

Stat 134 Lec 22

Warmup 10:00-10:10

Suppose customers are arriving at a ticket booth at rate of five per minute, according to a Poisson arrival process. Find the probability that:

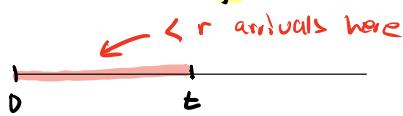
- (a) Starting from time 0, the 9th customer doesn't arrive within 5 minutes;

$$P(T_9 > 5) = P(N_5 \leq 9)$$

$$\sum_{k=0}^8 \frac{e^{-25} 25^k}{k!}$$

$$N_t \sim \text{Pois}(\lambda t)$$

$$P(T_r > t) = P(N_t \leq r) \quad \text{where } N_t \sim \text{Pois}(\lambda t)$$



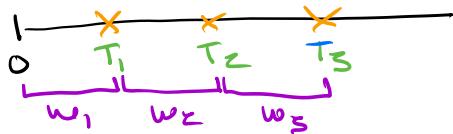
Announcements

Nice job on challenging problem!



Monday lecture on Moment Generating Functions (not in book)

Last time sec 4.2 Gamma Distribution



$$T_i \sim \text{Exp}(\lambda), \lambda > 0$$

$$f(t) = \begin{cases} \lambda e^{-\lambda t} & t \geq 0 \\ 0 & \text{else} \end{cases}$$

← Variable part

$$T_r \sim \text{Gamma}(r, \lambda), r > 0$$

$$f(t) = \begin{cases} \frac{1}{\Gamma(r)} \lambda^r t^{r-1} e^{-\lambda t} & t \geq 0 \\ 0 & \text{else} \end{cases}$$

where $\Gamma(r) = \int_0^\infty t^{r-1} e^{-t} dt$
 $r \in \{1, 2, 3, \dots\}$
 then $\Gamma(r) = (r-1)!$

$$T_r = w_1 + w_2 + \dots + w_r, w_i \stackrel{iid}{\sim} \text{Exp}(\lambda)$$

$$E(w_1) = \frac{1}{\lambda} \Rightarrow E(T_r) = \frac{r}{\lambda}$$

$$\text{Var}(w_1) = \frac{1}{\lambda^2} \Rightarrow \text{Var}(T_r) = \frac{r}{\lambda^2}$$

Ex

A random variable X has non negative values and density $Cx^4 e^{-3x}$ for $0 \leq x < \infty$, and some constant C .

What distribution is X ? $X \sim \text{Gamma}(5, 3)$

$$\text{Find } \text{Var}(X) = \boxed{\frac{5}{9}}$$

TODAY sec 4.4 (skip 4.3)

- ① Gamma example
- ② Change of Variable formula for densities.

① Gamma example

Suppose customers are arriving at a ticket booth at rate of five per minute, according to a Poisson arrival process. Find the probability that:

see warm up

- (a) Starting from time 0, the 9th customer doesn't arrive within 5 minutes;
- (b) At least one customer arrives within 40 seconds after the arrival of the 13th customer.

