

stat 134 lec 39

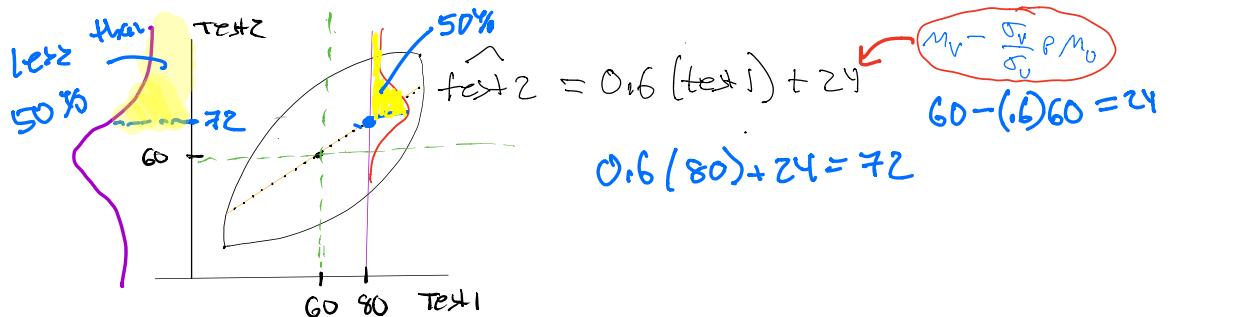
Warmup 8:00 - 8:10 AM

$$(\text{Test 1}, \text{Test 2}) \sim \text{BV}(60, 60, 20^2, 20^2, 0.6)$$

What is greater? \nwarrow mean \nwarrow variance

a) The chance you get greater than 72 on test 2 among students who get 80 on test 1 $\rightarrow 50\%$

b) The chance you get greater than 72 on test 2 among all students,



This is why you sit next to the successful kids in class. You have a higher chance of doing well on the next test!

Last time Sec 6.5 Bivariate normal distribution.

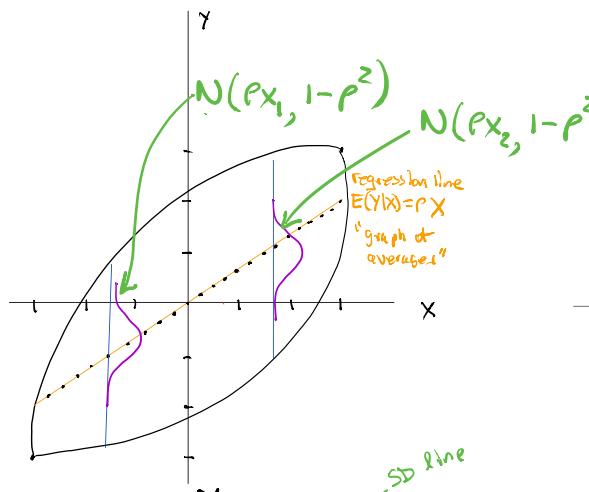
std bivariate normal construction

$$X, Z \sim \text{IID } N(0, 1), -1 \leq P \leq 1$$

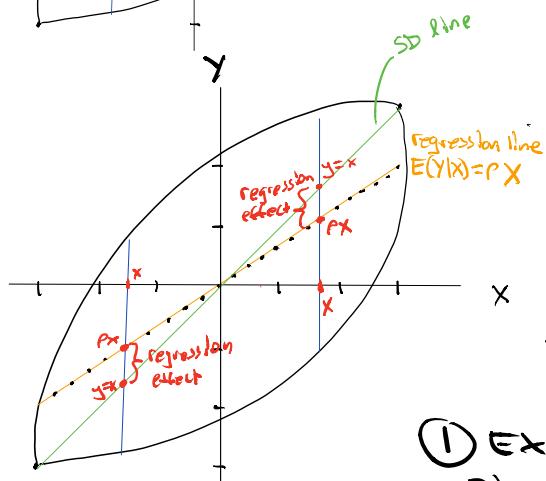
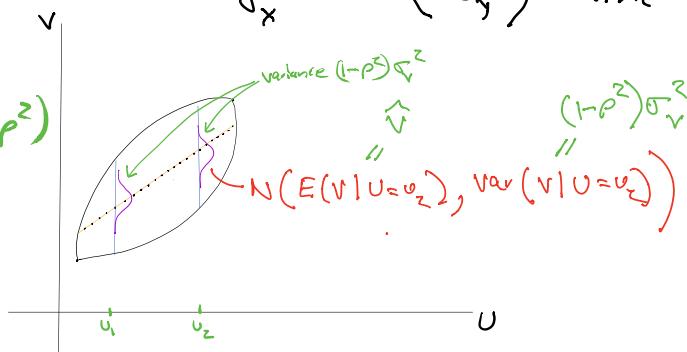
$$Y = PX + \sqrt{1-P^2}Z \sim N(0, 1)$$

