

Stat 134 Lec 16Last time

Sec 3.6 $X \sim \text{Bin}(n, p) \Rightarrow \text{Var}(X) = np(1-p)$

$X \sim \text{HG}(n, N, G) \Rightarrow \text{Var}(X) = np(1-p)\left(\frac{N-n}{N-1}\right)$

where $p = \frac{G}{N}$

Correction factor

Sec 3.1 Geometric distribution ($\text{Geom}(p)$)
 $X = \# p\text{-coin tosses until the first heads}$

→ takes values $1, 2, 3, \dots$

$P(X=k) = q^{k-1} p$

$E(X) = \frac{1}{p}$

$\text{Var}(X) = \frac{q}{p^2}$

Negative Binomial distribution ($\text{NegBin}(r, p)$) $T_r = W_1 + \dots + W_r$ where $W_1, W_2, \dots, W_r \stackrel{iid}{\sim} \text{Geom}(p)$

$P(T_r = k) = \binom{k-1}{r-1} p^r q^{k-r}$

$E(T_r) = \frac{r}{p}$

$\text{Var}(T_r) = \frac{r}{p^2}$

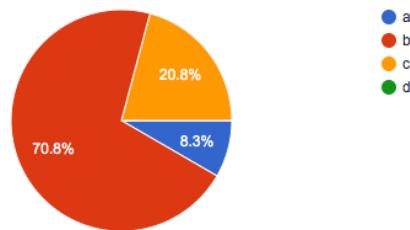
Today

- ① Review concept test from last time
- ② Finish Sec 3.4 Minimum of independent geometrics
- ③ Poisson distribution
- ④ Poisson random scatter (PRS) AKA
Poisson Process
- ⑤ Poisson thinning

(c)

The population of a small town is 1000 with 50% democrats. We wish to know what is a more accurate assessment of the % of democrats in the town, to randomly sample with or without replacement? The sample size is 10.

- a** with replacement
- b** without replacement
- c** same accuracy with or without replacement
- d** not enough info to answer the question



c

When you calculate the var. It's the same

b

The sample size is 10. The population is 1000 with 50% demos. We are expecting 5 out of 10 are demos. We need a tighter percentage. If we do with replacement, the chance of getting demo people will be higher. So we need to do without replacement.

① sec 3.4 Minimum of independent geometrics

Adam, Beth and John independently flip a P_1, P_2, P_3 coin respectively.
 let $X = \# \text{ trials until Adam, Beth or John get a heads.}$

etk	A	TTT	$X_1 \sim \text{Geom}(P_1)$
	B	TTT	$X_2 \sim \text{Geom}(P_2)$
	J	TTH <u>H</u>	$X_3 \sim \text{Geom}(P_3)$
		$X = 3$	

a) what is probability Adam, Beth or John get a head?

$$P = \text{Prob}(A \text{ or } B \text{ or } J \text{ get head})$$

$$= 1 - \text{Prob}(A, B, J \text{ don't get head})$$

$$\boxed{1 - q_1 q_2 q_3}$$

b) what distribution is X ?

$$X = \min(X_1, X_2, X_3) \sim \boxed{\text{Geom}(1 - q_1 q_2 q_3)}$$

② Sec 3.5 Poisson distribution ($\text{Pois}(\mu)$)

$$X \sim \text{Pois}(\mu)$$

$$P(X=k) = e^{-\mu} \frac{\mu^k}{k!} \quad k=0, 1, 2, \dots$$

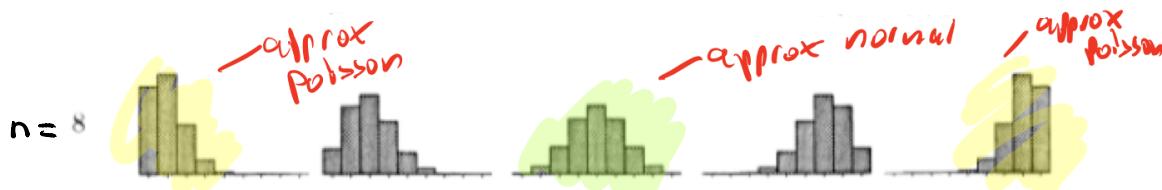
Find $E(X), V(X)$.