b) $1 - P(\text{no arrival in } \frac{2}{3} \text{ min})$

$$\frac{e^{-\lambda} \lambda^0}{0!}$$

$$\lambda = \lambda t = 5 \frac{2}{3}$$

$$= 1 - e^{-\frac{10}{3}}$$

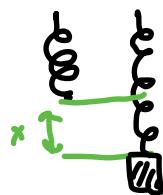
$$T \sim \text{Exp}(5)$$

Also

$$P(T_1 < \frac{2}{3}) = 1 - e^{-5(\frac{2}{3})}$$

$$\xrightarrow[T_{13}]{} X$$

Hooke's Law $F = kX$

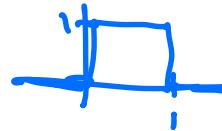


displacement of a Spring

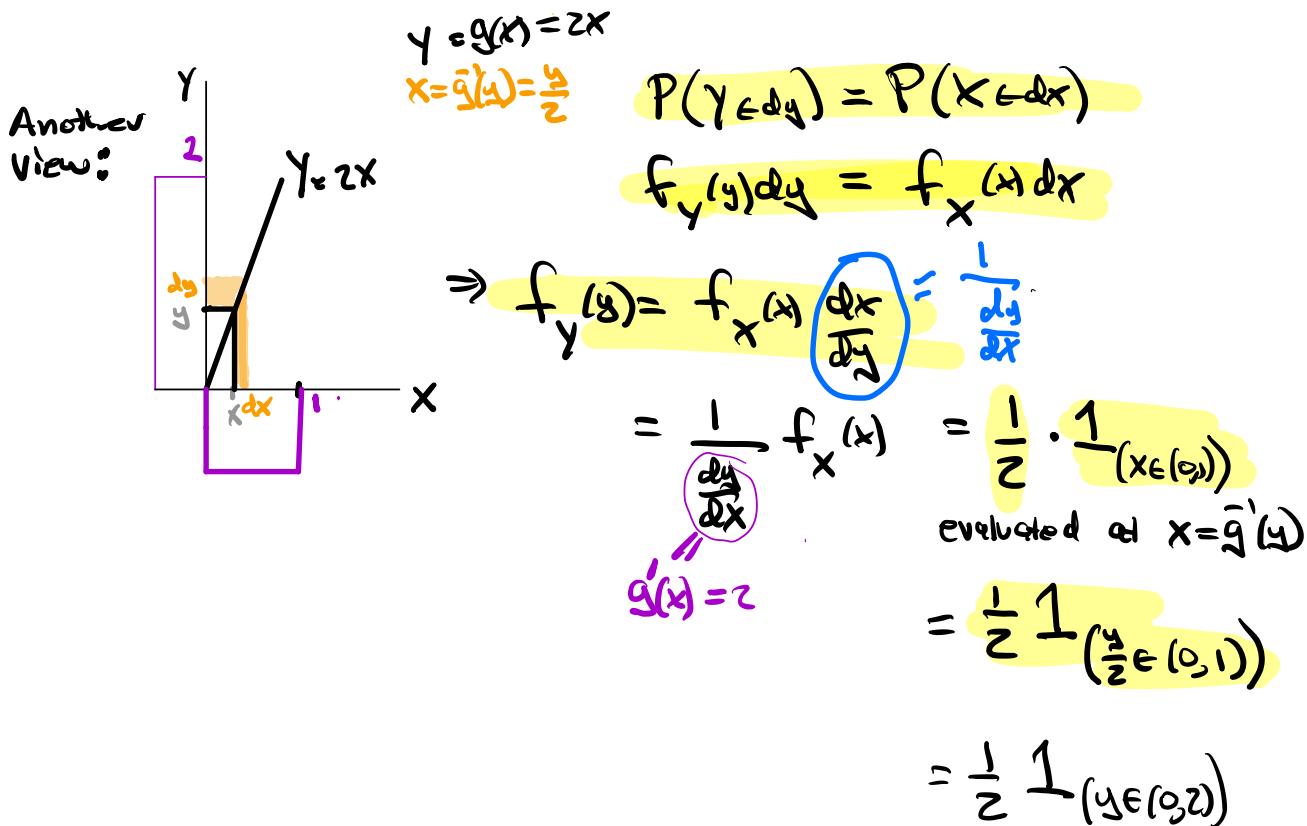
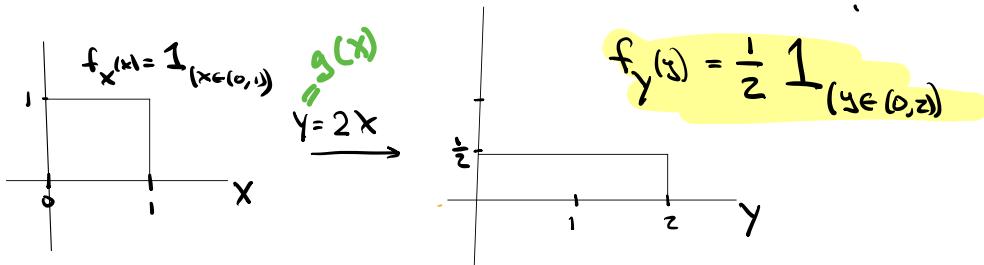
② Sec 4.1 Change of Variable formula for densities

