

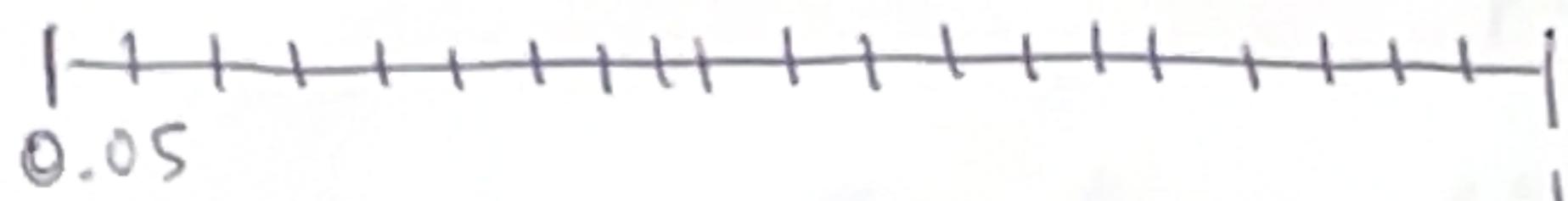
①

Bayesian hierarchical regression models =
mixed effects models

From Haigis et al 2004:

21 mice

intestine:



0.05



divided into 20 equally spaced regions

count the # of tumors in each region for
each mouse.

In regression, recall we're always interested
in $EY|X$.

- * From figure, tumor counts look more similar
within a mouse than b/wn mice.

Let $\boxed{Y_{x,j}}$ be mouse j's tumor count at location x.

$$Y_{x,j} \sim \text{Poisson}(\theta_j(x))$$

$$p(Y_{x,j} | \theta_j) = \frac{\theta_j^{y_{x,j}} e^{-\theta_j(x)}}{y_{x,j}!}$$

(2)

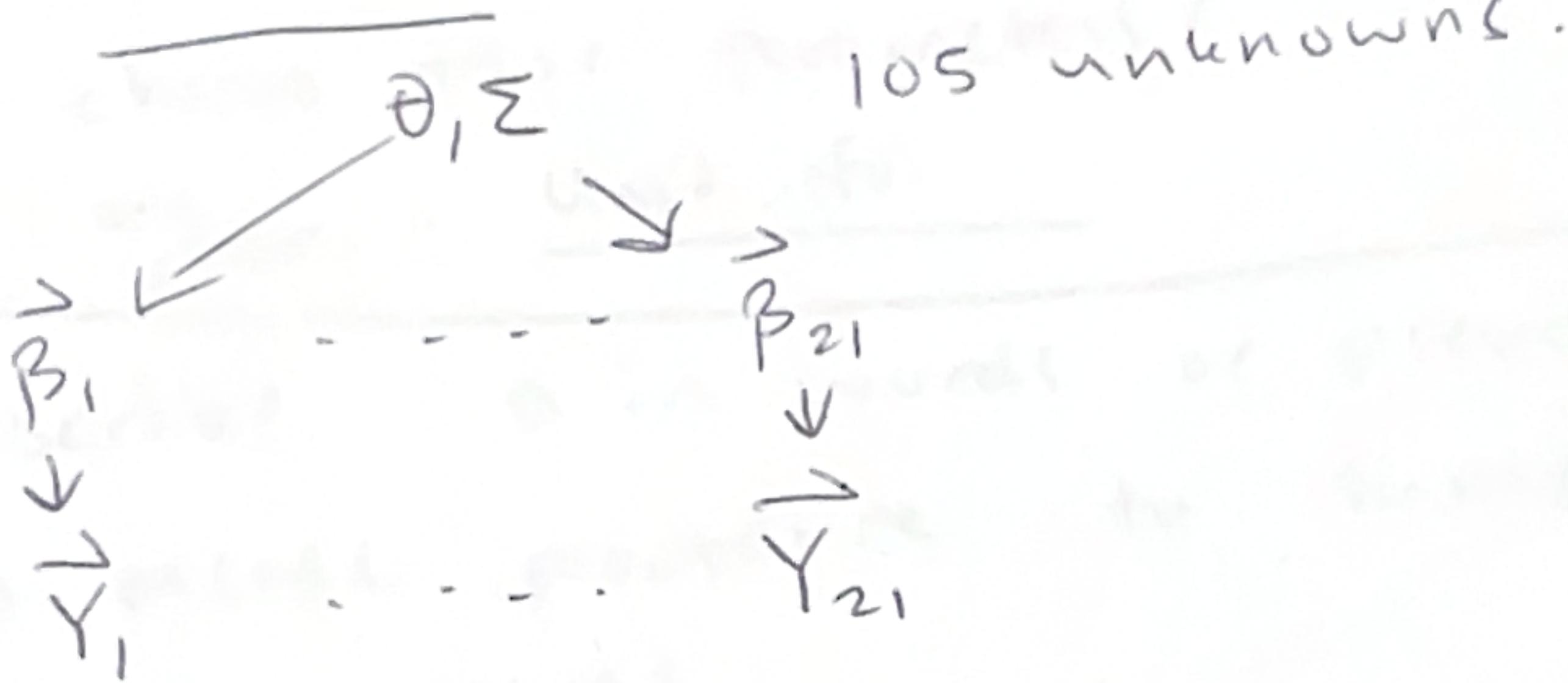
$$Y_{x,j} \sim \text{Poisson}(\theta_j(x))$$

