

Friday — I am out of town.

Brian will give in class review

Today

- ① MGF review (Dec 25)
- ② Conditional expectation. (Dec 34)
- ③ Gamma/Poisson distributions. (Dec 22)

$$M_X(t) = E(e^{tX})$$

Recall

$$\underset{X}{M'_X}(0) = E(X)$$

$$\underset{X}{M''_X}(0) = E(X^2)$$

$$\vdots \underset{X}{M^{(k)}_X}(0) = E(X^k)$$

Consider a Pois(λ) RV X

$$P(X=x) = \frac{e^{-\lambda} \lambda^x}{x!} \quad x=0, 1, 2, \dots$$

a) Compute its MGF

b) Compute $E(X)$ by calculating the appropriate derivative of the MGF.

Recall $e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$

~~Ex~~ Consider a Pois(λ) RV X

$$P(X=k) = \frac{e^{-\lambda} \lambda^k}{k!} \quad k=0, 1, 2, \dots$$

a) Compute its MGF

b) Compute $E(X)$ by calculating the appropriate derivative of the MGF.

Soln
a) $M_X(t) = E(e^{tX}) = \sum_{k=0}^{\infty} e^{tk} \frac{e^{-\lambda} \lambda^k}{k!}$
 $= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^t)^k}{k!}$

$$= e^{-\lambda} e^{\lambda e^t} = e^{\lambda e^t - \lambda} = \boxed{e^{\lambda(e^t - 1)}}$$

b) $M'_X(t) = e^{\lambda(e^t - 1)} \cdot \underbrace{\frac{d}{dt} \lambda(e^t - 1)}_{\lambda e^t} = \lambda e^t \cdot e^{\lambda(e^t - 1)}$

$$\Rightarrow M'_X(0) = \boxed{\lambda}$$

Review of conditional expectation Sec 6.4

Properties (from lec 34)

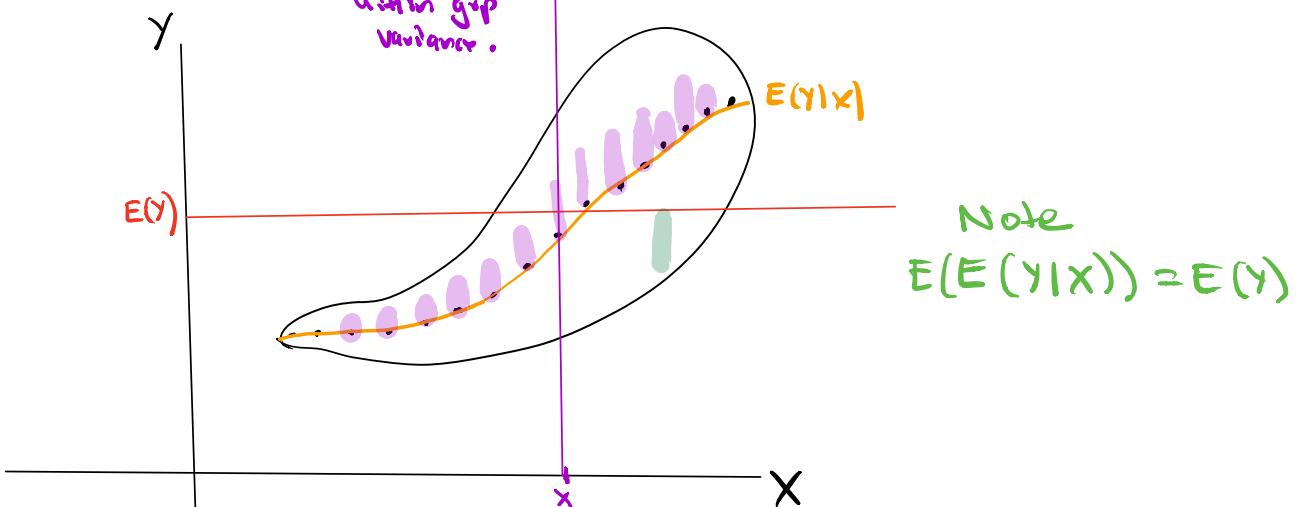
- ① $E(Y) = E(E(Y|X))$ iterated expectation
- ② $E(aY+b|X) = aE(Y|X) + b$
- ③ $E(Y+Z|X) = E(Y|X) + E(Z|X)$
- ④ $E(g(X)|X) = g(X)$
- ⑤ $E(g(X)Y|X) = g(X)E(Y|X)$
- ⑥ $\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X))$

$$\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X))$$

total variance decomposition (sec 6.2.18.)

