

Asymptotic properties of maximum likelihood estimators

Ciaran Evans

Big results so far

- ▶ **WLLN:** Under certain conditions, $\bar{X}_n \xrightarrow{P} \mu$
- ▶ **CLT:** Under certain conditions, $\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} Z \sim N(0, 1)$

These are nice properties! But I don't just want to estimate the mean. Can we say something similar about maximum likelihood estimates in general?

Key results for the MLE

range of values for θ

(for example, $\theta \in [0, 1]$
for Bernoulli(θ))

Definition: Let Y_1, Y_2, \dots, Y_n be a sample from some distribution with parameter $\theta \in \Theta$. Let $\hat{\theta}_n$ be an estimator constructed from the sample Y_1, \dots, Y_n . We say that $\hat{\theta}_n$ is a **consistent** estimator of θ if, for every $\varepsilon > 0$ and every $\theta \in \Theta$,

$$\lim_{n \rightarrow \infty} P_\theta(|\hat{\theta}_n - \theta| \geq \varepsilon) = 0 \quad \leftarrow \text{convergence in probability}$$

probability if θ is true
parameter

i.e., regardless of true value of θ , $\hat{\theta}_n \xrightarrow{P} \theta$

Key results for the MLE

Let Y_1, Y_2, \dots be iid from a distribution with probability function $f(y|\theta)$, where $\theta \in \mathbb{R}^d$ is the parameter(s) we are trying to estimate.
Let

$$\ell_n(\theta) = \sum_{i=1}^n \log f(Y_i|\theta) \quad \leftarrow \text{log likelihood}$$
$$\hat{\theta}_n = \operatorname{argmax}_{\theta} \ell_n(\theta) \quad \leftarrow \text{MLE}$$

Theorem: Under certain regularity conditions (to be discussed later),

(a) $\hat{\theta}_n \xrightarrow{P} \theta$ (consistency)

(b) $\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(\mathbf{0}, v(\theta))$ (asymptotic normality)

(where we still need to determine what the variance $v(\theta)$ should be!)

Proofs make use of WLLN and CLT!

Key results for the MLE

Theorem (consistency of the MLE): Let Y_1, Y_2, \dots be iid from a distribution with parameter θ , and let $\hat{\theta}_n$ be the MLE constructed from Y_1, \dots, Y_n . Under certain conditions (to be discussed later), $\hat{\theta}_n$ is a **consistent** estimator of θ .

Idea: ① $\frac{1}{n} \ln(\theta) = \frac{1}{n} \sum_{i=1}^n \log f(Y_i | \theta)$
 $\xrightarrow{(w/w)} \mathbb{E} [\log f(Y_i | \theta)]$

- ② Let θ_0 be true value of θ
 $\mathbb{E} [\log f(Y_i | \theta)] < \mathbb{E} [\log f(Y_i | \theta_0)]$ if $\theta \neq \theta_0$
(expected log-likelihood is maximized at true θ_0 !)
- ③ $\frac{1}{n} \ln(\theta)$ is close to $\mathbb{E} [\log f(Y_i | \theta)]$
so maximizing $\frac{1}{n} \ln(\theta)$ is close to
maximizing $\mathbb{E} [\log f(Y_i | \theta)]$
 $\Rightarrow \hat{\theta}_n$ close to θ_0

Key results for the MLE

Theorem (asymptotic normality of the MLE): Let Y_1, Y_2, \dots be iid from a distribution with parameter θ , and let $\hat{\theta}_n$ be the MLE constructed from Y_1, \dots, Y_n . Under certain conditions (to be discussed later),

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(0, v(\theta))$$

(where we still need to determine what the variance $v(\theta)$ should be!)

Idea: $\hat{\theta}_n$ maximizes $l_n(\theta)$. If l_n is differentiable,
 $\hat{\theta}_n$ solves $l'_n(\theta) = \frac{\partial}{\partial \theta} l_n(\theta) = 0$

Consider case when $\theta \in \mathbb{R}$. Taylor expansion:

$$l'_n(\hat{\theta}_n) \approx l'_n(\theta) + l''_n(\theta)(\theta - \hat{\theta}_n)$$

" ($\hat{\theta}_n \rightarrow \theta$ so $\hat{\theta}_n$ close to θ for big n)

$$0 \approx l'_n(\theta) + l''_n(\theta)(\theta - \hat{\theta}_n)$$

$$\Rightarrow \hat{\theta}_n - \theta \approx \frac{l'_n(\theta)}{-l''_n(\theta)}$$

Class activity

Work on the class activity, then we will discuss as a group.

https://sta711-s26.github.io/class_activities/ca_12.html

Class activity

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Bernoulli}(p) \quad \mathbb{E}[Y_i] = p \quad \text{var}(Y_i) = p(1-p)$$

$$\ell_n(p) = \log(p) \left(\sum_{i=1}^n Y_i \right) + \log(1-p) \left(n - \sum_{i=1}^n Y_i \right)$$

$$\ell'_n(p) = \frac{\sum_{i=1}^n Y_i}{p} - \frac{\left(n - \sum_{i=1}^n Y_i \right)}{1-p} = \frac{\sum_{i=1}^n Y_i}{p(1-p)} - \frac{n}{1-p}$$

$$\text{var}(\ell'_n(p)) = \frac{n p (1-p)}{(p(1-p))^2} = \frac{n}{p(1-p)} \quad \mathbb{E}[\ell'_n(p)] = \frac{np}{p(1-p)} - \frac{n}{1-p} = 0$$

$$\ell''_n(p) = -\frac{\sum_{i=1}^n Y_i}{p^2} - \frac{\left(n - \sum_{i=1}^n Y_i \right)}{(1-p)^2} - \mathbb{E}[\ell'_n(p)] = \text{var}(\ell'_n(p))$$

$$-\mathbb{E}[\ell''_n(p)] = -\left(-\frac{np}{p^2} - \frac{(n-np)}{(1-p)^2}\right) = \frac{n}{p} + \frac{n}{(1-p)} = \frac{n}{p(1-p)}$$

$$\hat{p} = \frac{1}{n} \sum_{i=1}^n Y_i \stackrel{\text{var}}{=} \frac{1}{n^2} \cdot \sum_i \text{var}(Y_i) = \frac{1}{n^2} \cdot n \cdot p(1-p) = \frac{p(1-p)}{n} = \frac{1}{\text{var}(\ell'_n(p))}$$