

Last time

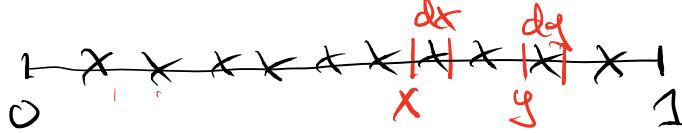
Sec 5.4 Convolution formula for density of sum

Let  $X, Y$  be RVs and  $S = X + Y$ ,

$$f_S(s) = \int_{x=0}^{x=s} f(x, s-x) dx \stackrel{X, Y \text{ independent}}{\underset{x=0}{\underset{x=s}{=}}} \int_x^s f(x) f(s-x) dx$$

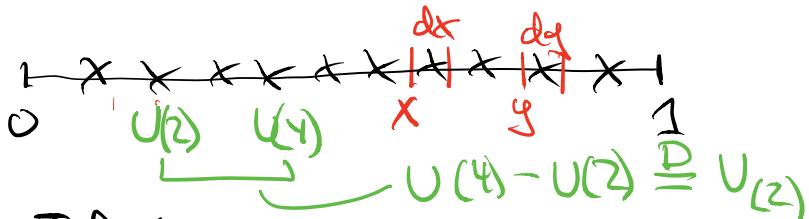
$\cong$  (Uniform spacing) (See #13 p 355)

Throw 10 darts.  $X = U_{(7)}, Y = U_{(9)}$



We showed  $Y - X \sim \text{Beta}(2, 9)$  ( $k, n-k+1$ )

Notice  $U_{(9)} - U_{(7)} \stackrel{D}{=} U_{(2)} \sim \text{Beta}(2, 10-2+1)$



Todays

① Go over student comments from current test last time.

Sec 5.4

② triangular density  $X, Y$  iid  $U(0, 1)$   
 $S = X + Y$ .

③ Convolution formula of ratio  $Y/X$

(1)

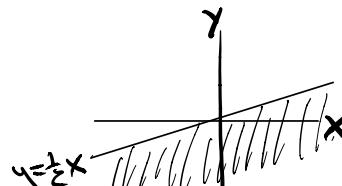
1. Let  $X \sim N(0, 1)$  and  $Y \sim N(0, 1)$  be independent.  $P(X > 2Y)$  equals

- a**  $1 - \Phi(0)$
- b**  $1 - \Phi(\frac{1}{\sqrt{3}})$
- c**  $1 - \Phi(\sqrt{3})$
- d** none of the above

Discuss with neighbour how you did this for 1 minute.

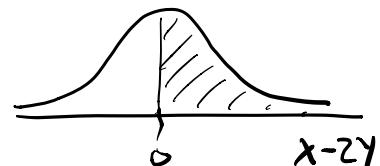
a

Wait, the image is symmetry on the plane!



a

the mean of  $X-2Y$  will still be 0



a

we know the dist of  $X-2Y$  so it's easier to solve that way

a

Adding together the normals,  $P(X-2Y > 0)$

$$X-2Y \sim N(0, 5)$$

$$1-\Phi(0/5) = 1-\Phi(0)$$

$$\begin{aligned} P(X-2Y > 0) &= P\left(\frac{X-2Y}{\sqrt{5}} > \frac{0}{\sqrt{5}}\right) = P(Z > 0) \\ &\approx 1 - \Phi(0) \end{aligned}$$

② Sec 5.4 Triangular density ← convolution  
of sum of 2 iid  $U(0,1)$

Let  $X, Y \sim U(0,1)$

$S = X + Y$  — takes values between 0 and 2

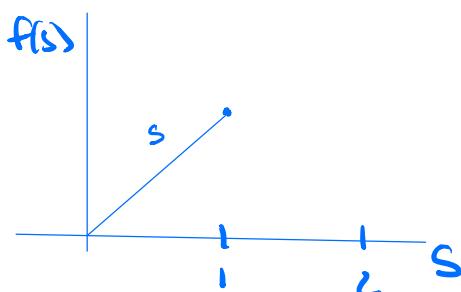
recall, if  $X > 0, Y > 0$

$$f_S(s) = \int_{x=0}^{x=s} f(x, s-x) dx = \int_{x=0}^{x=s} f(x) f(s-x) dx$$

if  $X, Y$  independent

a) for  $0 < s < 1$  find  $f_S(s)$ .

$$f_S(s) = \int_{x=0}^{x=s} 1 \cdot 1 dx = s$$



if  $1 < s < 2$ ,

$$\text{let } T = 1-X + 1-Y = 2-S$$

$$1-X \stackrel{D}{=} X, \quad 1-Y \stackrel{D}{=} Y$$

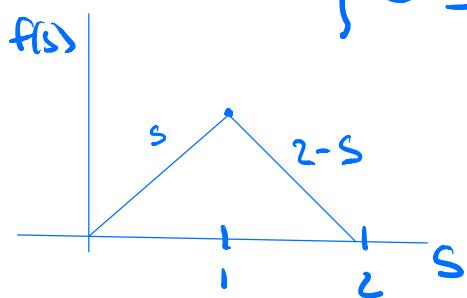
$\Rightarrow T$  is sum of 2 iid  $U(0,1)$  and  $0 < T < 1$   
since  $1 < s < 2$ .

