

Maximum likelihood estimation

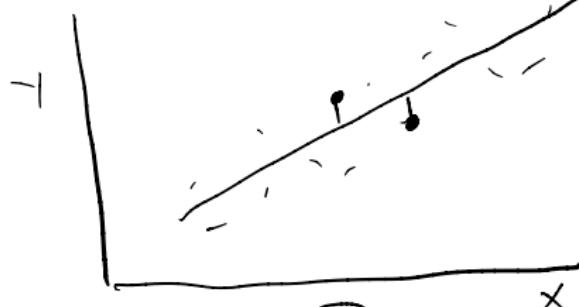
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Fitting a *logistic* regression model?

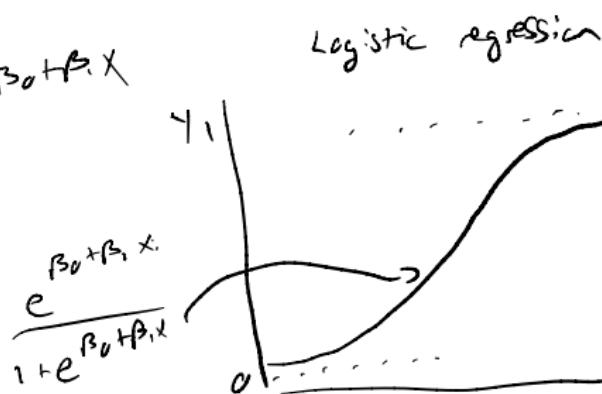
Linear regression: minimize $\sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_{i1} - \cdots - \beta_k X_{ik})^2$

Question: Should we minimize a similar sum of squares for a *logistic* regression model? no

Linear regression



$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$
additive error term



$Y_i | X_i \sim \text{Bernoulli}(p_i)$
 $\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 X_i$

Motivation: likelihoods and estimation

Let $Y \sim \text{Bernoulli}(p)$ be a Bernoulli random variable, with $p \in [0, 1]$. We observe 5 independent samples from this distribution:

$$Y_1 = 1, Y_2 = 1, Y_3 = 0, Y_4 = 0, Y_5 = 1$$

The true value of p is unknown, so two friends propose different guesses for the value of p : 0.3 and 0.7. Which do you think is a "better" guess?

$$\hat{p} = 0.6 \quad (\text{closer to } 0.7)$$

more 1s than 0s \Rightarrow 0.7 better guess than 0.3

$$P(\text{data} \mid p = 0.3) = (0.3)^3 (1-0.3)^2 = 0.013$$

$$P(\text{data} \mid p = 0.7) = (0.7)^3 (1-0.7)^2 = 0.031$$

Intuition: choose value of p that makes data "more likely"

Likelihood

Definition: Let $\mathbf{y} = (Y_1, \dots, Y_n)$ be a sample of n observations, and let $f(\mathbf{y}|\theta)$ denote the joint pdf or pmf of \mathbf{y} , with parameter(s) θ .

The *likelihood function* is

$$\underbrace{L(\theta|\mathbf{y})}_{\substack{\text{function of } \theta, \\ \text{given observed data } \mathbf{y}}} = f(\mathbf{y}|\theta)$$

"probability" of the observed data, if θ is the true parameter

$$L(\theta|\mathbf{y}) \geq 0 \quad \text{since} \quad f(\mathbf{y}|\theta) \geq 0$$

Special case: Y_1, \dots, Y_n are iid

$$\Rightarrow L(\theta|\mathbf{y}) = \prod_{i=1}^n f(Y_i|\theta)$$

Example: Bernoulli data

Let $\gamma_1, \dots, \gamma_n \stackrel{iid}{\sim} \text{Bernoulli}(\rho)$

$$f(y_i | \rho) = \rho^{y_i} (1-\rho)^{1-y_i} \quad y_i \in \{0, 1\}$$

$$\begin{aligned} L(\rho | \gamma_1, \dots, \gamma_n) &= \prod_{i=1}^n f(\gamma_i | \rho) \\ &= \prod_{i=1}^n \rho^{\gamma_i} (1-\rho)^{1-\gamma_i} \\ &= \rho^{\sum_i \gamma_i} (1-\rho)^{n - \sum_i \gamma_i} \end{aligned}$$

