

Bayesian Estimation in Linear Models

STA721 Linear Models Duke University

Merlise Clyde

September 18, 2019

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

In the Bayesian paradigm describe uncertainty about unknown parameters using probability distributions

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

In the Bayesian paradigm describe uncertainty about unknown parameters using probability distributions

- Prior Distribution $p(\beta, \phi)$ describes uncertainty about parameters prior to seeing the data

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

In the Bayesian paradigm describe uncertainty about unknown parameters using probability distributions

- ▶ Prior Distribution $p(\beta, \phi)$ describes uncertainty about parameters prior to seeing the data
- ▶ Posterior Distribution $p(\beta, \phi \mid \mathbf{Y})$ describes uncertainty about the parameters after updating beliefs given the observed data

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

In the Bayesian paradigm describe uncertainty about unknown parameters using probability distributions

- ▶ Prior Distribution $p(\beta, \phi)$ describes uncertainty about parameters prior to seeing the data
- ▶ Posterior Distribution $p(\beta, \phi \mid \mathbf{Y})$ describes uncertainty about the parameters after updating beliefs given the observed data
- ▶ updating rule is based on Bayes Theorem

$$p(\beta, \phi \mid \mathbf{Y}) \propto \mathcal{L}(\beta, \phi)p(\beta, \phi)$$

Bayesian Estimation

Model $\mathbf{Y} = \mathbf{X}\beta + \epsilon$ with $\epsilon \sim N(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$ is equivalent to

$$\mathbf{Y} \sim N(\mathbf{X}\beta, \mathbf{I}_n/\phi)$$

$\phi = 1/\sigma^2$ is the *precision*.

In the Bayesian paradigm describe uncertainty about unknown parameters using probability distributions

- ▶ Prior Distribution $p(\beta, \phi)$ describes uncertainty about parameters prior to seeing the data
- ▶ Posterior Distribution $p(\beta, \phi \mid \mathbf{Y})$ describes uncertainty about the parameters after updating beliefs given the observed data
- ▶ updating rule is based on Bayes Theorem

$$p(\beta, \phi \mid \mathbf{Y}) \propto \mathcal{L}(\beta, \phi) p(\beta, \phi)$$

reweight prior beliefs by likelihood of parameters under observed data

Posterior

Posterior is obtained by conditional distribution theory

Posterior

Posterior is obtained by conditional distribution theory

Let $\theta = (\beta, \phi)^T$

Posterior

Posterior is obtained by conditional distribution theory

Let $\theta = (\beta, \phi)^T$

$$p(\theta \mid \mathbf{Y}) = \frac{p(\mathbf{Y} \mid \theta)p(\theta)}{\int_{\Theta} p(\mathbf{Y} \mid \theta)p(\theta) d\theta}$$

Posterior

Posterior is obtained by conditional distribution theory

Let $\theta = (\beta, \phi)^T$

$$\begin{aligned} p(\theta \mid \mathbf{Y}) &= \frac{p(\mathbf{Y} \mid \theta)p(\theta)}{\int_{\Theta} p(\mathbf{Y} \mid \theta)p(\theta) d\theta} \\ &= \frac{p(\mathbf{Y}, \theta)}{p(\mathbf{Y})} \end{aligned}$$

Posterior

Posterior is obtained by conditional distribution theory

Let $\theta = (\beta, \phi)^T$

$$\begin{aligned} p(\theta \mid \mathbf{Y}) &= \frac{p(\mathbf{Y} \mid \theta)p(\theta)}{\int_{\Theta} p(\mathbf{Y} \mid \theta)p(\theta) d\theta} \\ &= \frac{p(\mathbf{Y}, \theta)}{p(\mathbf{Y})} \end{aligned}$$

$p(\mathbf{Y})$, the normalizing constant, is the marginal distribution of the data.

Posterior

Posterior is obtained by conditional distribution theory

Let $\theta = (\beta, \phi)^T$

$$\begin{aligned} p(\theta \mid \mathbf{Y}) &= \frac{p(\mathbf{Y} \mid \theta)p(\theta)}{\int_{\Theta} p(\mathbf{Y} \mid \theta)p(\theta) d\theta} \\ &= \frac{p(\mathbf{Y}, \theta)}{p(\mathbf{Y})} \end{aligned}$$

$p(\mathbf{Y})$, the normalizing constant, is the marginal distribution of the data.

Easiest to work with Bayes Theorem in proportional form and then identify the normalizing constant.