$$\text{Recall } e^\mu = 1 + \mu + \frac{\mu^2}{2!} + \dots = \sum_{k=0}^{\infty} \frac{\mu^k}{k!}$$

Taylor Series.

$$\begin{aligned} E(X) &= \sum_{k=0}^{\infty} k \cdot P(X=k) = \sum_{k=0}^{\infty} k e^{-\mu} \frac{\mu^k}{k!} \\ &= \sum_{k=1}^{\infty} k e^{-\mu} \frac{\mu^{k-1} \mu}{(k-1)! k} \quad (\text{note } 0 \cdot e^{-\mu} \frac{\mu^0}{0!} = 0) \\ &= \mu e^{-\mu} \sum_{k=1}^{\infty} \frac{\mu^{k-1}}{(k-1)!} \\ &= \mu e^{-\mu} \left(1 + \mu + \frac{\mu^2}{2!} + \dots \right) = \boxed{\mu} \end{aligned}$$

This makes sense since
 $B(n, p) \rightarrow \text{Pois}(\mu)$ when $n \rightarrow \infty$
 $p \rightarrow 0$,
 $np \rightarrow \mu$.



Also we expect

$$npq \rightarrow mq \approx m \text{ since } q \rightarrow 1$$

so $\text{Var}(X)$ should be M .

lets check:

$$\begin{aligned}\text{Var}(X) &= E(X^2) - E(X)^2 \\ &= E(X^2) - E(X) + E(X) - E(X)^2 \\ &= E(X(X-1)) + E(X) - E(X)^2\end{aligned}$$

$$E(X(X-1)) = \sum_{k=0}^{\infty} k(k-1) P(X=k)$$

$\frac{e^{-\mu} \mu^k}{k(k-1)(k-2)!}$

$$\begin{aligned}&= e^{-\mu} \sum_{k=2}^{\infty} \frac{\mu^k}{(k-2)!} e^{\mu} \\ &= e^{-\mu} \mu^2 \sum_{k=2}^{\infty} \frac{\mu^{k-2}}{(k-2)!} = \mu^2\end{aligned}$$

$$\Rightarrow \text{Var}(X) = \mu^2 + \mu - \mu^2 = \boxed{\mu}$$

ex Let $X \sim \text{Pois}(\mu)$

Find $E(X(X+1))$

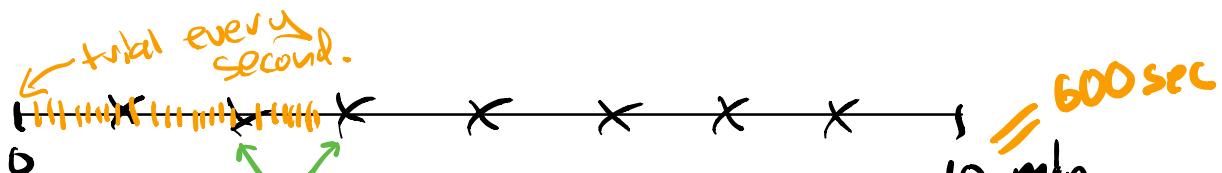
$$\begin{aligned} E(X^2 + X) &= E(X^2) + E(X) \\ &= \text{Var}(X) + (E(X))^2 + E(X) \\ &= \mu + \mu^2 + \mu = \boxed{2\mu + \mu^2} \end{aligned}$$

(3) Poisson Random Scatter (PRS)

$\text{Pois}(np) \approx \text{Bin}(n, p)$ with large n and small p
models situations with low probability, independent trials.

ex X = number of calls coming into a hotel reservation center in 10 minutes,

Say $\mu = 5$ call in 10 min



The distribution of calls
should look random not
clustered since we have
independent trials w/ same p.

PRS assumptions

- 1) No time interval gets more than one call
- 2) Have n iid Bernoulli p trials with $\mu = np$ large n , small p .
(i.e all calls are independent of each other with the same probability)

As $n \rightarrow \infty$ and $p \rightarrow 0$ and $np \rightarrow \lambda$

Let $X = \# \text{ calls in } t \text{ minutes}$,

time of n trials.

