Logistic regression

Inference

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Risk of coronary heart disease

This dataset is from an ongoing cardiovascular study on residents of the town of Framingham, Massachusetts. We want to examine the relationship between various health characteristics and the risk of having heart disease in the next 10 years.

high_risk: 1 = High risk, 0 = Not high risk

age: Age at exam time (in years)

education: 1 = Some High School; 2 = High School or GED; 3 = Some College or Vocational School; 4 = College

currentSmoker: 0 = nonsmoker; 1 = smoker



Modeling risk of coronary heart disease

term	estimate	std.error	statistic	p.value	conf.low	conf.high
(Intercept)	-5.385	0.308	-17.507	0.000	-5.995	-4.788
age	0.073	0.005	13.385	0.000	0.063	0.084
education2	-0.242	0.112	-2.162	0.031	-0.463	-0.024
education3	-0.235	0.134	-1.761	0.078	-0.501	0.023
education4	-0.020	0.148	-0.136	0.892	-0.317	0.266

$$\log\left(\frac{\hat{\pi}}{1-\hat{\pi}}\right) = -5.385 + 0.073 \text{ age} - 0.242 \text{ ed2} - 0.235 \text{ ed3} - 0.020 \text{ ed4}$$



Hypothesis test for β_j

Hypotheses: $H_0: \beta_j = 0$ vs $H_a: \beta_j \neq 0$

Test Statistic:

$$z = \frac{\hat{\beta}_j - 0}{SE(\hat{\beta}_j)}$$

P-value: P(|Z| > |z|),

where $Z \sim N(0, 1)$, the Standard Normal distribution



Confidence interval for β_j

We can calculate the **C% confidence interval** for β_j as the following:

$$\hat{\beta}_j \pm z^* SE(\hat{\beta}_j)$$

where z^* is calculated from the N(0, 1) distribution



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This is an interval for the change in the log-odds for every one unit increase in x_i .



Interpretation in terms of the odds

The change in **odds** for every one unit increase in x_j .

$$\exp\{\hat{\beta}_j \pm z^* SE(\hat{\beta}_j)\}\$$



Interpretation in terms of the odds

The change in **odds** for every one unit increase in x_j .

$$\exp\{\hat{\beta}_j \pm z^* SE(\hat{\beta}_j)\}\$$

Interpretation: We are C% confident that for every one unit change in x_j , the odds multiply by a factor of $\exp\{\hat{\beta}_j - z^*SE(\hat{\beta}_j)\}$ to $\exp\{\hat{\beta}_j + z^*SE(\hat{\beta}_j)\}$, holding all else constant.



