

Asymptotic properties of maximum likelihood estimators

Ciaran Evans

Key results for the MLE

Let Y_1, Y_2, \dots be iid from a distribution with probability function $f(y|\boldsymbol{\theta})$, where $\boldsymbol{\theta} \in \mathbb{R}^d$ is the parameter(s) we are trying to estimate.
Let

$$\ell_n(\boldsymbol{\theta}) = \sum_{i=1}^n \log f(Y_i|\boldsymbol{\theta})$$

$$\hat{\boldsymbol{\theta}}_n = \operatorname{argmax}_{\boldsymbol{\theta}} \ell_n(\boldsymbol{\theta})$$

Theorem: Under certain regularity conditions,

(a) $\hat{\boldsymbol{\theta}}_n \xrightarrow{P} \boldsymbol{\theta}$

(b) $\sqrt{n}(\hat{\boldsymbol{\theta}}_n - \boldsymbol{\theta}) \xrightarrow{d} N(\mathbf{0}, \mathcal{I}_1^{-1}(\boldsymbol{\theta}))$

Application to regression models

Suppose that $(\mathbf{x}_1, Y_1), \dots, (\mathbf{x}_n, Y_n)$ are iid from the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mu_i, \sigma^2)$$

$$\mu_i = \mathbf{x}_i^T \boldsymbol{\beta}$$

The MLE is $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$. Asymptotic normality of the MLE means that

$$\sqrt{n}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}) \xrightarrow{d} N(\mathbf{0}, \mathcal{I}_1^{-1}(\boldsymbol{\beta}))$$

Fisher information for the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mathbf{x}_i^T \beta, \sigma^2)$$

$$\text{Score: } U(\beta) = \frac{1}{\sigma^2} \mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta) = \frac{1}{\sigma^2} \sum_{i=1}^n (y_i - \mathbf{x}_i^T \beta) \mathbf{x}_i$$

$$\text{let } u_i = (y_i - \mathbf{x}_i^T \beta)$$

$$\mathcal{L}_i(\beta) = \text{Var}\left(\frac{1}{\sigma^2} u_i\right) = \frac{1}{\sigma^4} \text{Var}(u_i)$$

$$\text{Var}(u_i) = \text{Var}(\mathbb{E}[u_i | \mathbf{x}_i]) + \mathbb{E}[\text{Var}(u_i | \mathbf{x}_i)]$$

$$\mathbb{E}[u_i | \mathbf{x}_i] = 0$$

$$\text{Var}(u_i | \mathbf{x}_i) = \mathbb{E}[(u_i - \mathbb{E}[u_i | \mathbf{x}_i])(u_i - \mathbb{E}[u_i | \mathbf{x}_i])^T | \mathbf{x}_i]$$

$$= \mathbb{E}[u_i u_i^T | \mathbf{x}_i]$$

$$= \mathbb{E}[(y_i - \mathbf{x}_i^T \beta)^2 \mathbf{x}_i \mathbf{x}_i^T | \mathbf{x}_i]$$

$$= \mathbb{E}[(y_i - \mathbf{x}_i^T \beta)^2 | \mathbf{x}_i] \mathbf{x}_i \mathbf{x}_i^T$$

$$= \text{Var}(y_i | \mathbf{x}_i) \mathbf{x}_i \mathbf{x}_i^T$$

$$= \sigma^2 \mathbf{x}_i \mathbf{x}_i^T$$

Fisher information for the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mathbf{x}_i^T \beta, \sigma^2)$$

$$\text{Score: } U(\beta) = \frac{1}{\sigma^2} \mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta)$$

$$\begin{aligned} \text{let } u_i &= (y_i - \mathbf{x}_i^T \beta) \mathbf{x}_i \\ \mathcal{I}_1(\beta) &= \text{Var}\left(\frac{1}{\sigma^2} u_i\right) = \frac{1}{\sigma^4} \text{Var}(u_i) \end{aligned}$$

$$\text{Var}(u_i) = \text{Var}(\mathbb{E}[u_i | \mathbf{x}_i]) + \mathbb{E}[\text{Var}(u_i | \mathbf{x}_i)]$$

$$\begin{aligned} \text{Var}(u_i) &= \text{Var}(0) + \mathbb{E}[\sigma^2 \mathbf{x}_i \mathbf{x}_i^T] \\ &= \sigma^2 \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T] \end{aligned}$$

$$\Rightarrow \mathcal{I}_1(\beta) = \frac{1}{\sigma^2} \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$$

Fisher information for the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mathbf{x}_i^T \beta, \sigma^2)$$

Fisher information: $\mathcal{I}_1(\beta) = \frac{1}{\sigma^2} \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$

$$\sqrt{n}(\hat{\beta} - \beta) \xrightarrow{d} N(0, \sigma^2 (\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T])^{-1})$$

$$\hat{\beta} \approx N(\beta, \frac{\sigma^2}{n} (\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T])^{-1})$$

need to estimate $\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$
Estimate $\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$ by $\frac{1}{n} \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T = \frac{1}{n} X^T X$
(X = design matrix)

$$\Rightarrow \hat{\beta} \approx N(\beta, \hat{\sigma}^2 (X^T X)^{-1})$$