

# Lecture 15: Asymptotic normality of the MLE

# Some more theorems about convergence

## Continuous mapping theorem:

random variables, and let

① If  $X_n \xrightarrow{d} X$  then

② If  $X_n \xrightarrow{P} X$  then

③ If  $X_n \xrightarrow{a.s.} X$  then

Let  $X_1, X_2, \dots$  be a sequence of  
 $g$  be a continuous function,

$$g(X_n) \xrightarrow{d} g(X)$$

$$g(X_n) \xrightarrow{P} g(X) \leftarrow \text{HW}$$

$$g(X_n) \xrightarrow{a.s.} g(X)$$

## Slutsky's theorem:

Let  $\{X_n\}$  and  $\{Y_n\}$  be sequences of  
 random variables, and suppose  
 $Y_n \xrightarrow{P} c$  ( $c$  is a constant). Then

that  $X_n \xrightarrow{d} X$ , and

$$X_n + Y_n \xrightarrow{d} X + c$$

$$X_n - Y_n \xrightarrow{d} X - c$$

$$\frac{X_n}{Y_n} \xrightarrow{d} \frac{X}{c}$$

(provided  $c$  is invertible)

# Convergence of the MLE

Suppose that  $\gamma_1, \gamma_2, \gamma_3, \dots$  are iid with probability function  $f(\gamma | \theta)$ ,  $\theta \in \mathbb{R}^d$

Let  $l_n(\theta) = \sum_{i=1}^n \log f(\gamma_i | \theta)$ , and  $\hat{\theta}_n$  be the MLE of  $\theta$  using first  $n$  observations. Let

$$\mathcal{I}_1(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(\gamma_i | \theta)\right]$$

Theorem: Under regularity,

$$(a) \quad \hat{\theta}_n \xrightarrow{P} \theta \quad \text{as } n \rightarrow \infty \quad (\text{consistency})$$

$$(b) \quad \sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, \mathcal{I}_1^{-1}(\theta)) \quad \text{as } n \rightarrow \infty$$

(asymptotic normality)

We will prove (b) when  $d=1$

Proof atline ( $\theta=1$ );

$$\textcircled{1} \quad \sqrt{n}(\hat{\theta}_n - \theta) \approx \frac{\frac{1}{\sqrt{n}} l'_n(\theta)}{-\frac{1}{n} l''_n(\theta)} \quad (\text{Taylor expansion})$$

$$\textcircled{2} \quad \frac{1}{n} l''_n(\theta) \xrightarrow{P} -\mathcal{I}_1(\theta) \quad (\text{WLLN})$$

$$\textcircled{3} \quad \frac{1}{\sqrt{n}} l'_n(\theta) \xrightarrow{d} N(0, \mathcal{I}_1(\theta)) \quad (\text{CLT})$$

$$\begin{aligned} \textcircled{4} \quad \text{Apply Slutsky's:} \\ \sqrt{n}(\hat{\theta}_n - \theta) &\xrightarrow{d} \frac{1}{\mathcal{I}_1(\theta)} N(0, \mathcal{I}_1(\theta)) \\ &= N(0, \mathcal{I}_1^{-1}(\theta)) \end{aligned}$$

$$X \sim N(\mu, \sigma^2)$$

$$aX \sim N(a\mu, a^2\sigma^2)$$

$$\text{CLT: } \sqrt{n} \frac{\bar{X}_n - \mu}{\sigma} \xrightarrow{d} N(0, 1)$$

$\bar{X}_n =$  mean of  $n$  iid observations

# Intermediate steps

Using the WLLN and the CLT, argue that:

$$\textcircled{2} \bullet \frac{1}{n} \ell''_n(\theta) \xrightarrow{p} -I_1(\theta)$$

$$\textcircled{3} \bullet \frac{1}{\sqrt{n}} \ell'_n(\theta) \xrightarrow{d} N(0, I_1(\theta))$$

Pf of  $\textcircled{2}$ :  $\ell_n(\theta) = \sum_{i=1}^n \log f(Y_i | \theta)$

$$\frac{1}{n} \ell''_n(\theta) = \frac{1}{n} \sum_{i=1}^n \frac{\partial^2}{\partial \theta^2} \log f(Y_i | \theta)$$

sample mean of  $\frac{\partial^2}{\partial \theta^2} \log f(Y_i | \theta)$

$$\mathbb{E} \left[ \frac{\partial^2}{\partial \theta^2} \log f(Y_i | \theta) \right] = -\hat{I}_1(\theta)$$

$\Rightarrow$  WLLN:  $\frac{1}{n} \ell''_n(\theta) \xrightarrow{p} -\hat{I}_1(\theta)$   
(or  $-\frac{1}{n} \ell''_n(\theta) \xrightarrow{p} \hat{I}_1(\theta)$ )

