

Stat 134    lec 28

Weeks up 10:00-10:10

Let  $X \sim \text{Exp}(\lambda)$ ,  $Y \sim \text{Exp}(\mu)$   
 (recall,  $f_X(x) = \lambda e^{-\lambda x}$ )

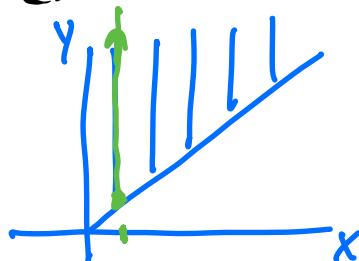
be independent lifetimes of two bulbs.

Find  $P(X < Y)$ .

Hint: use  $f(x,y) = f_X(x)f_Y(y)$

$$f(x,y) = \lambda e^{-\lambda x} \mu e^{-\mu y}$$

$$P(X < Y) = \lambda \int_{x=0}^{\infty} e^{-\lambda x} dx \int_{y=x}^{\infty} e^{-\mu y} dy$$



dy/dx  
dy/dt

$$= \lambda \int_{x=0}^{\infty} e^{-(\mu+\lambda)x} dx = \boxed{\frac{\lambda}{\lambda+\mu}}$$



Last time.

Sec 4.6 Beta Distribution

Let  $r, s > 0$

$P \sim \text{Beta}(r, s)$  if

$$f(p) = \frac{\Gamma(r+s)}{\Gamma(r)\Gamma(s)} p^{r-1} (1-p)^{s-1} \quad \text{for } 0 < p < 1$$

Applications

a)  $\text{Beta}(r, s)$  takes values between 0 and 1 and commonly models the prior distribution of a probability in Bayesian statistics.

b) generalization of standard uniform ordered statistic  
If throw  $n$  darts at  $[0, 1]$   
 $U_{(k)} \sim \text{Beta}(k, n-k+1)$

Today

- ① Sec 5.1, 5.2 Independent RVs
- ② Sec 5.2 Competing exponentials
- ③ Sec 5.2 Marginal density

① Sec 5.1, 5.2  
Independent RVs

Defn  $X$  and  $Y$  are independent if

$$P(X \in dx, Y \in dy) = P(X \in dx) P(Y \in dy)$$

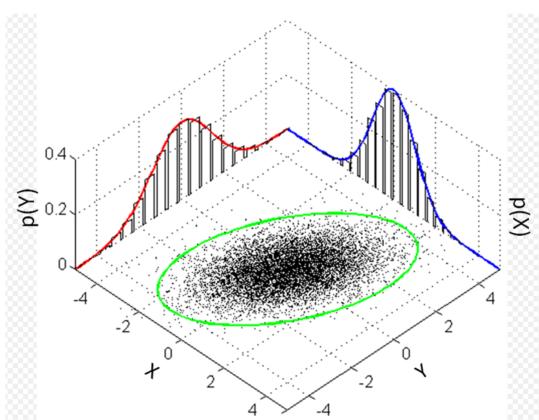
$$\quad \quad \quad f(x,y) dx dy \quad \quad \quad " \quad " \quad " \quad f(x) dx \quad f(y) dy$$

$$\Leftrightarrow f(x,y) = f(x)f(y)$$

Ex  $X, Y \sim_{iid} N(0, 1)$

$$f(x,y) = \phi(x)\phi(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2}$$

$$= \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)}$$



Not a great picture because the oval in green should be a circle. This is the picture of a correlated bivariate normal from chapter 6 instead of an uncorrelated bivariate normal.

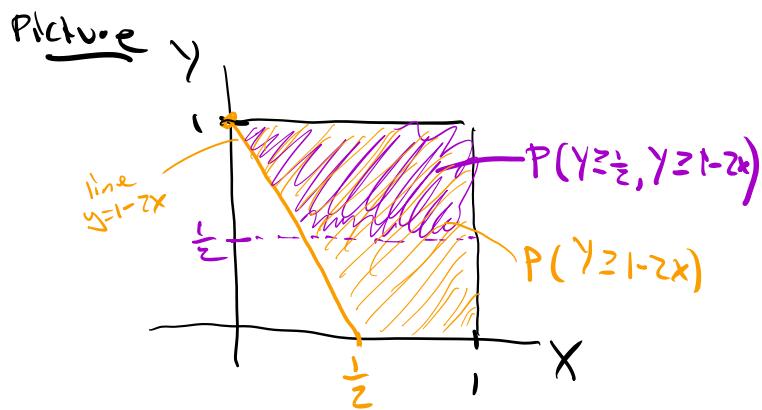
ex If  $X, Y \sim \text{ifd } U(0, 1)$

$$\text{Find } P(Y \geq \frac{1}{2} \mid Y \geq 1 - 2x)$$

Soln

$$f(x, y) = f(x)f(y) = 1 \text{ for } 0 < x, y < 1, 0 \text{ else.}$$

$$P(Y \geq \frac{1}{2} \mid Y \geq 1 - 2x) = \frac{P(Y \geq \frac{1}{2}, Y \geq 1 - 2x)}{P(Y \geq 1 - 2x)} \quad \text{Bayes' rule}$$



so,

$$\frac{P(Y \geq \frac{1}{2}, Y \geq 1 - 2x)}{P(Y \geq 1 - 2x)} = \frac{\frac{1}{2} - \frac{1}{16} = \frac{7}{16}}{1 - \frac{1}{4} = \frac{3}{4}} = \boxed{\frac{7}{12}}$$

