

# Convergence of random variables

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## Last time: Class activity

Suppose that  $X_1, X_2, \dots \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ , and let  $X_{(n)} = \max\{X_1, \dots, X_n\}$ . Then  $X_{(n)} \xrightarrow{P} 1$ .

PF : wTS  $\forall \varepsilon > 0$ ,  $P(|X_{(n)} - 1| \geq \varepsilon) \rightarrow 0$  as  $n \rightarrow \infty$

Let  $0 < \varepsilon < 1$

$$\begin{aligned} P(|X_{(n)} - 1| \geq \varepsilon) &= P(X_{(n)} \leq 1 - \varepsilon) + \underbrace{P(X_{(n)} \geq 1 + \varepsilon)}_{(X_{(n)} = 1)} \\ &= P(X_{(n)} \leq 1 - \varepsilon) \\ &= P(X_1 \leq 1 - \varepsilon, X_2 \leq 1 - \varepsilon, \dots, X_n \leq 1 - \varepsilon) \\ &= P(X_1 \leq 1 - \varepsilon) P(X_2 \leq 1 - \varepsilon) \cdots P(X_n \leq 1 - \varepsilon) \\ &= (1 - \varepsilon)^n \\ &\rightarrow 0 \quad \text{as } n \rightarrow \infty \end{aligned}$$

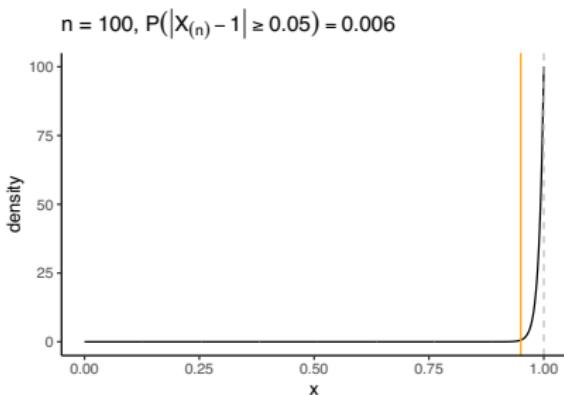
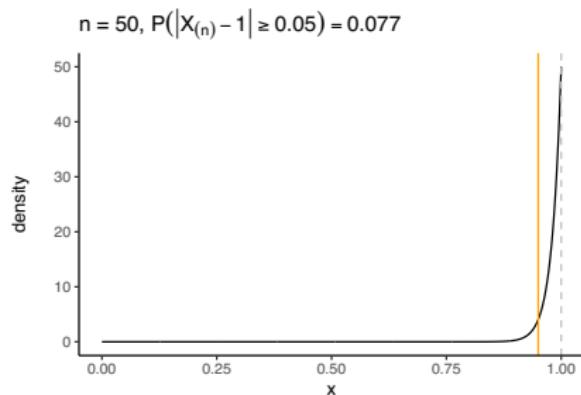
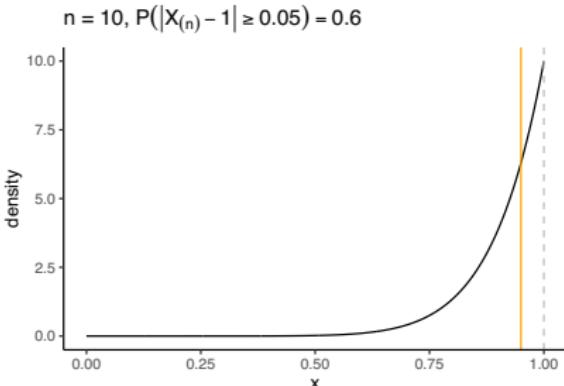
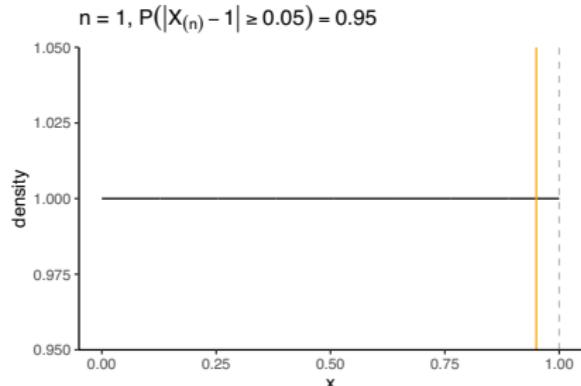
If  $\varepsilon \geq 1$   
 $P(|X_{(n)} - 1| \geq \varepsilon) \leq P(|X_{(n)} - 1| \geq 1) \rightarrow 0$

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## Last time: Class activity

pdf:  $\propto x^{n-1}$

Suppose that  $X_1, X_2, \dots \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ . Then  $X_{(n)} \sim \text{Beta}(n, 1)$



## Warmup

Work on the warmup activity (handout), then we will discuss as a class.