$$\text{tot var} = \text{avg}(\text{purple}) + \text{green}$$

— var of $E(Y|X)$
(between group variance)



$$X|U=u \sim \text{Exp}\left(\frac{1}{u}\right)$$

i.e. $E(X|U=u) = u$

Ex Let $U \sim \text{Unif}(0,1)$

Given $U=u$, X is exponential with mean u

Find $E(X)$, $E(UX)$ and $\text{Var}(X)$.

Soln

$$\text{Know } E(X|U=u) = u$$

$$E(X|U) = U$$

$$E(X) = E(E(X|U)) = E(U) = \frac{1}{2}$$

$$E(UX) = E(E(UX|U)) = E(UE(X|U))$$

$$= E(U^2) = \frac{1}{3}, \frac{1}{12}, \frac{1}{2}$$

$$\text{Var}(U) + E(U)^2$$

$$\text{Var}(X) = E(\text{Var}(X|U)) + \text{Var}(E(X|U))$$

$$U^2 \quad \frac{1}{3}$$

$$U$$

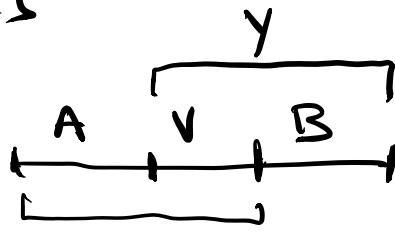
$$= \frac{1}{3} + \frac{1}{12}$$

$$= \boxed{\frac{5}{12}}$$

Ex Toss a fair coin 30 times

$X = \# \text{ heads first 20}$

$Y = \# \text{ heads last 20}$



$$A \sim \text{Bin}(10, \frac{1}{2})$$

$$V \sim \text{Bin}(10, \frac{1}{2})$$

$$B \sim \text{Bin}(10, \frac{1}{2})$$

indep.

$$X = A + V$$

$$Y = V + B$$

Find $\text{Corr}(X, Y)$?

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\text{SD}(X)\text{SD}(Y)}$$

$$= \text{Cov}(A+V, V+B) = \text{Cov}(V, V)$$

$$= \text{Var}(V) = 10 \left(\frac{1}{2}\right)\left(\frac{1}{2}\right)$$

$$= \frac{10}{4}$$

$$\text{Cov}(X, Y) = \frac{10}{4} = \frac{1}{2}$$

$$\text{SD}(X)\text{SD}(Y) = \sqrt{\frac{20}{4}} \cdot \sqrt{\frac{20}{4}}$$

$$\text{Find } E(Y|X) = E(V+B|X)$$

$$= E(V|X) + E(B|X)$$

$$E(B) = 10 \cdot \frac{1}{2} = 5$$

$V|X$ is hypergeometric ($N=20, n=10, G=X$)



