

Covariance Functions

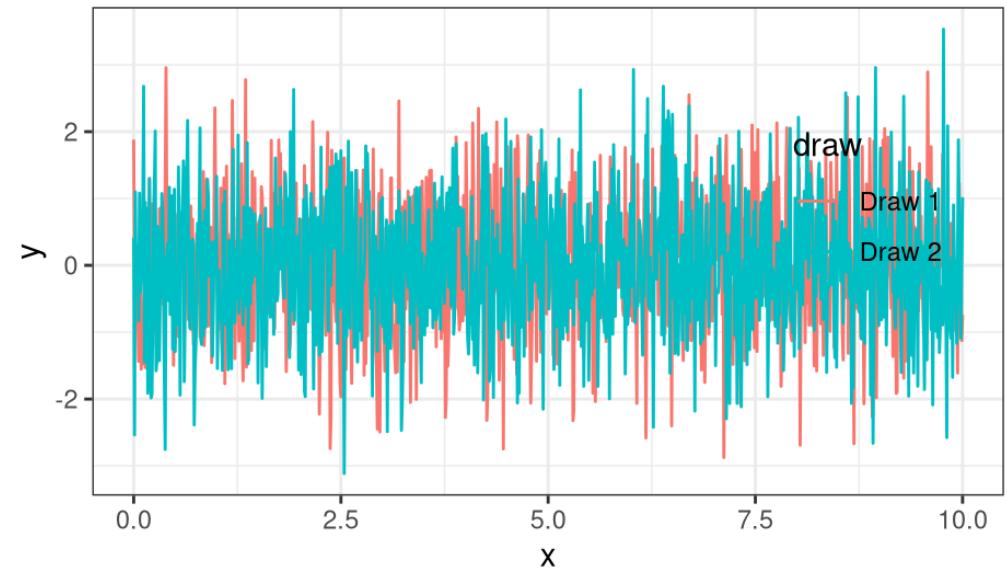
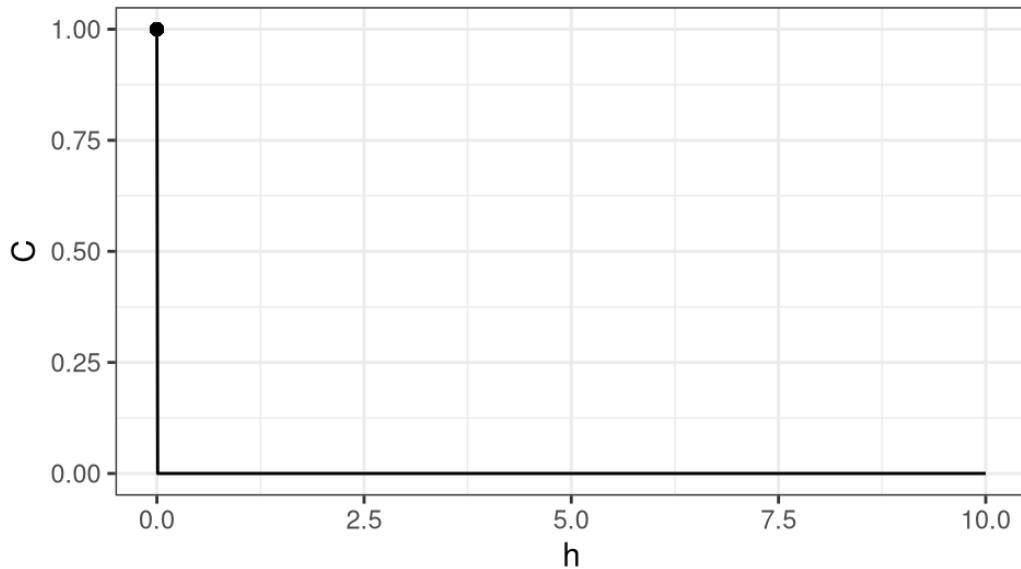
Lecture 15

Dr. Colin Rundel

More Covariance Functions

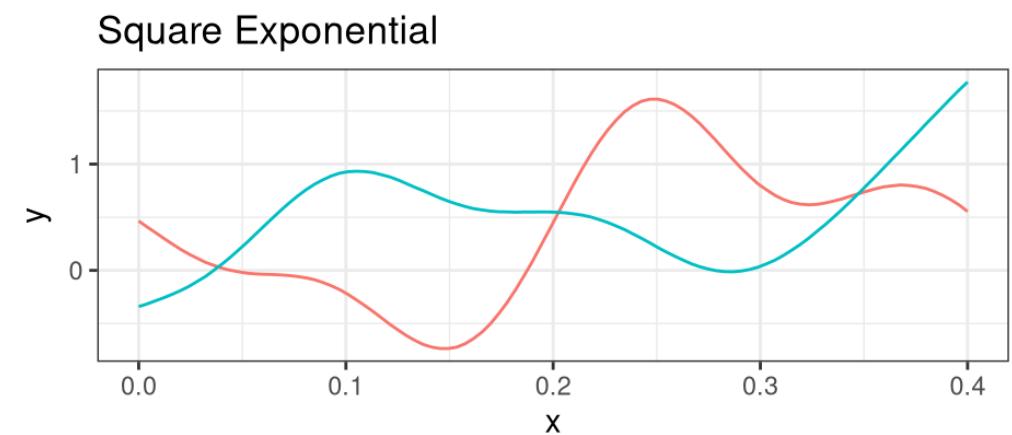
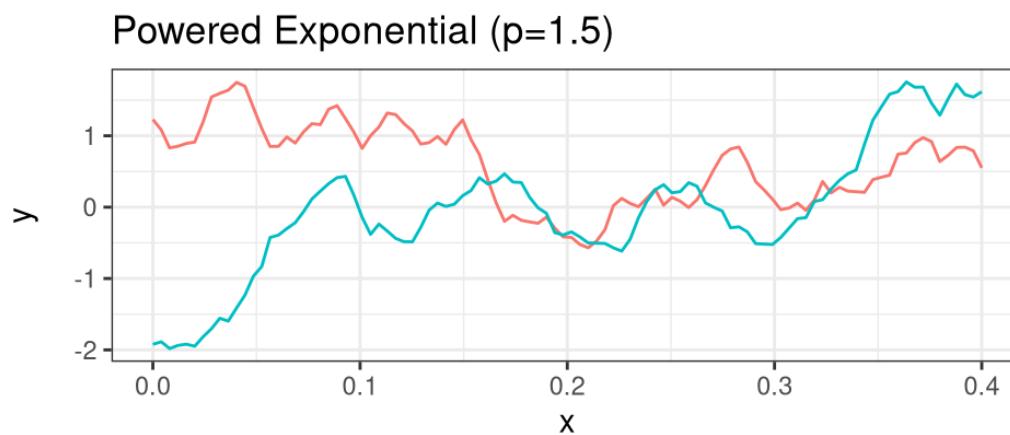
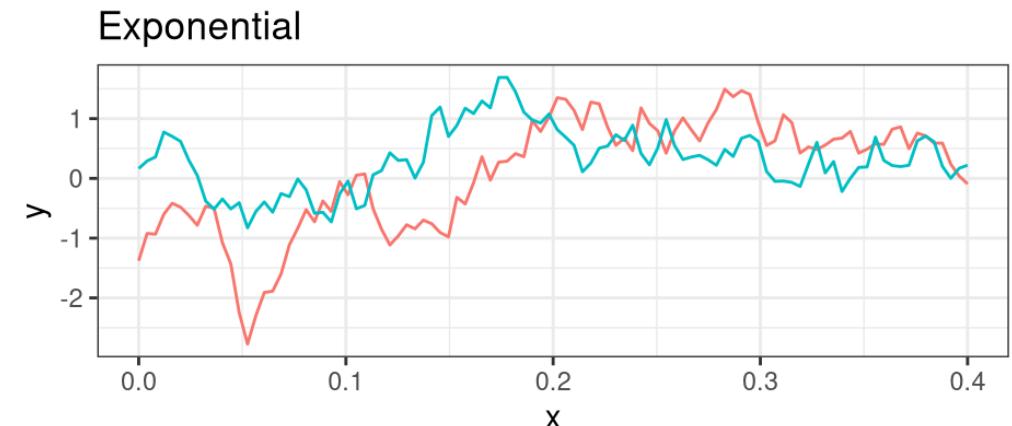
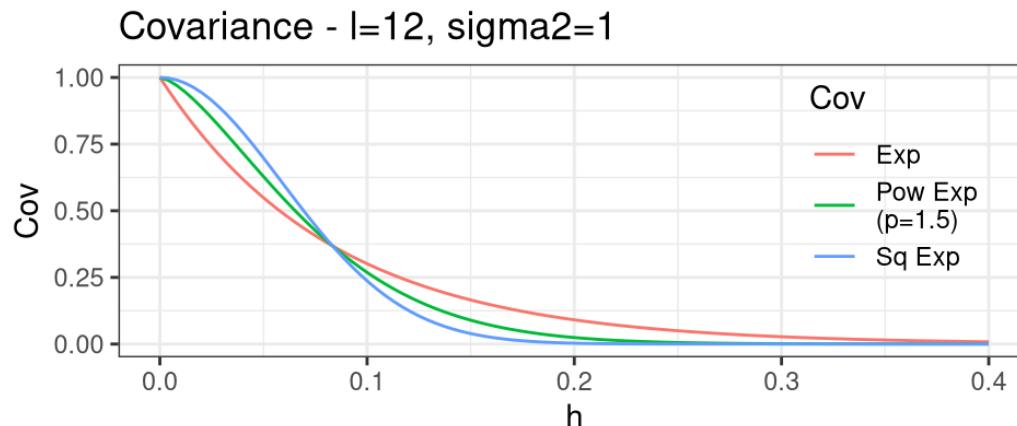
Nugget Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma^2 1_{\{h=0\}} \text{ where } h = |t_i - t_j|$$



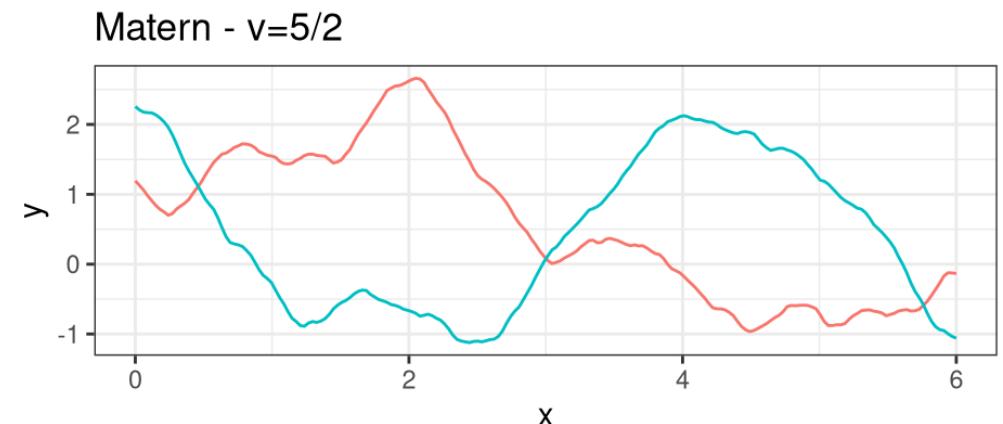
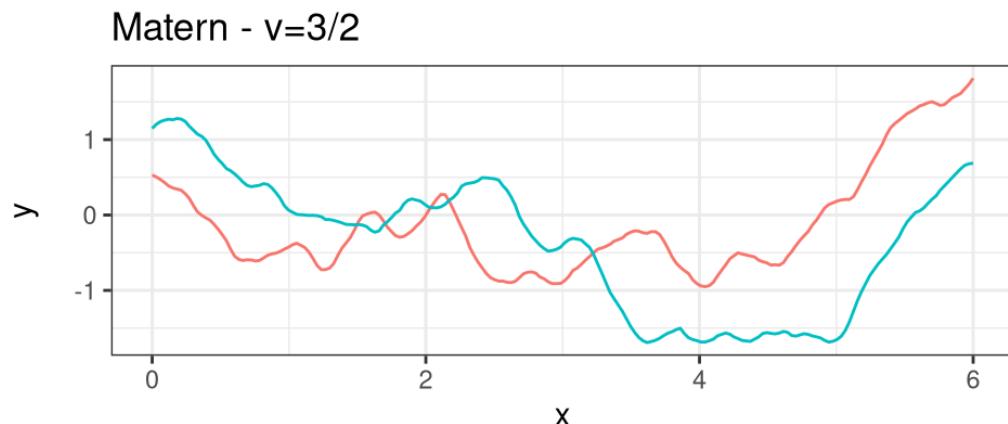
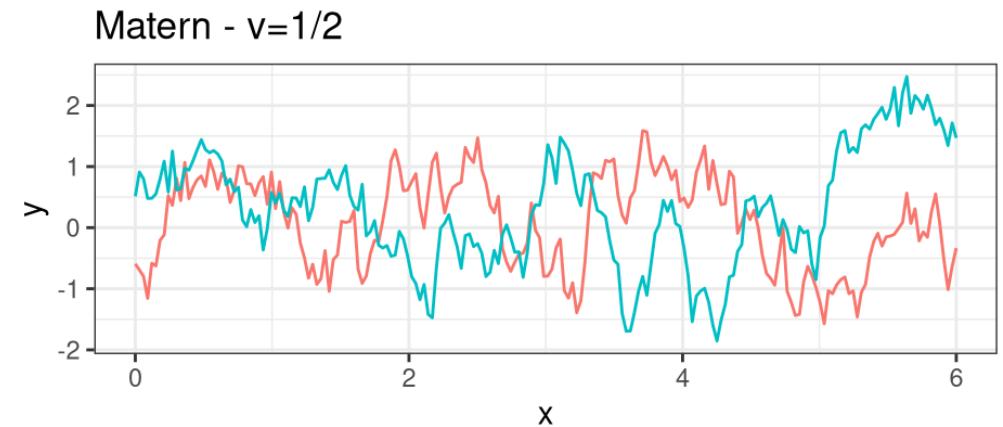
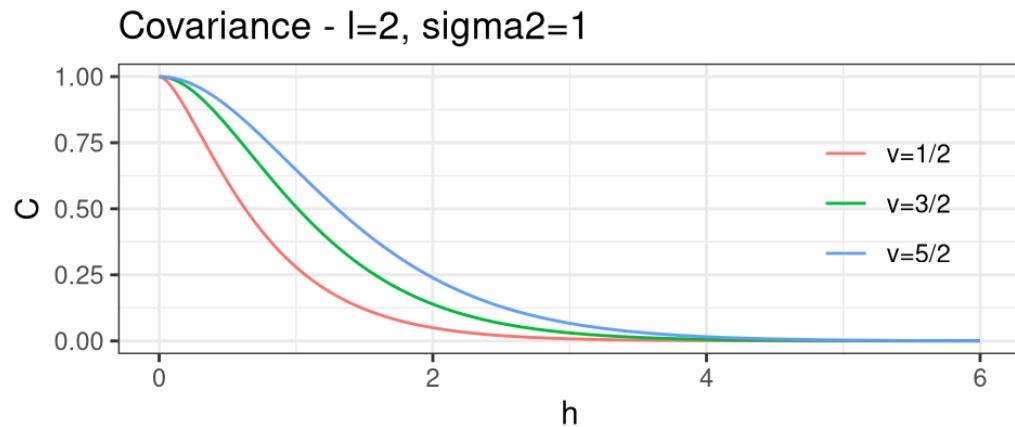
(- / Powered / Square) Exponential Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma^2 \exp(-(h l)^p) \text{ where } h = |t_i - t_j|$$



