

# Lecture 16: Wald tests

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## Wald test for one parameter

Let  $\theta \in \mathbb{R}$  be a parameter of interest, and let  $\hat{\theta}_n$  be an estimator s.t.  $\frac{\hat{\theta}_n - \theta}{S_n} \xrightarrow{d} N(0,1)$

for some sequence  $S_n$  ( $S_n^2 \approx \text{var}(\hat{\theta}_n)$ )

To test  $H_0: \theta = \theta_0$  vs.  $H_A: \theta \neq \theta_0$

• let  $Z_n = \frac{\hat{\theta}_n - \theta_0}{S_n}$

• reject when  $|Z_n| > z_{\frac{\alpha}{2}}$   $\leftarrow$  upper  $\frac{\alpha}{2}$  quantile of  $N(0,1)$

• p-value:  $2\Phi(-|Z_n|)$   
 $\uparrow$   
 $N(0,1)$  cdf

## Wald test for one parameter: examples

- Population mean:  $H_0 : \mu = \mu_0$

$$\hat{\Theta}_n = \bar{X}_n, \quad \Theta_0 = \mu_0, \quad S_n = \frac{\sqrt{\sigma}}{\sqrt{n}} \quad \text{or} \quad \frac{s}{\sqrt{n}}$$
$$s = \sqrt{\frac{1}{n-1} \sum_i (x_i - \bar{x}_n)^2}$$

- Population proportion:  $H_0 : p = p_0$

$$\hat{\Theta}_n = \hat{P}_n, \quad \Theta_0 = p_0, \quad S_n = \sqrt{\frac{p_0(1-p_0)}{n}} \quad \text{or} \quad \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

- Regression coefficient:  $H_0 : \beta_j = 0$

$$\hat{\Theta}_n = \hat{\beta}_j, \quad \Theta_0 = 0, \quad S_n = \sqrt{\underbrace{[\hat{I}^{-1}(\beta)]_{jj}}_{j^{\text{th}} \text{ diagonal entry}}}$$

## Testing multiple parameters

Logistic regression model for the dengue data:

$$Y_i \sim \text{Bernoulli}(p_i)$$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 WBC_i + \beta_2 PLT_i$$

Researchers want to know if there is any relationship between white blood cell count or platelet count, and the probability a patient has dengue.

**Question:** What hypotheses should they test?

$$H_0: \beta_1 = \beta_2 = 0$$

$$H_A: \text{at least one of } \beta_1, \beta_2 \neq 0$$

## Testing multiple parameters

##	Estimate	Std. Error	z value	Pr(> z )
## (Intercept)	2.642	0.121	21.772	0
## WBC	-0.289	0.013	-21.533	0
## PLT	-0.007	0.001	-11.061	0

$$H_0 : \beta_1 = \beta_2 = 0$$

Can the researchers test their hypotheses using this output?

No! each p-value only tests  $H_0: \beta_j = 0$

## Wald tests for multiple parameters

$$\mathcal{I}(\beta) = X^T W X$$

(Fisher info for the full data)

For the dengue example:

$$\hat{\beta} = \begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} \approx N(\beta, \mathcal{I}^{-1}(\beta))$$

```
m1 <- glm(Dengue ~ WBC + PLT, data = dengue, family = binom)
coef(m1)  $\hat{\beta}$ 
```

```
## (Intercept)          WBC          PLT
## 2.641506279 -0.289290446 -0.006561464
```

```
vcov(m1)  $\hat{\mathcal{I}}^{-1}(\beta)$   $X^T \hat{W} X$ 
```

```
## (Intercept)          WBC          PLT
## (Intercept) 1.471934e-02 -4.937020e-04 -5.125888e-05
## WBC -4.937020e-04 1.804972e-04 -3.221337e-06
## PLT -5.125888e-05 -3.221337e-06 3.518938e-07
```

# Multivariate normal distribution

**Definition:** Let  $X = (X_1, \dots, X_k)^T$ . We say that  $X \sim N(\mu, \Sigma)$  if for any  $\mathbf{a} \in \mathbb{R}^k$ ,  $\mathbf{a}^T X$  follows a (univariate) normal distribution.

