

6. Applicazioni pratiche

Massimiliano Pastore
Università di Padova

Contenuti

- 1 t test
- 2 Studio clinico
- 3 Meta analisi
- 4 Power analysis
- 5 The dark side of the priors

Stime bayesiane in generale

*When you have eliminated the impossible
all that remain, no matter how improbable,
must be the truth.
(Sherlock Holmes)*

- L'inferenza bayesiana può essere considerata una riallocazione di credibilità in uno spazio di possibili candidati.
- In pratica vuol dire che riconsideriamo i *belief* dei valori dei parametri che sono consistenti con i dati ed eliminiamo i valori non consistenti.

t -test

Ripensare il t -test

- A differenza di NHST, il metodo bayesiano consente anche di accettare l'ipotesi nulla, non solo rigettarla.
- Il metodo bayesiano tiene conto di tutte le informazioni relative alle distribuzioni delle medie e delle deviazioni standard dei due gruppi.
- Inoltre consente di gestire outliers e distribuzioni asimmetriche piuttosto che normali.
- In particolare ci occuperemo di stima dei parametri, così da ottenere un risultato molto più ricco ed informativo rispetto al semplice t -test.

Esempio 1

- Uno psicologo vuole valutare se una nuova tecnica di meditazione sia efficace per la riduzione dello stress.
- A tal fine recluta un campione di 60 soggetti stressati, li divide in due gruppi (controllo e trattamento).
- Al termine dell'esperimento calcola i punteggi di stress dei soggetti nei due gruppi e li mette a confronto.

Esempio 1: t test

Le ipotesi:

- H_0 :

Esempio 1: t test

Le ipotesi:

- $H_0 : \mu_t = \mu_c$
- $H_1 :$

Esempio 1: t test

Le ipotesi:

- $H_0 : \mu_t = \mu_c$
- $H_1 : \mu_t < \mu_c$

I risultati:

Esempio 1: t test

Le ipotesi:

- $H_0 : \mu_t = \mu_c$
- $H_1 : \mu_t < \mu_c$

I risultati:

- Gruppo di controllo: $\bar{x}_c = 80.8$, $s_c = 10.05$
- Gruppo di trattamento: $\bar{x}_t = 75.6$, $s_t = 14.76$

Esempio 1: t test

Le ipotesi:

- $H_0 : \mu_t = \mu_c$
- $H_1 : \mu_t < \mu_c$

I risultati:

- Gruppo di controllo: $\bar{x}_c = 80.8$, $s_c = 10.05$
- Gruppo di trattamento: $\bar{x}_t = 75.6$, $s_t = 14.76$

$$t_{(58)} = \frac{(75.6 - 80.8)}{\sqrt{14.76^2/30 + 10.05^2/30}} = -1.59$$

Esempio 1: t test

Le ipotesi:

- $H_0 : \mu_t = \mu_c$
- $H_1 : \mu_t < \mu_c$

I risultati:

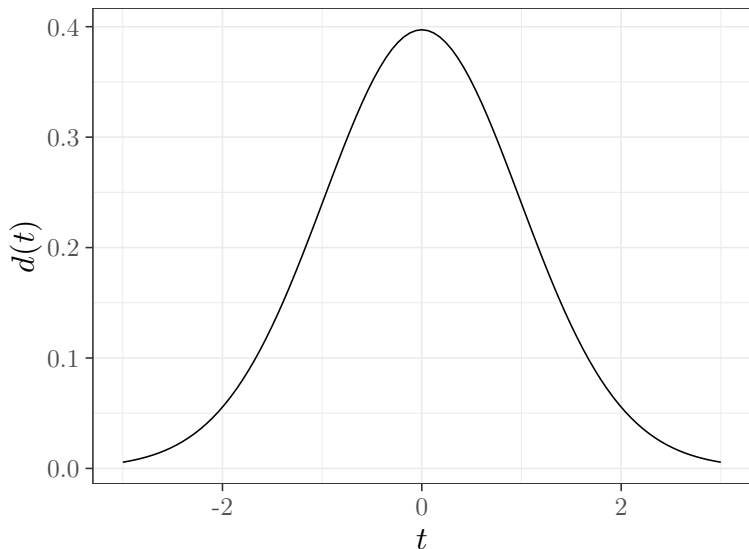
- Gruppo di controllo: $\bar{x}_c = 80.8$, $s_c = 10.05$
- Gruppo di trattamento: $\bar{x}_t = 75.6$, $s_t = 14.76$

$$t_{(58)} = \frac{(75.6 - 80.8)}{\sqrt{14.76^2/30 + 10.05^2/30}} = -1.59$$

$$p = 0.06$$

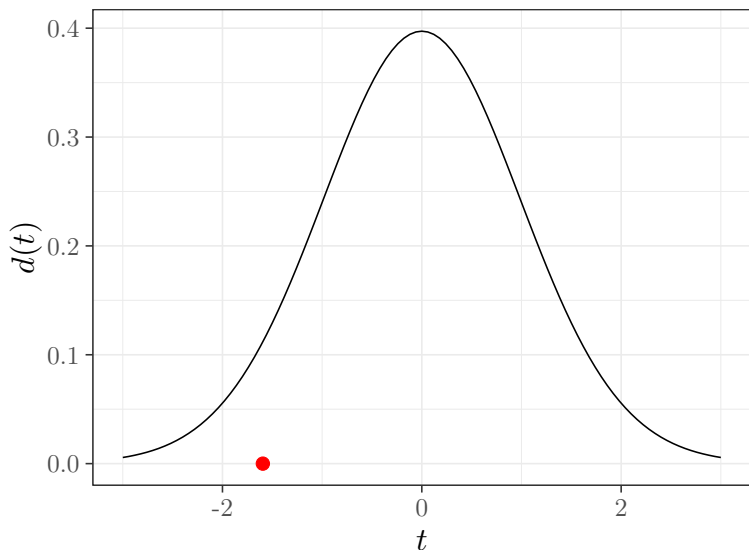
```
curve( dt( x, 58 ), -3, 3 )
```

```
curve( dt( x, 58 ), -3, 3 )
```



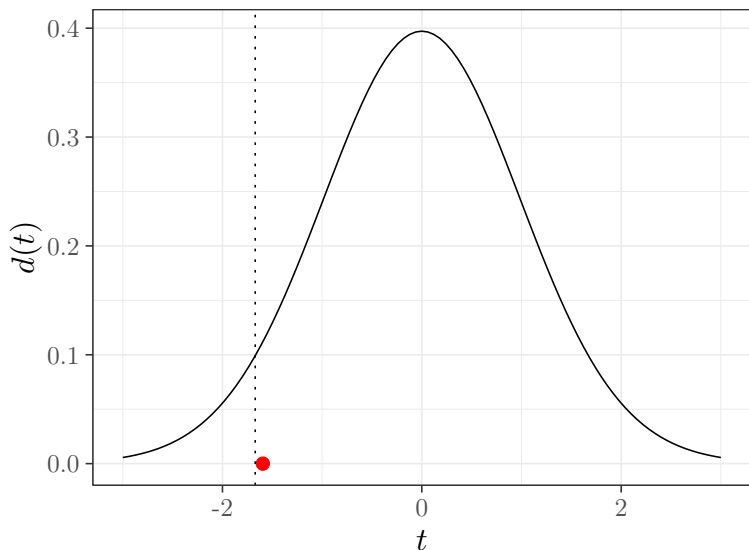
```
points( -1.59, 0, pch = 3, col = 'red')
```

```
points( -1.59, 0, pch = 3, col = 'red')
```



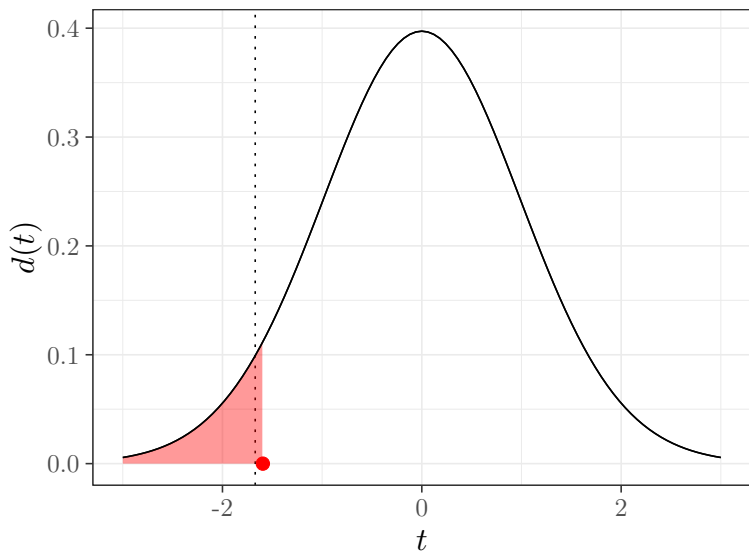

```
abline( v = qt( .05, 58 ), lty = 3 )
```

```
abline( v = qt( .05, 58 ), lty = 3 )
```



```
pt( -1.59, 58 )
```

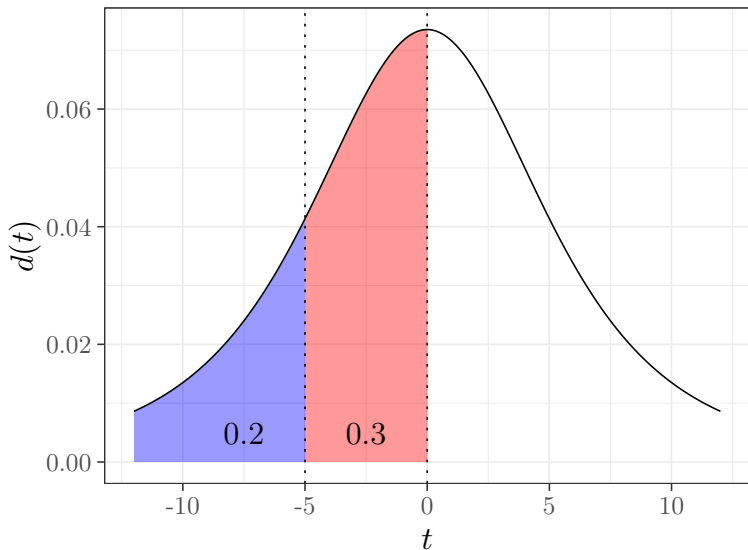
```
pt( -1.59, 58 )
```



Esempio 1: t -test bayesiano

- Supponiamo che il trattamento, per essere efficace, deve produrre una differenza di almeno 5 punti.
- In altre parole, una differenza compresa nell'intervallo $[-5, 0]$ è una differenza sostanzialmente nulla.
- Quindi, scegliamo come prior per la differenza una $t(3, 0, 5)$.
- Sulla base di questa prior, abbiamo che $p(-5 < \delta < 0) = 0.3$ e $p(\delta < -5) = 0.2$.