$\Rightarrow f_T(t) = t$  by part 1.

b) For  $1 < s < 2$  find  $f_s(s)$

$$S = 2-T$$
$$f_s(s) = \frac{1}{t-1} \cdot f_T(t) \Big|_{t=2-s} = t \Big|_{t=2-s} = 2-s$$

$$\Rightarrow f(s) = \begin{cases} s & 0 < s < 1 \\ 2-s & 1 < s < 2 \end{cases}$$



(3)

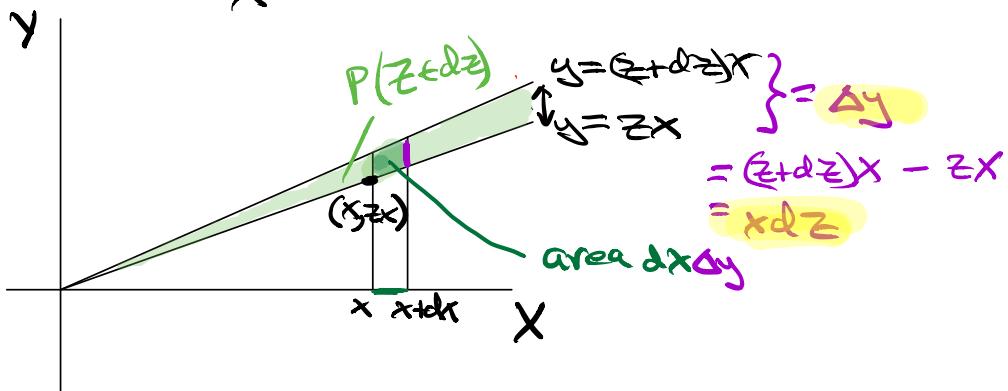
Convolution formula for density of ratio  $y/x$

$$X > 0, Y > 0$$

$$\text{let } z = \frac{y}{x}.$$

Find  $f_z(z)$ .

Picture  $z = \frac{y}{x} \Rightarrow y = zx$  slope.



$$P(z \in dz) = \int_{x=0}^{x=\infty} P(z \in dz, x \in dx)$$

$$f(z)dz$$

$$= \int_{x=0}^{x=\infty} f(x, zx) x dz dx$$

$$\Rightarrow f(z) = \int_{x=0}^{x=\infty} \int_{x=0}^{x=\infty} f(x, zx) x dx = \text{if } x, y \text{ indep.}$$

$$= \int_{x=0}^{x=\infty} f(x) f_y(zx) x dx$$

Convolution formula.

Ex

Recall gamma density

$$X \sim \text{Gamma}(r, \lambda)$$

$$f_X(x) = \frac{\lambda^r}{\Gamma(r)} x^{r-1} e^{-\lambda x}$$

so  $\int_0^\infty x^{r-1} e^{-\lambda x} = \frac{\Gamma(r)}{\lambda^r}$

Let  $X, Y \stackrel{iid}{\sim} \text{Exp}(1)$ .  $Z = \frac{Y}{X}$   
Find  $f_Z(z)$ .

Soln

$$f_Z(z) = \int_{x=0}^{\infty} f_X(x) f_Y(zx) x dx$$

$$= \int_{x=0}^{\infty} e^{-x} e^{-zx} x dx$$

$$= \int_{x=0}^{\infty} x e^{-(1+z)x} dx$$

gamma( $r=z, \lambda=1+z$ )

$$= \frac{\Gamma(z)}{(1+z)^z} = \boxed{\frac{1}{(1+z)^z} \text{ for } z < \infty.}$$

tinyurl

<http://tinyurl.com/april15-pt1>

<http://tinyurl.com/april15-pt2>

## Stat 134

Monday April 15 2019

- Let  $X \sim U(0, 1)$  and  $Y \sim U(0, 1)$  be independent. The density of  $Z = Y/X$  for  $0 < z < 1$  is:

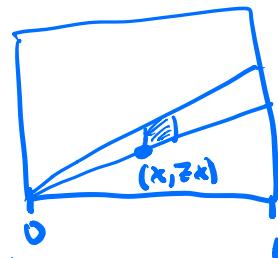
a  $1/(2z)$

b  $1/2$

c  $1/(2z^2)$

d none of the above

$\alpha z < 1$



$z = \frac{y}{x}$

$$f_Z(z) = \int_{x=0}^{x=1} f_X(x) f_Y(zx) x dx = \frac{1}{z} \int_0^1 \frac{1}{x} dx = \frac{1}{z}$$