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \text{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \mathbf{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$
- ▶ $\phi \sim \mathbf{G}(\nu_0/2, SS_0/2)$ with $E(\sigma^2) = SS_0/(\nu_0 - 2)$

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \mathbf{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$
- ▶ $\phi \sim \mathbf{G}(\nu_0/2, SS_0/2)$ with $E(\sigma^2) = SS_0/(\nu_0 - 2)$

$$p(\phi) = \frac{1}{\Gamma(\nu_0/2)} \left(\frac{SS_0}{2} \right)^{\nu_0/2} \phi^{\nu_0/2-1} e^{-\phi SS_0/2}$$

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \mathbf{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$
- ▶ $\phi \sim \mathbf{G}(\nu_0/2, SS_0/2)$ with $E(\sigma^2) = SS_0/(\nu_0 - 2)$

$$p(\phi) = \frac{1}{\Gamma(\nu_0/2)} \left(\frac{SS_0}{2} \right)^{\nu_0/2} \phi^{\nu_0/2-1} e^{-\phi SS_0/2}$$

- ▶ $(\boldsymbol{\beta}, \phi)^T \sim \mathbf{NG}(\mathbf{b}_0, \Phi_0, \nu_0, SS_0)$

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \mathbf{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$
- ▶ $\phi \sim \mathbf{G}(\nu_0/2, SS_0/2)$ with $E(\sigma^2) = SS_0/(\nu_0 - 2)$

$$p(\phi) = \frac{1}{\Gamma(\nu_0/2)} \left(\frac{SS_0}{2} \right)^{\nu_0/2} \phi^{\nu_0/2-1} e^{-\phi SS_0/2}$$

- ▶ $(\boldsymbol{\beta}, \phi)^T \sim \mathbf{NG}(\mathbf{b}_0, \Phi_0, \nu_0, SS_0)$
- ▶ Conjugate “Normal-Gamma” family implies

Prior Distributions

Factor joint prior distribution

$$p(\boldsymbol{\beta}, \phi) = p(\boldsymbol{\beta} \mid \phi)p(\phi)$$

Convenient choice is to take

- ▶ $\boldsymbol{\beta} \mid \phi \sim \mathbf{N}(b_0, \Phi_0^{-1}/\phi)$ where b_0 is the prior mean and Φ_0^{-1}/ϕ is the prior covariance of $\boldsymbol{\beta}$
- ▶ $\phi \sim \mathbf{G}(\nu_0/2, SS_0/2)$ with $E(\sigma^2) = SS_0/(\nu_0 - 2)$

$$p(\phi) = \frac{1}{\Gamma(\nu_0/2)} \left(\frac{SS_0}{2} \right)^{\nu_0/2} \phi^{\nu_0/2-1} e^{-\phi SS_0/2}$$

- ▶ $(\boldsymbol{\beta}, \phi)^T \sim \mathbf{NG}(\mathbf{b}_0, \Phi_0, \nu_0, SS_0)$
- ▶ Conjugate “Normal-Gamma” family implies

$$(\boldsymbol{\beta}, \phi)^T \mid \mathbf{Y} \sim \mathbf{NG}(\mathbf{b}_n, \Phi_n, \nu_n, SS_n)$$

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})}$

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} \\ e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})} e^{-\frac{\phi}{2} (\beta - \mathbf{b}_0)^T \Phi (\beta - \mathbf{b}_0)}$$

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0)} \\ e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})} e^{-\frac{\phi}{2}(\beta - \mathbf{b}_0)^T \Phi(\beta - \mathbf{b}_0)}$$

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2}(\beta - \mathbf{b})^T \Phi(\beta - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2}(\beta^T \Phi \beta - 2\beta^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0)} e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})} e^{-\frac{\phi}{2}(\beta - \mathbf{b}_0)^T \Phi(\beta - \mathbf{b}_0)}$$

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2}(\beta - \mathbf{b})^T \Phi(\beta - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2}(\beta^T \Phi \beta - 2\beta^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

- Expand quadratics and regroup terms

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})} e^{-\frac{\phi}{2} (\beta - \mathbf{b}_0)^T \Phi (\beta - \mathbf{b}_0)}$$

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\beta - \mathbf{b})^T \Phi (\beta - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\beta^T \Phi \beta - 2\beta^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

- ▶ Expand quadratics and regroup terms
- ▶ Read off posterior precision from Quadratic in β

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} e^{-\frac{\phi}{2} (\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X}) (\beta - \hat{\beta})} e^{-\frac{\phi}{2} (\beta - \mathbf{b}_0)^T \Phi (\beta - \mathbf{b}_0)}$$

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\beta - \mathbf{b})^T \Phi (\beta - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\beta^T \Phi \beta - 2\beta^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

- ▶ Expand quadratics and regroup terms
- ▶ Read off posterior precision from Quadratic in β
- ▶ Read off posterior mean from Linear term in β

Finding the Posterior Distribution

Express Likelihood: $\mathcal{L}(\beta, \phi) \propto \phi^{n/2} e^{-\phi \frac{\text{SSE}}{2}} e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})}$

$$p(\beta, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0)} e^{-\frac{\phi}{2}(\beta - \hat{\beta})^T (\mathbf{X}^T \mathbf{X})(\beta - \hat{\beta})} e^{-\frac{\phi}{2}(\beta - \mathbf{b}_0)^T \Phi(\beta - \mathbf{b}_0)}$$

