Lecture 9: Fisher information and Fisher scoring

imagine U(B) = XT(Y-p) Recap: Fisher information

Pi= eptx: For agiven B, Heptx: UB) will very L(B/Y) be a log-livelihood from deteset Def. and $U(\theta) = \frac{\partial l}{\partial \theta}$. The Fisher information is 2(0) = Var (U(0) (0)

> depends on doserred time, In (and so U(O) is a random variable)

Lost time: Storted to see how I-1(0) is related to variance of MLE ô

Properties

Under appropriate regularity conditions,

$$\mathbb{E}[\mathbf{U}(\theta)|\theta] = 0$$

Proof:
$$E[u(0)|0] = E[\frac{2}{20}l(0|1)|0] = E[\frac{2}{20}logf(1|0)|0]$$

$$= E[\frac{1}{20}logf(1|0)|0] = [\frac{2}{20}f(y|0)] \cdot \frac{1}{20}logf(1|0)|0]$$

$$= [\frac{1}{20}f(y|0)]dy$$

= $\frac{2}{20}f(y|0)dy$

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(See Casella & Berger 2.4)

Properties
$$-E\left[\frac{3^2}{56^2}f(x_16)\right] = -\frac{1}{5}$$

$$\int = -\int \left(\frac{3^2}{3G} f(16)\right) \cdot \frac{1}{f(16)} \cdot \frac{1}{f(16)}$$

Properties

Under appropriate regularity conditions,

$$\frac{\partial^2}{\partial \theta^2} \ell(\theta | \mathbf{Y}) | \theta$$

Figure 1. So that the second in the se

Proof:
$$-\mathbb{E}\left[\frac{3^2}{36^2} \mathcal{L}(\Theta|Y) \mid \Theta\right] = -\mathbb{E}\left[\frac{3^2}{36^2} \mid og f(Y|Y) \mid \Theta\right]$$

$$= - \mathbb{E} \left[\frac{3^{2}}{30^{2}} \frac{f(10)}{f(10)} - \left(\frac{3}{30} \frac{f(10)}{f(10)} \right)^{2} | \theta \right]$$

$$= - \mathbb{E} \left[\frac{3^{2}}{30^{2}} \frac{f(10)}{f(10)} | \theta \right] + \mathbb{E} \left[\left(\frac{3}{30} \log \frac{f(10)}{f(10)} \right)^{2} | \theta \right]$$

$$= \mathbb{E} \left[(u(0))^{2} | \theta \right] = V_{cr}(u(0) | \theta)_{H}$$

$$- \mathbb{E} \left[\left(\frac{3}{36} \log f(Y|B) \right)^2 |B| \right]$$

Fisher scoring

Newton's method: B(r) = B(r) - H-1(B(r))U(B(r))

replaced w/ I-1(BUI) under

regularity

1) Start w/ initial gress B(0)

3) Stop wer B(rti) ~ B(r)

(Heratively reneighted least squares)

IRLS for logistic regression

Logistic regression:
$$\beta^{(1)} = \beta^{(1)} + (\chi^T \psi^{(1)} \chi)^T \chi^T (Y-p^{(1)})$$

Recall: linear regression: $\hat{\beta} = (\chi^T \chi)^T \chi^T Y$

Rewrite: $\beta^{(1)} = (\chi^T \psi^{(1)} \chi)^T (\chi^T$

Weighted least squares

Suppose
$$\chi = \chi \beta + \xi$$

Suppose
$$\gamma = \lambda \beta + 2$$

$$w^{\frac{1}{2}} = w^{\frac{1}{2}} \times \beta + w^{\frac{1}{2}} \times \xi$$

$$= W^2 \times \beta^{-1} \times Z$$

$$= V + \beta + \beta + \beta$$

$$= 7 \hat{\beta} = (X_{\omega}^{T} X_{\omega})^{T} X_{\omega}^{T} Y_{\omega}$$

$$X^T W X)^{-1} X^T W Y$$

$$(X^T W X)^{-1} X^T W Y$$

$$\forall w = xw\beta + \xi w$$

$$V_{cr}(\xi_{i}) = \frac{1}{w_{i}}$$

$$W^{\frac{1}{2}} = \lambda i \left(\sqrt{w} \right)$$

2 ~ N(O, W-1)

W = diag (w,,..., w,)

usual assumption & NN(O, 0°I)

If I have non-constant variance:

$$\mathcal{E}_{w} \sim N(0, T)$$

$$V_{cr}(w^{\frac{1}{2}} \mathcal{E}) = w^{\frac{1}{2}} w^{-1} w^{\frac{1}{2}} = \overline{I}$$