Discrete Time Series

Lecture 07

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Random variable review

Mean and variance of RVs

Expected Value

$$E(X) = \begin{cases} \sum_{x} x \cdot P(X = x) & X \text{ is discrete} \\ \int_{-\infty}^{\infty} x \cdot f(x) \, dx & X \text{ is continuous} \end{cases}$$

Variance

$$Var(X) = E((X - E(X))^{2}) = E(X^{2}) - E(X)^{2}$$

$$= \begin{cases} \sum_{x} (x - E(X))^{2} \cdot P(X = x) & X \text{ is discrete} \\ \int_{-\infty}^{\infty} (x - E(X))^{2} \cdot f(x) dx & X \text{ is continuous} \end{cases}$$

Covariance of RVs

$$\begin{split} Cov(X,Y) &= E\Big(\Big(X-E(X)\Big)\Big(Y-E(Y)\Big)\Big) = E(XY)-E(X)E(Y) \\ &= \begin{cases} \sum_x \Big(x-E(X)\Big)\Big(y-E(Y)\Big) \cdot P(X=x,Y=y) & X \text{ is discrete} \\ \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \Big(x-E(X)\Big)\Big(y-E(Y)\Big) \cdot f(x,y) \text{ dx dy} & X \text{ is continuous} \end{cases} \end{split}$$

$$Corr(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}}$$

Properties of Expected Value

Constant

$$E(c) = c$$
 if c is constant

• Constant Multiplication

$$E(cX) = cE(X)$$

• Constant Addition

$$E(X + c) = E(X) + c$$

Addition

$$E(X + Y) = E(X) + E(Y)$$

Subtraction

$$E(X - Y) = E(X) - E(Y)$$
\$

• Multiplication

$$E(XY) = E(X) E(Y)$$

if X and Y are independent

Properties of Variance

- Constant Var(c) = 0 if c is constant
- Constant Multiplication $Var(cX) = c^2 Var(x)$
- Constant Addition
 Var(X + c) = Var(X)

Addition

$$Var(X + Y) = Var(X) + Var(Y)$$

if X and Y are independent.

Subtraction

$$Var(X - Y) = Var(X) + Var(Y)$$

if X and Y are independent.

Properties of Covariance

Constant

$$Cov(X, c) = 0$$
 if c is constant

• Identity

$$Cov(X, X) = Var(X)$$

• Symmetric

$$Cov(X, Y) = Cov(Y, X)$$

• Constant Multiplication

$$Cov(aX, bY) = ab Cov(X, Y)$$

Constant Addition

$$Cov(X + a, Y + b) = Cov(X, Y)$$

Distribution

$$Cov(aX + bY, cV + dW) = ac Cov(X, V) + ad Cov(X, W) + bc Cov(Y, V) + bd$$

Discrete Time Series

Stationary Processes

A stocastic process (i.e. a time series) is considered to be *strictly stationary* if the properties of the process are not changed by a shift in origin.

In the time series context this means that the joint distribution of $\{y_{t_1},\ldots,y_{t_n}\}$ must be identical to the distribution of $\{y_{t_1+k},\ldots,y_{t_n+k}\}$ for any value of n and k.

Weakly Stationary

Strict stationary is unnecessarily strong / restrictive for many applications, so instead we often opt for weak stationary which requires the following,

1. The process must have finite variance / second moment

$$E(y_t^2) < \infty$$
 for all t

2. The mean of the process must be constant

$$E(y_t) = \mu$$
 for all t

3. The cross moment (covariance) may only depends on the lag (i.e. t-s for y_t and y_s)

$$Cov(y_t, y_s) = Cov(y_{t+k}, y_{s+k})$$
 for all t, s, k

When we say stationary in class we will almost always mean *weakly stationary*.

Autocorrelation

For a stationary time series, where $E(y_t)=\mu$ and $Var(y_t)=\sigma^2$ for all t, we define the autocorrelation at lag k as

$$\rho_{k} = Cor(y_{t}, y_{t+k})$$

$$= \frac{Cov(y_{t}, y_{t+k})}{\sqrt{Var(y_{t})Var(y_{t+k})}}$$

$$= \frac{E((y_{t} - \mu)(y_{t+k} - \mu))}{\sigma^{2}}$$

this is also sometimes written in terms of the autocovariance function (γ_k) as

$$\begin{split} \gamma_k &= \gamma(t,t+k) = Cov(y_t,y_{t+k}) \\ \rho_k &= \frac{\gamma(t,t+k)}{\sqrt{\gamma(t,t)\gamma(t+k,t+k)}} = \frac{\gamma(k)}{\gamma(0)} \end{split}$$