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2}(\beta - \mathbf{b})^T \Phi(\beta - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2}(\beta^T \Phi \beta - 2\beta^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

- ▶ Expand quadratics and regroup terms
- ▶ Read off posterior precision from Quadratic in β
- ▶ Read off posterior mean from Linear term in β
- ▶ will need to complete the quadratic in the posterior mean

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})^T(\mathbf{X}^T\mathbf{X})(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_0)^T\Phi_0(\boldsymbol{\beta}-\mathbf{b}_0)}$$

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})^T(\mathbf{X}^T\mathbf{X})(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_0)^T\Phi_0(\boldsymbol{\beta}-\mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \end{aligned}$$

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})^T(\mathbf{X}^T\mathbf{X})(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_0)^T\Phi_0(\boldsymbol{\beta}-\mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T\mathbf{X}+\Phi_0)\boldsymbol{\beta})} \end{aligned}$$

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})^T(\mathbf{X}^T\mathbf{X})(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_0)^T\Phi_0(\boldsymbol{\beta}-\mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T\mathbf{X}+\Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T(\mathbf{X}^T\mathbf{X}\hat{\boldsymbol{\beta}}+\Phi_0\mathbf{b}_0))} \end{aligned}$$

Expand and Regroup

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})^T(\mathbf{X}^T\mathbf{X})(\boldsymbol{\beta}-\hat{\boldsymbol{\beta}})} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_0)^T\Phi_0(\boldsymbol{\beta}-\mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T\mathbf{X}+\Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T(\mathbf{X}^T\mathbf{X}\hat{\boldsymbol{\beta}}+\Phi_0\mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\hat{\boldsymbol{\beta}}^T\mathbf{X}^T\mathbf{X}\hat{\boldsymbol{\beta}}+\mathbf{b}_0^T\Phi_0\mathbf{b}_0)} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0)}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} \\ e^{-\frac{\phi}{2} (\boldsymbol{\beta}^T (\mathbf{X}^T \mathbf{X} + \Phi_0) \boldsymbol{\beta})}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T (\mathbf{X}^T \mathbf{X} + \Phi_0) \boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1} (\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \boldsymbol{\Phi} (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \boldsymbol{\Phi} \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \boldsymbol{\Phi} \mathbf{b} + \mathbf{b}^T \boldsymbol{\Phi} \mathbf{b}) \right\}$$

Let $\boldsymbol{\Phi}_n = \mathbf{X}^T \mathbf{X} + \boldsymbol{\Phi}_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2} (\boldsymbol{\beta}^T (\mathbf{X}^T \mathbf{X} + \boldsymbol{\Phi}_0) \boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2} (-2\boldsymbol{\beta}^T \boldsymbol{\Phi}_n \boldsymbol{\Phi}_n^{-1} (\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \boldsymbol{\Phi}_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2} (\mathbf{b}_n^T \boldsymbol{\Phi}_n \mathbf{b}_n - \mathbf{b}_n^T \boldsymbol{\Phi}_0 \mathbf{b}_n)} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \boldsymbol{\Phi} (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \boldsymbol{\Phi} \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \boldsymbol{\Phi} \mathbf{b} + \mathbf{b}^T \boldsymbol{\Phi} \mathbf{b}) \right\}$$

Let $\boldsymbol{\Phi}_n = \mathbf{X}^T \mathbf{X} + \boldsymbol{\Phi}_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2} (\text{SSE} + \text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2} (\boldsymbol{\beta}^T (\mathbf{X}^T \mathbf{X} + \boldsymbol{\Phi}_0) \boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2} (-2\boldsymbol{\beta}^T \boldsymbol{\Phi}_n \boldsymbol{\Phi}_n^{-1} (\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \boldsymbol{\Phi}_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2} (\mathbf{b}_n^T \boldsymbol{\Phi}_n \mathbf{b}_n - \mathbf{b}_n^T \boldsymbol{\Phi}_0 \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2} (\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \boldsymbol{\Phi}_0 \mathbf{b}_0)} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T \mathbf{X} + \Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\mathbf{b}_n^T \Phi_n \mathbf{b}_n - \mathbf{b}_n^T \Phi_0 \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T \mathbf{X} + \Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\mathbf{b}_n^T \Phi_n \mathbf{b}_n - \mathbf{b}_n^T \Phi_0 \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T (\Phi_n) \boldsymbol{\beta})} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \Phi (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \Phi \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \Phi \mathbf{b} + \mathbf{b}^T \Phi \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T \mathbf{X} + \Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\mathbf{b}_n^T \Phi_n \mathbf{b}_n - \mathbf{b}_n^T \Phi_0 \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T (\Phi_n) \boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \end{aligned}$$

