

(1)

"Completing the square"

Let $Y_i | \theta, \sigma^2 \stackrel{\text{iid}}{\sim} \text{normal}(\theta, \sigma^2)$

* note: $\vec{y} = \{y_1, \dots, y_n\}$

Let $\theta | \sigma^2 \sim \text{normal}(\mu_0, \tau_0^2)$

We want to prove that the full conditional posterior
 $\theta | \sigma^2, \vec{y}$ is normal with some posterior mean μ_n
 and posterior variance τ_n^2 . i.e. we want to prove:
 $\boxed{\theta | \sigma^2, \vec{y} \sim \text{normal}(\mu_n, \tau_n^2)}$

PROOF:

STEP 1: Bayes' thm:

$$p(\theta | \sigma^2, \vec{y}) \propto \underbrace{p(\vec{y} | \theta, \sigma^2)}_{\prod_{i=1}^n \text{dnorm}(y_i, \theta, \sigma)} p(\theta | \sigma^2)$$

$$\propto \exp \left\{ -\frac{1}{2\sigma^2} \sum (y_i - \theta)^2 \right\} \cdot \exp \left\{ -\frac{1}{2\tau_0^2} (\theta - \mu_0)^2 \right\}$$

STEP 2: simplify by expanding quadratic terms and absorbing what we can into the normalizing constant. *Note: $\sum y_i = n\bar{y}$

$$\propto \exp \left\{ -\frac{1}{2\sigma^2} (n\theta^2 - 2n\bar{y}\theta) - \frac{1}{2\tau_0^2} (\theta^2 - 2\mu_0\theta) \right\}$$

combine like terms:

$$\propto \exp \left\{ -\frac{1}{2} \underbrace{\left(\frac{n}{\sigma^2} + \frac{1}{\tau_0^2} \right)}_{\text{"a"}} \theta^2 + \underbrace{\left(\frac{n\bar{y}}{\sigma^2} + \frac{\mu_0}{\tau_0^2} \right)}_{\text{"b"}} \theta \right\}$$

define "a" & "b":

$$\propto \exp \left\{ -\frac{1}{2} a \theta^2 + b \theta \right\}$$

(*) Notice: this is the kernel of a normal with $\text{mean} = \frac{b}{a}$
 and $\text{variance} = a^{-1}$

therefore, $\theta | \sigma^2, \vec{y} \sim \text{normal}(\mu_n, \tau_n^2)$ where

$$\mu_n = \frac{b}{a} \quad \text{and} \quad \tau_n^2 = a^{-1}$$

and a, b are defined above. \square

Proof of (*) on previous page: This answers the question: (2)
How did I recognize $\exp\left\{-\frac{1}{2}a\theta^2 + b\theta\right\}$ as the
kernel of a normal w/ mean $\frac{b}{a}$ and variance a^{-1} ?

Assume $\theta | \vec{y}, \sigma^2 \sim \text{normal}\left(\frac{b}{a}, a^{-1}\right)$.

Then $p(\theta | \vec{y}, \sigma^2) \propto \exp\left\{-\frac{1}{2}a(\theta - \frac{b}{a})^2\right\}$
 $\propto \exp\left\{-\frac{1}{2}a\theta^2 + b\theta - \frac{1}{2}a\frac{b^2}{a^2}\right\}$
 $\propto \exp\left\{-\frac{1}{2}a\theta^2 + b\theta\right\}$ ~~$\propto \frac{b^2}{a^2}$~~ constant in θ .



