

Last timeSec 4.1 Continuous distributions,① Probability density function

$$f(x) \geq 0 \text{ for } x \in X$$

and $\int_{-\infty}^{\infty} f(x) dx = 1$

$$P(a \leq X \leq b) = \int_a^b f(x) dx$$

$$P(X=a) = \int_a^a f(x) dx = 0$$

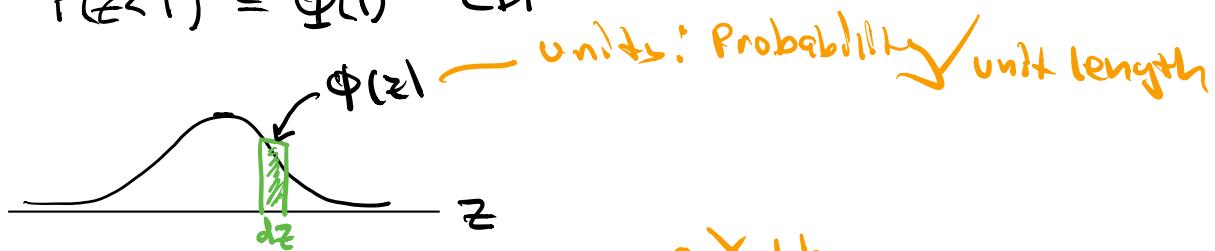
$$\text{so } P(X \geq a) = P(X > a)$$

$$\Leftrightarrow Z \sim N(0, 1)$$

$$\Phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

density function

$$P(Z < 1) = \Phi(1) \quad \text{CDF}$$



$$\textcircled{3} \text{ Change of scale of RV } X - aX + b$$

$$\Leftrightarrow X \sim N(\mu, \sigma^2)$$

$$\text{Find } P(X > a)$$

$$\text{Make change of scale } z = \frac{X-\mu}{\sigma} = \frac{1}{\sigma}X + \frac{-\mu}{\sigma}$$

$$\Rightarrow X = \sigma z + \mu$$

$$\begin{aligned} P(X > a) &= P(\sigma z + \mu > a) = P\left(z > \frac{a-\mu}{\sigma}\right) \\ &= \boxed{1 - \Phi\left(\frac{a-\mu}{\sigma}\right)} \end{aligned}$$

To do Sec 4.5 Cumulative Distribution Function (CDF)
Sec 4.2 Exponential Distribution

Sec 4.5 The Cumulative Distribution Function (CDF).

Let X be a RV. The CDF of X is $F(x) = P(X \leq x)$. — a number between 0 and 1.

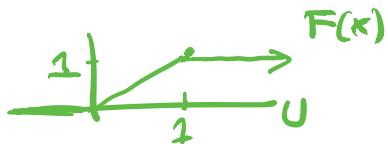
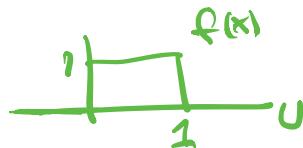
If $f(x)$ is the density of X ,

$$P(X \leq x) = \int_{-\infty}^x f(x) dx$$

e.g. $X \sim \text{Unif}(0, 1)$

$$f(x) = \begin{cases} 1 & 0 < x < 1 \\ 0 & \text{else} \end{cases}$$

$$F(x) = \int_0^x 1 dx = x$$



By the FTC, $F'(x) = f(x)$.

Consequently a density functions and cdf are equivalent descriptions of a RV.

Sec 4.2 Exponential and Gamma Distribution

Memoryless Property of the geometric distribution

$$X \sim \text{Geom}(p)$$

$X = \# \text{ trials until your first success.}$

$$P(X=k) = q^{k-1} p$$

$$P(X>k) = q^k$$

$$\text{Note that } P(X=k) = P(X \geq k-1) - P(X \geq k)$$

$$= q^{k-1} \cancel{q^k} - \cancel{q^{k-1}} q^k = q^{k-1} p$$

$$\text{so } P(X \geq k) = q^k \Rightarrow X \text{ is Geom}(p).$$

Theorem

the geometric distribution is the only discrete distribution with values $1, 2, 3, \dots$ having the memoryless property

$$P(X > k+j | X > j) = P(X > k)$$

In words:

If it takes you more than $j=10$ coin tosses to get your first heads, what is the chance it will take you more than $k+j=13$ coin tosses to get your first heads?