Esempio 1: prior



Esempio 1: t con Stan

```
data {  
  int<lower=0> N;  
  vector[N] y;  
  vector[N] x;  
}  
parameters {  
  real beta0;  
  real beta1;  
  real<lower=0> sigma;  
}
```

Esempio 1: t con Stan

```
data {
  int<lower=0> N;
  vector[N] y;
  vector[N] x;
}
parameters {
  real beta0;
  real beta1;
  real<lower=0> sigma;
}
transformed parameters {
  vector[N] mu;
  mu = beta0 + beta1 * x;
}
model {
  target += student_t_lpdf( beta0 | 3, 0, 10 );
  target += student_t_lpdf( beta1 | 3, 0, 5 );
  target += student_t_lpdf( sigma | 3, 0, 10 );
  target += normal_lpdf( y | mu, sigma);
}
```


Esempio 1: t con Stan

```
> library(rstan)
> fit1 <- stan( ... , data = dataList, seed = 3, iter = 3000 )
> print( fit, pars = c('beta0','beta1' ),
+       probs = c( .025, .5, .975 ) )
```

Esempio 1: t con Stan

```
> library(rstan)
> fit1 <- stan( ... , data = dataList, seed = 3, iter = 3000 )
> print( fit, pars = c('beta0','beta1' ),
+       probs = c( .025, .5, .975 ) )
```

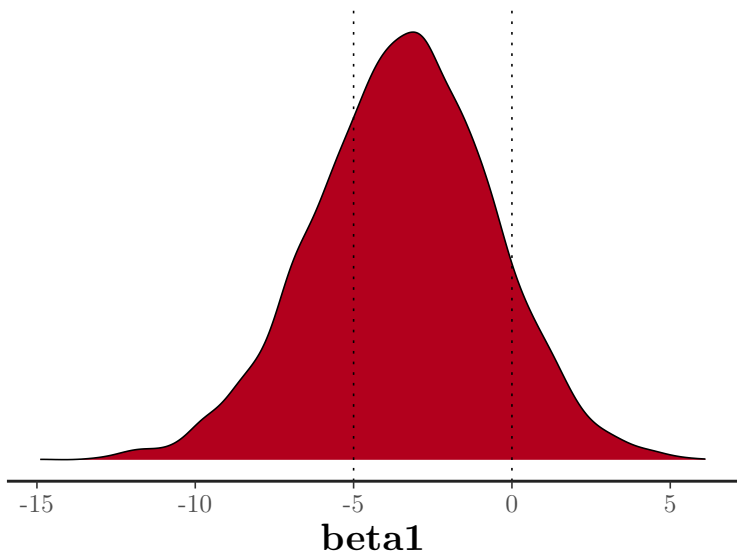
Inference for Stan model: anon_model.

4 chains, each with iter=3000; warmup=1500; thin=1;
post-warmup draws per chain=1500, total post-warmup draws=6000.

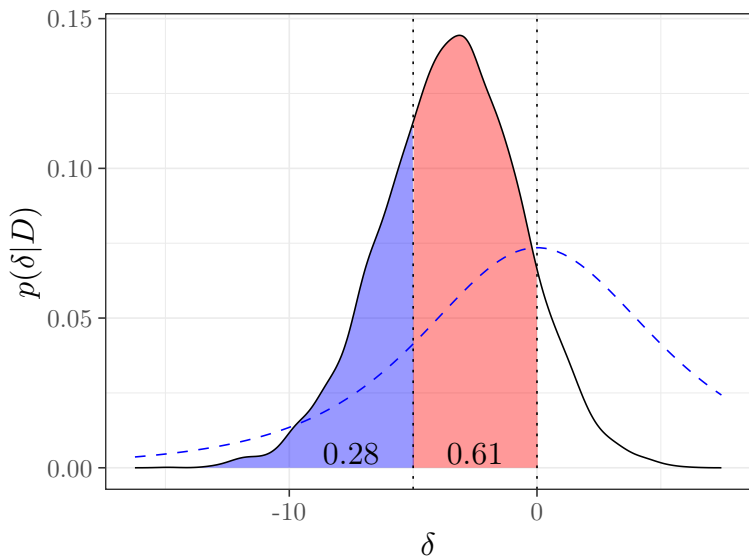
	mean	se_mean	sd	2.5%	50%	97.5%	n_eff	Rhat
beta0	79.79	0.04	2.14	75.57	79.80	84.0	3468	1
beta1	-3.45	0.05	2.81	-9.17	-3.39	1.9	3556	1

Samples were drawn using NUTS(diag_e) at Sat Mar 29 09:26:22 202
For each parameter, n_eff is a crude measure of effective sample
and Rhat is the potential scale reduction factor on split chains
convergence, Rhat=1).

```
plot( fit1, pars = "beta1", plotfun = "dens" ) +  
  geom_vline( xintercept = ROPE, lty = 3 )
```



```
plot( fit1, pars = "beta1", plotfun = "dens" ) +  
  geom_vline( xintercept = ROPE, lty = 3 )
```



Uno studio clinico

I dati

- Il set di dati include 318 soggetti \times 7 variabili:
 - Tre variabili endogene:
EHQ_PROBLEMI, EHQ_CONVINZIONI,
EHQ_EMOZIONI
 - Quattro variabili esogene:
GROUP, MPS_POS, MPS_NEG, OCI_TOT
- In particolare ci sono tre gruppi di soggetti:
clinical (144), diet (85), normal (89).

Le ipotesi

EHQ – item a 4 punti:

- Scala Problemi (12 items) Per i clinici attendiamo punteggi superiori ai 32 punti e 4/5 punti sopra i soggetti degli altri due gruppi.
- Scala Convinzioni (5 items) Per i clinici ed i diet attendiamo punteggi superiori ai 10 punti e circa 2 punti sopra i normali.
- Scala Emozioni (4 items). Per i clinici ed i diet attendiamo punteggi superiori a 8 punti e circa 2 punti sopra i normali.

I modelli

Utilizzando queste variabili definiamo un insieme di 6 modelli multivariati:

- M00, modello nullo:

$$\mathbf{Y} \sim 1$$

- M01, modello delle differenze tra i gruppi:

$$\mathbf{Y} \sim \text{GROUP}$$

- M02, effetti additivi GRUPPO + MPS:

$$\mathbf{Y} \sim \text{GROUP} + \text{MPS_striving} + \text{MPS_concern}$$

- M03, effetti additivi GRUPPO + MPS + OCI:

$$\mathbf{Y} \sim \text{GROUP} + \text{MPS_striving} + \text{MPS_concern} + \text{OCI_TOT}$$

- M04, modello con interazione MPS_NEG × OCI:

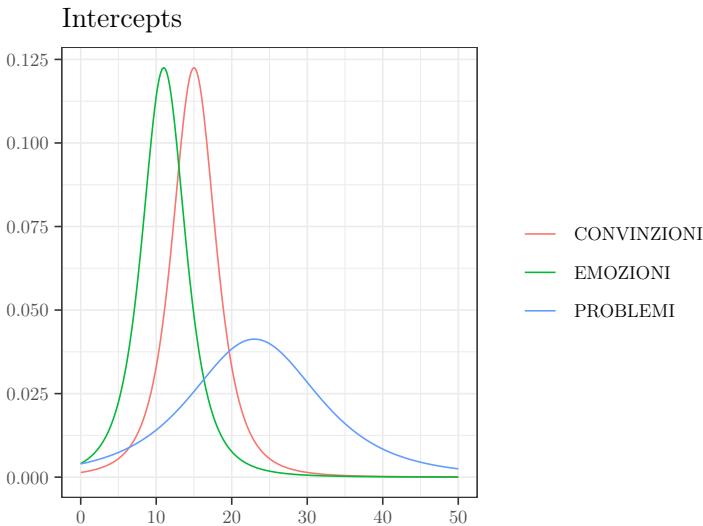
$$\mathbf{Y} \sim \text{GROUP} + \text{MPS_striving} + \text{MPS_concern} \times \text{OCI_TOT}$$

- M05, modello con interazioni rispetto al GRUPPO:

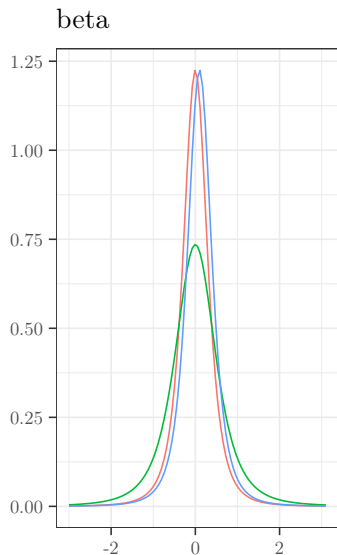
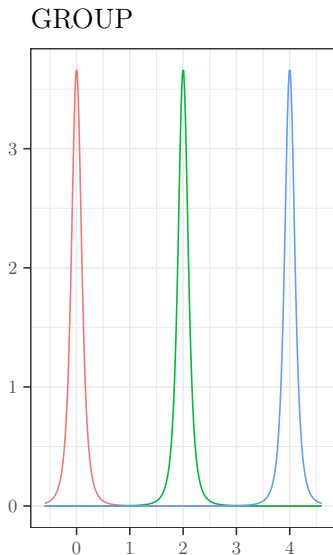
$$\mathbf{Y} \sim \text{GROUP} \times \text{MPS_striving} + \text{GROUP} \times \text{MPS_concern} + \text{GROUP} \times \text{OCI_TOT}$$

in cui \mathbf{Y} rappresenta la matrice delle variabili endogene EQ: PROBLEMI, CONVINZIONI, EMOZIONI. Ciascuno di questi modelli viene specificato con i residui correlati.