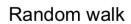
Covariance Structure

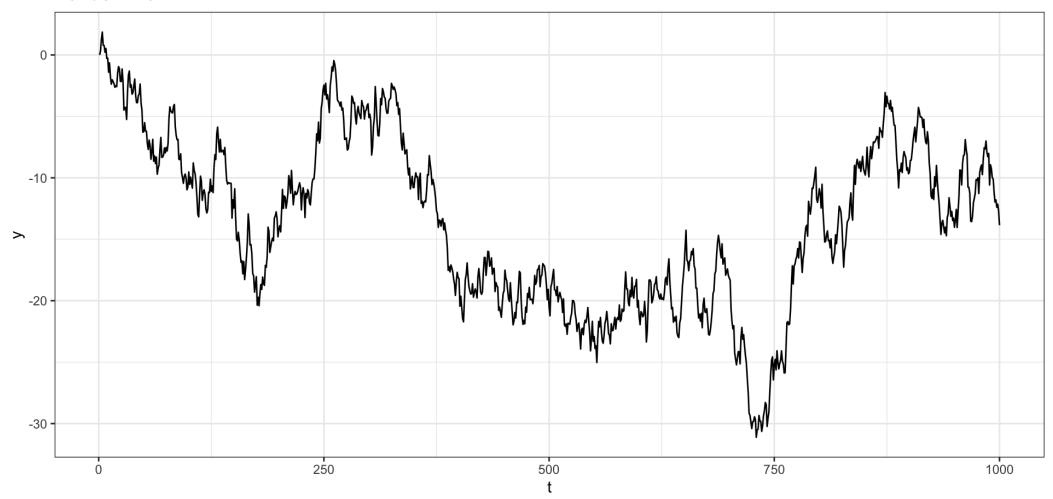
Based on our definition of a (weakly) stationary process, it implies a covariance of the following structure,

$$\Sigma = \begin{pmatrix} \gamma(0) & \gamma(1) & \gamma(2) & \gamma(3) & \cdots & \gamma(n-1) & \gamma(n) \\ \gamma(1) & \gamma(0) & \gamma(1) & \gamma(2) & \cdots & \gamma(n-2) & \gamma(n-1) \\ \gamma(2) & \gamma(1) & \gamma(0) & \gamma(1) & \cdots & \gamma(n-3) & \gamma(n-2) \\ \gamma(3) & \gamma(2) & \gamma(1) & \gamma(0) & \cdots & \gamma(n-4) & \gamma(n-3) \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ \gamma(n-1) & \gamma(n-2) & \gamma(n-3) & \gamma(n-4) & \cdots & \gamma(0) & \gamma(1) \\ \gamma(n) & \gamma(n-1) & \gamma(n-2) & \gamma(n-3) & \cdots & \gamma(1) & \gamma(0) \end{pmatrix}$$

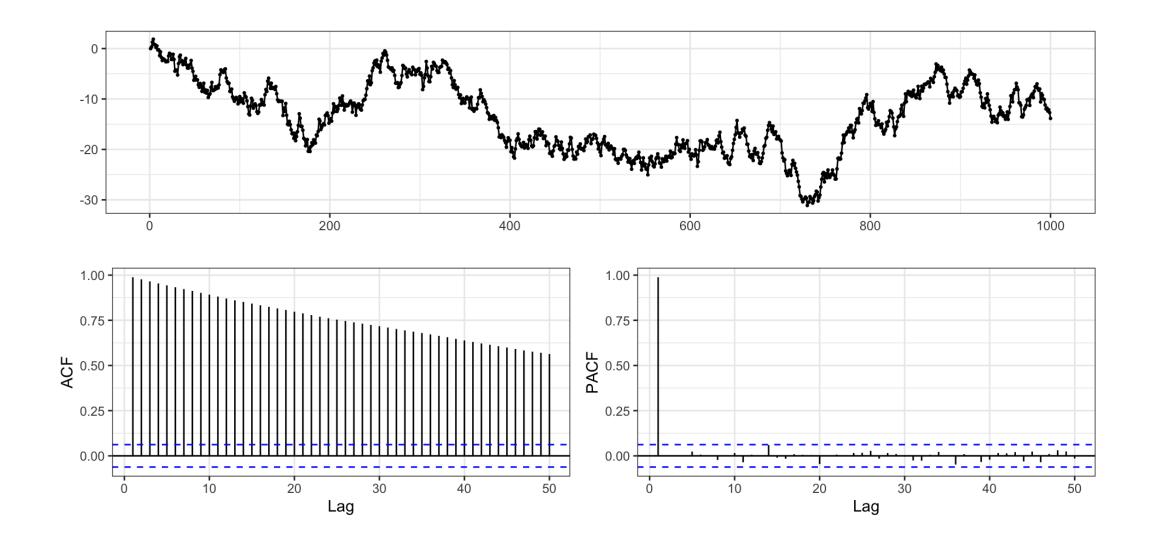
Example - Random walk

Let $y_t = y_{t-1} + w_t$ with $y_0 = 0$ and $w_t \sim N(0, 1)$.





ACF + PACF



Stationary?

Is y_t stationary?

Partial Autocorrelation - pACF

Given these type of patterns in the autocorrelation we often want to examine the relationship between y_t and y_{t+k} with the (linear) dependence of y_t on y_{t+1} through y_{t+k-1} removed.

