

Exact Inference

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Outline



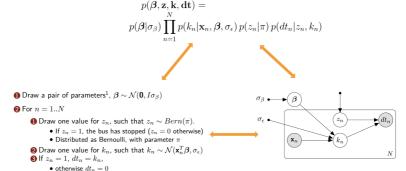
- Recap
- Analytical Derivation (Bayes' theorem)
- Variable Elimination
- Belief Propagation



- Representation (weeks 1-4)
- ullet Modelling toolbox (weeks 5-8, +12)
- Inference (weeks 9-11)

DTU

- Representation (weeks 1-4)
 - PGM basics
 - Conditional independence
 - Generative processes
 - Joint probability distribution
 - Priors and likelihood
 - Factorization





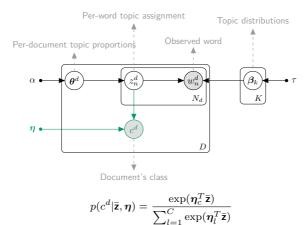
- Representation (weeks 1-4)
- \bullet Modelling toolbox (weeks 5-8, +12)
 - Mixture models
 - Different likelihoods (Gaussian, Poisson, etc.)
 - Link functions (log, softmax, Probit, etc.)
 - Non-linear relationships (e.g. using neural networks)
 - Discrete vs. continuous target variables
 - Heteroscedastic models
 - Hierarchical models
 - Temporal models (continuous and discrete)
 - Topic modelling (discrete data; e.g. text corpora)
 - Bayesian non-parametric models (week 12)
- Inference (weeks 9-11)



- Representation (weeks 1-4)
- Modelling toolbox (weeks 5-8, +12)
 - Bayesian Gaussian Mixture models
 - Bayesian Linear regression
 - Poisson regression
 - Heteroscedastic regression
 - Bayesian Logistic regression
 - Bayesian Probit regression
 - Hierarchical Logistic regression
 - Autoregressive models
 - Linear dynamical systems (e.g. Kalman filter)
 - Hidden Markov models
 - Latent Dirichlet allocation
 - Gaussian processes (week 12)
- Inference (weeks 9-11)



- Representation (weeks 1-4)
- Modelling toolbox (weeks 5-8, +12)
 - Mix & Match (e.g. LDA + (Logistic) Regression = Supervised LDA)

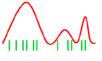


• Inference (weeks 9-11)



- Representation (weeks 1-4)
- Modelling toolbox (weeks 5-8, +12)
- Inference (weeks 9-11)
 - Exact inference
 - Analytical (Bayes' rule)
 - Variable elimination
 - Belief propagation
 - Approximate inference
 - Stochastic methods Markov chain Monte Carlo (MCMC)
 - Deterministic methods Variational inference (VI)







Monte Carlo

Variational

Exact inference



- Computation of the exact posterior probability distribution over the variables of interest
 - Best possible solution, given data and specification
- Analytical derivations:
 - High computational efficiency
 - However, quite often, not possible at all (when it is not possible, we say that it is intractable)
- Algorithmic methods
 - Variable elimination
 - Message passing



- The Gaussian form has very nice properties¹
- A multivariate Gaussian with d dimensions is:

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2} |\boldsymbol{\Sigma}|^{1/2}} e^{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})}$$

- ullet The inverse of the covariance matrix is $oldsymbol{\Lambda} = oldsymbol{\Sigma}^{-1}$ is called *precision matrix*
- If you have a marginal Gaussian distribution for x and a conditional Gaussian distribution of y given x in the form

$$\begin{split} p(\mathbf{x}) &= \mathcal{N}(\mathbf{x}|\pmb{\mu}, \pmb{\Lambda}^{-1}) \\ p(\mathbf{y}|\mathbf{x}) &= \mathcal{N}(\mathbf{y}|\mathbf{A}\mathbf{x} + \mathbf{b}, \mathbf{L}^{-1}) \end{split}$$

ullet Then the marginal distribution for ${f y}$ and a conditional Gaussian distribution of ${f x}$ given ${f y}$ have the form

$$p(\mathbf{y}) = \mathcal{N}(\mathbf{y}|\mathbf{A}\boldsymbol{\mu} + \mathbf{b}, \mathbf{L}^{-1} + \mathbf{A}\boldsymbol{\Lambda}^{-1}\mathbf{A}^{\mathsf{T}})$$
$$p(\mathbf{x}|\mathbf{y}) = \mathcal{N}(\mathbf{x}|\mathbf{\Gamma}[\mathbf{A}^{\mathsf{T}}\mathbf{L}(\mathbf{y} - \mathbf{b})] + \mathbf{\Gamma}\boldsymbol{\mu}, \mathbf{\Gamma})$$

• where $\Gamma = (\mathbf{\Lambda} + \mathbf{A}^\mathsf{T} \mathbf{L} \mathbf{A})^{-1}$

¹Check appendix of Bishop's book for many more useful properties!



