}\mathbf{1}_{\log(X_T/F_t)>\log(K/F_t)}] \\
 &= \int_{\log(K/F_t)}^{\infty} e^y \frac{1}{\sqrt{2\pi\sigma^2(T-t)}} \exp\left(-\frac{(y+(\sigma^2/2)(T-t))^2}{2\sigma^2(T-t)}\right) dy \\
 &= \int_{\log(K/F_t)}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2(T-t)}} \exp\left(-\frac{(y-(\sigma^2/2)(T-t))^2}{2\sigma^2(T-t)}\right) dy \\
 &= \int_{\frac{\log(K/F_t)-(\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{z^2}{2}\right) dz \\
 &= N\left(-\frac{\log(K/F_t)-(\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}\right) = N(d_1)\Big|_{X=X_t},
 \end{aligned}$$

as claimed. The substitution was  $z = \frac{y-(\sigma^2/2)(T-t)}{\sigma\sqrt{T-t}}$  so  $dz = \frac{1}{\sigma\sqrt{T-t}}dy$

# Probabilistic interpretation of $N(d_1)$ and $N(d_2)$

In summary,

$$\begin{aligned} C_t &= e^{-r(T-t)} \left( F_t \mathbb{E}_t \left[ \frac{X_T}{F_t} \mathbf{1}_{S_T > K} \right] - K \mathbb{P}_t(S_T > K) \right) \\ &= e^{-r(T-t)} \left( F_t N(d_1) - K N(d_2) \right) \end{aligned}$$

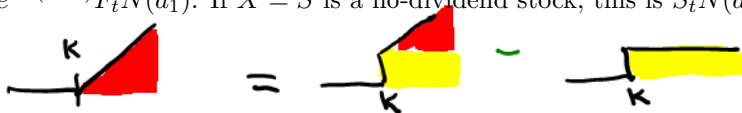
The  $N(d_2)$  is the  $\mathbb{P}_t$  that the call option expires in the money.

(So  $e^{-r(T-t)} N(d_2) =$  time- $t$  price of  $K$ -strike  $T$ -expiry binary call.)

The  $N(d_1)$  is the “share measure”  $\mathbb{P}_t^S$  that the call expires ITM.

And time- $t$  price of asset-or-nothing call paying  $X_T \mathbf{1}_{X_T > K}$  is

$e^{-r(T-t)} F_t N(d_1)$ . If  $X = S$  is a no-dividend stock, this is  $S_t N(d_1)$



## Probabilistic analysis of effect of $r$

Increasing  $r$ , while keeping everything else fixed, has what effect on time-0 call prices?

Martingale methods make the answer clear:

$$\begin{aligned}
 C_0 &= e^{-rT} \mathbb{E} C_T \\
 &= e^{-rT} \mathbb{E} (S_T - K)^+ \\
 &= e^{-rT} \mathbb{E} (S_0 e^{(r-\sigma^2/2)T + \sigma \tilde{W}_T} - K)^+ \\
 &= \mathbb{E} (S_0 e^{(-\sigma^2/2)T + \sigma \tilde{W}_T} - K e^{-rT})^+
 \end{aligned}$$

If  $r \uparrow$  then  $C_0 \uparrow$

# Probabilistic intuition about impact of $\sigma$

The *vega*, at time- $t$ , of an asset or portfolio with value  $C_t = C(S_t, t; \sigma)$  is  $\frac{\partial C}{\partial \sigma}(S_t, t; \sigma)$ . For a call or put in the B-S model,

$$\text{vega} := \frac{\partial C^{BS}}{\partial \sigma} = S\sqrt{T-t}N'(d_1) > 0$$

Why positive? Let's take  $r = 0$ .

- ▶ For a linear payout  $a + bS_T$ , the time-0 value is  $a + bS_0$  regardless of  $\sigma$ . So vega for a linear payout is zero.

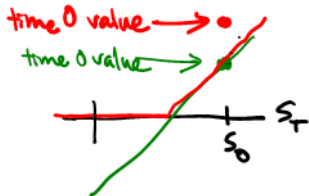
- ▶ For a **convex** payout  $f(S_T)$ , such as a call, the payout

dominates the linear tangent to  $f$  at  $S_0$ . So the

contract's time-0 value is bigger than  $f(S_0)$ . By how much? Depends on  $\sigma$ . The

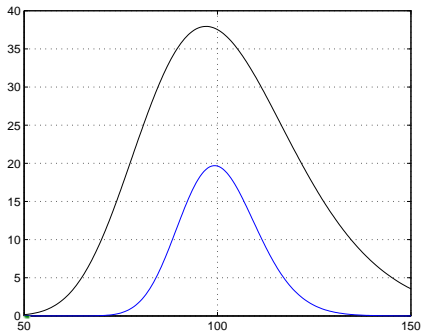
larger the  $\sigma$ , the larger the **convex** contract's time-0 value, because the larger the

chance that  $S$  goes to where  $f > \text{linear}$ . Large  $\sigma \Rightarrow$  more **convexity** is reachable.



# Vega of a call

Vega of a call, plotted against  $S_t$ , for  $T - t = 1$ , and  $T - t = 0.25$ .



$K=100$

$S_t$

Under B-S dynamics, vega of a call is positive.

# Probabilistic analysis of joint effect of $\sigma$ and $T$

Halving  $\sigma$  and quadrupling  $T$ , while keeping everything else fixed, has what effect on time-0 call prices?

$$\log \frac{S_T}{S_0} \sim \text{Normal} \left( (r - \frac{\sigma^2}{2})T, \sigma^2 T \right)$$

$$\sigma \rightarrow \frac{\sigma}{2}$$

$$T \rightarrow 4T$$

$$\begin{aligned} & \left( r - \frac{(\sigma/2)^2}{2} \right) 4T \\ &= \left( 4r - \frac{\sigma^2}{2} \right) T \end{aligned}$$

$$\left( \frac{\sigma}{2} \right)^2 (4T) = \sigma^2 T$$

$$\text{Also } e^{-rT} \rightarrow e^{-4rT}$$

Equivalent to  $r \rightarrow 4r$ .  $C_0 \uparrow$  if  $r > 0$



# Appendix

On an interval  $I$  a function  $f : I \rightarrow \mathbb{R}$  is said to be *convex* if its graph lies on or below all of its chords: for all  $x, y \in I$ , all  $\alpha \in [0, 1]$ ,

$$f(\alpha x + (1 - \alpha)y) \leq \alpha f(x) + (1 - \alpha)f(y)$$

Properties:

- ▶ A twice differentiable function  $f$  is convex iff  $f'' \geq 0$  everywhere.
- ▶ A convex function's graph lies on or above all of its tangents.



## Jensen's inequality:

If  $f$  is convex on  $I$  and  $X$  is an integrable random variable taking values in  $I$  then  $\mathbb{E}f(X) \geq f(\mathbb{E}X)$ .

Johan Jensen (1859-1925)