

Wald tests

Recap

want to test $H_0: \beta_1 = \beta_2 = 0$ H_A : at least one of $\beta_1, \beta_2 \neq 0$

$$\begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} = \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{bmatrix} = C\beta$$

$$\hat{\beta} \sim N(\beta, \hat{\Sigma}^{-1}(\beta))$$

$$\Rightarrow C\hat{\beta} \sim N(C\beta, C\hat{\Sigma}^{-1}(\beta)C^T)$$

$$\Rightarrow (C\hat{\Sigma}^{-1}(\beta)C^T)^{-\frac{1}{2}} (C\hat{\beta} - C\beta) \sim N(0, \overset{\substack{\uparrow \\ \text{identity matrix}}}{I})$$

$$\text{If } Z \sim N(0, I) \quad Z \in \mathbb{R}^q, \text{ then } Z^T Z \sim \chi^2_q$$

$$\Rightarrow (C\hat{\beta} - C\beta)^T (C\hat{\Sigma}^{-1}(\beta)C^T)^{-1} (C\hat{\beta} - C\beta) \sim \chi^2_q$$

$q = \text{length of } C\beta$

logistic regression: $V_n = \hat{I}^{-1}(\hat{\beta})$ (using MLE $\hat{\beta}$)

General Wald test

Let $\theta \in \mathbb{R}^p$ and let $\hat{\theta}_n$ be an estimator such that

$$V_n^{-\frac{1}{2}} (\hat{\theta}_n - \theta) \xrightarrow{d} N(0, I)$$

$$(\Rightarrow \hat{\theta}_n \approx N(\theta, V_n))$$

Let $C \in \mathbb{R}^{q \times p}$. To test

$$H_0: C\theta = \gamma_0$$

$$H_A: C\theta \neq \gamma_0$$

$$Z_n = (CV_n C^T)^{-\frac{1}{2}} (C\hat{\theta}_n - \gamma_0) \quad Z_n \approx N(0, I) \text{ under } H_0$$

$$\text{Let } W_n = (C\hat{\theta}_n - \gamma_0)^T (CV_n C^T)^{-1} (C\hat{\theta}_n - \gamma_0) = Z_n^T Z_n$$

$$\text{Under } H_0, W_n \approx \chi^2_q$$

Wald test: specifies $\alpha \in [0, 1]$

rejects H_0 when $W_n > \chi^2_{q, \alpha}$

$\chi^2_{q, \alpha}$ = upper α quantile of χ^2_q



Class activity, Part I

$$H_0: C\beta = \underbrace{\begin{bmatrix} 0 \\ 0 \end{bmatrix}}_{\delta_0}$$

$$\hat{\beta} \Rightarrow C\hat{\beta}$$

https://sta711-s23.github.io/class_activities/ca_lecture_18.html

$$\beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \end{bmatrix}$$

$$\begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} = C\beta$$

$$V_n = X^{-1}(\beta) = \begin{bmatrix} \text{Var}(\hat{\beta}_0) & \text{Var}(\hat{\beta}_1) & \text{Var}(\hat{\beta}_2) & \text{Var}(\hat{\beta}_3) & \text{Var}(\hat{\beta}_4) \\ \text{Cov}(\hat{\beta}_0, \hat{\beta}_1) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_2) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_3) & \text{Cov}(\hat{\beta}_0, \hat{\beta}_4) & \text{Cov}(\hat{\beta}_1, \hat{\beta}_2) \\ \text{Cov}(\hat{\beta}_1, \hat{\beta}_3) & \text{Cov}(\hat{\beta}_1, \hat{\beta}_4) & \text{Cov}(\hat{\beta}_2, \hat{\beta}_3) & \text{Cov}(\hat{\beta}_2, \hat{\beta}_4) & \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) \\ \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) & \text{Cov}(\hat{\beta}_4, \hat{\beta}_4) & & & \end{bmatrix}$$