• If we have a joint Gaussian distribution $N(\mathbf{x}|\boldsymbol{\mu},\boldsymbol{\Sigma})$, with $\boldsymbol{\Lambda}=\boldsymbol{\Sigma}^{-1}$, and we define the following partitions:

$$egin{aligned} \mathbf{x} &= egin{pmatrix} \mathbf{x}_a \ \mathbf{x}_b \end{pmatrix}, oldsymbol{\mu} &= egin{pmatrix} oldsymbol{\mu}_a \ oldsymbol{\mu}_b \end{pmatrix} \ oldsymbol{\Sigma} &= egin{pmatrix} oldsymbol{\Sigma}_{aa} oldsymbol{\Sigma}_{ab} \ oldsymbol{\Sigma}_{ba} oldsymbol{\Sigma}_{bb} \end{pmatrix}, oldsymbol{\Lambda} &= egin{pmatrix} oldsymbol{\Lambda}_{aa} oldsymbol{\Lambda}_{ab} \ oldsymbol{\Lambda}_{ba} oldsymbol{\Lambda}_{bb} \end{pmatrix} \end{aligned}$$

• Then the conditional distribution $p(\mathbf{x}_a|\mathbf{x}_b)$ is given by

$$p(\mathbf{x}_a|\mathbf{x}_b) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_{a|b},\boldsymbol{\Lambda}_{aa}^{-1})$$

$$oldsymbol{\mu}_{a|b} = oldsymbol{\mu}_a - oldsymbol{\Lambda}_{aa}^{-1} oldsymbol{\Lambda}_{ab} (oldsymbol{\mathsf{x}}_b - oldsymbol{\mu}_b)$$

• And the marginal distribution $p(\mathbf{x}_a)$ is given by

$$p(\mathbf{x}_a) = \mathcal{N}(\mathbf{x}_a | \boldsymbol{\mu}_a, \boldsymbol{\Sigma}_{aa})$$



• If we have two (univariate) Gaussian distributions (for the **same** variable x):

$$p(x) = \mathcal{N}(x|\mu_a, \sigma_a^2), \quad p(x) = \mathcal{N}(x|\mu_b, \sigma_b^2)$$

Their product is:

$$p(x) = \mathcal{N}(x|\mu_{ab}, \sigma_{ab}^2)$$

where

$$\mu_{ab} = \frac{\mu_a \sigma_b^2 + \mu_b \sigma_a^2}{\sigma_a^2 + \sigma_b^2}, \quad \sigma_{ab}^2 = \frac{\sigma_a^2 \sigma_b^2}{\sigma_a^2 + \sigma_b^2}$$

 This is very useful to directly combine models into a single prediction (e.g. ensemble models)!



- Example of Bayesian linear regression
- The joint probability of our model is given by:

$$p(\mathbf{y}, \boldsymbol{\beta} | \mathbf{X}, \lambda, \sigma) = p(\boldsymbol{\beta} | \mathbf{0}, \lambda \mathbf{I}) \prod_{n=1}^{N} p(y_n | \boldsymbol{\beta}^T \mathbf{x}_n, \sigma^2)$$

