Gaussian Process Models Part 2

Lecture 15

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EDA and GPs

Variogram

When fitting a Gaussian process model, it is often difficult to fit the covariance parameters (hard to identify). Today we will discuss some EDA approaches for getting a sense of the values for the scale, range and nugget parameters.

From the spatial modeling literature the typical approach is to examine an *empirical variogram*, first we will define the *theoretical variogram* and its connection to the covariance.

Variogram:

$$2\gamma(t_i, t_j) = Var(y(t_i) - y(t_j))$$

where $\gamma(t_i, t_j)$ is known as the semivariogram.

Properties of the Variogram / Semivariogram

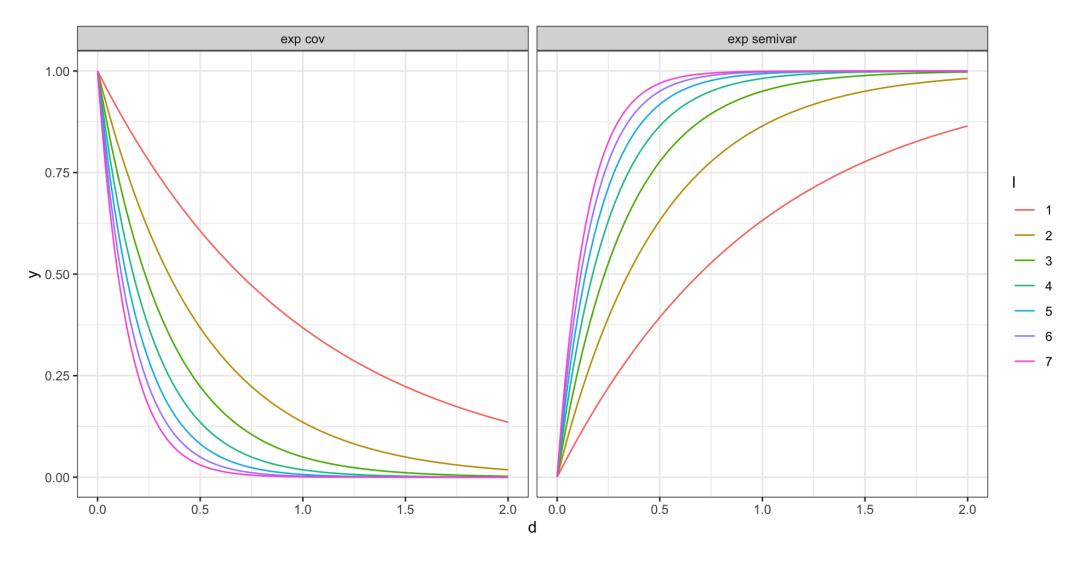
- are non-negative $\gamma(t_i, t_i) \ge 0$
- are equal to 0 at distance 0 $\gamma(t_i, t_i) = 0$
- are symmetric $\gamma(t_i, t_j) = \gamma(t_j, t_i)$
- if observations are independent $2\gamma(t_i, t_j) = Var(y(t_i)) + Var(y(t_j)) \quad \text{for all } i \neq j$
- if the process *is not* stationary $2\gamma(t_i, t_j) = Var(y(t_i)) + Var(y(t_j)) 2Cov(y(t_i), y(t_j))$
- if the process is stationary $2\gamma(t_i,t_j) = 2 \operatorname{Var}(y(t_i)) 2 \operatorname{Cov}(y(t_i),y(t_j))$

Connection to Covariance

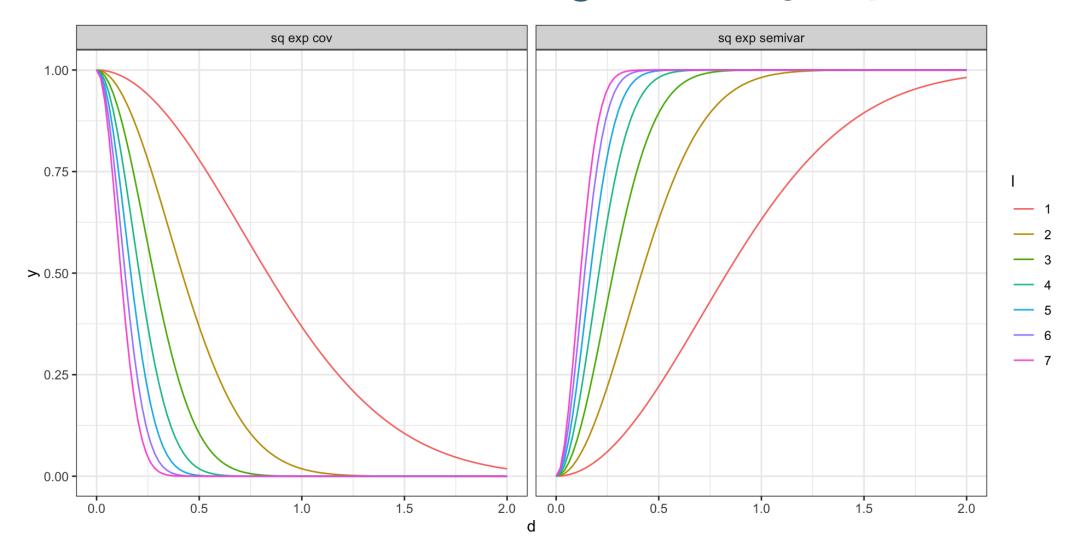
Assuming a squared exponential covariance structure,