Ans: The same as the unconditional chance of needing more than $k=3$ coin tosses to get your first head,

PF/ Let $X \sim \text{Geom}(p)$, $X=1, 2, 3, \dots$

$$P(X > k+j \mid X > j) = \frac{P(X > k+j, X > j)}{P(X > j)}$$

$$= \frac{P(X > k+j)}{P(X > j)} = \frac{q^{k+j}}{q^j} = q^k$$
$$= P(X > k),$$

Conversely,

for positive integers k, j ,
Suppose $P(X > k+j \mid X > j) = P(X > k)$.

$$\frac{P(X > k+j)}{P(X > j)}$$

$$\Rightarrow P(X > k+j) = P(X > j)P(X > k)$$

$$\text{let } q = P(X > 1)$$

we show that $P(X > j) = q^j$ by induction.

base case:

$$P(X > 1) = q^1 \checkmark$$

$$\text{Assume } P(X > j-1) = q^{j-1}$$

$$P(X > j) = P(X > \boxed{j-1} + 1) = P(X > j-1)P(X > 1) = q^j$$

$$\Rightarrow X \sim \text{Geom}(p),$$

□

Only 2 distributions are memoryless

For discrete ($X: 1, 2, 3, \dots$) — geometric

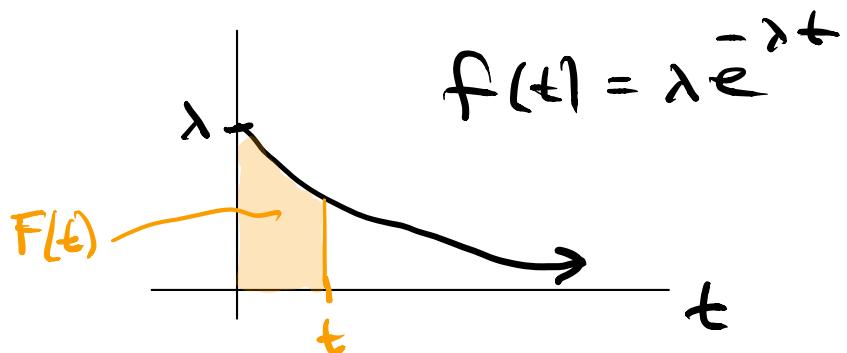
For continuous ($T > 0$) — exponential

Exponential distribution

Proof
Similar to
geometric

Def'n A random time T has exponential distribution with rate $\lambda > 0$

$T \sim \text{expon}(\lambda)$, if T has density $f(t) = \lambda e^{-\lambda t}$ $t \geq 0$



T = time until your first success
where λ is rate of success,

$\Leftrightarrow T$ = time until a lightbulb burns out,

CDF and survival function

$T \sim \text{expon}(\lambda)$

Compute the CDF $F(t) = P(T \leq t)$

$$F(t) = \int_0^t f(s)ds = \int_0^t \lambda e^{-\lambda s} ds = \frac{\lambda e^{-\lambda s}}{-\lambda} \Big|_0^t \\ = -e^{-\lambda t} + 1 = \boxed{1 - e^{-\lambda t}}$$

$P(T > t)$ is called the survival function

$$\boxed{P(T > t) = e^{-\lambda t}}$$

Since $F(t) = 1 - P(T > t)$ and $f(t) = F'(t)$,

$$P(T > t) = e^{-\lambda t} \Rightarrow T \sim \text{expon}(\lambda)$$

Hence, $T \sim \text{expon}(\lambda)$ iff $P(T > t) = e^{-\lambda t}$.

Can we think of geometric as a discrete version of exponential? - yes.

let $X \sim \text{Geom}(p)$ $X=1, 2, 3, \dots$

let ρX be a change of scale of X

$$\rho X = p, 2p, 3p, \dots$$

We will show that $\rho X \rightarrow \text{expon}(1)$

as $p \rightarrow 0$.

