

# Lecture 12: Convergence in distribution

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# Logistics

- ▶ Reminder: Exam 1 released February 21 (covers HW 1–4)
- ▶ Early-semester feedback form sent out

## Recap: Convergence in probability

**Definition:** A sequence of random variables  $X_1, X_2, \dots$  *converges in probability* to a random variable  $X$  if, for every  $\varepsilon > 0$ ,

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \varepsilon) = 0$$

We write  $X_n \xrightarrow{p} X$ .

## Example

Suppose that  $X_1, X_2, \dots \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ , and let  $X_{(n)} = \max\{X_1, \dots, X_n\}$ . Then  $X_{(n)} \xrightarrow{P} 1$ .

pf: WTS  $P(|X_{(n)} - 1| > \varepsilon) \rightarrow 0$  as  $n \rightarrow \infty \forall \varepsilon > 0$

Let  $\varepsilon > 0$

$$\begin{aligned} P(|X_{(n)} - 1| > \varepsilon) &= 1 - P(|X_{(n)} - 1| \leq \varepsilon) \\ &= 1 - P(1 - \varepsilon \leq X_{(n)} \leq 1 + \varepsilon) \end{aligned}$$

we know  $X_{(n)} \leq 1 \Rightarrow X_{(n)} \leq 1 + \varepsilon$

$$\Rightarrow P(1 - \varepsilon \leq X_{(n)} \leq 1 + \varepsilon) = P(1 - \varepsilon \leq X_{(n)})$$

$$P(1 - \varepsilon \leq X_{(n)}) = 1 - P(X_{(n)} \leq 1 - \varepsilon)$$

$$\Rightarrow P(|X_{(n)} - 1| > \varepsilon) = 1 - (1 - P(X_{(n)} \leq 1 - \varepsilon)) = P(X_{(n)} \leq 1 - \varepsilon)$$

$$= (P(X_i \leq 1 - \varepsilon))^n = (1 - \varepsilon)^n \rightarrow 0 \text{ as } n \rightarrow \infty$$

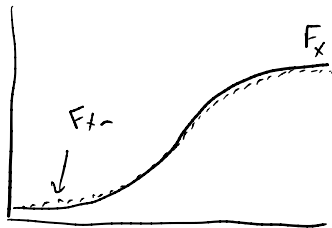
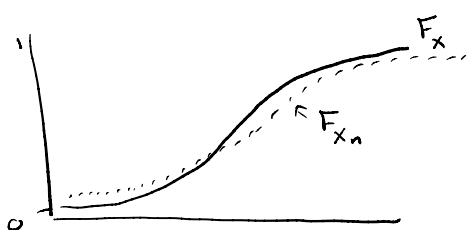
$$\Rightarrow P(|X_{(n)} - 1| > \varepsilon) \rightarrow 0 \text{ as } n \rightarrow \infty //$$

# Convergence in distribution

**Definition:** A sequence of random variables  $X_1, X_2, \dots$  converges in distribution to a random variable  $X$  if

$$\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$$

at all points where  $F_X(x)$  is continuous. We write  $X_n \xrightarrow{d} X$ .



as  $n$  increases,  
 $F_{X_n}$  gets closer to  $F_X$

$\hookrightarrow n$  increases

## Example

Exponential ( $\theta$ )      cdf =  $1 - e^{-\theta t}$

Suppose that  $X_1, X_2, \dots \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ . Let

$X_{(n)} = \max\{X_1, \dots, X_n\}$ . Then  $n(1 - X_{(n)}) \xrightarrow{d} Y$ , where  
 $Y \sim \text{Exp}(1)$ .

WTS :  $F_{n(1-X_{(n)})}(t) \rightarrow F_Y(t) \quad \forall t \text{ where } F_Y \text{ is continuous}$

$$F_Y(t) = 1 - e^{-t}$$

Pf :  $F_{n(1-X_{(n)})}(t) = P(n(1-X_{(n)}) \leq t)$

$$= P(1-X_{(n)} \leq \frac{t}{n}) = P(X_{(n)} \geq 1 - \frac{t}{n})$$
$$= 1 - P(X_{(n)} \leq 1 - \frac{t}{n})$$
$$= 1 - \left(1 - \frac{t}{n}\right)^n$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{t}{n}\right)^n = e^{-t}$$

$$\Rightarrow 1 - \left(1 - \frac{t}{n}\right)^n \rightarrow 1 - e^{-t}$$

in general:

$$\lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n = e^x$$

as  $n \rightarrow \infty$

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# Convergence in distribution: Central Limit Theorem

Let  $X_1, X_2, \dots$  be iid random variables, whose mgf exists in a neighborhood of 0. Let  $\mu = \mathbb{E}[X_i]$  and  $\sigma^2 = \text{Var}(X_i) < \infty$ . Then

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} Z$$

where  $Z \sim N(0, 1)$ .

intuition:  $\bar{X}_n - \mu \xrightarrow{P} 0$

need to multiply by something increasing  
to "balance out" convergence to 0

why  $\sqrt{n}$ ?  $\text{SD}(\bar{X}_n) = \frac{\sigma}{\sqrt{n}}$

$\text{SD}(\sqrt{n} \bar{X}_n) = \sigma$  (not increasing  
or decreasing  
as  $n \rightarrow \infty$ )

# Activity

Simulations to explore convergence in distribution:

[https://sta711-s25.github.io/class\\_activities/ca\\_lecture\\_12.html](https://sta711-s25.github.io/class_activities/ca_lecture_12.html)