$$\blacktriangleright \mu = \mathbb{E}[X] = \begin{bmatrix} \mathbb{E}[X_1] \\ \mathbb{E}[X_2] \\ \vdots \\ \mathbb{E}[X_n] \end{bmatrix} \in \mathbb{R}^n$$

$$\blacktriangleright \Sigma = \text{Var}(X)$$

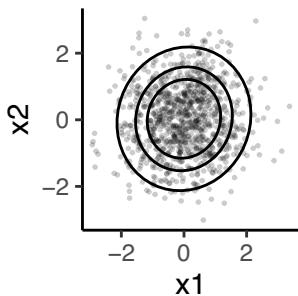
(variance-covariance matrix)

$$= \begin{bmatrix} \text{Var}(X_1) & \text{Cov}(X_1, X_2) & \dots & \text{Cov}(X_1, X_n) \\ \text{Cov}(X_2, X_1) & \text{Var}(X_2) & & \\ \vdots & & \ddots & \\ \text{Cov}(X_n, X_1) & \dots & & \text{Var}(X_n) \end{bmatrix}$$
$$\in \mathbb{R}^{n \times n}$$

# Multivariate normal distribution

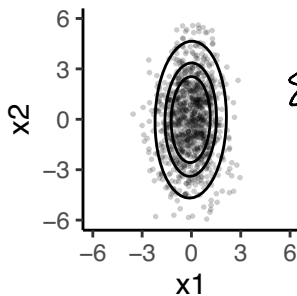
$$\mathcal{N}(\mathbf{0}, \mathbf{I})$$

$$\mathbf{I} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



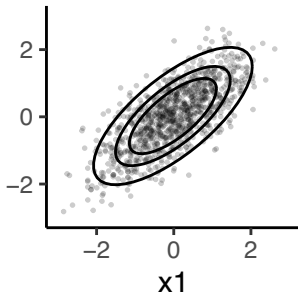
$$\mathcal{N}(\mathbf{0}, \Sigma)$$

$$\Sigma = \begin{pmatrix} 1 & 0 \\ 0 & 5 \end{pmatrix}$$



$$\mathcal{N}(\mathbf{0}, \Sigma)$$

$$\Sigma = \begin{pmatrix} 1 & 0.75 \\ 0.75 & 1 \end{pmatrix}$$





## Wald tests for multiple parameters

For the dengue example:  $\hat{\beta} = \begin{pmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} \approx \text{Normal}$

We want to test:  $\begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad H_0: C\beta = 0$

rewrite:

$$\begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix} = \begin{pmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}$$
$$= C\beta$$

$$\begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix} = C\hat{\beta}$$

In general: hypotheses of form  $H_0: C\beta = \gamma_0$

vector  
under  
 $H_0$

# Wald tests for multiple parameters

want:  
 $\text{var}(C\hat{\beta})$

$$H_0 : C\beta = \gamma_0$$

For the dengue example:

```
C <- matrix(c(0, 1, 0,  
              0, 0, 1), nrow=2, byrow=T)
```

```
C
```

```
##      [,1] [,2] [,3]  
## [1,]    0    1    0  
## [2,]    0    0    1
```

```
C %*% coef(m1)
```

```
##      [,1]  
## [1,] -0.289290446  $\hat{\beta}_1$   
## [2,] -0.006561464  $\hat{\beta}_2$ 
```

## Wald tests for multiple parameters

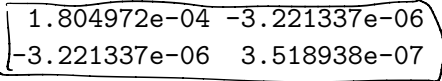
For the dengue example:

```
C
```

```
##           [,1] [,2] [,3]
## [1,]         0    1    0
## [2,]         0    0    1
```

```
vcov(m1)
```

```
##              (Intercept)              WBC              PLT
## (Intercept)  1.471934e-02 -4.937020e-04 -5.125888e-05
## WBC          -4.937020e-04  1.804972e-04 -3.221337e-06
## PLT          -5.125888e-05 -3.221337e-06  3.518938e-07
```



```
C %>% vcov(m1) %>% t(C)
```

```
##           [,1] [,2]
## [1,] 1.804972e-04 -3.221337e-06
## [2,] -3.221337e-06  3.518938e-07
```

