

Lecture 28: Wald vs. LRT

Equivalence of the Wald and LRT statistics

$H_0: \theta = \theta_0$ vs. $H_A: \theta \neq \theta_0$ $\theta \in \mathbb{R}^2$

(X_1, \dots, X_n) from distribution with parameter θ

Suppose we want to look performance for a given value of θ

Fixed alternative: true $\theta = \theta_0 + d$ $d \in \mathbb{R}^2$
↑
difference between null & alternative hypotheses

Local alternative: true $\theta = \theta_0 + \frac{d}{\sqrt{n}}$

Intuition: under a fixed alternative $\theta = \theta_0 + d$,
expect $\text{Power}(\theta) \rightarrow 1$ as $n \rightarrow \infty$

under a local alternative, maybe $\text{Power}(\theta) \rightarrow ? \in (0, 1)$

Key points

- Under H_0 , Wald & LRT are asymptotically equivalent as $n \rightarrow \infty$
- For a fixed alternative $\theta = \theta_0 + d$, Wald and LRT are not equivalent (if H_A is true)
- For a local alternative $\theta = \theta_0 + \frac{d}{\sqrt{n}}$, Wald and LRT are asymptotically equivalent as $n \rightarrow \infty$ (if H_A is true)

For a local alternative $\theta = \theta_0 + \frac{d}{\sqrt{n}}$, or under $H_0: \theta = \theta_0$, Wald & LRT statistics are asymptotically equivalent

why? Consider $\theta \in \mathbb{R}$

From previous class: if $\hat{\theta} \approx \theta_0$ (either H_0 is true, or $\theta = \theta_0 + \frac{d}{\sqrt{n}}$)

then

$$2\ell(\hat{\theta}) - 2\ell(\theta_0) \approx \underbrace{-\frac{1}{2} \ell''(\hat{\theta})}_{\mathcal{I}_1(\theta)} (\sqrt{n}(\hat{\theta} - \theta_0))^2$$

more generally,

$$\begin{aligned} 2\ell(\hat{\theta}) - 2\ell(\theta_0) &\approx \mathcal{I}_1(\theta_0) n(\hat{\theta} - \theta_0)^2 \\ &\approx (\hat{\theta} - \theta_0)^T n \mathcal{I}_1(\theta_0) (\hat{\theta} - \theta_0) \\ &\quad \text{(Wald test statistic)} \\ &= \left(\frac{\sqrt{n}(\hat{\theta} - \theta_0)}{\mathcal{I}_1^{-1}(\theta_0)} \right)^2 \\ &= \text{Wald test statistic!} \end{aligned}$$

Asymptotic normality of MLE:

$$\hat{\theta} \approx N(\theta, \mathcal{I}^{-1}(\theta))$$

$$\mathcal{I}(\theta) = n \mathcal{I}_1(\theta)$$

Test $H_0: \theta = \theta_0$ vs. $H_A: \theta \neq \theta_0$

$$W = (\hat{\theta} - \theta_0)^T n \mathcal{I}_1(\theta_0) (\hat{\theta} - \theta_0)$$

Under H_0 , $W \approx \chi^2_q$

$q = \text{dimension of } \theta$
 $\theta \in \mathbb{R}^q$

what happens if H_A is true?

$$\hat{\theta} \approx N(\theta, \mathcal{I}^{-1}(\theta))$$

$$\Rightarrow \hat{\theta} - \theta_0 \approx N(\theta - \theta_0, \mathcal{I}^{-1}(\theta))$$

Def: If $Z \sim N(\mu, \mathcal{I})$

$$\Rightarrow \mathcal{I}^{\frac{1}{2}}(\theta)(\hat{\theta} - \theta_0) \approx N(\mathcal{I}^{\frac{1}{2}}(\theta)(\theta - \theta_0), \mathcal{I})$$

$$Z^T Z \sim \chi^2_q(\lambda)$$

$$\Rightarrow (\hat{\theta} - \theta_0)^T n \mathcal{I}_1(\hat{\theta}) (\hat{\theta} - \theta_0) \approx$$

$$W \approx \chi^2_q(\lambda)$$

$$\lambda = \mu^T \mu$$

Under $\theta = \theta_0 + d$: $\lambda = n d^T \mathcal{I}_1(\theta) d \rightarrow \infty$ as $n \rightarrow \infty$
 Under $\theta = \theta_0 + \frac{d}{\sqrt{n}}$: $\lambda = d^T \mathcal{I}_1(\theta_0) d \in \mathbb{R}$

$$\mathcal{I}^{\frac{1}{2}}(\theta)(\hat{\theta} - \theta_0) \approx N(\mathcal{I}^{-\frac{1}{2}}(\theta)(\theta - \theta_0), \mathcal{I})$$

$$(\hat{\theta} - \theta_0)^T \sim \mathcal{I}_1(\hat{\theta})(\hat{\theta} - \theta_0)$$

$$\approx \chi^2_2(\lambda)$$

$$\lambda = (\theta - \theta_0)^T \mathcal{I}_1(\theta)(\theta - \theta_0)$$

$$\left. \begin{array}{l} Z \sim N(\mu, \mathcal{I}) \\ \Rightarrow Z^T Z \sim \chi^2_2(\lambda) \end{array} \right\}$$

if $\theta - \theta_0 = d$ (fixed alternative):

$$\lambda = d^T \mathcal{I}_1(\theta) d \rightarrow \infty$$

if $\theta - \theta_0 = \frac{d}{\sqrt{n}}$

$$\lambda = n \left(\frac{d}{\sqrt{n}} \right)^T \underbrace{\mathcal{I}_1\left(\theta_0 + \frac{d}{\sqrt{n}}\right)}_{\approx \mathcal{I}_1(\theta_0)} \left(\frac{d}{\sqrt{n}} \right)$$

$$= d^T \mathcal{I}_1(\theta_0) d \quad \text{as } n \rightarrow \infty$$

$\nrightarrow \infty$

