

Last time

Stat 134 lec 27

sec 4.5 Expectation of a non-negative RV using CDF

$$E(X) = \int_0^{\infty} (1 - F(x)) dx$$

e.g. let $X \sim \text{Geom}\left(\frac{1}{2}\right)$

$$P(X=1) = \frac{1}{2}$$

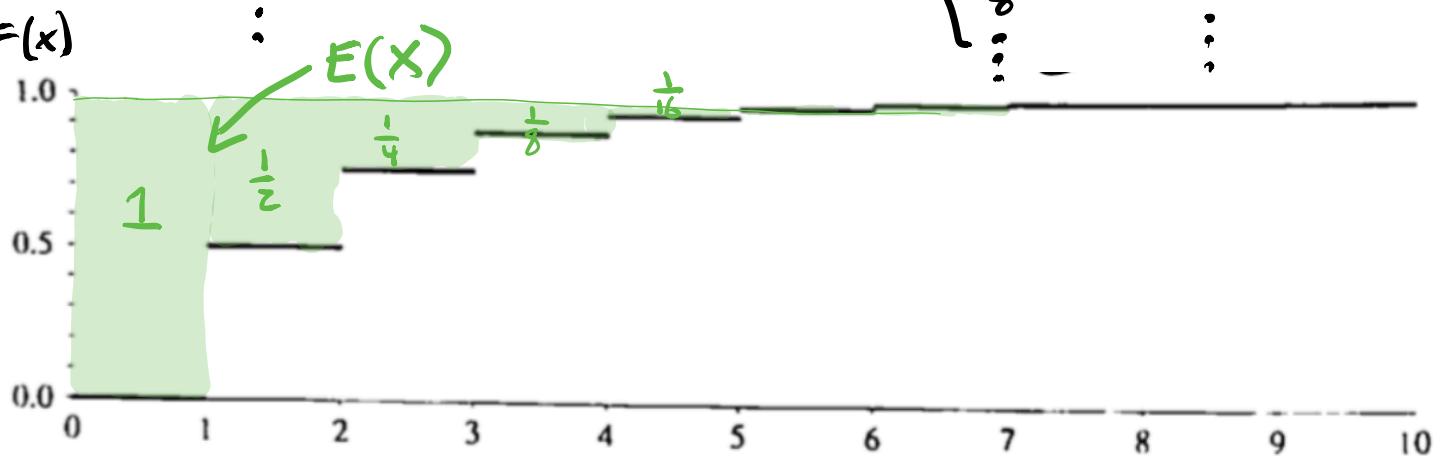
$$P(X=2) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

$$P(X=3) = \left(\frac{1}{2}\right)^2 \cdot \frac{1}{2} = \frac{1}{8}$$

Picture

$F(x)$

$$F(x) = \begin{cases} 0 & x < 1 \\ \frac{1}{2} & 1 \leq x < 2 \\ \frac{3}{4} & 2 \leq x < 3 \\ \frac{7}{8} & 3 \leq x < 4 \\ \vdots & \vdots \end{cases}$$



$$E(X) = \int_0^{\infty} (1 - F(x)) dx$$

$$E(X) = 1 + \frac{1}{2} + \frac{1}{4} + \frac{1}{8} + \dots = \frac{1}{1 - \frac{1}{2}} = \frac{1}{\frac{1}{2}} = 2$$

$$= \sum_{j=0}^{\infty} P(X > j) = \sum_{j=0}^{\infty} \left(\frac{1}{2}\right)^j \quad \leftarrow \text{tail sum formula, } q_j$$

Today

- ① Review student explanations of MGF concept test,
- ② sec 4.6 Order statistics
- ③ sec 5.1 continuous joint distributions,

① Student response

Let X have density $f(x) = xe^{-x}$ for $x > 0$.

The MGF is?

