Linear Models

Lecture 02

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Linear Models Basics

Pretty much everything we a going to see in this course will fall under the umbrella of either linear or generalized linear models.

In previous classes most of your time has likely been spent with the simple iid case,

$$y_i = eta_0 + eta_1 \, x_{i1} + \dots + eta_p \, x_{ip} + \epsilon_i$$
 $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$

these models can also be expressed simply as,

$$y_i \stackrel{iid}{\sim} N(eta_0 + eta_1 \ x_{i1} + \dots + eta_p \ x_{ip}, \ \sigma^2)$$

Some notes on notation

- Observed values and scalars will usually be lower case letters, e.g. $x_i, y_i, z_{ij}.$
- Parameters will usually be greek symbols, e.g. μ, σ, ρ .
- Vectors and matrices will be shown in bold, e.g. μ, X, Σ .
- Elements of a matrix (or vector) will be referenced with {}s, e.g. $\{{m Y}\}_i, \{{m \Sigma}\}_{ij}$
- $oldsymbol{\cdot}$ Random variables will be indicated by ~, e.g. $x \sim \mathrm{Norm}(0,1), z \sim \mathrm{Gamma}(1,1)$
- Matrix / vector transposes will be indicated with ', e.g. $m{A}', (1-m{B})'$

Linear model - matrix notation

We can also express a linear model using matrix notation as follows,

$$egin{aligned} oldsymbol{Y} &= oldsymbol{X} oldsymbol{eta} + oldsymbol{\epsilon} \ & oldsymbol{n} imes 1 &\sim N(oldsymbol{0}, \ \sigma^2 \ 1_n) \ & n imes n \end{aligned}$$

or alternative as,

$$oldsymbol{X}_{n imes 1} \sim N\left(oldsymbol{X}_{n imes p}oldsymbol{eta},\; \sigma^2 1_n top n imes n
ight)$$

Where possible I will include the dimensions of matrices and vectors as these provide a useful sanity check

Multivariate Normal Distribution - Review

For an n-dimension multivate normal distribution with covariance Σ (positive semidefinite) can be written as

$$oldsymbol{Y}_{n imes 1} \sim N(oldsymbol{\mu}_{n imes 1}, oldsymbol{\Sigma}_{n imes n})$$

where
$$\left\{ oldsymbol{\Sigma}
ight\} _{ij}=
ho _{ij}\sigma _{i}\sigma _{j}$$

$$egin{pmatrix} y_1 \ y_2 \ dots \ y_n \end{pmatrix} \sim N \left(egin{pmatrix} \mu_1 \ \mu_2 \ dots \ \mu_n \end{pmatrix}, egin{pmatrix}
ho_{11}\sigma_1\sigma_1 &
ho_{12}\sigma_1\sigma_2 & \cdots &
ho_{1n}\sigma_1\sigma_n \
ho_{21}\sigma_2\sigma_1 &
ho_{22}\sigma_2\sigma_2 & \cdots &
ho_{2n}\sigma_2\sigma_n \ dots & dots & dots & dots \
ho_{n1}\sigma_n\sigma_1 &
ho_{n2}\sigma_n\sigma_2 & \cdots &
ho_{nn}\sigma_n\sigma_n \end{pmatrix}
ight)$$

Multivariate Normal Distribution - Density

For the n dimensional multivate normal given on the last slide, its density is given by

$$fig(oldsymbol{Y}|oldsymbol{\mu},oldsymbol{\Sigma}ig)=(2\pi)^{-n/2} \ \det(oldsymbol{\Sigma})^{-1/2} \ \exp\left(-rac{1}{2}(oldsymbol{Y}-oldsymbol{\mu})'oldsymbol{\Sigma}_{n imes n}^{-1}(oldsymbol{Y}-oldsymbol{\mu})
ight)$$

and its log density is given by

$$\log fig(m{Y}|m{\mu},m{\Sigma}ig) = -rac{n}{2} \log 2\pi - rac{1}{2} \log \det(m{\Sigma}) - rac{1}{2} (m{Y}_{1 imes n}m{\mu})' rac{m{\Sigma}^{-1}}{n imes n} (m{Y}_{n imes 1}m{\mu})'$$