Identify Hyperparameters and Complete the Quadratic

Quadratic in Normal

$$\exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta} - \mathbf{b})^T \boldsymbol{\Phi} (\boldsymbol{\beta} - \mathbf{b}) \right\} = \exp \left\{ -\frac{\phi}{2} (\boldsymbol{\beta}^T \boldsymbol{\Phi} \boldsymbol{\beta} - 2\boldsymbol{\beta}^T \boldsymbol{\Phi} \mathbf{b} + \mathbf{b}^T \boldsymbol{\Phi} \mathbf{b}) \right\}$$

Let $\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$

$$\begin{aligned} p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) &\propto \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T(\mathbf{X}^T \mathbf{X} + \Phi_0)\boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\mathbf{b}_n^T \Phi_n \mathbf{b}_n - \mathbf{b}_n^T \Phi_0 \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0)} \\ &= \phi^{\frac{n+p+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ &\quad e^{-\frac{\phi}{2}(\boldsymbol{\beta}^T (\Phi_n) \boldsymbol{\beta})} \\ &\quad e^{-\frac{\phi}{2}(-2\boldsymbol{\beta}^T \Phi_n \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0))} \\ &\quad e^{-\frac{\phi}{2}(\mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \end{aligned}$$

Posterior Distribution

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ \phi^{\frac{p}{2}} e^{-\frac{\phi}{2}(\boldsymbol{\beta} - \mathbf{b}_n)^T \Phi_n (\boldsymbol{\beta} - \mathbf{b}_n)}$$

Posterior Distribution

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0+\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}}+\mathbf{b}_0^T \Phi_0 \mathbf{b}_0-\mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ \phi^{\frac{p}{2}} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_n)^T \Phi_n (\boldsymbol{\beta}-\mathbf{b}_n)}$$

$$\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$$

Posterior Distribution

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ \phi^{\frac{p}{2}} e^{-\frac{\phi}{2}(\boldsymbol{\beta} - \mathbf{b}_n)^T \Phi_n (\boldsymbol{\beta} - \mathbf{b}_n)}$$

$$\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$$

$$\mathbf{b}_n = \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0)$$

Posterior Distribution

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE}+\text{SS}_0+\hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}}+\mathbf{b}_0^T \Phi_0 \mathbf{b}_0-\mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ \phi^{\frac{p}{2}} e^{-\frac{\phi}{2}(\boldsymbol{\beta}-\mathbf{b}_n)^T \Phi_n (\boldsymbol{\beta}-\mathbf{b}_n)}$$

$$\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$$

$$\mathbf{b}_n = \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0)$$

Posterior Distribution

$$\boldsymbol{\beta} \mid \phi, \mathbf{Y} \sim \text{N}(\mathbf{b}_n, (\phi \Phi_n)^{-1})$$

Posterior Distribution

$$p(\boldsymbol{\beta}, \phi \mid \mathbf{Y}) \propto \phi^{\frac{n+\nu_0}{2}-1} e^{-\frac{\phi}{2}(\text{SSE} + \text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n)} \\ \phi^{\frac{p}{2}} e^{-\frac{\phi}{2}(\boldsymbol{\beta} - \mathbf{b}_n)^T \Phi_n (\boldsymbol{\beta} - \mathbf{b}_n)}$$

$$\Phi_n = \mathbf{X}^T \mathbf{X} + \Phi_0$$

$$\mathbf{b}_n = \Phi_n^{-1}(\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \Phi_0 \mathbf{b}_0)$$

Posterior Distribution

$$\boldsymbol{\beta} \mid \phi, \mathbf{Y} \sim \mathbf{N}(\mathbf{b}_n, (\phi \Phi_n)^{-1})$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{n + \nu_0}{2}, \frac{\text{SSE} + \text{SS}_0 + \hat{\boldsymbol{\beta}}^T \mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} + \mathbf{b}_0^T \Phi_0 \mathbf{b}_0 - \mathbf{b}_n^T \Phi_n \mathbf{b}_n}{2}\right)$$

Marginal Distribution from Normal–Gamma

Theorem

Let $\boldsymbol{\theta} \mid \phi \sim N(m, \frac{1}{\phi}\Sigma)$ and $\phi \sim \mathbf{G}(\nu/2, \nu\hat{\sigma}^2/2)$. Then $\boldsymbol{\theta}$ ($p \times 1$) has a p dimensional multivariate t distribution

$$\boldsymbol{\theta} \sim t_{\nu}(m, \hat{\sigma}^2\Sigma)$$

with density

$$p(\boldsymbol{\theta}) \propto \left[1 + \frac{1}{\nu} \frac{(\boldsymbol{\theta} - m)^T \Sigma^{-1} (\boldsymbol{\theta} - m)}{\hat{\sigma}^2} \right]^{-\frac{p+\nu}{2}}$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi)p(\phi) d\phi$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi) p(\phi) d\phi$