- a $M_X(t) = \frac{1}{1-t}$ for $t < 1$
- b $M_X(t) = \frac{1}{(1-t)^2}$ for $t < 1$
- c $M_X(t) = \frac{1}{(1+t)^2}$ for $t > -1$
- d none of the above

b

This is Gamma(2,1), so the MGF is $(1/(1-t))^2$

match xe^{-x} with
 $x^{r-1}e^{-\lambda x}$

$$\Rightarrow X \sim \text{Gamma}(2, 1)$$

$$\text{Know } M_X(t) = \left(\frac{\lambda}{\lambda-t}\right)^r, t < \lambda$$

b

Using wolfram alpha I know that the integral of $x^r e^{tx} e^{-x}$ dx from zero to infinity is $1/(1-t)^r$ for $t < 1$ thus b is the answer

b

Computing the moment generating function:
 $M_X(t) = E(e^{tX}) = \int_0^\infty e^{tx} xe^{-x} dx = \int_0^\infty xe^{t-x} dx.$
 Rewriting the integral, we produce the following:
 $\int_0^\infty x e^{t-x} dx = \int_0^\infty x e^{-(1-t)x} dx.$
 Looking closely at the above expression, we can clearly see that it is the expectation of an Exponential Distribution with $\lambda = 1-t$ (or equivalently Gamma Distribution with $r = 1$ and $\lambda = 1-t$), multiplied by a factor t . Using this knowledge, we can conclude that $M_X(t) = t^2$ with $t < 1$. The constraint of $t < 1$ is enforced to preserve convergence.

$$M_X(t) = E(e^{tx})$$

$$= \int_0^\infty e^{tx} xe^{-x} dx$$

$$= \int_0^\infty x e^{t-x} dx$$

$$= \frac{1}{1-t} \int_0^\infty x(1-t)e^{-x} dx$$

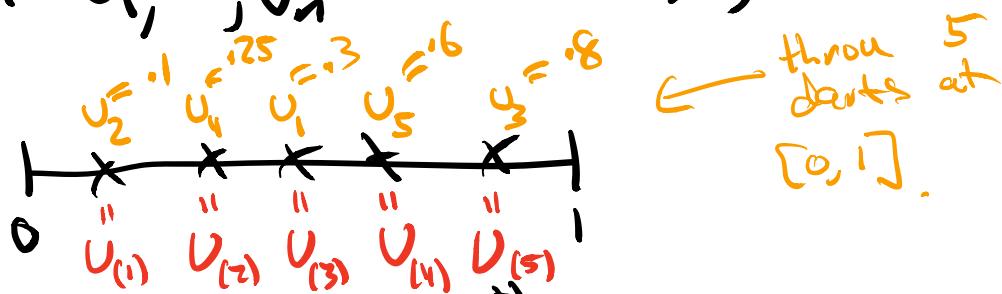
$$= \frac{1}{1-t} \int_0^\infty x \lambda e^{-\lambda x} dx$$

$$= \frac{1}{1-t} \cdot \frac{1}{1-t}$$

expectation of
Exp(1-t)

② Sec 4.6 order statistic at $U(0,1)$

let $U_1, \dots, U_n \sim \text{Unit}(0,1)$ iid



let $U_{(k)}$ = called the k^{th} order statistic
 = k^{th} largest value of U_1, \dots, U_n
 (assuming no ties)

$$U_{(1)} = \min(U_1, \dots, U_n)$$

$$U_{(n)} = \max(U_1, \dots, U_n)$$

Review counting
 You have 3 red, 2 green and 5 blue marbles,
 How many orderings of these 10 marbles are there?

ex rrr ggg bbb bb

grrr g bb b bb

ggrrr bb b bb

:

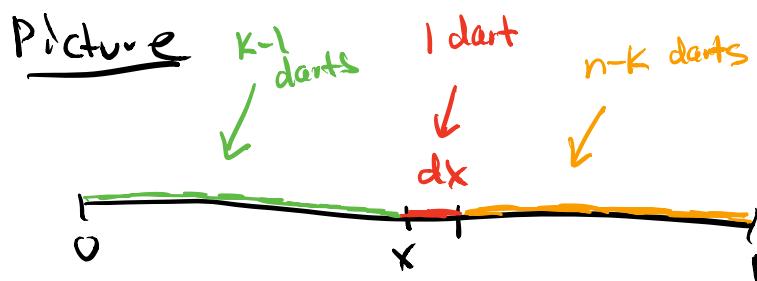
Answer

$$\binom{10}{3,2,5} = \binom{10}{3} \binom{7}{2} \binom{5}{5}$$

$$\frac{10!}{3!2!5!}$$

Next, find density of $U_{(k)}$

$$\text{write } P(U_{(k)} \in dx) = f(k)dx$$



$U_{(k)} \in dx$ means that $k-1$ darts are between 0 and x , and one is in dx , and $n-k$ darts are between x and 1