$$\begin{aligned} 2\gamma(t_i, t_j) &= 2V \operatorname{ar} \left(y(t_i) \right) - 2\operatorname{Cov} \left(y(t_i), y(t_j) \right) \\ \gamma(t_i, t_j) &= V \operatorname{ar} \left(y(t_i) \right) - \operatorname{Cov} \left(y(t_i), y(t_j) \right) \\ &= \sigma^2 - \sigma^2 \exp \left(- \left(|t_i - t_j| \, 1 \right)^2 \right) \end{aligned}$$

Covariance vs Semivariogram - Exponential



Covariance vs Semivariogram - Sq. Exp.



Nugget variance

Very often in the real world we will observe that $\gamma(t_i,t_i)=0$ is not true - there will be an initial discontinuity in the semivariogram at $|t_i-t_j|=0$. Why is this?

We can think about Gaussian process regression in the following way,

$$y(t) = \mu(t) + w(t) + \epsilon(t)$$

where

$$\mu(t) = X\beta$$

$$w(t) \sim N(0, \Sigma)$$

$$\epsilon(t) \stackrel{\text{iid}}{\sim} N(0, \sigma_{w}^{2})$$

Implications

With the inclusion of the $\epsilon(t)$ terms in the model we now have,

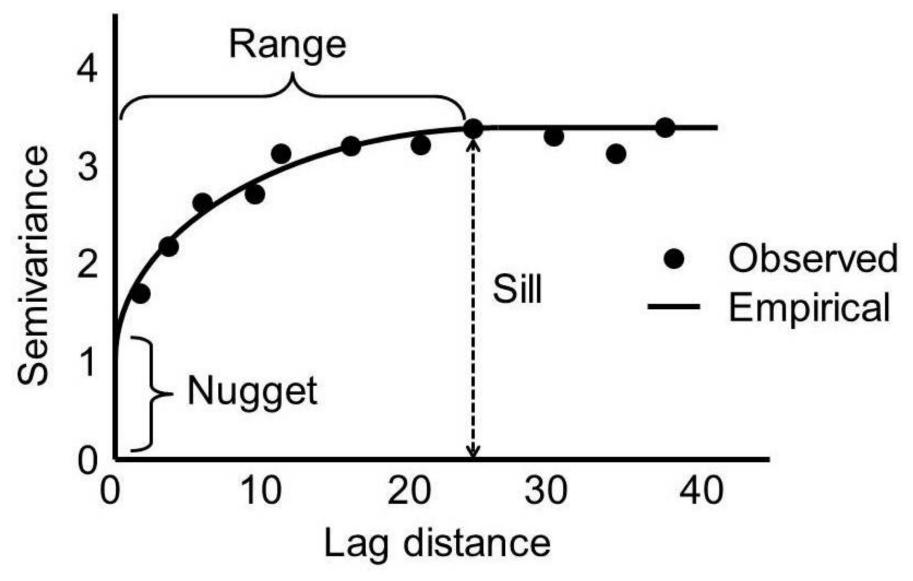
$$Var(y(t_i)) = \sigma_w^2 + \Sigma_{ii}$$

$$Cov(y(t_i), y(t_j)) = \Sigma_{ij}$$

Therefore, for a squared exponential covariance model with a nugget component the semivariogram is given by,

$$\gamma(t_i, t_j) = (\sigma^2 + \sigma_w^2) - \sigma^2 \exp(-(|t_i - t_j| 1)^2)$$

Semivariogram features



Empirical Semivariogram

We will assume that our process of interest is stationary, in which case we will parameterize the semivariagram in terms of $d = |t_i - t_i|$.

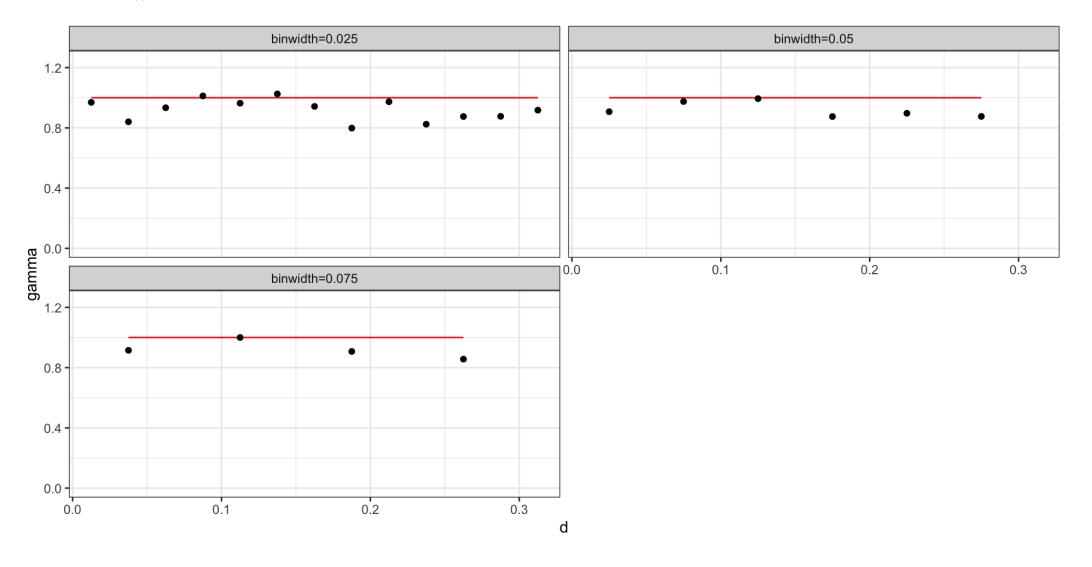
Empirical Semivariogram:

