

Stat 134 Lec 33

Warmup 11:00 - 11:10

ex Let $X \sim U_{(7)}$, $Y \sim U_{(9)}$ for 10 iid $U(0,1)$.

The joint density $f_{X,Y}(x,y) = C x^6 (y-x)(1-y)$ where $C = \binom{10}{6,1,1,1,1}$
for $0 \leq x < y < 1$.

Find the density of $Z = Y - X$
what distribution is Z ?

Hint: Use the convolution formula $f_Z(z) = \int_0^{1-z} f(x, x+z) dx$

$$C = \binom{10}{6,1,1,1,1}$$

$$f_Z(z) = \int_0^{1-z} C x^6 (x+z-x)(1-(x+z)) dx$$

$$= C z \int_0^{1-z} ((1-z)x^6 - x^7) dx$$

$$= C z \left[(1-z) \frac{x^7}{7} - \frac{x^8}{8} \right] \Big|_{x=0}^{x=1-z}$$

$$= C z \left(\frac{(1-z)^8}{7} - \frac{(1-z)^8}{8} \right) = \frac{C}{56} z (1-z)^8$$

$$Z \sim \text{Beta}(2, 9)$$

$U_{(9)} - U_{(7)}$ should have the same distribution as
 $U_{(7)} \sim \text{Beta}(k, n-k+1)$

$$U_{(2)} - 0 = U_{(2)} \sim \text{Beta}(2, \underbrace{10-2+1}_9) \checkmark$$

Announcement: MT2 Friday 11/19 (take home)
 approx 50 min. M6F, chap 4 (skip sec 4.3),
 Chap 5,
 review materials coming.

Last time

Sec 5.4 Density Convolution Formula of $S = X + Y$
 Assume $X > 0, Y > 0$

$$f_S(s) = \int_{x=0}^{x=s} f(x, s-x) dx$$

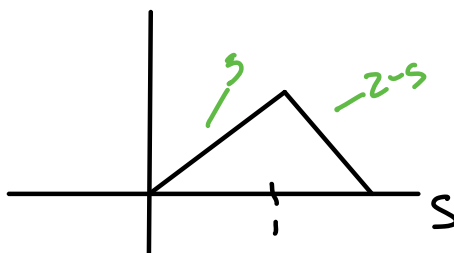
convolution formula for densities.

ex (triangular density)

let $X, Y \stackrel{iid}{\sim} U(0,1)$

$$S = X + Y$$

$$f_S(s) = \begin{cases} s & \text{for } 0 < s < 1 \\ 2-s & \text{for } 1 < s < 2 \end{cases}$$



Convolution formula for $Z = Y - X$ for $0 < X < Y$

$$f_Z(z) = \int_0^{1-z} f(x, x+z) dz$$

Today ① (see #13 p 355) Uniform Spacing

② Sec 5.4 more Convolution Formulas

③ sec 6.1, 6.2 Conditional Distribution, Expectation

discrete case

④ sec 5.4 General Convolution formula

① (see #13 p 355) Uniform Spacing

We saw above

Let $X \sim U_{(7)}$, $Y \sim U_{(9)}$ for iid $U(0,1)$.

then $Z = Y - X \sim \text{Beta}(2, 9)$

We know $U_{(9)} - U_{(7)}$ and $U_{(2)}$

both are $\text{Beta}(2, 9)$

More generally (Uniform Spacing)

You randomly throw n darts at $[0, 1]$.

For $0 \leq k \leq n$, $U_{(k+1)} - U_{(k)}$ is?