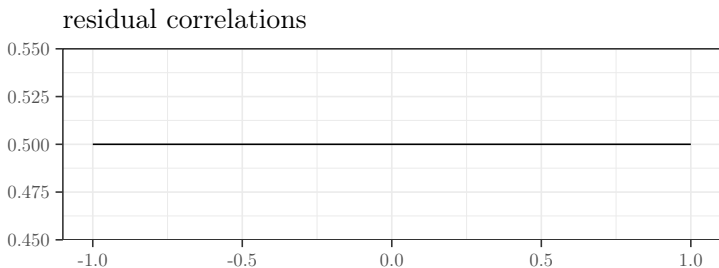
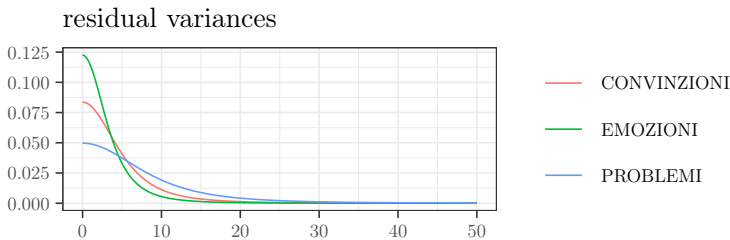
Prior specification



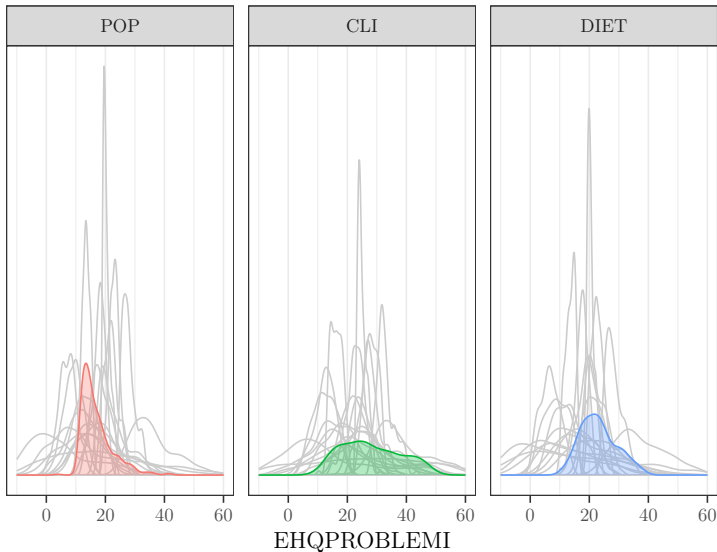
Prior specification



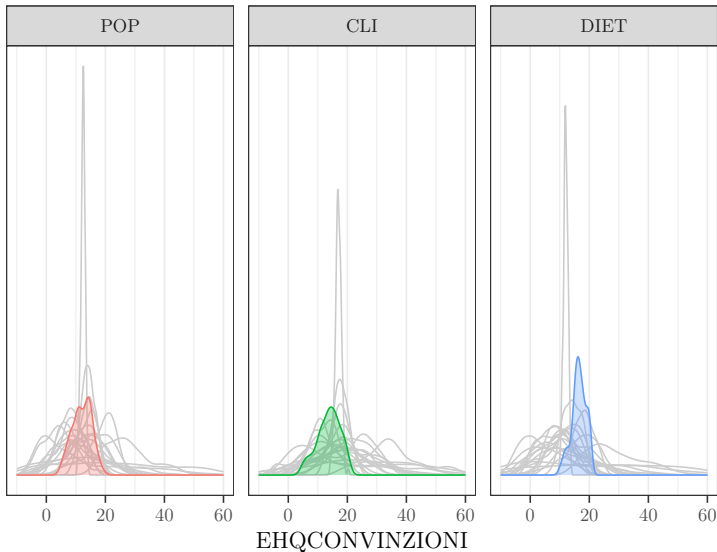
Prior specification



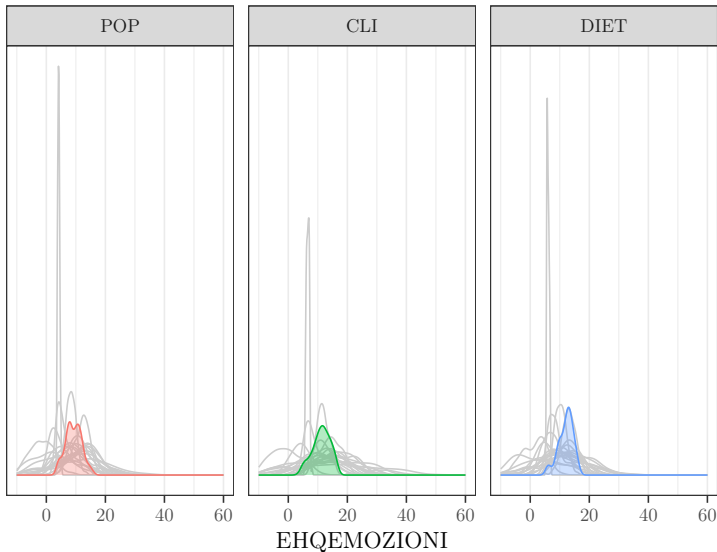
Prior Predictive Check



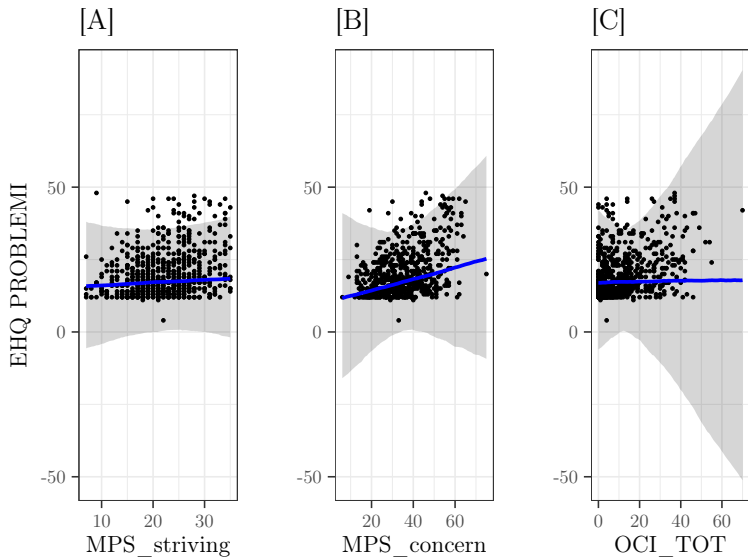
Prior Predictive Check



Prior Predictive Check



Prior Predictive Check



Modello finale (solo su PROBLEMI)

Le prior

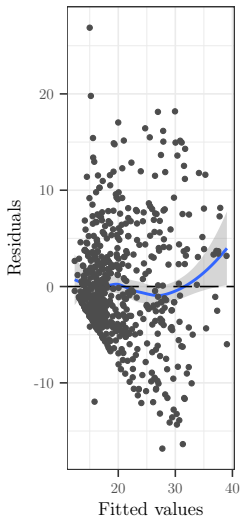
	prior	class	coef
26	student_t(3, 4, 0.1)	b	GROUPCLI
27	student_t(3, 0, 0.5)	b	GROUPCLI:MPS_concern
28	student_t(3, 0, 0.5)	b	GROUPCLI:MPS_striving
29	student_t(3, 0, 0.5)	b	GROUPCLI:OCI_TOT
30	student_t(3, 0, 0.1)	b	GROUPDIET
31	student_t(3, 0, 0.5)	b	GROUPDIET:MPS_concern
32	student_t(3, 0, 0.5)	b	GROUPDIET:MPS_striving
33	student_t(3, 0, 0.5)	b	GROUPDIET:OCI_TOT
34	student_t(3, 0.2, 0.3)	b	MPS_concern
35	student_t(3, 0.1, 0.3)	b	MPS_striving
36	student_t(3, 0, 0.5)	b	OCI_TOT
39	student_t(3, 18, 7.4)	Intercept	
43	student_t(3, 0, 7.4)	sigma	

Modello finale (solo su PROBLEMI)

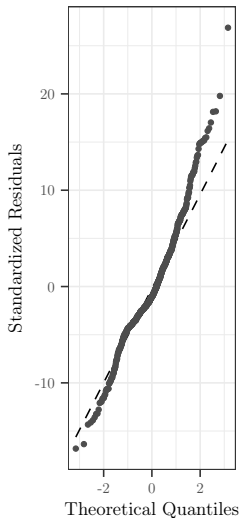
```
fit2 <- brm( EHQ_PROBLEMI ~ GROUP * MPS_concern +
             GROUP * MPS_striving + GROUP * OCI_TOT,
             data = Z, prior = myPrior, cores = 4 )
```