(X, Y) is std bivariate normal with $\text{cov}(X, Y) = P$.

$$Y|X \sim N(PX, 1-P^2)$$



$$\frac{\hat{v} - \mu_x}{\sigma_x} = P \left(\frac{U - \mu_x}{\sigma_y} \right) \text{ regression line}$$



Picture of regression effect.

Today

- ① Examples with Bivariate normal (Sec 6.5)
- ③ MGF of Bivariate normal
- ③ Properties of Bivariate normal
- ④ Review of Conditional Expectation (Sec 6.4)

① Examples with bivariate normal,

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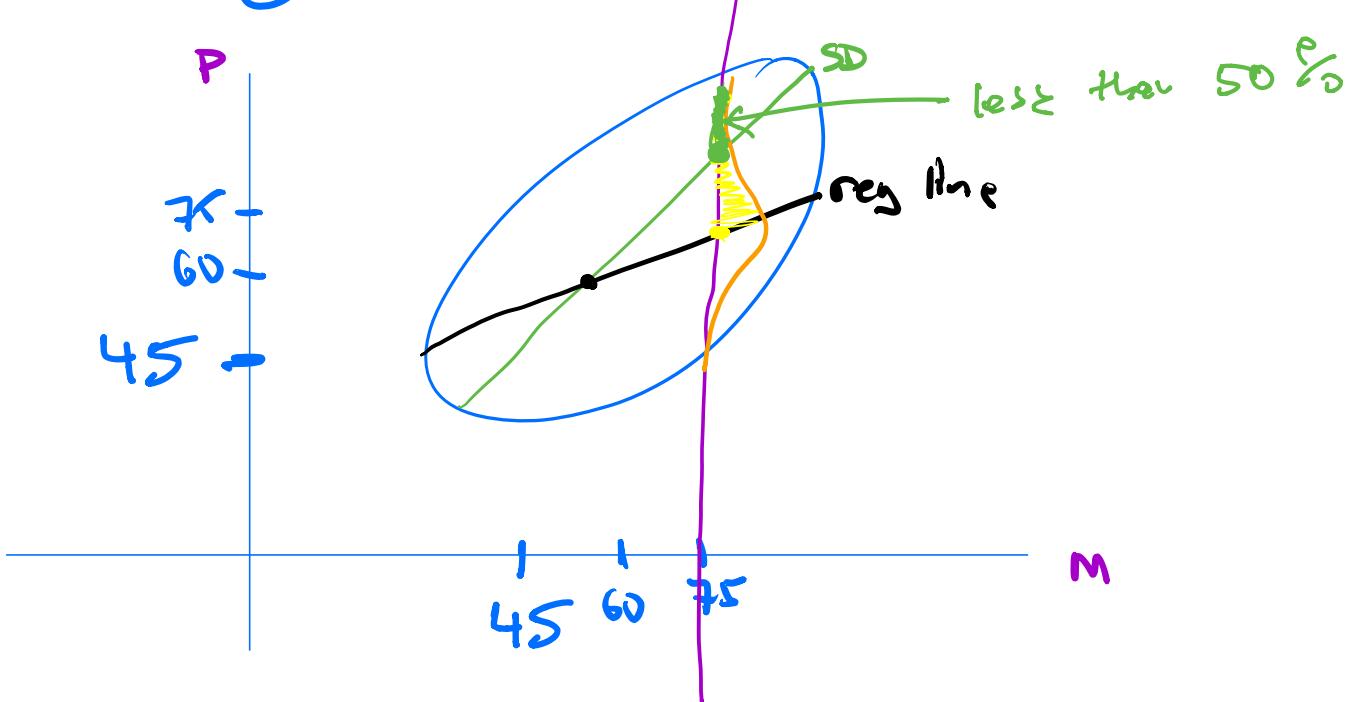
<http://tinyurl.com/dec6-pt1>

