

Fisher information

Recap: Newton's method

To find β^* such that $U(\beta^*) = 0$, when there is no closed-form solution we use Newton's method:

- + Begin with an initial guess $\beta^{(0)}$
- + Iteratively update: $\beta^{(r+1)} = \beta^{(r)} - \mathbf{H}^{-1}(\beta^{(r)})U(\beta^{(r)})$
- + Stop when the algorithm converges

For logistic regression: $U(\beta) = \mathbf{x}^T(\mathbf{y} - \mathbf{p})$

$$\mathbf{H}(\beta) = -\mathbf{x}^T \mathbf{w} \mathbf{x}$$

$$\mathbf{w} = \text{diag}(\mathbf{p}(1-\mathbf{p}))$$

Some intuition about Hessians

Example: Suppose that $\beta = (\beta_0, \beta_1)^T \in \mathbb{R}^2$, and

$$\ell(\beta) = -\beta_0^2 - 100\beta_1^2$$

Calculate the score function

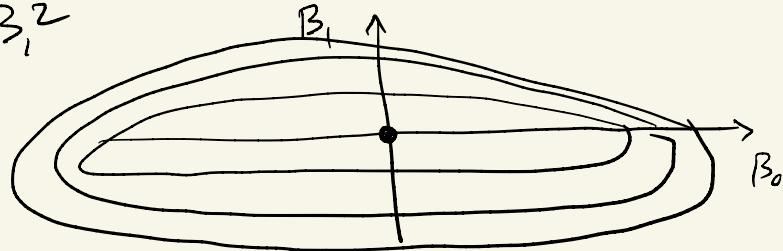
$$U(\beta) = \begin{bmatrix} \frac{\partial \ell}{\partial \beta_0} \\ \frac{\partial \ell}{\partial \beta_1} \end{bmatrix} = \begin{bmatrix} -2\beta_0 \\ -200\beta_1 \end{bmatrix}$$

and the Hessian

$$\mathbf{H}(\beta) = \begin{bmatrix} \frac{\partial^2 \ell}{\partial \beta_0^2} & \frac{\partial^2 \ell}{\partial \beta_0 \partial \beta_1} \\ \frac{\partial^2 \ell}{\partial \beta_1 \partial \beta_0} & \frac{\partial^2 \ell}{\partial \beta_1^2} \end{bmatrix} = \begin{bmatrix} -2 & 0 \\ 0 & -200 \end{bmatrix}$$

$$l(\beta) = -\beta_0^2 - 100\beta_1^2$$

$$H(\beta) = - \begin{bmatrix} 2 & 0 \\ 0 & 200 \end{bmatrix}$$



captures information
about curvature of $l(\beta)$

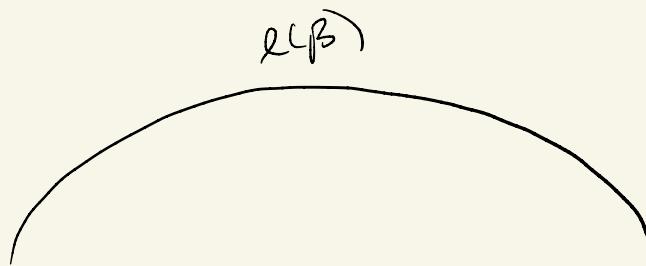
$$H^{(r)}(\beta) = - \begin{bmatrix} \frac{1}{2} & 0 \\ 0 & \frac{1}{200} \end{bmatrix}$$

depends on curvature of $l(\beta)$

$$\begin{aligned} \beta^{(r+1)} &= \beta^{(r)} - \underbrace{H^{(r)}(\beta^{(r)})}_{\text{more along gradient}} \underbrace{U(\beta^{(r)})}_{\text{U}} \\ &= \begin{bmatrix} 0.1 \\ 0.1 \end{bmatrix} + \begin{bmatrix} \frac{1}{2} & 0 \\ 0 & \frac{1}{200} \end{bmatrix} \begin{bmatrix} -2(0.1) \\ -200(0.1) \end{bmatrix} \end{aligned}$$

More intuition

which $\ell(\beta)$ is "easier" to maximize?



Harder to maximize
(lots of β s have similar $\ell(\beta)$)

second derivative is close to 0

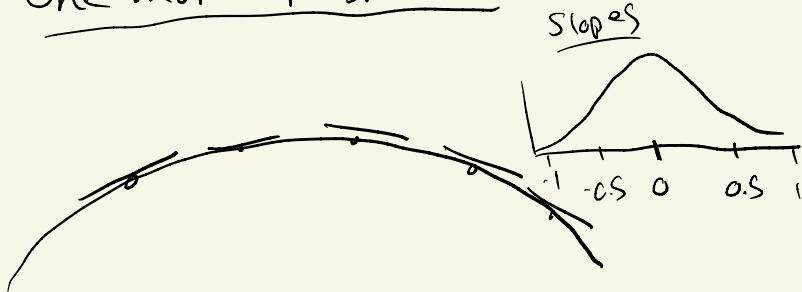


easier to maximize
(small changes in $\beta \Rightarrow$
large changes in $\ell(\beta)$)

