

STA261 Lecture 1 — 2017-07-05

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admin

contact, notes

date format	YYYY-MM-DD – <i>All Hail ISO8601!!!</i>
instructor	Neil Montgomery
email	neilmontg@gmail.com
office	TBA
office hours	Monday and Wednesday 17:00 – 18:00
website	portal (announcements, grades, suggested exercises, etc.)
github	https://github.com/sta261-summer-2017 (lecture material, code, etc.)

evaluation, book(s), tutorials

what	when	how much
midterm 1	Probably 2017-07-19 18:00 to 20:00 (short lecture after)	25%
midterm 2	Probably 2017-08-02 18:00 to 20:00 (short lecture after)	25%
exam	TBA	50%

Book: Mathematical Statistics and Data Analysis, 3rd ed. by John Rice

Tutorials start Monday.

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The importance of a **Theorem** to you is that each test will ask you to prove one of the theorems from the previous four lectures. The final exam will ask you to prove one of the theorems from the entire course.

what we're going to do, and why

probability versus statistics

The important objects from STA257:

- ▶ Random variable

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In *statistics* we don't know the parameter values, so we'll (... imagine we can...) use a *dataset* to make *inferences* about the parameter values.

mathematical model for the idea of *sample*

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Abbreviation: **i.i.d.**

We might refer to a “parent” or “population” random variable X , with some distribution we might call an “underlying distribution”, and the sample is considered to be i.i.d. “replications” of X .

data, dataset

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The way the dataset was collected will dictate the method of analysis.

probabilistic model for the notion of “dataset” - I

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There could be non-random columns with categorical or numerical information.

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X
X_1
X_2
X_3
X_4
\vdots
X_n

probabilistic model for the notion of “dataset” - II

X	Y
X_1	Y_1
X_2	Y_2
X_3	Y_3
X_4	Y_4
\vdots	\vdots
X_n	Y_n

probabilistic model for the notion of “dataset” - II

"Subject ID"	X	Y
ID345	X_1	Y_1
ID952	X_2	Y_2
ID826	X_3	Y_3
ID118	X_4	Y_4
\vdots	\vdots	\vdots
ID503	X_n	Y_n

probabilistic model for the notion of “dataset” - II

"Subject ID"	X	Y	"Group ID"
ID345	X_1	Y_1	A
ID952	X_2	Y_2	A
ID826	X_3	Y_3	B
ID118	X_4	Y_4	B
\vdots	\vdots	\vdots	\vdots
ID503	X_n	Y_n	A

probabilistic model for the notion of “dataset” - II

"Subject ID"	X	Y	"Group ID"	"InputVariable"
ID345	X_1	Y_1	A	w_1
ID952	X_2	Y_2	A	w_2
ID826	X_3	Y_3	B	w_3
ID118	X_4	Y_4	B	w_4
\vdots	\vdots	\vdots	\vdots	\vdots
ID503	X_n	Y_n	A	w_n

a snippet of real dataset from the wild

Ident	Date	WorkingAge	TakenBy	Fe	Al
448574	11579	3112	EMPL_9134	11	6
448574	11082	1562	EMPL_0592	15	6
448576	12435	5124	EMPL_0917	8	2
448579	11894	1208	EMPL_4925	10	4
448590	11005	210	EMPL_9375	11	6
448595	11039	170	EMPL_0917	12	3
448572	11434	2436	EMPL_0592	13	7
448594	11292	1230	EMPL_0917	24	5

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This explains the centrality of the normal distributions to statistics.

back to probability - distributions of functions of random
variables

single variable case - a general formula

You might (or might not, which is fine!) recall from STA257, when you have a random variable X with density $f_X(x)$, and a monotone, differentiable function g , the density of $Y = g(X)$ is:

$$f_Y(y) = f_X\left(g^{-1}(y)\right) \left| \frac{d}{dy} g^{-1}(y) \right|$$

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Note 1.0: This result is really not all that grand, or mysterious. It is just an application of the “change of variables” or “substitution” strategy from single variable calculus.

bivariate transformations - I

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At some point in MAT237 you will (or will have) learned how to change both variables in such an integral. We'll use this technique in STA261 for a few specific cases.

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We might want to define random variables U and V in terms of X and Y , in general as:

$$U = g_1(X, Y)$$

$$V = g_2(X, Y)$$

Question: what is the joint density for U and V ?

bivariate transformations - II

In the simplest case the transformation is smooth and invertible, so that one can determine inverse functions:

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The “differential” term $g^{-1}(y)$ in the single variable case is played by the *Jacobian*, which is the determinant of the matrix of partial derivatives.

