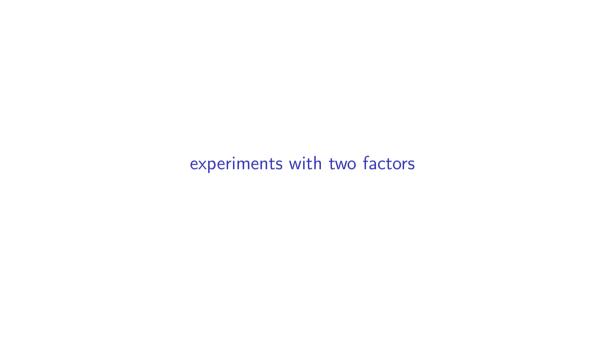
STA221

Neil Montgomery

Last edited: 2017-07-27 17:57



fire retardant example

There aren't many datasets that come with Chapter 26, so I found another to use as a motivating example.

In this dataset a new fire retardant treatment of cotton fabric is being tested. The name of this factor is Bath. There are two levels names I and II, which stand for "received the treatment" and "did not receive the new treatment" respectively.

(It's called "Bath" because the fabric is bathed in the treatment solution.)

fire retardant example

There aren't many datasets that come with Chapter 26, so I found another to use as a motivating example.

In this dataset a new fire retardant treatment of cotton fabric is being tested. The name of this factor is Bath. There are two levels names I and II, which stand for "received the treatment" and "did not receive the new treatment" respectively.

(It's called "Bath" because the fabric is bathed in the treatment solution.)

The amount of time it takes each cloth sample to start to burn is recorded.

Here is a numerical summary of the results:

Bath	n	Means	SD
	24	16.875	5.921
Ш	24	9.154	5.670

but there is also another variable

The new treatment works. But there is also the matter of the efficiency with which the treatment can be applied.

There is another factor variable in this dataset: the number of "launderings", which is the way a retardant treatment is applied.

This variable has two levels named 5 and 10, corresponding to the actual number of launderings.

Here is a summary of the results with respect to this variable:

Launderings	n	Means	SD
10	24	15.067	5.609
5	24	10.963	7.620

The new treatment is better than the old treatment (all else being equal). More laundering is better than less laundering (all else being equal).

But what about the combination of both variables? There are a few different ways they might combine to affect the results. We'll be concerned with these possibilities:

The new treatment is better than the old treatment (all else being equal). More laundering is better than less laundering (all else being equal).

But what about the combination of both variables? There are a few different ways they might combine to affect the results. We'll be concerned with these possibilities:

Their effects could simply be additive.

The new treatment is better than the old treatment (all else being equal). More laundering is better than less laundering (all else being equal).

But what about the combination of both variables? There are a few different ways they might combine to affect the results. We'll be concerned with these possibilities:

Their effects could simply be additive.

The effect of one factor could depend on the level of the other factor. In this case we say there is an *interaction* between the two factor variables.

The new treatment is better than the old treatment (all else being equal). More laundering is better than less laundering (all else being equal).

But what about the combination of both variables? There are a few different ways they might combine to affect the results. We'll be concerned with these possibilities:

Their effects could simply be additive.

The effect of one factor could depend on the level of the other factor. In this case we say there is an *interaction* between the two factor variables.

A good graphical method to evaluate the relationship is called an *interaction plot*.

The new treatment is better than the old treatment (all else being equal). More laundering is better than less laundering (all else being equal).

But what about the combination of both variables? There are a few different ways they might combine to affect the results. We'll be concerned with these possibilities:

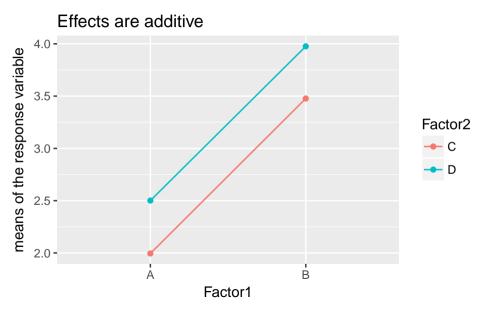
Their effects could simply be additive.

