

Stat 134 Lec 40

Plan

Friday 4/27 Finish 6.5

Practice final on
website

Monday 4/28 Adam has extended OH
10-12 SLC

Class time review.

Review 6.5 in section
(have 6.5 suggested problems)
for final

Wed, Fri TBA.

Sec 6.5 Bivariate normal

Last time:

Let $X, Z \sim N(0, 1)$ and $Y = \rho X + \sqrt{1-\rho^2}Z$, $-\leq \rho \leq 1$

We call the joint distribution X, Y the

Standard bivariate normal w/ correlation ρ

Properties

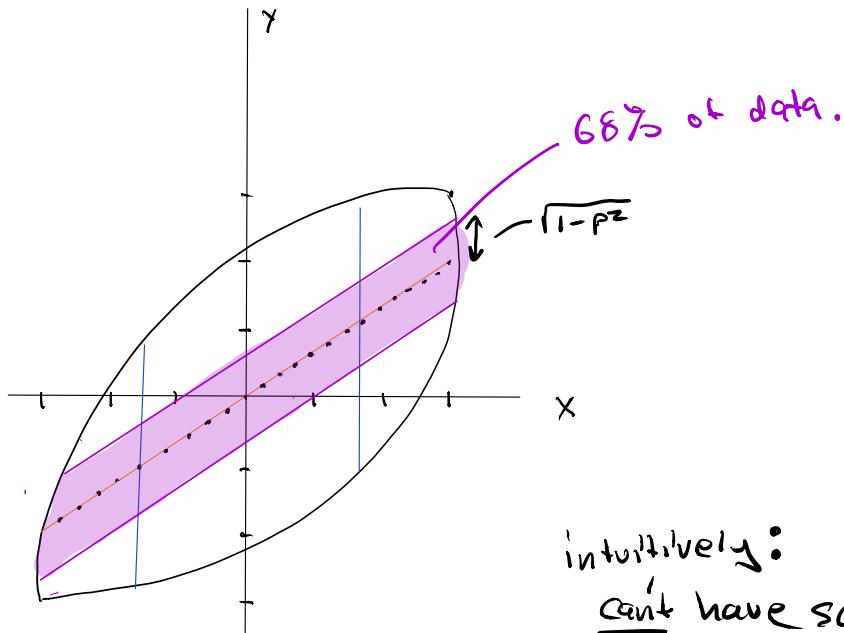
$$1) Y \sim N(0, 1)$$

$$2) \text{Corr}(X, Y) = \rho$$

$$3) Y|X=x \sim N(\rho x, 1-\rho^2)$$

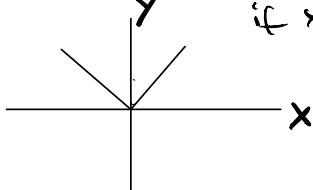
$\text{SD}(Y|X=x)$

So 68% of data is within $\sqrt{1-\rho^2}$ of regression line.



- 4) X, Y indep iff $\rho=0$
but see next page for proof.

intuitively:
can't have scatter diagram
if X, Y bivariate normal,



Theorem

If X, Y are standard bivariate w/ $\text{corr}(x, y) = \rho$

X, Y are independent iff $\rho = 0$

Pf /

$$\text{Recall if } W \sim N(\mu, \sigma^2), f_W(w) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(w-\mu)^2}$$

multiplication rule for density says

$$\begin{aligned} f(x, y) &= f_X(x) f_{Y|X=x}(y|x=x) \\ &= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot \frac{1}{\sqrt{2\pi} \cdot \sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}(y-\rho x)^2} \\ &= \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{x^2}{2} - \frac{1}{2(1-\rho^2)}(y-\rho x)^2} \end{aligned}$$

Algebra

$$\begin{aligned} -\frac{x^2}{2} - \frac{1}{2(1-\rho^2)}(y-\rho x)^2 &= -\frac{x^2(1-\rho^2) - y^2 + 2\rho xy - \rho^2 x^2}{2(1-\rho^2)} \\ &= -\frac{x^2 + x^2\rho^2 - y^2 + 2\rho xy - \rho^2 x^2}{2(1-\rho^2)} \\ &= -\frac{(x^2 - 2\rho xy + y^2)}{2(1-\rho^2)} \end{aligned}$$

$$\text{so } f(x, y) = \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}(x^2 - 2\rho xy + y^2)}$$

is joint density for bivariate normal. $-\frac{1}{2}(x^2 + y^2)$

$$\begin{aligned} \text{If } \rho = 0 \text{ then } f(x, y) &= \frac{1}{2\pi} e^{-\frac{1}{2}(x^2 + y^2)} \\ &= \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} \\ &= f_X(x) f_Y(y). \\ \Rightarrow X, Y \text{ indep.} \end{aligned}$$

Conversely, If X, Y indep, $\rho = 0$ for any RV x, y

□

Theorem Let $X, Y \sim N(0, 1)$

and $\text{Corr}(X, Y) = \rho$.

