

# Hypothesis testing framework

$$H_0: \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad \leftarrow \chi^2_{\text{length of part of } \beta \text{ that we're testing}}$$

## Last time

We have the model

$$Y_i \sim Bernoulli(p_i)$$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 Sex_i + \beta_2 Age_i + \beta_3 SecondClass_i + \beta_4 ThirdClass_i$$

We want to test whether there is a difference in the chance of survival for second and third class passengers, holding age and sex fixed.

What hypotheses should we test?

$$H_0: \beta_3 = \beta_4$$

$$\beta_4 - \beta_3 = 0$$

$$H_A: \beta_3 \neq \beta_4$$

$$\beta_4 - \beta_3 \neq 0$$

$$\text{Want: } C \text{ st } CB = \beta_4 - \beta_3$$

## Contrasts

$$\beta_4 - \beta_3 = \underbrace{\begin{bmatrix} 0 & 0 & 0 & -1 & 1 \end{bmatrix}}_{a^T} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_4 \end{bmatrix} = a^T \beta$$
$$= c \beta$$
$$c = a^T$$

Def: Let  $\theta \in \mathbb{R}^q$  and  $a \in \mathbb{R}^q$  such that  $\sum_i a_i = 0$   
Then  $a^T \theta$  is called contrast

$$H_0: a^T \theta = \theta_0$$

$$H_A: a^T \theta \neq \theta_0$$

$$\text{ex: } H_0: a^T \beta = 0$$

$$H_A: a^T \beta \neq 0$$

Test:  $Z_n = \frac{(a^T \Sigma^{-1}(\beta) a)^{-\frac{1}{2}} (a^T \hat{\beta} - 0)}{\sqrt{a^T \Sigma^{-1}(\beta) a}}$

$\approx N(0, 1)$  under  $H_0$

reject  $H_0$  when  $|Z_n| > z_{\alpha/2}$  (or  $Z_n^2 > \chi^2_{1, \alpha}$ )

# Class activity

Work on Part II from last class:

[https://sta711-s23.github.io/class\\_activities/ca\\_lecture\\_18.html](https://sta711-s23.github.io/class_activities/ca_lecture_18.html)

## Class activity

$$\mathcal{Z}^{-1}(\beta)$$

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

$\hat{\beta}$

$a^\top \hat{\beta}$

$\sqrt{a^\top \mathcal{Z}^{-1}(\beta) a}$

```
##          [,1]
## [1,] -5.207289
```

$$z_n = -5.21$$

$$\alpha = 0.05 \Rightarrow \text{reject} \quad \text{when} \quad |z_n| > z_{\frac{\alpha}{2}} = \underbrace{z_{0.025}}_{1.96}$$

# Class activity

```
a <- c(0, 0, 0, -1, 1)
test_stat <- (t(a) %*% coef(m1))/sqrt(t(a) %*% vcov(m1) %*% a)
test_stat
```

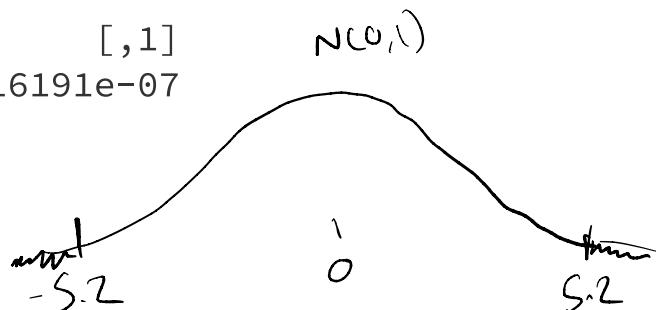
```
## [1] -5.207289
```

```
# rejection region for alpha = 0.05
qnorm(0.025, lower.tail=F)
```

```
## [1] 1.959964
```

```
# p-value
2*pnorm(abs(test_stat), lower.tail=F)
```

```
## [1] 1.916191e-07
```