$$p(\boldsymbol{\theta}) \propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m})} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi) p(\phi) d\phi$

$$\begin{aligned} p(\boldsymbol{\theta}) &\propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m})} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{p/2} \phi^{\nu/2-1} e^{-\phi \frac{(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m}) + \nu \hat{\sigma}^2}{2}} d\phi \end{aligned}$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi)p(\phi) d\phi$

$$\begin{aligned} p(\boldsymbol{\theta}) &\propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m})} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{p/2} \phi^{\nu/2-1} e^{-\phi \frac{(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m}) + \nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{\frac{p+\nu}{2}-1} e^{-\phi \frac{(\boldsymbol{\theta}-\mathbf{m})^T \Sigma^{-1}(\boldsymbol{\theta}-\mathbf{m}) + \nu \hat{\sigma}^2}{2}} d\phi \end{aligned}$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi)p(\phi) d\phi$

$$\begin{aligned} p(\boldsymbol{\theta}) &\propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m)} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{p/2} \phi^{\nu/2-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{\frac{p+\nu}{2}-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &= \Gamma((p+\nu)/2) \left(\frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2} \right)^{-\frac{p+\nu}{2}} \end{aligned}$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi)p(\phi) d\phi$

$$\begin{aligned} p(\boldsymbol{\theta}) &\propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m)} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{p/2} \phi^{\nu/2-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{\frac{p+\nu}{2}-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &= \Gamma((p+\nu)/2) \left(\frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2} \right)^{-\frac{p+\nu}{2}} \\ &\propto \left((\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2 \right)^{-\frac{p+\nu}{2}} \end{aligned}$$

Derivation

Marginal density $p(\boldsymbol{\theta}) = \int p(\boldsymbol{\theta} \mid \phi)p(\phi) d\phi$

$$\begin{aligned} p(\boldsymbol{\theta}) &\propto \int |\Sigma/\phi|^{-1/2} e^{-\frac{\phi}{2}(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m)} \phi^{\nu/2-1} e^{-\phi \frac{\nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{p/2} \phi^{\nu/2-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &\propto \int \phi^{\frac{p+\nu}{2}-1} e^{-\phi \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2}} d\phi \\ &= \Gamma((p+\nu)/2) \left(\frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2}{2} \right)^{-\frac{p+\nu}{2}} \\ &\propto \left((\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m) + \nu \hat{\sigma}^2 \right)^{-\frac{p+\nu}{2}} \\ &\propto \left(1 + \frac{1}{\nu} \frac{(\boldsymbol{\theta}-m)^T \Sigma^{-1}(\boldsymbol{\theta}-m)}{\hat{\sigma}^2} \right)^{-\frac{p+\nu}{2}} \end{aligned}$$

Marginal Posterior Distribution of β

$$\beta \mid \phi, \mathbf{Y} \sim \mathbf{N}(\mathbf{b}_n, \phi^{-1} \Phi_n^{-1})$$

Marginal Posterior Distribution of β

$$\beta \mid \phi, \mathbf{Y} \sim \mathbf{N}(\mathbf{b}_n, \phi^{-1} \Phi_n^{-1})$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{SS_n}{2}\right)$$

Marginal Posterior Distribution of β

$$\beta \mid \phi, \mathbf{Y} \sim \mathbf{N}(\mathbf{b}_n, \phi^{-1} \Phi_n^{-1})$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{SS_n}{2}\right)$$

Let $\hat{\sigma}^2 = SS_n/\nu_n$ (Bayesian MSE)

Marginal Posterior Distribution of β

$$\beta \mid \phi, \mathbf{Y} \sim \mathbf{N}(\mathbf{b}_n, \phi^{-1} \Phi_n^{-1})$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{SS_n}{2}\right)$$

Let $\hat{\sigma}^2 = SS_n/\nu_n$ (Bayesian MSE)

Then the marginal posterior distribution of β is

$$\beta \mid \mathbf{Y} \sim t_{\nu_n}(\mathbf{b}_n, \hat{\sigma}^2 \Phi_n^{-1})$$

Marginal Posterior Distribution of β

$$\begin{aligned}\beta \mid \phi, \mathbf{Y} &\sim \mathbf{N}(\mathbf{b}_n, \phi^{-1} \Phi_n^{-1}) \\ \phi \mid \mathbf{Y} &\sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{SS_n}{2}\right)\end{aligned}$$

Let $\hat{\sigma}^2 = SS_n/\nu_n$ (Bayesian MSE)

Then the marginal posterior distribution of β is

$$\beta \mid \mathbf{Y} \sim t_{\nu_n}(\mathbf{b}_n, \hat{\sigma}^2 \Phi_n^{-1})$$

Any linear combination $\lambda^T \beta$

$$\lambda^T \beta \mid \mathbf{Y} \sim t_{\nu_n}(\lambda^T \mathbf{b}_n, \hat{\sigma}^2 \lambda^T \Phi_n^{-1} \lambda)$$

has a univariate t distribution with ν_n degrees of freedom

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \beta, \phi \sim \mathcal{N}(\mathbf{X}^*\beta, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given β and ϕ