Some useful identities

The following come from the Matrix Cookbook Chapters 1 & 2.

$$(oldsymbol{A}oldsymbol{B}' = oldsymbol{B}'oldsymbol{A}'$$
 $\partial oldsymbol{A} = 0$ (where $oldsymbol{A}$ is $(oldsymbol{A} + oldsymbol{B})' = oldsymbol{A}' + oldsymbol{B}'$ $\partial (oldsymbol{A}X) = a(\partial oldsymbol{A})$ $\partial (oldsymbol{X} + oldsymbol{Y}) = \partial oldsymbol{X} + \partial oldsymbol{Y}$ $(oldsymbol{A}oldsymbol{B}C \ldots)^{-1} = \ldots oldsymbol{C}^{-1}oldsymbol{B}^{-1}oldsymbol{A}^{-1}$ $\partial (oldsymbol{X} + oldsymbol{Y}) = \partial oldsymbol{X} + oldsymbol{X}(\partial oldsymbol{Y})$ $\partial (oldsymbol{X}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{X}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{A}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{X}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{A}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{A}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{X}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{A}') = (\partial oldsymbol{X})'$ $\partial (oldsymbol{X}') = (\partial oldsymbol{X})'$ $\partial (oldsymbo$

Maximum Likelihood - β (iid)

Maximum Likelihood - σ^2 (iid)

A Quick Example

Parameters -> Synthetic Data

Lets generate some simulated data where the underlying model is known and see how various regression procedures function.

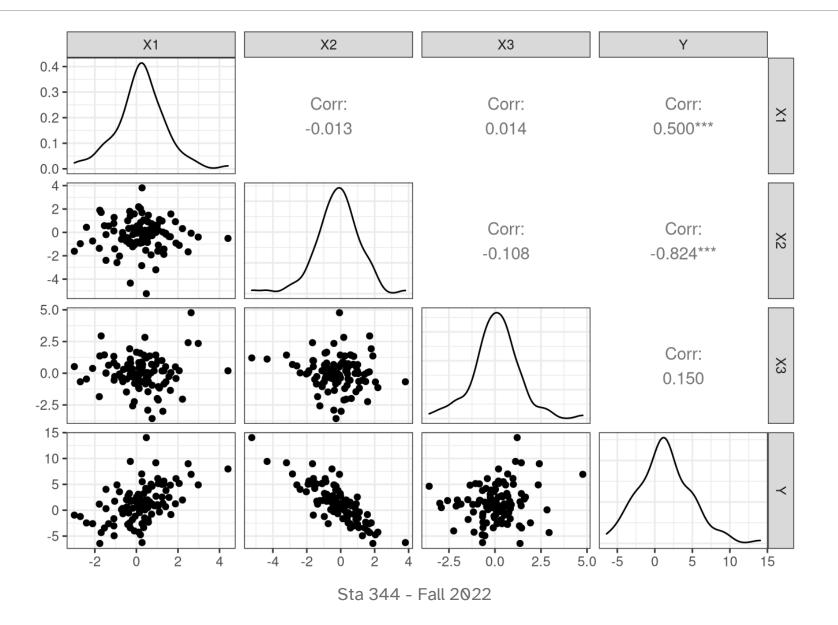
$$y_i = eta_0 + eta_1 x_{1i} + eta_2 x_{2i} + eta_3 x_{3i} + \epsilon_i \ eta_0 = 0.7, \quad eta_1 = 1.5, \quad eta_2 = -2.2, \quad eta_3 = 0.1 \ \epsilon_i \sim N(0,1)$$

