

Robust Weighting and Matching Techniques for Causal Inference in Observational Studies with Continuous Treatment

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Chapter 1

Introduction

Chapter 2

Chapter One Title

Assumption 1. Assume, the following conditions hold:

1.1. The minimizer $\lambda_0 = \arg \min_{\lambda \in \Theta} \mathbb{E} [-Tn\rho (B(X)^T \lambda) + B(X)^T \lambda]$ is unique, where $\Theta \subseteq \mathbb{R}^n$ is the parameter space for λ .

1.2. The parameter space $\Theta \subseteq \mathbb{R}^n$ is compact.

1.3. $\lambda_0 \in \text{int}(\Theta)$, where $\text{int}(\cdot)$ stands for the interior of a set.

1.4. There exists $\lambda_1^* \in \Theta$ such that $\|m^*(\cdot) - B(\cdot)^T \lambda_1^*\|_\infty \leq \varphi_{m^*}$, where $m^*(\cdot) := (\rho')^{-1} \left(\frac{1}{n\pi(\cdot)} \right)$.

1.5. There exists a constant $\varphi_\pi \in (0, \frac{1}{2})$ such that $\pi(x) \in (\varphi_\pi, 1 - \varphi_\pi)$ for all $x \in \mathcal{X}$

1.6. There exists $\varphi_{\rho''} > 0$ such that $-\rho'' \geq \varphi_{\rho''} > 0$

1.7. There exists $\varphi_{B(x)B(x)^T} > 0$ such that $B(x)B(x)^T \succcurlyeq \varphi_{B(x)B(x)^T} I$

1.8. There exists $\varphi_{\|B\|} > 0$ such that $\sup_{x \in \mathcal{X}} \|B(x)\|_2 \leq \varphi_{\|B\|}$.

We study the following problem:

$$\begin{aligned} & \underset{w \in \mathbb{R}^n}{\text{minimize}} && \sum_{i=1}^n T_i f(w_i) \\ & \text{subject to} && \left| \sum_{i=1}^n w_i T_i B_k(X_i) - \frac{1}{n} \sum_{i=1}^n B_k(X_i) \right| \leq \delta_k, \quad k = 1, \dots, K \end{aligned} \tag{2.1}$$

We aim to prove that the solution to Problem (2.1) is asymptotical consistent with the propensity score, i.e.

Theorem 2.1. *Under some (non-optimal) Assumptions, there exist constants $c_1, c_2 > 0$ and decreasing sequences $(\varepsilon_n^1), (\varepsilon_n^2) \subset (0, 1]$ that converge to 0 such that for all $\tau \in (0, 1]$ there exists a constant $c_\tau \in [0, \infty)$ only depending on τ such that for all $n \geq 1$ and $\tau \in (0, 1]$ it holds*

$$\mathbb{P} \left(\left\| w_i^* - \frac{1}{n\pi(X_i)} \right\|_\infty \leq c_1 c_\tau \varepsilon_n^1 \right) \geq 1 - \tau, \\ \left\| w_i^* - \frac{1}{n\pi(X_i)} \right\|_{\mathbb{P}, 2} \leq c_2 \varepsilon_n^2,$$

where w^* is the solution to Problem (2.1).

Plan of Proof

It is easier to study the dual of Problem (2.1). Thus we employ results from convex analysis [1] to establish

Proposition 2.1. *The dual of Problem (2.1) is equivalent to the unconstrained optimization problem*

$$\underset{\lambda \in \mathbb{R}^K}{\text{minimize}} \quad \frac{1}{n} \sum_{j=1}^n [-T_j n \rho(B(X_j)^T \lambda) + B(X_j)^T \lambda] + |\lambda|^T \delta, \quad (2.2)$$

where $B(X_j) = (B_k(X_j))_{1 \leq k \leq K}$ denotes the K basis functions of the covariates, $\rho(t) := \frac{t}{n} - t(h')^{-1}(t) + h((h')^{-1}(t))$ with $h(x) := f\left(\frac{1}{n} - x\right)$ and $|\lambda| := (|\lambda_k|)_{1 \leq k \leq K}$. Moreover, the primal solution w_j^* satisfies

$$w_j^* = \rho' \left(B(X_j)^T \lambda^\dagger \right) \quad (2.3)$$

for $j = 1, \dots, n$, where λ^\dagger is the solution to the dual optimization problem.