$$\hat{\gamma(d)} = \frac{1}{2 N(d)} \sum_{\substack{|t_i - t_j| \in (d - \epsilon, d + \epsilon)}} (y(t_i) - y(t_j))^2$$

Practically, for any data set with n observations there are $\binom{n}{2} + n$ possible data pairs to examine. Each individually is not very informative, so we aggregate into bins and calculate the empirical semivariogram for each bin.

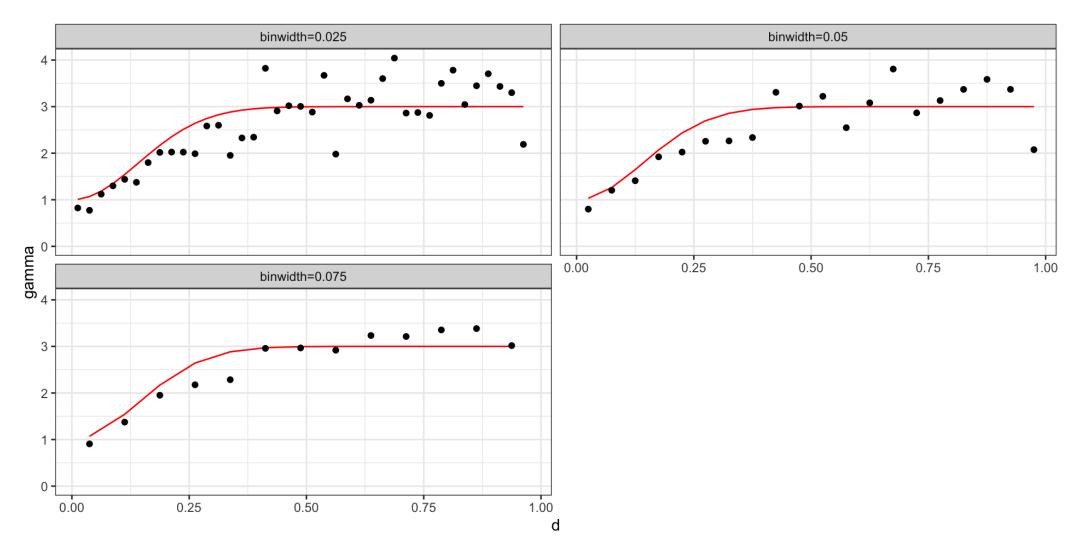
Empirical semivariogram of WN

Where $\sigma_w^2 = 1$,



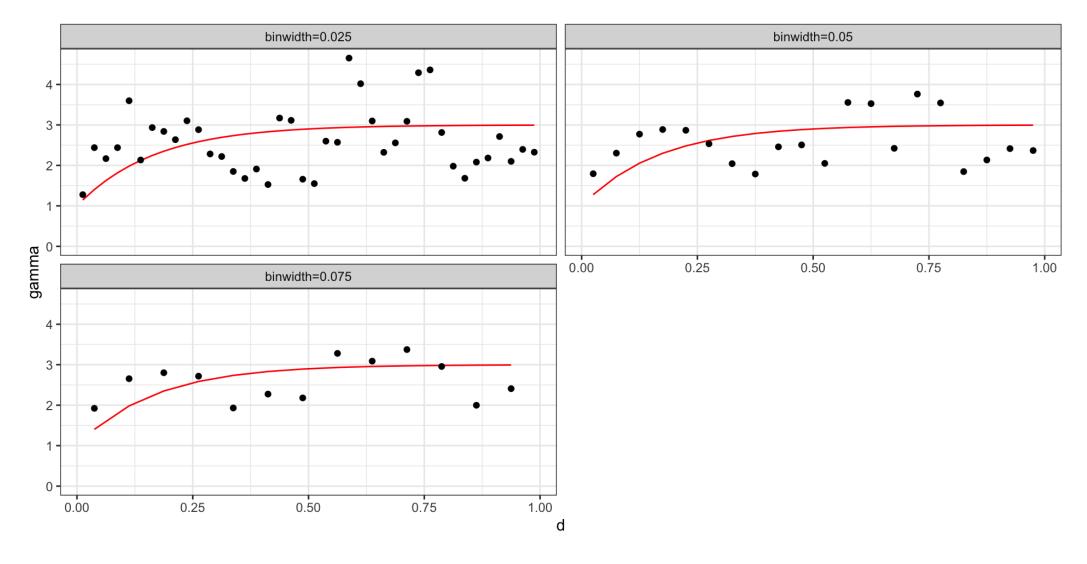
Empirical Variogram of GP w/ Sq Exp

Where $\sigma^2 = 2$, 1 = 5, and $\sigma_w^2 = 1$,

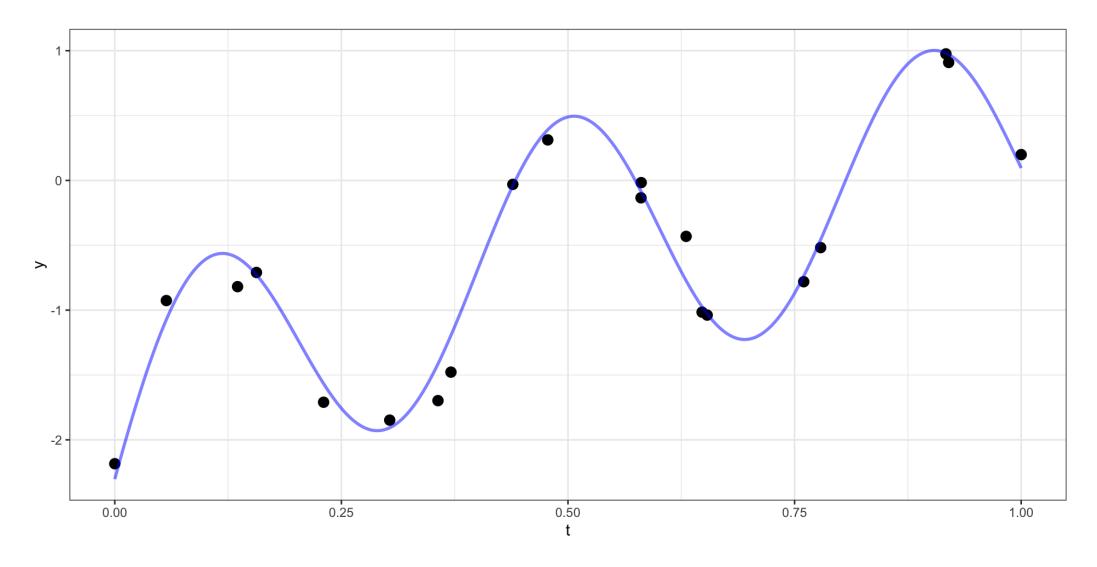


Empirical Variogram of GP w/ Exp

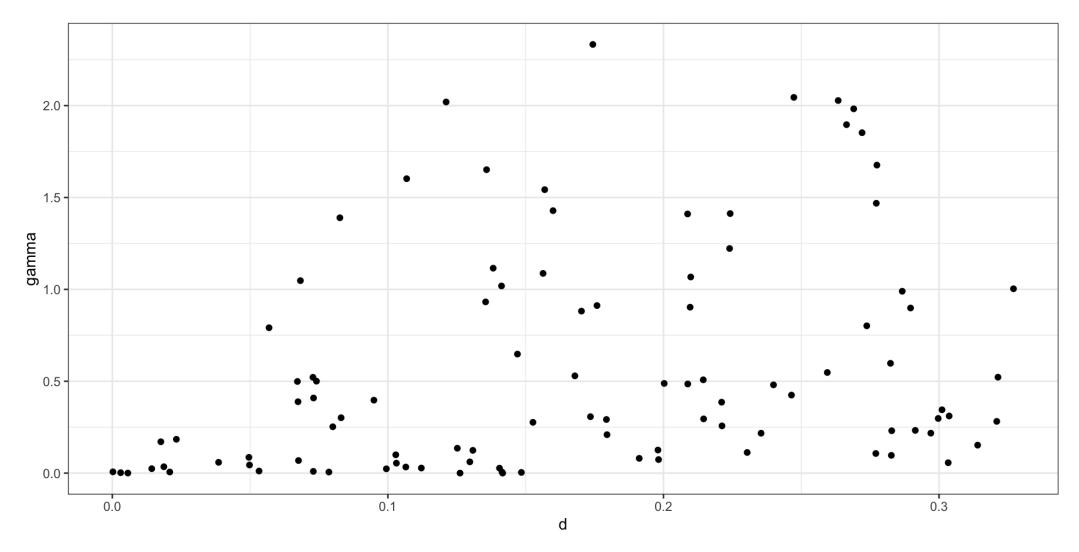
Where $\sigma^2 = 2$, 1 = 6, and $\sigma_w^2 = 1$,