$$\mathbb{E} Y_{x,j} | x = \theta_j(x)$$

$$\log(\theta_j(x)) = x\beta = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \beta_4 x^4$$

Data

unknowns: β_s 5 β s for each mouse!



$\beta_j \stackrel{\text{iid}}{\sim} \text{MVN}(\theta, \Sigma) \leftarrow \text{btwn group sampling model}$.

This hierarchical regression model is a ~~a~~ mixed effects model.

$$\beta_j = \theta + \gamma_j \quad \gamma_j \sim \text{MVN}(0, \Sigma)$$

\uparrow \uparrow
 fixed effect random effect

$$x\beta = x_j(\theta + \gamma_j) = \theta_0 + \theta_1 x_j + \theta_2 x_j^2 + \theta_3 x_j^3 + \dots + \theta_n x_j^n$$

Aside ex: General random effects model

6

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$$Y_{ij} = \theta^T \gamma_{ij} + \gamma_j^T \varepsilon_{ij} + \varepsilon_{ij}$$

\uparrow \uparrow

fixed effects random effects

(4)

Putting it all together,

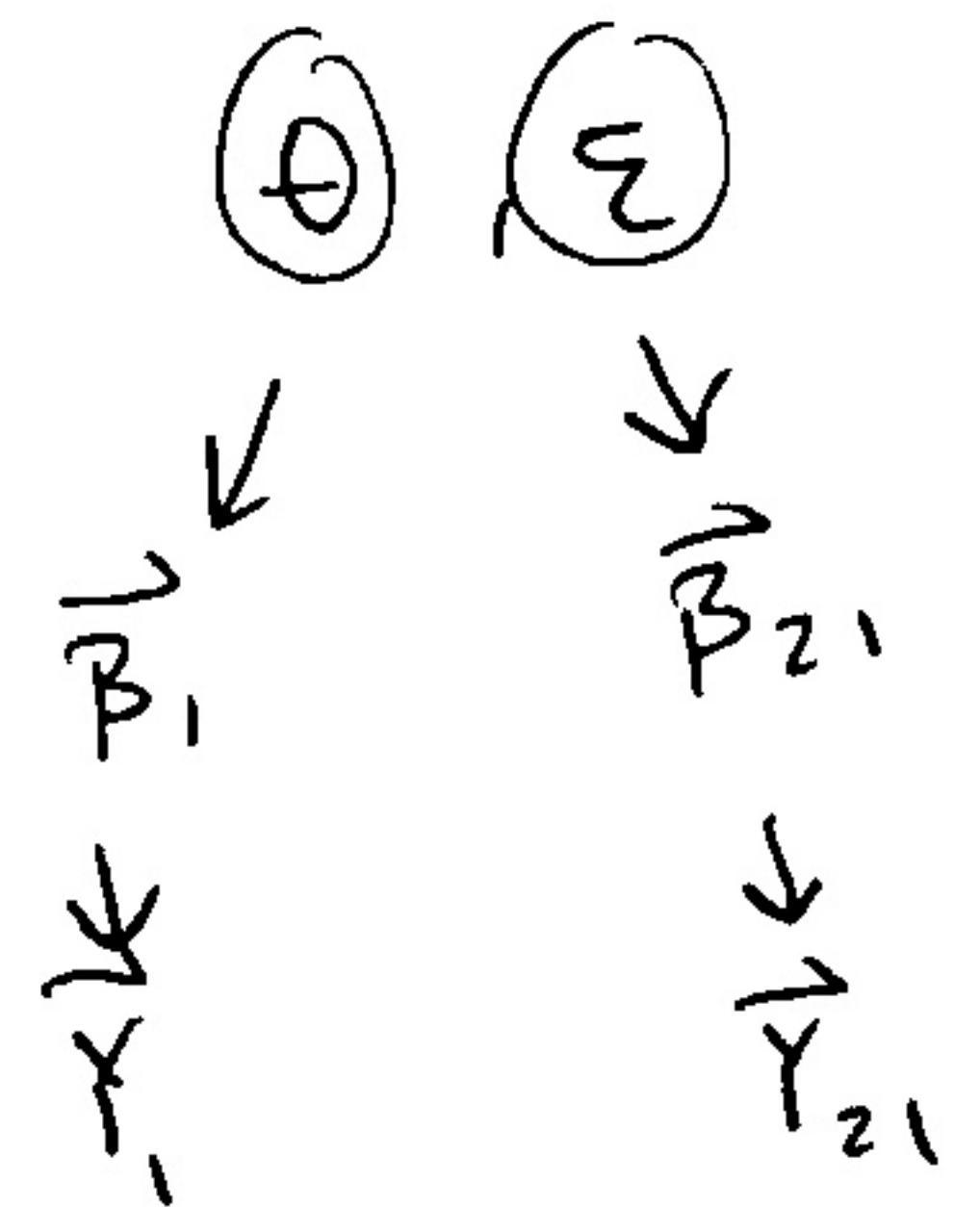
$$Y_{x,j} \mid \text{on } X, \beta \sim \text{Poisson}(\theta_j(x))$$