Stat 134

Monday April 29 2019

1. A test score in Math and Physics is bivariate normal, $\rho > 0$. The average is 60 on both tests and the SDs are the same. Of students scoring 75 on the Math test:

- a about half scored over 75 on Physics
- b more than half scored over 75 on Physics
- c less than half scored over 75 on Physics



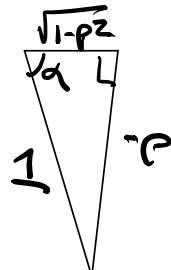
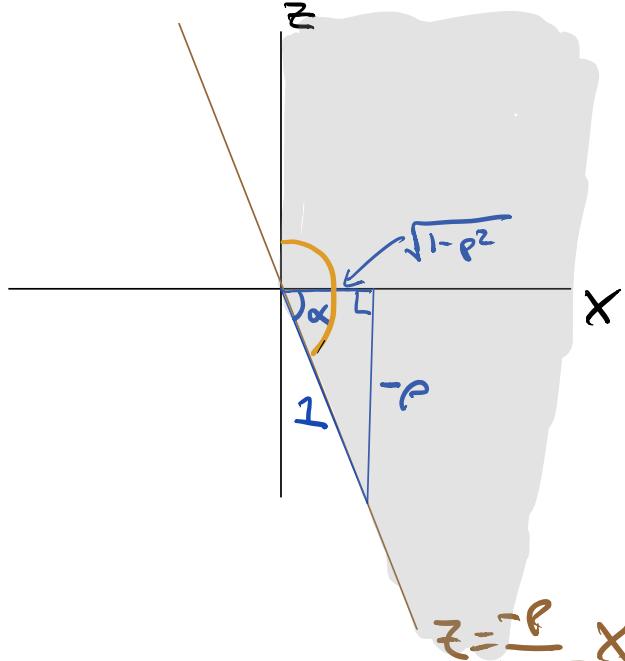
$\Leftrightarrow X, Y$ std bivariate normal, $\rho > 0$

Find $P(X > 0, Y > 0)$

$$P(X > 0, Y > 0) = P(X > 0, \rho X + \sqrt{1-\rho^2} Z > 0)$$

$$= P(X > 0, Z > -\frac{\rho}{\sqrt{1-\rho^2}} X) \quad \rightarrow \text{Note } X \text{ and } Z \text{ are uncorrelated}$$

so joint (X, Z) is a symmetric bell over X, Z Plane,



$$Z = \frac{-\rho}{\sqrt{1-\rho^2}} X$$

$$\tan \alpha = \frac{-\rho}{\sqrt{1-\rho^2}}$$

$$\alpha = \tan^{-1} \left(\frac{-\rho}{\sqrt{1-\rho^2}} \right)$$

$$P(X > 0, Z > -\frac{\rho}{\sqrt{1-\rho^2}} X) = \frac{90 + |\alpha|}{360} = \boxed{\frac{90 + \tan^{-1} \left(\frac{-\rho}{\sqrt{1-\rho^2}} \right)}{360}}$$