The core of the subsequent analysis is based on Assumption 1.4, i.e. the existence of an oracle parameter λ_1^* in a sieve estimate of the true propensity score (or a transformation). It is then natural to enquire about the convergence of the dual solution λ^\dagger to λ_1^* . Making certain assumptions and employing matrix concentration inequalitys [2] we can establish

Proposition 2.2. *Under some (non-optimal) Assumptions, there exists a constant $c_3 > 0$ and a decreasing sequence $(\varepsilon_n^3) \subset (0, 1]$ that converges to 0 such that for all $\tau \in (0, 1]$ there exists a constant $\tilde{c}_\tau \in [0, \infty)$ only depending on τ such that for all $n \geq 1$ and $\tau \in (0, 1]$ it holds*

$$\mathbb{P} \left(\|\lambda^\dagger - \lambda_1^*\|_2 \leq c^3 \tilde{c}_\tau (\varepsilon_n^3) \right) \geq 1 - \tau. \quad (2.4)$$

It is then straightforward to prove a more general result then Theorem 2.1.

Theorem 2.2. *Under some (non-optimal) Assumptions, there exist constants $c_1, c_2 > 0$ and decreasing sequences $(\varepsilon_n^1), (\varepsilon_n^2) \subset (0, 1]$ that converge to 0 such that for all $\tau \in (0, 1]$ there exists a constant $c_\tau \in [0, \infty)$ only depending on τ such that for all $n \geq 1$ and $\tau \in (0, 1]$ it holds*

$$\begin{aligned} \mathbb{P} \left(\left\| w^*(\cdot) - \frac{1}{n\pi(\cdot)} \right\|_\infty \leq c_1 c_\tau \varepsilon_n^1 \right) &\geq 1 - \tau, \\ \left\| w^*(X) - \frac{1}{n\pi(X)} \right\|_{\mathbb{P}, 2} &\leq c_2 \varepsilon_n^2, \end{aligned}$$

where $w^*(X)$ is as in (2.3) without the index.

Proof of theorem 2.2

Proof. Motivated by Proposition 5.1 we set $\|\Delta\|_2 = C$ and consider

$$G(\lambda) := \frac{1}{n} \sum_{j=1}^n [-T_j n \rho(B(X_j)^T \lambda) + B(X_j)^T \lambda] + |\lambda|^T \delta. \quad (2.5)$$

Since $\rho \in C^2(\mathbb{R})$ we can employ Proposition 5.1, Corollary 5.1.1 and Proposition 5.2 to get

$$\begin{aligned}
& G(\lambda_1^* + \Delta) - G(\lambda_1^*) \\
& \geq \frac{1}{n} \sum_{j=1}^n \left[-T_j n \rho' (B(X_j)^T \lambda_1^*) + 1 \right] \Delta^T B(X_j) \\
& + \frac{1}{2} \sum_{j=1}^n -T_j \rho'' (B(X_j)^T (\lambda_1^* + \xi \Delta)) \Delta^T (B(X_j) B(X_j)^T) \Delta \\
& - |\Delta|^T \delta \\
& \geq -\|\Delta\|_2 \left(\left\| \frac{1}{n} \sum_{j=1}^n \left[-T_j n \rho' (B(X_j)^T \lambda_1^*) + 1 \right] B(X_j) \right\|_2 + \|\delta\|_2 \right) \\
& + n \|\Delta\|_2^2 \varphi_{\rho''} \varphi_{aa^T} \\
& := -\|\Delta\|_2 (I_1 + \|\delta\|_2) + \|\Delta\|_2^2 I_2.
\end{aligned} \tag{2.6}$$

The second inequality is due to the Cauchy-Schwarz-Inequality and Assumptions 1.6 and 1.7 .

