# STAT 88: Lecture 38

### Contents

Section 12.2: The Distribution of the Estimated Slope

## Warm up:

Let (X,Y) be a random pair and we observe  $(x_1,Y_1),\ldots,(x_n,Y_n)$  from the linear regression model

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i.$$

(a) We defined the mean squared error of a linear function of X as

$$MSE(a, b) = E((Y - (aX + b))^{2}).$$

How would you estimate MSE(a, b) from  $(x_1, Y_1), \ldots, (x_n, Y_n)$ ?

(b) How can you estimate the best regression line,  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ ?

(c) Compare the regression line in (b) with the population regression line  $\widehat{Y} = \widehat{a}X + \widehat{b}$ .

#### Last time

## The simple regression model

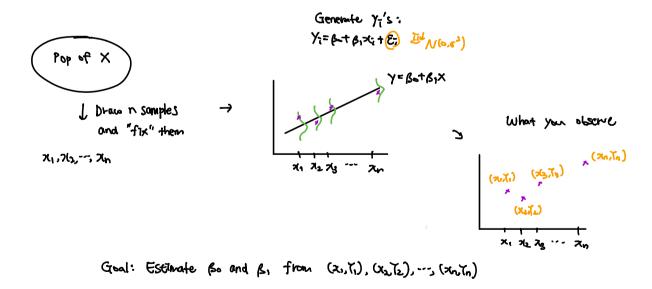
Y = response and x = predictor variable/covariate/feature

We assume for each of n observations

$$Y_i = \underbrace{\beta_0 + \beta_1 x_i}_{\text{signal}} + \underbrace{\epsilon_i}_{\text{noise}},$$

where

- $\beta_0$  and  $\beta_1$  are unobservable constant parameters.
- $x_i$  is the value of the predictor variable for individual i and is assumed to be constant (that is, not random).
- The errors  $\epsilon_1, \epsilon_2, \dots, \epsilon_n$  are i.i.d. normal  $\mathcal{N}(0, \sigma^2)$  random variables.
- The error variance  $\sigma^2$  is an unobservable constant parameter, and is assumed to be the same for all individuals i.



# 12.2. The Distribution of the Estimated Slope

# **Estimated Slope**

The least-squares estimate of the true slope  $\beta_1$  is the slope of the regression line, given by

$$\widehat{\beta}_1 = \frac{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})(Y_i - \bar{Y})}{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}.$$

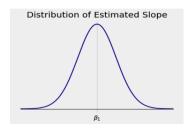
Is  $\widehat{\beta}_1$  random or constant? What distribution does  $\widehat{\beta}_1$  follow?

## **Expectation of the Estimated Slope**

Let's find  $E(\widehat{\beta}_1)$ .

First, what is  $E(Y_i)$  and  $E(\bar{Y})$ ?

Hence  $\widehat{\beta}_1$  is an unbiased estimator of  $\beta_1$ .



### Variance of the Estimated Slope

FACT:

$$\operatorname{Var}(\widehat{\beta}_1) = \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}.$$

Note that as  $n \to \infty$ ,  $\sum_{i=1}^{n} (x_i - \bar{x})^2$  gets very large and  $\text{Var}(\widehat{\beta}_1) \to 0$  so the difference between  $\widehat{\beta}_1$  and  $\beta_1$  becomes very small with high probability. Hence

$$\widehat{\beta}_1 \sim \mathcal{N}\left(\beta_1, \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}\right).$$

Example: (Exercise 12.4.1) Recall that the intercept of the regression line is given by the average of Y minus the slope times the average of x. That is,  $\widehat{\beta}_0 = \overline{Y} - \widehat{\beta}_1 \overline{x}$ . Is  $\widehat{\beta}_0$  an unbiased estimate of  $\beta_0$ ?

#### Standard Error of the Estimated Slope

We have

$$SD(\widehat{\beta}_1) = \frac{\sigma}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}.$$

 $\sigma$  is unknown so we have to estimate it. Recall  $\sigma$  is the SD of the error,  $SD(\epsilon_1) = \sigma$ . So we estimate  $\sigma$  with the SD of the residuals. If

$$D_i = Y_i - \widehat{Y}_i = Y_i - \widehat{\beta}_0 - \widehat{\beta}_1 x_i,$$

then

$$\hat{\sigma} = SD(D_1, \dots, D_n) = \frac{1}{n} \sum_{i=1}^{n} (D_i - \bar{D})^2.$$

This can be calculated in Python.

When the SD of an estimator is approximated by the data, it is called the SE (standard error):

$$SE(\widehat{\beta}_1) = \frac{\widehat{\sigma}}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}.$$

When n is large, it can be shown  $SE(\widehat{\beta}_1)$  converges to  $SD(\widehat{\beta}_1)$  so

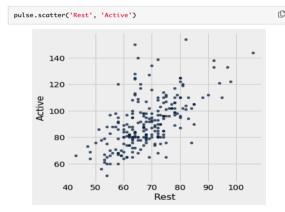
$$T = \frac{\widehat{\beta}_1 - \beta_1}{\operatorname{SE}(\widehat{\beta}_1)}$$

is approximately  $\mathcal{N}(0,1)$  for large n.

## **Pulse Rates**

We wish to predict active pulse rates from resting pulse rates.





/ Assumptions of sample tracor regression model?

```
active = pulse.column(0)
resting = pulse.column(1)

stats.linregress(x=resting, y=active)

Output:

(1.142879681904831,
13.182572776013345,
0.6041870881060092,
1.7861044071652305e-24,
0.09938884436389145)
```

n = 232 is large so

$$T = \frac{\widehat{\beta}_1 - \beta_1}{\operatorname{SE}(\widehat{\beta}_1)} \sim \mathcal{N}(0, 1).$$

A 95% CI for  $\beta_1$  is

$$(\widehat{\beta}_1 \pm 2 \cdot SE(\widehat{\beta}_1)) = (0.944, 1.342).$$

A fundamentally important question is whether the true slope  $\beta_1$  is 0. If it is 0, then the resting pulse rate isn't involved in the prediction of the active pulse rate, according to the regression model. Our testing problem is

$$H_0: \beta_1 = 0 \text{ vs } H_A: \beta_1 \neq 0.$$

T is our test statistic. Under  $H_0$ ,

$$T = \frac{\widehat{\beta}_1}{\operatorname{SE}(\widehat{\beta}_1)} \sim \mathcal{N}(0, 1).$$

The observed value of the test statistic is 11.5. So the p-value is

p-value = 
$$P(T \ge 11.5) + P(T \le -11.5) \approx 0$$
.

We reject  $H_0$  at 5% level.