② MGF of Bivariate normal

The single and multi-variate MGF is defined as :

$$M_y(t) = E(e^{tY})$$

$$M_{(x,y)}(s,t) = E(e^{sx+ty}) \quad \text{multi-variate MGF}$$

Note that

$$M_y(t) = E(e^{tY}) = M_{tY}(1)$$

$$M_{(x,y)}(s,t) = E(e^{sx+ty}) = M_{sx+ty}(1)$$

recall,

$$Z \sim N(\mu, \sigma^2) \text{ iff } M_Z(\omega) = e^{\mu\omega} e^{\sigma^2 \frac{\omega^2}{2}}$$

thus $M_{(X,Y)}(s,t) = M_X(s) M_Y(t)$ iff
 X, Y are independent,

PFY

$$M_{(X,Y)}(s,t) = E(e^{sX+tY}) \quad \text{mu H: variable = MGF}$$

$$= E(e^{sX} \cdot e^{tY})$$

if X, Y
indep.

$$= E(e^{sX}) E(e^{tY}) = M_X(s) M_Y(t)$$

□

Thm Let (X, Y) be standard bivariate normal.

The MGF of (X, Y) is

$$M_{(X,Y)}(s,t) = e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}$$

Pf/ Recall that $X \sim N(0, 1)$ so $M_X(s) = e^{\frac{s^2}{2}}$

$$\begin{aligned} M_{X,Y}(s,t) &= E[e^{sx+ty}] \\ &= E[e^{sx+t(\rho X + \sqrt{1-\rho^2}Z)}] \\ &= E\left[e^{(s+t\rho)X} \cdot e^{t\sqrt{1-\rho^2}Z}\right] \\ &\stackrel{\text{independence}}{=} E\left[e^{(s+t\rho)X}\right] E\left[e^{t\sqrt{1-\rho^2}Z}\right] \\ &= M_X(s+t\rho) \cdot M_Z(t\sqrt{1-\rho^2}) \end{aligned}$$

Finish proof.

$$\begin{aligned} &= e^{\frac{(s+t\rho)^2}{2}} \cdot e^{\frac{(t)\sqrt{1-\rho^2}}{2}} \\ &= e^{\frac{s^2}{2} + st\rho + \frac{t^2\rho^2}{2}} \cdot e^{\frac{t^2}{2} - \frac{t^2\rho^2}{2}} \\ &= \boxed{e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}} \quad \square \end{aligned}$$

③ Properties of Standard Normal

Recall that if 2 RVs X, Y are independent, $\text{Cov}(X, Y) = 0$
 $\Rightarrow \text{Corr}(X, Y) = 0$

However the converse is not true
 in general. ($\text{Corr}(X, Y) = 0 \not\Rightarrow X, Y \text{ indep.}$)

Let $X \sim N(0, 1)$

$$Y = X^2$$

X and Y are dependent

we show X, Y are uncorrelated.

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$= E(X^3) - E(X)E(X^2)$$

$$= E(X^3) - 0$$

Find $E(X^3)$

$$X \sim N(0, 1) \quad Z \sim N(0, 1)$$

$$M_X(t) = e^{t^2/2}$$

$$M_X'''(t) = t^2 e^{t^2/2} + 2t e^{t^2/2} + t^3 e^{t^2/2} \Big|_{t=0} = 0$$

$$\Rightarrow E(X^3) = 0$$

$$\Rightarrow \text{Corr}(X, Y) = 0$$

Then If (X, Y) is ^{std} bivariate normal then

$\rho = \text{Corr}(X, Y) = 0$ iff X, Y are independent.

Pf

$$M_{(X,Y)}^{(s,t)} = M_{sX+tY}^{(1)} = M_{sX}^{(1)} \cdot M_{tY}^{(1)} = M_X^{(s)} M_Y^{(t)}$$

iff X and Y are independent.

Since (X, Y) is ^{std} bivariate normal,

$$M_{(X,Y)}^{(s,t)} = e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho}$$

Finish proof.

$$= M_X^{(s)} M_Y^{(t)} \text{ iff } \rho = 0$$

$$\text{since } M_X^{(s)} = e^{\frac{s^2}{2}}$$

$$M_Y^{(t)} = e^{\frac{t^2}{2}}$$

□

□

Recall from Lec 30 that the sum of independent normal random variables is normal.

What about dependent normal random variables?

Then Let $X, Y \sim N(0, 1)$ and $\text{Corr}(X, Y) = \rho$.