ex: $\gamma = (1, 1, 0, 0, 1)$

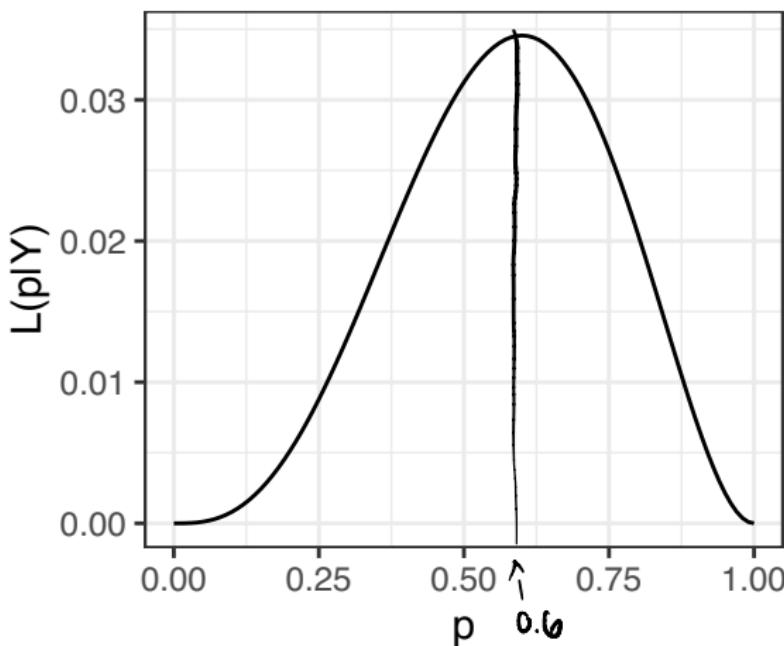
$$L(\rho | \gamma) = \rho^3 (1-\rho)^2$$

Example: Bernoulli data

$Y_1, \dots, Y_5 \stackrel{iid}{\sim} \text{Bernoulli}(p)$, with observed data

$$Y_1 = 1, Y_2 = 1, Y_3 = 0, Y_4 = 0, Y_5 = 1$$

$$L(p|\mathbf{y}) = p^3(1-p)^2$$



Maximum likelihood estimator

Definition: Let $\mathbf{y} = (Y_1, \dots, Y_n)$ be a sample of n observations. The *maximum likelihood estimator* (MLE) is

$$\hat{\theta} = \operatorname{argmax}_{\theta} L(\theta|\mathbf{y})$$

↑
value of θ that maximizes...

Example: $Bernoulli(p)$

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$$

$$L(p|Y) = p^{\sum_i Y_i} (1-p)^{n - \sum_i Y_i}$$

maximize to estimate p :

① Take \log to make life easier:

$$\begin{aligned} l(p|Y) &= \log L(p|Y) \\ &= (\sum_i Y_i) \log p + (n - \sum_i Y_i) \log (1-p) \end{aligned}$$

② Differentiate wrt parameter of interest:

$$\begin{aligned} \frac{\partial}{\partial p} l(p|Y) &= \frac{\sum_i Y_i}{p} - \frac{(n - \sum_i Y_i)}{1-p} \stackrel{\text{set } 0}{=} 0 \\ \Rightarrow \frac{\sum_i Y_i}{p} &= \frac{n - \sum_i Y_i}{1-p} \end{aligned}$$

$$\Rightarrow \hat{p} = \frac{1}{n} \sum_i Y_i = \bar{Y}$$

(sample proportion!)

Example: $Bernoulli(p)$

$$\frac{\partial}{\partial p} \ell(p|y) = \frac{\sum_i y_i}{p} - \frac{(n - \sum_i y_i)}{1-p}$$

check second derivative:

$$\frac{\partial^2}{\partial p^2} \ell(p|y) \bigg|_{p=\bar{p}} = -\frac{\sum_i y_i}{p^2} - \frac{(n - \sum_i y_i)}{(1-p)^2} \bigg|_{p=\bar{p}} < 0$$

$$\Rightarrow \hat{p} = \frac{1}{n} \sum_i y_i = \bar{p} \quad \text{maximizes } \ell(p|y)$$

Example: $N(\theta, 1)$

$$y_1, \dots, y_n \stackrel{iid}{\sim} N(\theta, 1) \quad \theta \in (-\infty, \infty)$$

$$f(y_i | \theta) = \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{1}{2} (y_i - \theta)^2 \right\}$$

$$L(\theta | y) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{1}{2} (y_i - \theta)^2 \right\}$$

$$= (2\pi)^{-\frac{n}{2}} \exp \left\{ -\frac{1}{2} \sum_i (y_i - \theta)^2 \right\}$$

$$\ell(\theta | y) = -\frac{n}{2} \log(2\pi) - \frac{1}{2} \sum_{i=1}^n (y_i - \theta)^2$$

$$\frac{\partial}{\partial \theta} \ell(\theta | y) = -\frac{1}{2} \sum_{i=1}^n 2(y_i - \theta)(-1) = \sum_{i=1}^n (y_i - \theta) = 0$$

$$\Rightarrow \sum_{i=1}^n y_i = n\theta$$

$$\Rightarrow \hat{\theta} = \frac{1}{n} \sum_{i=1}^n y_i = \bar{y}$$

$$\frac{\partial^2}{\partial \theta^2} \ell(\theta | y) = -n < 0 \quad \checkmark$$

\Rightarrow unique maximum