

Interval estimation

Motivation

Suppose we have data $(X_1, Y_1), \dots, (X_n, Y_n)$ with

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1 - p_i}\right) = \beta^T X_i$$

So far, we have discussed:

- + Finding point estimates $\hat{\beta}$
- + Testing hypotheses about the true (but unknown) parameters β

What are the limitations of point estimates and hypothesis tests for inference about β ?

Confidence interval

```
...  
##                Estimate Std. Error z value Pr(>|z|)  
## (Intercept) 2.6415063 0.1213233 21.77 <2e-16 ***  
## WBC         -0.2892904 0.0134349 -21.53 <2e-16 ***  
## PLT         -0.0065615 0.0005932 -11.06 <2e-16 ***  
## ---  
...  
...
```

How would I calculate a 95% confidence interval for β_1 (the change in the log odds of dengue for a one-unit increase in WBC, holding PLT fixed)?

$$\hat{\beta}_1 \pm \frac{z_{\alpha/2} \text{SE}(\hat{\beta}_1)}{\sqrt{2}} \leftarrow (-\alpha \text{ Wald CI}\right)$$

95% CI : $-0.289 \pm 1.96 (0.0134)$
 $= (-0.315, -0.262)$

β_1 is either in $(-0.315, -0.262)$ or not
(no "95% probability" for a specific interval)

Confidence interval

...

	Estimate	Std. Error	z value	Pr(> z)	
## (Intercept)	2.6415063	0.1213233	21.77	<2e-16	***
## WBC	-0.2892904	0.0134349	-21.53	<2e-16	***
## PLT	-0.0065615	0.0005932	-11.06	<2e-16	***
## ---					

... coverage probability

95% confidence interval for β_1 : $(-0.315, -0.262)$

How do I interpret this confidence interval?

95% confident : if we take many samples and we calculate many intervals, 95% should contain the true (unknown) parameter

"we are 95% confident that a one-unit increase in wbc is associated w/ a decrease in log odds of dengue by between 0.262 and 0.315" 4/9

$$\hat{\theta} \sim N(\theta, \text{var}(\hat{\theta})) \Rightarrow \frac{\hat{\theta} - \theta}{\text{SE}(\hat{\theta})} \sim N(0, 1)$$

Deriving the coverage probability

($1-\alpha$) Wald interval: $\hat{\theta} \pm z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})$

$$\hookrightarrow P(\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \theta \leq \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})) = 1 - \alpha$$

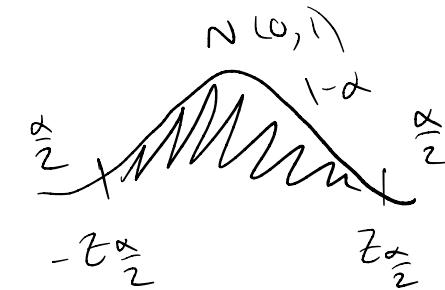
\hookrightarrow endpoints are random (function of data)

$$P(\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \theta \leq \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}))$$

$$= P(-z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}) \leq \hat{\theta} - \theta \leq z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}))$$

$$= P(-z_{\frac{\alpha}{2}} \leq \frac{\hat{\theta} - \theta}{\text{SE}(\hat{\theta})} \leq z_{\frac{\alpha}{2}}) = 1 - \alpha$$

$\sim N(0, 1)$



Note: $\theta_0 \in [\hat{\theta} - z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta}), \hat{\theta} + z_{\frac{\alpha}{2}} \text{SE}(\hat{\theta})]$

$\hookrightarrow \left| \frac{\hat{\theta} - \theta_0}{\text{SE}(\hat{\theta})} \right| \leq z_{\frac{\alpha}{2}}$ i.e. the α -level Wald test of $H_0: \theta = \theta_0$ vs. $H_A: \theta \neq \theta_0$ fails to reject 5/9