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \boldsymbol{\beta}, \phi \sim \mathcal{N}(\mathbf{X}^* \boldsymbol{\beta}, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given $\boldsymbol{\beta}$ and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \boldsymbol{\beta}, \phi \sim \mathcal{N}(\mathbf{X}^*\boldsymbol{\beta}, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given $\boldsymbol{\beta}$ and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

$\mathbf{Y}^* = \mathbf{X}^*\boldsymbol{\beta} + \boldsymbol{\epsilon}^*$ and $\boldsymbol{\epsilon}^*$ is independent of \mathbf{Y} given ϕ

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \boldsymbol{\beta}, \phi \sim \mathcal{N}(\mathbf{X}^*\boldsymbol{\beta}, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given $\boldsymbol{\beta}$ and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

$\mathbf{Y}^* = \mathbf{X}^*\boldsymbol{\beta} + \boldsymbol{\epsilon}^*$ and $\boldsymbol{\epsilon}^*$ is independent of \mathbf{Y} given ϕ

$$\mathbf{X}^*\boldsymbol{\beta} + \boldsymbol{\epsilon}^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^*\mathbf{b}_n, (\mathbf{X}^*\boldsymbol{\Phi}_n^{-1}\mathbf{X}^{*T} + \mathbf{I})/\phi)$$

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \beta, \phi \sim \mathcal{N}(\mathbf{X}^*\beta, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given β and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

$\mathbf{Y}^* = \mathbf{X}^*\beta + \epsilon^*$ and ϵ^* is independent of \mathbf{Y} given ϕ

$$\mathbf{X}^*\beta + \epsilon^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^*\mathbf{b}_n, (\mathbf{X}^*\Phi_n^{-1}\mathbf{X}^{*T} + \mathbf{I})/\phi)$$

$$\mathbf{Y}^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^*\mathbf{b}_n, (\mathbf{X}^*\Phi_n^{-1}\mathbf{X}^{*T} + \mathbf{I})/\phi)$$

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \beta, \phi \sim \mathcal{N}(\mathbf{X}^*\beta, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given β and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

$\mathbf{Y}^* = \mathbf{X}^*\beta + \epsilon^*$ and ϵ^* is independent of \mathbf{Y} given ϕ

$$\mathbf{X}^*\beta + \epsilon^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^*\mathbf{b}_n, (\mathbf{X}^*\Phi_n^{-1}\mathbf{X}^{*T} + \mathbf{I})/\phi)$$

$$\mathbf{Y}^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^*\mathbf{b}_n, (\mathbf{X}^*\Phi_n^{-1}\mathbf{X}^{*T} + \mathbf{I})/\phi)$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{\hat{\sigma}^2\nu_n}{2}\right)$$

Predictive Distribution

Suppose $\mathbf{Y}^* \mid \boldsymbol{\beta}, \phi \sim \mathcal{N}(\mathbf{X}^* \boldsymbol{\beta}, \mathbf{I}/\phi)$ and is conditionally independent of \mathbf{Y} given $\boldsymbol{\beta}$ and ϕ

What is the predictive distribution of $\mathbf{Y}^* \mid \mathbf{Y}$?

$\mathbf{Y}^* = \mathbf{X}^* \boldsymbol{\beta} + \boldsymbol{\epsilon}^*$ and $\boldsymbol{\epsilon}^*$ is independent of \mathbf{Y} given ϕ

$$\mathbf{X}^* \boldsymbol{\beta} + \boldsymbol{\epsilon}^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^* \mathbf{b}_n, (\mathbf{X}^* \boldsymbol{\Phi}_n^{-1} \mathbf{X}^{*T} + \mathbf{I})/\phi)$$

$$\mathbf{Y}^* \mid \phi, \mathbf{Y} \sim \mathcal{N}(\mathbf{X}^* \mathbf{b}_n, (\mathbf{X}^* \boldsymbol{\Phi}_n^{-1} \mathbf{X}^{*T} + \mathbf{I})/\phi)$$

$$\phi \mid \mathbf{Y} \sim \mathbf{G}\left(\frac{\nu_n}{2}, \frac{\hat{\sigma}^2 \nu_n}{2}\right)$$

$$\mathbf{Y}^* \mid \mathbf{Y} \sim t_{\nu_n}(\mathbf{X}^* \mathbf{b}_n, \hat{\sigma}^2 (\mathbf{I} + \mathbf{X}^* \boldsymbol{\Phi}_n^{-1} \mathbf{X}^T))$$

Alternative Derivation

Conditional Distribution:

$$f(\mathbf{Y}^* | \mathbf{Y}) = \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})}$$

