

Comparing estimators

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Course so far

- ▶ Maximum likelihood estimation
- ▶ Tools for asymptotic results
- ▶ Asymptotic properties of MLEs

Questions:

- ▶ How else could we estimate parameters?
- ▶ What makes a good estimator? How do we compare estimators?
- ▶ Why is maximum likelihood estimation so widespread?
- ▶ What happens to MLEs if our model is *wrong*?

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}(0, \theta)$. How could I estimate θ ?

$$\textcircled{1} \quad \text{MLE} : \quad \hat{\theta} = X_{(n)}$$

$$\textcircled{2} \quad \mathbb{E}[X] = \frac{\theta}{2} \quad \bar{x} \approx \mathbb{E}[X] \\ \Rightarrow \hat{\theta} = 2\bar{x}$$

$$\textcircled{3} \quad \mathbb{E}[X^2] = \frac{\theta^2}{3} \quad \Rightarrow \quad \hat{\theta} = \sqrt{\frac{3}{n} \sum_{i=1}^n x_i^2}$$

$$\textcircled{4} \quad \hat{\theta} = S \quad (\text{probably a terrible estimate})$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}(a, b)$. How could I estimate a and b ?

① MLE: $\hat{a} = X_{(1)}, \quad \hat{b} = X_{(n)}$

② $E[X] = \frac{a+b}{2} = \mu_1 \quad \hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n X_i = \bar{X}$

$E[X^2] = \frac{1}{3} (a^2 + ab + b^2) = \mu_2 \quad \hat{\mu}_2 = \frac{1}{n} \sum_{i=1}^n X_i^2$

$$b = 2\mu_1 - a$$

$$\begin{aligned} \Rightarrow \mu_2 &= \frac{1}{3} (a^2 + a(2\mu_1 - a) + (2\mu_1 - a)^2) \\ &= \frac{1}{3} (a^2 - 2a\mu_1 + 4\mu_1^2) \end{aligned}$$

∴ (algebra)

$$a = \mu_1 - \sqrt{3(\mu_2 - \mu_1^2)}$$

$$\hat{a} = \hat{\mu}_1 - \sqrt{3(\hat{\mu}_2 - \hat{\mu}_1^2)}$$

$$b = \mu_1 + \sqrt{3(\mu_2 - \mu_1^2)}$$

$$\hat{b} = \hat{\mu}_1 + \sqrt{3(\hat{\mu}_2 - \hat{\mu}_1^2)}$$

Method of moments

Let X_1, \dots, X_n be a sample from a distribution with probability function $f(x|\theta_1, \dots, \theta_k)$, with k parameters $\theta_1, \dots, \theta_k$.

$$\begin{array}{llll} \text{Let} & \mu_1 = E[X] & = g_1(\theta_1, \dots, \theta_k) & \hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n x_i \\ & \mu_2 = E[X^2] & = g_2(\theta_1, \dots, \theta_k) & \hat{\mu}_2 = \frac{1}{n} \sum_{i=1}^n x_i^2 \\ & \vdots & \vdots & \vdots \\ & \mu_k = E[X^k] & = g_k(\theta_1, \dots, \theta_k) & \hat{\mu}_k = \frac{1}{n} \sum_{i=1}^n x_i^k \end{array}$$

WLLN: $\hat{\mu}_i$ close to μ_i for large n

The method of moments (MOM) approach estimates $\theta_1, \dots, \theta_k$ by solutions to

$$\hat{\mu}_1 = g_1(\hat{\theta}_1, \dots, \hat{\theta}_k)$$

$$\vdots$$
$$\hat{\mu}_k = g_k(\hat{\theta}_1, \dots, \hat{\theta}_k)$$

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$.

$$\mu_1 = \mathbb{E}[X] = \mu$$

$$\mu_2 = \mathbb{E}[X^2] = \sigma^2 + \mu^2$$

want: $\hat{\mu}$ and $\hat{\sigma}^2$

$$\hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\hat{\mu}_2 = \frac{1}{n} \sum_{i=1}^n x_i^2$$

$$\hat{\mu} = \bar{X} \quad \checkmark$$

$$\sigma^2 = \mu_2 - \mu^2 \quad \Rightarrow \quad \hat{\sigma}^2 = \hat{\mu}_2 - \hat{\mu}_1^2$$

$$= \frac{1}{n} \sum_{i=1}^n x_i^2 - (\bar{X})^2$$

$$= \frac{1}{n} \sum_{i=1}^n (x_i - \bar{X})^2 \quad \checkmark$$

In this case, $\hat{\mu}_{\text{mom}} = \hat{\mu}_{\text{MLE}} = \bar{X}$

$$\hat{\sigma}_{\text{mom}}^2 = \hat{\sigma}_{\text{MLE}}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{X})^2$$

What makes a good estimator?

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}(0, \theta)$. Two possible estimates:

$$\text{MLE: } \hat{\theta} = X_{(n)} \quad \text{MoM: } \hat{\theta} = 2\bar{X}$$

Question: How would I choose between these estimators?

nice properties:

consistency: $\hat{\theta} \xrightarrow{P} \theta$ as $n \rightarrow \infty$
(true here for both $\hat{\theta}_{MLE} = X_{(n)}$ and $\hat{\theta}_{MoM} = 2\bar{X}$)

Low variance: $\text{var}(\hat{\theta})$

unbiased: $E[\hat{\theta}] = \theta$

Bias, variance, and MSE (mean squared error)

MSE: Let $\hat{\theta}$ be an estimator of θ

The MSE of $\hat{\theta}$ is $E_{\theta}[(\hat{\theta} - \theta)^2]$

θ is the value of the parameter

$$E_{\theta}[(\hat{\theta} - \theta)^2] = \text{Var}(\hat{\theta}) + \underbrace{(E_{\theta}[\hat{\theta}] - \theta)^2}_{\text{Bias}^2}$$

$$= \text{Var}(\hat{\theta}) + \text{Bias}^2(\hat{\theta})$$

From HW: if $\text{MSE}(\hat{\theta}) \rightarrow 0$, then $\hat{\theta} \xrightarrow{q.m.} \theta$ and
so $\hat{\theta} \xrightarrow{P} \theta$

One approach to choosing $\hat{\theta}$: minimize MSE

Another approach: restrict to unbiased estimators,
minimize variance

Example

Suppose $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Uniform}(0, \theta)$.