Matern Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma^2 \frac{2^{1-v}}{\Gamma(v)} (\sqrt{2v} h \cdot 1)^v K_v(\sqrt{2v} h \cdot 1) \text{ where } h = |t_i - t_j|$$



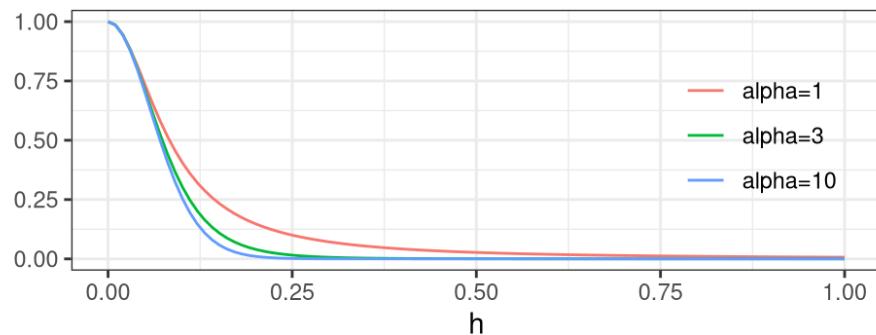
Matern Covariance

- $K_v()$ is the modified Bessel function of the second kind.
- A Gaussian process with Matérn covariance has sample functions that are $\lceil v - 1 \rceil$ times differentiable.
- When $v = 1/2 + p$ for $p \in \mathbb{N}^+$ then the Matern has a simplified form
- When $v = 1/2$ the Matern is equivalent to the exponential covariance.
- As $v \rightarrow \infty$ the Matern converges to the squared exponential covariance.

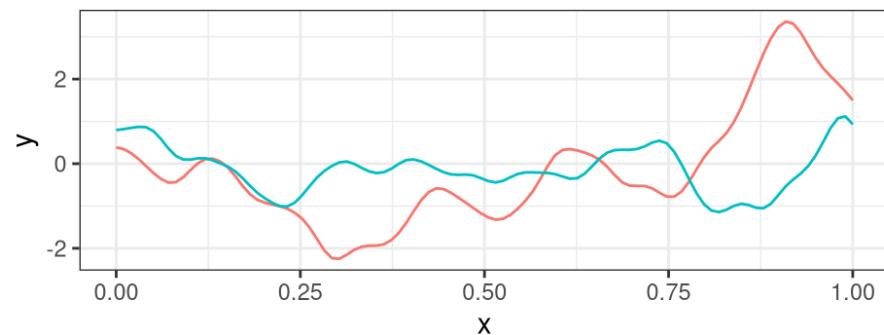
Rational Quadratic Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma^2 \left(1 + \frac{h^2 l^2}{\alpha} \right)^{-\alpha} \quad \text{where } h = |t_i - t_j|$$

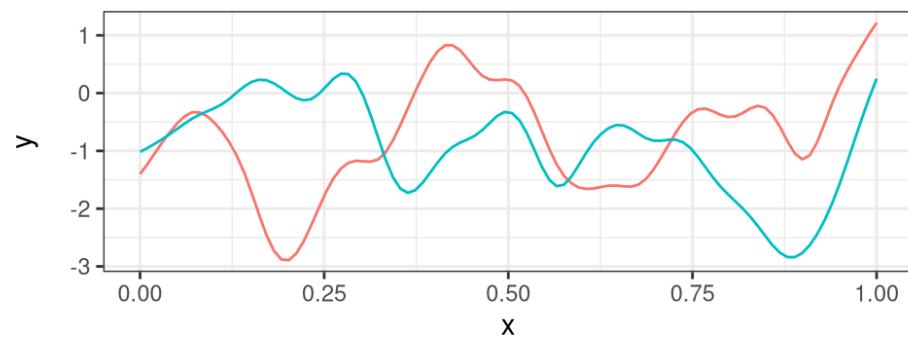
Covariance - $l=12$, $\sigma^2=1$



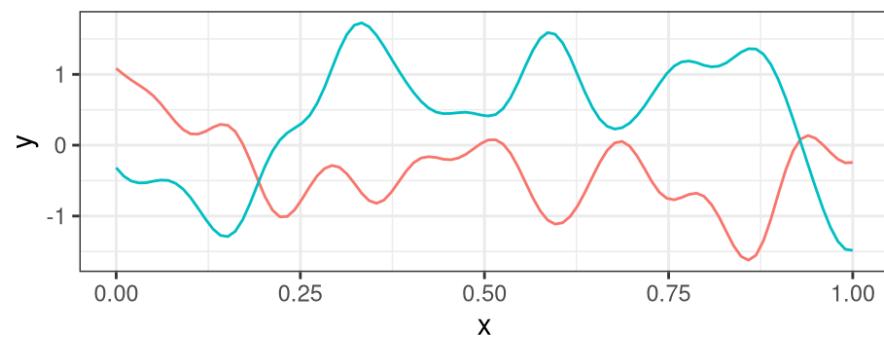
Rational Quadratic - $\alpha=1$



Rational Quadratic - $\alpha=3$



Rational Quadratic - $\alpha=10$

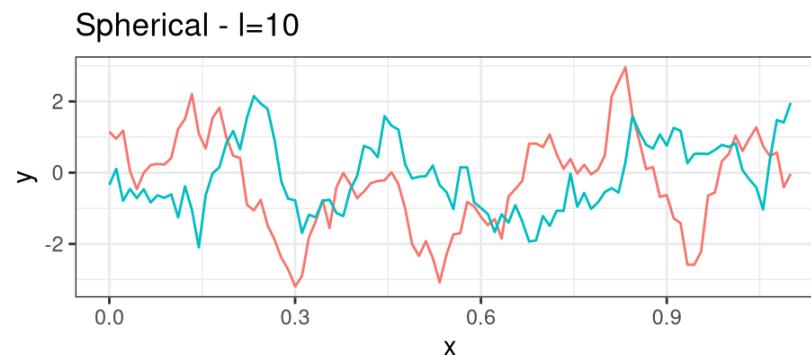
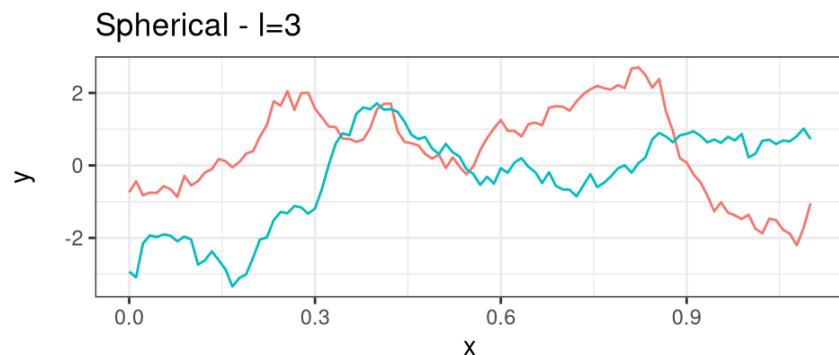
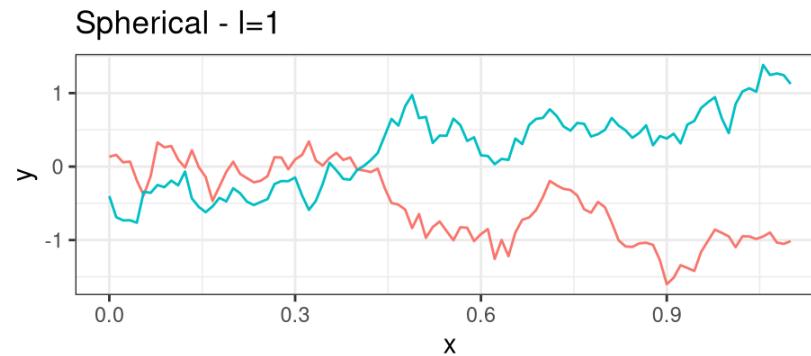
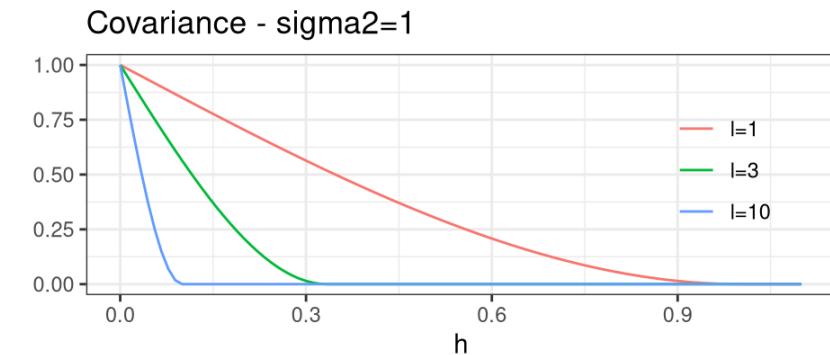


Rational Quadratic Covariance

- is a scaled mixture of squared exponential covariance functions with different characteristic inverse length-scales (1).
- As $\alpha \rightarrow \infty$ the rational quadratic converges to the square exponential covariance.
- Has sample functions that are infinitely differentiable for any value of α