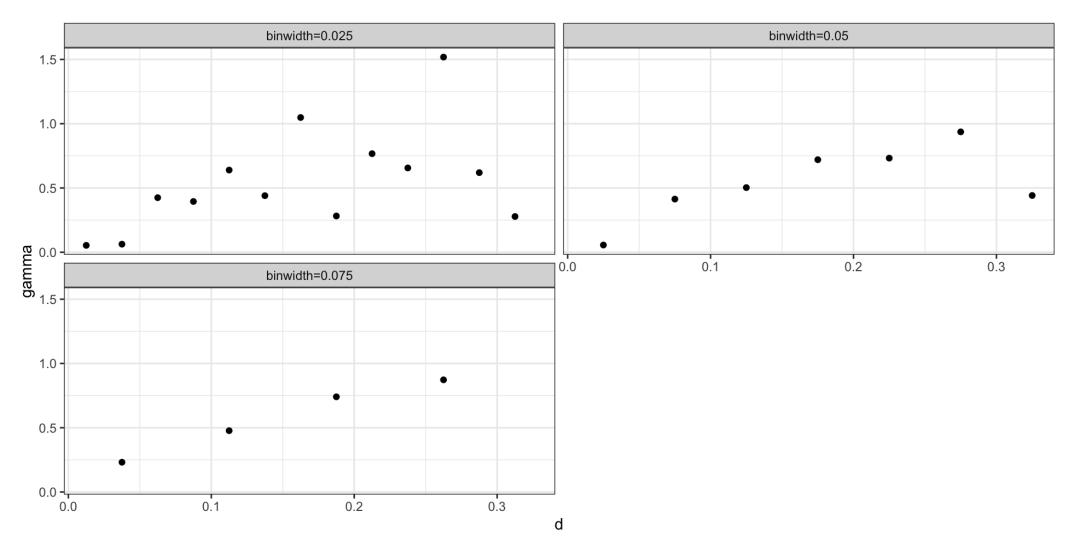
From last time



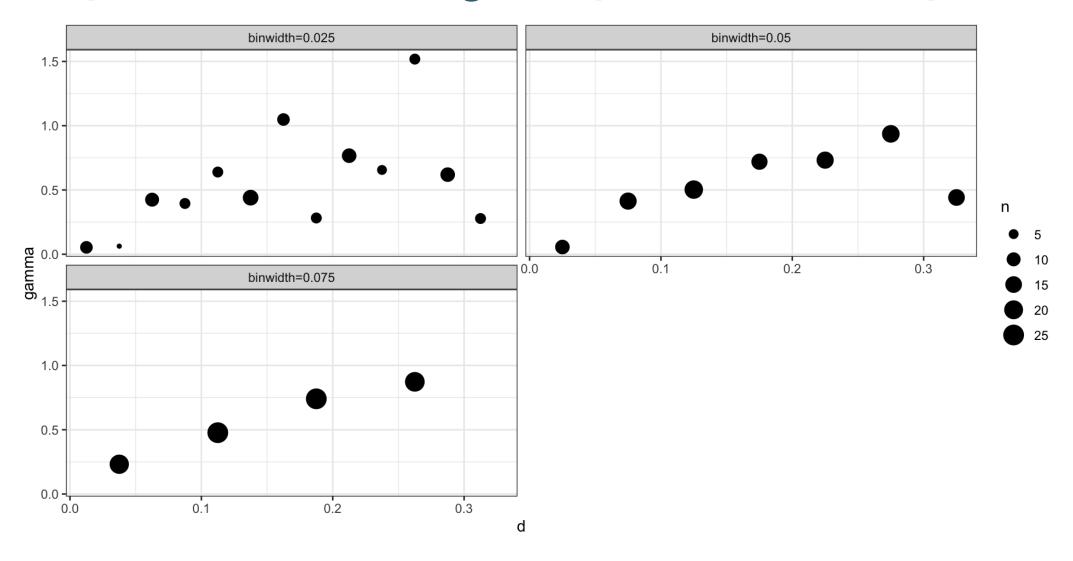
Empirical semivariogram - no bins / cloud



Empirical semivariogram (binned)



Empirical semivariogram (binned w/ size)