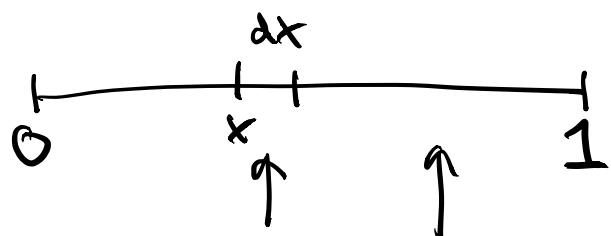
$$\begin{aligned}
 P(U_{(k)} \in dx) &= P(k-1 \text{ darts} \in (0, x), 1 \text{ dart} \in dx, n-k \text{ darts} \in (x, 1)) \\
 &= P(k-1 \text{ darts} \in (0, x)) \cdot P(1 \text{ dart} \in dx \mid k-1 \text{ darts} \in (0, x)) \\
 &\quad \cdot P(n-k \text{ darts} \in (x, 1) \mid 1 \text{ dart} \in dx, k-1 \text{ darts} \in (0, x)) \\
 &= \binom{n}{k-1} x^{k-1} \binom{n-k+1}{1} dx \binom{n-k}{n-k} (1-x)^{n-k} \\
 &= \binom{n}{k-1, 1, n-k} x^{k-1} (1-x)^{n-k+1-1} dx \\
 &\qquad\qquad\qquad f_{U_{(k)}}(x)
 \end{aligned}$$

$$\Rightarrow f_{U_{(k)}}(x) = \binom{n}{k-1, 1, n-k} x^{k-1} (1-x)^{(n-k+1)-1}$$

for
 $0 < x < 1$

\Leftarrow Let $U_1, \dots, U_n \stackrel{\text{iid}}{\sim} \text{Unif}(0, 1)$

Find the density of $U_{(1)}$



$$f_{U_{(1)}}(x)dx = \binom{n}{1, n-1} dx (1-x)^{n-1}$$

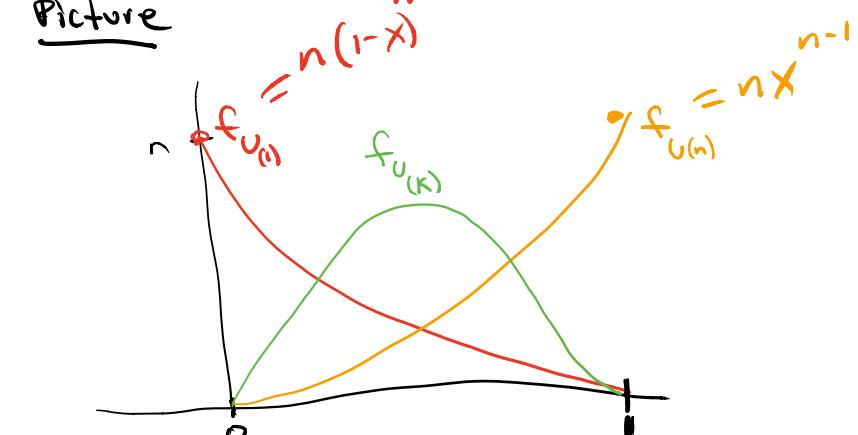
$$= n (1-x)^{n-1} dx$$

\Leftarrow Let $U_1, \dots, U_n \stackrel{iid}{\sim} \text{Unif}(0, 1)$

Find the density of $U_{(n)}$

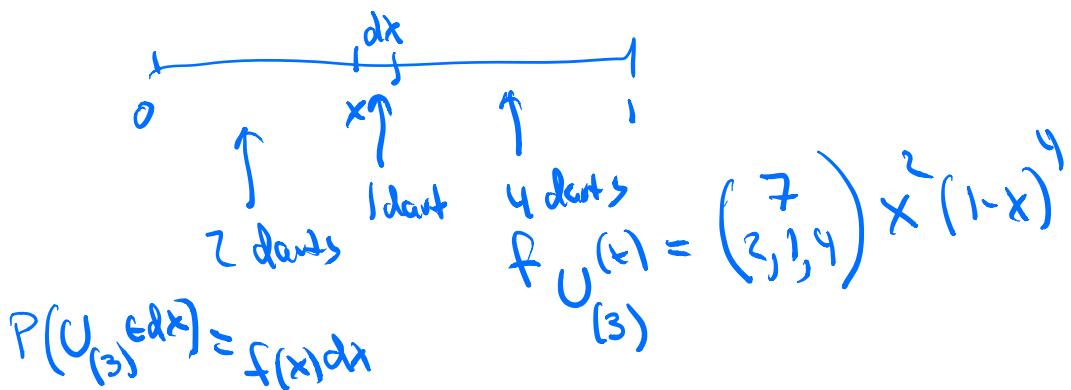
$$f_{U_{(n)}}(x) dx = \binom{n}{n-1} x^{n-1} dx = nx^{n-1} dx$$