• Since both our prior $p(\beta|\lambda)$ and likelihood $p(y_n|\mathbf{x}_n, \beta, \sigma)$ are both Gaussian, we apply Bayes' theorem to compute posterior over β :

$$\underbrace{\frac{p(\boldsymbol{\beta}|\mathbf{y},\mathbf{X},\boldsymbol{\lambda},\boldsymbol{\sigma})}{\text{posterior}}} \propto \underbrace{\frac{p(\boldsymbol{\beta}|\mathbf{0},\boldsymbol{\lambda}\mathbf{I})}{\text{prior}}} \underbrace{\prod_{n=1}^{N} p(y_n|\boldsymbol{\beta}^T\mathbf{x}_n,\boldsymbol{\sigma}^2)}_{\text{likelihood}}$$
$$= \mathcal{N}(\boldsymbol{\beta}|\mathbf{0},\boldsymbol{\lambda}\mathbf{I}) \prod_{n=1}^{N} \mathcal{N}(y_n|\boldsymbol{\beta}^T\mathbf{x}_n,\boldsymbol{\sigma}^2)$$
$$= \mathcal{N}(\boldsymbol{\beta}|\boldsymbol{\mu}=?,\boldsymbol{\Sigma}=?)$$

• Can you do it? $\mu = ?$, $\Sigma = ?$ (Hint: look at slide 4...)



$$\underbrace{p(\boldsymbol{\beta}|\mathbf{y},\mathbf{X},\boldsymbol{\lambda},\sigma)}_{\text{posterior}} \propto \underbrace{p(\boldsymbol{\beta}|\mathbf{0},\boldsymbol{\lambda}\mathbf{I})}_{\text{prior}} \underbrace{\prod_{n=1}^{N} p(y_n|\boldsymbol{\beta}^T\mathbf{x}_n,\sigma^2)}_{\text{likelihood}}$$
$$= \mathcal{N}(\boldsymbol{\beta}|\mathbf{0},\boldsymbol{\lambda}\mathbf{I}) \underbrace{\prod_{n=1}^{N} \mathcal{N}(y_n|\boldsymbol{\beta}^T\mathbf{x}_n,\sigma^2)}_{n=1}$$

ullet Notice that, because our observations y_n are i.i.d., we can re-write:

$$= \mathcal{N}(\boldsymbol{\beta}|\mathbf{0}, \lambda \mathbf{I}) \, \mathcal{N}(\mathbf{y}|\mathbf{X}\boldsymbol{\beta}, \sigma^2 \mathbf{I})$$
$$= \mathcal{N}(\boldsymbol{\beta}|\boldsymbol{\mu}=?, \boldsymbol{\Sigma}=?)$$

- ullet where $\mathbf{y} = \{y_1, \dots, y_N\}$ and $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$
- Thus, we can apply the properties in slide 4, yielding:

$$\boldsymbol{\mu} = \boldsymbol{\Sigma} \left(\sigma^{-2} \mathbf{X}^T \mathbf{y} \right)$$
$$\boldsymbol{\Sigma} = (\lambda^{-1} \mathbf{I} + \sigma^{-2} \mathbf{X}^T \mathbf{X})^{-1}$$



• A chain graph



$$p(z_1, z_2, z_3, z_4) = p(z_4|z_3) p(z_3|z_2) p(z_2|z_1) p(z_1)$$

- Notice that: $z_4 \perp \{z_1, z_2\}|z_3$
- ullet We want to make inference on z_4

$$p(z_4) = \sum_{z_1, z_2, z_3} p(z_1, z_2, z_3, z_4)$$





A trivial solution is to go over all possible combinations of values!

$$p(z_4) = \sum_{z_1} \sum_{z_2} \sum_{z_3} p(z_1, z_2, z_3, z_4)$$

- \bullet Generally, if each of the m variables has k possible values, we'd have a complexity of $O(k^{m-1})$
- ullet This quickly becomes intractable (just remember our $\it trivial$ example with a mixture model) ullet NP-hard problem



• We can take advantage of the PGM structure

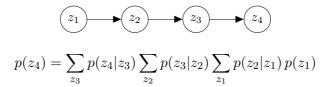
$$p(z_4) = \sum_{z_1, z_2, z_3} p(z_1, z_2, z_3, z_4)$$

$$= \sum_{z_1, z_2, z_3} p(z_4|z_3) p(z_3|z_2) p(z_2|z_1) p(z_1)$$

$$= \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) \sum_{z_1} p(z_2|z_1) p(z_1)$$

ullet In a chain graph, the complexity reduces to $O(mk^2)$









$$p(z_4) = \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) \sum_{z_1} p(z_2|z_1) p(z_1)$$
$$= \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) f_a(z_2)$$