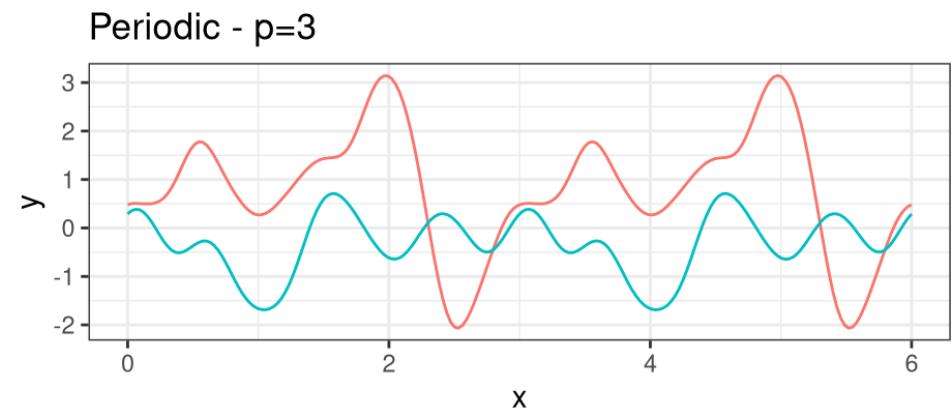
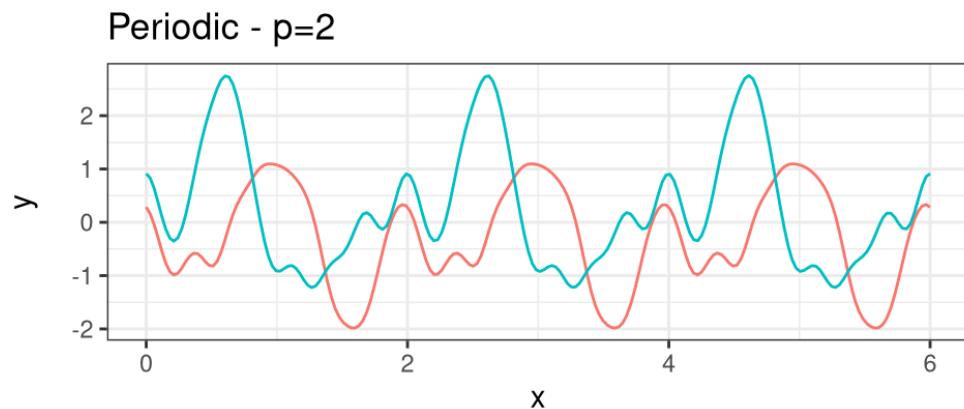
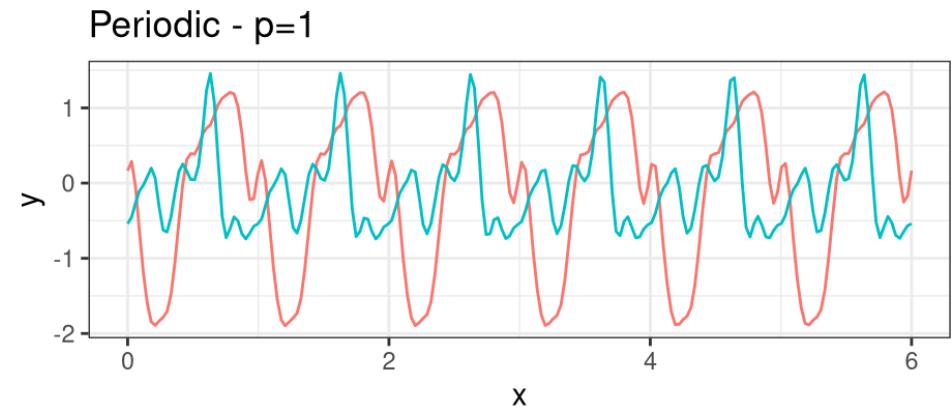
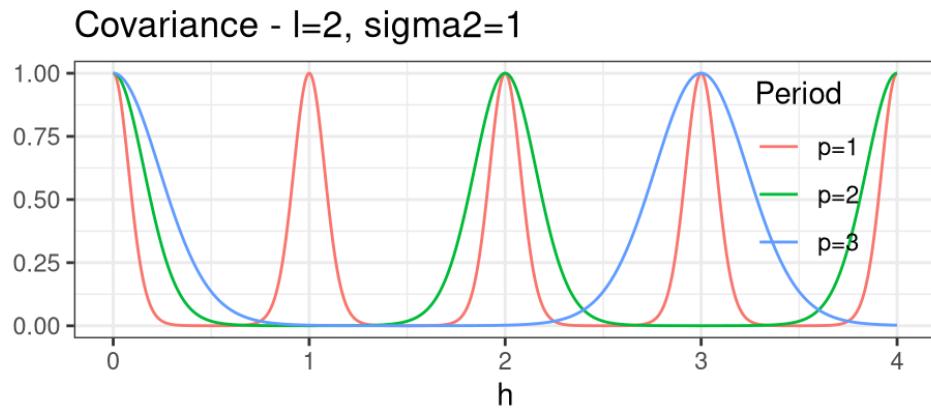
Spherical Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \begin{cases} \sigma^2 \left(1 - \frac{3}{2}h \cdot 1 + \frac{1}{2}(h \cdot 1)^3 \right) & \text{if } 0 < h < 1 \\ 0 & \text{otherwise} \end{cases}$$



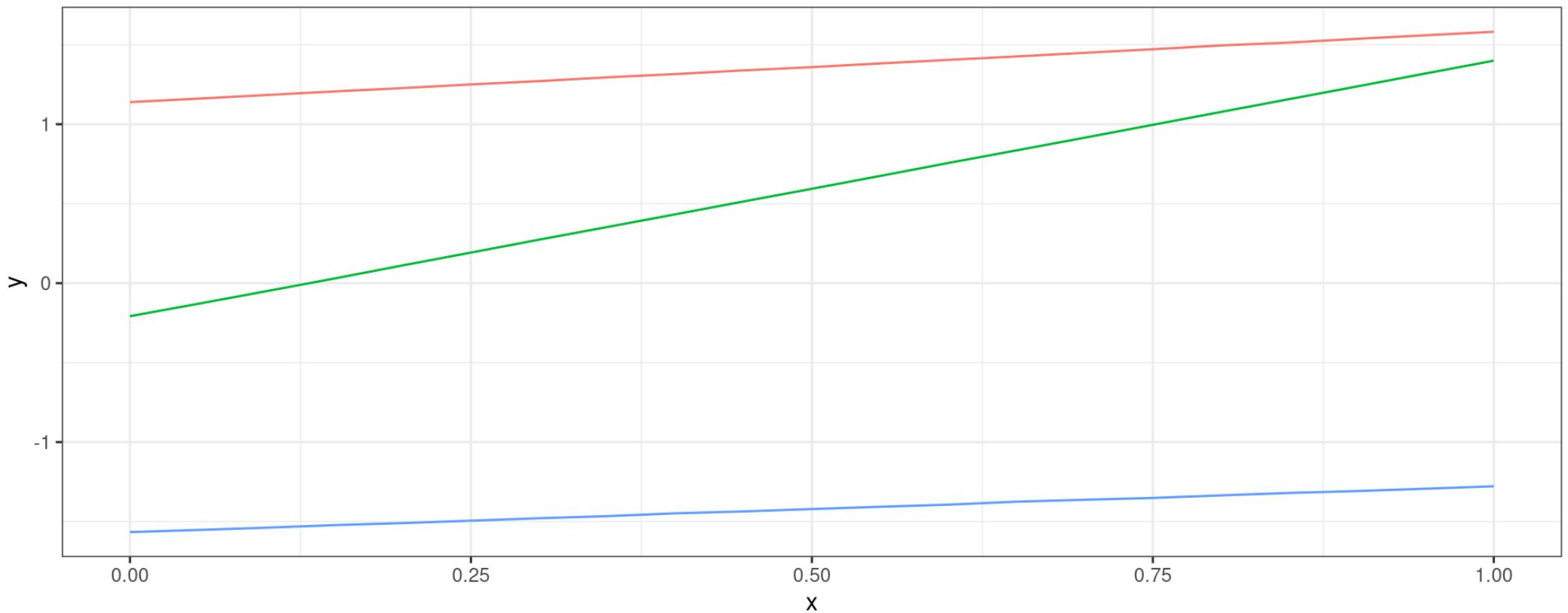
Periodic Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma^2 \exp\left(-2 l^2 \sin^2\left(\pi \frac{h}{p}\right)\right) \text{ where } h = |t_i - t_j|$$



Linear Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma_b^2 + \sigma_v^2 (t_i - c)(t_j - c)$$



Combining Covariances

If we definite two valid covariance functions, $\text{Cov}_a(y_{t_i}, y_{t_j})$ and $\text{Cov}_b(y_{t_i}, y_{t_j})$ then the following are also valid covariance functions,

$$\text{Cov}_a(y_{t_i}, y_{t_j}) + \text{Cov}_b(y_{t_i}, y_{t_j})$$

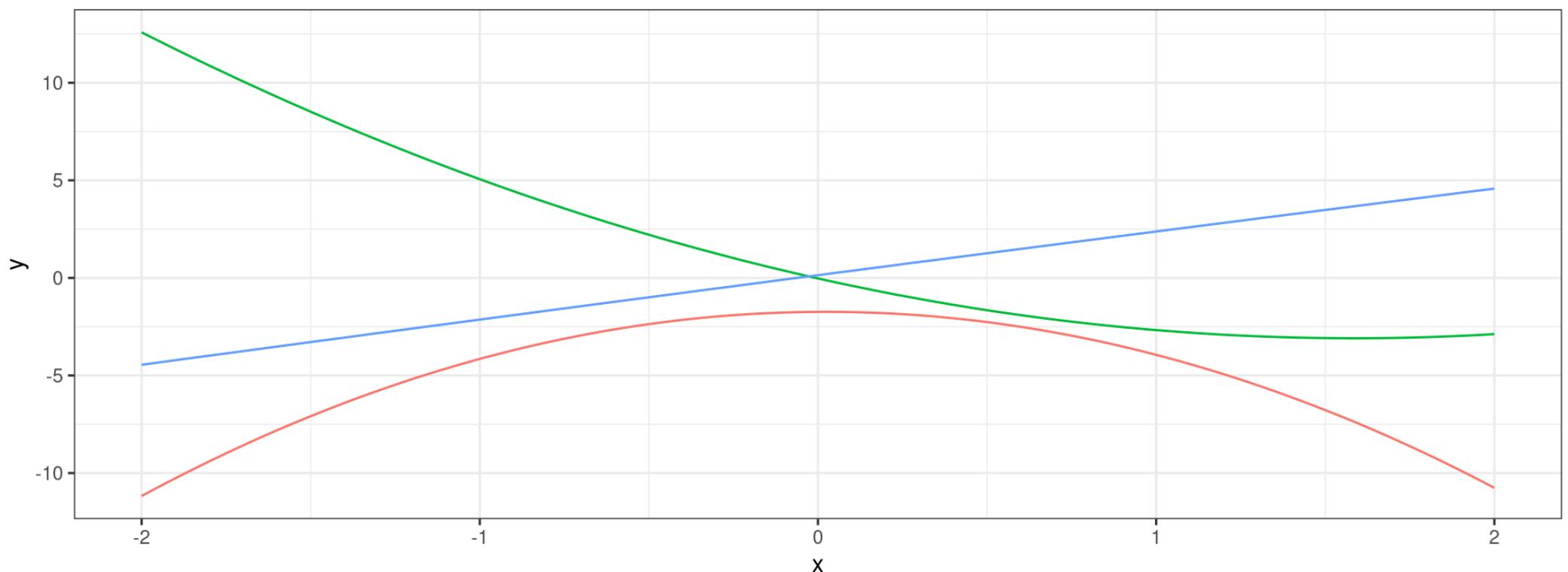
$$\text{Cov}_a(y_{t_i}, y_{t_j}) \times \text{Cov}_b(y_{t_i}, y_{t_j})$$

Linear \times Linear \rightarrow Quadratic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + 2(t_i \times t_j)$$

$$\text{Cov}_b(y_{t_i}, y_{t_j}) = 2 + 1(t_i \times t_j)$$

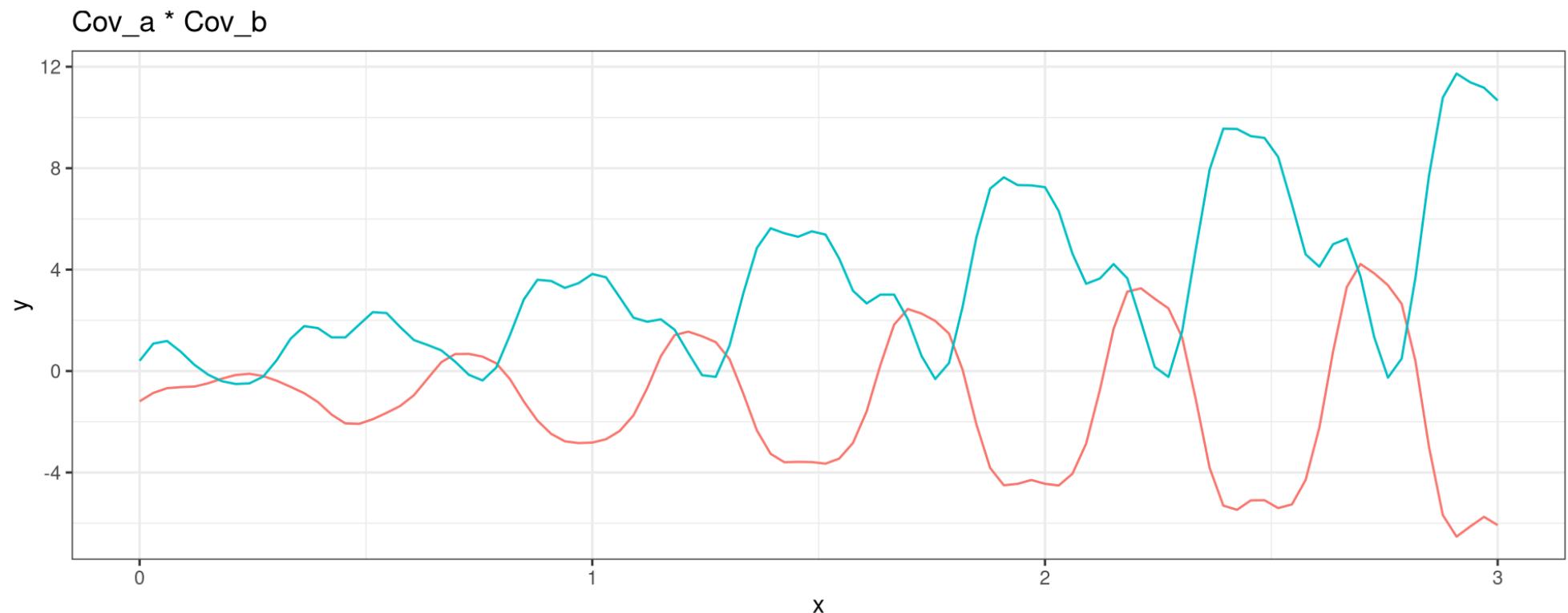
`Cov_a * Cov_b`



Linear × Periodic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + \frac{1}{2} (t_i \times t_j)$$