Analysis of I_1

We want to use Assumption 1.3. Thus we perform the following split:

$$I_1 \leq \left\| \sum_{j=1}^n T_j \left[\rho' (B(X_j)^T \lambda_1^*) - \frac{1}{n\pi(X_j)} \right] B(X_j) \right\|_2 \tag{2.7}$$

$$+ \left\| \frac{1}{n} \sum_{j=1}^n \left[\frac{T_j}{\pi(X_j)} - 1 \right] B(X_j) \right\|_2 \tag{2.8}$$

$$=: J_1 + J_2 \tag{2.9}$$

Analysis of J_1

By the Lipschitz-continuity of ρ' , Assumption 1.8 and Assumption 1.4, $T \in \{0, 1\}$ and the triangle inequality we have

$$J_1 \leq n L_{\rho'} \varphi_{\|B(x)\|} \varphi_{m^*} \tag{2.10}$$

Analysis of J_2

We employ Bernstein Inequality for matrices To this end we define

$$A_j := \frac{1}{n} \left[\frac{T_j}{\pi(X_j)} - 1 \right] B(X_j) \quad (2.11)$$

$$\mathbb{E}A_j = 0$$

It holds

$$\mathbb{E} \left[\frac{T_j}{\pi(X_j)} B(X_j) \right] = \mathbb{E} \left[\mathbb{E}[T_j | X_j] \frac{1}{\pi(X_j)} B(X_j) \right] = \mathbb{E}[B(X_j)]. \quad (2.12)$$

Thus $\mathbb{E}[A_j] = 0$.

L

Since

$$\left| \frac{T_j}{\pi(X_j)} - 1 \right| \leq 1 + \frac{1 - \varphi_\pi}{\varphi_\pi} = \frac{1}{\varphi_\pi} \quad (2.13)$$

by Assumption 1.5, we can employ Assumption 1.8 to get

$$\|A_j\|_2 \leq \frac{\varphi_{\|B\|}}{n\varphi_\pi} =: L. \quad (2.14)$$

v(S)

Since

$$\mathbb{E} [A_j A_j^T] \leq \left(\frac{1}{n\varphi_\pi} \right)^2 \mathbb{E} [B(X) B(X)^T] \quad (2.15)$$

and

$$\mathbb{E} [A_j^T A_j] \leq \left(\frac{\varphi_{\|B\|}}{n\varphi_\pi} \right)^2 \quad (2.16)$$

we have

$$v(S) \leq \frac{|\lambda_{\max}| + \varphi_{\|B\|}^2}{n\varphi_\pi^2}, \quad (2.17)$$

where λ_{\max} is the maximal eigenvalue of $\mathbb{E}[B(X)B(X)^T]$. Then by Bernsteins inequality 4.1 we get

$$\mathbb{E}[J_2] \leq \sqrt{\frac{2 \log(K+1) (|\lambda_{\max}| + \varphi_{\|B\|}^2)}{n\varphi_{\pi}^2}} + \frac{\log(K+1)\varphi_{\|B\|}}{3n\varphi_{\pi}} \quad (2.18)$$

and by the Markov-inequality

$$\mathbb{P}\left(J_2 \leq \frac{1}{\tau} \mathbb{E}[J_2]\right) \geq 1 - \tau \quad (2.19)$$

Finish

If we choose for $\gamma > 0$

$$\|\Delta\|_2 = \frac{\frac{1}{\tau} \mathbb{E}[J_2] + nL_{\rho'} \varphi_{\|B(x)\|} \varphi_{m^*} + \|\delta\|_2}{\varphi_{\rho''} \varphi_{BB^T}} (1 + \gamma) \quad (2.20)$$

$$=: C \quad (2.21)$$

we have

$$\mathbb{P}(\|\lambda^{\dagger} - \lambda_1^*\|_2 \leq C) = \mathbb{P}\left(\inf_{\|\Delta\|_2 = C} G(\lambda_1^* + \Delta) - G(\lambda_1^*) > 0\right) \quad (2.22)$$

$$\geq 1 - \tau \quad (2.23)$$

Finish 2

$$\left\|w^*(X) - \frac{1}{n\pi(X)}\right\|_{\mathbb{P},2} \leq L_{\rho'} \left[\|B(X)^T (\lambda^{\dagger} - \lambda_1^*)\|_{\mathbb{P},2}\right. \quad (2.24)$$

