

# Confidence intervals

## Recap: Pivotal quantities

Let  $X_1, \dots, X_n$  be a sample and  $\theta$  be an unknown parameter. A function  $Q(X_1, \dots, X_n, \theta)$  is called a pivot if the distribution of  $Q(X_1, \dots, X_n, \theta)$  does not depend on  $\theta$ .

To construct a  $1-\alpha$  confidence set:

Find  $a, b$  st  $P_\theta(a \leq Q(X_1, \dots, X_n, \theta) \leq b) = 1-\alpha$

The  $1-\alpha$  confidence set for  $\theta$  is

$$\{ \theta : a \leq Q(X_1, \dots, X_n, \theta) \leq b \}$$

## Example

$$P_{\theta} \left( \underbrace{\alpha^{\frac{1}{n}} \leq \frac{X_{(n)}}{\theta} \leq 1}_{1 \leq \frac{\theta}{X_{(n)}} \leq \frac{1}{\alpha^{\frac{1}{n}}}} \right) = 1 - \alpha \Rightarrow \text{CI} = \left[ X_{(n)}, \frac{X_{(n)}}{\alpha^{\frac{1}{n}}} \right]$$

Suppose  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}[0, \theta]$ . We want to construct a  $1 - \alpha$  confidence set for  $\theta$  using a pivot.

Last time:

$$\underbrace{Q(X_1, \dots, X_n, \theta)}_{\text{pivot!}} = \frac{X_{(n)}}{\theta} \Rightarrow P\left(\frac{X_{(n)}}{\theta} \leq t\right) = \left(\frac{t \cdot \theta}{\theta}\right)^n = t^n \quad \begin{matrix} \nwarrow \\ \in [0, 1] \\ \nearrow \end{matrix}$$

does not depend on  $\theta$ !

next: choose  $a, b$  st  $P_{\theta} \left( a \leq \frac{X_{(n)}}{\theta} \leq b \right) = 1 - \alpha$

e.g.  $b = 1$ ,  $P_{\theta} \left( a \leq \frac{X_{(n)}}{\theta} \leq 1 \right) = P_{\theta} \left( a \leq \frac{X_{(n)}}{\theta} \right)$

$$= 1 - P_{\theta} \left( \frac{X_{(n)}}{\theta} \leq a \right) \stackrel{\text{set}}{=} 1 - \alpha$$

$$\Rightarrow a = \alpha^{\frac{1}{n}}$$

## Example

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exponential}(\theta)$ , with density  $f(x|\theta) = \theta e^{-\theta x}$ .

Find a pivotal quantity  $Q(X_1, \dots, X_n, \theta)$  and construct a  $1 - \alpha$  confidence interval for  $\theta$  using the pivotal quantity.

Hints:

- + Begin with the maximum likelihood estimate of  $\theta$ , which is

$$\hat{\theta} = \frac{n}{\sum_{i=1}^n X_i}$$

- + If  $X \sim \text{Exponential}(\theta)$ , then  $cX \sim \text{Exponential}\left(\frac{\theta}{c}\right)$
- +  $\text{Exponential}\left(\frac{1}{2}\right) = \chi^2_2$

$$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exp}(\theta)$$

$$Q(X_1, \dots, X_n, \theta) = 2\theta \sum_i X_i$$

$$= \sum_i 2\theta X_i$$

$$2\theta X_i \sim \chi^2_2$$

$$\sim \chi^2_{2n}$$

want  $a, b$  st

$$P_\theta \left( a \leq 2\theta \sum_i X_i \leq b \right) = 1 - \alpha$$

or

$$a = \chi^2_{2n, 1-\frac{\alpha}{2}}$$

$$b = \chi^2_{2n, \frac{\alpha}{2}}$$

$$a=0$$

$$b = \chi^2_{2n, \alpha}$$

etc.

$$\Rightarrow \left[ \frac{a}{2 \sum_i X_i}, \frac{b}{2 \sum_i X_i} \right]$$

$$\mathcal{I}(\theta) \cdot \mathcal{I}^{-1}(\theta) = 1$$

## Wald CI

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Exponential}(\theta)$ , with density  $f(x|\theta) = \theta e^{-\theta x}$ .

$$\text{MLE: } \hat{\theta} = \frac{1}{\bar{x}} = \frac{n}{\sum_i X_i}$$

$$\hat{\theta} \approx N(\theta, \mathcal{I}^{-1}(\theta))$$

$$\approx N(\theta, \frac{\theta^2}{n})$$

$$\mathcal{I}(\theta) = \frac{n}{\theta^2}$$

$$\rightarrow \mathcal{I}^{-1}(\theta) = \frac{\theta^2}{n}$$

Wald CI for  $\theta$ :

$$\hat{\theta} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\theta^2}{n}} \leftarrow \text{don't know } \theta$$

variance depends on  $\theta$

$$\hat{\theta} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{\theta}^2}{n}}$$

Can we find a transformation  $g(\hat{\theta})$  such that the variance  $\text{Var}(g(\hat{\theta}))$  does not depend on  $\theta$ ?

Exponential:

$$\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{d} N(0, \theta^2)$$

## Delta method

$$\sqrt{n}(g(\hat{\theta}) - g(\theta)) \xrightarrow{d} N(0, \theta^2 \cdot [g'(\theta)]^2)$$

Suppose  $\hat{\theta}$  is an estimate of  $\theta \in \mathbb{R}$ , such that

$$\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{d} N(0, \sigma^2)$$

for some  $\sigma^2$ , and  $g$  is a continuously differentiable function with  $g'(\theta) \neq 0$ .  
Then

$$\sqrt{n}(g(\hat{\theta}) - g(\theta)) \xrightarrow{d} N(0, \sigma^2 [g'(\theta)]^2)$$

**Proof sketch:**

- + First-order Taylor expansion of  $g(\hat{\theta})$  around  $\theta$
- + Slutsky's theorem

$$\begin{aligned} \text{want } g'(\theta) &= \frac{1}{\theta} & g(\theta) &= \log(\theta) \\ \sqrt{n}(\log(\hat{\theta}) - \log(\theta)) &\xrightarrow{d} N(0, 1) \end{aligned}$$

# Variance stabilizing transformations