#### STA286 Lecture 30

**Neil Montgomery** 

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How to misuse hypothesis testing to destroy the universe:

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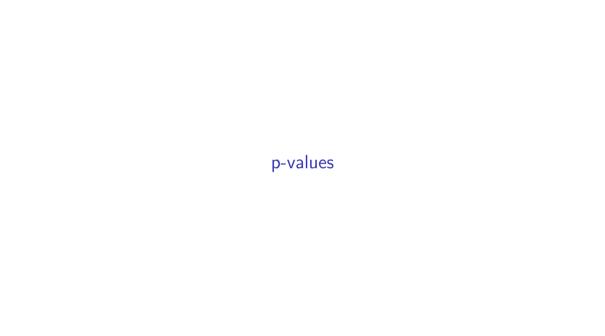
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- 7. Use "one-sided alternatives" because you think you really know what you are doing.
- 8. Think that you really know what you are doing.



Are the doors the right width or not?

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 $H_1: \mu \neq 700$ 

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Use population model  $N(\mu, \sigma = 0.5)$ . Set  $\alpha = 0.05$ . Plan to collect a sample of size n = 10. The rejection region is:

$$\left\{ \overline{X} < 700 - 1.96 \frac{0.5}{\sqrt{10}} \right\} \cup \left\{ \overline{X} > 700 + 1.96 \frac{0.5}{\sqrt{10}} \right\}$$
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You actually measure 10 doors. The sample average is  $\overline{x} = 699.68$ .

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You actually measure 10 doors. The sample average is  $\overline{x} = 699.68$ . We have a 2319! PUSH THE RED BUTTON!!! **REJECT THE NULL!** 

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Everthing has *completely changed*.  $H_0$  is not rejected. FAIL TO REJECT! FAIL TO REJECT! Situation is niner-niner-zero.

But it's too late. The market has decided. Lives are destroyed. Demagogues rise to power in the wake of mass disillusionment.

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Another option is to use something called a *p-value*.

A p-value is the probability (calculated using the  $H_0$  parameter value) of observing a value of the test statistic "more extreme" that what was actually observed.

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In the Doors example,  $\overline{x}$  was thought at first to be 699.68, which is 0.32mm away from 700. The probability of being *more than* 0.32mm away from 700 is:

$$P(\overline{X} < 699.68) + P(\overline{X} > 700.32) = 0.021 + 0.021 = 0.043$$

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After the correction,  $\overline{x}$  is now 699.70. The p-value is now:

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Too bad you didn't know about p-values before causing WWIII.

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Think in terms of orders of magnitude.

# things that annoy me

When people ask me how small a p-value "has" to be.

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When people ask me how small a p-value "has" to be.

When people compute a p-value, and then say "The p-value is smaller than 0.05, so I reject the null hypothesis."

A C.I. formula and a rejection region formula for  $H_0: \mu = \mu_0$  versus  $H_1: \mu \neq \mu_0$  will be based on (something like):

$$P\left(-z_{\alpha/2} < \frac{\overline{X} - \mu}{\sigma/\sqrt{n}} < z_{\alpha/2}\right) = 1 - \alpha$$

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For the C.I., unwrap to isolate  $\mu$  in the middle. For the R.R., put  $\mu=\mu_0$  and unwrap to isolate  $\overline{X}$  in the middle.

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For the C.I., unwrap to isolate  $\mu$  in the middle. For the R.R., put  $\mu=\mu_0$  and unwrap to isolate  $\overline{X}$  in the middle.

The following is true:

Reject 
$$H_0$$
 at level  $\alpha \iff 100 \cdot (1-\alpha)\%$  C.I. contains  $\mu_0$ 

#### the "one-sample t test"

You'll never know  $\sigma$ , so use the data to estimate  $\sigma$  with s, as usual.

From the Doors example where n=10 and  $\overline{x}=699.70$ , suppose also that that s=0.33.

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From the Doors example where n=10 and  $\overline{x}=699.70$ , suppose also that that s=0.33.

The p-value is now calculated based on:

$$\frac{\overline{X}-700}{s/\sqrt{n}}\sim t_9$$

$$P(\overline{X} < 699.70) + P(\overline{X} > 700.30) = P(t_9 < -2.872) + P(t_9 > 2.872) = 0.009 + 0.009 = 0.018$$

#### the "two-sample t-test"

A more realistic hypothesis testing scenario.

Two populations:  $N(\mu_1, \sigma_1)$  and  $N(\mu_2, \sigma_2)$ . The obvious hypotheses will always be:

$$H_0: \mu_1 = \mu_2$$
  
 $H_1: \mu_1 \neq \mu_2$ 

The "parameter" is  $\theta = \mu_1 - \mu_2$ , estimated (as usual) by  $\overline{X_1} - \overline{X_2}$  from samples of sizes  $n_1$  and  $n_2$ .

Two possibilities:

$$rac{\overline{X_1} - \overline{X_2}}{s_p \sqrt{rac{1}{n_1} + rac{1}{n_2}}} \sim t_{n_1 + n_2 - 2} \qquad ext{or} \qquad rac{\overline{X_1} - \overline{X_2}}{\sqrt{rac{s_1^2}{n_1} + rac{s_2^2}{n_2}}} \sim t_
u$$

#### two-sample t-test example

Modified from 10.106. Can nutritional counselling change blood cholesterol level? A group of 15 people received counseling for 8 weeks. A group of 18 people did not.

The readings are made available by the textbook in the following terrible manner:

_	
Treat	Control
129	151
131	132
154	196
172	195
115	188
126	198
175	187
191	168
122	115
238	165
159	137
156	208
176	133
175	217
126	191
	193
	140
	146

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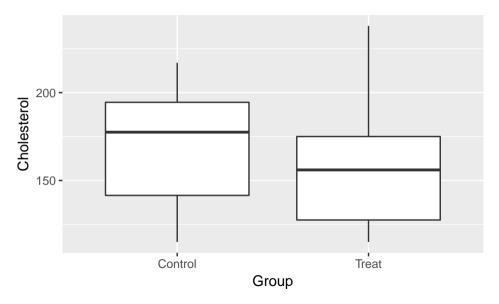
Croup Chologtorol

#### A Real Dataset:

##

##		Group	Cholesterol
##	1	${\tt Treat}$	129
##	2	${\tt Treat}$	131
##	3	${\tt Treat}$	154
##	4	Treat	172
##	5	${\tt Treat}$	115
##	6	Treat	126
##	7	Treat	175
##	8	Treat	191
##	9	Treat	122
##	10	${\tt Treat}$	238
##	11	${\tt Treat}$	159
##	12	${\tt Treat}$	156
##	13	Treat	176
##	14	Treat	175

# two-sample t-test example - plot



# two-sample t-test example - equal variance version

170 00

##

```
## # A tibble: 2 × 4
## Group n X bar S
##
     <fctr> <int> <dbl> <dbl>
## 1 Control 18 170.00 30.788
## 2 Treat 15 156.33 33.090
##
##
   Two Sample t-test
##
## data: Cholesterol by Group
## t = 1.23, df = 31, p-value = 0.23
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -9.0417 36.3750
## sample estimates:
## mean in group Control mean in group Treat
```

156 33

# two-sample t-test example - no variance assumption version

```
##
##
   Welch Two Sample t-test
##
## data: Cholesterol by Group
## t = 1.22, df = 29, p-value = 0.23
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -9.2584 36.5917
## sample estimates:
## mean in group Control mean in group Treat
##
                  170.00
                                        156.33
```