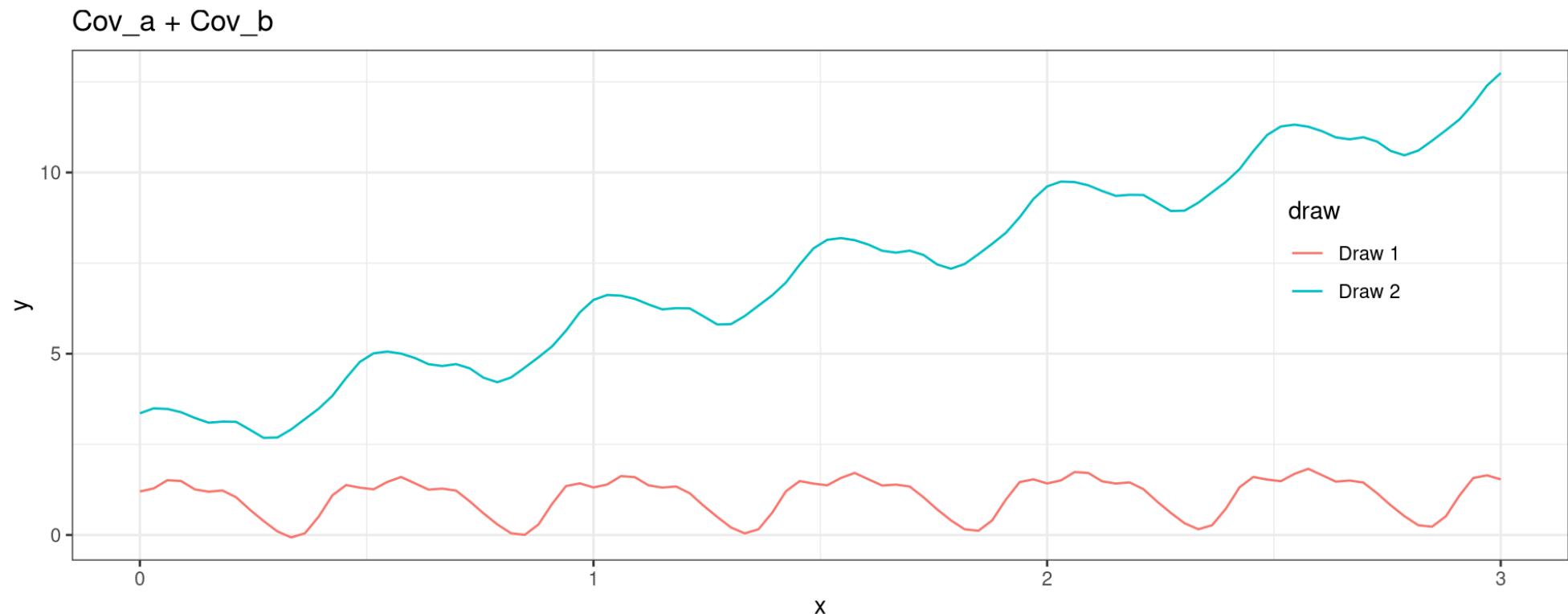
$$\text{Cov}_b(y_{t_i}, y_{t_j}) = \exp(-2 \sin^2(2\pi h))$$



Linear + Periodic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + 1(t_i \times t_j)$$

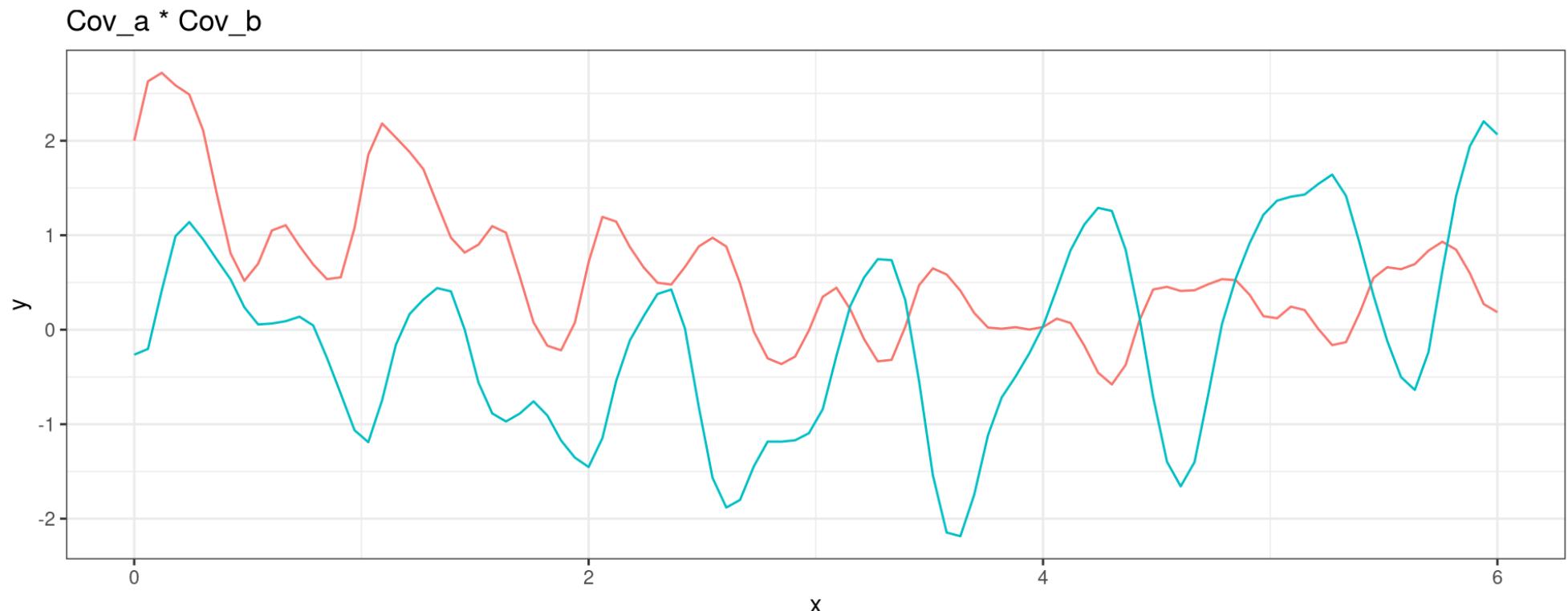
$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-2 \sin^2(2\pi h))$$



Sq Exp \times Periodic \rightarrow Locally Periodic

$$\text{Cov}_a(h = |t_i - t_j|) = \exp(-(1/3)h^2)$$

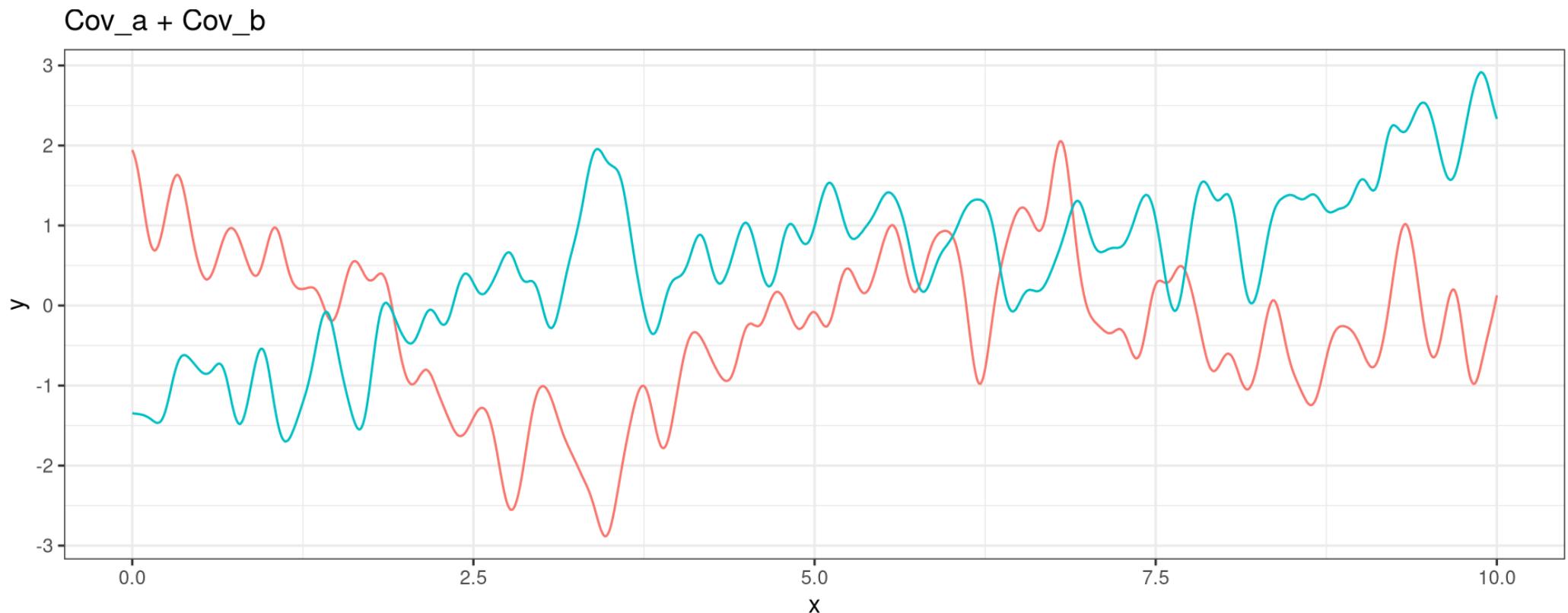
$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-2 \sin^2(\pi h))$$



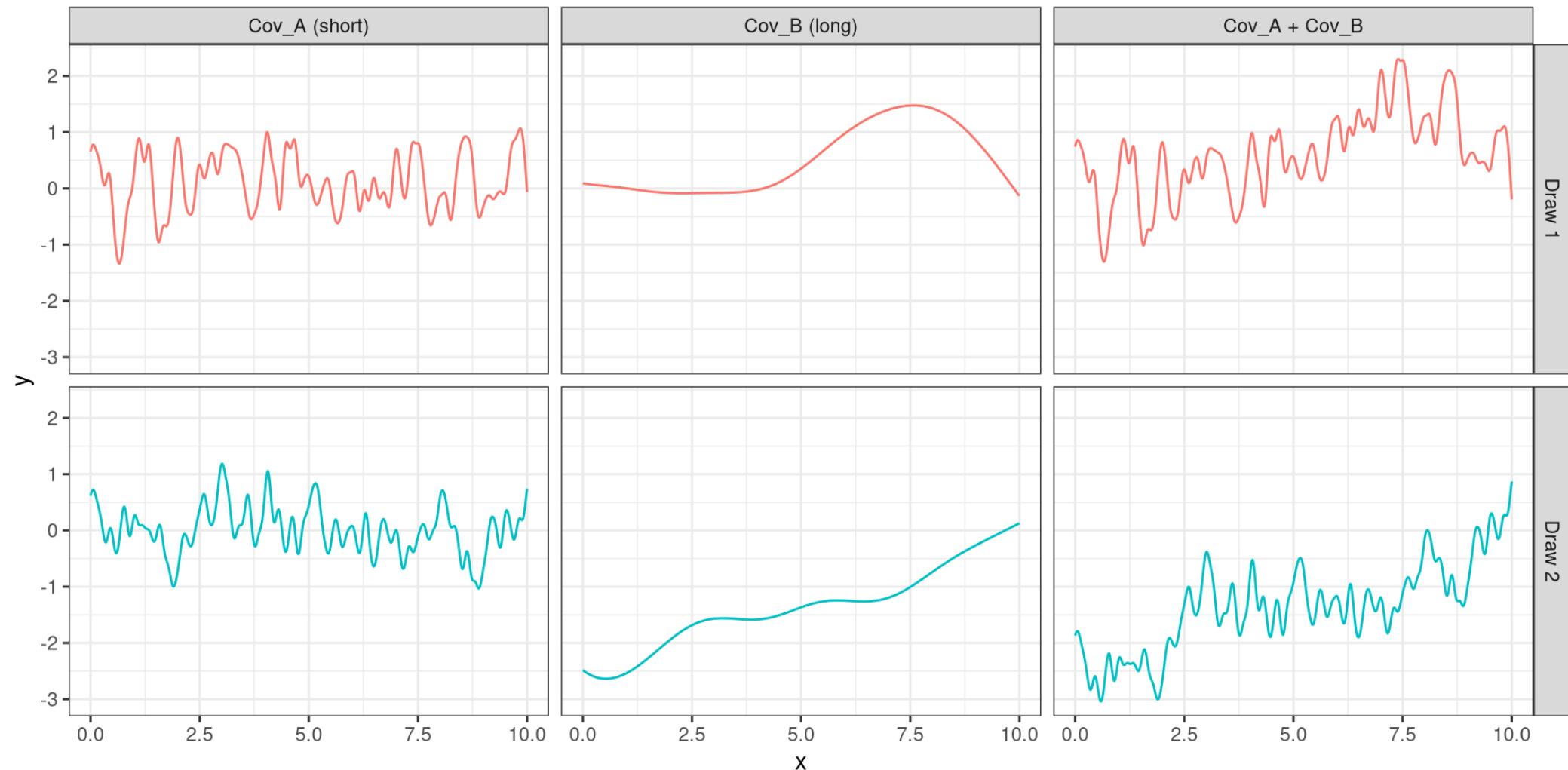
Sq Exp (short) + Sq Exp (long)