Diagnostics

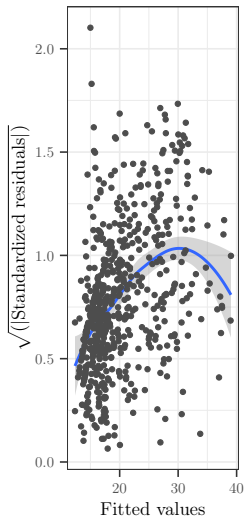
Residuals vs Fitted



Normal Q-Q

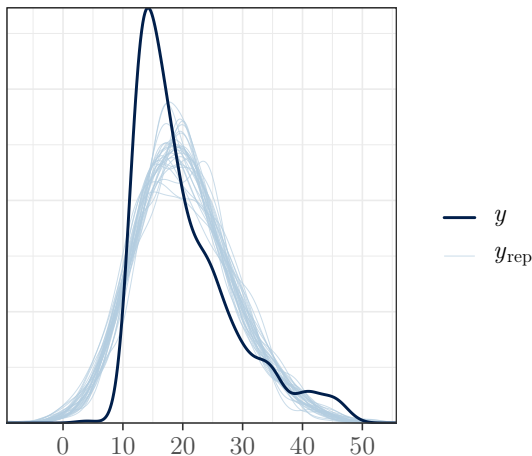


Scale-Location

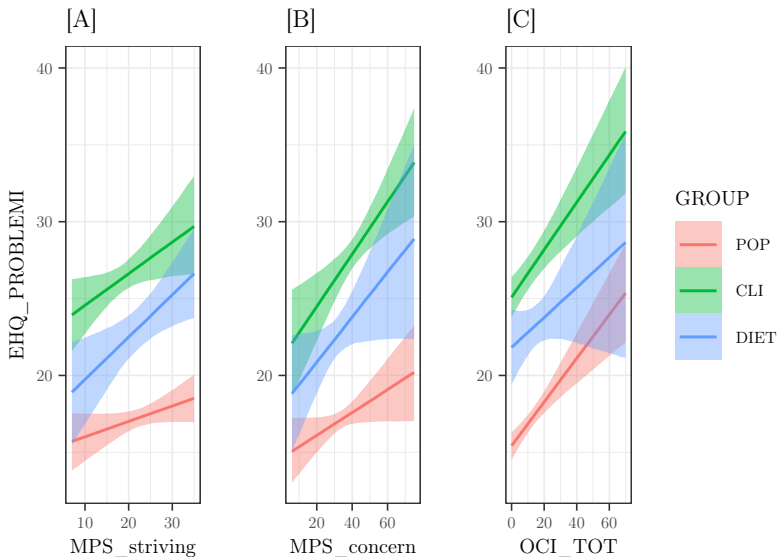


Posterior Predictive Check

```
pp_check( fit2, ndraws = 30 )
```



Model predictions



Analisi delle ipotesi

- Per i clinici
attendiamo punteggi
superiori ai 32 punti.

Analisi delle ipotesi

- Per i clinici
attendiamo punteggi
superiori ai soggetti
degli altri due
gruppi.

[B]

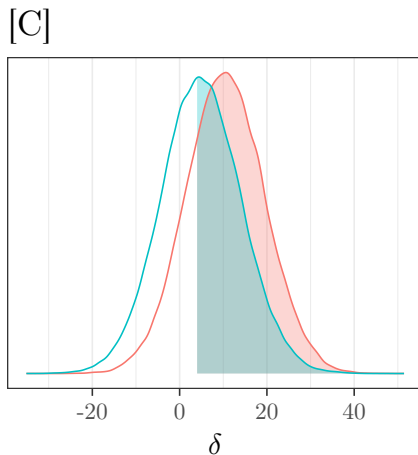
-

Analisi delle ipotesi

- Per i clinici
attendiamo punteggi
superiori di circa 4
punti rispetto a quelli
degli altri due gruppi.

Analisi delle ipotesi

- Per i clinici attendiamo punteggi superiori di circa 4 punti rispetto a quelli degli altri due gruppi.
- *Posterior predictive distributions* delle differenze tra i punteggi CLI-POP (in rosso) e tra i punteggi CLI-DIET (in azzurro);
 $Pr(\delta_{\text{CLI-POP}} > 4) = 0.77$, $Pr(\delta_{\text{CLI-DIET}} > 4) = 0.54$

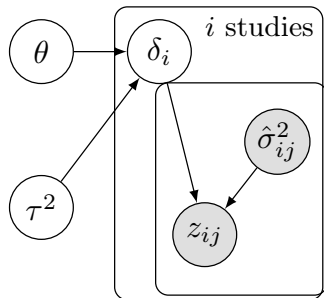


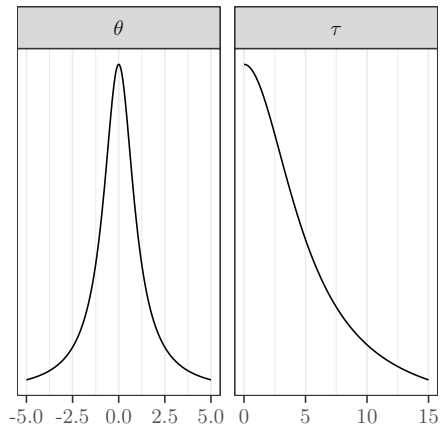
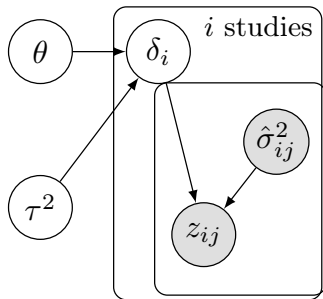
Meta analisi

Abstract

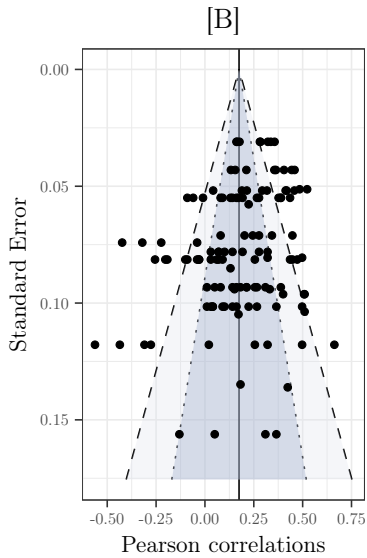
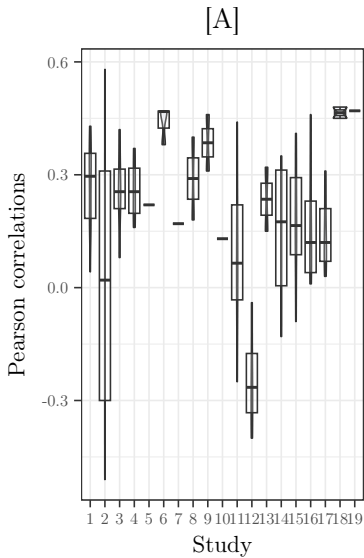
In two Bayesian meta-analyses, we investigated associations between Sensory Processing Sensitivity (SPS) and the Big Five personality traits (MA1) as well as both Positive and Negative Affect (MA2). Moderators were age and the three SPS subscales. In MA1 (8 papers, 6790 subjects), SPS in children correlated with Neuroticism ($r = 0.42$) but did not with Extraversion, Openness, Agreeableness or Conscientiousness. In adults, SPS correlated with Openness ($r = 0.14$) and Neuroticism ($r = 0.40$) but did not with Extraversion, Agreeableness or Conscientiousness. In MA2 (19 papers, 5326 subjects), SPS in children correlated with Negative ($r = 0.29$) and Positive Affect ($r = 0.21$), but only with Negative Affect ($r = 0.34$) in adults. Developmental and measurement aspects are discussed.

Lionetti, F., Pastore, M., Moscardino, U., Nocentini, A., Pluess, K., Pluess, M. (2019). Sensory Processing Sensitivity and its Association with Personality Traits and Affect: A Meta-Analysis. *Journal of Research in Personality*, 81, 138-152.



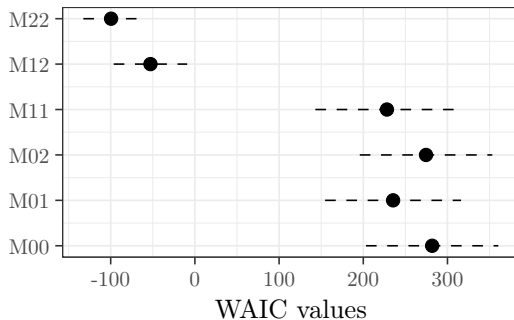


	authors	r	mean	sd	Zmean	Zsd	Nmean	Nsd
1	1 Nocentini	24	0.28	0.11	0.29	0.12	654.33	290.75
2	10 Pluess	1	0.13		0.13		141.00	
3	11 Sobocko	24	0.08	0.19	0.09	0.20	142.00	17.34
4	12 Booth	4	-0.24	0.15	-0.25	0.16	185.00	0.00
5	13 Meredith	2	0.23	0.12	0.24	0.13	116.00	0.00
6	14 Kiellgren	4	0.14	0.22	0.15	0.23	44.00	0.00
7	15 Pluess	16	0.18	0.16	0.18	0.16	334.00	0.00
8	16 Moscardino	12	0.16	0.14	0.16	0.15	100.00	0.00
9	17 Listou	8	0.15	0.10	0.15	0.10	167.00	0.00
10	18 Benham	2	0.46	0.02	0.50	0.03	383.00	0.00
11	19 Aron	1	0.47		0.51		96.00	
12	2 Evers	9	0.01	0.40	0.02	0.44	75.00	0.00
13	3 Liss	6	0.26	0.12	0.27	0.12	201.00	0.00
14	4 Ahadi	4	0.26	0.09	0.27	0.10	118.00	0.00
15	5 Gearhart	1	0.22		0.22		303.00	
16	6 Bakker	3	0.44	0.05	0.47	0.06	111.00	0.00
17	7 Wachs	1	0.17		0.17		94.00	
18	8 Jonsson	2	0.29	0.16	0.30	0.17	57.50	0.71
19	9 Brindle	2	0.39	0.11	0.41	0.12	157.00	0.00
20	TOTAL	126	0.17	0.21	0.18	0.22	259.17	243.51

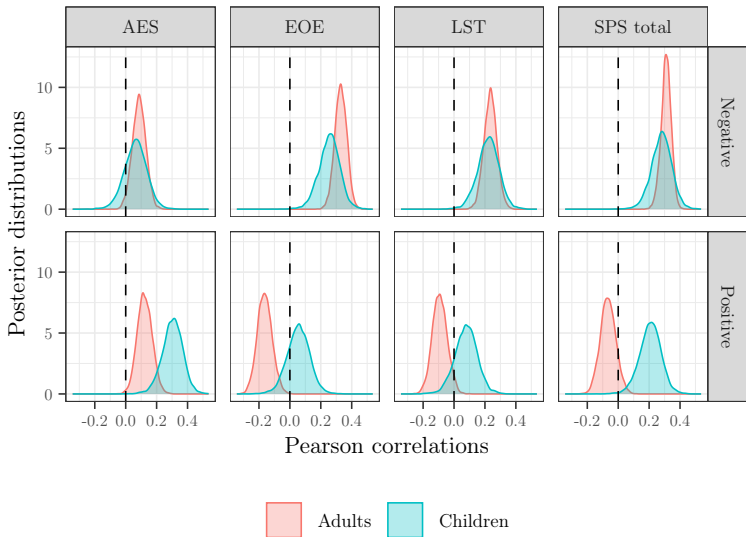


Model comparison

	n	WAIC	se	delta	weight
M00	126	282	79	-381	0.00
M01	126	235	81	-335	0.00
M02	126	275	79	-374	0.00
M11	126	228	85	-328	0.00
M12	126	-53	44	-47	0.00
M22	126	-100	33	0	0.99



