

WED, JAN 29

MORE ON BAYES' RULE:

$$P(B_i | A) = \frac{P(A | B_i) P(B_i)}{\sum_{i=1}^n P(A | B_i) P(B_i)}$$

(P.49)

IMPORTANT DISCUSSION, BOTTOM OF P.48:

1st STAGE: ONE OF B_1, \dots, B_n OCCURS (UNKNOWN)

2nd STAGE: RESULT A IS OBSERVED.

$P(B_i | A) \leftarrow$ POSTERIOR PROBS.

$P(B_i) \leftarrow$ PRIOR PROBS.

$P(A | B_i) \leftarrow$ LIKELIHOODS.

OFTEN
EASIER TO
CALCULATE.

§1.6 SEQUENCES OF EVENTS.

* OFTEN RANDOM EXPERIMENTS/
PROCESSES HAVE MULTIPLE STEPS / STAGES.

GROUPS LET A, B, C BE EVENTS.

SHOW THAT:

$$P(A \cap B \cap C) = P(A)P(B|A)P(C|A \cap B)$$

ANS

$$\begin{aligned}P(A \cap B \cap C) &= P(A \cap B) P(A \cap B \cap C | A \cap B) \\&= P(A \cap B) P(C | A \cap B) \\&= P(A) P(B | A) P(C | A \cap B).\end{aligned}$$

IN GENERAL:

$$P\left(\bigcap_{i=1}^n A_i\right) = P(A_1)P(A_2|A_1)P(A_3|A_1 \cap A_2) \\ \dots P(A_n | \bigcap_{i=1}^{n-1} A_i).$$

EG GEOMETRIC (p) DISTRIBUTION.

COIN: $P(T) = p$, $P(H) = 1-p$.

$X = \#$ OF FLIPS UNTIL A T.

WRITE T_i AND H_i FOR EVENT THAT
THE FLIP IS A T/H.

$$\begin{aligned} P(X=k) &= P(H_1 H_2 \dots H_{k-1} T_k) \\ &= P(H_1) P(H_2) \dots P(H_{k-1}) P(T_k) \\ &= (1-p)^{k-1} p \end{aligned}$$

*FLIPS
INDEP.

($k \geq 1$).

INDEPENDENCE, IN GENERAL :

A, B, C ARE INDEPENDENT IF

$$\textcircled{1} P(B|A) = P(B|A^c) = P(B)$$

$$\textcircled{2} P(C|AB) = P(C|A^cB) = P(C|AB^c) \\ = P(C|A^cB^c) = P(C).$$

$$\Rightarrow P(ABC) = P(A)P(B)P(C).$$

* THIS GENERALIZES TO $n \geq 2$ EVENTS
 A_1, \dots, A_n

→ SEE DISCUSSION ON p. 67.

PAIRWISE INDEPENDENCE:

A_1, \dots, A_n SUCH THAT A_i, A_j

ARE INDEPENDENT (i.e. $P(A_i A_j) = P(A_i)P(A_j)$)

FOR ALL $1 \leq i \neq j \leq n$.

GROUPS FIND AN EXAMPLE THAT

SHOWS THAT

PW INDEP. \neq INDEP.

ANS FLIP A FAIR COIN TWICE

S = EVENT THAT BOTH FLIPS
ARE THE SAME.

① CLEARLY S , 1st FLIP, 2nd FLIP ARE
NOT INDEPENDENT.

② BUT

$$P(H_1 \cap S) = P(H_1 \cap T_2) = \frac{1}{4} = P(H_1)P(S)$$

so S , 1st FLIP ARE INDEP.

(SEE P. 70 FOR ALL DETAILS)

EG (GAMBLER'S RUIN)

SUPPOSE YOU PLAY n INDEPENDENT GAMES. IN EACH YOU HAVE A $1/N$ CHANCE OF WINNING.

HOW MANY TIMES n MUST YOU PLAY TO HAVE $\geq 50\%$ CHANCE OF WINNING AT LEAST ONCE?