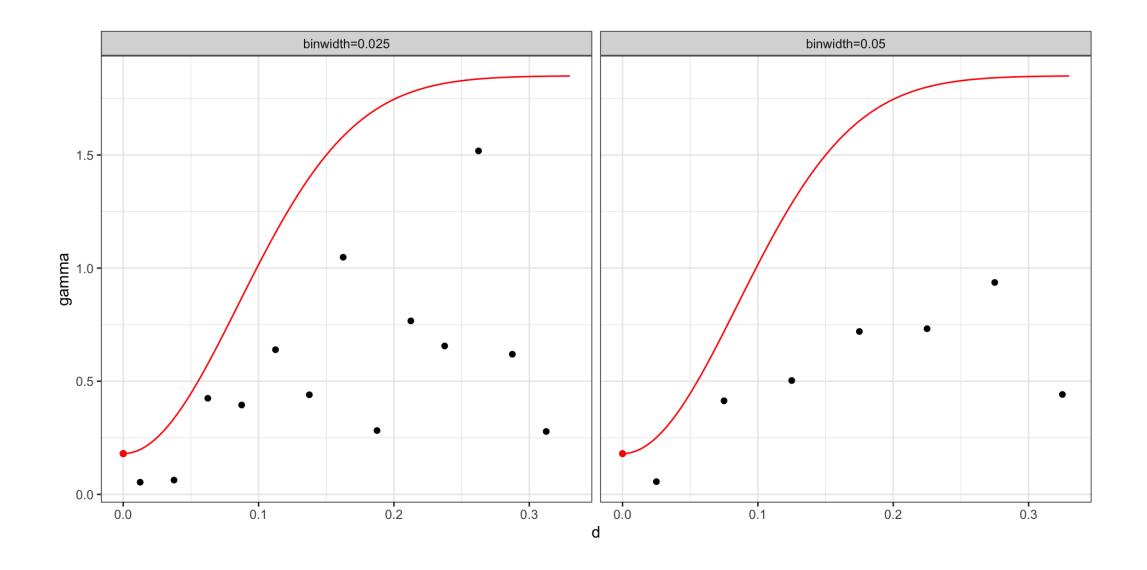


Theoretical vs empirical semivariogram

After fitting the model last time we came up with a posterior mean of $\sigma^2 = 1.67$, 1 = 8.33, and $\sigma_w^2 = 0.18$ for a square exponential covariance.

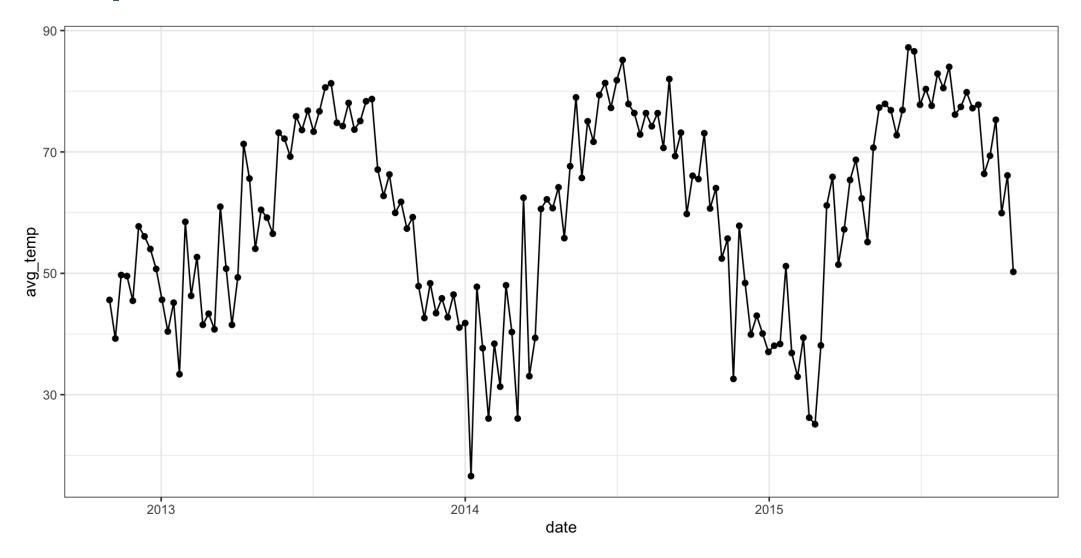
Cov(d) =
$$\sigma^2 \exp(-(d l)^2) + \sigma_w^2 \mathbf{1}_{h=0}$$

 $\gamma(h) = (\sigma^2 + \sigma_w^2) - \sigma^2 \exp(-(h l)^2)$
= $(1.67 + 0.18) - 1.67 \exp(-(8.33 h)^2)$

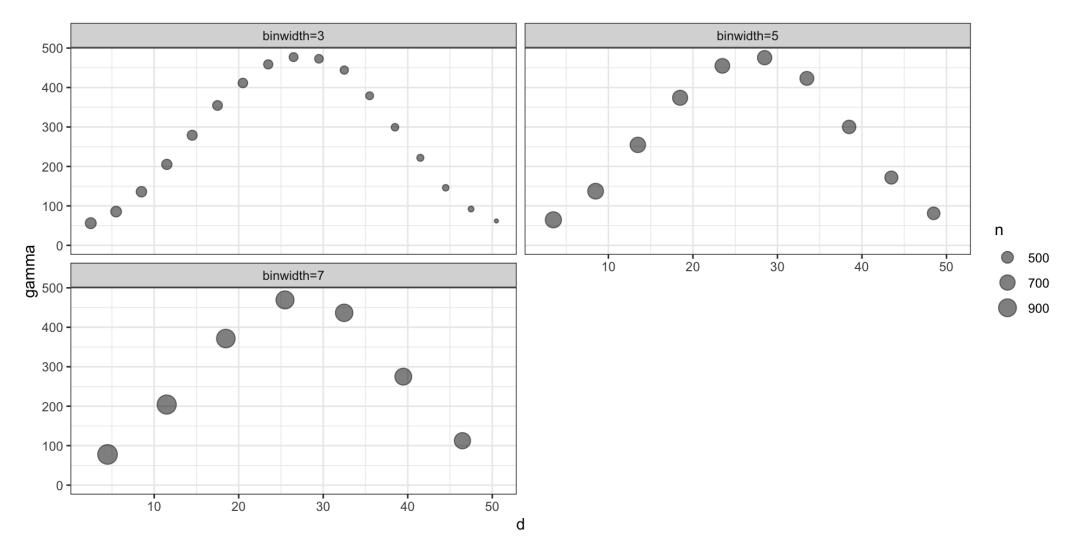


Durham Average Daily Temperature

Temp Data



Empirical semivariogram



Model

What does the model we are trying to fit actually look like?