Picture

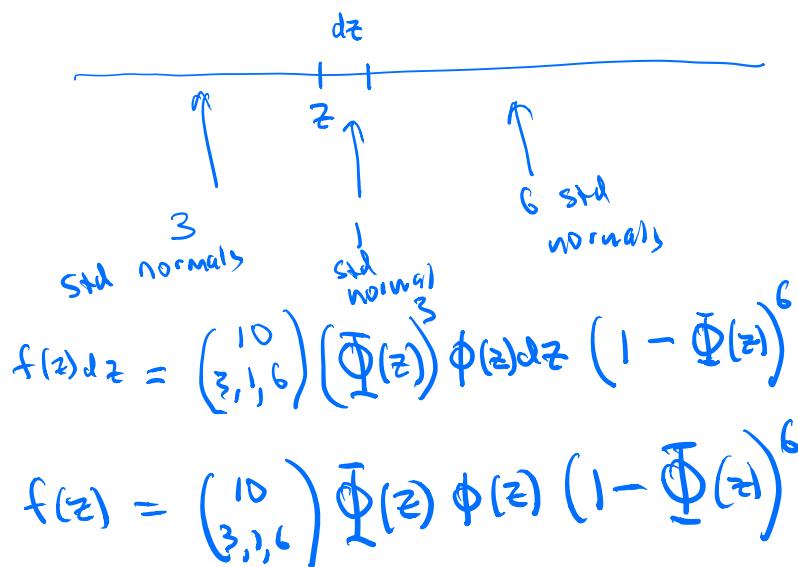


Order statistic of $U(0,1)$ provides a family of densities on the unit interval.

Ex $x^2(1-x)^4$ for $0 < x < 1$ is the
variable part of what RV? How many
darts do you throw?



Ex Let $Z_{(1)}, \dots, Z_{(10)}$ be the values of 10 independent standard normal variables arranged in increasing order. Find the density of $Z_{(4)}$



Chap 5 Continuous Joint Dist

sec 5.1, 5.2

x, y have joint density $f(x, y)$
means f must satisfy
 $f(x, y) \geq 0$ ← think of this as a surface over the plane