ex let $X = \text{displacement of a spring}$

The density of $X \sim U(0,1) \Rightarrow f_X(x) = 1_{(x \in (0,1))}$

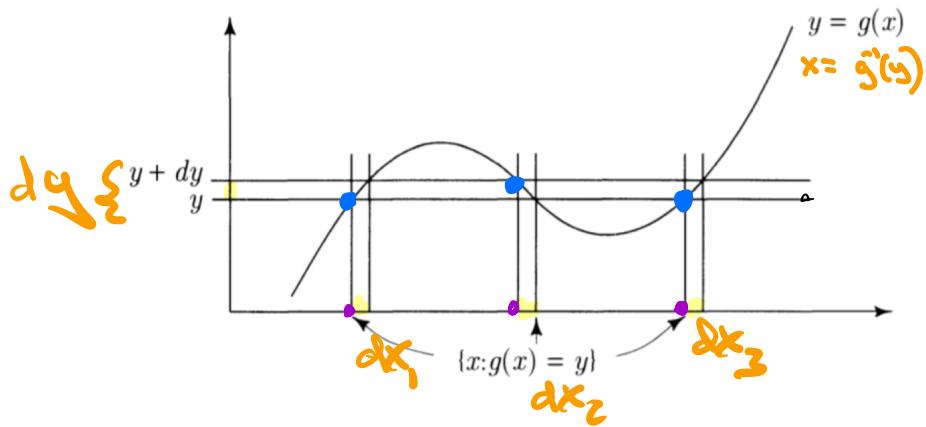


What is density of $Y=2X$, the force on the spring?



Here the transformation of X to $Y=g(X)$ is linear and one to one.

What if $Y=g(X)$ isn't linear and one-one?



$y \in dy$ iff $x \in dx_1$, or $x \in dx_2$ or $x \in dx_3$

$$P(y \in dy) = P(x \in dx_1) + P(x \in dx_2) + P(x \in dx_3)$$

$$f_y(y) dy = f_x(x_1) dx_1 + f_x(x_2) dx_2 + f_x(x_3) dx_3$$

$$f_y(y) = f_x(x_1) \frac{dx_1}{dy} + f_x(x_2) \frac{dx_2}{dy} + f_x(x_3) \frac{dx_3}{dy}$$

$$= \frac{f_x(x_1)}{|g'(x_1)|} + \frac{f_x(x_2)}{|g'(x_2)|} + \frac{f_x(x_3)}{|g'(x_3)|}$$

evaluated
at $x = g^{-1}(y)$

$$\frac{dx}{dy} = \frac{1}{\frac{dy}{dx}} = \frac{1}{g'(x)}$$

$$\nwarrow P(x \in dx_2) \geq 0$$

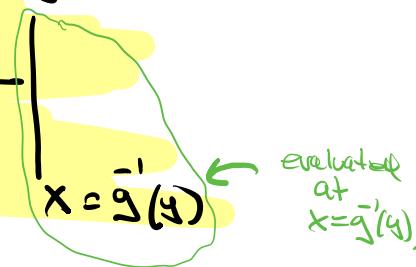
mutually exclusive events.

Thm (P307) Change of Variable Formula for densities

Let X be a continuous RV with density $f_X(x)$.

Let $y = g(x)$ have a derivative that is zero at only finitely many pts.

then $f_Y(y) = \sum_{\{x | g(x)=y\}} \frac{f_X(x)}{|g'(x)|}$



ex

let $X = N(0, 1)$, $f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$

* — we will show later in the semester that this is a density



Find the density of $Y = \sigma X + \mu$ where $\sigma > 0$, $\mu \in \mathbb{R}$

Steps

1) Find $g(x) = \sigma x + \mu$

2) Find $g'(x) = \sigma$

3) Find $x = g^{-1}(y) = \frac{y - \mu}{\sigma}$

4) Find $f_Y(y) = \frac{f_X(x)}{|g'(x)|} \Big|_{x=\frac{y-\mu}{\sigma}}$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \Big|_{x=\frac{y-\mu}{\sigma}}$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2}$$