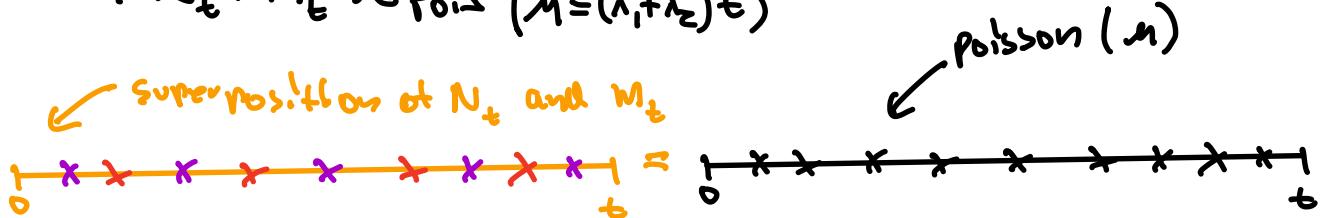
(2) sec 5.2

## Competing exponentials

Superposition of Poisson random scatters:

Let  $N_t \sim \text{Pois}(\lambda_1 t)$  and  $M_t \sim \text{Pois}(\lambda_2 t)$   
be independent PRS corresponding to the number of arrivals of red and purple cars in time  $t$ .

Then  $N_t + M_t \sim \text{Pois}(\lambda = (\lambda_1 + \lambda_2)t)$



Competing exponentials:

Let  $X = \text{time until the first red car}$   
 $Y = \text{time until the first purple car}$

What is the chance the first car is red?

$$P(X < Y) = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

ex

(Exponential distributions) You are first in line to have your question answered by either of the 2 uGSIs, Brian and Yiming, who are busy with their current students. Your question will be answered as soon as the first uGSI finishes. The times spent with student by Brian and Yiming are independent and exponentially distributed with (positive) rates  $\lambda_B$  and  $\lambda_Y$  respectively, i.e. Brian's distribution is Exponential( $\lambda_B$ ), and Yiming's is Exponential( $\lambda_Y$ ).

- (a) Find the probability that Yiming will be the one answering your questions.

$$P(Y < B) = \frac{\lambda_Y}{\lambda_Y + \lambda_B}$$

- (b) What is the distribution of your wait time? Your answer should not include integrals.

$$W = \min(B, Y) \sim \text{Exp}(\lambda_Y + \lambda_B)$$

$$\begin{aligned} P(W > w) &= P(\min(B, Y) > w) \\ &= P(Y > w, B > w) = P(Y > w)P(B > w) \\ &= e^{-\lambda_Y w} \cdot e^{-\lambda_B w} \\ &= e^{-(\lambda_Y + \lambda_B)w} \end{aligned}$$

Think of  $\min(Y, B)$  as a superposition of Poisson processes with rate  $\lambda_Y + \lambda_B$ .

$$\text{So } P(Y = \min(Y, B)) = P(Y < B) = \frac{\lambda_y}{\lambda_y + \lambda_B}$$

If  $X_1, \dots, X_n$  are independent exponentials with rates  $\lambda_1, \dots, \lambda_n$

$$P(X_i = \min(X_1, \dots, X_n)) = ? \quad \frac{\lambda_i}{\lambda_1 + \dots + \lambda_n}$$

Now have 3 GSI, Yilmaz, Bitan and Rowen.  
What's chance Yilmaz done first, then Bitan and then Rowen (independent exponentials with rates  $\lambda_Y, \lambda_B, \lambda_R$ )?

$$\text{i.e. } P(Y < B < R)$$

$$\begin{array}{c} \gamma \quad \gamma \quad B \quad B \quad Y \quad R \\ \hline 0 \end{array}$$

$$P(Y = \min(Y, B, R), B = \min(B, R)) \\ = P(Y = \min(Y, B, R)) \cdot P(B = \min(B, R) \mid Y = \min(Y, B, R))$$

$$\boxed{\frac{\lambda_Y}{\lambda_Y + \lambda_B + \lambda_R} \cdot \frac{\lambda_B}{\lambda_B + \lambda_R}}$$

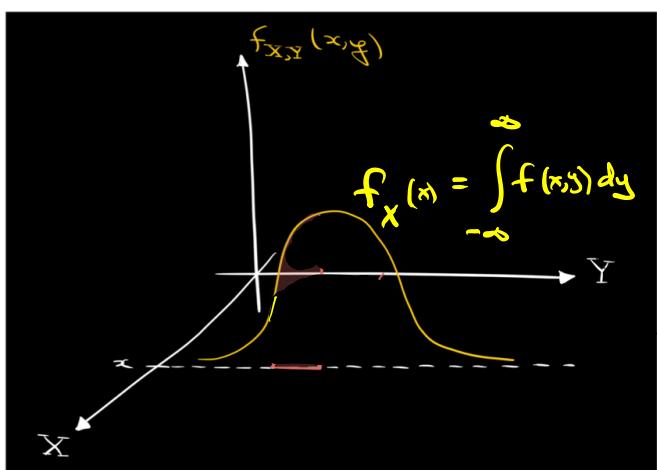
### (3) Sec 5.2 Marginal densities

Recall marginal probability:

discrete picture

		Marginal probability of X	
		$P(x) = \sum_{y \in Y} P(x,y)$	
		0	1
Y	2	0	$\frac{1}{4}$
	1	$\frac{1}{4}$	$\frac{1}{4}$
	0	$\frac{1}{4}$	0
		Marginal Prob of Y	$P(y) = \sum_{x \in X} P(x,y)$
		$\frac{1}{2}$	$\frac{1}{4}$
		$\frac{1}{4}$	$\frac{1}{4}$
		0	1

Continuous Picture: marginal density



$f_X(x)$  is the integral from  $y = -\infty$  to  $y = \infty$  under the yellow curve. You can do this for every  $x$  value.

