

# STA257

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DOFORV - distributions of functions  
of (continuous) random variables -  
continued

## DOFORV method 2 - direct "theorem"

I am delighted that the book downplays this method as not as easy to use. Nor do I recommend it for practical use.

**The cdf approach is usually the cleanest and least error prone.**

Theorem: Given  $X$  with density  $f_x(x)$  and  $g$  monotone and differentiable with inverse  $g^{-1}$  where  $f_x(x) > 0$ , let  $Y = g(X)$ . Then:

$$f_y(y) = f_x(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|$$

This theorem can be extended to non-monotonic  $g$ .

# DOFORV - three proofs

The theorem can be proved using the cdf approach and two other ways that are straight outta calculus.

The proofs indeed all look very similar.

Proof 2: uses the "change of variables" method from integration (emphasizes my advice to always think of a density as living in an integral.)

Proof 3: uses the fundamental theorem of calculus.

# a seemingly strange example

The techniques apply to any continuous r.v.  $X$  and to any differentiable, invertible  $g(x)$ .

So let's consider  $X \sim \text{Exp}(\lambda)$  and let  $g(x) = 1 - e^{-\lambda x}$ . It turns out  $Y \sim \text{Unif}[0, 1]$ .

The function  $g$  was not chosen by accident—it is precisely the cdf  $F_X(x)$  of  $X$ .

Theorem: If  $X$  is continuous and has cdf  $F_X(x)$  then  $Y = F_X(X)$  will have a uniform distribution on  $[0, 1]$ .

Proof: ...

## another DOFORV example

Suppose  $X \sim \text{Gamma}(\alpha, \lambda)$  and

$$g(x) = \begin{cases} \frac{1}{x} & : x > 0, \\ 0 & : \text{otherwise.} \end{cases}$$

Determine the distribution of  $Y = g(X)$  by finding its density.

probabilities involving more than  
one random variable at a time

# motivation

A random variable is a function of a sample space, and we care about its distribution.

So far we've focussed on  $X : S \rightarrow \mathbb{R}$ .

Now we will look at  $X : S \rightarrow \mathbb{R}^n$ , which arise quite naturally. We've actually been doing this already sometimes, without saying so explicitly (see Case 2).

Case 1: you *actually observe* multiple things about a particular random outcome (e.g. you measure the weight and blood pressure of a randomly selected study participant.)

Case 2: you are considering a sequence of random variables that "replicate" the same "experiment" (e.g. repeat a Bernoulli( $p$ ) trial  $n$  times...)



# discrete motivating example

Toss two fair six-sided dice. The sample space has 36 elements.

Observe the total and the difference.

Denote the total by  $X_1$  and the difference by  $X_2$ , and  $X = (X_1, X_2)$

$X : S \rightarrow \mathbb{R}^2$  is a random variable and we can consider probabilities of the form  $P(X \in A_1 \times A_2) = P(X_1 \in A_1, X_2 \in A_2)$  for  $A_i \subset \mathbb{R}$ . All such probabilities together form the "distribution" of  $X$ .

But this is an excess of formality. Normally we work directly with the components, in this case  $X_1$  and  $X_2$ , which are two discrete rvs and we can put probabilities of combinations of outcomes in a table.

# table of probabilities

		$X_1$										
		2	3	4	5	6	7	8	9	10	11	12
$X_2$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$
	1	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0
	2	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0
	3	0	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0	0
	4	0	0	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0	0	0
	5	0	0	0	0	0	$\frac{2}{36}$	0	0	0	0	0

# "joint" distribution

*<Insert cannabis joke here>*

Such a table summarizes the distribution of  $X$ , which we'll just call the *joint distribution* of  $X_1$  and  $X_2$ .

The table has all the values of the form  $p(x_1, x_2) = P(X_1 = x_1, X_2 = x_2)$  and is called the *joint* probability mass function.

A joint pmf is non-negative, and its positive values sum to 1 (just like before).

The joint cdf is defined as:  $P(X_1 \leq x_1, X_2 \leq x_2)$  for all  $(x_1, x_2) \in \mathbb{R}^2$ .

It is non-decreasing and right-continuous in both variables, and goes to 0 and 1 as both dimensions go to  $\pm\infty$  respectively.

# marginal distributions

The pmf of any of the component random variables can be recovered by summing over all the others, e.g.:

$$p_{x_1}(x_1) = \sum_{x_2} p(x_1, x_2),$$

(where  $\sum_{x_2}$  denotes summing over all values that  $X_2$  takes on.)

This is called a *marginal* pmf, which characterizes the distribution of that random variable.

example: marginal for  $X_2$

		$X_1$											$p_{X_2}(x_2)$
		2	3	4	5	6	7	8	9	10	11	12	
$X_2$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	0	$\frac{1}{36}$	$\frac{6}{36}$
	1	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{10}{36}$
	2	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0	$\frac{8}{36}$
	3	0	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0	0	$\frac{6}{36}$
	4	0	0	0	0	$\frac{2}{36}$	0	$\frac{2}{36}$	0	0	0	0	$\frac{4}{36}$
	5	0	0	0	0	0	$\frac{2}{36}$	0	0	0	0	0	$\frac{2}{36}$