$$y(t) = \mu(t) + w(t) + \epsilon(t)$$

where

$$\mu(t) = \beta_0$$

$$w(t) \sim (0, \Sigma)$$

$$\epsilon(t) \sim (0, \sigma_w^2)$$

$$\{\Sigma\}_{ij} = \text{Cov}(t_i, t_j) = \sigma^2 \exp(-(|t_i - t_j| 1)^2)$$

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BRMS Model

```
1 library(brms)
 2 	 (m = brm(
   avg temp ~ 1+ gp(week), data=temp,
 4 cores = 4, refresh=0
 5 ) )
Family: gaussian
 Links: mu = identity; sigma = identity
Formula: avg temp ~ 1 + gp(week)
  Data: temp (Number of observations: 156)
 Draws: 4 chains, each with iter = 2000; warmup = 1000; thin = 1;
        total post-warmup draws = 4000
Gaussian Process Terms:
             Estimate Est. Error 1-95% CI u-95% CI Rhat Bulk ESS Tail ESS
sdgp(gpweek) 14.41 6.80 2.56 19.33 4.69
                                                                  12
lscale(gpweek) 0.22 0.10 0.05 0.33 3.44 4
                                                                  12
```

BRMS Alternatives

The BRMS model (and hence Stan) took >10 minutes (per chain) to attempt to fit the model and failed spectacularly.

We could improve things slightly by tweaking the priors and increasing iterations but this wont solve the slowness issue.

The stop gap work around - using spBayes

- Interface is old and clunky (inputs and outputs)
- Designed for spatial GPs
- Super fast (~10 seconds for 20k iterations)
- I am working on a wrapper to make the interface / usage not as terrible (more next week)

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Fitting a model

```
(m = gplm(
     avg temp~1,
     data = d, coords = cbind(d$week, 0),
     starting=list(
 4
 5
       "phi"=sqrt(3)/4, "sigma.sq"=1, "tau.sq"=1
 6
     ),
     priors=list(
       "phi.unif"=c(sqrt(3)/52, sqrt(3)/1),
8
       "sigma.sq.ig"=c(2, 1),
9
       "tau.sq.iq"=c(2, 1)
10
11
    ),
    thin=10
12
13 ) )
```

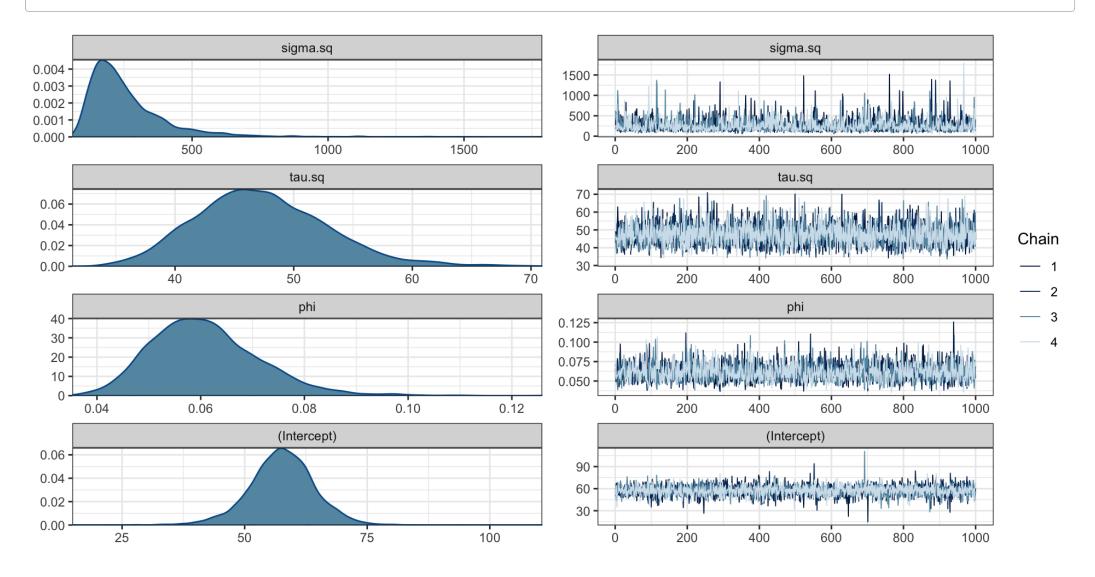
```
# A gplm model (spBayes spLM) with 4 chains, 4 variables, and 4000 iterations.
# A tibble: 4 \times 10
 variable mean median
                           mad
                                  q5 q95 rhat ess b...¹ ess t...²
                      sd
 <dbl>
                                                     <dbl>
1 sigma.sq 2.58e+2 2.18e+2 1.50e+2 9.85e+1 1.11e+2 5.34e+2 1.00
                                               2837.
                                                     3108.
2 tau.sq 4.73e+1 4.69e+1 5.56e+0 5.42e+0 3.89e+1 5.68e+1 0.999
                                               4055.
                                                     3893.
    3 phi
                                               2988.
                                                     3158.
```

4 (Interc... 5.76e+1 5.77e+1 6.75e+0 6.12e+0 4.64e+1 6.83e+1 1.00 4081. 3535.

