

STA 712 Homework 6

Due: Tuesday, November 14, 12:00pm (noon) on Canvas.

Instructions: Submit your work as a single PDF.

The EM algorithm

In this problem, we will use the EM algorithm to estimate the parameters in a mixture of two univariate Gaussian distributions.

Let $\theta \in \mathbb{R}^d$ be an unknown parameter we want to estimate. Let $Y = Y_1, \dots, Y_n$ be a set of observed data, and $Z = Z_1, \dots, Z_n$ a set of unobserved latent data. To estimate θ , we want to maximize the likelihood

$$L(\theta; Y) = f_Y(Y|\theta) = \int f_{Y|Z=z}(Y|\theta) f_Z(z) dz$$

However, maximizing this likelihood is challenging when Z is unobserved. Our solution is to alternate between the E and M steps of the EM algorithm:

E step: Let $\theta^{(k)}$ be the current estimate of θ . Calculate

$$Q(\theta|\theta^{(k)}) = \mathbb{E}_{Z|Y, \theta^{(k)}} [\log L(\theta; Z, Y)]$$

M step: $\theta^{(k+1)} = \operatorname{argmax}_{\theta} Q(\theta|\theta^{(k)})$

1. Let $Z_i \sim \text{Bernoulli}(\alpha)$, and $Y_i|Z_i = j \sim N(\mu_j, \sigma_j^2)$. Then our parameter vector of interest is $\theta = (\alpha, \mu_0, \mu_1, \sigma_0^2, \sigma_1^2)$, and the conditional density of $Y_i|Z_i = j$ is

$$f_{Y_i|Z_i=j}(y|\theta) = \frac{1}{\sqrt{2\pi\sigma_j^2}} \exp \left\{ -\frac{1}{2\sigma_j^2} (y - \mu_j)^2 \right\}.$$

We observe data Y_1, \dots, Y_n , and our goal is to estimate θ . We will use the EM algorithm to estimate these parameters.

- (a) Show that the *complete-data* likelihood (i.e., if we were able to observe Z_i) is

$$L(\theta; Z, Y) = \prod_{i=1}^n \alpha^{Z_i} (1 - \alpha)^{1-Z_i} \frac{1}{\sqrt{2\pi\sigma_{Z_i}^2}} \exp \left\{ -\frac{1}{2\sigma_{Z_i}^2} (Y_i - \mu_{Z_i})^2 \right\}$$

- (b) Using (a), show that

$$Q(\theta|\theta^{(k)}) = \sum_{i=1}^n \sum_{j=0}^1 [\log \alpha_j - \frac{1}{2} \log(2\pi\sigma_j^2) - \frac{1}{2\sigma_j^2} (Y_i - \mu_j)^2] P(Z_i = j|Y_i, \theta^{(k)}),$$

where $\alpha_1 = \alpha$ and $\alpha_0 = 1 - \alpha$.

(c) Differentiate $Q(\theta|\theta^{(k)})$ with respect to μ_j to show that

$$\mu_j^{(k+1)} = \frac{\sum_{i=1}^n Y_i P(Z_i = j|Y_i, \theta^{(k)})}{\sum_{i=1}^n P(Z_i = j|Y_i, \theta^{(k)})}$$

(d) Calculate similar update rules for σ_j^2 and α_j .

Fisher information for ZIP models

Recall that for a ZIP model,

$$P(Y_i = y|\gamma, \beta) = \begin{cases} e^{-\lambda_i}(1 - p_i) + p_i & y = 0 \\ \frac{e^{-\lambda_i} \lambda_i^y}{y!} (1 - p_i) & y > 0 \end{cases}$$

with

$$\begin{aligned} \log\left(\frac{p_i}{1 - p_i}\right) &= \gamma^T X_i \\ \log(\lambda_i) &= \beta^T X_i \end{aligned}$$

2. Suppose we observe data $(X_1, Y_1), \dots, (X_n, Y_n)$ and fit a ZIP model, estimating γ and β . One option for testing hypotheses about coefficients in γ and β is to use a Wald test. This relies on the fact that the distribution of $(\hat{\gamma}, \hat{\beta})^T$ is approximately normal, and requires us to calculate the observed information. In this problem, we will calculate the observed information matrix for the ZIP model.