Formal definition

Let $\theta \in \mathbb{H}$ be a parameter of interest, and x_1, \dots, x_n a sample. Let $C(x_1, \dots, x_n) \subseteq \mathbb{H}$ be a set constructed from x_1, \dots, x_n ($\Rightarrow C(x_1, \dots, x_n)$ is a random set). $C(x_1, \dots, x_n)$ is a $1-\alpha$ confidence set for θ if

$$\inf_{\theta \in \mathbb{H}} P_\theta (\theta \in C(x_1, \dots, x_n)) = 1 - \alpha$$

(i.e. $\forall \theta \in \mathbb{H}, P_\theta (\theta \in C(x_1, \dots, x_n)) \geq 1 - \alpha$)

α -level test of $H_0: \Theta = \Theta_0$: $P_{\Theta_0}((X_1, \dots, X_n) \in R(\Theta_0)) \leq \alpha$

Inverting a test

↑
rejection
region

Theorem: Let $\Theta \in \mathbb{H}$ be a parameter of interest.

For each value of $\Theta_0 \in \mathbb{H}$, consider testing $H_0: \Theta = \Theta_0$ vs. $H_A: \Theta \neq \Theta_0$, and let $R(\Theta_0)$ be the rejection region for a level α test.

Let $C(X_1, \dots, X_n) = \{\Theta_0 : (X_1, \dots, X_n) \notin R(\Theta_0)\}$

Then $C(X_1, \dots, X_n)$ is a $1-\alpha$ confidence set for Θ .

Pf: $\Theta \in C(X_1, \dots, X_n) \Leftrightarrow (X_1, \dots, X_n) \notin R(\Theta)$

$$\begin{aligned}\Rightarrow P_\Theta(\Theta \in C(X_1, \dots, X_n)) &= P_\Theta((X_1, \dots, X_n) \notin R(\Theta)) \\ &= 1 - P_\Theta((X_1, \dots, X_n) \in R(\Theta)) \\ &\geq 1 - \alpha \quad (\alpha\text{-level test})\end{aligned}$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} Uniform[0, \theta]$. We want to test

$$H_0 : \theta = \theta_0 \quad H_A : \theta \neq \theta_0$$

Find the LRT statistic for this test.

Reject H_0 when

$$\frac{\sup_{\theta} L(\theta | X)}{L(\theta_0 | X)} > k$$

$$L(\theta | X) = \left(\frac{1}{\theta}\right)^n \mathbb{1}\{X_{(n)} \leq \theta\} \quad \hat{\theta}_{MLE} = X_{(n)}$$

\Rightarrow reject H_0 when

$$\frac{\theta_0^n}{X_{(n)}^n \mathbb{1}\{\theta_0 \geq X_{(n)}\}} > k$$

$$\Rightarrow \text{reject } H_0 \text{ when } \theta_0 < X_{(n)} \text{ or when } \frac{\theta_0}{X_{(n)}} > k^{\frac{1}{n}}$$

reject H_0 when $\theta_0 < x_{(n)}$ or $\frac{\theta_0}{x_{(n)}} > K^{\frac{1}{n}}$

\Rightarrow fail to reject H_0 when $x_{(n)} \leq \theta_0 \leq x_{(n)} K^{\frac{1}{n}}$

confidence set = $[x_{(n)}, x_{(n)} K^{\frac{1}{n}}] = [x_{(n)}, x_{(n)} K']$

so need K' st $P_\theta (\theta \in [x_{(n)}, x_{(n)} K']) \geq 1 - \alpha$

$$\begin{aligned}
 P_\theta (\theta \in [x_{(n)}, x_{(n)} K']) &= P_\theta (x_{(n)} K' \geq \theta) \\
 &= 1 - P_\theta (x_{(n)} K' < \theta) \\
 &= 1 - P_\theta (x_{(n)} < \frac{\theta}{K'}) \\
 &= 1 - \left(\frac{\theta/K'}{\theta} \right)^n = 1 - \left(\frac{1}{K'} \right)^n
 \end{aligned}$$

\Rightarrow choose K' st $\left(\frac{1}{K'}\right)^n = \alpha$ for $\theta: [x_{(n)}, \frac{x_{(n)}}{\alpha^{1/n}}]$