

# STA257

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## from last time

Theorem:  $X \perp Y$  if and only if the joint cdf  $F(x, y) = F_x(x)F_y(y)$  is the product of the marginal cdfs.

Proof:  $\Leftarrow$  ("only if") too hard;  $\Rightarrow$  left as exercise.

Corollary:  $X \perp Y$  if and only if the joint  $f(x, y) = f_x(x)f_y(y)$

To verify, in practice check two things:

1. The density factors. **Note: enough to factor into a function of  $x$  and a function of  $y$ .**
2. The non-zero region is a rectangle (possibly infinite in either direction.) **Note: technically a "cross product" is all that is needed, but in almost all practical cases it will be a rectangle.**

# other important independence results (advanced)

Theorem: If  $X$  and  $Y$  are independent, so are  $g(X)$  and  $h(Y)$  for any\* functions  $g$  and  $h$ .

Sketch of proof: ...

Definition of independence extends to any number of random variables. We say  $X, Y, \dots, X_n$  are independent if:

$$P(X_1 \in A_1, X_2 \in A_2, \dots, X_n \in A_n) = P(X \in A_1) \cdots P(X_n \in A_n)$$

for any\* subsets  $A_i \in \mathbb{R}$ .

# conditional distributions

Recall the sum  $X$  and the absolute difference  $X$  of two dice:

|       |   | $X_1$          |                |                |                |                |                |                |                |                |                |                |
|-------|---|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|
|       |   | 2              | 3              | 4              | 5              | 6              | 7              | 8              | 9              | 10             | 11             | 12             |
| $X_2$ | 0 | $\frac{1}{36}$ | 0              | $\frac{1}{36}$ | 0              | $\frac{1}{36}$ | 0              | $\frac{1}{36}$ | 0              | $\frac{1}{36}$ | 0              | $\frac{1}{36}$ |
|       | 1 | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              |
|       | 2 | 0              | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | 0              |
|       | 3 | 0              | 0              | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | 0              | 0              |
|       | 4 | 0              | 0              | 0              | 0              | $\frac{2}{36}$ | 0              | $\frac{2}{36}$ | 0              | 0              | 0              | 0              |
|       | 5 | 0              | 0              | 0              | 0              | 0              | $\frac{2}{36}$ | 0              | 0              | 0              | 0              | 0              |

# discrete case

Given a joint pmf for  $X$  and  $Y$  denoted by  $p(x, y)$ , define:

$$p_{x|y}(x|y) = \begin{cases} \frac{p(x,y)}{p_y(y)} & : \text{ where } p_y(y) > 0 \\ 0 & : \text{ otherwise} \end{cases}$$

For any fixed  $Y$  with  $p_y(y) > 0$ , this is a valid pmf.

This pmf describes what is called "the conditional distribution of  $X$  given  $Y = y$ ."

Useful result:

$$p(x, y) = p_{x|y}(x|y)p_y(y)$$
$$p_x(x) = \sum_y p_{x|y}(x|y)p_y(y)$$

# classic example

At home my phone rings  $Y$  times with  $Y \sim \text{Poisson}(\lambda)$  in one hour. I answer the phone with probability  $p$  when it rings. What is the distribution of the  $X$ , the number of times I answer the phone in an hour?

(In fact  $p$  so  $X = 0$  always. Note for people not in attendance...this is a joke about me not answering the phone.)

## continuous case

The concept is similar. We examine a "slice" of the joint density at, say  $X = x$  and consider the distribution of  $Y$  at that fixed value of  $x$ .

The *conditional density of  $Y$  given  $X = x$*  is defined as:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)}$$

wherever  $f_X(x) > 0$ .

Examples:

1.  $f(x, y) = \frac{1}{\pi}$  on  $x^2 + y^2 \leq 1$ .
2.  $f(x, y) = \lambda^2 e^{-\lambda y}$  on  $0 < x < y$ .

the bivariate normal distributions -  
an important class of joint  
distributions



# since civilization is over anyway...

Let's do something **crazy**. Recall  $X \sim N(\mu, \sigma^2)$  has density for all  $x \in \mathbb{R}$ :

$$f_x(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2} \frac{(x - \mu)^2}{\sigma^2}\right)$$

Re-imagine  $x$  and  $\mu$  as a column vectors with one element each:  $\mathbf{x} = \begin{pmatrix} x \end{pmatrix}$  and  $\mu = \begin{pmatrix} \mu \end{pmatrix}$ . Re-imagine  $\sigma^2$  as a  $1 \times 1$  matrix  $\Sigma = \begin{pmatrix} \sigma^2 \end{pmatrix}$ .

Note that  $\det \Sigma = |\Sigma| = \sigma^2$  and  $\Sigma^{-1} = \frac{1}{\sigma^2}$ , and:

$$f_x(x) = \frac{1}{|\Sigma|^{1/2} (2\pi)^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \mu)^t \Sigma^{-1} (\mathbf{x} - \mu)\right)$$

# bivariate normal

The random variables  $X_1$  and  $X_2$  have a bivariate normal distribution with parameters  $\mu_1, \mu_2, \sigma_1^2 > 0, \sigma_2^2 > 0$ , and  $-1 < \rho < 1$  if:

$$f(x_1, x_2) = \frac{1}{|\mathbf{\Sigma}|^{1/2} (2\pi)^{2/2}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^t \mathbf{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)$$

where:

$$\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \quad \boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} \quad \mathbf{\Sigma} = \begin{pmatrix} \sigma_1^2 & \rho\sigma_1\sigma_2 \\ \rho\sigma_1\sigma_2 & \sigma_2^2 \end{pmatrix}$$

# book version of density

Some work will reveal that this is equivalent to the formula given in the textbook.

These densities actually look like "bells".

What if  $\rho = 0$ ?

The densities have the interesting properties that the *marginal distributions* are normal, and the *conditional distributions* are also normal.