Model Matrix

```
1 X = model.matrix(\sim X1+X2+X3, d)
 2 \text{ tbl\_df}(X)
# A tibble: 100 × 4
   `(Intercept)`
                             X2
                                     X3
                     X1
          <dbl> <dbl> <dbl>
                                  <dbl>
                 0.557 0.897
                                -1.46
1
 2
                       0.375
                 0.758
                                -0.945
 3
                 0.273
                       3.81
                                 -0.675
                 1.41
                        -0.0745
                                0.514
 4
              1
 5
                 1.01 0.623
                                 -1.99
 6
              1 0.942
                        -0.00618 0.700
              1
                 1.66
                        1.57
                             0.0478
8
              1 -1.09 0.766
                                1.33
 9
              1 - 0.296
                         1.40
                                 -0.0914
```

Pairs plot

1 GGally::ggpairs(d, progress = FALSE)



Least squares fit

Let $\hat{m{Y}}$ be our estimate for $m{Y}$ based on our estimate of $m{eta}$,

$$\hat{m{Y}} = \hat{eta}_0 + \hat{eta}_1 \, m{X}_1 + \hat{eta}_2 \, m{X}_2 + \hat{eta}_3 \, m{X}_3 = m{X} \, \hat{m{eta}}_3$$

The least squares estimate, $\hat{m{eta}}_{ls}$, is given by

$$rgmin_{oldsymbol{eta}} \sum_{i=1}^n \left(Y_i - oldsymbol{X}_{i\cdot}oldsymbol{eta}
ight)^2$$

Previously we showed that,

$$\hat{oldsymbol{eta}}_{ls} = (oldsymbol{X}'oldsymbol{X})^{-1}oldsymbol{X}'oldsymbol{Y}$$

Frequentist Fit

```
1 (beta_hat = solve(t(X) %*% X, t(X)) %*% d$Y)
                  [,1]
(Intercept) 0.5522298
X1
            1.4708769
X2
            -2.1761159
X3
            0.1535830
 1 l = lm(Y \sim X1 + X2 + X3, data=d)
 2 l$coefficients
(Intercept)
                    X1
                                 X2
                                             X3
  0.5522298 1.4708769 -2.1761159 0.1535830
```

Bayesian regression model

Basics of Bayes

brms

We will be using a package called brms for most of our Bayesian model fitting

- it has a convenient model specification syntax
- mostly sensible prior defaults
- supports most of the model types we will be exploring
- uses Stan behind the scenes

brms + linear regression

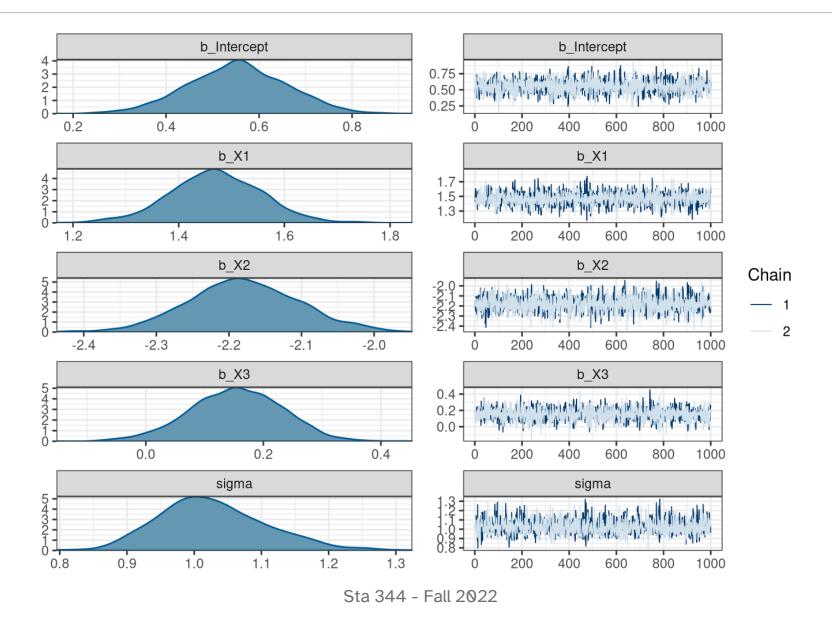
```
1 b = brms::brm(Y \sim X1 + X2 + X3, data=d, chains = 2)
Running /usr/lib64/R/bin/R CMD SHLIB foo.c
gcc -m64 -I"/usr/include/R" -DNDEBUG
I"/usr/lib64/R/library/Rcpp/include/"
I"/usr/lib64/R/library/RcppEigen/include/" -
I"/usr/lib64/R/library/RcppEigen/include/unsupported" -
I"/usr/lib64/R/library/BH/include" -
I"/usr/lib64/R/library/StanHeaders/include/src/"
I"/usr/lib64/R/library/StanHeaders/include/" -
I"/usr/lib64/R/library/RcppParallel/include/" -
I"/usr/lib64/R/library/rstan/include" -DEIGEN_NO_DEBUG
DBOOST_DISABLE_ASSERTS - DBOOST_PENDING_INTEGER_LOG2_HPP
                                                          -DSTAN THREADS
DBOOST NO AUTO PTR -include
```