This is done through the calculation of a partial autocorrelation ($\alpha(k)$), which is defined as follows:

$$\begin{split} \alpha(0) &= 1 \\ \alpha(1) &= \rho(1) = Cor(y_t, y_{t+1}) \\ &\vdots \\ \alpha(k) &= Cor(y_t - P_{t,k}(y_t), \ y_{t+k} - P_{t,k}(y_{t+k})) \end{split}$$

where $P_{t,k}(y)$ is the project of y onto the space spanned by $y_{t+1}, \dots, y_{t+k-1}$.

pACF - Calculation

Let $\rho(k)$ be the autocorrelation for the process at lag k then the partial autocorrelation at lag k will be $\phi(k,k)$ given by the Durbin-Levinson algorithm,

$$\phi(k,k) = \frac{\rho(k) - \sum_{t=1}^{k-1} \phi(k-1,t) \rho(k-t)}{1 - \sum_{t=1}^{k-1} \phi(k-1,t) \rho(t)}$$

where

$$\phi(k,t) = \phi(k-1,t) - \phi(k,k) \phi(k-1,k-t)$$

Starting with $\phi(1,1) = \rho(1)$ we can solve iteratively for $\phi(2,2),\ldots,\phi(k,k)$.

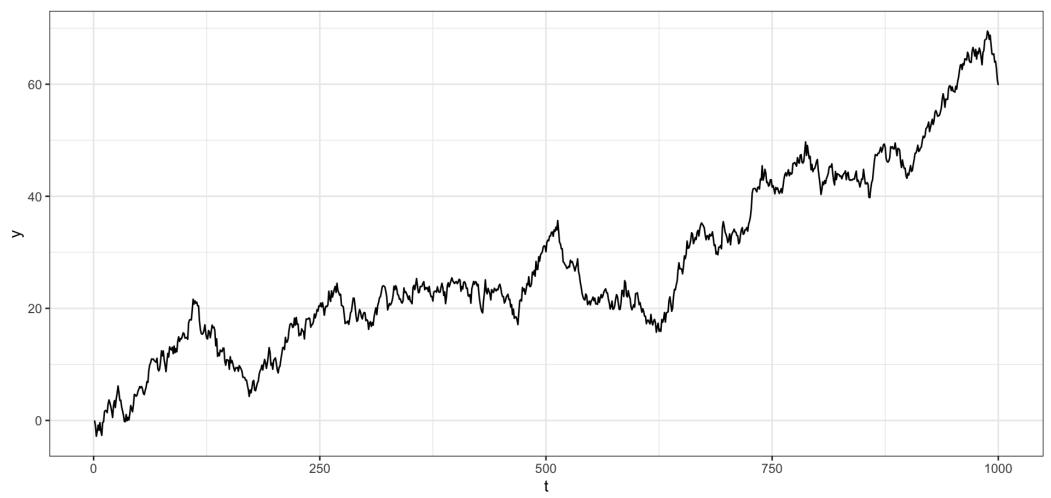
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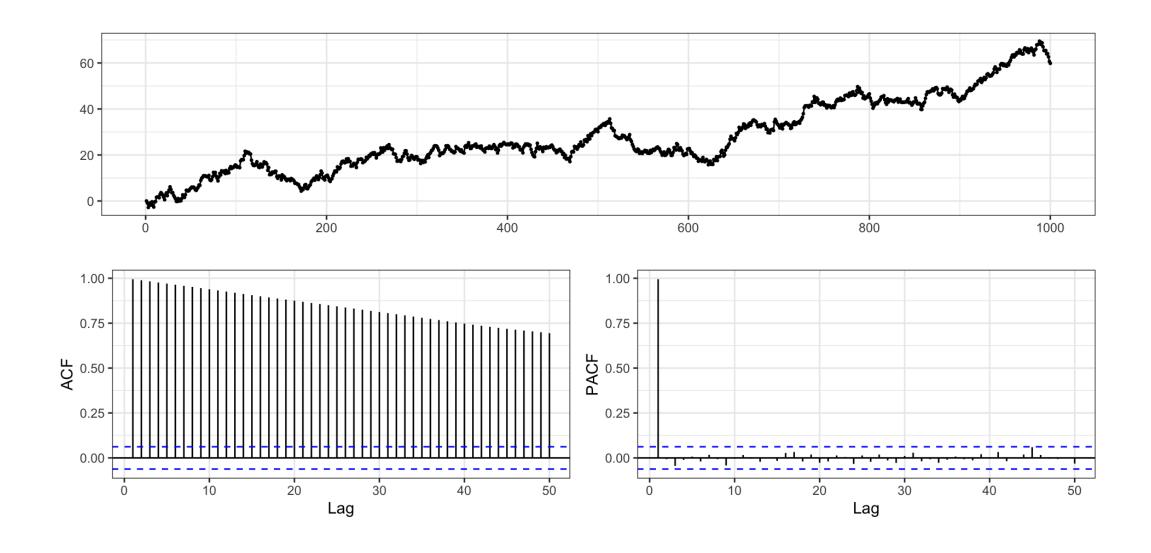
Example - Random walk with drift

Let $y_t = \delta + y_{t-1} + w_t$ with $y_0 = 0$ and $w_t \sim N(0, 1)$.