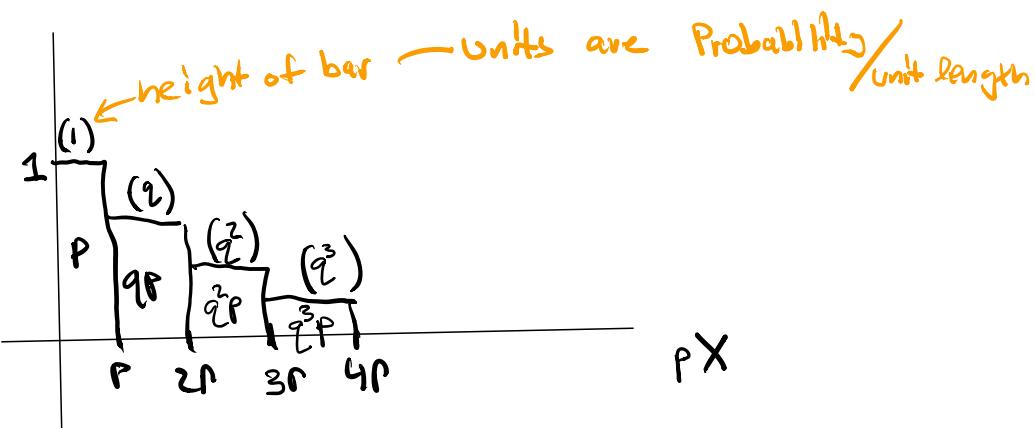
$$P(\rho X = p) = P(X=1) = p$$

$$P(\rho X = 2p) = P(X=2) = q_1 p$$

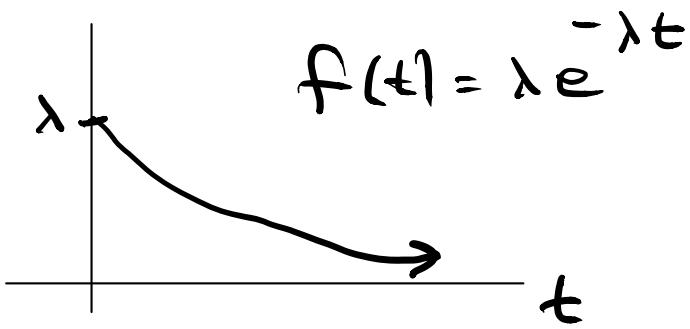
$$P(\rho X = 3p) = P(X=3) = q_1^2 p$$

⋮

Draw the prob mass function for ρX ?



Compare this with the density of expon (λ).



Review Taylor series for

$$e^x = 1 + x + \frac{x^2}{2!} + \dots \approx 1 + x \quad \text{for small } x.$$
$$\Rightarrow \boxed{e^{-x} \approx 1 - x} \quad \text{for small } x$$

The height of bar at time $t = kp$ sec

$$\text{is } q^k = (1-p)^k \approx \underbrace{(e^{-p})^k}_{p \text{ small}} = e^{-kp} = e^{-t}$$

$$\text{as } p \rightarrow 0 \quad f(t) = 1 e^{-t}$$

$\Rightarrow pX \sim \text{expon}(\lambda=1)$ for very small p .

The height of the first bar is always 1 (no matter what p is).

The units of the height is probability

As $p \rightarrow 0$, the height of the first bar stays at 1 and we get $\lambda = \frac{1}{p}$.

λ is called the "instantaneous death rate".

It is the y intercept in the graph of the density function.

Interpretation: λ is the prob of success
during an infinitesimally small time interval,

In summary: For $X \sim \text{geom}(p)$, $\lambda X \rightarrow \text{exp}(\lambda)$ as $p \rightarrow 0$.

ex

- Suppose X follows an exponential distribution with parameter $\lambda = 5$. Let $Y = \lceil X \rceil$. $\lceil X \rceil$ is called the ceiling function of X : it gives the smallest integer $Y \geq X$. Example: $\lceil 3.14 \rceil = 4$. Identify the distribution Y . Please show all of your work below. (Hint: Can you find $P(Y = y)$?)

sln

$y \in \mathbb{N}$

$$\begin{aligned} P(Y=y) &= P(X \in (y-1, y]) \\ &= P(X > y-1) - P(X > y) \\ &= e^{-5(y-1)} - e^{-5y} \end{aligned}$$

$$= e^{-5y+5} - e^{-5y}$$

$$= e^{-5y} (e^5 - 1)$$

$$= e^{-5y} e^5 (1 - e^{-5})$$

$$= \underbrace{e^{-5(y-1)}}_q \underbrace{(1 - e^{-5})}_p$$

$$= q^{y-1} p$$

$$\rightarrow Y \sim \text{Geom}(1 - e^{-5})$$