Logistic regression MLE:  $\text{Var}(\hat{\beta}) = V_n = I^{-1}(\beta)$

## Summary of Wald tests

Let  $\theta \in \mathbb{R}^p$  be some parameter of interest. We wish to test the hypotheses

$$H_0 : C\theta = \gamma_0 \quad H_A : C\theta \neq \gamma_0 \quad \text{Specify hypothesis}$$

for some  $C \in \mathbb{R}^{q \times p}$ . Given an estimator  $\hat{\theta}_n$  such that

$$V_n^{-\frac{1}{2}}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, I),$$

the Wald test rejects when

$$(C\hat{\theta}_n - \gamma_0)^T (CV_n C^T)^{-1} (C\hat{\theta}_n - \gamma_0) > \chi_{q,\alpha}^2$$

calculate a test statistic

determine when to reject  $H_0$

# General framework for hypothesis tests

**Definition:** Let  $\theta \in \mathbb{R}^p$  be some parameter of interest. A **hypothesis** is a statement about  $\theta$ . The goal of a hypothesis test is to compare two competing hypotheses:

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

- here  $\Theta_0 \cap \Theta_1 = \emptyset$
- If  $\Theta_0 = \{\theta_0\}$ ,  $H_0$  is a simple hypothesis  
otherwise,  $H_0$  is a composite hypothesis (likewise for  $H_A$ )

Example :  $H_0 : \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$        $H_A : \begin{bmatrix} \beta_3 \\ \beta_4 \end{bmatrix} \neq 0$

↑  
simple  
null

↑  
composite  
alternative

# Outcomes

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

The outcome of the test is a decision to either **reject  $H_0$**  or **fail to reject  $H_0$** .

Four possibilities

Decision

Fail to reject

Reject

Truth

$H_0$  is true

$H_A$  is true

	‘ <u>ay!</u>	Type II error (false negative)
Type I error (false positive)		‘ <u>ay!</u>

- Goal:
- 1) control type I error
  - 2) minimize type II error

# Constructing a test

$$H_0 : \theta \in \Theta_0 \quad H_A : \theta \in \Theta_1$$

Observe data  $x_1, \dots, x_n$

- ① Calculate a test statistic  $T_n = \bar{T}(x_1, \dots, x_n)$
- ② Choose a rejection region  $R = \{(x_1, \dots, x_n) : \text{reject } H_0\}$
- ③ Reject  $H_0$  if  $(x_1, \dots, x_n) \in R$

Example :  $x_1, \dots, x_n$  iid from population w/ mean  $\mu$ ,  
variance  $\sigma^2$ .  $H_0: \mu = \mu_0$      $H_A: \mu \neq \mu_0$

①  $T(x_1, \dots, x_n) = \frac{\bar{x}_n - \mu_0}{\sigma/\sqrt{n}} = Z_n$

②  $R = \{(x_1, \dots, x_n) : |\frac{\bar{x}_n - \mu_0}{\sigma/\sqrt{n}}| > z_{\frac{\alpha}{2}}\}$

③ Reject  $H_0$  if  $(x_1, \dots, x_n) \in R$ , i.e. if  $|Z_n| > z_{\frac{\alpha}{2}}$

Prob. of type I error =  $P(\text{reject } H_0 \mid H_0 \text{ is true})$

Under  $H_0$ ,  $Z_n \sim N(0,1) \Rightarrow \text{Prob. of type I error} = P(|Z| > z_{\frac{\alpha}{2}}) = \alpha$

## Power function

Suppose we reject  $H_0$  when  $(X_1, \dots, X_n) \in R$ . The **power function**  $\beta(\theta)$  is

$$\beta(\theta) = P_\theta((X_1, \dots, X_n) \in R)$$

## Example

$X_1, \dots, X_n$  iid from a population with mean  $\mu$  and variance  $\sigma^2$ .