$$\left. + \|m^*(X) - B(X)^T \lambda_1^*\|_{\mathbb{P},2}\right] \quad (2.25)$$

$$\leq L_{\rho'} \left(\varphi_{\|B\|} \sqrt{C^2(1 - \tau) + \text{diam}(\Theta)^2 \tau} + \varphi_{m^*}\right) \quad (2.26)$$

$$\left\|w^*(\cdot) - \frac{1}{n\pi(\cdot)}\right\|_{\infty} \leq L_{\rho'} \left[\|B(\cdot)^T (\lambda^{\dagger} - \lambda_1^*)\|_{\infty}\right. \quad (2.27)$$

$$\left. + \|m^*(\cdot) - B(\cdot)^T \lambda_1^*\|_{\infty}\right] \quad (2.28)$$

$$\leq L_{\rho'} (\varphi_{\|B\|} C + \varphi_{m^*}) \quad (2.29)$$

with probability greater than $1 - \tau$. \square

Chapter 3

Convex Analysis

We begin by defining convex sets

Definition 3.1. A subset $\Omega \subseteq \mathbb{R}^n$ is called CONVEX if we have $\lambda x + (1 - \lambda)y \in \Omega$ for all $x, y \in \Omega$ and $\lambda \in (0, 1)$.

Clearly, the line segment $[a, b] := \{\lambda a + (1 - \lambda)b \mid \lambda \in [0, 1]\}$ is contained in Ω for all $a, b \in \Omega$ if and only if Ω is a convex set.

Next we define convex functions.

The concept of convex functions is closely related to convex sets.

The line segment between two points on the graph of a convex function lies on or above and does not intersect the graph.

In other words: The area above the graph of a convex function f is a convex set, i.e. the *epigraph* $\text{epi}(f) := \{(x, \alpha) \in \mathbb{R}^n \times \mathbb{R} \mid f(x) \leq \alpha\}$ is a convex set in \mathbb{R}^{n+1} .

Often an equivalent characterisation of convex functions is more useful.

Theorem 3.1. The convexity of a function $f : \mathbb{R}^n \rightarrow \overline{\mathbb{R}}$ on \mathbb{R}^n is equivalent to the following statement:

For all $x, y \in \mathbb{R}^n$ and $\lambda \in (0, 1)$ we have

$$f(\lambda x + (1 - \lambda)y) \leq \lambda f(x) + (1 - \lambda)f(y). \quad (3.1)$$

Definition 3.2. proper convex function

Definition 3.3. convex conjugate

Given proper convex functions $f, g : \mathbb{R}^n \rightarrow \overline{\mathbb{R}}$ and a matrix $A \in \mathbb{R}^{n \times n}$, we define the primal minimization problem as follows:

$$\text{minimize } f(x) + g(Ax) \quad \text{subject to } x \in \mathbb{R}^n. \quad (3.2)$$

The Fenchel dual problem is then

$$\text{maximize } -f^*(A^T y) - g^*(-y) \quad \text{subject to } y \in \mathbb{R}^n. \quad (3.3)$$

Theorem 3.2. *Let $f, g : \mathbb{R}^n \rightarrow \overline{\mathbb{R}}$ be proper convex functions and $0 \in \text{ri}(\text{dom}(g) - A(\text{dom}(f)))$. Then the optimal values of (3.2) and (3.3) are equal, i.e.*

$$\inf_{x \in \mathbb{R}^n} \{f(x) + g(Ax)\} = \sup_{y \in \mathbb{R}^n} \{-f^*(A^T y) - g^*(-y)\}. \quad (3.4)$$

Chapter 4

Random Matrix Inequality

Theorem 4.1. *Let $(A_k)_{1 \leq k \leq n} \subseteq \mathbb{R}^{d_1 \times d_2}$ be a finite sequence of independent, random matrices. Assume that*

$$\mathbb{E}(A_k) = 0 \quad \text{and} \quad \|A_k\| \leq L \quad \text{for each } k \in \{1, \dots, n\}. \quad (4.1)$$