↓ draw n w/o replacement

$V|X$ is number of good in sample size n

$$E(V|X) = n \cdot \frac{6}{N} = 10 \cdot \frac{X}{20} = \frac{1}{2}X$$

$$E(Y|X) = \frac{1}{2}X + 5$$

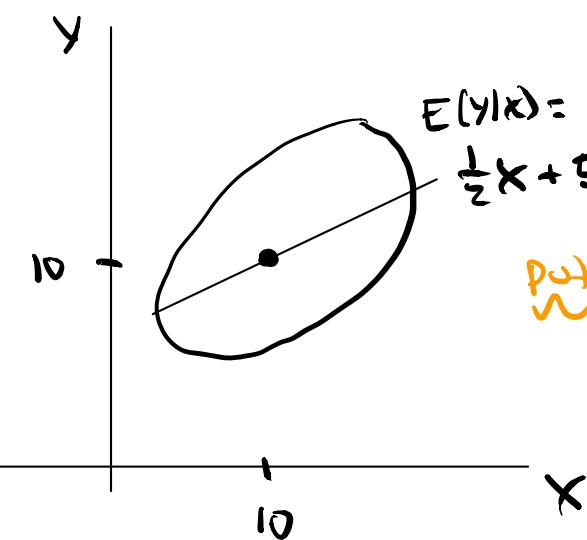
$$X \sim N(10, 5)$$

$$Y \sim N(10, 5)$$

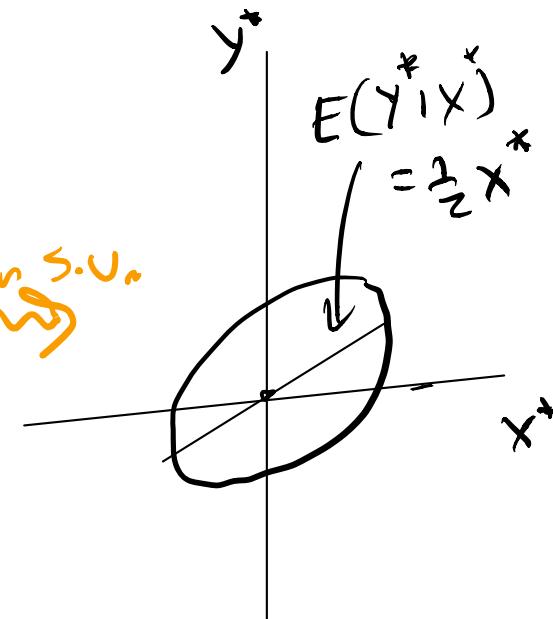
$$\text{Corr}(X, Y) = \frac{1}{2}$$

binomial is approximately normal by CLT

Picture



(x, y) bivariate
normal



(x^*, y^*) std
bivariate normal

Let N have the Poisson distribution with mean μ .
 Let $U_1, U_2, \dots \stackrel{\text{iid}}{\sim} U(0,1)$ independent of N .
 Let $M = \min(U_1, \dots, U_N)$. If $N=0$, define
 M to be 1.

a) Find $E(M|N)$

b) Find $E(M)$

Soln

a) If $n=0$, $M=1$ so $E(M|N=0)=1$

else $M|N \sim \text{beta}(1, N)$

$$\Rightarrow E(M|N) = \frac{1}{N+1}$$

b) $E(M) = E(E(M|N)) = E\left(\frac{1}{N+1}\right)$

$$= \sum_{n=0}^{\infty} \frac{1}{n+1} \bar{e}^{-\mu} \frac{\mu^n}{n!} = \frac{1}{\mu} \sum_{n=0}^{\infty} \bar{e}^{-\mu} \frac{\mu^{n+1}}{(n+1)!}$$

$$= \frac{1}{\mu} \left[\frac{\bar{e}^{-\mu} \mu^0}{0!} + \frac{\bar{e}^{-\mu} \mu^1}{1!} + \dots \right]$$

$$= \frac{1}{\mu} \left[\sum_{n=1}^{\infty} \bar{e}^{-\mu} \frac{\mu^n}{n!} \right] = \frac{1 - \bar{e}^{-\mu}}{\mu}$$

$$\stackrel{!!}{P(N \geq 1)} = 1 - P(N=0)$$

$$= 1 - \bar{e}^{-\mu}$$

Let N have the Poisson distribution with mean μ .
 Let $U_1, U_2, \dots \stackrel{\text{iid}}{\sim} N(0, 1)$ independent of N .
 Let $M = \min(U_1, \dots, U_N)$, If $N=0$, define
 M to be 1.