- Notice that $f_a(z_2)$ is a function of z_2 (also called a *factor*). For example, a CPT with the probabilities of the k values of z_2
- We just "got rid" of z_1 by marginalizing over its values
- Same as in last lecture, when we marginalized over z for implementing LDA in STAN...





$$p(z_4) = \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) \sum_{z_1} p(z_2|z_1) p(z_1)$$

$$= \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) f_a(z_2)$$

$$= \sum_{z_3} p(z_4|z_3) f_b(z_3)$$



 (z_4)

$$p(z_4) = \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) \sum_{z_1} p(z_2|z_1) p(z_1)$$

$$= \sum_{z_3} p(z_4|z_3) \sum_{z_2} p(z_3|z_2) f_a(z_2)$$

$$= \sum_{z_3} p(z_4|z_3) f_b(z_3)$$

$$= f_c(z_4)$$

Variable Elimination (VE)



- Time complexity is exponential in size of largest factor
 - Each f_i is a factor
 - Size of factor is number of variables it depends on
- Order is vital (and sometimes a complicated problem)!
- Observed variables

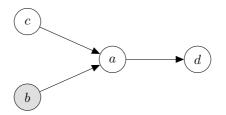
$$p(z|x=k) = \frac{p(z, x=k)}{p(x=k)}$$

- \bullet Perform VE on p(z,x=k) (i.e. we fix x=k) and then on $p(x=k)=\sum_z p(z,x=k)$
- Only acyclic graphs!

Playtime!



- Apply the Variable Elimination algorithm to the following graph
- ullet We want to infer p(d|b=1)



p(a|b,c)

	a = 0	a = 1
b = 0, c = 0	0.7	0.3
b = 0, c = 1	0.3	0.7
b = 1, c = 0	0.5	0.5
b = 1, c = 1	0.1	0.9

p(c)		
c = 0	c = 1	
0.7	0.3	
p(b)		
b = 0	b = 1	
0.4	0.6	

	p(d a)	
	d = 0	d = 1
a = 0	0.6	0.4
a = 1	0.2	0.8



Pearl's algorithm

ullet A graph is a polytree if (and only if) there is at most one simple path between any two nodes, v_i and v_k

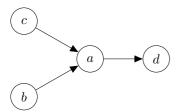


Figure: Polytree

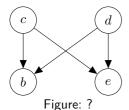
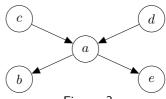


Figure: Not polytree



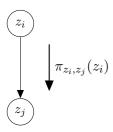


Pearl's algorithm

- We want to solve $p(\mathbf{z}|\mathbf{x})$, i.e. the conditional probability of all variables $\mathbf{z} = \{z_1, z_2, ..., z_V\}$, given evidence $\mathbf{x} = \{x_1, x_2, ..., x_E\}$
- Graph structure allows for incremental inference steps
 - Like in Variable Flimination...
- Some nodes are clearly specified from beginning
 - Evidence $x_i = k$
 - Priors
- The other nodes can build on them
- Information flows as messages between nodes



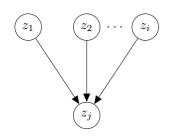
Pearl's algorithm



ullet π messages - from parent to child



Pearl's algorithm

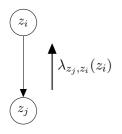


- \bullet π messages from parent to children
- Represent the current belief about the parent causal evidence
- After receiving all parents' messages, one can know more of the child

$$\pi(z_j) = \sum_{z_i \in \mathsf{parent}(z_j)} p(z_j|z_1, z_2, \ldots) \prod_{z_i \in \mathsf{parent}(z_j)} \pi_{z_i, z_j}(z_i)$$



Pearl's algorithm



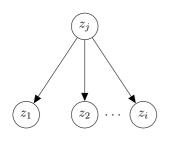
ullet λ messages - from child to parent

$$p(z_j|z_i)$$

 Represents the belief about the parent's value diagnostic evidence



Pearl's algorithm



ullet λ messages - from child to parent

$$p(z_i|z_j)$$

- Represents the belief about the parent's value diagnostic evidence
- After receiving all children's messages, one can know more of the parent

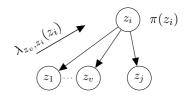
$$\lambda(z_j) = \prod_{z_i \in \mathsf{child}(z_j)} \lambda_{z_i, z_j}(z_j)$$



Pearl's algorithm

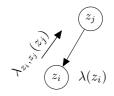
• But we still need the formulas for the messages!