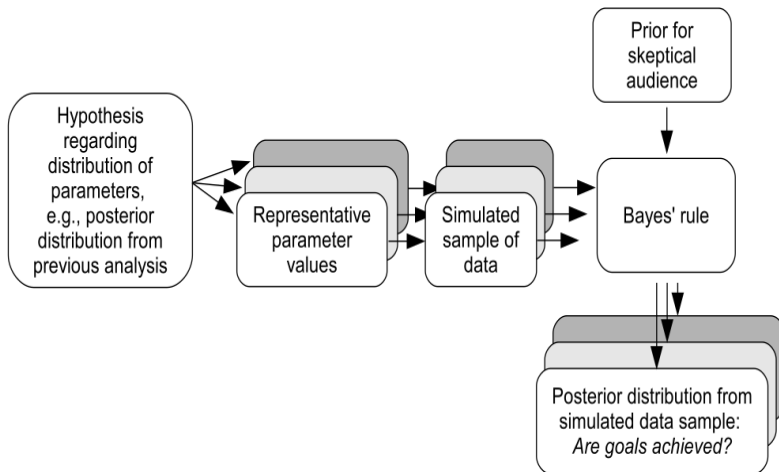
Best model



mtype	spstype	age	mean	hpd.lb	hpd.ub	range	<i>d</i>
neg	AES	children	0.06	-0.05	0.18	0.23	0.13
neg	EOE	children	0.25	0.14	0.36	0.21	0.53
neg	LST	children	0.22	0.11	0.33	0.22	0.45
neg	total	children	0.28	0.17	0.38	0.21	0.58
neg	AES	adults	0.09	0.02	0.15	0.14	0.17
neg	EOE	adults	0.33	0.27	0.39	0.12	0.70
neg	LST	adults	0.24	0.17	0.30	0.13	0.48
neg	total	adults	0.31	0.26	0.37	0.10	0.65
pos	AES	children	0.30	0.20	0.41	0.21	0.63
pos	EOE	children	0.05	-0.06	0.16	0.23	0.11
pos	LST	children	0.08	-0.03	0.20	0.23	0.16
pos	total	children	0.21	0.10	0.32	0.22	0.42
pos	AES	adults	0.12	0.05	0.20	0.15	0.25
pos	EOE	adults	-0.16	-0.24	-0.09	0.15	-0.33
pos	LST	adults	-0.09	-0.17	-0.01	0.15	-0.18
pos	total	adults	-0.07	-0.15	0.01	0.16	-0.13

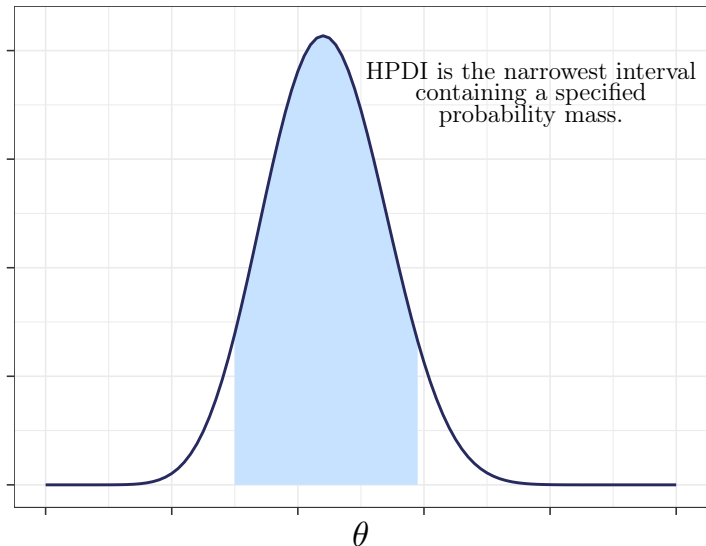
Power analysis

Esempio 4: (Bayesian) power analysis

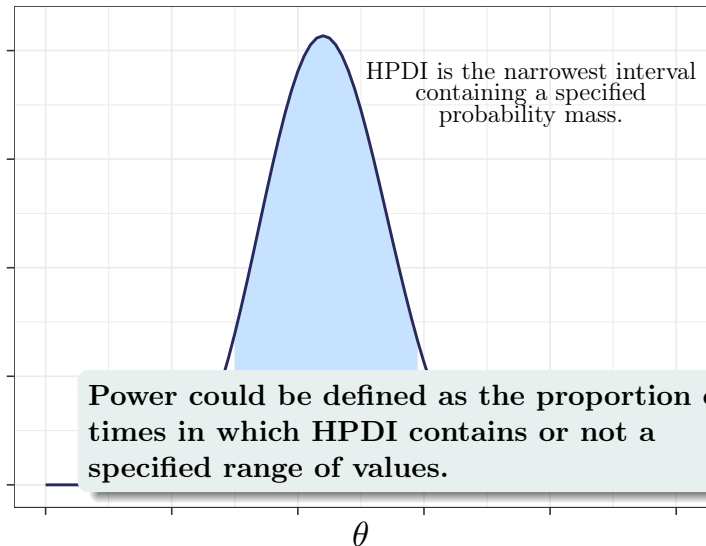


Kruschke, J. K. (2015). Doing Bayesian Data Analysis. Academic Press, UK.

Highest Posterior Density Interval



Highest Posterior Density Interval



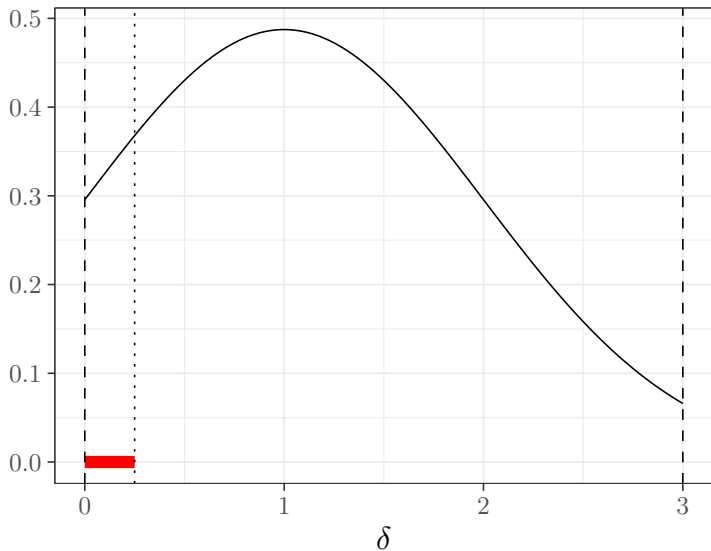
Example 4.1

- Let's suppose to want to test the mean scores standardized difference (δ) between two groups (Control and Experimental).
- A difference less than 0.25 is considered trivial in our particular application, i.e. null difference; consequently, the interval $[0, 0.25]$ represents the *Region of Practical Equivalence* (ROPE; Kruschke, 2015).
- How many subjects (per group) do we need for detecting a difference greater than 0.25?

Example 4.1: prior definition

- We know that, empirically, the difference could range between 0 to 3.
- We also believe that the most plausible value of the difference should be around 1.
- Consequently, we assume for the difference a truncated normal distribution in the interval $[0, 3]$, with mean equal to 1 and a variance value related to our degree of beliefs (for example $\sigma^2 = 1$).

Example 4.1

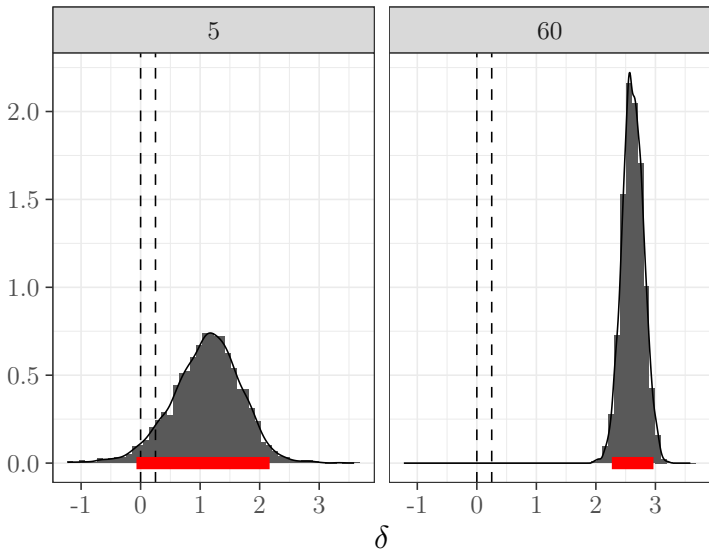


Example 4.1

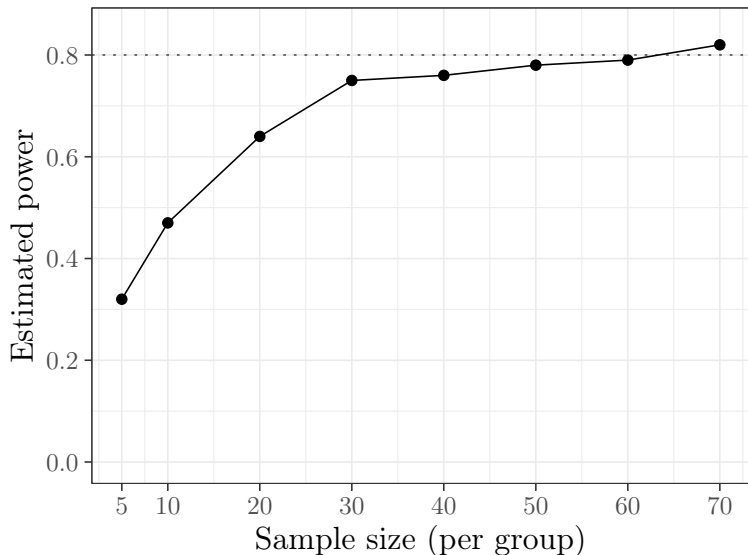
The following steps were repeated 1000 times for each specified value of sample size $n = (5, 10, 20, 30, 40, 50, 60, 70, 80, 90, 100)$.