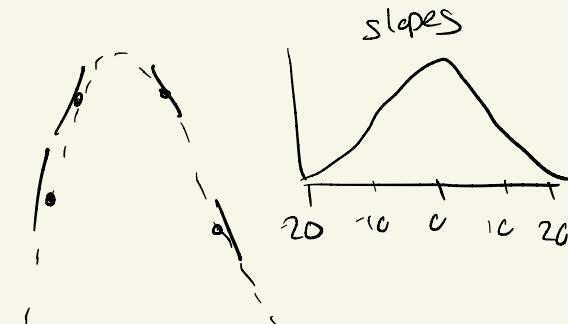
second derivative far from 0

\Rightarrow Hessian tells us how easy function is to maximize

One more perspective:



Slopes $u(\beta)$ are close to 0



slopes $u(\beta)$ are far from 0

more variability $u(\beta)$

\Rightarrow look @ variability in $u(\beta)$. Larger variance \Rightarrow easier to maximize

Logistic regression:

$$u(\beta) = X^T(Y - p)$$

$$\text{var}(u(\beta)) = X^T \text{Var}(Y - p) X$$

$$= X^T \text{Var}(Y) X$$

$$= X^T \begin{bmatrix} p_1(1-p_1) & & \\ & p_2(1-p_2) & \\ & & \ddots p_n(1-p_n) \end{bmatrix} X$$

$$= X^T W X = -H(\beta)$$

$$\text{Var}(Y_i) = p_i(1-p_i)$$

Fisher information

Def : Let $\ell(\theta | Y)$ be a log likelihood , and $U(\theta) = \frac{\partial \ell}{\partial \theta}$ the score function. The Fisher information is

$$I(\theta) = \text{Var}(U(\theta) | \theta)$$

i.e. the variance of the score, if θ is the true parameter

$$Y_i \sim \text{Bernoulli}(\rho_i)$$

$$\log\left(\frac{\rho_i}{1-\rho_i}\right) = \beta^T x_i$$

$$\begin{aligned} I(\beta) &= \text{Var}(U(\beta)) \\ &= \text{Var}(X^T(Y - \rho)) \\ &= X^T W X \end{aligned}$$

$$W = \text{diag}(\rho_i(1-\rho_i))$$

$$\text{Var}(Y_i) = p(1-p)$$

$$\text{Var}\left(\sum_{i=1}^n Y_i\right) = np(1-p)$$

(independent)

Example: Bernoulli sample

Suppose that $Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$.

$$L(p|Y) = p^{\sum_i Y_i} (1-p)^{n - \sum_i Y_i}$$

$$l(p|Y) = \left(\sum_{i=1}^n Y_i\right) \log p + (n - \sum_{i=1}^n Y_i) \log(1-p)$$

$$u(p) = \frac{\sum_i Y_i}{p} - \frac{(n - \sum_i Y_i)}{1-p}$$

$$\text{Var}(u(p)|p) = \text{Var}\left(\frac{\sum_i Y_i}{p} - \frac{(n - \sum_i Y_i)}{1-p}\right)$$

$$= \text{Var}\left(\frac{(\sum_i Y_i)(1-p)}{p(1-p)} - \frac{(n - \sum_i Y_i)p}{p(1-p)}\right) = \text{Var}\left(\frac{\sum_i Y_i}{p(1-p)}\right)$$

$$= \frac{n p (1-p)}{p^2 (1-p)^2} = \frac{n}{p (1-p)}$$

$\hat{Z}(p) = \frac{n}{p(1-p)}$

Properties

Under certain regularity conditions, we have the following:

$$\textcircled{1} \quad \mathbb{E}(u(\theta) | \theta) = 0$$

$$\begin{aligned} \text{(this implies that } \chi(\theta) &= \text{Var}(u(\theta) | \theta) \\ &= \mathbb{E}(u^2(\theta) | \theta) \quad) \end{aligned}$$

$$\textcircled{2} \quad \hat{\chi}(\theta) = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \ell(\theta | Y) | \theta\right]$$

(Fisher information captures curvature of log likelihood)

Example: Bernoulli sample

$$\mathbb{E}[\sum_i Y_i] = \\ n \mathbb{E}[Y_i] = np$$

Suppose that $Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$.

$$u(p) = \frac{\sum_i Y_i}{p} - \frac{(n - \sum_i Y_i)}{1-p}$$

$$1) \mathbb{E}[u(p)|p] = \mathbb{E}\left[\frac{\sum_i Y_i}{p}\right] - \mathbb{E}\left[\frac{n - \sum_i Y_i}{1-p}\right] \\ = \frac{np}{p} - \frac{n(1-p)}{1-p} = n - n = 0 \quad \checkmark$$

$$2) \frac{\partial^2 \ell}{\partial p^2} = -\frac{\sum_i Y_i}{p^2} - \frac{(n - \sum_i Y_i)}{(1-p)^2}$$

$$-\mathbb{E}\left[\frac{\partial^2 \ell}{\partial p^2}\right] = \mathbb{E}\left[\frac{\sum_i Y_i}{p^2}\right] + \mathbb{E}\left[\frac{(n - \sum_i Y_i)}{(1-p)^2}\right] \quad \checkmark$$

$$= \frac{np}{p^2} + \frac{n(1-p)}{(1-p)^2} = \frac{n}{p} + \frac{n}{1-p} = p \frac{1}{p(1-p)} \quad \checkmark$$