

Lecture 5: Maximum likelihood estimation for logistic regression

Invariance of the MLE

Last time: $\gamma_1, \dots, \gamma_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (\gamma_i - \bar{\gamma})^2$$

Q: what if we want $\hat{\sigma}$?

$$\hat{\sigma} = \sqrt{\hat{\sigma}^2} = \sqrt{\frac{1}{n} \sum_{i=1}^n (\gamma_i - \bar{\gamma})^2}$$

Theorem (invariance of the MLE): (see Thm 7.2.10 in CB)

Let $\hat{\theta}$ be the MLE of θ . For any function $\gamma(\theta)$, the MLE of $\gamma(\theta)$ is $\gamma(\hat{\theta})$.

Logistic regression

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1 - p_i}\right) = \beta_0 + \beta_1 X_{i,1} + \cdots + \beta_k X_{i,k}$$

Suppose we observe independent samples $(X_1, Y_1), \dots, (X_n, Y_n)$. Write down the likelihood function

$(X_1, Y_1), \dots, (X_n, Y_n) \stackrel{\text{iid}}{\sim}$ from joint distribution

Y_i are not iid
(b/c of distribution of Y_i depends on X_i)

$$L(\beta | \mathbf{X}, \mathbf{Y}) \propto \prod_{i=1}^n f(Y_i | \beta, X_i)$$

for the logistic regression problem.

$$L(\beta | X, Y) = f(X, Y | \beta) = \hat{\prod}_{i=1} \underbrace{f(X_i, Y_i | \beta)}_{\underbrace{f(X_i | \beta) f(Y_i | X_i, \beta)}} = \left(\hat{\prod}_{i=1} f(X_i) \right) \left(\hat{\prod}_{i=1} f(Y_i | X_i, \beta) \right)$$

$$\Rightarrow L(\beta | X, Y) \propto \hat{\prod}_{i=1} f(Y_i | X_i, \beta) = \hat{\prod}_{i=1} p_i^{Y_i} (1-p_i)^{1-Y_i}$$

$$= \hat{\prod}_{i=1} \left(\frac{e^{\beta^T X_i}}{1 + e^{\beta^T X_i}} \right)^{Y_i} \left(\frac{1}{1 + e^{\beta^T X_i}} \right)^{1-Y_i} \quad p_i = \frac{e^{\beta^T X_i}}{1 + e^{\beta^T X_i}}$$

$$\Rightarrow \ell(\beta | X, Y) = \sum_{i=1}^n \left\{ Y_i \log \left(\frac{e^{\beta^T X_i}}{1 + e^{\beta^T X_i}} \right) + (1 - Y_i) \log \left(\frac{1}{1 + e^{\beta^T X_i}} \right) \right\}$$

$$= \sum_{i=1}^n \left\{ Y_i \beta^T X_i - \log(1 + e^{\beta^T X_i}) \right\}$$

$$L(\beta | X, Y) = \sum_{i=1}^n \{ y_i \beta^T X_i - \log(1 + e^{\beta^T X_i}) \} \quad \beta = \begin{pmatrix} \beta_0 \\ \vdots \\ \beta_n \end{pmatrix}$$

$$\frac{\partial L}{\partial \beta} = \begin{pmatrix} \frac{\partial L}{\partial \beta_0} \\ \frac{\partial L}{\partial \beta_1} \\ \vdots \\ \frac{\partial L}{\partial \beta_n} \end{pmatrix}$$

Rule for matrix derivatives:

$$\frac{\partial}{\partial \beta} \beta^T X_i = X_i$$

$$X_i = \begin{pmatrix} 1 \\ x_{i1} \\ x_{i2} \\ \vdots \\ x_{in} \end{pmatrix}$$

$$\begin{aligned} \frac{\partial L}{\partial \beta} &= \sum_{i=1}^n \left\{ \underbrace{\frac{\partial L}{\partial \beta} y_i \beta^T X_i}_{y_i X_i} - \underbrace{\frac{\partial}{\partial \beta} \log(1 + e^{\beta^T X_i})}_{\frac{1}{1 + e^{\beta^T X_i}} \cdot e^{\beta^T X_i} \cdot X_i} \right\} \\ &= \sum_{i=1}^n \left\{ y_i X_i - \frac{e^{\beta^T X_i}}{1 + e^{\beta^T X_i}} X_i \right\} = \sum_{i=1}^n \left(y_i - \frac{e^{\beta^T X_i}}{1 + e^{\beta^T X_i}} \right) X_i \end{aligned}$$

$$= \sum_i (y_i - p_i) X_i = X^T (Y - P)$$

$$X = \begin{bmatrix} 1 & x_{11} & \dots & x_{1n} \\ 1 & x_{21} & \dots & x_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & \dots & x_{nn} \end{bmatrix} \quad Y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad P = \begin{bmatrix} p_1 \\ \vdots \\ p_n \end{bmatrix}$$