c) find the survival function of M

$$P(M \geq m)$$

$$\underline{P(M \geq m)} = \sum_{n=0}^{\infty} P(M \geq m | N=n) P(N=n)$$

$$\text{For } n \geq 0, P(M \geq m | N=n) = P(U_1 \geq m, U_2 \geq m, \dots, U_n \geq m) \\ = (1-m)^n$$

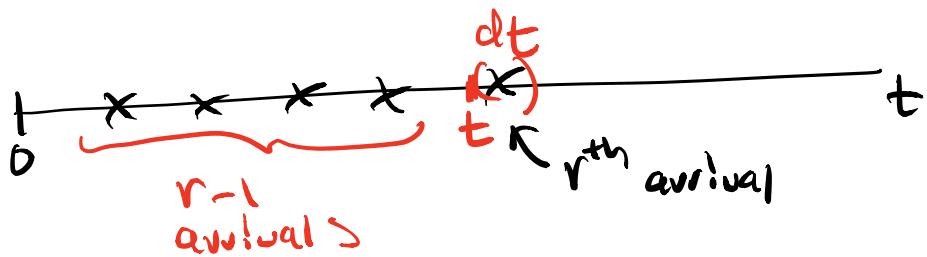
$$\Rightarrow P(M \geq m) = 1 - e^{-m} + \sum_{n=1}^{\infty} (1-m)^n e^{-m} \frac{m^n}{n!}$$

$$= \sum_{n=0}^{\infty} (1-m)^n e^{-m} \frac{m^n}{n!} \\ = e^{-m} \left[\sum_{n=0}^{\infty} \frac{(1-m)^n m^n}{n!} \right] \\ = e^{-m + m - m^2} = e^{-m^2}$$

$$\text{So } P(M \geq m) = \begin{cases} 1 & m < 0 \\ e^{-m^2} & 0 \leq m < 1 \\ 0 & m \geq 1 \end{cases}$$

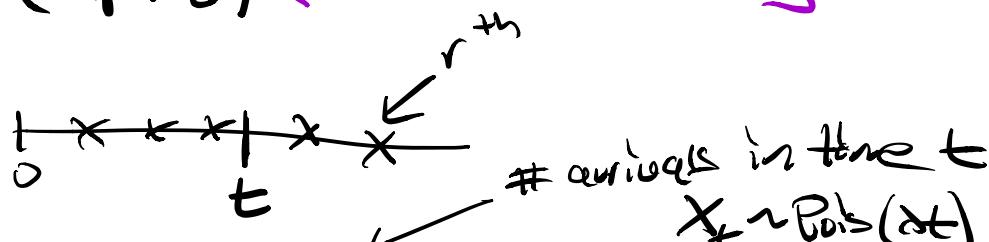
Sec 4.2 Gamma (r, λ) = distribution for r^{th} arrival time of a Poisson (λ) process.

T_r = arrival time of r^{th} call.



Let $T_r \sim \text{Gamma}(r, \lambda)$

Find $P(T_r > t)$ (right tail probability)



$$P(T_r > t) = P(X_t \leq r-1)$$

$$= P(X_t = 0) + P(X_t = 1) + \dots + P(X_t = r-1)$$

$$= e^{-\lambda t} + e^{-\lambda t} \frac{\lambda t}{1!} + \frac{e^{-\lambda t} (\lambda t)^2}{2!} + \dots + \frac{e^{-\lambda t} (\lambda t)^{r-1}}{(r-1)!}$$

Ex

Let $T \sim \text{Gamma}(r=4, \lambda=2)$

Find $P(T > 7)$

Ex
Let $T \sim \text{Gamma}(n=4, \lambda=2)$

Find $P(T > 7)$

Soln

$$P(T > 7) = P(N_7 \leq 3)$$

where
 $N_7 \sim \text{Pois}(14)$

$$P(T > 7) = e^{-14} \left(1 + \frac{14}{1} + \frac{14^2}{2!} + \frac{14^3}{3!} \right)$$