Random walk with trend



ACF + PACF



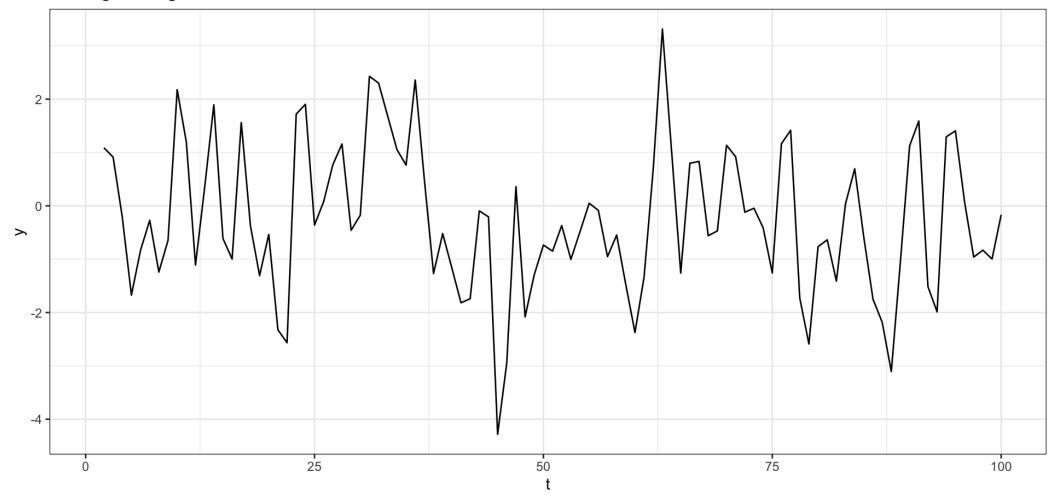
Stationary?

Is y_t stationary?

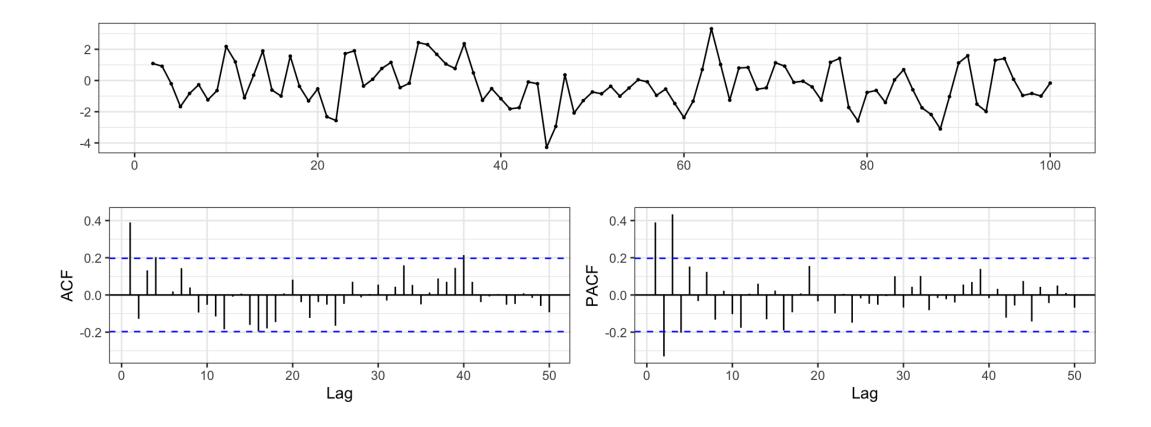
Example - Moving Average

Let $w_t \sim N(0, 1)$ and $y_t = w_{t-1} + w_t$.

Moving Average



ACF + PACF

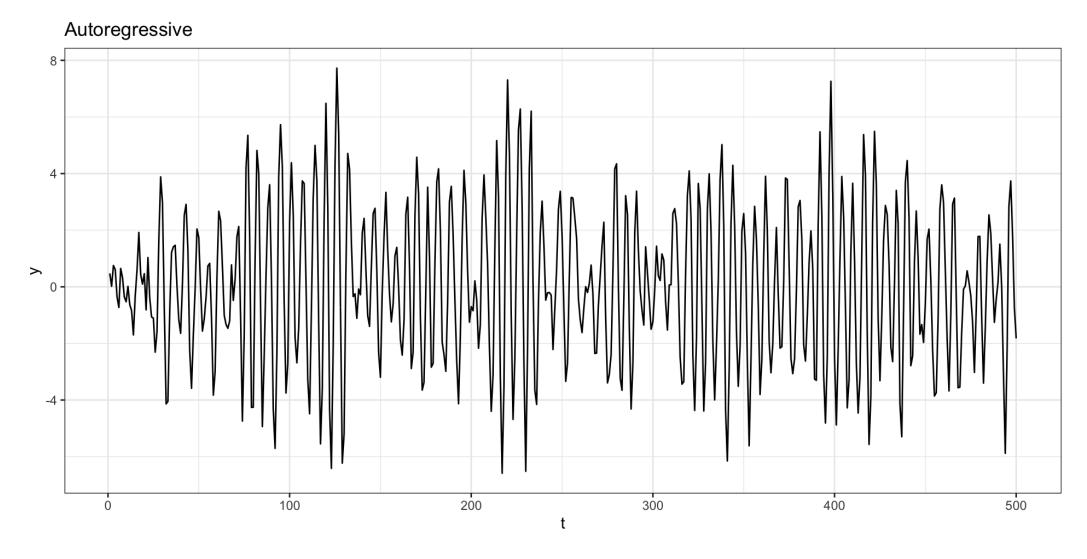


Stationary?