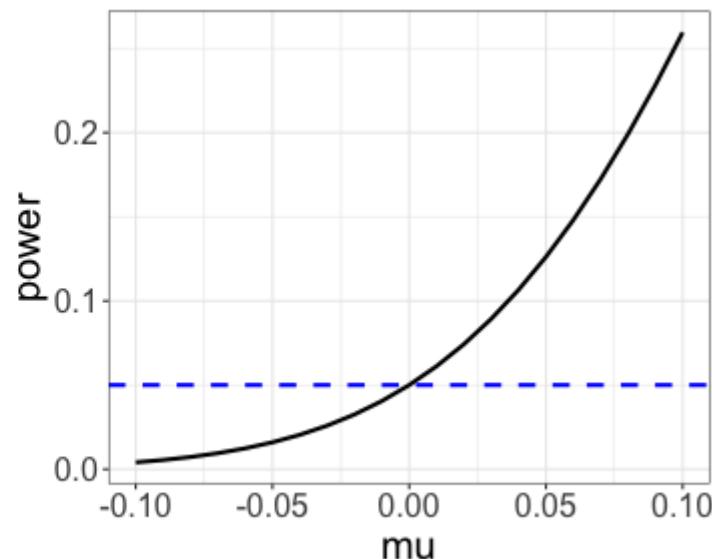
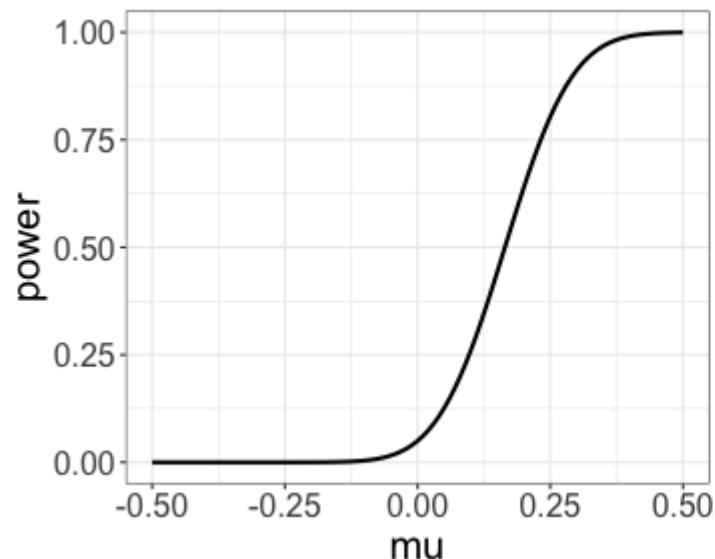
$$H_0 : \mu = \mu_0 \quad H_A : \mu > \mu_0$$

## Class activity

$$\beta(\mu) \approx 1 - \Phi \left( z_\alpha - \frac{(\mu - \mu_0)}{\sigma/\sqrt{n}} \right)$$

- + Suppose that  $\mu_0 = 0$ ,  $n = 100$ , and  $\sigma = 1$ . Make a plot of  $\beta(\mu)$  vs.  $\mu$  for  $\alpha = 0.05$ .
- + Now consider testing  $H_0 : \mu \leq \mu_0$  vs.  $H_A : \mu > \mu_0$ . Will this change our rejection region if we want a size  $\alpha$  test?

# Class activity



# Simulation verification

```
n <- 100
sigma <- 1
mu0 <- 0
nreps <- 5000
test_stats <- rep(0, nreps)
for(i in 1:nreps){
  x <- rnorm(n, mu0, sigma)
  test_stats[i] <- (mean(x) - mu0)/(sigma/sqrt(n))
}
mean(test_stats > qnorm(0.05, lower.tail=F))

## [1] 0.0548
```

# Simulation verification

```
n <- 100
sigma <- 1
mu0 <- 0
nreps <- 5000
test_stats <- rep(0, nreps)
for(i in 1:nreps){
  x <- rnorm(n, 0.1, sigma)
  test_stats[i] <- (mean(x) - mu0)/(sigma/sqrt(n))
}
mean(test_stats > qnorm(0.05, lower.tail=F))

## [1] 0.2646
```