$$\pi_{z_i,z_j}(z_i) = \pi(z_i) \prod_{v \neq j} \lambda_{z_v,z_i}(z_i)$$



ullet λ message when we have $p(z_i|z_j)$ (only one parent, z_j)

$$\lambda_{z_i,z_j}(z_j) = \sum_{z_i} \lambda(z_i) \, p(z_i|z_j)$$

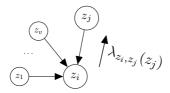




Pearl's algorithm

• Generic formula for λ messages:

$$\lambda_{z_i,z_j}(z_j) = \sum_{z_i} \lambda(z_i) \sum_{z_v \in \mathbf{z} \setminus z_j} p(z_i|z_1,z_2,...z_j) \prod_{v \neq j} \pi_{z_v,z_i}(z_v)$$



Pearl's algorithm - Initialization step



- For all $x_i \in \mathbf{x}$ (evidence nodes):
 - $\lambda(x_i) = 1$, wherever $x_i = e_i$, 0 otherwise (i.e. change the CPD tables, or dirac delta function)
 - $\pi(x_i) = 1$, wherever $x_i = e_i$, 0 otherwise (i.e. change the CPD tables, or dirac delta function)
- For all nodes, z_i , without parents:
 - $\pi(z_i) = p(z_i)$, the prior
- For all nodes, z_i , without children:
 - $\lambda(z_i) = 1$

Pearl's algorithm

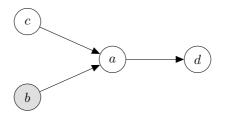


- Iterate until no change occurs
 - **1** For each node z_i , if it has received all π_{*,z_i} messages, calculate $\pi(z_i)$
 - **2** For each node z_i , if it has received all λ_{*,z_i} messages, calculate $\lambda(z_i)$
 - **3** For each node z_i , if $\pi(z_i)$ is calculated, and all λ_{*,z_i} messages have been received (except from z_j), calculate $\pi_{z_i,z_j}(z_i)$ and send the message to z_j
 - **4** For each node z_i , if $\lambda(z_i)$ is calculated, and all π_{*,z_i} messages have been received (except from z_j), calculate $\lambda_{z_i,z_j}(z_j)$ and send the message to z_j
- ullet Compute $\mathsf{Bel}(z_i) \propto \lambda(z_i) \, \pi(z_i)$ and normalize, for all desired nodes

Example



- Let's apply Belief Propagation to the following graph
- ullet We want to infer p(d|b=1)



p(a|b,c)

	a = 0	a = 1
b = 0, c = 0	0.7	0.3
b = 0, c = 1	0.3	0.7
b = 1, c = 0	0.5	0.5
b = 1, c = 1	0.1	0.9

p(c)		
c = 0	c = 1	
0.7	0.3	
p(b)		
b = 0	b = 1	
0.4	0.6	

-(-)

	p(u a)	
	d = 0	d = 1
a = 0 a = 1	0.6 0.2	0.4 0.8

n(d|a)

Example



• From the initialization, we have

$$\pi(c) = p(c)$$

$$c = 0 \quad c = 1$$

$$0.7 \quad 0.3$$

$$\frac{\pi(b)}{b=0 \quad b=1}$$

$$\frac{\lambda(b)}{b=0 \quad b=1}$$

$$0 \quad 1$$

$$\lambda(d)$$

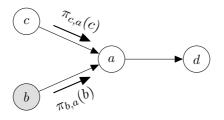
$$b = 0 \quad b = 1$$

$$1 \quad 1$$

Example



• We start with the evidence and prior



$$\pi_{c,a}(c) = \pi(c) \prod_{x \neq a} \lambda_{x,c}(c) = \pi(c)$$

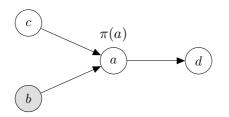
$$\pi_{b,a}(b) = \pi(b)$$

$$\frac{\pi_{c,a}(c)}{\frac{c = 0 \quad c = 1}{0.7 \quad 0.3}}$$

$$\frac{b = 0 \quad b = 1}{0 \quad 1}$$



• Now, $\pi(a)$



$$\pi(a) = \sum_{b,c} p(a|b,c) \, \pi_{b,a}(b) \, \pi_{c,a}(c)$$

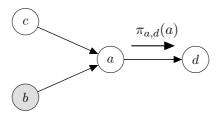
For a=1:

37

$$\begin{split} \pi(a=1) &= \sum_{b=\{0,1\}} \sum_{c=\{0,1\}} p(a=1|b,c) \, \pi_{b,a}(b) \, \pi_{c,a}(c) \\ &= \sum_{c=\{0,1\}} p(a=1|b=1,c) \, \pi_{c,a}(c) \\ &= p(a=1|b=1,c=0) \pi_{c,a}(c=0) + p(a=1|b=1,c=1) \, \pi_{c,a}(c=1) \\ &= 0.5 \times 0.7 + 0.9 \times 0.3 = 0.62 \text{, so } \pi(a=1) = 0.62 \text{ and } \pi(a=0) = 0.38 \end{split}$$



Finally, we reach d



Since there are no more parents, we have $\pi_{a,d}(a)=\pi(a)$

$$\pi_{a,d}(a)$$

a = 0	a = 1
0.38	0.62



$$\pi(d) = \sum_{a=\{0,1\}} p(d|a) \, \pi_{a,d}(a)$$

p(d|a) d = 0 d = 1 a = 0 0.6 0.4 a = 1 0.2 0.8

$$\begin{array}{c|c}
 & \pi_{a,d}(a) \\
\hline
 & a = 0 \quad a = 1 \\
\hline
 & 0.38 \quad 0.62
\end{array}$$

• Trivially becomes:

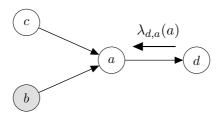
$$\pi(d)$$
 $b = 0$ $b = 1$
0.352 0.648

- $Bel(d) = \pi(d)\lambda(d)$
- In this case, the solution is trivially:

$$p(d|b=1) = \pi(d)$$

 But why stop here? We can propagate back, and also get p(a|b=1) and p(c|b=1)





$$\lambda_{d,a}(a) = \sum_{d=\{0,1\}} \lambda(d)p(d|a)$$

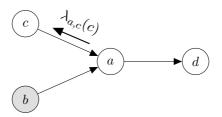
• Since $\lambda(d) = 1$, we have:

$$\lambda_{d,a}(a) = \sum_{d=\{0,1\}} p(d|a) = 1$$

• Notice that we have a unnormalized table!

$\lambda_{d,a}(a)$		
	a = 0	a = 1
	1	1





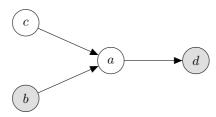
• Since $\lambda(a) = 1$, we have:

$$\lambda(c) = \lambda_{a,c}(c) = \sum_{a=\{0,1\}} \lambda(a) \, p(a|b=1,c) = \sum_{a=\{0,1\}} p(a|b=1,c) = 1$$

- Therefore:
 - Bel $(c) = \pi(c) \lambda(c) = \pi(c) = p(c)$
 - Bel $(a)=\pi(a)$ $\lambda(a)=\pi(a)=\sum_c p(a|b=1,c)\,\pi_{c,a}(c)$, as calculated before
- Makes sense, right (notice the independence of the graph!)?



• But, what if there's a new evidence, d=0?



- We only need to update d and propagate back!
- A new initialization gives:

$$\lambda(d)$$

$$d = 0 \qquad d = 1$$

$$1 \qquad 0$$

$$\begin{array}{c} \lambda(a) \\ \hline a = 0 \quad a = 1 \\ \hline 0.6 \quad 0.2 \end{array}$$

$$\begin{array}{c} \lambda_{d,a}(a) = \sum_{d=\{0,1\}} \lambda(d) p(d|a) = p(d=0|a) \\ \\ \lambda(a) = \lambda_{d,a}(a) \end{array}$$