$$\iint_{\mathbb{R}^2} f(x, y) dx dy = \iint_{\mathbb{R}^2} f(x, y) dy dx = 1$$

the total volume under the surface is 1

Let A be a subset of the plane

$$P((x, y) \in A) = \iint_A f(x, y) dx dy$$

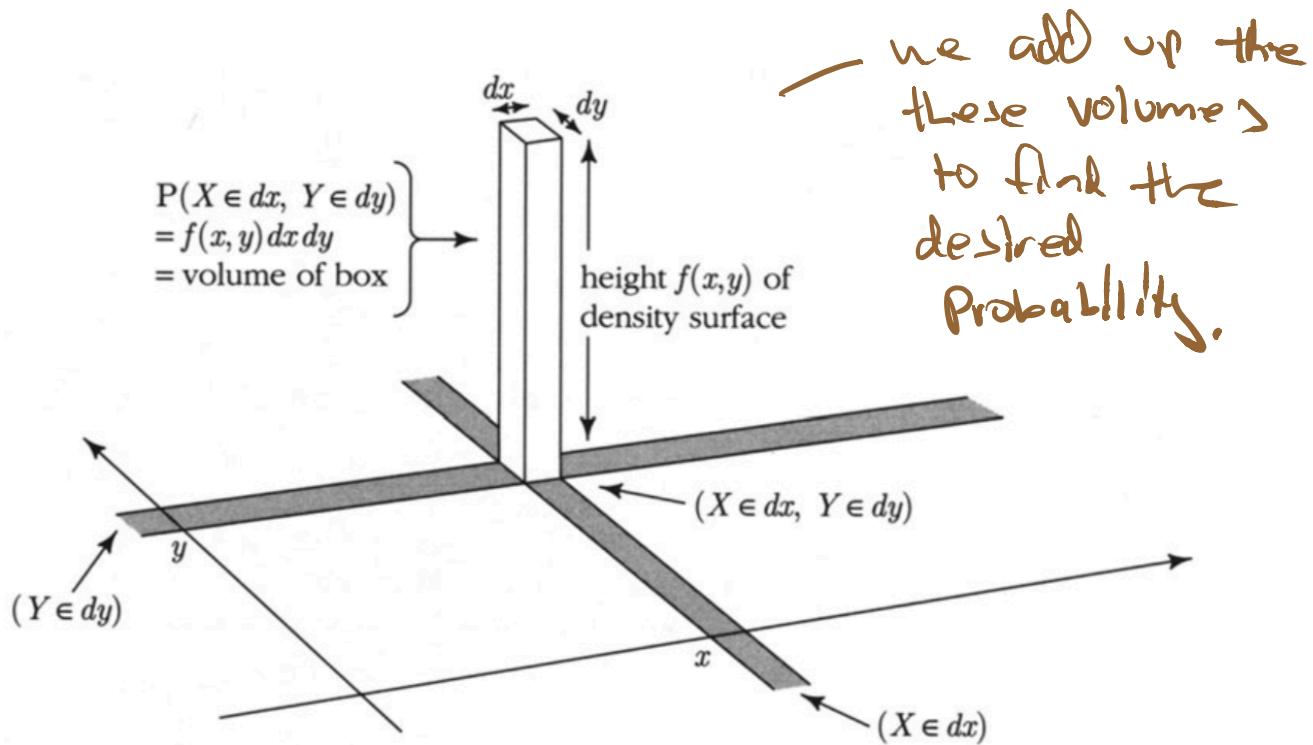
A

the volume of the surface over region A . This is a number between 0 and 1.

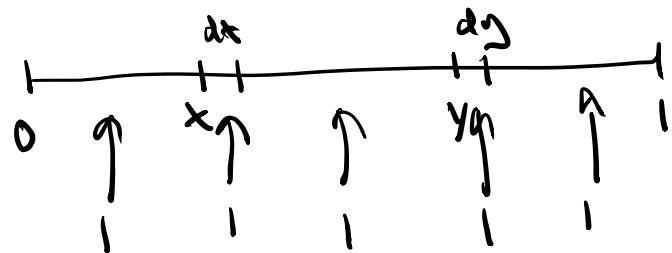
$$P(X \in dx, Y \in dy) = f(x, y) dx dy$$

the volume of the surface over a little rectangle in the plane.

Picture



Ex Throw down 5 darts on $(0,1)$.
 Find the joint density of
 $X \in U_{(2)}$ and $Y \in U_{(4)}$.



$$f(x,y) = \binom{5}{1,1,1,1,1} \times dx \times (y-x) \times dy \times (1-y)$$

$$= \binom{5}{1,1,1,1,1} \times (y-x)(1-y)$$

"5!" $0 < x < y < 1$

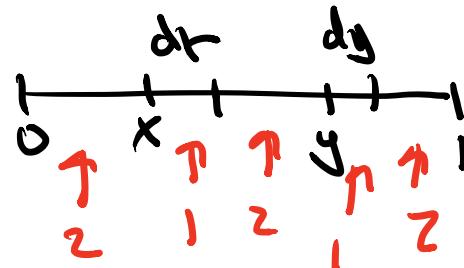
$$\Rightarrow \boxed{f(x,y) = 5! \times (y-x)(1-y) \quad \text{for } 0 < x < y < 1}$$

Stat 134

Wednesday April 3 2019

1. I throw down 8 darts on $(0, 1)$. The variable part of the joint density of $X = U_{(3)}$ and $Y = U_{(6)}$ is:

- a $x(y - x)^5(1 - y)^2$
- b $x^2(y - x)^2(1 - y)^2$
- c $x^4(y - x)^2(1 - y)^2$
- d none of the above



2. Is $f(x, y) = \binom{6}{1,4,1}(y - x)^4$ on $0 < x < y < 1$ a joint density function?

a yes

b no

c not enough info to decide