Jacobian, and density formula

$$J = \begin{vmatrix} \frac{\partial h_1}{\partial u} & \frac{\partial h_1}{\partial v} \\ \frac{\partial h_2}{\partial u} & \frac{\partial h_2}{\partial v} \end{vmatrix}$$

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$$f_{U,V}(u, v) = f_{X,Y}(h_1(u, v), h_2(u, v)) |J|$$

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Examples are easier than the formula itself!

bivariate transformation examples

Example 1.1: Suppose X and Y are independent with $N(0, 1)$ distributions. What is the joint density of $U = X + Y$ and $V = X - Y$?

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Example 1.2: Suppose X and Y are independent with joint density $f_{X,Y}(x, y)$. What is the density of $U = X + Y$?

In this example we only have a $g_1(x, y)$. The technique is to add your own $g_2(x, y)$, and find the relevant marginal at the end.

multivariate transformations - we'll do this only once!

This technique works to transform any number of X_1, \dots, X_n into the same number U_1, \dots, U_n using functions g_1, \dots, g_n .

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Find the inverse transformations h_1, \dots, h_n . The Jacobian is now:

$$J = \begin{vmatrix} \frac{\partial h_1}{\partial u_1} & \dots & \frac{\partial h_1}{\partial u_n} \\ \vdots & & \vdots \\ \frac{\partial h_n}{\partial u_1} & \dots & \frac{\partial h_n}{\partial u_n} \end{vmatrix}$$

and the new density is:

$$f_{U_1, \dots, U_n}(u_1, \dots, u_n) = f_{X_1, \dots, X_n}(h_1(u_1, \dots, u_n), \dots, h_n(u_1, \dots, u_n)) |J|$$

other techniques for functions of random variables

Those general formulae are used as a last resort.

We will also make free use of moment generating functions:

$$M_X(t) = E\left(e^{tX}\right)$$

and their most important properties, which are:

1. If the moment generating function exists, it uniquely describes the distribution.

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and their most important properties, which are:

1. If the moment generating function exists, it uniquely describes the distribution.
2. If $X \perp Y$, then $M_{X+Y}(t) = M_X(t)M_Y(t)$

distributions of statistics from normal samples

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It also turns out the central limit theorem has some friends that also converge quickly in practice.

sums of i.i.d. normals, and variations

Proposition 1.3: If X_1, \dots, X_n are i.i.d. $N(\mu, \sigma^2)$, then:

1. $\sum X_i \sim N(n\mu, n\sigma^2)$

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2. $\bar{X} \sim N(\mu, \sigma^2/n)$

sums of i.i.d. normals, and variations

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1. $\sum X_i \sim N(n\mu, n\sigma^2)$
2. $\bar{X} \sim N(\mu, \sigma^2/n)$
3. $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$

squares of i.i.d. standard normals, and their sums

Straightforward result from STA257: $Z \sim N(0, 1)$ then $Z^2 \sim \chi_1^2$.

χ_ν^2 is a nickname for a Gamma $\left(\frac{\nu}{2}, \frac{1}{2}\right)$ distribution, which has density and m.g.f. respectively:

$$f(x) = \frac{(1/2)^{\nu/2}}{\Gamma(\nu/2)} x^{\nu/2-1} e^{-x/2} \quad M(t) = (1 - 2t)^{-\nu/2}$$

Proposition 1.4: If X_1, \dots, X_n are i.i.d. $N(\mu, \sigma^2)$, then:

$$\sum \left(\frac{X_i - \mu}{\sigma} \right)^2 \sim \chi_n^2$$

another χ^2_ν result

Proposition 1.5: If X and Y are independent with $X \sim \chi^2_n$ and $X + Y \sim \chi^2_{n+m}$, then

$$Y \sim \chi^2_m$$

t distributions

Theorem 1.6: Suppose $Z \sim N(0, 1)$ and $U \sim \chi^2_\nu$. Define $T = Z/\sqrt{U/\nu}$. The density of T is:

$$f_T(t) = \frac{\Gamma[(\nu + 1)/2]}{\sqrt{\nu\pi} \Gamma(\nu/2)} \left(1 + \frac{t^2}{\nu}\right)^{-(\nu+1)/2}$$