$$\text{Cov}_a(h = |t_i - t_j|) = (1/4) \exp(-4\sqrt{3}h^2)$$

$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-(\sqrt{3}/2)h^2)$$

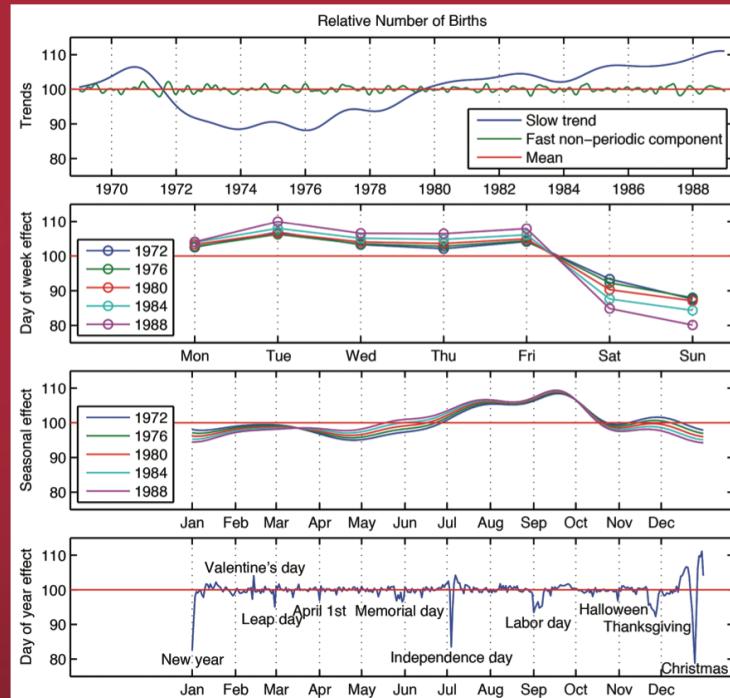


Seen another way



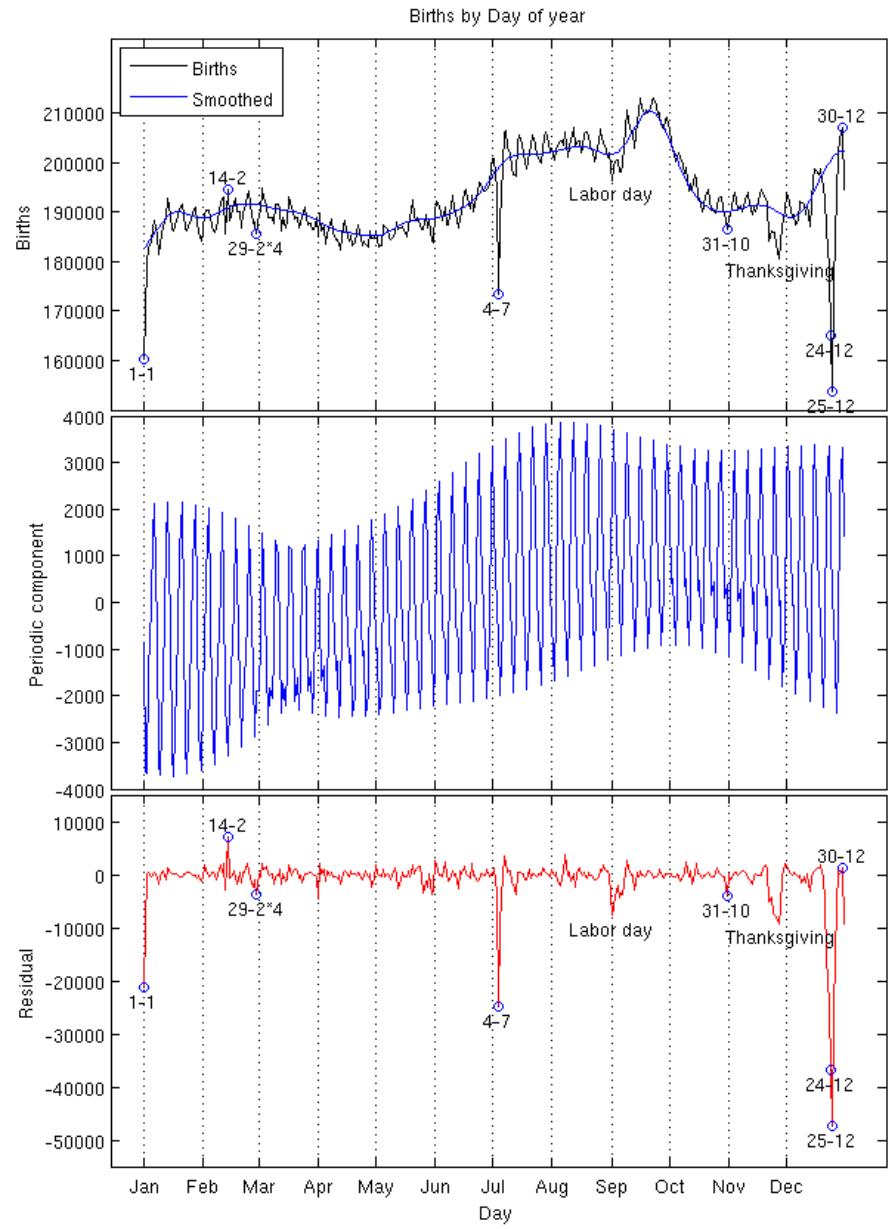
BDA3 example

Bayesian Data Analysis Third Edition



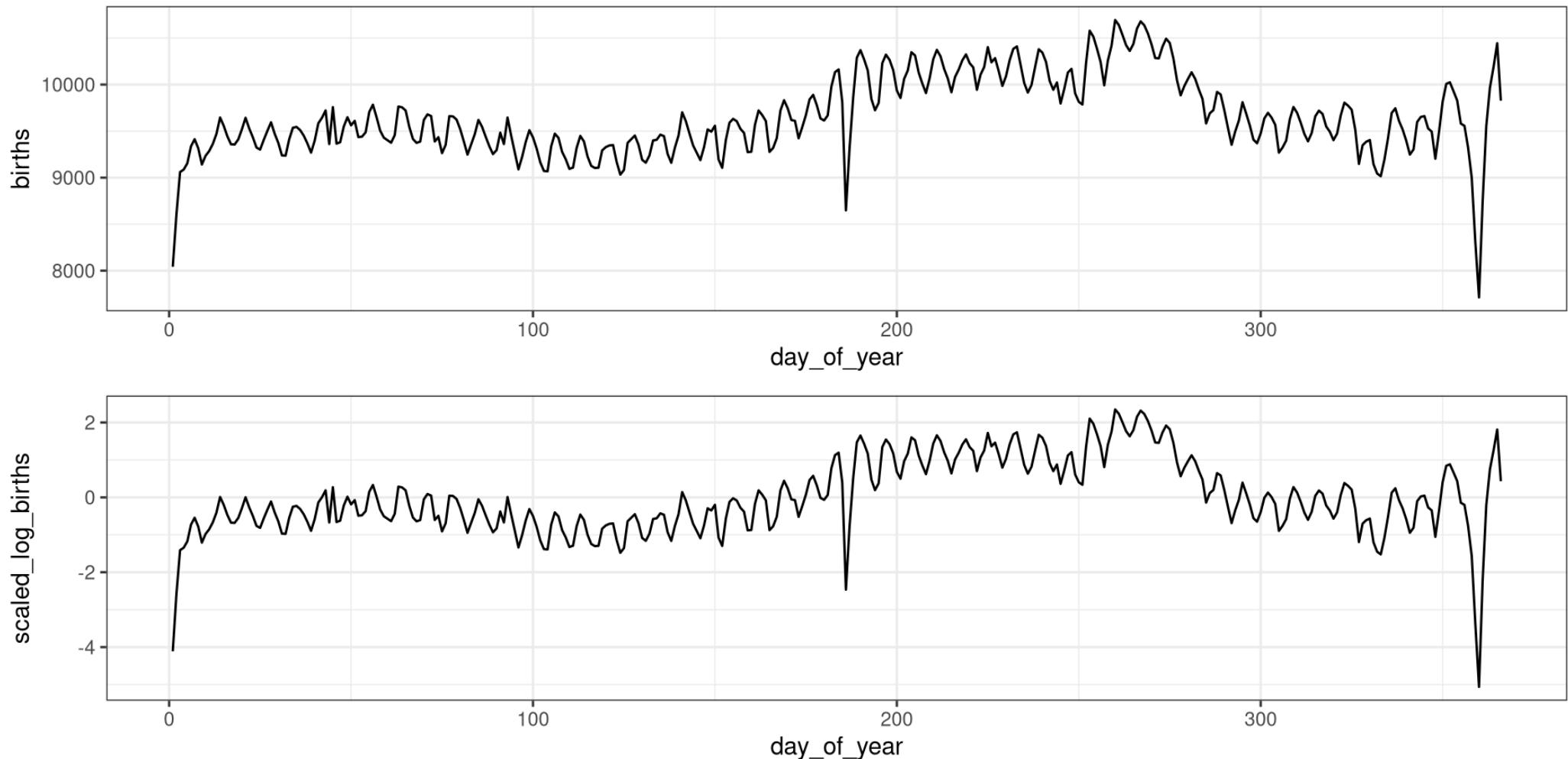
Andrew Gelman, John B. Carlin, Hal S. Stern,
David B. Dunson, Aki Vehtari, and Donald B. Rubin

Births (one year)



1. Smooth long term trend
(*sq exp cov*)
2. Seven day periodic trend with decay (*periodic x sq exp cov*)
3. Constant mean

Birth data



Model (JAGS)

```
1 model = "model{
2     y ~ dmnorm(rep(0,N), inverse(Sigma))
3
4     for (i in 1:(length(y)-1)) {
5         for (j in (i+1):length(y)) {
6             k1[i,j] <- sigma2[1] * exp(- pow(l[1] * d[i,j],2))
7             k2[i,j] <- sigma2[2] * exp(- pow(l[2] * d[i,j],2) - 2 * pow(l[3] * sin(pi*d[i,j] / per), 2
8
9             Sigma[i,j] <- k1[i,j] + k2[i,j]
10            Sigma[j,i] <- Sigma[i,j]
11        }
12    }
13
14    for (i in 1:length(y)) {
15        Sigma[i,i] <- sigma2[1] + sigma2[2] + sigma2[3]
16    }
17
18    for(i in 1:3){
19        sigma2[i] ~ dt(0, 2.5, 1) T(0,)
20        l[i] ~ dt(0, 2.5, 1) T(0,)
21    }
22}"
```

Model fitting

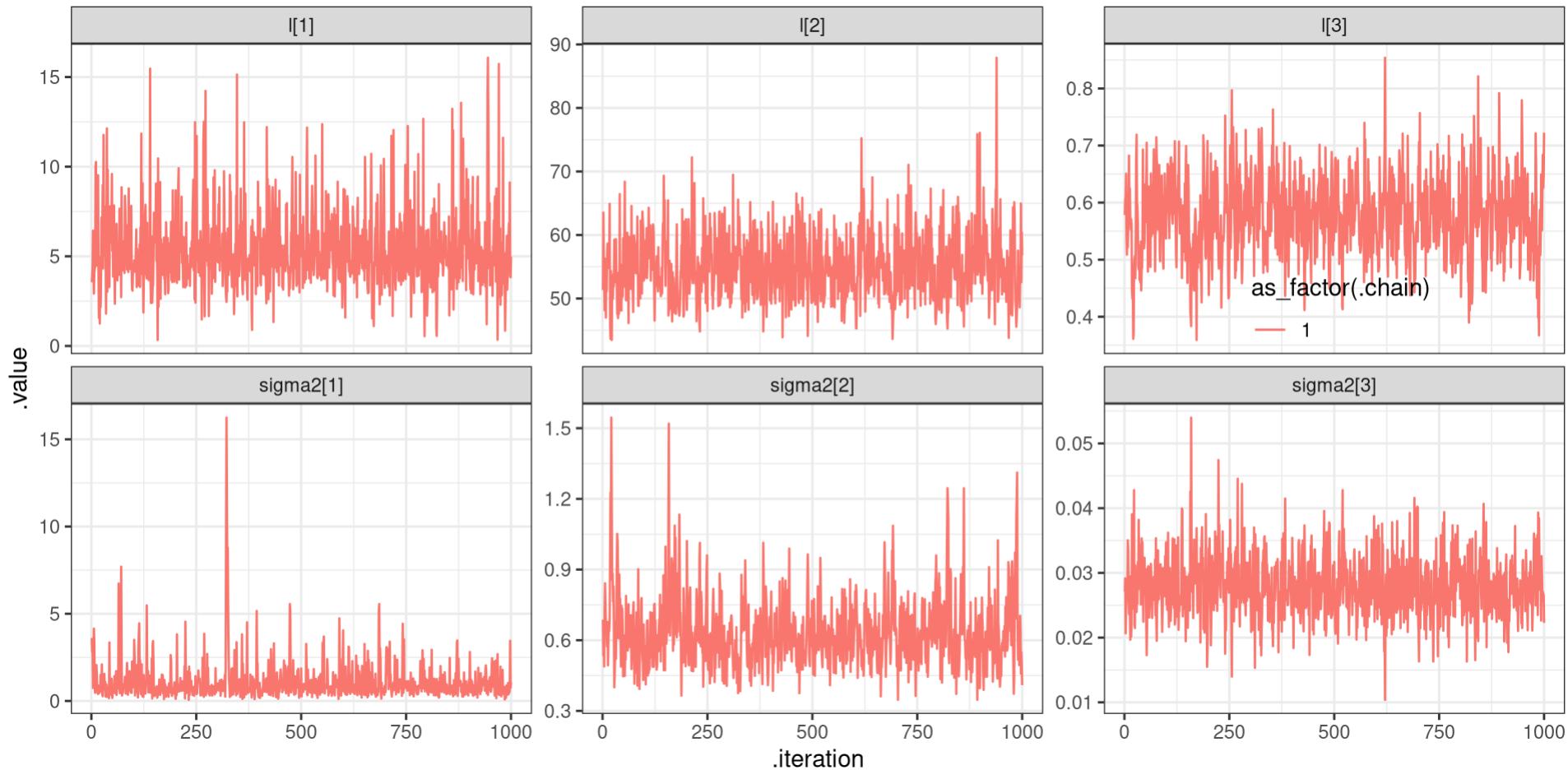
```
1 m = rjags::jags.model(  
2   textConnection(model),  
3   data = list(  
4     y = births$scaled_log_births,  
5     d = fields::rdist(births$day_of_year / ma  
6     per = 7 / max(births$day_of_year),  
7     pi = pi,  
8     N = nrow(births)  
9   ),  
10  n.adapt=5000,  
11  n.chains = 1  
12 )
```

```
1 gp_coda = rjags::coda.samples(  
2   m, variable.names=c("sigma2", "l"),  
3   n.iter=5000,  
4   thin=5  
5 )
```