Alternative Derivation

Conditional Distribution:

$$\begin{aligned} f(\mathbf{Y}^* | \mathbf{Y}) &= \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})} \\ &= \frac{\int \int f(\mathbf{Y}^*, \mathbf{Y} | \beta, \phi) p(\beta, \phi) d\beta d\phi}{f(\mathbf{Y})} \end{aligned}$$

Alternative Derivation

Conditional Distribution:

$$\begin{aligned}f(\mathbf{Y}^* | \mathbf{Y}) &= \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^*, \mathbf{Y} | \beta, \phi) p(\beta, \phi) d\beta d\phi}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^* | \beta, \phi) f(\mathbf{Y} | \beta, \phi) p(\beta, \phi) d\beta d\phi}{f(\mathbf{Y})}\end{aligned}$$

Alternative Derivation

Conditional Distribution:

$$\begin{aligned}f(\mathbf{Y}^* | \mathbf{Y}) &= \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^*, \mathbf{Y} | \beta, \phi) p(\beta, \phi) d\beta d\phi}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^* | \beta, \phi) f(\mathbf{Y} | \beta, \phi) p(\beta, \phi) d\beta d\phi}{f(\mathbf{Y})} \\&= \iint f(\mathbf{Y}^* | \beta, \phi) p(\beta, \phi | \mathbf{Y}) d\beta d\phi\end{aligned}$$

Alternative Derivation

Conditional Distribution:

$$\begin{aligned}f(\mathbf{Y}^* | \mathbf{Y}) &= \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^*, \mathbf{Y} | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi) d\boldsymbol{\beta} d\phi}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^* | \boldsymbol{\beta}, \phi) f(\mathbf{Y} | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi) d\boldsymbol{\beta} d\phi}{f(\mathbf{Y})} \\&= \iint f(\mathbf{Y}^* | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi | \mathbf{Y}) d\boldsymbol{\beta} d\phi\end{aligned}$$

$$\mathbf{Y}^* = \mathbf{X}^* \boldsymbol{\beta} + \boldsymbol{\epsilon}^* | \mathbf{Y}, \phi \sim N(\mathbf{X}^* \mathbf{b}_n, \phi^{-1}(\mathbf{I} + \mathbf{X}^* \boldsymbol{\Phi}_n \mathbf{X}^{*T}))$$

Alternative Derivation

Conditional Distribution:

$$\begin{aligned}f(\mathbf{Y}^* | \mathbf{Y}) &= \frac{f(\mathbf{Y}^*, \mathbf{Y})}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^*, \mathbf{Y} | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi) d\boldsymbol{\beta} d\phi}{f(\mathbf{Y})} \\&= \frac{\iint f(\mathbf{Y}^* | \boldsymbol{\beta}, \phi) f(\mathbf{Y} | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi) d\boldsymbol{\beta} d\phi}{f(\mathbf{Y})} \\&= \iint f(\mathbf{Y}^* | \boldsymbol{\beta}, \phi) p(\boldsymbol{\beta}, \phi | \mathbf{Y}) d\boldsymbol{\beta} d\phi\end{aligned}$$

$$\mathbf{Y}^* = \mathbf{X}^* \boldsymbol{\beta} + \boldsymbol{\epsilon}^* | \mathbf{Y}, \phi \sim \mathcal{N}(\mathbf{X}^* \mathbf{b}_n, \phi^{-1}(\mathbf{I} + \mathbf{X}^* \boldsymbol{\Phi}_n \mathbf{X}^{*T}))$$

Use result about Marginals of Normal-Gamma family to integrate out ϕ

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations
- ▶ Simple updating in terms of sufficient statistics “weighted average”

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations
- ▶ Simple updating in terms of sufficient statistics “weighted average”
- ▶ Interpretation as prior samples - prior sample size

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations
- ▶ Simple updating in terms of sufficient statistics “weighted average”
- ▶ Interpretation as prior samples - prior sample size
- ▶ Elicitation of prior through imaginary or historical data

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations
- ▶ Simple updating in terms of sufficient statistics “weighted average”
- ▶ Interpretation as prior samples - prior sample size
- ▶ Elicitation of prior through imaginary or historical data
- ▶ limiting “non-proper” form recovers MLEs

Conjugate Priors

Definition

A class of prior distributions \mathcal{P} for θ is conjugate for a sampling model $p(y \mid \theta)$ if for every $p(\theta) \in \mathcal{P}$, $p(\theta \mid \mathbf{Y}) \in \mathcal{P}$.

Advantages:

- ▶ Closed form distributions for most quantities; bypass MCMC for calculations
- ▶ Simple updating in terms of sufficient statistics “weighted average”
- ▶ Interpretation as prior samples - prior sample size
- ▶ Elicitation of prior through imaginary or historical data
- ▶ limiting “non-proper” form recovers MLEs

Choice of conjugate prior?