(AN OLD GAMBLERS' RULE OF THUMB SAYS $n \approx \frac{2}{3}N$.)

ANS (p.60)

$$\begin{aligned} P(\text{WIN} \geq \text{ONCE}) &= 1 - P(\text{NO WIN}) \\ &= 1 - \left(1 - \frac{1}{N}\right)^n \stackrel{(\text{SET})}{=} \frac{1}{2}. \end{aligned}$$

$$\Rightarrow \frac{1}{2} = \left(1 - \frac{1}{N}\right)^n$$

$$\Rightarrow \log\left(\frac{1}{2}\right) = n \log\left(1 - \frac{1}{N}\right)$$

$$\Rightarrow n = \log\left(\frac{1}{2}\right) / \log\left(1 - \frac{1}{N}\right)$$

$$\Rightarrow n \approx N \log(2) \left[\log(1+z) \approx z, z \rightarrow 0 \right]$$

$$\approx 0.69 N \approx \frac{2}{3} N.$$

GROUPS DECK OF CARDS IS

WELL-SHUFFLED, AND 5 CARDS

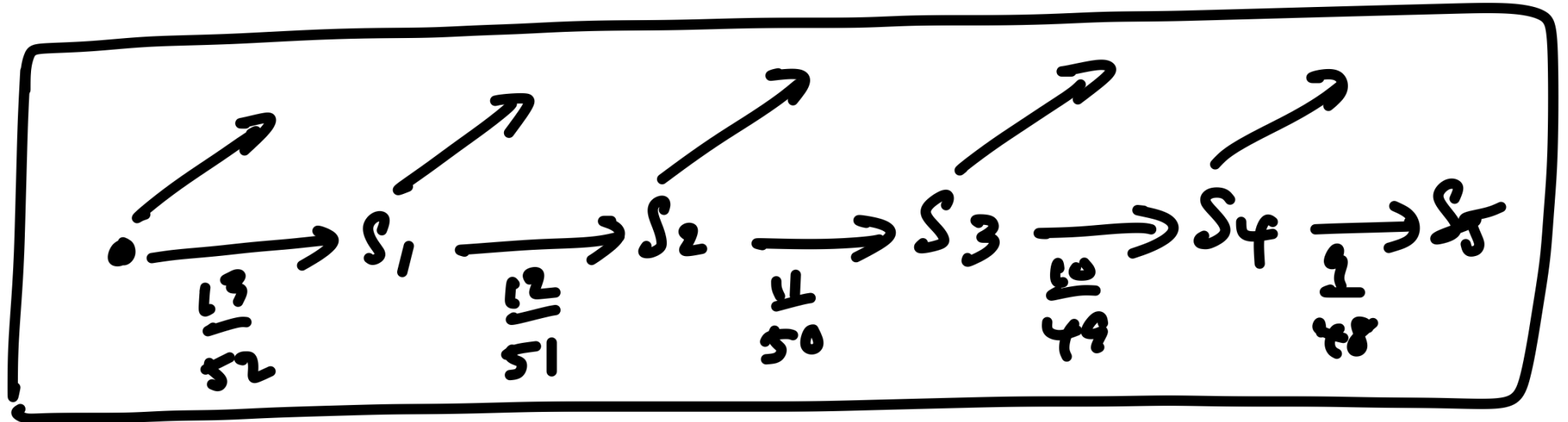
ARE DEALT.

$P(\text{FLUSH})$

$= P(\text{ALL 5 OF SAME SUIT}) = ?$

ANS LET S_i, H_i, D_i, C_i BE EVENT
THAT i TH CARD IS A S, H, D, C.

TREE
DIAGRAM



$$P(\text{SPADE FLUSH}) = \frac{13 \cdot 12 \cdot 11 \cdot 10 \cdot 9}{52 \cdot 51 \cdot 50 \cdot 49 \cdot 48} = \frac{(13)5}{(52)5}$$

$$P(\text{FLUSH}) = 4 \cdot P(\text{SPADE FLUSH})$$

$$= 4 \cdot \frac{(13)5}{(52)5} \approx 0.198\%.$$