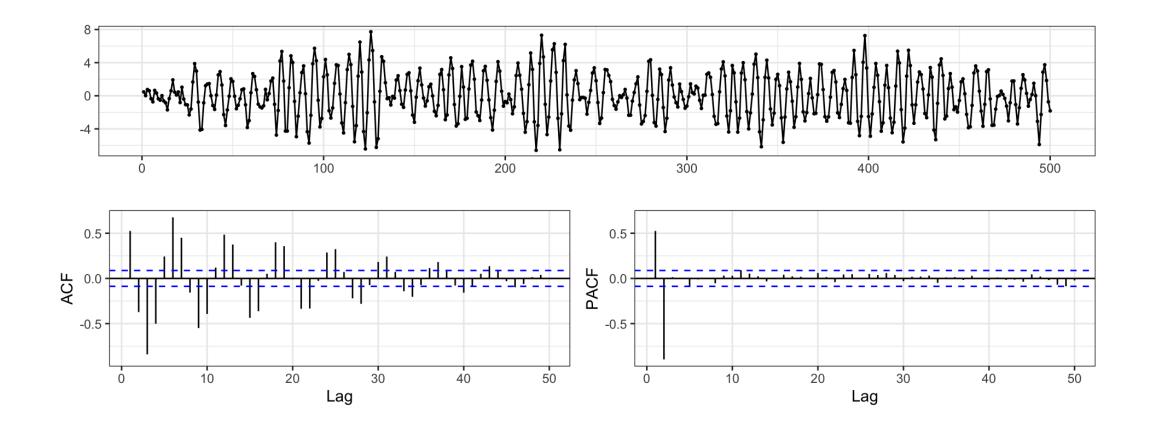
Is y_t stationary?

Autoregressive

Let $w_t \sim N(0, 1)$ and $y_t = y_{t-1} - 0.9y_{t-2} + w_t$ with $y_t = 0$ for t < 1.