Introduce the random matrix

$$S := \sum_{k=1}^n A_k. \quad (4.2)$$

Let $v(S)$ be the matrix variance statistic of the sum:

$$v(S) := \max \left\{ \|\mathbb{E}(SS^T)\|, \|\mathbb{E}(S^T S)\| \right\} \quad (4.3)$$

$$= \max \left\{ \left\| \sum_{k=1}^n \mathbb{E}(A_k A_k^T) \right\|, \left\| \sum_{k=1}^n \mathbb{E}(A_k^T A_k) \right\| \right\}. \quad (4.4)$$

Then

$$\mathbb{E} \|S\| \leq \sqrt{2v(S) \log(d_1 + d_2)} + \frac{1}{3} L \log(d_1 + d_2). \quad (4.5)$$

Furthermore, for all $t \geq 0$,

$$\mathbb{P}(\|S\| \geq t) \leq (d_1 + d_2) \exp \left(\frac{-t^2/2}{v(S) + Lt/3} \right). \quad (4.6)$$

Chapter 5

Simple yet useful Calculations

Proposition 5.1. *Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ be continuous such that a minimum x^* exists and is unique. Then for all $y \in \mathbb{R}^n$ and $C > 0$ it follows*

$$\inf_{\|\Delta\|=C} f(y + \Delta) - f(y) > 0 \quad \Rightarrow \quad \|x^* - y\| \leq C. \quad (5.1)$$

Proof. Since $\mathcal{C} := \{\|\Delta\| \leq C\}$ is compact and

$$f(x^*) \leq f(y) < \inf_{\|\Delta\|=C} f(y + \Delta),$$

the continuous function $f(y + \cdot)$ has a minimum in $\text{int}(\mathcal{C}) := \{\|\Delta\| < C\}$. Since x^* is the unique minimum of f there exists $\Delta^* \in \text{int}(\mathcal{C})$ such that $x^* - y = \Delta^*$. We conclude that $\|x^* - y\| \leq C$. \square

Theorem 5.1. (Multivariate Taylor Theorem) *Let $f \in C^2(\mathbb{R}^n, \mathbb{R})$. Then for all $x, \Delta \in \mathbb{R}^n$ there exists $\xi \in [0, 1]$ such that it holds*

$$\begin{aligned} f(x + \Delta) = f(x) &+ \sum_{i=1}^n \frac{\partial f(x)}{\partial x_i} \Delta_i + \sum_{\substack{i,j=1 \\ i \neq j}}^n \frac{\partial^2 f(x + \xi \Delta)}{\partial x_i \partial x_j} \Delta_i \Delta_j \\ &+ \frac{1}{2} \sum_{i=1}^n \frac{\partial^2 f(x + \xi \Delta)}{\partial x_i^2} \Delta_i^2 \end{aligned} \quad (5.2)$$

Corollary 5.1.1. *Let $f \in C^2(\mathbb{R})$. Then for all $a, x, \Delta \in \mathbb{R}^n$ there exist $\xi \in [0, 1]$ such that it holds*

$$f(a^T(x + \Delta)) - f(a^T x) = f'(a^T x) \Delta^T a + \frac{1}{2} f''(a^T(x + \xi \Delta)) \Delta^T A \Delta, \quad (5.3)$$

where $A := aa^T \in \mathbb{R}^{n \times n}$.

Proof. By the chain rule we have for all $a, x, \Delta \in \mathbb{R}^n$ and $\xi \in [0, 1]$

$$\frac{\partial^2 f(a^T(x + \xi\Delta))}{\partial x_i \partial x_j} = f''(a^T(x + \xi\Delta)) a_i a_j. \quad (5.4)$$

Since $A := aa^T$ is symmetric we have

$$\Delta^T A \Delta = 2 \sum_{\substack{i,j=1 \\ i \neq j}}^n a_i a_j \Delta_i \Delta_j + \sum_{i=1}^n a_i^2 \Delta_i^2. \quad (5.5)$$

Plugging (5.4) and (5.5) into (5.2) yields (5.3). \square

Proposition 5.2. *For all $x, y \in \mathbb{R}$ it holds*

$$|x + y| - |x| \geq -|y| \quad (5.6)$$

Proof. Checking all 6 combinations of $x+y, x, y$ being nonnegative or negative yields the result. \square

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