Note:

$$E(Y) = E(\sigma X + \mu) = \sigma E(X) + \mu = \mu$$

$$\text{Var}(Y) = \text{Var}(\sigma X + \mu) = \sigma^2 \text{Var}(X) = \sigma^2$$

$$\Rightarrow Y \sim N(\mu, \sigma^2)$$

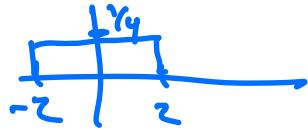
Change of variable formula:

$$f_Y(y) = \sum_{\substack{x \\ \{x \mid g(x)=y\}}} \frac{f_X(x)}{|g'(x)|} \Big|_{x=g^{-1}(y)}$$

ex $X \sim \text{Unif } (-2, 2)$, $f_X(x) = \frac{1}{4} \mathbf{1}_{(x \in (-2, 2))}$

Find density of $Y = X^2$

note: $\begin{aligned} g(y) &= x^2 \\ x &= g^{-1}(y) = \pm \sqrt{y} \end{aligned}$



$$f_Y(y) = \sum_{\substack{x \\ x \in g^{-1}(y) = \{\pm \sqrt{y}\}}} \frac{f_X(x)}{|g'(x)|}$$

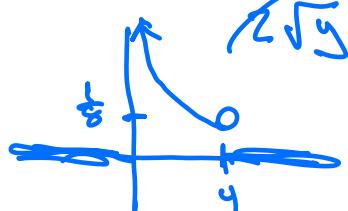
$$\frac{\frac{1}{4} \mathbf{1}_{x \in (-2, 2)}}{|2x|} \Big|_{x=\sqrt{y}} +$$

$$\Big|_{x=-\sqrt{y}}$$

$$\frac{\frac{1}{4} \mathbf{1}_{x \in (-2, 2)}}{|2x|} \Big|_{x=-\sqrt{y}}$$

$$\frac{\frac{1}{4} \mathbf{1}_{\sqrt{y} \in (0, 2)}}{2\sqrt{y}} + \frac{\frac{1}{4} \mathbf{1}_{-\sqrt{y} \in (-2, 0)}}{2\sqrt{y}} \quad \begin{matrix} \sqrt{y} \in (0, 2) \\ -\sqrt{y} \in (-2, 0) \end{matrix}$$

$$= \sum \frac{\frac{1}{4} \mathbf{1}_{\sqrt{y} \in (0, 2)}}{2\sqrt{y}}$$



$$= \frac{\frac{1}{4} \mathbf{1}_{y \in (0, 4)}}{2\sqrt{y}}$$

$$= \begin{cases} \frac{1}{4\sqrt{y}} & 0 < y < 4 \\ 0 & \text{else} \end{cases}$$

Ex (extra problem)

(3 pts) Suppose the random variable X , which measures the magnitude of an earthquake (on the Richter scale) in the Bay Area, follows the Exponential (λ) distribution. Since the Richter scale is logarithmic, we want to study the distribution of the total energy of earthquakes. Find the distribution of $Y = e^X$.

Change of variable formula:

$$f_Y(y) = \sum_{\{x | g(x)=y\}} \frac{f_X(x)}{|g'(x)|} \quad \begin{matrix} \text{evaluated} \\ \text{at} \\ x=g^{-1}(y) \end{matrix}$$

$$\text{Find } g(x) = e^x$$

$$g'(x) = e^x$$

$$g'(x) = \ln y$$

$$f_Y(y) = \frac{\lambda e^{-\lambda x}}{e^x} \quad \left. \begin{matrix} = \lambda e^{(-\lambda-1)x} \\ x = \ln y \end{matrix} \right|_{x=\ln y}$$

$$= \lambda e^{(-\lambda-1)\ln y} = \lambda \left(e^{\ln y} \right)^{(-\lambda-1)} = \boxed{\lambda y^{(-\lambda-1)}, y > 1}$$