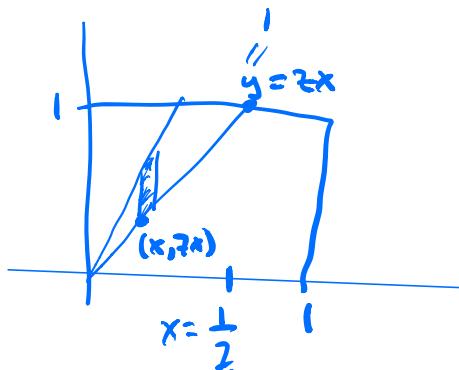
2. Let  $X \sim U(0, 1)$  and  $Y \sim U(0, 1)$  be independent. The density of  $Z = Y/X$  for  $z > 1$  is:

a  $1/(2z)$

b  $1/2$

c  $1/(2z^2)$

d none of the above

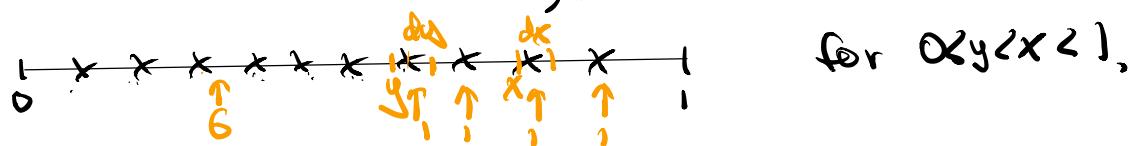


$$f_Z(z) = \int_{x=0}^{x=\frac{1}{z}} f_X(x) f_Y(zx) x dx$$

$$= \int_0^{\frac{1}{z}} 1 \cdot 1 x dx = \frac{1}{z} \left[ x \right]_0^{\frac{1}{z}} = \boxed{\frac{1}{z^2}}$$

Ex Let  $Y \sim U_{(0,1)}, X \sim U_{(0,1)}$  for 10 iid  $U(0,1)$ .

The joint density  $f_{X,Y}(x,y) = \binom{10}{6,1,1,1} y^6 (x-y)(1-x)$



for  $0 < y < x < 1$ ,

Let  $Z = \frac{Y}{X}$  — values between 0 and 1.

Prove via the convolution formula that the density of  $Z = \frac{Y}{X}$  is a Beta distribution and find the parameters.

$$f_Z(z) = \int_{x=0}^{x=1} f_{X,Y}(x, zx) x dx$$

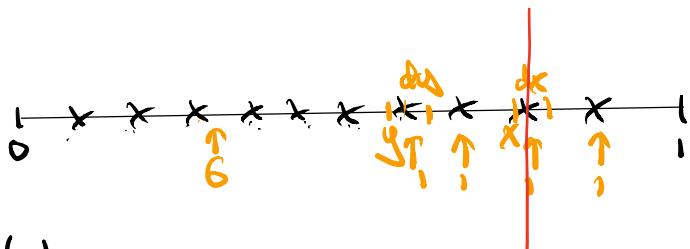
$$C = \binom{10}{6,1,1,1}$$

$$\begin{aligned} f(z) &= C \int_{x=0}^{x=1} (zx)^6 (x-zx)(1-x) dx \\ &= C (1-z)^6 \int_0^1 x^7 (1-x) dx \\ &= C \underbrace{\int_0^1 x^7 (1-x) dx}_{\text{constant}} \cdot \underbrace{[z^6 (1-z)]}_{\Rightarrow Z \sim \text{Beta}(7,2)} \end{aligned}$$

$Z \stackrel{D}{=} U_{(7)}$  when throw 8 darts.

chk  $U_{(7)} \sim \text{Beta}(7, 8-7+1) = \text{Beta}(7,2)$

interpretation:



$y \sim \frac{U_{(7)} \text{ of } 10}{U_{(9)} \text{ of } 10}$  Make the location of the  
 $x = U_{(9)} + 10$  9<sup>th</sup> dant 1 (see red line). There  
are 8 dants before it.  
The location of the 7<sup>th</sup> of the 8 dants  
is  $U_{(7)} \text{ of } 8$ .

we have  $\frac{U_{(7)} \text{ of } 10}{U_{(9)} \text{ of } 10} \stackrel{D}{=} U_{(7)} \text{ of } 8$

You can check that they are both  
 $\text{Beta}(7, 2)$  since we know

$$U_{(k)} \text{ of } n \sim \text{Beta}(k, n-k+1) \quad \text{so}$$

$$U_{(7)} \text{ of } 8 \sim \text{Beta}\left(7, \frac{8-7+1}{2}\right)$$

Ex You have 10 darts  
What distribution is

$$\frac{U_{(4)} \text{ of } 10}{U_{(8)} \text{ of } 10} ?$$

$$\boxed{\frac{U_{(4)} \text{ of } 10}{U_{(8)} \text{ of } 10} \stackrel{D}{=} U_{(4)} \text{ of } 7}$$

More generally  
If  $U_1, \dots, U_n \stackrel{iid}{\sim} U(0, 1)$ .

Then for  $k < m$

$$\boxed{\frac{U_{(k)} \text{ of } n}{U_{(m)} \text{ of } n} \stackrel{D}{=} U_{(k)} \text{ of } m-1}$$