Model results

```
1 b
Family: gaussian
 Links: mu = identity; sigma = identity
Formula: Y \sim X1 + X2 + X3
  Data: d (Number of observations: 100)
 Draws: 1 chains, each with iter = 2000; warmup = 1000; thin = 1;
        total post-warmup draws = 1000
Population-Level Effects:
        Estimate Est.Error l-95% CI u-95% CI Rhat Bulk_ESS Tail_ESS
                     0.11 0.34
                                    0.76 1.00
                                                 1269
                                                          627
Intercept
            0.55
X1
            1.47 0.09 1.29 1.64 1.00
                                                 1273
                                                          723
X2
    -2.18 0.08 -2.32 -2.02 1.00
                                                 1178
                                                          663
```

Model visual summary

1 plot(b)



What about the priors?

```
1 brms::prior_summary(b)
                                           1 brms::get_prior("Y ~ X1 + X2 + X;
                                                             prior
                   prior
                              class
                                                                        class
                                          coef group resp dpar nlpar lb ub
coef group resp dpar nlpar lb ub
source
                                          source
                  (flat)
                                  b
                                                            (flat)
                                                                            b
default
                                          default
                  (flat)
                                                            (flat)
                                  b
                                                                            b
X1
                                          X1
(vectorized)
                                          (vectorized)
                  (flat)
                                  b
                                                            (flat)
                                                                            b
X2
                                          X2
(vectorized)
                                          (vectorized)
                  (flat)
                                                            (flat)
                                  b
                                                                            b
```

tidybayes

```
(post = b \%)
      tidybayes::gather_draws(b_Intercept, b_X1, b_X2, b_X3, sigma)
 3
# A tibble: 10,000 × 5
# Groups: .variable [5]
   .chain .iteration .draw .variable
                                     .value
                                <dbl>
    <int> <int> <int> <chr>
                  1
                        1 b_Intercept
                                       0.440
 1
        1
 2
                        2 b_Intercept 0.524
       1
 3
                        3 b_Intercept 0.544
       1
                  4
                        4 b_Intercept
                                       0.502
 4
        1
 5
       1
                  5
                        5 b_Intercept
                                       0.567
 6
       1
                  6
                        6 b_Intercept 0.569
        1
                        7 b_Intercept 0.373
 8
       1
                        8 b_Intercept 0.731
```

