

①

What about nonconjugate priors?

$$\overline{p(\theta | \vec{y})} \propto \boxed{\vec{y} = y_1, \dots, y_n}$$

$$\prod_{i=1}^n \theta^{y_i} \frac{e^{-\theta}}{y_i!}$$

$y_i | \theta \sim \text{Poisson}$

$\theta > 0$

likelihood

$$p(y_1, \dots, y_n | \theta)$$

$$\frac{1}{\Theta \sigma \sqrt{2\pi}} \exp \left\{ \frac{-(\ln \theta - \mu)^2}{2\sigma^2} \right\}$$

Prior (log-normal)
 $p(\theta)$

$$\mathbb{E} \theta | y_1, \dots, y_n = ?$$

Exercise 1

$$p(\theta) = \frac{1}{2} \quad \boxed{0 < \theta < 2}$$

$$\phi = \log \theta \Rightarrow \theta = e^\phi$$

want $p(\phi)$

$$\rightarrow p(\phi) d\phi = p(\theta) d\theta$$

$$p(\phi) = \underbrace{p(\theta)}_{\text{constant}} \underbrace{\left| \frac{d\theta}{d\phi} \right|}_{=1}$$

$$p(\phi) = \frac{1}{2} e^\phi \quad -\infty < \phi < \log 2$$

(2)

Exercise 2

We want $\vec{y}_2 = \{y_{21}, \dots, y_{2n_2}\}$

$$p(\tilde{y}_1 < \tilde{y}_2 | \vec{y}_1, \vec{y}_2)$$

↙ ↓ ↘

new obs
of # children
to a woman
w/o bachelors same
but
"w/
bachelors"

$\vec{y}_1 = \{y_{11}, \dots, y_{1n_1}\}$

To compute w/ Monte Carlo sampling
we need to be able to sample from

$$p(\tilde{y}_1 | \vec{y}_1, \vec{y}_2) \quad \text{and}$$

$$p(\tilde{y}_2 | \vec{y}_1, \vec{y}_2)$$

By independence assumed in our model:

$$p(\tilde{y}_1 | \vec{y}_1, \vec{y}_2) = p(\tilde{y}_1 | \vec{y}_1)$$

$$p(\tilde{y}_2 | \vec{y}_1, \vec{y}_2) = p(\tilde{y}_2 | \vec{y}_2)$$

In general, when \tilde{y} is a new obs.
from the same population, we call

$p(\tilde{y} | \vec{y})$ the "posterior
predictive distribution"

$$\vec{y}_1 \rightarrow \Theta_1 \rightarrow \tilde{y}_1$$

$$\vec{y}_2 \rightarrow \Theta_2 \rightarrow \tilde{y}_2$$