- ➊ Sample a δ value from a $\text{TruncNorm}(0, 3, \mu = 1, \sigma = 1)$.
- ➋ Generate a couple of independent samples with sizes n from two populations in which the true difference $\mu_1 - \mu_2$ is δ .
- ➌ Estimate posterior distribution of $\hat{\delta} = \bar{x}_2 - \bar{x}_1$ by using a skeptical prior – e.g. $\delta \sim \text{Normal}(0, 1)$.
- ➍ Compute the 89% HPDI, and the number of cases in which the value 0.25 is outside the interval.

Example 4.1: results

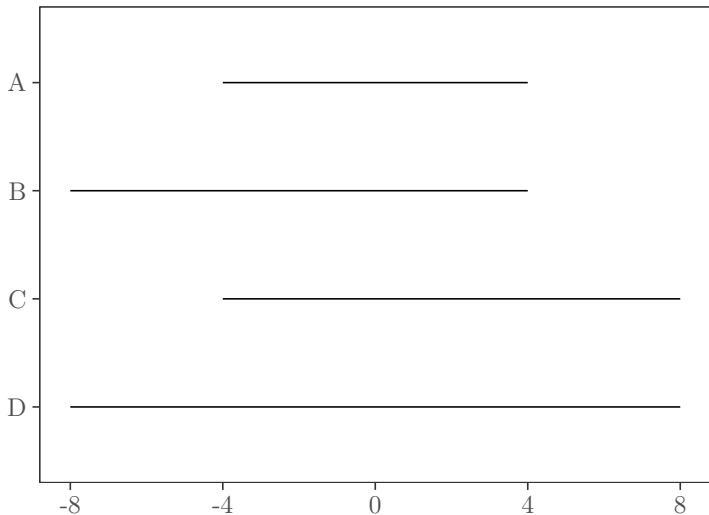


Example 4.1: results

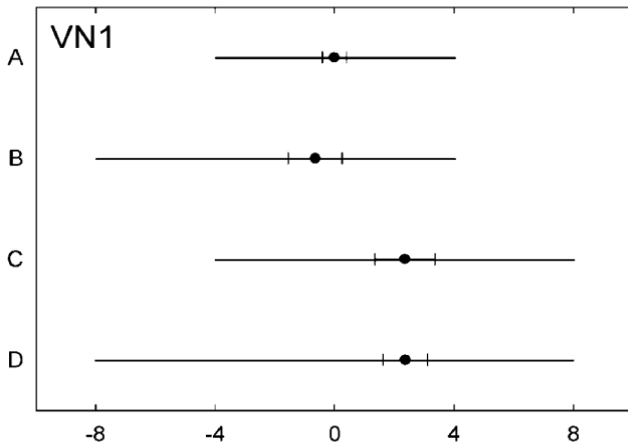


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Example 4.2: The task



Example of Individual data for a patient



McIntosh, R.D., Schindler, I., Birchall, D., Milner, A.D. (2005). Weights and measures: A new look at bisection behaviour in neglect. *Cognitive Brain Research*, 25, 833–850.

Example 4.2: The model

$$\left\{ \begin{array}{l} P \sim \mathcal{N}(\mu, \sigma) \end{array} \right.$$

general model

Example 4.2: The model

$$\left\{ \begin{array}{ll} P \sim \mathcal{N}(\mu, \sigma) & \text{general model} \\ \mu = k + dP_L \times L + dP_R \times R & \mu \text{ model} \end{array} \right.$$

Example 4.2: The model

$$\left\{ \begin{array}{ll} P \sim \mathcal{N}(\mu, \sigma) & \text{general model} \\ \mu = k + dP_L \times L + dP_R \times R & \mu \text{ model} \\ \sigma = \exp(\alpha + \beta \times (R - L)) & \sigma \text{ model} \end{array} \right.$$

Example 4.2: The model

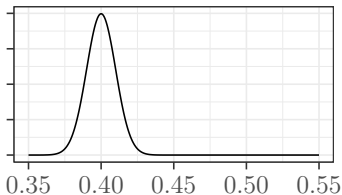
$$\left\{ \begin{array}{ll}
 P \sim \mathcal{N}(\mu, \sigma) & \text{general model} \\
 \mu = k + dP_L \times L + dP_R \times R & \mu \text{ model} \\
 \sigma = \exp(\alpha + \beta \times (R - L)) & \sigma \text{ model} \\
 k = k_0 + k_B[Block] & \text{elements of } \mu \\
 dP_L = dP_{L_0} + dP_{L_B}[Block] & \\
 dP_R = dP_{R_0} + dP_{R_B}[Block] & \\
 \alpha = \alpha_0 + \alpha_B[Block] & \text{elements of } \sigma \\
 \beta = \beta_0 + \beta_B[Block] &
 \end{array} \right. \quad (1)$$

Example 4.2: The question

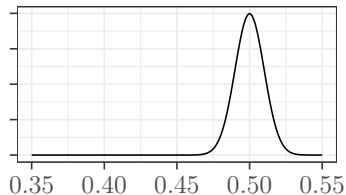
- We are interested in evaluating the experimental effects (for simplicity only for the mean), measured from the parameters dP_L and dP_R .
- In this case, we suppose that values lower than 0.2 can be considered equivalent to a NULL value.
- Unfortunately, we cannot have more than one patient and consequently our power analysis does not focus on “*how many subjects*” but instead on “*how many Blocks*” do we need.

Representative parameter values

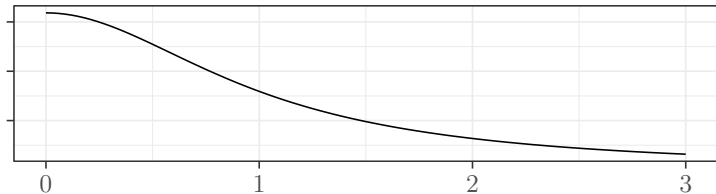
$$dP_{LB} = \mathcal{N}(0.4, 0.01)$$



$$dP_{RB} = \mathcal{N}(0.5, 0.01)$$



$$\sigma_{dP_L} = \sigma_{dP_R} = \text{HalfCauchy}(0, 1)$$

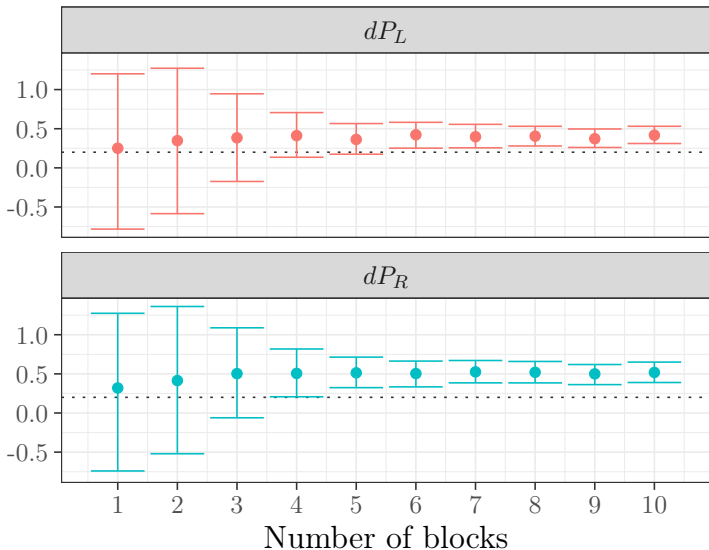


Example 4.2

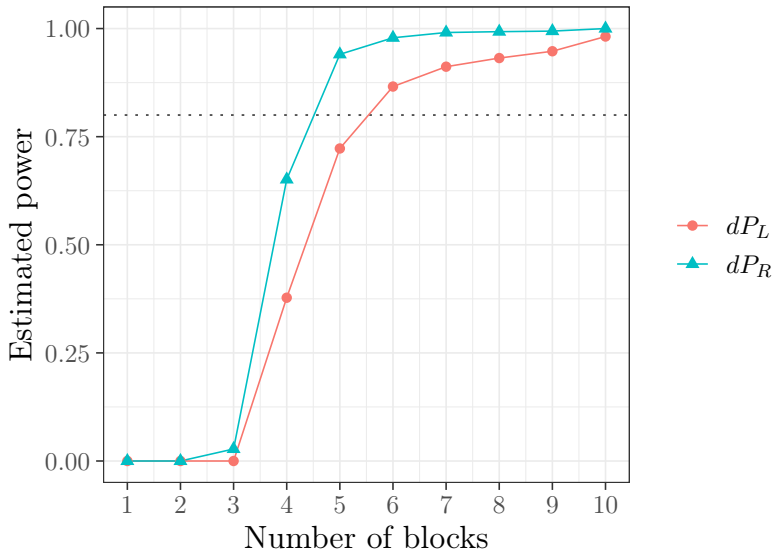
We replicated at least 1000 times for each specified number of Blocks (from Inf to -Inf) the following steps:

- 1 Generate 40 values of P for each Block.
- 2 Estimate posterior distributions of dP_L and dP_R .
- 3 Compute the 89% HPDI, and the number of cases in which the value 0.2 is outside the interval.