(X, Y) is std bivariate normal iff

$sX + tY$ is normal for all constants s, t .

Pf/

\Rightarrow : Suppose X, Y std bivariate normal.
 (i.e. $Y = \rho X + \sqrt{1-\rho^2} Z$ for $X, Z \sim N(0, 1)$)

$$\begin{aligned}sX + tY &= sX + t(\rho X + \sqrt{1-\rho^2} Z) \\ &= (s+t\rho)X + t\sqrt{1-\rho^2} Z\end{aligned}$$

is normal since X, Z are independent normals,

\Leftarrow : Suppose $sX + tY$ is normal,

note $E(sX + tY) = sE(X) + tE(Y) = 0$

$$\begin{aligned}\text{Var}(sX + tY) &= \text{Var}(sX) + \text{Var}(tY) + \text{Cov}(sX, tY) \\ &= s^2 + t^2 + 2st\rho\end{aligned}$$

hence, $sX + tY \sim N(0, s^2 + t^2 + 2st\rho)$

"
 $zst\text{Cov}(X, Y)$
 " "
 $\text{Corr}(X, Y)$

Recall,

$$\text{For } Z \sim N(\mu, \sigma^2) \Rightarrow M_Z^{(n)} = e^{\mu n} e^{\frac{\sigma^2 n^2}{2}}$$

$$\begin{aligned}\text{Then } M_{(X, Y)}^{(s, t)} &= M_{sX+tY}^{(1)} = e^{(s^2 + t^2 + 2st\rho)\frac{1}{2}} \\ &\quad \text{normal} \\ &= e^{\frac{s^2}{2} + \frac{t^2}{2} + st\rho} \quad \text{M6F of } (X, Y)\end{aligned}$$

$\Rightarrow (X, Y)$ is standard bivariate normal \square

we dont need to restrict
ourselves to $X, Y \sim N(0, 1)$

Corollary → proof at end of lecture

Let $U \sim N(\mu_u, \sigma_u^2)$

$V \sim N(\mu_v, \sigma_v^2)$

and $\text{Corr}(U, V) = \rho$

$(U, V) \sim BV(\mu_u, \mu_v, \sigma_u^2, \sigma_v^2, \rho)$ iff

$sU + tV$ is normal for all constants s, t .

Ex

Let M be a student's score on the midterms of a class and F the student's score on the final of the same class.

Suppose $(M, F) \sim \text{BV}(70, 65, 8^2, 10^2, 0.6)$.

$$\begin{matrix} M & F \\ \sigma_M^2 & \sigma_F^2 \\ \rho & \end{matrix}$$

Find the chance that the student scores higher on the final than the midterms.

Hint $P(F > M) = P(F - M > 0)$

What distribution is $F - M$? $\xrightarrow{\text{Normal}}$
 (since $(M, F) \xrightarrow{\text{Bivariate normal}}$)

$$E(F - M) = 65 - 70 = -5$$

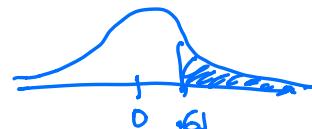
$$\begin{aligned} \text{Var}(F - M) &= \text{Var}(F) + \text{Var}(M) - 2\text{Cov}(F, M) \\ &\quad " \\ &\quad \text{Corr}(F, M) \text{SD}(M) \text{SD}(F) \end{aligned}$$

$$= 10^2 + 8^2 - 2(0.6)(8)(10) = 68$$

$$\text{SD}(F - M) = \sqrt{68}$$

$$\Rightarrow F - M \sim N(-5, 68)$$

$$\Rightarrow \frac{0 - E(F - M)}{\text{SD}(F - M)} = \frac{0 - (-5)}{\sqrt{68}} = \boxed{.61}$$



$$P(F > M) = P(F - M > 0) = 1 - \Phi(.61) = \boxed{.27}$$

Appendix

Corollary

Let $U \sim N(\mu_U, \sigma_U^2)$
 $V \sim N(\mu_V, \sigma_V^2)$

and $\text{Corr}(U, V) = \rho$

$(U, V) \sim BV(\mu_U, \mu_V, \sigma_U^2, \sigma_V^2, \rho)$ iff
 $sU + tV$ is normal for all constants s, t .