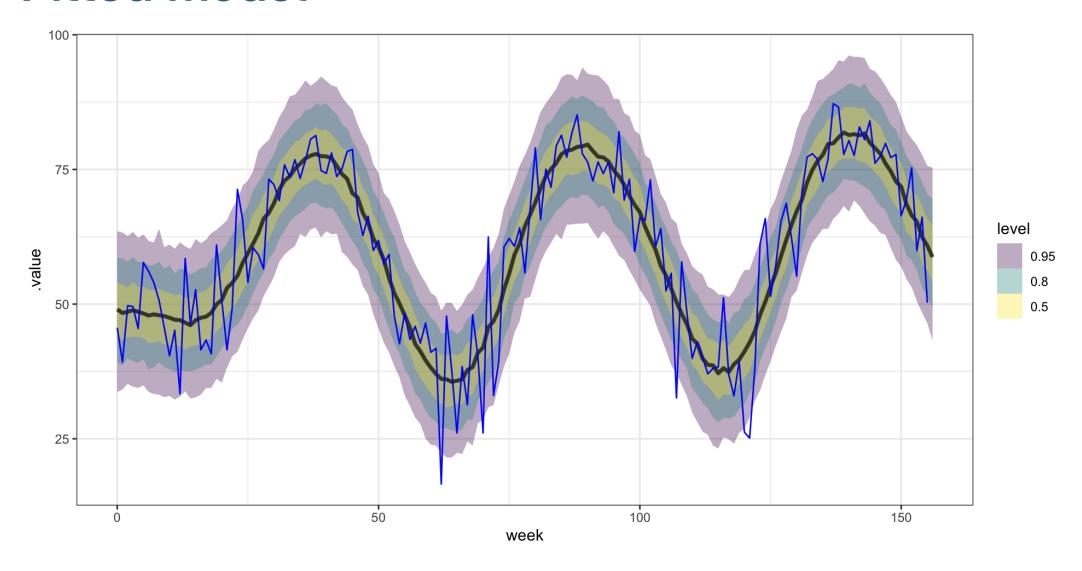
... with abbreviated variable names 'ess_bulk, 'ess_tail

Parameter posteriors

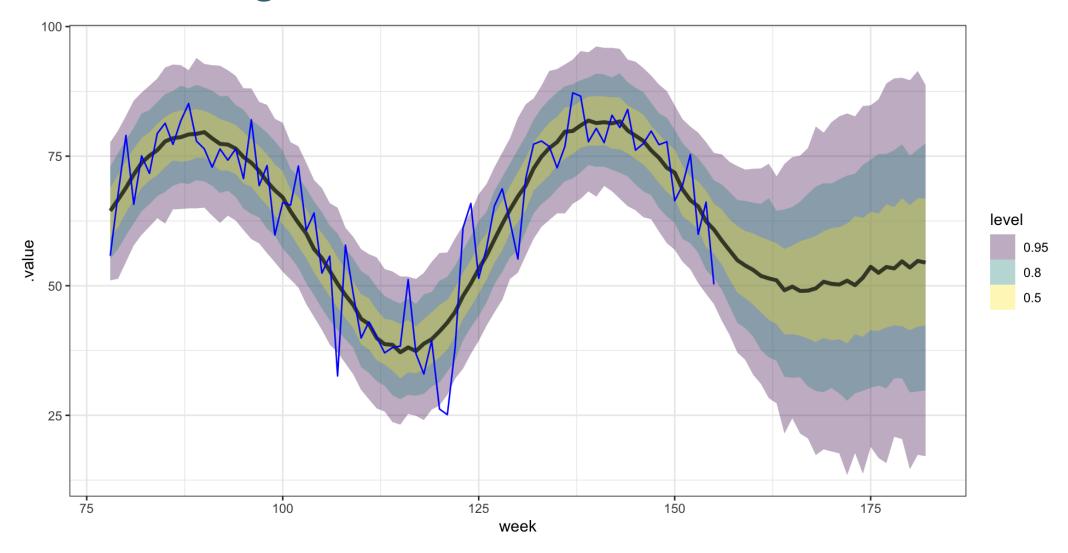
1 plot(m)



Fitted model



Forecasting



Empirical semivariogram vs. model

From the model summary we have the following,

- posterior means: $\sigma^2 = 258$, $\sigma_w^2 = 47.3$, 1 = 0.06
- posterior medians: $\sigma^2 = 218$, $\sigma_w^2 = 46.9$, 1 = 0.06