Posterior distributions (89% HPDI)



Estimated power



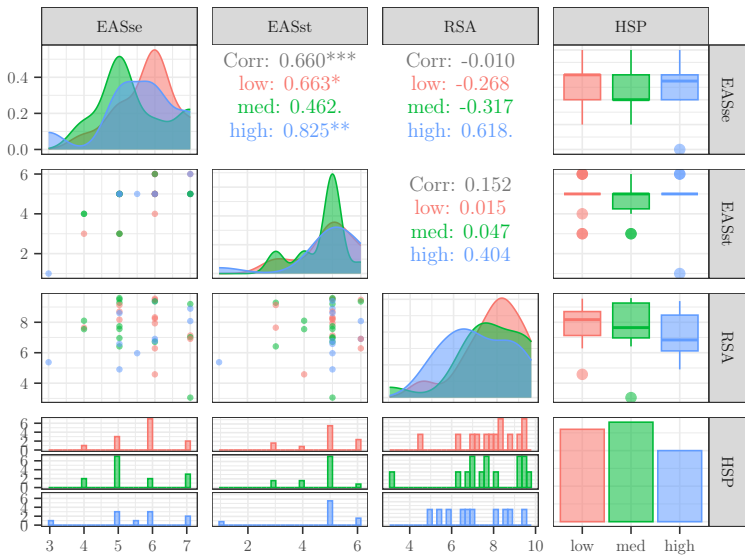
The dark side of the priors

Esempio 5: A Correlation Case

- In a study of the association between environmental sensitivity, parenting behavior and physiological arousal, 37 mothers with 3-month-old infants were recruited.
- The aim of the study was to examine whether the association between environmental sensitivity and parenting behavior is moderated by physiological arousal.
- Three variables were considered:
 - ❶ Environmental Sensitivity (ES, measured by the Highly Sensitive Person scale, Pluess et al., 2023)
 - ❷ Parenting Behavior (measured by the Emotional Availability Scales, EAS, Biringen, 2008)
 - ❸ Physiological Arousal (respiratory sinus arrhythmia, RSA)

Pluess, M., Lionetti, F., Aron, E. N., & Aron, A. (2023). People differ in their sensitivity to the environment: An integrated theory, measurement and empirical evidence. *Journal of Research in Personality*, 104, 104377.

Biringen, Z. (2008). *The Emotional Availability (EA) Scales Manual*, 4th Edn. Boulder, CO: International Center for Excellence in Emotional Availability.



Hypothetical correlation values

Pairs		SPS	r	μ
RSA	EASse	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30
RSA	EASst	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30

- Once we have defined the target intervals, we need to formalise the priors.

Hypothetical correlation values

Pairs		SPS	r	μ
RSA	EASse	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30
RSA	EASst	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30

- Since the correlations are between -1 and 1, the truncated Student's t is a good choice.

Hypothetical correlation values

Pairs		SPS	r	μ
RSA	EASse	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30
RSA	EASst	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30

- We set the degrees of freedom to 3 and the mean of the distribution to the mid-points of the target intervals (i.e. 0, 0.15, 0.3).

Hypothetical correlation values

Pairs		SPS	r	μ
RSA	EASse	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30
RSA	EASst	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30

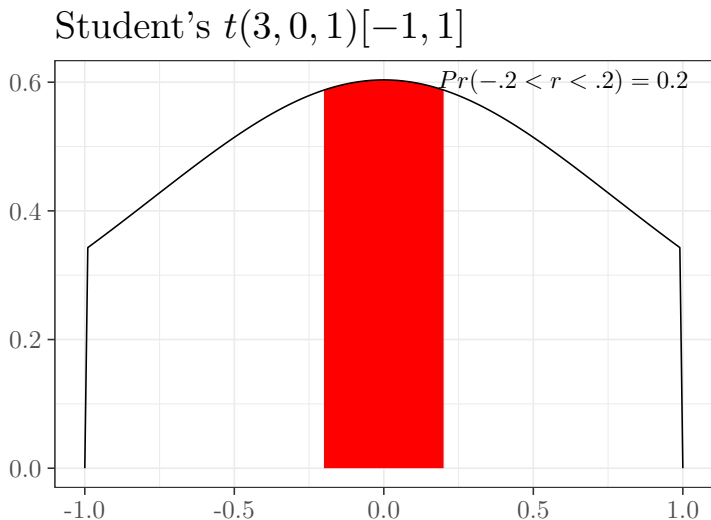
- We then need to determine the variance based on the knowledge/belief of the researcher.

Hypothetical correlation values

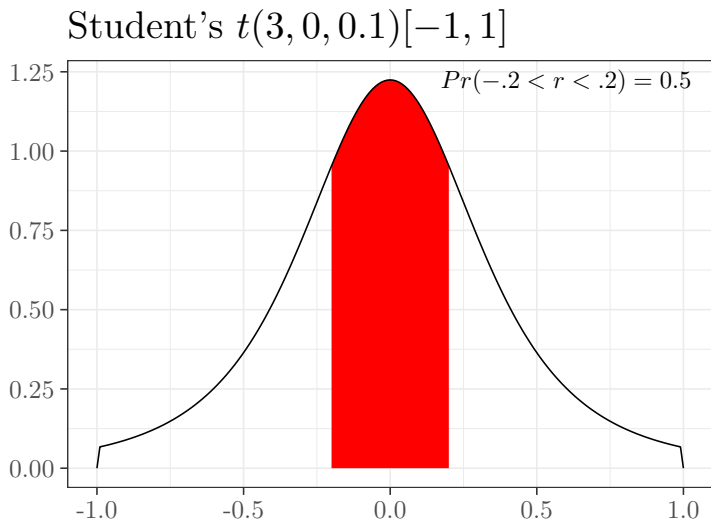
Pairs		SPS	r	μ
RSA	EASse	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30
RSA	EASst	Low	[-0.2; 0.2]	0.00
		Medium	[0; 0.3]	0.15
		High	[0.2; 0.4]	0.30

- A simple way of doing this is to ask the expert for the subjective probability that the correlation is in the target interval, and use this probability to calculate the appropriate variance of the distribution.

Truncated Student's t

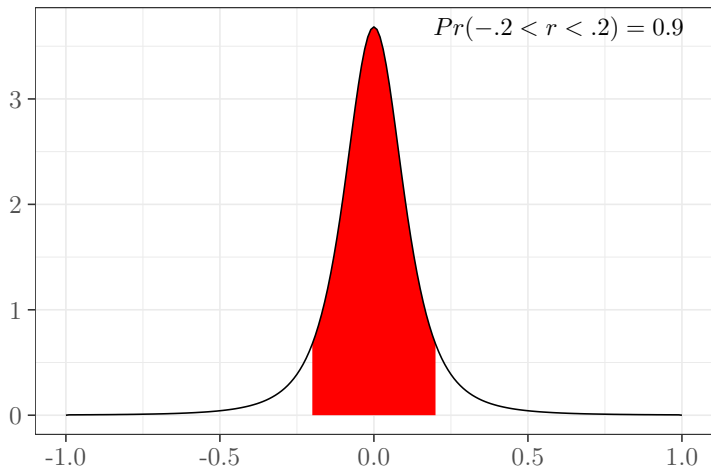


Truncated Student's t



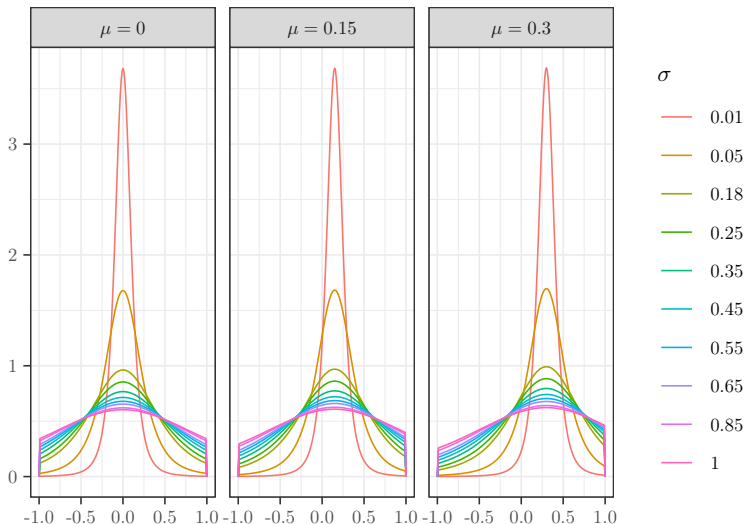
Truncated Student's t

Student's $t(3, 0, 0.01)[-1, 1]$



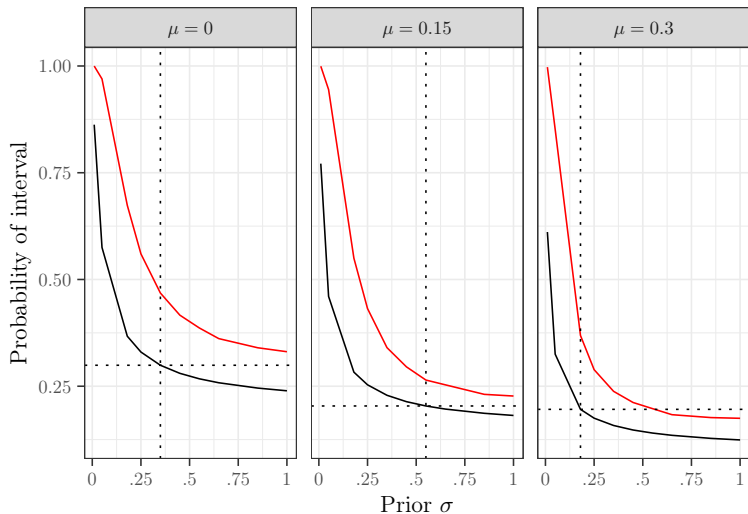
Sensitivity Analysis

- Given the small sample size, the choice of hypothesized intervals and related prior probabilities is crucial. In particular, we must consider that prior variance cannot be too small because – with small sample size – this means to obtain results depending only on prior distribution.
- Consequently we performed a sensitivity analysis for the three groups identified with the SPS scores. The objective of this analysis was to identify an optimal prior variance value that would enable us to evaluate our hypotheses without excessively influencing the results.

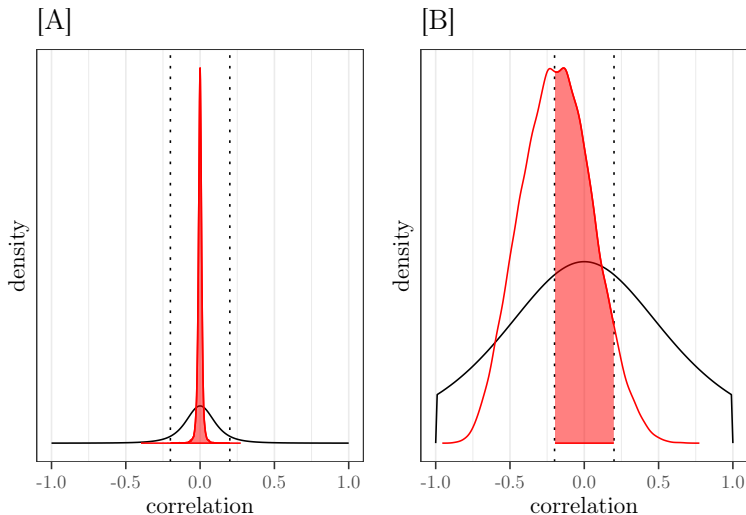


Truncated Student's t priors distributions in different scenarios. Panels refers to different means, colored lines to different standard deviations for fixed 3 degrees of freedom.