Unit Information Prior

Unit information prior $\beta \mid \phi \sim \mathbf{N}(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE
- ▶ “average information” in one observation is $\phi \mathbf{X}^T \mathbf{X} / n$

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE
- ▶ “average information” in one observation is $\phi \mathbf{X}^T \mathbf{X} / n$
- ▶ center prior at MLE and base covariance on the information in “1” observation

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1}/\phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE
- ▶ “average information” in one observation is $\phi \mathbf{X}^T \mathbf{X}/n$
- ▶ center prior at MLE and base covariance on the information in “1” observation
- ▶ Posterior mean

$$\frac{n}{1+n} \hat{\beta} + \frac{1}{1+n} \hat{\beta} = \hat{\beta}$$

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE
- ▶ “average information” in one observation is $\phi \mathbf{X}^T \mathbf{X} / n$
- ▶ center prior at MLE and base covariance on the information in “1” observation
- ▶ Posterior mean

$$\frac{n}{1+n} \hat{\beta} + \frac{1}{1+n} \hat{\beta} = \hat{\beta}$$

- ▶ Posterior Distribution

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\hat{\beta}, \frac{n}{1+n} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

Unit Information Prior

Unit information prior $\beta \mid \phi \sim N(\hat{\beta}, n(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

- ▶ Fisher Information is $\phi \mathbf{X}^T \mathbf{X}$ based on a sample of n observations
- ▶ Inverse Fisher information is covariance matrix of MLE
- ▶ “average information” in one observation is $\phi \mathbf{X}^T \mathbf{X} / n$
- ▶ center prior at MLE and base covariance on the information in “1” observation
- ▶ Posterior mean

$$\frac{n}{1+n} \hat{\beta} + \frac{1}{1+n} \hat{\beta} = \hat{\beta}$$

- ▶ Posterior Distribution

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\hat{\beta}, \frac{n}{1+n} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

Cannot represent real prior beliefs; double use of data

Zellner's g -prior

$$\text{Zellner's } g\text{-prior(s) } \beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$$

Zellner's g -prior

Zellner's g -prior(s) $\beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\frac{g}{1+g} \hat{\beta} + \frac{1}{1+g} \mathbf{b}_0, \frac{g}{1+g} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

Zellner's g -prior

Zellner's g -prior(s) $\beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\frac{g}{1+g} \hat{\beta} + \frac{1}{1+g} \mathbf{b}_0, \frac{g}{1+g} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

- Invariance: Require posterior of $\mathbf{X}\beta$ equal the posterior of $\mathbf{X}\mathbf{H}\alpha$

Zellner's g -prior

Zellner's g -prior(s) $\beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\frac{g}{1+g} \hat{\beta} + \frac{1}{1+g} \mathbf{b}_0, \frac{g}{1+g} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

- ▶ Invariance: Require posterior of $\mathbf{X}\beta$ equal the posterior of $\mathbf{X}\mathbf{H}\alpha$ ($\mathbf{a}_0 = \mathbf{H}^{-1}\mathbf{b}_0$) (take $\mathbf{b}_0 = \mathbf{0}$)
- ▶ Choice of g ?

Zellner's g -prior

Zellner's g -prior(s) $\beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\frac{g}{1+g} \hat{\beta} + \frac{1}{1+g} \mathbf{b}_0, \frac{g}{1+g} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

- ▶ Invariance: Require posterior of $\mathbf{X}\beta$ equal the posterior of $\mathbf{X}\mathbf{H}\alpha$ ($\mathbf{a}_0 = \mathbf{H}^{-1}\mathbf{b}_0$) (take $\mathbf{b}_0 = \mathbf{0}$)
- ▶ Choice of g ?
- ▶ $\frac{g}{1+g}$ weight given to the data

Zellner's g -prior

Zellner's g -prior(s) $\beta \mid \phi \sim N(\mathbf{b}_0, g(\mathbf{X}^T \mathbf{X})^{-1} / \phi)$

$$\beta \mid \mathbf{Y}, \phi \sim N \left(\frac{g}{1+g} \hat{\beta} + \frac{1}{1+g} \mathbf{b}_0, \frac{g}{1+g} (\mathbf{X}^T \mathbf{X})^{-1} \phi^{-1} \right)$$

- ▶ Invariance: Require posterior of $\mathbf{X}\beta$ equal the posterior of $\mathbf{X}\mathbf{H}\alpha$ ($\mathbf{a}_0 = \mathbf{H}^{-1}\mathbf{b}_0$) (take $\mathbf{b}_0 = \mathbf{0}$)
- ▶ Choice of g ?
- ▶ $\frac{g}{1+g}$ weight given to the data
- ▶ Fixed g effect does not vanish as $n \rightarrow \infty$
- ▶ Use $g = n$ or place a prior distribution on g

Shrinkage

Posterior mean under g -prior with $\mathbf{b}_0 = 0$ $\frac{g}{1+g}\hat{\beta}$