ACF + PACF

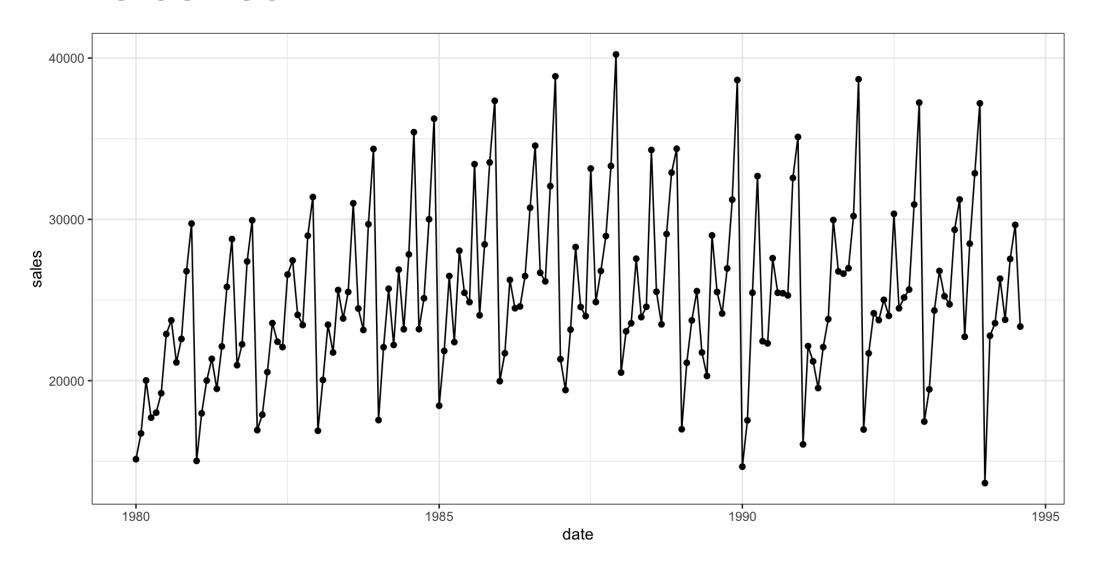


Example - Australian Wine Sales

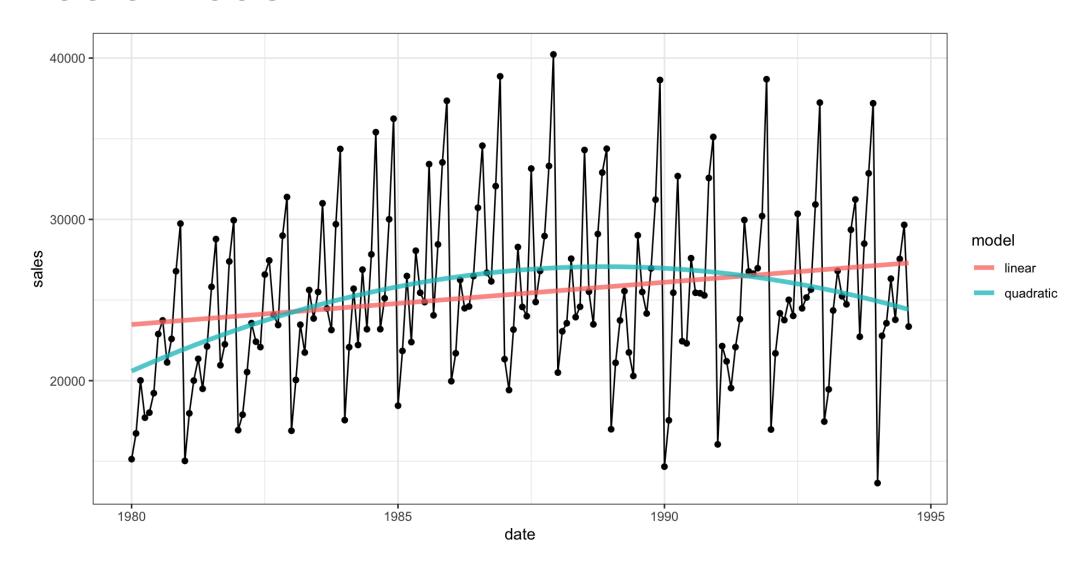
Australian total wine sales by wine makers in bottles <= 1 litre. Jan 1980 – Aug 1994.

```
aus wine = readRDS("data/aus wine.rds")
    aus_wine
# A tibble: 176 × 2
    date sales
   <dbl> <dbl>
 1 1980 15136
 2 1980, 16733
 3 1980. 20016
 4 1980, 17708
 5 1980. 18019
 6 1980, 19227
 7 1980, 22893
 8 1981, 23739
 9 1981, 21133
```

