

Sufficiency and minimal sufficiency

Where we are and where we're going

So far:

- We know how to calculate MLEs
 - We know that in terms of MSE, MLEs are asymptotically the "best" we can do (under regularity conditions)
 - asymptotically unbiased
 - asymptotic variance is CRLB
 - We also know how to Rao-Blackwellize an unbiased estimator to improve MSE
 - condition on sufficient statistics
 - sufficient statistics contain all the info we need to calculate MLEs
- Currently:
- how do we find sufficient statistics?

Up next:

- Are there more "compact" sufficient statistics?
- When do sufficient statistics give us a best unbiased estimator?

Recap: factorization theorem

Let x_1, \dots, x_n be a sample with joint probability function $f(x_1, \dots, x_n | \theta)$. A statistic $T = T(x_1, \dots, x_n)$ is sufficient for θ if and only if there exist function $g(T|\theta)$ and $h(x_1, \dots, x_n)$ such that, for all possible x_1, \dots, x_n and all possible θ ,

$$f(x_1, \dots, x_n | \theta) = g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)$$

joint distribution only depends on θ through the sufficient statistic

Pf: (we will prove for discrete distributions)

① Suppose the functions $g \notin h$ exist. WTS T is sufficient.

$$\begin{aligned} f(x_1, \dots, x_n | T, \theta) &= \frac{f(x_1, \dots, x_n, T(x_1, \dots, x_n) | \theta)}{f(T(x_1, \dots, x_n) | \theta)} \\ &= \frac{f(x_1, \dots, x_n | \theta)}{f(T(x_1, \dots, x_n) | \theta)} = \frac{g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)}{\sum_{y_1, \dots, y_n} f(y_1, \dots, y_n | \theta)} \\ &\quad \text{if } y_1, \dots, y_n = T(x_1, \dots, x_n) \end{aligned}$$

$$\begin{aligned}
 & \frac{g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)}{\sum_{y_1, \dots, y_n} f(y_1, \dots, y_n | \theta)} = \frac{g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)}{\sum_{y_1, \dots, y_n} g(T(y_1, \dots, y_n) | \theta) h(y_1, \dots, y_n)} \\
 & \text{y}_1, \dots, \text{y}_n \\
 & T(\text{y}_1, \dots, \text{y}_n) = T(x_1, \dots, x_n) \\
 \\
 & = \frac{g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)}{g(T(x_1, \dots, x_n) | \theta) \sum_{y_1, \dots, y_n} h(y_1, \dots, y_n)} = \frac{h(x_1, \dots, x_n)}{\sum_{y_1, \dots, y_n} h(y_1, \dots, y_n)} \quad (\text{does not depend on } \theta) \\
 & \text{y}_1, \dots, \text{y}_n \\
 & T(\text{y}_1, \dots, \text{y}_n) = T(x_1, \dots, x_n)
 \end{aligned}$$

✓

② Suppose T is a sufficient statistic.

$$\text{Let } g(T | \theta) = P_\theta(T(X_1, \dots, X_n) = t) \quad (\text{does not depend on } \theta)$$

$$h(x_1, \dots, x_n) = P(X_1, \dots, X_n = x_1, \dots, x_n | T(X_1, \dots, X_n) = T(x_1, \dots, x_n))$$

$$\begin{aligned}
 f(x_1, \dots, x_n | \theta) &= P_\theta(X_1, \dots, X_n = x_1, \dots, x_n) \\
 &= P_\theta(X_1, \dots, X_n = x_1, \dots, x_n \text{ and } T(X_1, \dots, X_n) = T(x_1, \dots, x_n)) \\
 &= P_\theta(T(X_1, \dots, X_n) = T(x_1, \dots, x_n)) P(X_1, \dots, X_n = x_1, \dots, x_n | T(X_1, \dots, X_n) = T(x_1, \dots, x_n)) \\
 &= g(T(x_1, \dots, x_n) | \theta) h(x_1, \dots, x_n)
 \end{aligned}$$

//

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ with μ and σ^2 unknown. Let $\theta = (\mu, \sigma^2)$.