## Wald tests for multiple parameters

univariate intuition:

$$X \sim N(\mu, \sigma^2)$$

$$aX \sim N(a\mu, a^2\sigma^2)$$

►  $H_0 : \mathbf{C}\beta = \gamma_0$

► Look at  $\mathbf{C}\hat{\beta}$

**Fact:** Suppose  $X \sim N(\mu, \Sigma)$  (multivariate normal), and  $\mathbf{A}$  is a matrix (not random). Then:

$$\mathbf{A}X \sim N(\mathbf{A}\mu, \mathbf{A}\Sigma\mathbf{A}^T) \quad (\text{pf: Hw?})$$

$$\hat{\beta} \approx N(\beta, \mathcal{I}^{-1}(\beta))$$

$$\mathbf{C}\hat{\beta} \approx N(\mathbf{C}\beta, \mathbf{C}\mathcal{I}^{-1}(\beta)\mathbf{C}^T)$$

under  $H_0$ :  $\mathbf{C}\hat{\beta} \approx N(\gamma_0, \mathbf{C}\mathcal{I}^{-1}(\beta)\mathbf{C}^T)$

## Test statistic and p-value

- ▶  $H_0 : \mathbf{C}\beta = \gamma_0$
- ▶  $\mathbf{C}\hat{\beta} \approx N(\mathbf{C}\beta, \mathbf{C}\mathbf{I}^{-1}(\beta)\mathbf{C}^T)$
- ▶ Want to turn  $\mathbf{C}\hat{\beta}$  into a scalar test statistic

Facts:

- If  $Z \sim N(0, 1)$  (scalar), then  $Z^2 \sim \chi^2_1$
- If  $Z_1, \dots, Z_H \stackrel{\text{iid}}{\sim} N(0, 1)$ , then  $\sum_{i=1}^H Z_i^2 \sim \chi^2_H$

$$\Rightarrow \sum_i Z_i^2 = [Z_1 \dots Z_H] \begin{bmatrix} Z_1 \\ \vdots \\ Z_H \end{bmatrix}$$

$\Rightarrow$  if  $Z \sim N(0, \mathbf{I})$  (multivariate normal),  
then  $Z^T Z \sim \chi^2_H$

## Test statistic and p-value

$$Z \sim N(0, I) \\ \Rightarrow Z^T Z \sim \chi^2_n$$

►  $H_0: C\beta = \gamma_0$

►  $C\hat{\beta} \approx N(C\beta, CI^{-1}(\beta)C^T)$

► Want to turn  $C\hat{\beta}$  into a scalar test statistic

under  $H_0$ :  $C\hat{\beta} \approx N(\gamma_0, CI^{-1}(\beta)C^T)$

$$C\hat{\beta} - \gamma_0 \approx N(0, CI^{-1}(\beta)C^T)$$

$$(CI^{-1}(\beta)C^T)^{-\frac{1}{2}} (C\hat{\beta} - \gamma_0) \approx N(0, I)$$

$$\Rightarrow (C\hat{\beta} - \gamma_0)^T (CI^{-1}(\beta)C^T)^{-\frac{1}{2}} (CI^{-1}(\beta)C^T)^{-\frac{1}{2}} (C\hat{\beta} - \gamma_0) \approx \chi^2_z$$

$$\Rightarrow (C\hat{\beta} - \gamma_0)^T (CI^{-1}(\beta)C^T)^{-1} (C\hat{\beta} - \gamma_0) \approx \chi^2_z$$

$z = \dim(C\beta)$   
= # parameters tested

## Example

$$H_0 : \mathbf{C}\beta = \gamma_0$$

Dengue example:

$$H_0 : \mathbf{C}\beta = \mathbf{0}$$

$$(\mathbf{C}\hat{\beta} - \gamma_0)^T (\mathbf{C}\mathcal{I}^{-1}(\beta)\mathbf{C}^T)^{-1} (\mathbf{C}\hat{\beta} - \gamma_0) \approx \chi_q^2 \quad \text{under } H_0$$

```
C %*% coef(m1)
                $\hat{\beta}_1$        $\hat{\beta}_2$ 
## [1] -0.289290446 -0.006561464

test_stat <- t(C %*% coef(m1)) %*% solve(C %*% vcov(m1) %*%
(C %*% coef(m1)))
test_stat

##           [,1]
## [1,] 930.5777

pchisq(test_stat, 2, lower.tail=F)

##           [,1]
## [1,] 8.464858e-203
```