

# STA257

Neil Montgomery | HTML is official | PDF versions good for in-class use only  
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# conditional distributions and independence

Suppose  $X$  and  $Y$  are independent so that  $f(x, y) = f_x(x)f_y(y)$ .

Then:

$$\begin{aligned} f_{x|y}(x|y) &= \frac{f(x, y)}{f_y(y)} \\ &= \frac{f_x(x)f_y(y)}{f_y(y)} \\ &= f_x(x) \end{aligned}$$

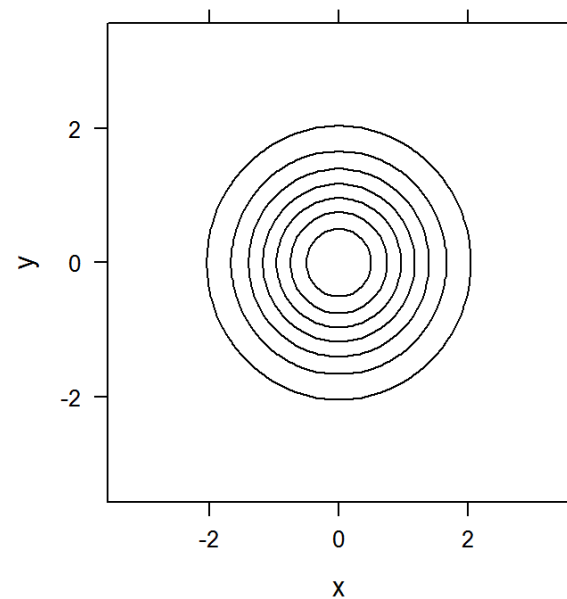
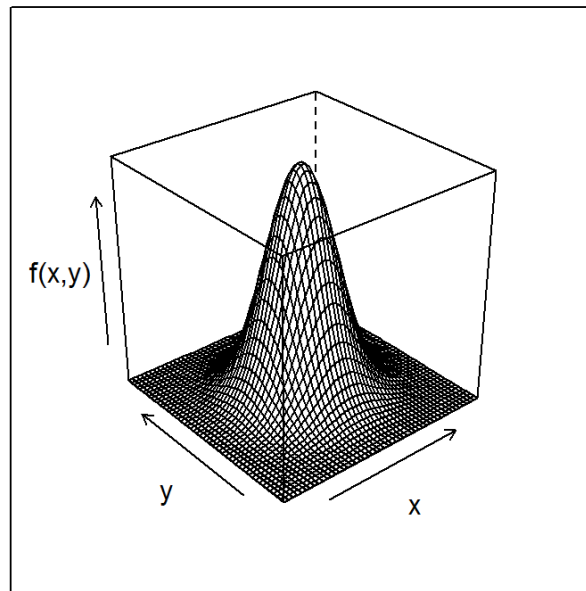
I would argue: *as expected*.

the bivariate normal distributions -  
a very fast introduction to an  
important class of joint distributions

# bivariate normal - independent $N(0, 1)$ case

Start with  $X \sim N(0, 1)$  and  $Y \sim N(0, 1)$  with  $X \perp Y$ . The joint density will be:

$$f(x, y) = \frac{1}{2\pi} \exp\left(-\frac{1}{2}(x^2 + y^2)\right)$$



# bivariate normal - independent general case

If we let  $X \sim N(\mu_x, \sigma_x^2)$  and  $Y \sim N(\mu_y, \sigma_y^2)$  with  $X \perp Y$  the joint density is:

$$\begin{aligned} f(x, y) &= \frac{1}{\sigma_x \sigma_y 2\pi} \exp \left( -\frac{1}{2} \left[ \left( \frac{x - \mu_x}{\sigma_x} \right)^2 + \left( \frac{y - \mu_y}{\sigma_y} \right)^2 \right] \right) \\ &= \frac{1}{\sigma_x \sigma_y 2\pi} \exp \left( -\frac{1}{2} [a^2 + b^2] \right) \end{aligned}$$

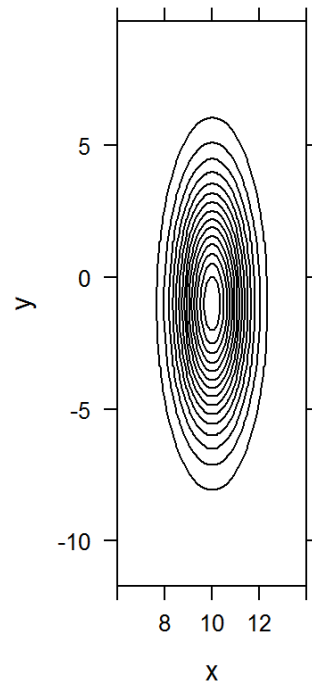
I would argue: "obvious" (?)