The effect of one factor could depend on the level of the other factor. In this case we say there is an *interaction* between the two factor variables.

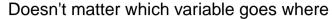
A good graphical method to evaluate the relationship is called an *interaction plot*.

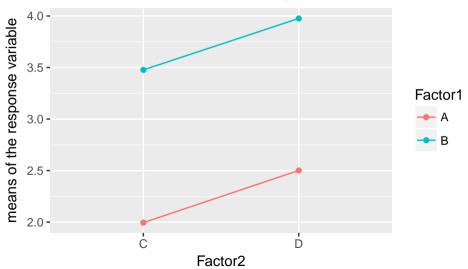
To illustrate, I have simulated a dataset with two variables. Factor1 has levels A and B while Factor2 has levels C and D.

interaction plot example 1a - additive

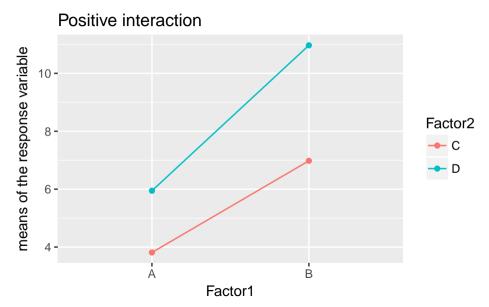


interaction plot example 1b - additive

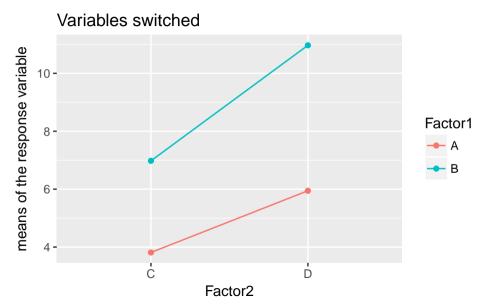




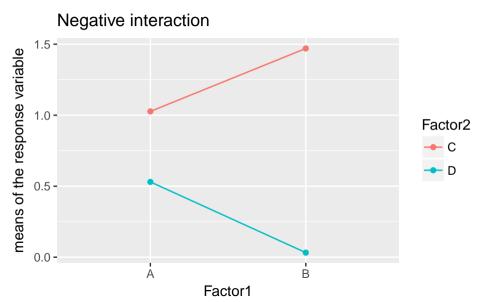
interaction plot example 2a - positive interaction



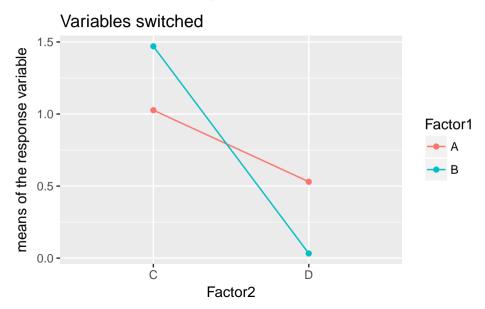
interaction plot example 2b - positive interaction



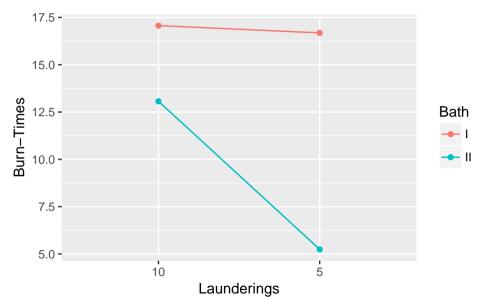
interaction plot example 3a - negative interaction



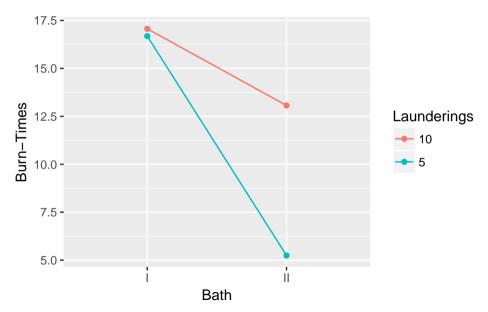
interaction plot example 3a - negative interaction



interaction plot of the fire retardant data



fire data - variables switched



It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

 $\boldsymbol{\mu}$ is the grand, overall average.