Use the factorization theorem to find a sufficient statistic for θ .
(The dimension of the statistic will be greater than 1).

$$T = (\bar{X}, S)$$

$$T = (\bar{X}, S^2)$$

$$T = (\sum_i X_i, \sum_i X_i^2)$$

$x_1, \dots, x_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

$$\begin{aligned} & \sum_i (x_i^2 - 2\mu x_i + \mu^2) \\ &= \sum_i x_i^2 - 2\mu \sum_i x_i + n\mu^2 \end{aligned}$$

$$\begin{aligned} f(x_1, \dots, x_n | \mu, \sigma^2) &= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left\{ -\frac{1}{2\sigma^2} \sum_i (x_i - \mu)^2 \right\} \\ &= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left\{ -\frac{1}{2\sigma^2} \left(\sum_i x_i^2 - 2\mu \sum_i x_i + n\mu^2 \right) \right\} \\ &= g(\sum_i x_i, \sum_i x_i^2 | \mu, \sigma^2) \\ &= g(\sum_i x_i, \sum_i x_i^2 | \mu, \sigma^2) \underbrace{h(x_1, \dots, x_n)}_{=1} \end{aligned}$$

Sufficient statistic = $(\sum_i x_i, \sum_i x_i^2)$

also do: $(\frac{1}{n} \sum_i x_i, \frac{1}{n-1} \sum_i (x_i - \bar{x})^2)$

$$\sum_i (x_i - \bar{x})^2 = \sum_i x_i^2 - n(\bar{x})^2$$

Is the dimension of a sufficient statistic always = dimension of Θ ?

NO : (a) $(x_1, x_2, x_3, \dots, x_n)$ is sufficient

(b) Ex: $x_1, \dots, x_n \stackrel{iid}{\sim} \text{Uniform}[\theta, \theta+1]$

minimal sufficient statistic : $(x_{(1)}, x_{(n)})$

Sufficient statistics as partitions

Suppose x_1, x_2, x_3 \sim Bernoulli(p), we know $T = x_1 + x_2 + x_3$ is sufficient

(x_1, x_2, x_3)	$T(x_1, x_2, x_3)$	$P(x_1, x_2, x_3 T)$	
$(0, 0, 0)$	0	1	\nwarrow
$(1, 0, 0)$	1	$\frac{1}{3}$	
$(0, 1, 0)$	1	$\frac{1}{3}$	
$(0, 0, 1)$	1	$\frac{1}{3}$	
$(1, 1, 0)$	2	$\frac{1}{3}$	
$(1, 0, 1)$	2	$\frac{1}{3}$	\swarrow
$(0, 1, 1)$	2	$\frac{1}{3}$	
$(1, 1, 1)$	3	1	\downarrow

$T(x_1, x_2, x_3)$ partitions space (x_1, x_2, x_3) into sets B_1, B_2, B_3, B_4
 $B_1 = \{(x_1, x_2, x_3) : T=0\}, B_2 = \{(x_1, x_2, x_3) : T=1\},$ etc.

T is sufficient $\Leftrightarrow f((x_1, x_2, x_3)) | (x_1, x_2, x_3) \in B_i)$ does not depend on $p^{5/6}$

Another partition

$x_1, x_2, x_3 \stackrel{\text{ iid }}{\sim} \text{Bernoulli}(p)$

$T = (x_1, x_2, x_3)$

(x_1, x_2, x_3)	$T(x_1, x_2, x_3)$	$P(x_1, x_2, x_3 T)$
$(0, 0, c)$	$(0, 0, c)$	1
$(1, 0, c)$	$(1, 0, c)$	1
$(0, 1, c)$	$(0, 1, c)$	1
		\vdots
		etc.

T partitions sample space into 8 graphs (one per
 (x_1, x_2, x_3))

Question: which statistic more efficiently summarizes the data?
(i.e., results in fewer partitions?)

Minimal sufficient statistics

Def: A statistic $T(X_1, \dots, X_n)$ is a minimal sufficient statistic if for any other sufficient statistic $T^*(X_1, \dots, X_n)$, $T(X_1, \dots, X_n)$ is a function of $T^*(X_1, \dots, X_n)$

Ex: $X_1, X_2, X_3 \stackrel{iid}{\sim} \text{Bernoulli}(\rho)$

$$T = \sum_i X_i \quad T^* = (X_1, X_2, X_3)$$

T is a function of T^* , but T^* is not a function of T