$$\lambda_{a,c}(c) = \sum_{a,b=\{0,1\}} \lambda(a) p(a|b,c) \pi_{b,a}(b) =$$

$$= \sum_{a=\{0,1\}} \lambda(a) p(a|b=1,c)$$

$$\lambda(c) = \lambda_{a,c}(c)$$



- So,
 - Bel $(a) = \alpha \lambda(a) \pi(a)$, with α the normalizing factor

$$\frac{\pi(a)}{a=0 \quad a=1}$$
0.38 0.62

$\lambda(a)$		
a = 0	a = 1	
0.6	0.2	

•
$$\alpha = (0.38 * 0.6 + 0.62 * 0.2)^{-1} = 2.84$$

$$p(a|b=1,d=0) = \mathsf{Bel}(a)$$

$$\begin{array}{ccc}
 & a = 0 & a = 1 \\
\hline
 & 0.65 & 0.35
\end{array}$$

$$p(c|b=1,d=0) = \mathsf{Bel}(c)$$

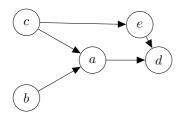
Belief Propagation

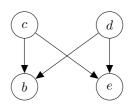


- A forward-backward pass of the BP algorithm gives us the exact inference for every variable
- Updating can be done incrementally
- Plates can be treated by expansion (given that it is still a polytree)
- We used discrete variables, but the reasoning is the same for continuous
 - Summations become integrals
 - Need to derive a new function for each step (easy in some cases, particularly with exponential family)
- In theory, it is still NP-Hard, but in practice, it is often linear time
- All of the above is valid for polytrees

The problem of non-polytrees







- Creates an (infinite) cycle
 - Clique Trees
 - Loopy belief propagation

Clique trees



- We aggregate nodes together (maximal cliques), and create a new graph, which is a polytree
- We may not get individual variable marginals, but instead joint distributions of subsets of variables
 - ...but you can marginalize individually in the end!
- One can create a clique tree just by following the Variable Elimination algorithm:
 - ullet Each factor f_i corresponds to a node in the tree
 - When factor f_i uses the result of factor f_j , then there is an undirected edge between the two.
 - It becomes a polytree

Clique trees



- Algorithm for creating a polytree
 - f 0 Given a set of factors, construct the undirected graph H_ϕ
 - **2** Triangulate H_{ϕ} , to construct a chordal graph² H^*
 - $\ensuremath{\mathfrak{3}}$ Find cliques in H^* , and make each one a node in a cluster graph
 - 4 Run the maximum spanning tree algorithm on the cluster graph to construct a tree



²In a chordal graph, all cycles of four or more vertices have a chord, which is an edge that is not part of the cycle but connects two vertices of the cycle.

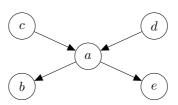
Loopy Belief Propagation



- Apply BP on the original graph in the following way:
 - 1 Initialize all messages to 1
 - 2 Run BP algorithm as before (start with evidence/priors)
 - 3 Each node sends messages in parallel (i.e. when it sends to one chlid/parent, it sends to all children/parents)
 - 4 Loop until convergence
- Works very well when there are some loops, but not a fully connected graph
- With one complete loop, the MAP should be correct
- \bullet If $p(\mathbf{z})$ is jointly Gaussian, Loopy Belief Propagation will converge to the correct marginals

Playtime!





 p(c)	
c = 0	c = 1
0.7	0.3

	p(b a)	
	b = 0	b = 1
a = 0 $a = 1$	0.3 0	0.7 1

p(a|c,d)

	a = 0	a = 1
c = 0, d = 0	0.5	0.5
c = 0, d = 1	0.9	0.1
c = 1, d = 0	0.1	0.9
c = 1, d = 1	0	1

d = 0	d = 1
0.8	0.2

p(e|a)

	e = 0	e = 1
a = 0 $a = 1$	0.2 0.7	0.8 0.3

- Consider the graph and CPTs above
- \bullet Perform forward-backward Belief Propagation, assuming the evidence that d=1, e=1

Some final notes



- There are actually many tools that do BP
 - Just check: https://www.cs.ubc.ca/~ murphyk/Software/bnsoft.html
- BP is often applied on the tractable sub-parts of your model, for example combining belief from different sub-models:
 - A Random Forest, RF provides $p_{\rm RF}(y)$, and a Logistic Regression model, LR, provides $p_{\rm LR}(y)$:

$$\mathsf{Bel}(y) = \alpha \, p_{\mathsf{RF}}(y) \, p_{\mathsf{LR}}(y)$$

- Also called Bayesian Model Averaging
- If you have gentle analytical forms for the sub-models (e.g. normal distribution for outputs), the combination is straightforward
- You can even treat sub-models as "priors" that provide to your random variables, and combine many in a complex model!

References



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