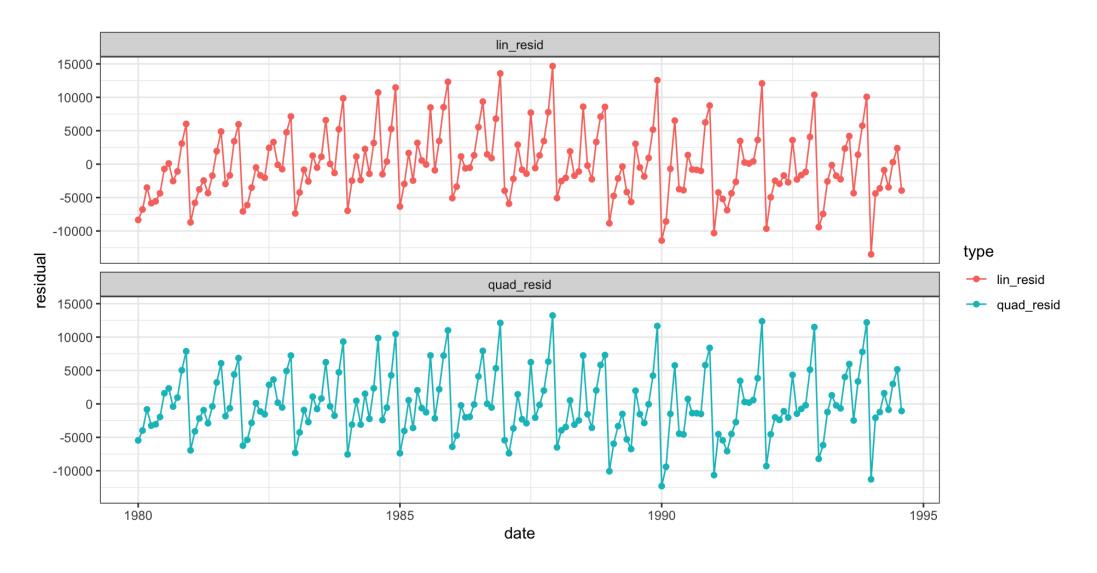
Time series



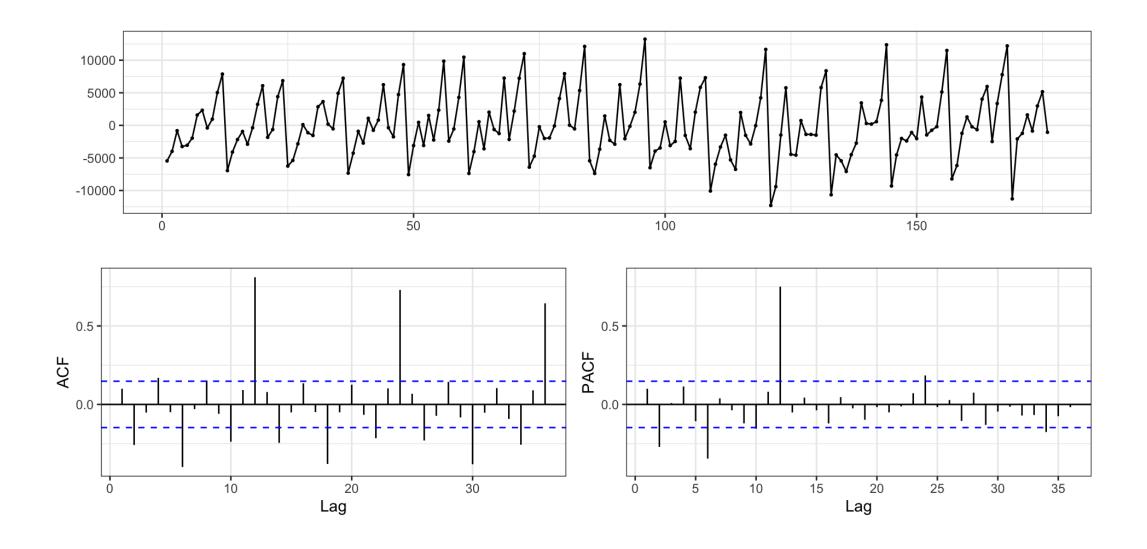
Basic Model Fit

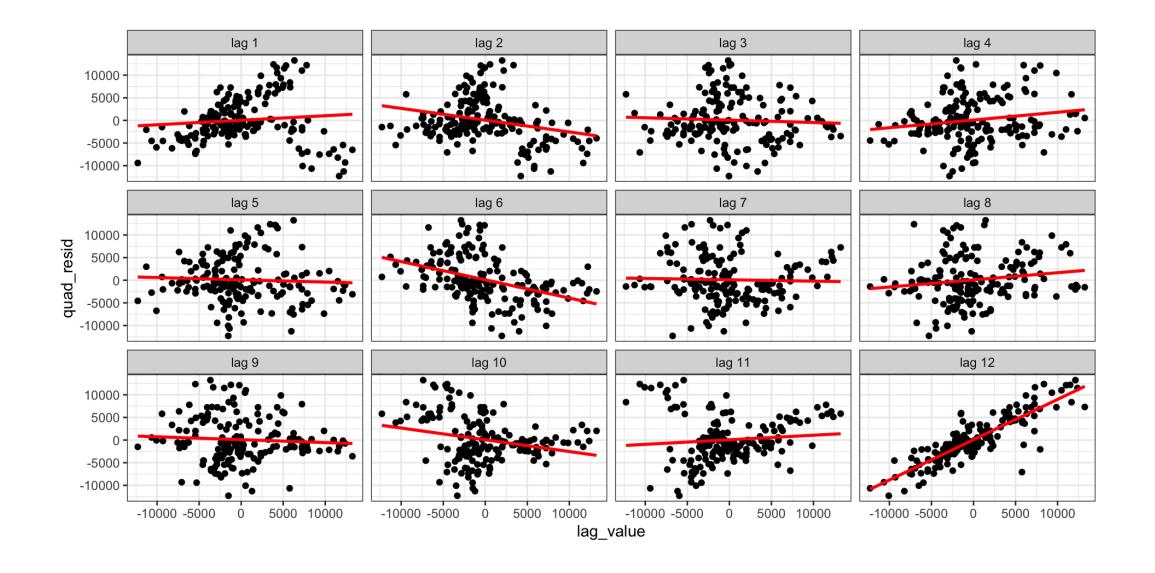


Residuals



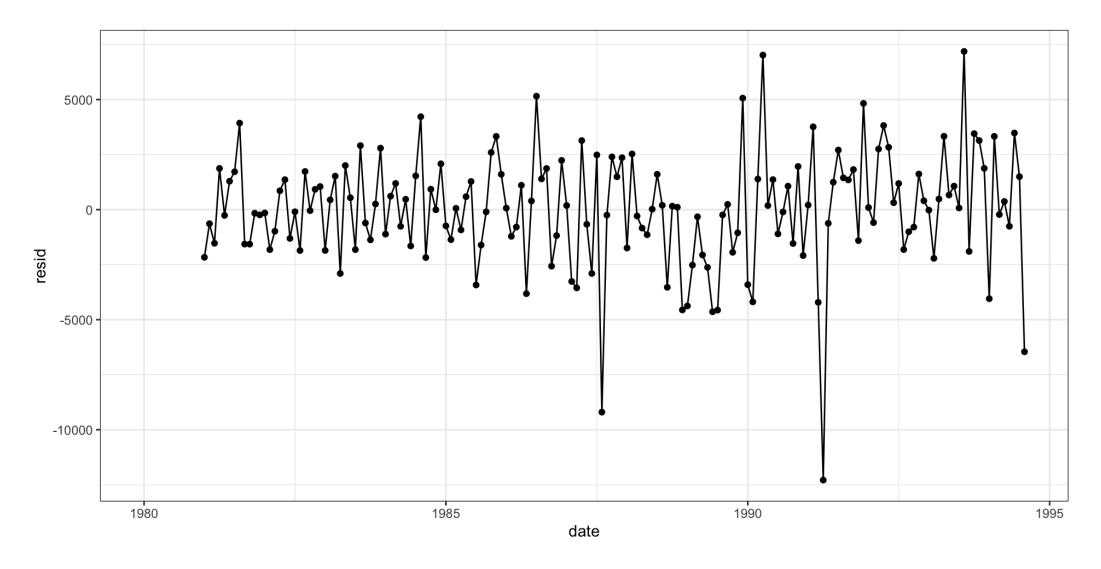
Autocorrelation Plot



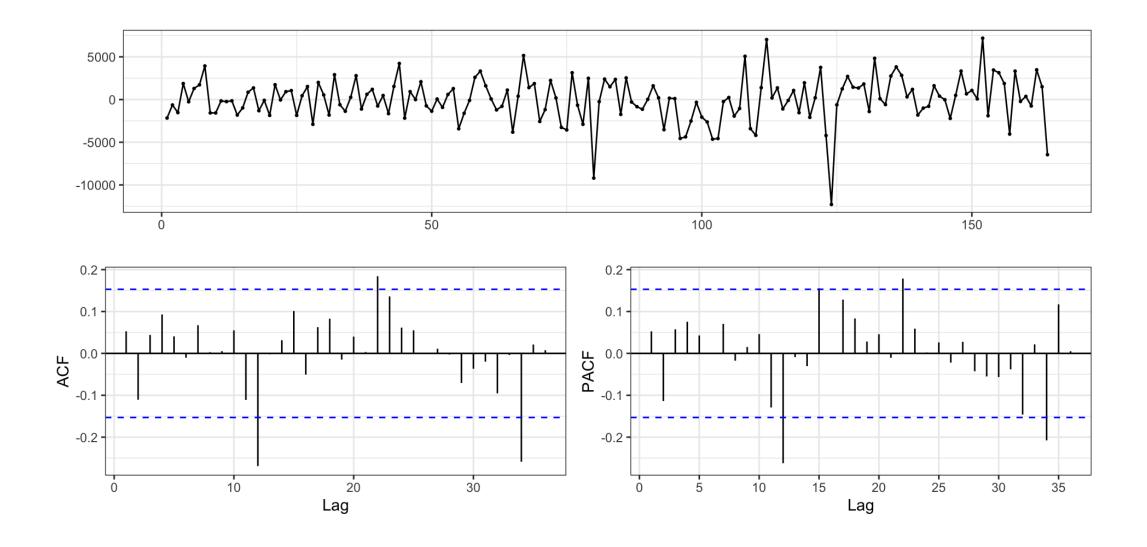


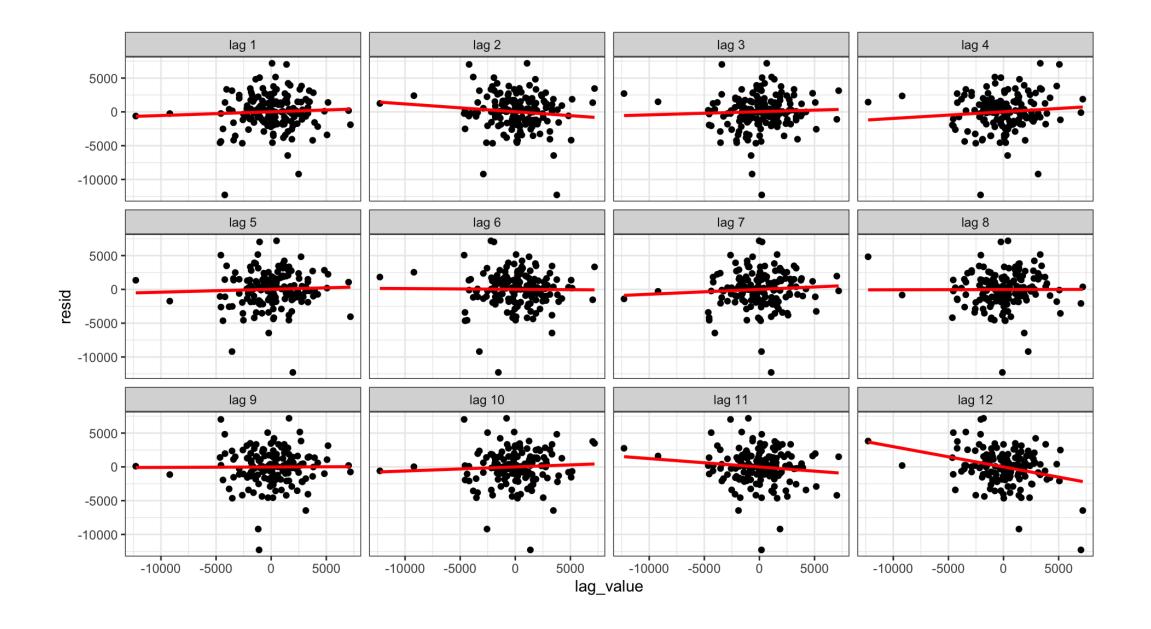
Auto regressive errors

Residual residuals



Residual residuals - acf





Writing down the model?

So, is our EDA suggesting that we fit the following model?

$$sales_t = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 sales_{t-12} + \epsilon_t$$

the model we actually fit is,

$$sales_t = \beta_0 + \beta_1 t + \beta_2 t^2 + w_t$$

where

$$\mathbf{w}_{t} = \delta \mathbf{w}_{t-12} + \epsilon_{t}$$