Aside:

$$X^T(Y - \mu) \stackrel{\text{set}}{=} 0$$

(linear regression)

$$\mu = XB$$

$$X^T(Y - XB) = 0$$

$$X^TY - X^T XB = 0$$

$$\Rightarrow (X^T X)^{-1} X^T Y = \hat{\beta}$$

$$\underbrace{u(\beta)}_{\text{score function}} = \frac{\partial \ell}{\partial \beta} = \begin{matrix} X^T(Y - \mu) \\ X^T\left(1 - \frac{e^{XB}}{1 + e^{XB}}\right) \end{matrix} \begin{matrix} \stackrel{\text{set}}{=} 0 \\ \stackrel{\text{set}}{=} 0 \end{matrix}$$

score
function

want to find β^* st $u(\beta^*) = 0$

no closed-form solution for logistic regression model

- Idea:
- 1) Start w/ an initial guess $\beta^{(0)}$
 - 2) update guess to $\beta^{(1)}$, which is (hopefully!) closer to β^*
 - 3) Iterate!

Iterative methods for maximizing likelihood

Newton's method

want β^* st $u(\beta^*) = 0$, given initial guess $\beta^{(0)}$

First-order Taylor expansion around $\beta^{(0)}$

$$u(\beta^*) \approx u(\beta^{(0)}) + \frac{\partial u(\beta^{(0)})}{\partial \beta^{(0)}} (\beta^* - \beta^{(0)})$$

||
0

$$\Rightarrow u(\beta^{(0)}) + \frac{\partial u(\beta^{(0)})}{\partial \beta^{(0)}} (\beta^* - \beta^{(0)}) \approx 0$$

$$\Rightarrow \beta^* \approx \beta^{(0)} - \underbrace{\left(\frac{\partial u(\beta^{(0)})}{\partial \beta^{(0)}} \right)^{-1} u(\beta^{(0)})}_{\text{we can evaluate this!}}$$

$$\Rightarrow \beta^{(1)} = \beta^{(0)} - \left(\frac{\partial u(\beta^{(0)})}{\partial \beta^{(0)}} \right)^{-1} u(\beta^{(0)})$$

$$u(\beta) = \begin{pmatrix} \frac{\partial \ell}{\partial \beta_0} \\ \frac{\partial \ell}{\partial \beta_1} \\ \vdots \end{pmatrix} = \frac{\partial \ell}{\partial \beta}$$

$$\frac{\partial u(\beta)}{\partial \beta} = \frac{\partial^2 \ell}{\partial \beta^2} = \begin{bmatrix} \frac{\partial^2 \ell}{\partial \beta_0^2} & \frac{\partial^2 \ell}{\partial \beta_0 \partial \beta_1} & \dots & \frac{\partial^2 \ell}{\partial \beta_0 \partial \beta_n} \\ \vdots & & & \\ \frac{\partial^2 \ell}{\partial \beta_n \partial \beta_0} & \dots & \dots & \frac{\partial^2 \ell}{\partial \beta_n^2} \end{bmatrix}$$

Hessian of log likelihood $H(\beta)$

$$\beta^{(1)} = \beta^{(0)} - (H(\beta^{(0)}))^{-1} u(\beta^{(0)})$$

Newton's method for logistic regression

Example

Suppose that $\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 X_i$, and we have

$$\beta^{(r)} = \begin{bmatrix} -3.1 \\ 0.9 \end{bmatrix}, \quad U(\beta^{(r)}) = \begin{bmatrix} 9.16 \\ 31.91 \end{bmatrix},$$

$$\mathbf{H}(\beta^{(r)}) = - \begin{bmatrix} 17.834 & 53.218 \\ 53.218 & 180.718 \end{bmatrix}$$

Use Newton's method to calculate $\beta^{(r+1)}$ (you may use R or a calculator, you do not need to do the matrix arithmetic by hand).