(all the covariances)

call the covariances

Var($\hat{\beta}_3$)

Var($\hat{\beta}_4$)

$$C V_n C^T = \begin{bmatrix} \text{Var}(\hat{\beta}_3) & \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) \\ \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) & \text{Var}(\hat{\beta}_4) \end{bmatrix}$$

$$\text{Test statistic : } \begin{bmatrix} \hat{\beta}_3 \\ \hat{\beta}_4 \end{bmatrix} \begin{bmatrix} \text{Var}(\hat{\beta}_3) & \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) \\ \text{Cov}(\hat{\beta}_3, \hat{\beta}_4) & \text{Var}(\hat{\beta}_4) \end{bmatrix} \begin{bmatrix} \hat{\beta}_3 \\ \hat{\beta}_4 \end{bmatrix}$$

(W_n)

$$\mathcal{I}^{-1}(\beta) = \text{vcov}(\text{model})$$

Class activity

```
betahat <- m1$coefficients[4:5]
V <- vcov(m1)[4:5, 4:5]
test_stat <- t(betahat) %*% solve(V) %*% betahat
test_stat
```

```
##           [,1]
## [1,] 85.60437 ← W_n
```

Under H_0 , $W_n \approx \chi^2_2 \leftarrow \# \text{ parameters tested}$

Reject when $W_n > \chi^2_{2,\alpha}$

Class activity

```
betahat <- m1$coefficients[4:5]
V <- vcov(m1)[4:5, 4:5]
test_stat <- t(betahat) %*% solve(V) %*% betahat
test_stat
```

```
##           [,1]
## [1,] 85.60437
```

rejection region for alpha = 0.05

```
qchisq(0.05, df=2, lower.tail=F)
```

```
## [1] 5.991465
```

$\chi^2_{2,0.05}$

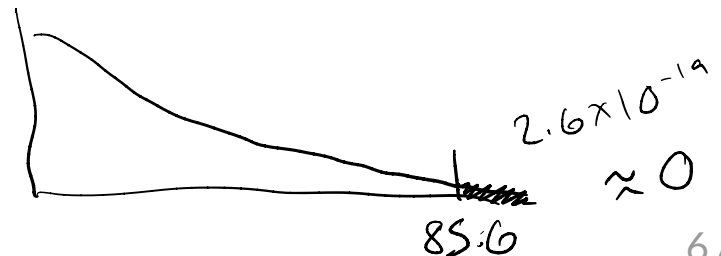
upper quantile

$W_n = 85.6 > \chi^2_{2,0.05}$
 \Rightarrow reject $H_0!$

p-value

```
pchisq(test_stat, df=2, lower.tail=F)
```

```
##           [,1]
## [1,] 2.577787e-19
```



A different question

We have the model

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 \text{Sex}_i + \beta_2 \text{Age}_i + \beta_3 \text{SecondClass}_i + \beta_4 \text{ThirdClass}_i$$

We want to test whether there is a difference in the chance of survival for second and third class passengers, holding age and sex fixed.

What hypotheses should we test?

$$H_0: \beta_3 = \beta_4$$

$$\beta_4 - \beta_3 = 0$$

$$H_A: \beta_3 \neq \beta_4$$

$$\beta_4 - \beta_3 \neq 0$$

idea: choose C st $C\beta = \beta_4 - \beta_3$

Contrasts

Class activity, Part II

https://sta711-s23.github.io/class_activities/ca_lecture_18.html

Class activity

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

```
##           [,1]
## [1,] -5.207289
```

Class activity

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

```
##           [,1]
## [1,] -5.207289
```

```
# rejection region for alpha = 0.05
qnorm(0.025, lower.tail=F)
```

```
## [1] 1.959964
```

```
# p-value
2*pnorm(abs(test_stat), lower.tail=F)
```

```
##           [,1]
## [1,] 1.916191e-07
```

A two-sample test for a difference in means