- For each scenario, we repeated the correlation estimation using these priors – the mean remains fixed across the groups, specifically at 0, 0.15, 0.3, while σ varies based on the 10 selected values – with four chains of 15000 effective replications each.
- Additionally, we calculated the prior and posterior probabilities for the target intervals, namely $[-0.2; 0.2]$, $[0; 0.3]$, and $[0.2; 0.4]$ respectively.



On the x -axis, the chosen values of σ are plotted. The black lines represent the prior probability of the target intervals, while the red lines represent the posterior probability. The vertical dashed lines indicate the selected values of σ for the analyses, while the horizontal dashed lines represent their corresponding prior probabilities.

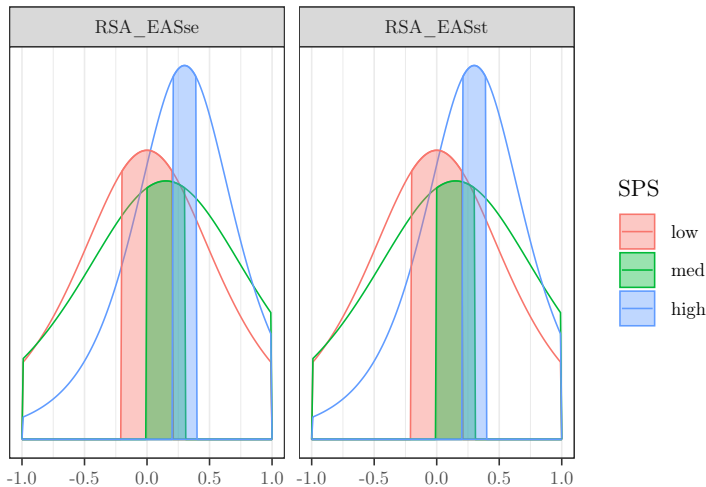


Examples of Scenarios for the case with $\mu = 0$; the black curves represent the priors, and the red curves represent the posteriors. [A] Priors is a truncated Student's $t(3, 0, 0.01)[-1, 1]$. $Pr(\rho \in [-.2, .2]) = 0.86$; $Pr(\rho \in [-.2, .2]|D) = 1$. [B] Priors is a truncated Student's $t(3, 0, 0.35)[-1, 1]$. $Pr(\rho \in [-.2, .2]) = 0.3$; $Pr(\rho \in [-.2, .2]|D) = 0.47$.

Correlation analysis

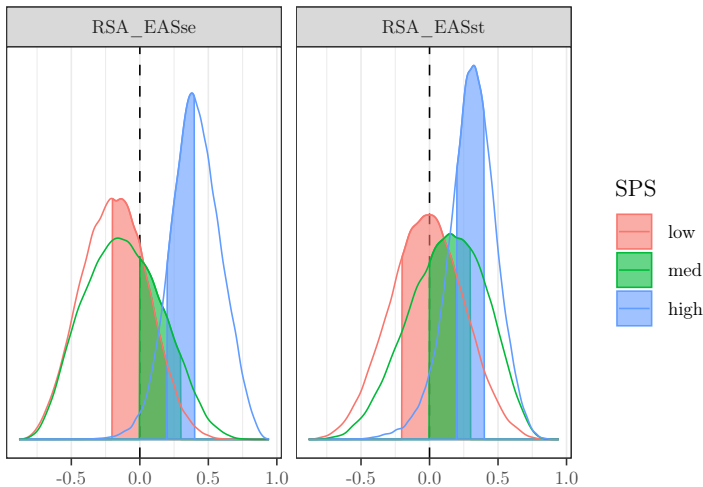
Pairs		SPS	r	μ	σ	Prob.
RSA	EASse	Low	[-0.2; 0.2]	0.0	0.3	0.30
		Medium	[0; 0.3]	0.2	0.6	0.20
		High	[0.2; 0.4]	0.3	0.2	0.20
RSA	EASst	Low	[-0.2; 0.2]	0.0	0.3	0.30
		Medium	[0; 0.3]	0.2	0.6	0.20
		High	[0.2; 0.4]	0.3	0.2	0.20

Priors



Priors distributions of correlations. Colors indicate the three SPS groups. Filled areas represent the prior probability of the target interval.

Posteriors



Posterior distributions of correlations. Colors refer to the SPS groups. Filled areas represent the posterior probability that correlation lies in the target interval.

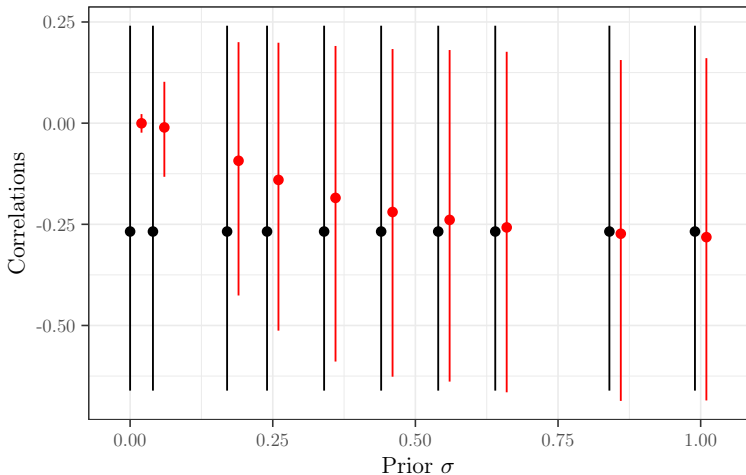
Pairs	HSP	<i>n</i>	Est <i>r</i>	89% HDI	Interval	Prior Pr.	Post Pr.
RSA EASse	low	13	-0.18	[-0.58; 0.2]	[-0.2; 0.2]	0.30	0.47
	med	14	-0.13	[-0.57; 0.33]	[0; 0.3]	0.20	0.27
	high	10	0.40	[0.12; 0.72]	[0.2; 0.4]	0.20	0.37
RSA EASst	low	13	-0.02	[-0.46; 0.42]	[-0.2; 0.2]	0.30	0.54
	med	14	0.14	[-0.32; 0.58]	[0; 0.3]	0.20	0.39
	high	10	0.31	[0.02; 0.61]	[0.2; 0.4]	0.20	0.45

Posterior table. Pairs = paired variables; HSP = SPS group; *n* = sample size; Est *r* = estimated correlation, i.e. the median of posterior distribution; 89% HDI = Highest Density Interval; Interval = target interval; Prior Pr. = prior probability of target interval; Post Pr. = posterior probability of target interval.

A frequentist remark

- To evaluate the advantages of the Bayesian approach over the simple frequentist method, we compare the frequentist confidence intervals with the Bayesian credible intervals.

A frequentist remark



Comparison of 89% confidence intervals (black lines) and 89% credibility intervals (red lines) as a function of the chosen values of σ in the prior.

- Obviously, the frequentist estimates (Pearson r) do not change – as the prior is not used – and the intervals are always quite wide, averaging 0.9, which is almost half of the available range $[-1, 1]$.
- On the other hand, the HDIs increase as σ increases, and when this parameter is too low – indicating a very strong prior – the HDIs become too narrow to be credible.
- In other words, if we were to adopt a frequentist approach (without using a prior), we would necessarily rely solely on the observed correlations and therefore have very wide confidence intervals. As a result, the credibility of the results would be certainly lower.
- Conversely, by adopting priors, we can reason in terms of probabilities (prior and posterior) and assess hypotheses regarding the differences between groups. This approach provides a level of credibility for the results that is certainly not less than that of the frequentist approach.

Summarizing

- The Bayesian approach allows incorporating prior information (priors) and updating it with the newly observed data to obtain the posterior distribution.
- With small sample sizes, the choice of the prior distribution is crucial and can significantly influence the results. Sensitivity analysis on the priors allows evaluating the impact of the prior choice on the results.
- In cases of small samples, the Bayesian approach can be more informative compared to frequentist confidence intervals alone. However, the use of overly informative priors can lead to results excessively influenced by the prior assumptions rather than the observed data.

Used R packages

- **bayesplot**. Gabry J, Mahr T (2025).
- **brms**. Bürkner P (2017).
- **cmdstanr**. Gabry J, Češnovar R, Johnson A, Bronder S (2025).
- **cowplot**. Wilke C (2024).
- **gamlss.dist**. Stasinopoulos M, Rigby R (2023).
- **GGally**. Schloerke B, Cook D, Larmarange J, Briatte F, Marbach M, Thoen E, Elberg A, Crowley J (2024).
- **ggdist**. Kay M (2024).
- **ggplot2**. Wickham H (2016).
- **HDInterval**. Meredith M, Kruschke J (2022).
- **knitr**. Xie Y (2025).
- **mnormt**. Azzalini A, Genz A (2022).
- **posterior**. Bürkner P, Gabry J, Kay M, Vehtari A (2025).
- **R**. R Core Team (2025).
- **Rcpp**. Eddelbuettel D, Francois R, Allaire J, Ushey K, Kou Q, Russell N, Ucar I, Bates D, Chambers J (2025).
- **report**. Makowski D, Lüdecke D, Patil I, Thériault R, Ben-Shachar M, Wiernik B (2023).
- **rethinking**. McElreath R (2024).
- **rstan**. Stan Development Team (2025).
- **rstanarm**. Goodrich B, Gabry J, Ali I, Brilleman S (2024).
- **StanHeaders**. Stan Development Team (2020).
- **truncnorm**. Mersmann O, Trautmann H, Steuer D, Bornkamp B (2023).
- **xtable**. Dahl D, Scott D, Roosen C, Magnusson A, Swinton J (2019).



massimiliano.pastore@unipd.it
<https://psicostat.dpss.psy.unipd.it/>



L^AT_EX