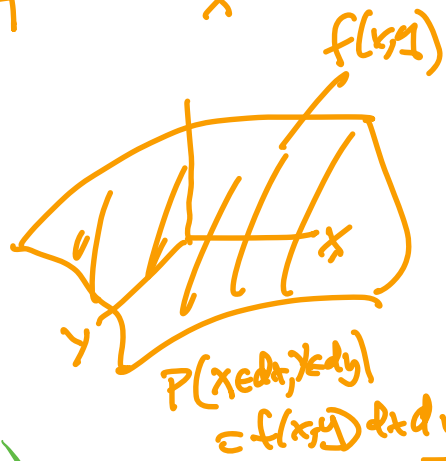
$$U_{(k)} \sim \text{Beta}(k, n-k+1)$$

② Convolution formula for density of ratio Y/X

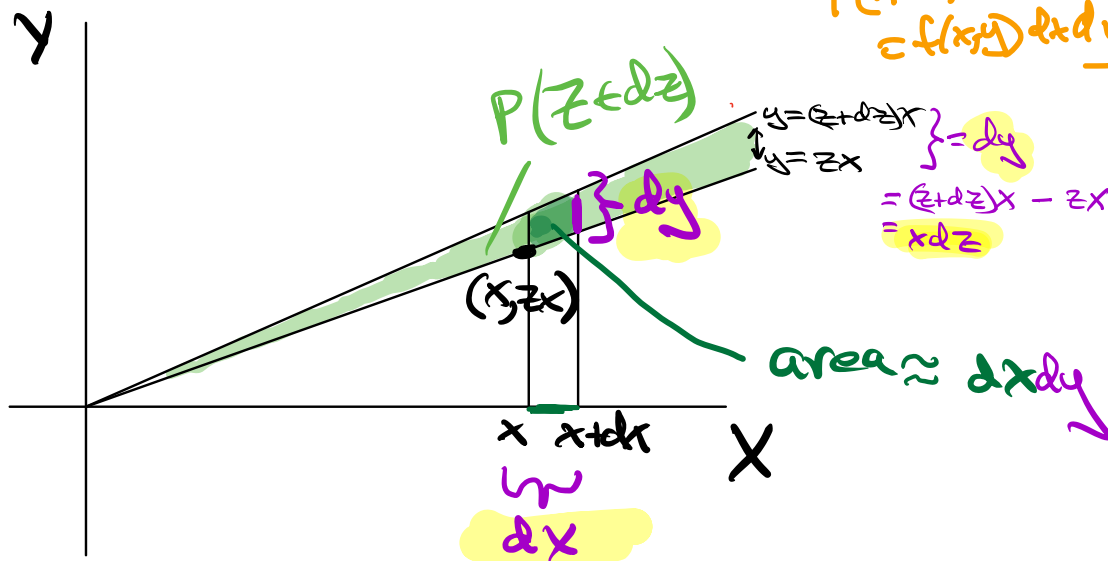
$$X > 0, Y > 0$$

$$\text{let } Z = \frac{Y}{X}$$

Find $f_Z(z)$.



Picture $Y = Z \cdot X$ ^{-slope}



$$\begin{aligned}
 P(Z \in dz) &= \int_{x=0}^{x=\infty} P(Z \in dz, X \in dx) \\
 &= \int_{x=0}^{x=\infty} \int_{y=0}^{y=\infty} f(x, y) dy dx \\
 &= \int_{x=0}^{x=\infty} \int_{y=zx}^{y=(z+dz)x} f(x, y) dy dx \\
 &= \int_{x=0}^{x=\infty} f(x, zx) x dz dx
 \end{aligned}$$

Convolution formula.

$$\Rightarrow f_z(z) = \int_{x=0}^{\infty} f(x, zx) x dx = \int_{x=0}^{\infty} f_x(x) f_y(zx) x dx$$

- if x, y indep.

ex
 Let $x, y \stackrel{iid}{\sim} \text{Exp}(1)$. $z = \frac{y}{x}$ $f_x(x) = e^{-x}$
 Find $f_z(z)$.

Hint: use Convolution formula

$$f_z(z) = \int_{x=0}^{\infty} f_x(x) f_y(zx) x dx$$

$= e^{-x} \cdot e^{-zx}$

$$= \int_0^{\infty} x e^{-(1+z)x} dx$$

variable part of Gamma ($r=z, \lambda=1+z$)

$x \sim \text{Gamma}(r, \lambda)$
 $f_x(x) = \frac{\lambda^r}{\Gamma(r)} x^{r-1} e^{-\lambda x}$

$$= \frac{1}{\text{Constant part of Gamma}(2, 1+z)} = \frac{\Gamma(2)}{(1+z)^2} \quad \text{for } 0 < z < \infty$$

$\frac{(1+z)^2}{\Gamma(2)}$

✓

③ sec 6.1 Conditional Distribution: Discrete case.

let X, N discrete RVs w/ joint distribution $P(X=x, N=n)$.

Bayes' rule

$$P(X=x | N=n) = \frac{P(X=x, N=n)}{P(N=n)}$$

$$\Rightarrow P(X=x, N=n) = P(X=x | N=n) P(N=n)$$

Rule of average conditional probabilities

marginal prob of X

$$\begin{aligned} P(X=x) &= \sum_n P(X=x, N=n) \\ &= \sum_n P(X=x | N=n) P(N=n) \end{aligned}$$

ex

buses in 1 min

green buses in 1 min

Let N have Poisson (λ) distribution. Let X be a random variable with the following property: for every n , the conditional distribution of X given ($N = n$) is binomial (n, p). Find the unconditional distribution of X and state its parameter(s). Show all your work for full credit.

$$P(N=n) = \frac{e^{-\lambda} \lambda^n}{n!}$$

$$P(X=x | N=n) = \frac{n!}{x!(n-x)!} p^x q^{n-x}$$

Find $P(X=x)$

$$P(X=x) = \sum_{n=x}^{\infty} P(X=x | N=n) P(N=n)$$

$$\begin{aligned} &= \sum_{n=x}^{\infty} \frac{n!}{x!(n-x)!} p^x q^{n-x} \frac{e^{-\lambda} \lambda^n}{n!} \\ &= \frac{e^{-\lambda} \lambda^x p^x}{x!} \sum_{n=x}^{\infty} \frac{\lambda^{n-x} q^{n-x}}{(n-x)!} \end{aligned}$$

$$\begin{aligned} &\text{Finish} \\ &\frac{e^{-\lambda(1-q)} (\lambda p)^x}{x!} \end{aligned}$$

$\times \times \times \times \times \times$

$X = \# \text{ green buses in 1 min}$

$p = \text{prob a bus is green}$

$N = \text{superposition of 2 poisson processes}$

$\Rightarrow X \sim \text{Pois}(\lambda p)$ by poisson thinning.

$$= 1 + \lambda q + \frac{(\lambda q)^2}{2!} + \dots = e^{\lambda q}$$

$$X \sim \text{Pois}(\lambda p)$$