Model

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Hypotheses

$$H_0: \beta_1 = 0 \text{ vs } H_a: \beta_1 \neq 0$$



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Test statistic

$$z = \frac{0.0733 - 0}{0.00547} = 13.4$$



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P-value

$$P(|Z| > |13.4|) \approx 0$$



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```
2 * pnorm(13.4, lower.tail = FALSE)
```

[1] 6.046315e-41



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Conclusion: The p-value is very small, so we reject H_0 . The data provide sufficient evidence that age is a statistically significant predictor of whether someone is high risk of having heart disease, *after accounting* for education.



Comparing models



Log likelihood

$$\log L = \sum_{i=1}^{n} [y_i \log(\hat{\pi}_i) + (1 - y_i) \log(1 - \hat{\pi}_i)]$$

- Measure of how well the model fits the data
- lacktriangle Higher values of $\log L$ are better
- Deviance = $-2 \log L$
 - $-2 \log L$ follows a χ^2 distribution with n-p-1 degrees of freedom



Comparing nested models

- Suppose there are two models:
 - Reduced Model includes predictors x_1, \ldots, x_q
 - Full Model includes predictors $x_1, \ldots, x_q, x_{q+1}, \ldots, x_p$
- We want to test the hypotheses

$$H_0: \beta_{q+1} = \dots = \beta_p = 0$$

$$H_a$$
: at least 1 β_i is not 0

 To do so, we will use the Drop-in-deviance test (also known as the Nested Likelihood Ratio test)



Drop-in-deviance test

Hypotheses:

$$H_0: \beta_{q+1} = \dots = \beta_p = 0$$

 H_a : at least 1 β_j is not 0

Test Statistic:

$$G = (-2 \log L_{reduced}) - (-2 \log L_{full})$$

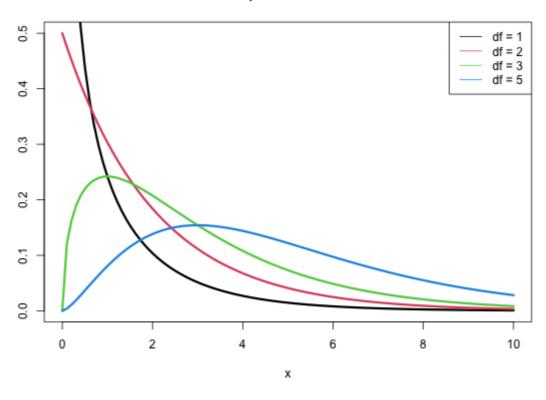
P-value: $P(\chi^2 > G)$,

calculated using a χ^2 distribution with degrees of freedom equal to the difference in the number of parameters in the full and reduced models



χ^2 distribution

Chi-square Distribution





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```
# Calculate deviance for each model
(dev_reduced <- glance(model_reduced)$deviance)

## [1] 3300.135

(dev_full <- glance(model_full)$deviance)

## [1] 3279.359</pre>
```



```
# Calculate deviance for each model
 (dev_reduced <- glance(model_reduced)$deviance)</pre>
## [1] 3300.135
 (dev_full <- glance(model_full)$deviance)</pre>
## [1] 3279.359
 # Drop-in-deviance test statistic
 (test_stat <- dev_reduced - dev_full)</pre>
```



```
# p-value
#1 = number of new model terms in model 2
pchisq(test_stat, 1, lower.tail = FALSE)
```

```
## [1] 5.162887e-06
```



```
# p-value
#1 = number of new model terms in model 2
pchisq(test_stat, 1, lower.tail = FALSE)
```

```
## [1] 5.162887e-06
```

Conclusion: The p-value is very small, so we reject H_0 . The data provide sufficient evidence that the coefficient of **currentSmoker** is not equal to 0. Therefore, we should add it to the model.



Drop-in-Deviance test in R

We can use the **anova** function to conduct this test

• Add test = "Chisq" to conduct the drop-in-deviance test

```
anova(model_reduced, model_full, test = "Chisq") %>%
  tidy()
```



Model selection

Use AIC or BIC for model selection

$$AIC = -2 * \log L - n \log(n) + 2(p+1)$$

$$BIC = -2 * \log L - n \log(n) + \log(n) \times (p+1)$$



AIC from glance function

Let's look at the AIC for the model that includes **age**, **education**, and **currentSmoker**

```
glance(model_reduced)$AIC
```

```
## [1] 3310.135
```



AIC from glance function

Let's look at the AIC for the model that includes **age**, **education**, and **currentSmoker**

```
glance(model_reduced)$AIC
```

```
## [1] 3310.135
```

Calculating AIC

```
- 2 * glance(model_reduced)$logLik + 2 * (4 + 1)
```

```
## [1] 3310.135
```



Comparing the models using AIC

Let's compare the full and reduced models using AIC.

```
glance(model_reduced)$AIC
```

```
## [1] 3310.135
```

```
glance(model_full)$AIC
```

```
## [1] 3291.359
```

Based on AIC, which model would you choose?



Comparing the models using BIC

Let's compare the full and reduced models using BIC

```
glance(model_reduced)$BIC
```

```
## [1] 3341.772
```

```
glance(model_full)$BIC
```

```
## [1] 3329.323
```

Based on BIC, which model would you choose?