Diagnostics

Traceplot

Density



Component Contributions

We can view our GP in the following ways (marginal form),

$$\mathbf{y} \sim N(\boldsymbol{\mu}, \Sigma_1 + \Sigma_2 + \sigma^2 \mathbf{I})$$

but with appropriate conditioning we can also think of \mathbf{y} as being the sum of multiple independent GPs (latent form)

$$\mathbf{y} = \boldsymbol{\mu} + \mathbf{w}_1(\mathbf{t}) + \mathbf{w}_2(\mathbf{t}) + \mathbf{w}_3(\mathbf{t})$$

where

$$\mathbf{w}_1(\mathbf{t}) \sim N(0, \Sigma_1) \quad (\text{sq exp covariance})$$

$$\mathbf{w}_2(\mathbf{t}) \sim N(0, \Sigma_2) \quad (\text{periodic x sq exp cov})$$

$$\mathbf{w}_3(\mathbf{t}) \sim N(0, \sigma^2 \mathbf{I}) \quad (\text{nugget cov / white noise})$$

Decomposition of Covariance Components

$$\begin{bmatrix} y \\ w_1 \\ w_2 \end{bmatrix} \sim N \left(\begin{bmatrix} \mu \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \Sigma_1 + \Sigma_2 + \sigma^2 I & \Sigma_1 & \Sigma_2 \\ \Sigma_1 & \Sigma_1 & 0 \\ \Sigma_2 & 0 & \Sigma_2 \end{bmatrix} \right)$$

therefore, if we want to know the contribution of w_1 we have the following

$$w_1 | y, \mu, \theta \sim N(\mu_{\text{cond}}, \Sigma_{\text{cond}})$$

$$\mu_{\text{cond}} = 0 + \Sigma_1 (\Sigma_1 + \Sigma_2 + \sigma^2 I)^{-1} (y - \mu)$$

$$\Sigma_{\text{cond}} = \Sigma_1 - \Sigma_1 (\Sigma_1 + \Sigma_2 + \sigma^2 I)^{-1} \Sigma_1^t$$

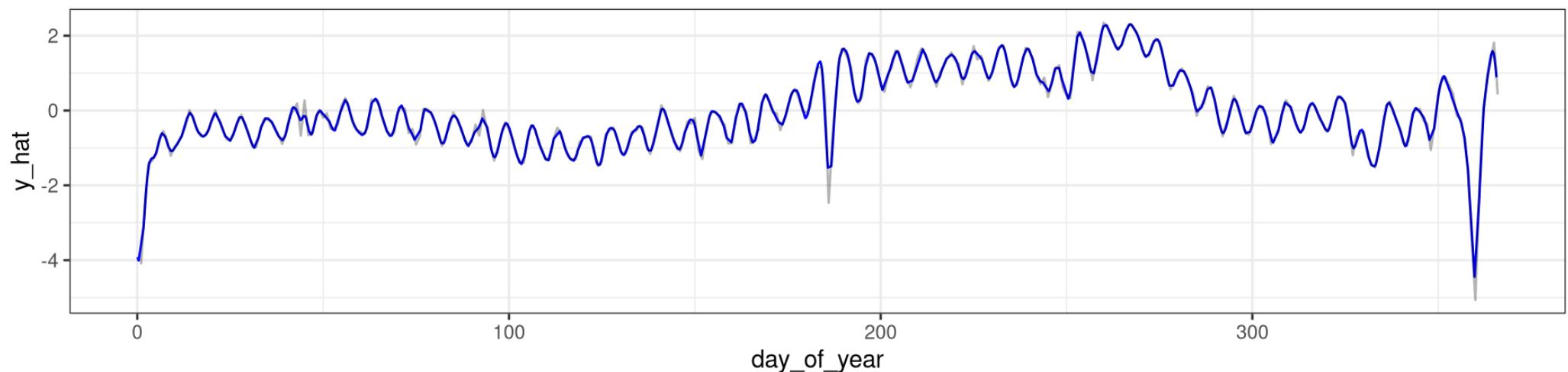
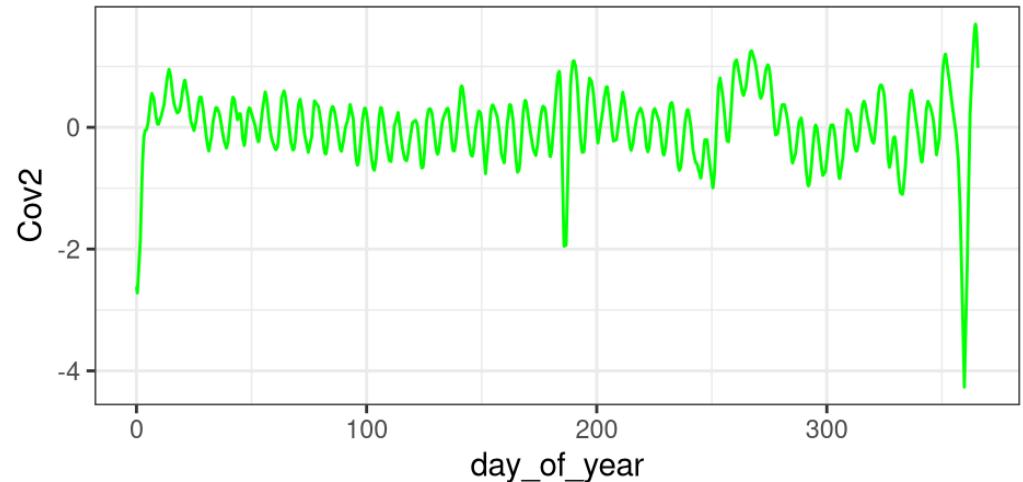
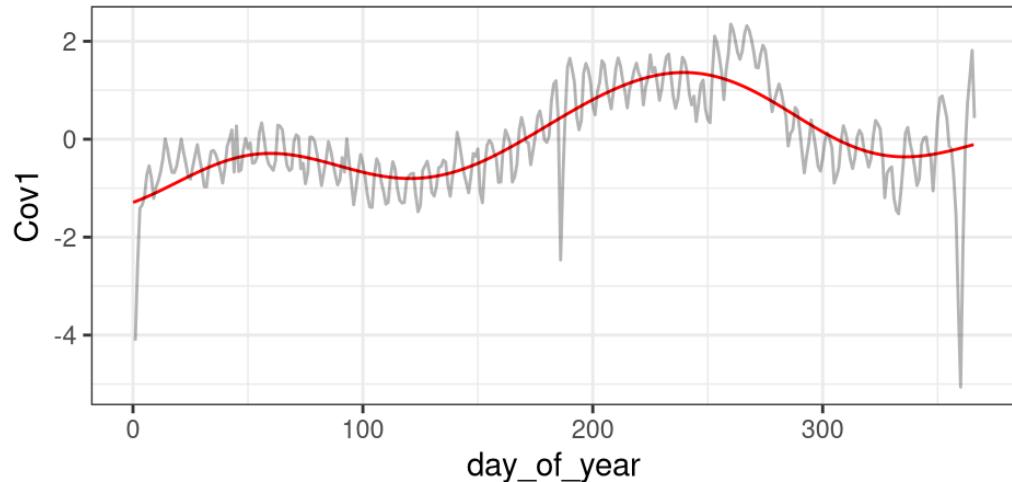
Covariance calculations

```
1 cov_C1 = function(d, sigma2, l, per) {  
2   sigma2[1] * exp(- (l[1] * d)^2)  
3 }  
4  
5 cov_C2 = function(d, sigma2, l, per) {  
6   sigma2[2] * exp(- (l[2] * d)^2 - 2 * (l[3] * sin(pi*d / per))^2)  
7 }  
8  
9 cov_full = function(d, sigma2, l, per) {  
10   cov_C1(d, sigma2, l, per) +  
11   cov_C2(d, sigma2, l, per) +  
12   ifelse(abs(d)<1e-6, sigma2[3] + 1e-6, 0)  
13 }
```

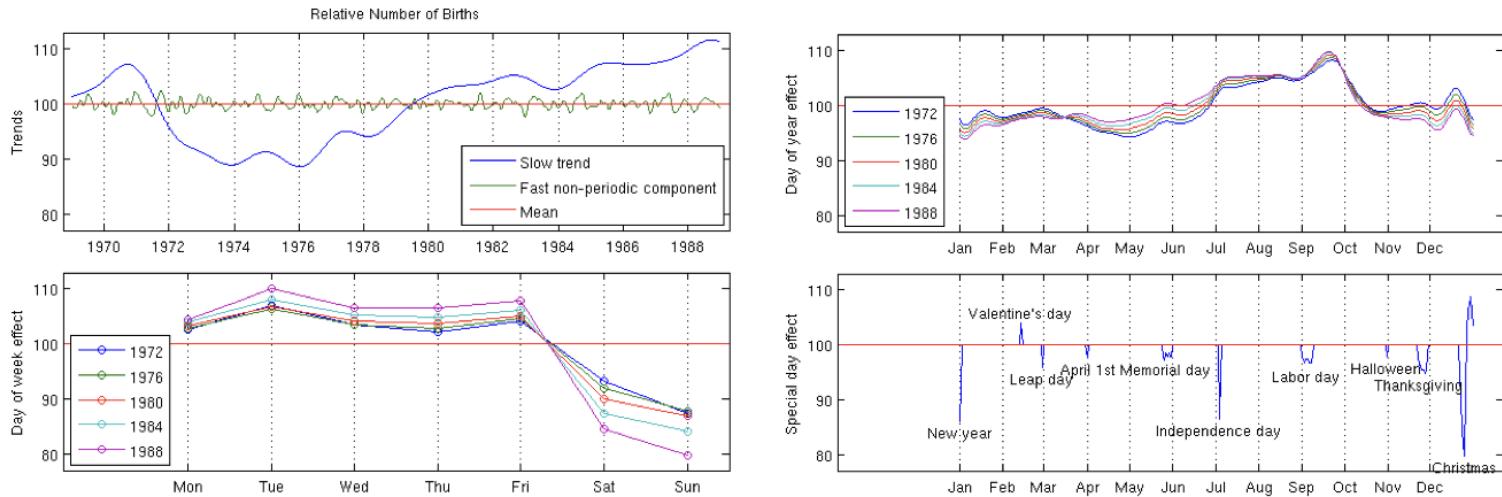
Conditional samples

```
1 full = dukestm::cond_predict(  
2   y=y, x=x, x_pred=x_pred, cov_full, sigma2 = post$sigma2, l=post$l, per=7 / max(births$day_of_year),  
3 )  
4  
5 comp_C1 = dukestm::cond_predict(  
6   y=y, x=x, x_pred=x_pred, sigma2 = post$sigma2, l=post$l, per=7 / max(births$day_of_year),  
7   cov_f_o = cov_full,  
8   cov_f_p = cov_C1,  
9   cov_f_po = cov_C1  
10 )  
11  
12 comp_C2 = dukestm::cond_predict(  
13   y=y, x=x, x_pred=x_pred, sigma2 = post$sigma2, l=post$l, per=7 / max(births$day_of_year),  
14   cov_f_o = cov_full,  
15   cov_f_p = cov_C2,  
16   cov_f_po = cov_C2  
17 )
```

Results



Births (multiple years)