$$a = \frac{x - \mu_x}{\sigma_x} \qquad b = \frac{y - \mu_y}{\sigma_y}$$

# free pictures of some joint independent normals

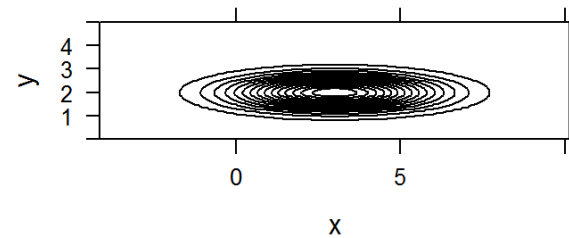
$$\mu_x = 10 \quad \sigma_x^2 = 1$$

$$\mu_y = -1 \quad \sigma_y^2 = 9$$



$$\mu_x = 3 \quad \sigma_x^2 = 4$$

$$\mu_y = 2 \quad \sigma_y^2 = \frac{1}{4}$$



# bivariate normals - general case

This is not obvious. Start with the bivariate independent case, re-jigged:

$$\begin{aligned} f(x, y) &= \frac{1}{\sigma_x \sigma_y 2\pi} \exp \left( -\frac{1}{2} [a^2 + b^2] \right) \\ &= \frac{1}{\sqrt{1 - 0^2} \sigma_x \sigma_y 2\pi} \exp \left( -\frac{1}{2} [a^2 - 2 \cdot 0 \cdot ab + b^2] \right) \end{aligned}$$

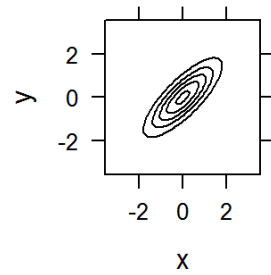
Why? Because I said so.

Now let plug in a new constant  $\rho$  with  $-1 < \rho < 1$  (Why? Because.) where the 0 lives to get the general case:

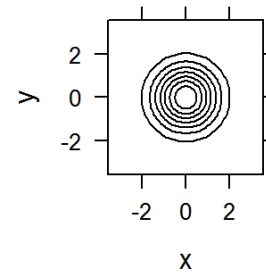
$$\frac{1}{\sqrt{1 - \rho^2} \sigma_x \sigma_y 2\pi} \exp \left( -\frac{1}{2} [a^2 - 2\rho ab + b^2] \right)$$

# free pictures

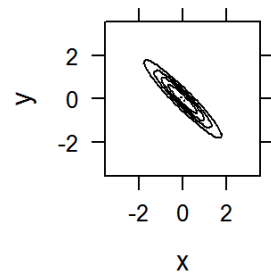
**$\rho=0.8$**



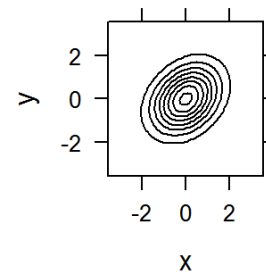
**$\rho=0$**



**$\rho=-0.95$**



**$\rho=0.3$**





## facts (some as exercises)

Fact: Marginals of bivariate normals have normal distribution. (Exercise: very carefully follow technique demonstrated on pp. 82-83 of the textbook.)

Fact: Conditional distributions  $X|Y = y$  and  $Y|X = x$  also have normal distributions. (Exercise: the book gives the "answer" but seems to think it's too hard. It isn't. Use the identity  $a^2 - 2\rho ab + b^2 = (1 - \rho^2)b^2 + (a - \rho b)^2$ )

Fact: Any "slice" along any line  $z = \alpha x + \beta y$  will also have a normal distribution (to be shown later).

Fact: These facts *characterize* bivariate normal distributions (advanced)

Opinion: Bi- and multi-variate normals are best worked with using vector and matrix notation...but that is for a different course.

distributions of functions of  
multiple random variables with  
focus on the continuous case

# stay woke

Probably way more important than the DOFORV situation, but we'll focus on the "classics" and leave the fully general case to other courses.

Here are the classics. Suppose  $X$  and  $Y$  have joint density  $f(x, y)$ . Consider:

$$g_1(x, y) = x + y \quad g_2(x, y) = x / y$$

These are smooth functions  $\mathbb{R}^2 \rightarrow \mathbb{R}$ .

What are the distributions of:

$$W_1 = g_1(X, Y) = X + Y \quad W_2 = g_2(X, Y) = Y / X$$

Technique: the cdf of the  $W_i$  is an integral of  $f(x, y)$  over a certain region.

the classic density formulae:  $W_1 = X + Y$

$$\begin{aligned} f_{w_1}(w) &= \int_{-\infty}^{\infty} f(x, w - x) dx \\ &= \int_{-\infty}^{\infty} f_x(x) f_y(w - x) dx \quad (\text{when } X \perp Y) \end{aligned}$$

Proof:...

Example:  $X \sim N(0, 1)$  and  $Y \sim N(0, 1)$  with  $X \perp Y$ .

Example:  $X \sim \text{Unif}[0, 1]$  and  $Y \sim \text{Unif}[0, 1]$  with  $X \perp Y$ .

the classic density formulae:  $W_2 = Y / X$

$$\begin{aligned} f_{w_2}(w) &= \int_{-\infty}^{\infty} f(x, wx) |x| dx \\ &= \int_{-\infty}^{\infty} f_x(x) f_y(wx) |x| dx \quad (\text{when } X \perp Y) \end{aligned}$$

Proof...

**Mandatory** exercise (done in book as example):  $X \sim N(0, 1)$  and  $Y \sim N(0, 1)$  with  $X \perp Y$ . This is a classic. We say  $W_2$  has a Cauchy distributions.

Example:  $X \sim \text{Exp}(\lambda)$  and  $Y \sim \text{Exp}(\lambda)$  with  $X \perp Y$ .