

# Asymptotic properties of maximum likelihood estimators

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## Key results for the MLE

Let  $Y_1, Y_2, \dots$  be iid from a distribution with probability function  $f(y|\theta)$ , where  $\theta \in \mathbb{R}^d$  is the parameter(s) we are trying to estimate.

Let

$$\ell_n(\theta) = \sum_{i=1}^n \log f(Y_i|\theta)$$

$$\hat{\theta}_n = \operatorname{argmax}_{\theta} \ell_n(\theta)$$

**Theorem:** Under certain regularity conditions,

- (a)  $\hat{\theta}_n \xrightarrow{p} \theta$
- (b)  $\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} N(\mathbf{0}, \mathcal{I}_1^{-1}(\theta))$

## Application to regression models

Suppose that  $(\mathbf{x}_1, Y_1), \dots, (\mathbf{x}_n, Y_n)$  are iid from the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mu_i, \sigma^2)$$

$$\mu_i = \mathbf{x}_i^T \boldsymbol{\beta}$$

The MLE is  $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$ . Asymptotic normality of the MLE means that

$$\sqrt{n}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}) \xrightarrow{d} N(\mathbf{0}, \mathcal{I}_1^{-1}(\boldsymbol{\beta}))$$

## Fisher information for the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mathbf{x}_i^T \boldsymbol{\beta}, \sigma^2)$$

Score:  $U(\boldsymbol{\beta}) = \frac{1}{\sigma^2} \mathbf{X}^T (\mathbf{y} - \mathbf{X}\boldsymbol{\beta}) = \frac{1}{\sigma^2} \sum_{i=1}^n (Y_i - \mathbf{x}_i^T \boldsymbol{\beta}) \mathbf{x}_i$

Let  $u_i = (Y_i - \mathbf{x}_i^T \boldsymbol{\beta}) \mathbf{x}_i$

$$\text{Var}(u_i) = \text{Var}\left(\frac{1}{\sigma^2} u_i\right) = \frac{1}{\sigma^4} \text{Var}(u_i)$$

$$\text{Var}(u_i) = \text{Var}(\mathbb{E}[u_i | \mathbf{x}_i]) + \mathbb{E}[\text{Var}(u_i | \mathbf{x}_i)]$$

$$\mathbb{E}[u_i | \mathbf{x}_i] = 0$$

$$\begin{aligned}\text{Var}(u_i | \mathbf{x}_i) &= \mathbb{E}[(u_i - \mathbb{E}[u_i | \mathbf{x}_i])(u_i - \mathbb{E}[u_i | \mathbf{x}_i])^T | \mathbf{x}_i] \\ &= \mathbb{E}[u_i u_i^T | \mathbf{x}_i] \\ &= \mathbb{E}[(Y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2 \mathbf{x}_i \mathbf{x}_i^T | \mathbf{x}_i] \\ &= \mathbb{E}[(Y_i - \mathbf{x}_i^T \boldsymbol{\beta})^2 | \mathbf{x}_i] \mathbf{x}_i \mathbf{x}_i^T \\ &= \text{Var}(Y_i | \mathbf{x}_i) \mathbf{x}_i \mathbf{x}_i^T \\ &\approx \sigma^2 \mathbf{x}_i \mathbf{x}_i^T\end{aligned}$$

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**Score:**  $U(\boldsymbol{\beta}) = \frac{1}{\sigma^2} \mathbf{X}^T (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$

Let  $u_i = (y_i - \mathbf{x}_i^T \boldsymbol{\beta}) \mathbf{x}_i$

$$\mathcal{L}(\boldsymbol{\beta}) = \text{Var}\left(\frac{1}{\sigma^2} u_i\right) = \frac{1}{\sigma^4} \text{Var}(u_i)$$

$$\text{Var}(u_i) = \text{Var}(\mathbb{E}[u_i | \mathbf{x}_i]) + \mathbb{E}[\text{Var}(u_i | \mathbf{x}_i)]$$

$$\begin{aligned}\text{Var}(u_i) &= \text{Var}(0) + \mathbb{E}[\sigma^2 \mathbf{x}_i \mathbf{x}_i^T] \\ &= \sigma^2 \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]\end{aligned}$$

$$\Rightarrow \mathcal{L}(\boldsymbol{\beta}) = \frac{1}{\sigma^2} \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$$

## Fisher information for the linear regression model

$$Y_i | \mathbf{x}_i \sim N(\mathbf{x}_i^T \boldsymbol{\beta}, \sigma^2)$$

**Fisher information:**  $I_1(\boldsymbol{\beta}) = \frac{1}{\sigma^2} \mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$

$$\sqrt{n}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta}) \xrightarrow{d} N(0, \sigma^2 (\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T])^{-1})$$

$$\hat{\boldsymbol{\beta}} \approx N(\boldsymbol{\beta}, \frac{\sigma^2}{n} (\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T])^{-1})$$

Need to estimate  $\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$   
Estimate  $\mathbb{E}[\mathbf{x}_i \mathbf{x}_i^T]$  by  $\hat{\sum}_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T = \frac{1}{n} \mathbf{X}^T \mathbf{X}$   
( $\mathbf{X}$  = design matrix)

$$\Rightarrow \hat{\boldsymbol{\beta}} \approx N(\boldsymbol{\beta}, \hat{\sigma}^2 (\mathbf{X}^T \mathbf{X})^{-1})$$