Then $X \sim \text{Pois}(\mu)$ ← limit of $\text{Bin}(n, p)$ as $n \rightarrow \infty$, $p \rightarrow 0$, $np \rightarrow \lambda$.

By independence of counts in disjoint time intervals there is a constant rate (intensity) λ of calls per unit time.

ex $\lambda = 5/10 = 1/2$ calls/min In above example,
and $\mu = \lambda t = \frac{1}{2} \cdot 10 = 5$ calls

Stat 134

Friday September 28 2018

1. Which of the following can be modeled as a Poisson Random Scatter with intensity $\lambda > 0$?

- a** The number of blueberries in a 3 cubic inch blueberry muffin
- b** The number of patients entering a doctor's office in a 24 hour period.
- c** The number of times a day a person feels hungry
- d** The number of air pulses counted every second from cars driving over an empty rubber hose lying across a highway between noon and 1pm.
- e** more than one of the above
- no patients outside regular hours*
- not hungry b/tz.*
- not air pulses since come in pairs.*

\leq (number of bb in a bb muffin)

A tub of bb muffin batter has

$$\lambda = 2 \text{ bb/in}^3.$$

A muffin is 3 in^3 .

On average how many bb are there per muffin?

$$\mu = \lambda \cdot 3 = 2 \cdot 3 = 6 \text{ bb}$$

Let $X_1 = \# \text{bb in } 1^{\text{st}} \text{ muffin}$

$$X_1 \sim \text{Pois}(6)$$

Another muffin is 4 in^3 (from the same batter)

Let $X_2 = \# \text{bb in } 2^{\text{nd}} \text{ muffin}$.

a) Find $P(5 \text{ bb in each muffin})$

$$X_1 \sim \text{Pois}(6)$$

$$X_2 \sim \text{Pois}(8)$$

$$P(X_1=5, X_2=5) = \frac{e^{-6} 6^5}{5!} \cdot \frac{e^{-8} 8^5}{5!}$$

b) Find $P(10 \text{ bb total in both muffins together})$

$$X_1 + X_2 \sim \text{Pois}(14)$$

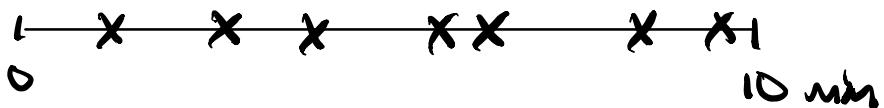
$$P(X_1 + X_2 = 10) = \frac{e^{-14} 14^{10}}{10!}$$

④

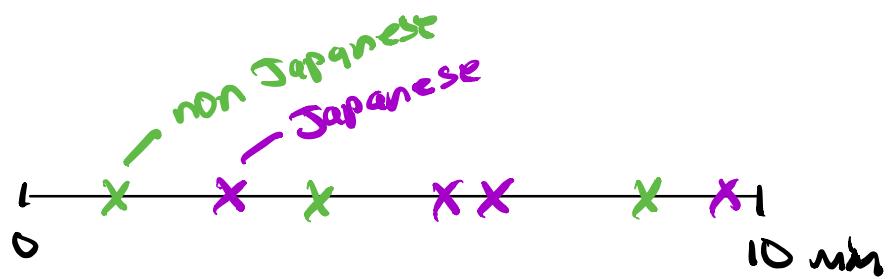
Sec 3.5 Poisson thinning

Cars arrive at a toll booth according to a Poisson process at a rate $\lambda = 3$ arrivals/min

$X = \# \text{ cars arriving at a toll booth}$
in 10 min. $X \sim \text{Pois}(\frac{\lambda \cdot 10}{30})$



Of cars arriving, it is known,
over the long term, that 60% are
Japanese imports.



Call Japanese cars a success and non Japanese a failure.

Each hit is a success with $P = .6$,
independent of all other hits,

Then the process of "success" hits in your
PRG is a PRG with intensity λp and
the process "failure" hits in your
PRG is a PRG with intensity λq