Pf/ let $X = \frac{U - \mu_U}{\sigma_U}$ ($U = \mu_U + X\sigma_U$)

$Y = \frac{V - \mu_V}{\sigma_V}$ ($V = \mu_V + Y\sigma_V$)

$(U, V) \sim BV(\mu_U, \mu_V, \sigma_U^2, \sigma_V^2, \rho)$

iff $(X, Y) \sim BV(0, 0, 1, 1, \rho)$

$$\begin{aligned} sU + tV &= s(\mu_U + X\sigma_U) + t(\mu_V + Y\sigma_V) \\ &= \underset{a}{s\sigma_U}X + \underset{b}{t\sigma_V}Y + \underset{\text{constant}}{s\mu_U + t\mu_V} \end{aligned}$$

this is normal for all constants s, t

if $ax + by$ is normal for all constants a, b ,

□

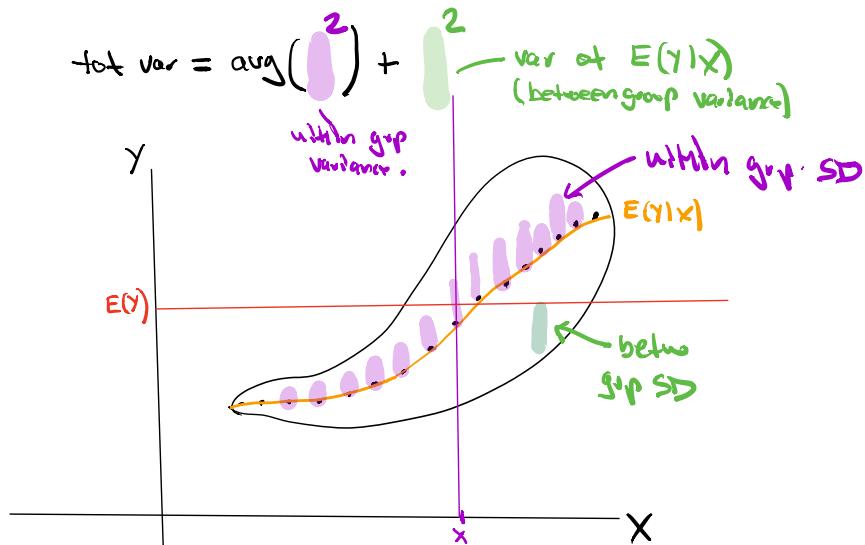
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Review of conditional expectation sec 6.4

Properties (from lec 34)

- (1) $E(Y) = E(E(Y|X))$ iterated expectations
- (2) $E(aY+b|X) = aE(Y|X) + b$
- (3) $E(Y+z|X) = E(Y|X) + E(z|X)$
- (4) $E(g(X)|X) = g(X)$
- (5) $E(g(X)Y|X) = g(X)E(Y|X)$
- (6) $\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X))$ total variance decomposition

$$\text{Var}(Y) = E(\text{Var}(Y|X)) + \text{Var}(E(Y|X)) \quad (\text{see 6.2.18})$$



Ex Let $U \sim \text{Unif}(0,1)$

Given $U=u$, X is exponential with mean u

i.e $X|U=u \sim \text{Exp}\left(\frac{1}{u}\right)$

$$E(X|U) = u$$

$$\text{Var}(X|U) = u^2$$

Find $E(X)$, $E(ux)$ and $\text{Var}(x)$.

$$E(X) = E(E(X|U)) = E(u) = \frac{1}{2}$$

$$E(ux) = E(E(ux|U)) = E(uE(x|U))$$

$$= E(u^2) = \text{Var}(u) + E(u)^2 = \frac{1}{4}$$

$$\text{Var}(x) = E(\underbrace{\text{Var}(x|U)}_{\frac{1}{12}}) + \underbrace{\text{Var}(E(x|U))}_{0} = \frac{5}{12}$$