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

 $\boldsymbol{\mu}$ is the grand, overall average.

The τ_i are the effects of the levels of the first factor.

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

 μ is the grand, overall average.

The τ_i are the effects of the levels of the first factor.

The γ_i are the effects of the levels of the second factor.

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

 μ is the grand, overall average.

The τ_i are the effects of the levels of the first factor.

The γ_i are the effects of the levels of the second factor.

The $(\tau \gamma)_{ij}$ are the effects of all combinations of the levels of hte two factors.

It seems that the new treatment is just as good, even after only 5 launderings, so it is also more efficient.

Here are the models we will consider:

$$y_{ijk} = \mu + \tau_i + \gamma_j + \varepsilon_{ijk}$$
 no interaction assumed

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

 μ is the grand, overall average.

The τ_i are the effects of the levels of the first factor.

The γ_i are the effects of the levels of the second factor.

The $(\tau \gamma)_{ii}$ are the effects of all combinations of the levels of hte two factors.

$$\varepsilon_{iik}$$
 is random noise, assumed to be $N(0, \sigma)$.

the analysis

We have new sums of squares decompositions. The details can get brutal - so we'll stay symbolic.

We'll call the factor variables A and B, with I and J levels respectively.

In the additive case we get, assuming a common sample size of n for each combination of factor levels (strongly encouraged in practice!):

$$SS_{Total} = SS_A + SS_B + SS_{Error}$$

the analysis

We have new sums of squares decompositions. The details can get brutal - so we'll stay symbolic.

We'll call the factor variables A and B, with I and J levels respectively.

In the additive case we get, assuming a common sample size of n for each combination of factor levels (strongly encouraged in practice!):

$$SS_{Total} = SS_A + SS_B + SS_{Error}$$

Everything has χ^2 distributions. The degrees of freedom add up (N is the grand sample size):

$$N-1 = (I-1) + (J-1) + (N-I-J+1)$$

fire example - no interaction (?!)

```
## Bath 1 715.3 715.3 23.954 0.0000131 ## Launderings 1 202.1 202.1 6.769 0.0125 ## Residuals 45 1343.8 29.9
```

sums of squares - a few details

The total sum of squares SS_{Total} is (as always) (N-1) times the sample varianace of the response variable:

$$\sum_{i,j,k} \left(y_{ijk} - \overline{\overline{y}} \right)^2$$

The treatment sums of squares will be:

$$SS_A = nI \sum_i (\overline{y}_{i..} - \overline{\overline{y}})^2$$

$$SS_B = nJ\sum_i \left(\overline{y}_{\cdot j \cdot} - \overline{\overline{y}}\right)^2$$

where the dots in the subscript mean "averaged over this index."

sums of squares - a few details

The total sum of squares SS_{Total} is (as always) (N-1) times the sample varianace of the response variable:

$$\sum_{i,j,k} \left(y_{ijk} - \overline{\overline{y}} \right)^2$$

The treatment sums of squares will be:

$$SS_A = nI \sum_i \left(\overline{y}_{i \cdot \cdot \cdot} - \overline{\overline{y}} \right)^2$$

$$SS_B = nJ \sum_{i} \left(\overline{y}_{.j.} - \overline{\overline{y}} \right)^2$$

where the dots in the subscript mean "averaged over this index."

The key thing to note is that each treatment sum of squares is computed using the average over all levels of the other treatment...

sums of squares - a few details

The total sum of squares SS_{Total} is (as always) (N-1) times the sample varianace of the response variable:

$$\sum_{i,j,k} (y_{ijk} - \overline{\overline{y}})^2$$

The treatment sums of squares will be:

$$SS_A = nI \sum_i \left(\overline{y}_{i..} - \overline{\overline{y}} \right)^2$$

$$SS_B = nJ\sum_i \left(\overline{y}_{\cdot j \cdot} - \overline{\overline{y}}\right)^2$$

where the dots in the subscript mean "averaged over this index."