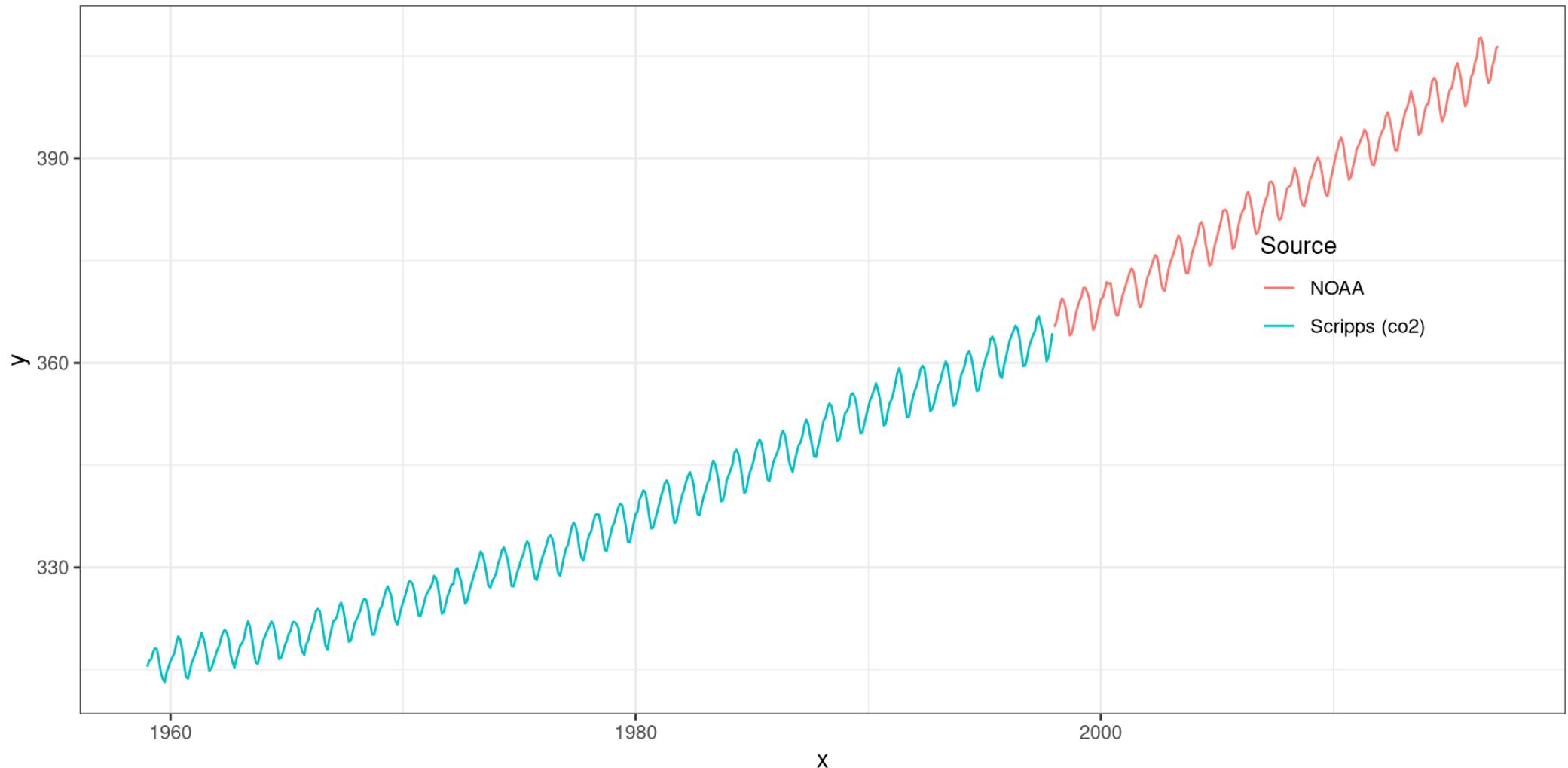


Full stan case study [here](#) with code [here](#)

1. slowly changing trend - yearly (*sq exp cov*)
2. small time scale trend - monthly (*sq exp cov*)
3. 7 day periodic - day of week effect (*periodic* \times *sq exp cov*)
4. 365.25 day periodic - day of year effect (*periodic* \times *sq exp cov*)
5. special days and interaction with weekends (*linear cov*)
6. independent Gaussian noise (*nugget cov*)
7. constant mean

Mauna Loa Example

Atmospheric CO₂



GP Model

Based on Rasmussen 5.4.3 (we are using slightly different data and parameterization)

$$\mathbf{y} \sim \square(\boldsymbol{\mu}, \Sigma_1 + \Sigma_2 + \Sigma_3 + \Sigma_4 + \sigma^2 I)$$

$$\{\boldsymbol{\mu}\}_i = y$$

$$\{\Sigma_1\}_{ij} = \sigma_1^2 \exp(-(l_1 \cdot d_{ij})^2)$$

smooth long term trend

$$\{\Sigma_2\}_{ij} = \sigma_2^2 \exp(-(l_2 \cdot d_{ij})^2) \exp(-2(l_3)^2 \sin^2(\pi d_{ij}/p))$$

seasonal trend w/ decay

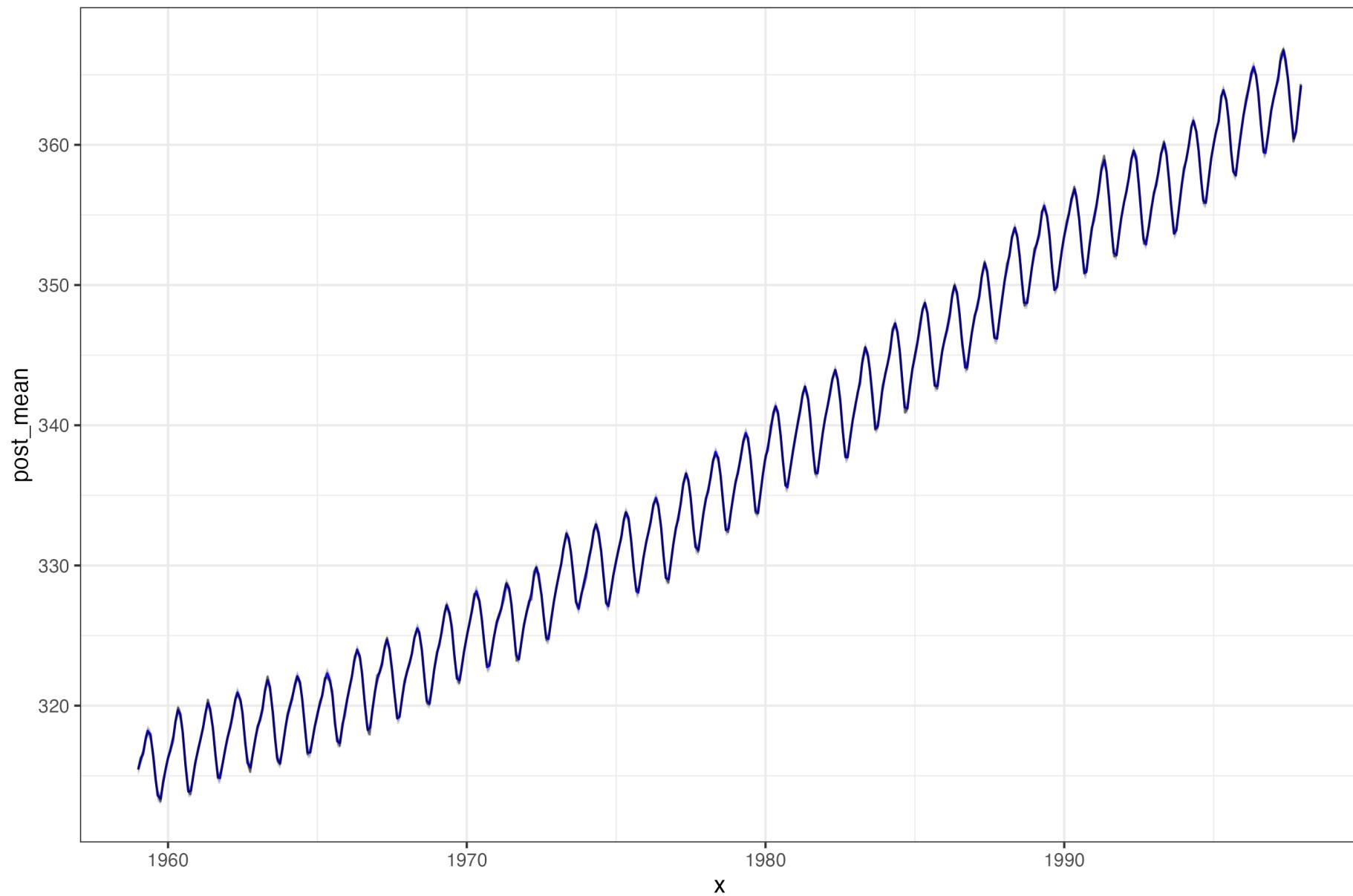
$$\{\Sigma_3\}_{ij} = \sigma_3^2 \left(1 + \frac{(l_4 \cdot d_{ij})^2}{\alpha}\right)^{-\alpha}$$

small / medium term trend

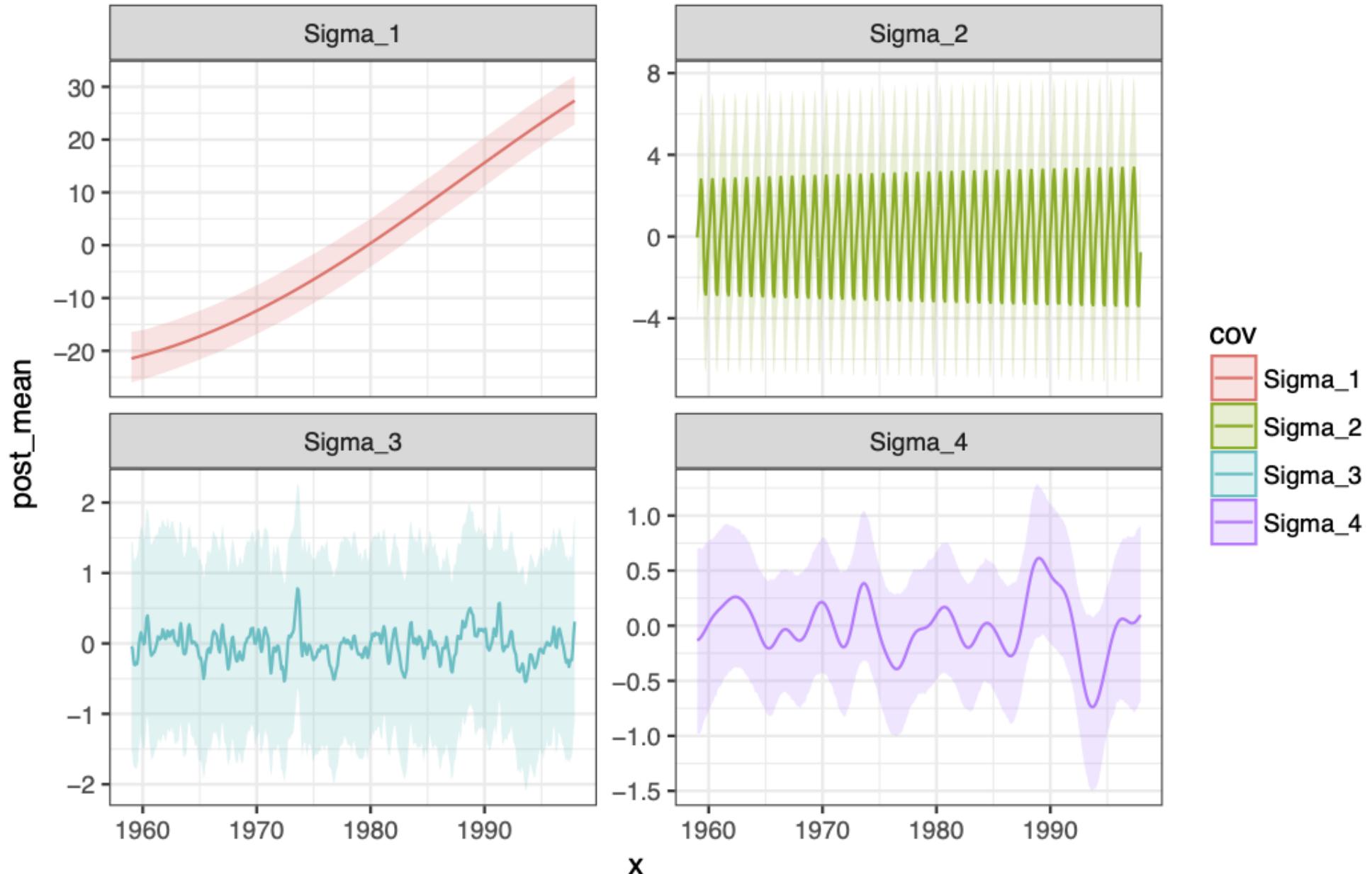
$$\{\Sigma_4\}_{ij} = \sigma_4^2 \exp(-(l_5 \cdot d_{ij})^2)$$