tidybayes + posterior summaries

2

6 b_X2

7 b X3

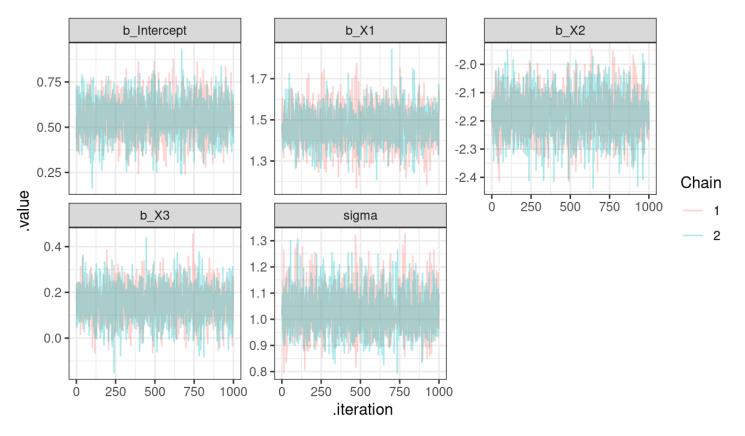
```
(post_sum = post %>%
     group_by(.variable, .chain) %>%
     summarize(
       post_mean = mean(.value),
       post_median = median(.value)
 5
# A tibble: 10 \times 4
# Groups: .variable [5]
  .variable .chain post_mean post_median
  <chr>
        <int> <dbl>
                                 <dbl>
1 b_Intercept
                 1 0.556 0.555
2 b_Intercept
             2 0.550 0.553
3 b_X1
                      1.47
                                 1.47
                      1.47 1.47
4 b_X1
5 b_X2
                 1
                      -2.18 -2.18
```

-2.18 -2.18

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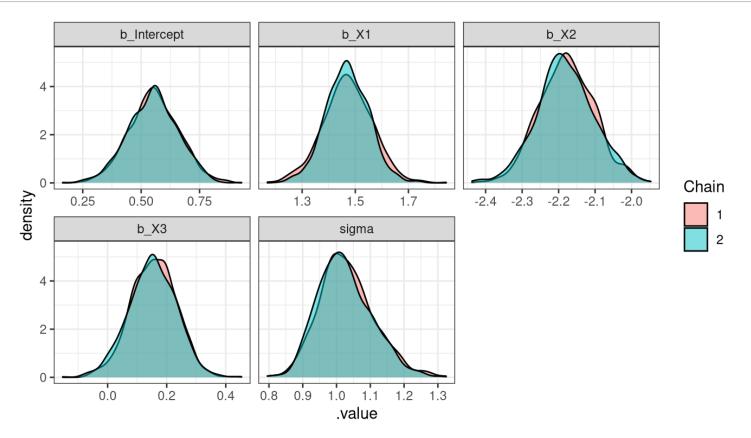
tidybayes + ggplot - traceplot

```
post %>%
ggplot(aes(x=.iteration, y=.value, color=as.character(.chain))) +
geom_line(alpha=0.33) +
facet_wrap(~.variable, scale="free_y") +
labs(color="Chain")
```



Tidy Bayes + ggplot - Density plot

```
post %>%
ggplot(aes(x=.value, fill=as.character(.chain))) +
geom_density(alpha=0.5) +
facet_wrap(~.variable, scale="free_x") +
labs(fill="Chain")
```



Comparing Approaches

3 b_X2 -2.2 -2.18 -2.18

4 b_X3 0.1 0.154 0.157

1 1.00

5 sigma

```
1 (pt_est = post_sum %>%
     filter(.chain == 1) %>%
     ungroup() %>%
     mutate(
   truth = c(beta, sigma),
 5
    ols = c(l$coefficients, sd(l$residuals))
    ) %>%
 8
     select(.variable, truth, ols, post_mean)
 9
# A tibble: 5 \times 4
  .variable truth ols post_mean
 <chr> <dbl> <dbl> <dbl>
1 b_Intercept 0.7 0.552 0.556
2 b_X1
      1.5 1.47 1.47
```

1.03

Comparing Approaches - code

```
post %>%
filter(.chain == 1) %>%
ggplot(aes(x=.value)) +
geom_density(alpha=0.5, fill="lightblue") +
facet_wrap(~.variable, scale="free_x") +
geom_vline(
data = pt_est %>% tidyr::pivot_longer(cols = truth:post_mean, names_aes(xintercept = value, color=pt_est),
alpha = 0.5, size=2
)
```

Comparing Approaches - plot