$$\log \theta_j(x) = \beta_{1,j} + \beta_{2,j} x + \dots + \beta_{5,j} x^4$$

$$\beta_j \mid \theta, \Sigma \sim \text{MVN}(\theta, \Sigma)$$

$$\theta \sim \text{MVN}(\mu, \Lambda_0)$$

$$\Sigma \sim \text{inv-Wishart}(n_0, S_0)$$



how to choose prior parameters?

one option : Unit info.

Quiz (1) Describe ~~a~~ in words or pseudo-code
an MCMC procedure to sample
all unknowns -

(5)

$$p(\beta, \theta, \varepsilon | y, x) \propto \underbrace{p(y | \beta, \theta, \varepsilon, x)}_{p(y | \beta, x) p(\beta | \theta, \varepsilon) p(\theta) p(\varepsilon)} p(\beta, \theta, \varepsilon | x)$$

$p(y | \beta, x) p(\beta | \theta, \varepsilon) p(\theta) p(\varepsilon)$

To facilitate Gibbs sampling, we need full cond'l post.

$$\underbrace{p(\theta | \cdot)}_{\text{MVN}} \propto \underbrace{p(\beta | \theta, \varepsilon)}_{\text{MVN}} \underbrace{p(\varepsilon)}_{\text{MVN}}$$

$$\underbrace{p(\varepsilon | \cdot)}_{\text{inv-wish}} \propto p(\beta | \theta, \varepsilon) p(\varepsilon)$$

$$p(\beta | \cdot) \propto p(y | \beta, x) \underbrace{p(\beta | \theta, \varepsilon)}_{\text{MVN}}$$

Recall $y_{x,j} \sim \text{Poisson}(\theta_j(x))$

$$p(y | \beta, x) = \prod_{j=1}^{21} \frac{\theta_j(x)^{y_{x,j}} e^{-\theta_j(x)}}{y_{x,j}!}$$

$$\prod_{j=1}^{21} \frac{(e^{x\beta_j})^{y_{x,j}} e^{-e^{x\beta_j}}}{y_{x,j}!}$$

can't identify full cond'l for $\beta | \cdot$
 \Rightarrow Metropolis algo-

(6)

initialize $\beta^{(0)}, \theta^{(0)}$
 for ($s = 1 : S$) {
 sample $\Sigma \sim \text{inv-wishart}(-, -)$
 sample $\theta \sim \text{MVN}(-, -)$
 propose $\beta^* \sim \text{MVN}(\beta^{(s)}, s_0)$
 accept/reject w/ prob. $\min(1, r)$
 $r = \frac{p(\beta^* | \cdot)}{p(\beta^{(s)} | \cdot)}$
 }