The key thing to note is that each treatment sum of squares is computed using the average over all levels of the other treatment...

... which only makes sense when there is no interaction.

error sum of squares

For the sake of completeness:

$$SS_{Error} = \sum_{i,i,k} \left(y_{ijk} - \overline{y}_{ij\cdot} \right)^2$$

Note that $y_{ijk} - \overline{y}_{ij}$ is also called a "residual".

Mostly the same as with one treatment factor, with the same verification techniques.

1. Observations are independent (this is assumed - only knowledge of the experiment itself is of any help in satisfying this assumption.)

Mostly the same as with one treatment factor, with the same verification techniques.

- 1. Observations are independent (this is assumed only knowledge of the experiment itself is of any help in satisfying this assumption.)
- 2. Equal variance for all combinations of levels of treatment factors. (Plots, or Levene's test when *n* is not too small.) (Fatal if violated.)

Mostly the same as with one treatment factor, with the same verification techniques.

- 1. Observations are independent (this is assumed only knowledge of the experiment itself is of any help in satisfying this assumption.)
- 2. Equal variance for all combinations of levels of treatment factors. (Plots, or Levene's test when *n* is not too small.) (Fatal if violated.)
- 3. Normal error (normal quantile plot of residuals) (Violation OK if sample size is large enough.)

Mostly the same as with one treatment factor, with the same verification techniques.

- 1. Observations are independent (this is assumed only knowledge of the experiment itself is of any help in satisfying this assumption.)
- 2. Equal variance for all combinations of levels of treatment factors. (Plots, or Levene's test when *n* is not too small.) (Fatal if violated.)
- 3. Normal error (normal quantile plot of residuals) (Violation OK if sample size is large enough.)

model assumptions

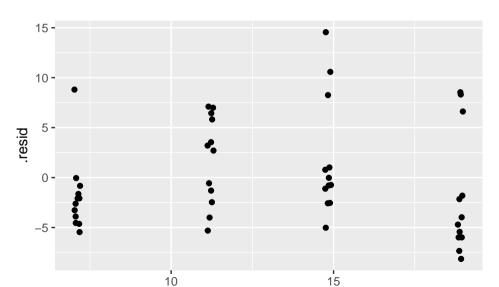
Mostly the same as with one treatment factor, with the same verification techniques.

- 1. Observations are independent (this is assumed only knowledge of the experiment itself is of any help in satisfying this assumption.)
- 2. Equal variance for all combinations of levels of treatment factors. (Plots, or Levene's test when n is not too small.) (Fatal if violated.)
- 3. Normal error (normal quantile plot of residuals) (Violation OK if sample size is large enough.)

When n = 1, the lack of interaction is also an assumption.

fire retardant model assumptions - equal variance

Plot of residuals versus "fitted values" (in this case, just the group averages):

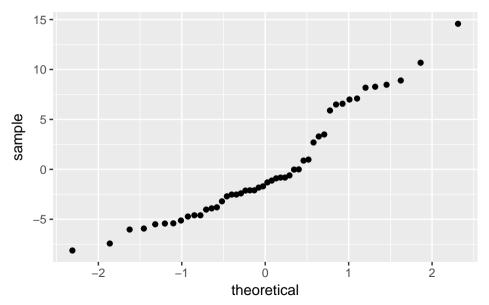


fire retardant model assumptions - equal variance

```
Since n = 12 Levene's test also works:
```

```
## Levene's Test for Homogeneity of Variance (center = median)
## Df F value Pr(>F)
## group 3 0.7138 0.549
## 44
```

fire retardant model assumptions - normality



the general model and analysis (with interaction)

This model has the $(\tau \gamma)$ interaction term:

$$y_{iik} = \mu + \tau_i + \gamma_i + (\tau \gamma)_{ii} + \varepsilon_{iik}$$
 general model

the general model and analysis (with interaction)

This model has the $(\tau \gamma)$ interaction term:

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

The sum of squares decomposition is now:

$$SS_{Total} = SS_A + SS_B + SS_{AB} + SS_{Error}$$

the general model and analysis (with interaction)

This model has the $(\tau \gamma)$ interaction term:

$$y_{ijk} = \mu + \tau_i + \gamma_j + (\tau \gamma)_{ij} + \varepsilon_{ijk}$$
 general model

The sum of squares decomposition is now:

$$SS_{Total} = SS_A + SS_B + SS_{AB} + SS_{Error}$$

The interaction sum of squares is:

$$n\sum_{i,j}\left(y_{ijk}-\overline{y}_{i..}-\overline{y}_{.j.}+\overline{\overline{y}}\right)^{2}$$

Small when additive; large when not.

degrees of freedom - balanced case

$$SS_{Total} = SS_A + SS_B + SS_{AB} + SS_{Error}$$

 $(N-1) = (I-1) + (J-1) + (I-1)(J-1) + IJ(n-1)$

Note: IJ(n-1) = N - IJ

We get (in addition):

$$\frac{SS_{AB}/(I-1)(J-1)}{SS_{Fror}/IJ(n-1)} \sim F_{(I-1)(J-1),IJ(n-1)}$$

If there is evidence for interaction, do not try to interpret the "main effects".

flame retardant with interaction

Flame retardant example, with interaction (and without):

```
##
                  Df Sum Sq Mean Sq F value
                                              Pr(>F)
## Bath
                      715.3
                             715.3 26.726 0.00000549
                      202.1 202.1 7.552
                                             0.00866
## Launderings
                   1 166.1 166.1 6.207
## Bath:Launderings
                                             0.01657
## Residuals
                  44 1177.7
                              26.8
##
             Df Sum Sq Mean Sq F value
                                        Pr(>F)
## Bath
              1 715.3
                        715.3 23.954 0.0000131
                        202.1 6.769
## Launderings
                 202.1
                                        0.0125
## Residuals
             45 1343.8 29.9
```

If n = 1, it is possible to proceed with the analysis, if you can assume there is no interaction.

If n = 1, it is possible to proceed with the analysis, if you can assume there is no interaction.

It isn't possible to test for interaction when n = 1. In principle, because there is no variation within each combination to estimate.

If n = 1, it is possible to proceed with the analysis, if you can assume there is no interaction.

It isn't possible to test for interaction when n = 1. In principle, because there is no variation within each combination to estimate.

Also, the ANOVA table wouldn't work!

If n = 1, it is possible to proceed with the analysis, if you can assume there is no interaction.

It isn't possible to test for interaction when n=1. In principle, because there is no variation within each combination to estimate.

Also, the ANOVA table wouldn't work!

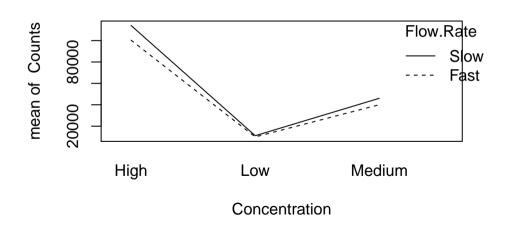
The technique is to use SS_{AB} as the error sum of squares.

another overall example

Chromatography example.

Two factors: flow rate (fast and slow); Concentration (low, med, high)

interaction plot



analysis

```
##
                            Df
                                    Sum Sq
                                               Mean Sq F value
## Flow Rate
                                 364008333
                                             364008333
                                                          29.645
## Concentration
                            2 48365460080 24182730040 1969.424
## Flow.Rate:Concentration 2
                                 203032027
                                             101516013
                                                           8.267
## Residuals
                            24
                                 294698040
                                              12279085
                                          Pr(>F)
##
## Flow Bate
                                       0.0000135
## Concentration
                            < 0.000000000000000000002
## Flow.Rate:Concentration
                                         0.00186
## Residuals
```

assumptions

```
## Levene's Test for Homogeneity of Variance (center = median)
## Df F value Pr(>F)
## group 5 1.5612 0.2089
## 24
```

assumptions