(a) Show that the log likelihood of γ and β is

$$\ell(\gamma, \beta; Y) = \sum_{i: Y_i=0} \log(e^{-\lambda_i}(1 - p_i) + p_i) + \sum_{i: Y_i>0} (Y_i \log \lambda_i - \lambda_i) + \sum_{i: Y_i>0} \log(1 - p_i) - \sum_{i: Y_i>0} \log(Y_i!)$$

(b) Rearrange (a) to show that

$$\begin{aligned} \ell(\gamma, \beta; Y) &= \sum_{i=1}^n \log(\exp\{-e^{\beta^T X_i}\} + \exp\{\gamma^T X_i\}) \mathbb{1}\{Y_i = 0\} + \sum_{i=1}^n (Y_i \beta^T X_i - \exp\{\beta^T X_i\}) \mathbb{1}\{Y_i > 0\} \\ &\quad - \sum_{i=1}^n \log(1 + \exp\{\gamma^T X_i\}) - \sum_{i: Y_i>0} \log(Y_i!) \end{aligned}$$

(c) The score function is

$$U(\gamma, \beta) = \begin{pmatrix} \frac{\partial \ell}{\partial \gamma} \\ \frac{\partial \ell}{\partial \beta} \end{pmatrix},$$

where both $\frac{\partial \ell}{\partial \gamma}$ and $\frac{\partial \ell}{\partial \beta}$ are vectors. Find $\frac{\partial \ell}{\partial \gamma}$ and $\frac{\partial \ell}{\partial \beta}$.

(d) The observed information matrix is

$$\mathcal{J}(\gamma, \beta) = - \begin{pmatrix} \frac{\partial^2 \ell}{\partial \gamma^2} & \frac{\partial^2 \ell}{\partial \gamma \partial \beta} \\ \frac{\partial^2 \ell}{\partial \beta \partial \gamma} & \frac{\partial^2 \ell}{\partial \beta^2} \end{pmatrix}$$

where each entry is itself a matrix. Calculate $\mathcal{J}(\gamma, \beta)$.

Multivariate EDMs

Recall that a multivariate EDM has probability function

$$f(y; \theta, \phi) = a(y, \phi) \exp \left\{ \frac{y^T \theta - \kappa(\theta)}{\phi} \right\},$$

where $\phi > 0$, $\theta, y \in \mathbb{R}^d$, and $\kappa : \mathbb{R}^d \rightarrow \mathbb{R}$. As in a univariate EDM,

$$\frac{\partial \kappa}{\partial \theta} = \mu \quad \frac{\partial \mu}{\partial \theta} = V(\mu),$$

with $\mu = \mathbb{E}[Y] \in \mathbb{R}^d$ and $V(\mu) = \frac{1}{\phi} \text{Var}(Y) \in \mathbb{R}^{d \times d}$.

3. Suppose that $Y \sim \text{Categorical}(\pi_1, \dots, \pi_J)$. Then $\mu = (\pi_1, \dots, \pi_{J-1})^T$,

$$\theta = \left(\log \left(\frac{\pi_1}{1 - \sum_{j=1}^{J-1} \pi_j} \right), \dots, \log \left(\frac{\pi_{J-1}}{1 - \sum_{j=1}^{J-1} \pi_j} \right) \right), \text{ and } \kappa(\theta) = -\log \left(1 - \sum_{j=1}^{J-1} \pi_j \right).$$

(a) By differentiating κ , confirm that $\frac{\partial \kappa}{\partial \theta} = \mu$ for the categorical distribution.

(b) For the categorical distribution, show that

$$V(\mu) = \begin{bmatrix} \pi_1(1 - \pi_1) & -\pi_1\pi_2 & \cdots & -\pi_1\pi_{J-1} \\ -\pi_2\pi_1 & \pi_2(1 - \pi_2) & \cdots & -\pi_2\pi_{J-1} \\ \vdots & \vdots & \ddots & \vdots \\ -\pi_{J-1}\pi_1 & -\pi_{J-1}\pi_2 & \cdots & \pi_{J-1}(1 - \pi_{J-1}) \end{bmatrix}$$