noise

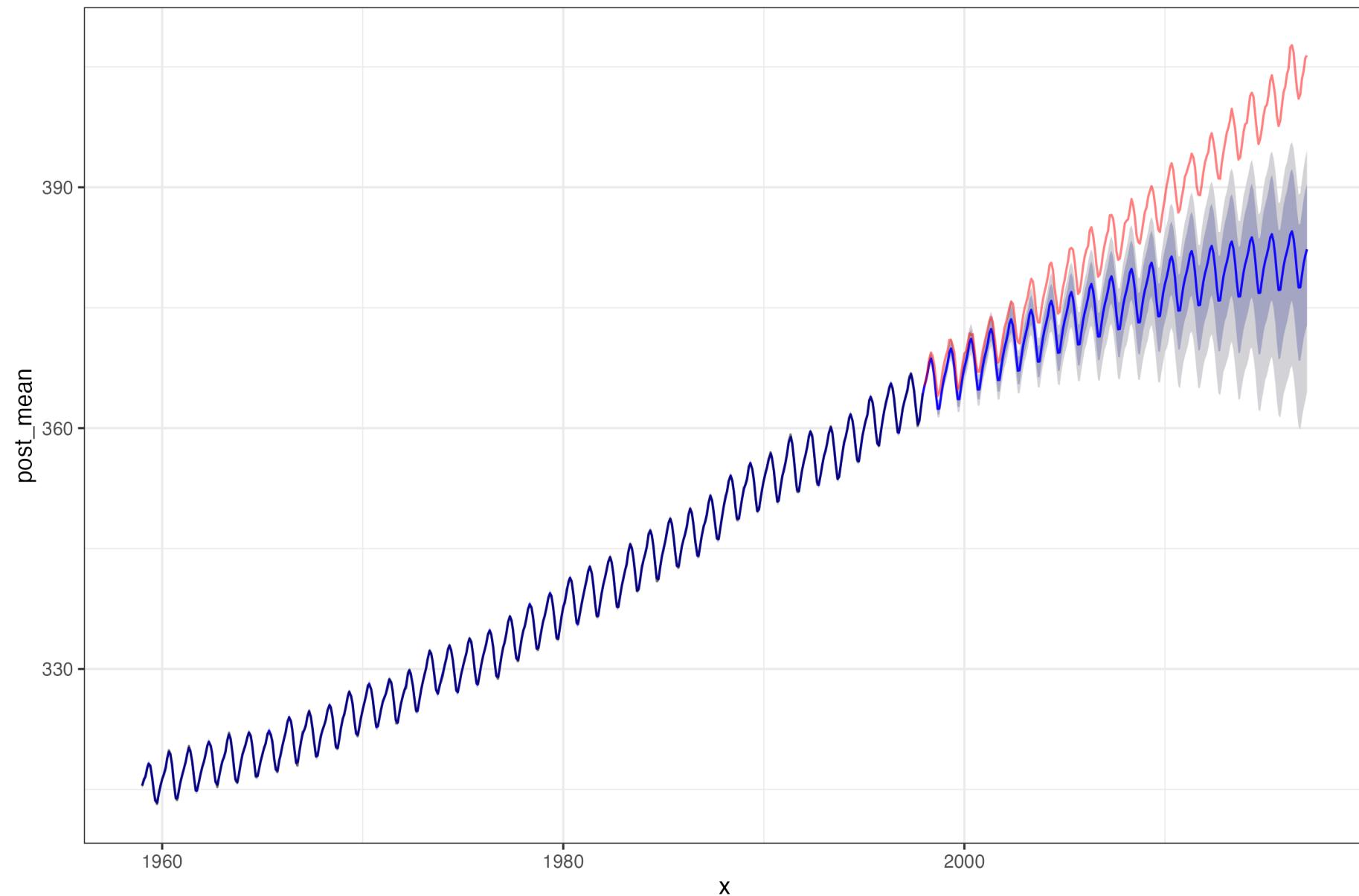
Model fit



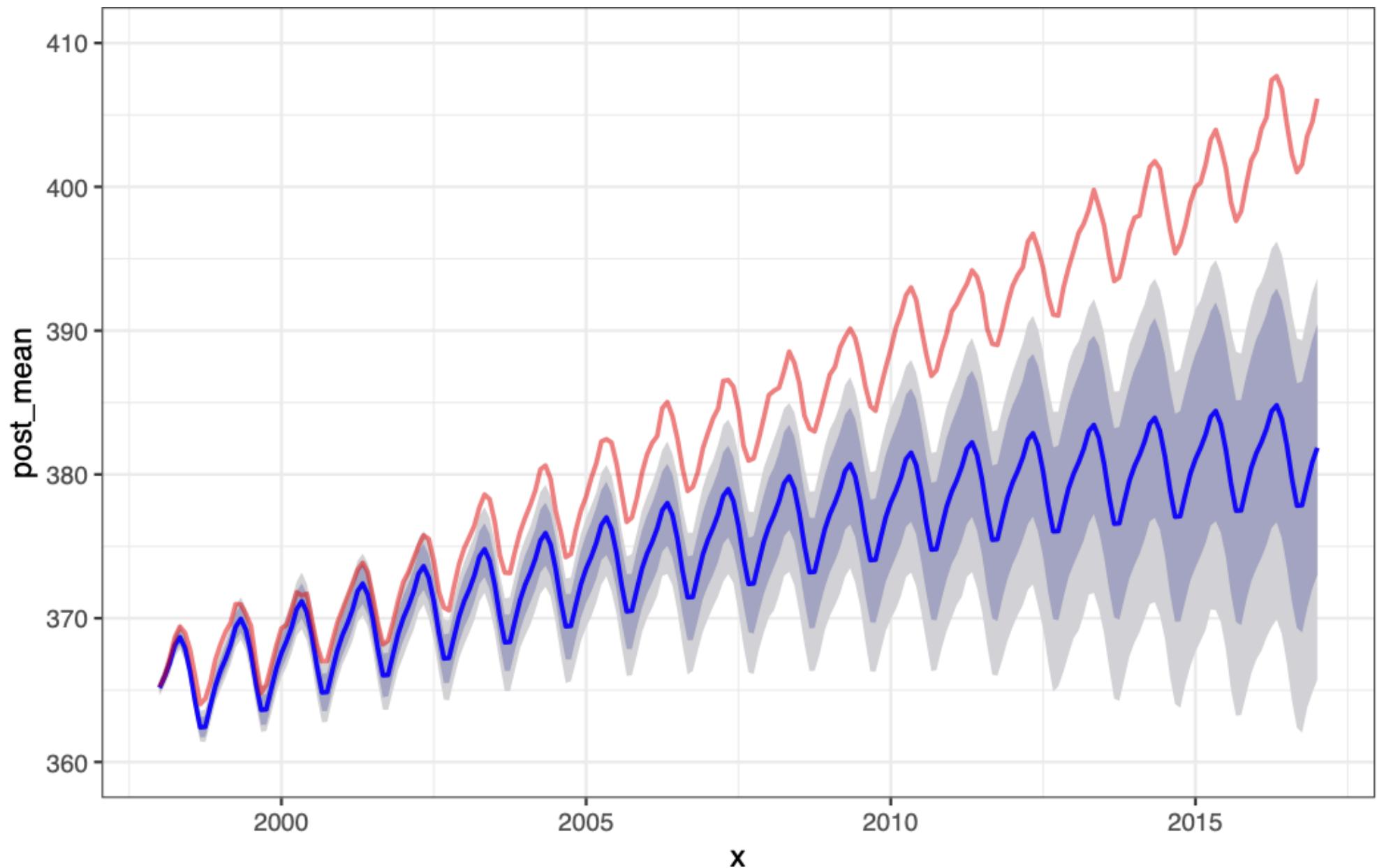
Fit Components



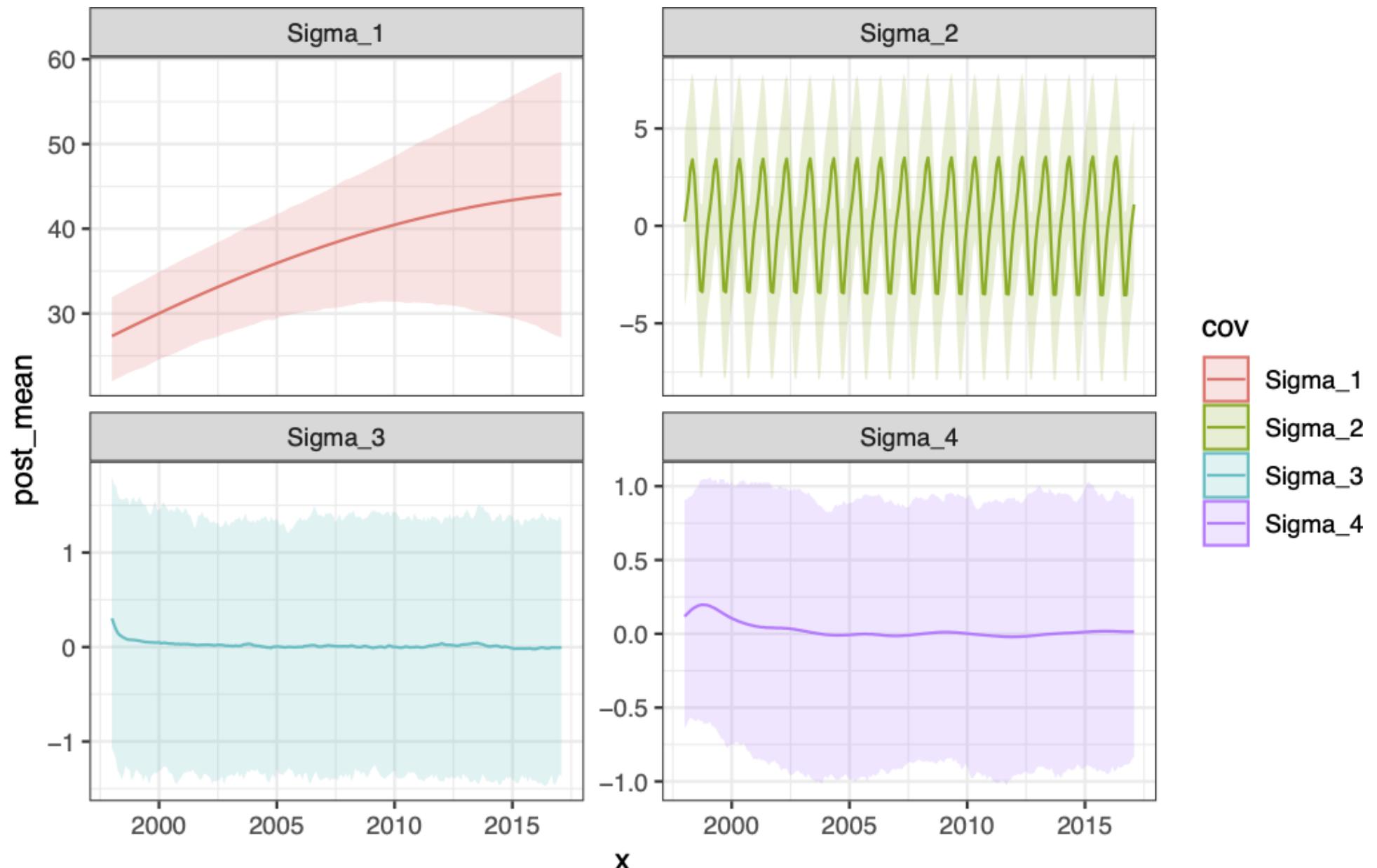
Model fit + forecast



Forecast

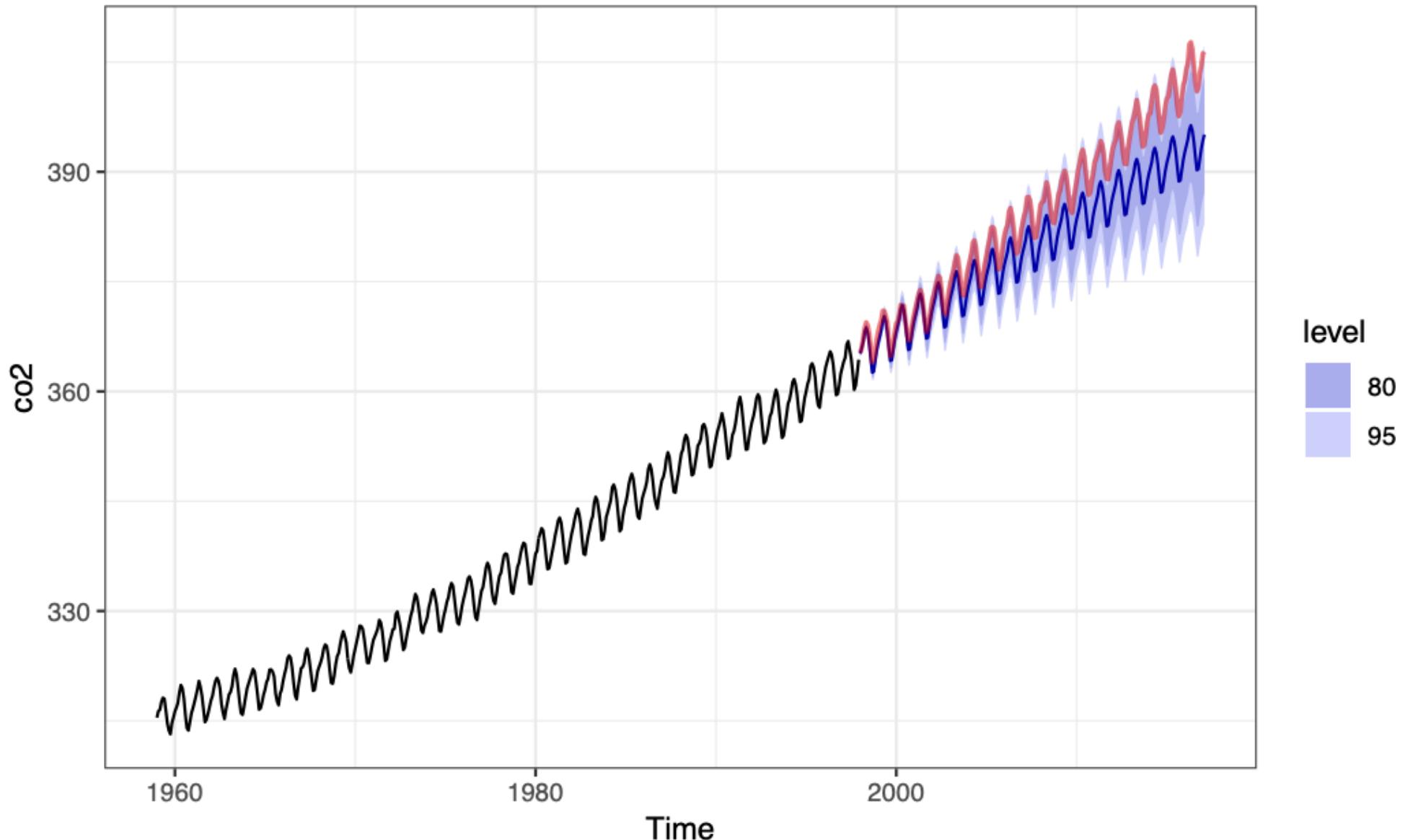


Forecast components



ARIMA forecast

Forecasts from ARIMA(1,1,1)(1,1,2)[12]



Model performance

Forecast dates	arima	gp
	RMSE	RMSE
Jan 1998 - Jan 2003	1.10	1.91
Jan 1998 - Jan 2008	2.51	4.58
Jan 1998 - Jan 2013	3.82	7.71
Jan 1998 - Mar 2017	5.46	11.40

