


Wald Test: use $N(0,1)$

$$\frac{\hat{\beta} - b}{SE\hat{\beta}} \approx N(0,1) \quad \text{(under } H_0)$$

$$H_0: \beta = b$$

$$H_a: \beta \neq b$$

$$\beta > b$$

etc.

(in R:)

	Estimate	Std Error
(Intercept)	$\hat{\beta}$	$SE\hat{\beta}$
variable	$\hat{\beta}$	$SE\hat{\beta}$

on question

1) write H_0 & H_A

2) calculate $z = \frac{\hat{\beta} - b}{SE\hat{\beta}}$

3) p-value

Likelihood ratio test (LRT)

(aka Drop in deviance test)

$$H_0: \beta_1 = \beta_2 = \beta_3 = 0$$

H_A : at least one of $\beta_1, \beta_2, \beta_3 \neq 0$

$$G = \begin{array}{l} \text{deviance for reduced model} \\ - \text{deviance for full model} \end{array}$$

Under H_0 :

$$G \sim \chi^2_{\# \text{parameters tested}}$$

R output: compare residual deviances
for full & reduced models

To know about nested F test
for LMMS:

- 1) Test fixed effects (β_s)
for full & reduced models
- 2) upper & lower bounds on χ^2 df
denominator

Poisson regression:

$$Y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \beta_0 + \beta_1 x_i$$

$$\text{Mean}(Y_i) = \text{Variance}(Y_i) = \lambda_i$$

$$\Rightarrow \phi = \frac{\text{Variance}(Y_i)}{\text{Mean}(Y_i)} = 1$$

\Rightarrow we assume $\phi = 1$

But what if $\phi > 1$?

Quasi-Poisson:

$$\text{Mean}(Y_i) = \lambda_i$$

$$\text{variance}(Y_i) = \phi \lambda_i$$

$$\text{St. Dev.}(Y_i) = \sqrt{\phi \lambda_i}$$

(linear function of λ_i)

we can estimate ϕ , & use estimate to adjust SEs:

$$SE_{\text{quasiPoisson}} = \underbrace{\sqrt{\hat{\phi}}}_{>1} \cdot SE_{\text{Poisson}}$$

\Rightarrow CIs wider,
test statistics smaller,

p-values are larger

Pro: easy to interpret, same $\hat{\beta}$ s as Poisson

Con: uses quasi-likelihood

Negative Binomial

$$Y_i \sim NB(\theta, p_i)$$

$$\text{Mean of } Y_i = \frac{p_i \theta}{1 - p_i} = \mu_i$$

$$\text{Variance of } Y_i = \mu_i + \frac{\mu_i^2}{\theta}$$

(quadratic function
of mean)

$$\log(\mu_i) = \beta_0 + \beta_1 X_i$$

$\hat{\mu}$ \nearrow
interpreted the same as
in (quasi-) Poisson regression

Lab 9: $Y_i \sim \text{NB}(\theta, p)$

$$P(Y_i = y) = \frac{(y + \theta - 1)!}{y! (\theta - 1)!} (1-p)^\theta p^y$$

Observe Y_1, \dots, Y_n . Want MLE \hat{p}

$$L(\hat{p}) = \prod_{i=1}^n \frac{(Y_i + \theta - 1)!}{Y_i! (\theta - 1)!} (1-\hat{p})^\theta \hat{p}^{Y_i}$$

$$\log L(\hat{p}) = \sum_{i=1}^n \left(\log \left(\frac{(Y_i + \theta - 1)!}{Y_i! (\theta - 1)!} \right) + \theta \log(1-\hat{p}) + Y_i \log(\hat{p}) \right)$$

$$\frac{d}{d\hat{p}} \log L(\hat{p}) = \sum_{i=1}^n \left(0 - \frac{\theta}{1-\hat{p}} + \frac{Y_i}{\hat{p}} \right) \stackrel{\text{set}}{=} 0$$

$$\sum_{i=1}^n \frac{Y_i}{\hat{p}} = \sum_{i=1}^n \frac{\theta}{1-\hat{p}} = \frac{n\theta}{1-\hat{p}}$$

$$\sum_{i=1}^n Y_i = \frac{n\theta \hat{p}}{1-\hat{p}} \Rightarrow \frac{\hat{p}}{1-\hat{p}} = \frac{\sum_{i=1}^n Y_i}{n\theta}$$

$$\frac{\hat{p}}{1-\hat{p}} = \frac{n\theta}{\sum_{i=1}^n y_i}$$

$$\Rightarrow \frac{1-\hat{p}}{\hat{p}} = \frac{\sum_{i=1}^n y_i}{n\theta}$$

$$\Rightarrow \frac{1}{\hat{p}} = \frac{\sum_{i=1}^n y_i}{n\theta} + 1 = \frac{\sum_{i=1}^n y_i + n\theta}{n\theta}$$

$$\Rightarrow \hat{p} = \frac{n\theta}{n\theta + \sum_{i=1}^n y_i}$$