## Warmup

Suppose that  $X_1, X_2, \dots \stackrel{iid}{\sim} \text{Uniform}(0, 1)$ ,  $X_{(n)} = \max\{X_1, \dots, X_n\}$ , and let  $Y_n = n(1 - X_{(n)})$ .

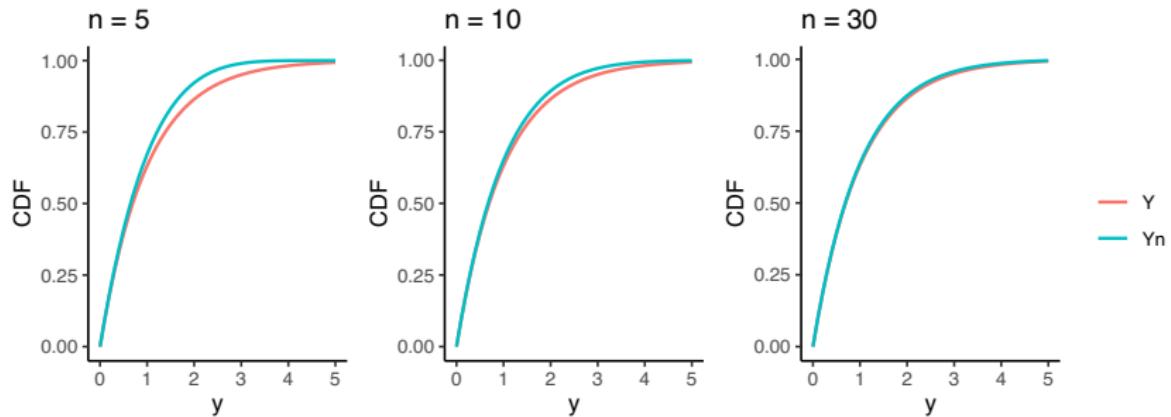
$$\begin{aligned} F_{Y_n}(t) &= P(Y_n \leq t) = P(n(1 - X_{(n)}) \leq t) \\ &= P(1 - X_{(n)} \leq \frac{t}{n}) \\ &= P(X_{(n)} \geq 1 - \frac{t}{n}) \\ &= 1 - P(X_{(n)} \leq 1 - \frac{t}{n}) \\ &= 1 - \left(P(X_i \leq 1 - \frac{t}{n})\right)^n \\ &= 1 - \left(1 - \frac{t}{n}\right)^n \\ \lim_{n \rightarrow \infty} F_{Y_n}(t) &= \lim_{n \rightarrow \infty} 1 - \left(1 - \frac{t}{n}\right)^n \\ &= 1 - e^{-t} \end{aligned}$$

$$\left(1 + \frac{x}{n}\right)^n \rightarrow e^x$$

Weibull distribution

# Warmup

$$F_{Y_n}(t) = 1 - \left(1 - \frac{t}{n}\right)^n \rightarrow F_Y(t) = 1 - e^{-t}$$



## Convergence in distribution

**Definition:** A sequence of random variables  $X_1, X_2, \dots$  converges in distribution to a random variable  $X$  if

$$\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$$

at all points where  $F_X(x)$  is continuous. We write  $X_n \xrightarrow{d} X$ .

## Class activity

Work on the class activity (handout), then we will discuss as a class.

## Class activity

Suppose that  $X_1, X_2, \dots$  are iid random variables with cdf

$$F(x) = \begin{cases} 1 - \left(\frac{1}{x}\right)^\alpha & x \geq 1 \\ 0 & x < 1 \end{cases}$$

Let  $X_{(n)} = \max\{X_1, \dots, X_n\}$ , and  $Y_n = n^{-1/\alpha} X_{(n)}$ .

$$\begin{aligned} F_{Y_n}(t) &= P(n^{-1/\alpha} X_{(n)} \leq t) = P(X_{(n)} \leq n^{\frac{1}{\alpha}} t) \\ &= \left(1 - \left(n^{\frac{1}{\alpha}} t\right)^\alpha\right)^n \\ &= \left(1 - \frac{1}{t^{\alpha/n}}\right)^n \\ &= \left(1 - \frac{t^{-\alpha}}{n}\right)^n \\ &\rightarrow e^{-t^{-\alpha}} \end{aligned}$$

(Frechet distribution)

## Convergence in distribution: Central Limit Theorem

Let  $X_1, X_2, \dots$  be iid random variables, with  $\mu = \mathbb{E}[X_i]$  and  $0 < \sigma^2 = \text{Var}(X_i) < \infty$ . Then

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} Z$$

where  $Z \sim N(0, 1)$ .

what if we don't know  $\sigma^2$ .

$$\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$

Does  $\frac{\sqrt{n}(\bar{X}_n - \mu)}{\hat{\sigma}} \xrightarrow{d} Z$  ?

Yes!

Key results:

- $\hat{\sigma}^2 \xrightarrow{P} \sigma^2$  (HW)

- Continuous mapping theorem: if  $g$  is continuous,

$$X_n \xrightarrow{P} X \Rightarrow g(X_n) \xrightarrow{d} g(X)$$

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