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Stat 430 - Karin Dorman
HW 1: Sep 2, 2010

1.

(a) The sample space is defined as

$$\Omega = \{000, 010, 001, 011, 100, 110, 101, 111\}$$

where each element $\omega \in \Omega$ corresponds to the 3 bits received.

(b) The event that 1 was transmitted and received correctly is $C = \{011, 110, 101, 111\}$.

(c) The set $A = \{A_1 = \{111\}, A_2 = \{101, 110, 011\}, A_3 = \{010, 001, 100\}, A_4 = \{000\}\}$ partitions the sample space Ω based on the number of mistransmitted bits.

(d) The number of mistransmitted bits follows a binomial distribution.

2.

(a) $P = \frac{7}{18} \cdot \frac{6}{17} + \frac{8}{18} \cdot \frac{7}{17} + \frac{9}{18} \cdot \frac{8}{17} = \frac{(7 \cdot 6) + (8 \cdot 7) + (9 \cdot 8)}{18 \cdot 17} = \frac{42 + 56 + 72}{306} = \frac{170}{306} = 0.5555556$.

(b) There field is $120 \cdot 53.3 = 6396$ square yards in size. Each 20x20 yard square contains 400 square yards, so there are $6396/400 = 15.99$ squares in the field. Since there are 18 boxes, we would expect to see $18/15.99 = 1.125704$ boxes in each 20x20 square of the field.

(c) It is reasonable to expect that the number X of boxes in a 20x20 square follows a Poisson distribution with $\lambda = 1.126$. The following simulation in R supports this choice.

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> rpois(16, 1.126)
[1] 2 0 0 0 2 1 1 3 1 2 1 1 2 1 1 1
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3.

Define the following events.

- C : The probability that a claim is submitted
- H : The probability that a client is high risk
- M : The probability that a client is medium risk
- L : The probability that a client is low risk

We are given the following probabilities.

- $P(H) = 0.1$
- $P(M) = 0.2$
- $P(L) = 0.7$
- $P(C|H) = 0.02$
- $P(C|M) = 0.01$
- $P(C|L) = 0.00125$

We are asked to find $P(H|C)$. Using Bayes' rule, we have

$$\begin{aligned}
 P(H|C) &= \frac{P(C|H)P(H)}{P(C|H)P(H) + P(C|M)P(M) + P(C|L)P(L)} \\
 &= \frac{(.02)(.1)}{(.02)(.1) + (.01)(.2) + (.00125)(.7)} \\
 &= \frac{0.002}{0.002 + 0.002 + 0.00875} \\
 &= 0.4102564
 \end{aligned}$$

4.

- (a) If events A and B are disjoint when $P(A) > 0$ and $P(B) > 0$, then A and B are independent.
- (b) The events A and B are not independent when $A \subset B$.
- (c) If $P(A \cup B) = P(A)P(B^C) + P(B)$, then events A and B are independent.

5.

Let the random variable X be defined as the probability that team A will win a "Best of n " tournament. X follows a binomial distribution. When $n = 5$, X is computed as follows.

$$X = \binom{n}{k} p^k (1-p)^{n-k} = \binom{5}{3} (.4)^3 (.6)^2 = 0.2304$$

When $n = 7$, X is computed as follows.

$$X = \binom{n}{k} p^k (1-p)^{n-k} = \binom{7}{4} (.4)^4 (.6)^3 = 0.193536$$

Therefore, team A would do well to elect a "Best of 5" tournament.

6.

To show that $F(x)$ is a CDF, we must demonstrate the following properties.

- The left limit of F is 0, or in other words, $\lim_{x \rightarrow 0} F(x) = 0$. Observe the following.

$$\lim_{x \rightarrow 0} 1 - \exp\{-\alpha x^\beta\} = 1 - e^0 = 1 - 1 = 0$$

- The right limit of F is 1, or in other words, $\lim_{x \rightarrow \infty} F(x) = 1$. Observe the following.

$$\lim_{x \rightarrow \infty} 1 - \exp\{-\alpha x^\beta\} = 1 - \frac{1}{e^\infty} = 1 - 0 = 1$$

- The function is increasing monotonically. Observe the following.

$$\frac{d}{dx} [1 - \exp\{-\alpha x^\beta\}] = 0 - [\exp\{-\alpha x^\beta\} \cdot (-\alpha \beta x^{\beta-1})] = \alpha \beta x^{\beta-1} \exp\{-\alpha x^\beta\}$$

This function, representing the slope of the F , is always positive. Therefore F is strictly monotonically increasing.

- The function is right continuous, or in other words, $\lim_{x \rightarrow x_o^+} F(x) = F(x_o)$. Because F is defined within the specified domain and is (infinitely) differentiable within the specified domain, it is also continuous.

The density, as derived above, is as follows.

$$\frac{d}{dx} F(x) = f(x) = \alpha \beta x^{\beta-1} \exp\{-\alpha x^\beta\}$$

7.

The exponential distribution has a PDF defined as follows.

$$f(x) = \lambda e^{-\lambda x}, x \geq 0$$

The corresponding CDF is defined as follows.

$$F(x) = \int_0^x \lambda e^{-\lambda u} du = 1 - e^{-\lambda x}$$

The upper quartile is the value $q_{.75}$ such that $F(q_{.75}) = .75$. We can determine this value by finding the inverse $F^{-1}(\cdot)$ of $F(\cdot)$ and plugging in .75. The inverse of the CDF is

$$F^{-1}(p) = \frac{-\ln(1-p)}{\lambda}$$

so the upper quartile is

$$q_{.75} = F^{-1}(.75) = \frac{-\ln(.25)}{\lambda} = \frac{1}{\lambda} \cdot 1.386294$$

8.

Observe the following.

$$\begin{aligned} P(|X - \mu| \leq 0.675\sigma) &= 0.5 \Leftrightarrow P(-0.675\sigma \leq X - \mu \leq 0.675\sigma) = 0.5 \\ &\Leftrightarrow P\left(-0.675 \leq \frac{X - \mu}{\sigma} \leq 0.675\right) = 0.5 \end{aligned}$$

The central term in the last equation follows a standard normal distribution, so we can calculate the probability as follows.

$$\begin{aligned} P\left(-0.675 \leq \frac{X - \mu}{\sigma} \leq 0.675\right) &= P\left(\frac{X - \mu}{\sigma} \leq 0.675\right) - P\left(\frac{X - \mu}{\sigma} \leq -0.675\right) \\ &= (0.75016) - (0.24984) \\ &= 0.50032 \end{aligned}$$

9.

Recall that the normal distribution is defined as follows.

$$f(y) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(y - \mu)^2}{2\sigma^2}\right\}$$

Now consider that $y(z) = e^z \Rightarrow z = y^{-1}(y) = \ln y$. Applying this change of variable we have

$$\begin{aligned} f_Y(y) &= f_Z(y^{-1}(y)) \cdot \frac{d}{dy} y^{-1}(y) \\ &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(\ln y - \mu)^2}{2\sigma^2}\right\} \cdot \frac{d}{dy} \ln y \\ &= \frac{1}{y\sqrt{2\pi}\sigma} \exp\left\{-\frac{(\ln y - \mu)^2}{2\sigma^2}\right\} \end{aligned}$$

which is the log-normal density.