Today's punchline:

Assume:

1 $Y_i | \theta, \sigma^2 \sim N(\theta, \sigma^2)$

2 $\theta | \sigma^2 \sim N(\mu_0, \sigma^2 / K_0)$

3 $\sigma^2 \sim \text{gamma}\left(\frac{v_0}{2}, \frac{v_0 \sigma_0^2}{2}\right)$

then the posterior

$$p(\theta, \sigma^2 | y_1, \dots, y_n)$$

$$= p(\theta | \sigma^2, y_1, \dots, y_n) p(\sigma^2 | y_1, \dots, y_n)$$



$$\text{dnorm}(\theta; \mu_n, \tau_n) \text{ dinvgamma}(\sigma^2; \frac{v_n}{2}, \frac{v_n \sigma_n^2}{2})$$

"full cond'l posterior of θ "

Today's agenda:

(1) sketch proof for $p(\sigma^2 | y_1, \dots, y_n)$

(2) sample from ^{joint} posterior

(3) sample from posterior predictive $p(\tilde{y} | y_1, \dots, y_n)$
 ~ (time permitting)

Interpretation:

μ_0 : prior guess for θ

σ_0^2 : prior guess for σ^2

K_0 : prior sample size for θ

v_0 : prior sample size for σ^2

Assume

$$\textcircled{1} \quad Y_i | \theta, \sigma^2 \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$$

$$\textcircled{2} \quad \theta | \sigma^2 \sim N(\mu_0, \sigma^2 / k_0)$$

$$\textcircled{3} \quad 1/\sigma^2 \sim \text{gamma}\left(\nu_0/2, \frac{\nu_0 \sigma_0^2}{2}\right)$$

then the joint posterior

$$p(\theta, \sigma^2 | \vec{y}) = p(\theta | \sigma^2, \vec{y}) p(\sigma^2 | \vec{y})$$
$$= \text{dnorm}(\theta, \mu_n, T_n) \cdot \text{dinvgamma}(\sigma^2, \frac{\nu_n}{2}, \frac{\nu_n \sigma_n^2}{2})$$

\checkmark shown previously

Goal: to prove $\sigma^2 | \vec{y} \sim \text{invgamma}$

Sketch proof:

$$p(\sigma^2 | \vec{y}) \propto p(\vec{y} | \sigma^2) p(\sigma^2)$$
$$\propto p(\sigma^2) \int p(\vec{y}, \theta | \sigma^2) d\theta$$
$$\propto p(\sigma^2) \int p(\vec{y} | \theta, \sigma^2) p(\theta | \sigma^2) d\theta \quad (+)$$

Known by assumption:

$$\textcircled{3} \quad \downarrow \quad \textcircled{1} \quad \downarrow \quad \textcircled{2}$$

The integral (+) above reduces to integrating a normal density.

Explicit bookkeeping:

attacking $\int p(\vec{y} | \theta, \sigma^2) p(\theta | \sigma^2) d\theta$

- (1) collect (but do not absorb into normalizing const!) terms
w/o θ outside integral:

$$\boxed{(2\pi\sigma^2)^{-n/2} \cdot \left(2\pi\sigma^2/K_0\right)^{-1/2} \cdot \exp\left\{-\frac{1}{2\sigma^2}(\sum y_i^2) - \frac{K_0}{2\sigma^2}(\mu_0)^2\right\}} \cdot \int \exp\left\{-\frac{1}{2\sigma^2}(-2(\sum y_i)\theta + n\theta^2) - \frac{K_0}{2\sigma^2}(-2\mu_0\theta + \theta^2)\right\} d\theta$$

call + his.
 $= c_1(\sigma^2)$

- (2) Compute the integral above using the kernel trick:
hint: it's the kernel of a normal, you must complete the square.

$$\int e^{-\frac{1}{2}a\theta^2 + b\theta} d\theta = \boxed{\left(2\pi/a\right)^{1/2} \exp\left\{\frac{1}{2}b^2/a\right\}}$$

where $a = \left(\frac{n}{\sigma^2} + \frac{K_0}{\sigma^2}\right)$

$b = \left(\frac{n\bar{y} + K_0\mu_0}{\sigma^2}\right)$

call + his
 $= c_2(\sigma^2)$

- (3) Return to (1) and

write down:

$$p(\sigma^2 | \vec{y}) \propto p(\sigma^2) \cdot c_1(\sigma^2) \cdot c_2(\sigma^2)$$

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