X, Y are standard bivariate normal correlation ρ

iff $aX + bY$ is normal for every constant a, b .

Proof (sketch)

\Rightarrow : Suppose X, Y std bivariate normal.

(i.e. $Y = \rho X + \sqrt{1-\rho^2} Z$ for $X, Z \sim N(0, 1)$)

$$\begin{aligned} aX + bY &= aX + b(\rho X + \sqrt{1-\rho^2} Z) \\ &= (a + b\rho)X + b\sqrt{1-\rho^2} Z \end{aligned}$$

is normal since X, Z are independent normals,

\Leftarrow : Suppose $aX + bY$ is normal for any a, b .

$$\text{let } Z = \frac{-\rho}{\sqrt{1-\rho^2}} X + \frac{1}{\sqrt{1-\rho^2}} Y$$

Z is normal? ? yes by assumption.

$$E(Z) = 0$$

$$\begin{aligned} \text{Var}(Z) &= \frac{\sigma^2}{1-\rho^2} \text{Var}(X) + \frac{1}{1-\rho^2} \text{Var}(Y) + 2\text{Cov}\left(\frac{-\rho}{\sqrt{1-\rho^2}} X, \frac{1}{\sqrt{1-\rho^2}} Y\right) \\ &= \frac{1-\rho^2}{1-\rho^2} = 1 \quad \frac{\rho^2}{1-\rho^2} \end{aligned}$$

It remains to show that X and Z are independent. This is surprising since X is in the definition of Z ! You can check that $\text{Cov}(X, Z) = 0$ but this doesn't prove independence. The proof is beyond the scope of this class. The proof involves

- (1) use the fact that $aX+bY$ is normal to find the joint density $f_{X,Y}$ (you need moment generating functions for this which you will learn about in Stat 155 or Stat 150),
- (2) Use the bivariate version of the change of variable rule to show that $f_{X,Z}(x,z) = \frac{1}{2\pi} e^{-(x^2+z^2)}$. It follows then that $f_{X,Z}(x,z) = f_X(x)f_Z(z)$ so X and Z are independent.

So finally we have, $X, Z \stackrel{\text{iid}}{\sim} N(0, 1)$ and $Y = \rho X + \sqrt{1-\rho^2} Z \Rightarrow X, Y$ are bivariate normal.



Example of $X, Y \sim N(0, 1)$ whose joint isn't std bivariate normal.

Ex $X \sim N(0, 1)$, $W = \begin{cases} 1 & \text{if prob } \frac{1}{2} \\ -1 & \text{if prob } \frac{1}{2} \end{cases}$
indep RVs.

$$\text{let } Y = wX = \begin{cases} X & \text{if } w=1 \\ -X & \text{if } w=-1 \end{cases}$$

Next we show that $Y \sim N(0, 1)$.

Note by change of variable rule that $-X \sim N(0, 1)$.

$$\begin{aligned} F_Y(y) &= P(Y \leq y) \\ &= P(Y \leq y | W=1)P(W=1) + P(Y \leq y | W=-1)P(W=-1) \\ &= \frac{1}{2}P(X \leq y | W=1) + \frac{1}{2}P(-X \leq y | W=-1) \\ &= \frac{1}{2}P(X \leq y) + \frac{1}{2}P(-X \leq y) \quad \text{since } X, W \text{ indep.} \\ &= \frac{1}{2}\Phi(y) + \frac{1}{2}\Phi(y) \quad \text{since } X, -X \sim N(0, 1) \\ &= \Phi(y). \end{aligned}$$

Thus $Y \sim N(0, 1)$.

$$\text{Then } X+Y = \begin{cases} 2X & \text{with prob } \frac{1}{2} \\ 0 & \text{with prob } \frac{1}{2} \end{cases}$$

is a mixed distribution, not normal.

So X, Y normal but $X+Y$ not normal.

Furthermore $\text{Corr}(X, Y) = 0$ since

$$E(YX) = E(wX^2) = E(w)E(X^2) \quad \text{by indep of } W \text{ and } X.$$

but X, Y are clearly dependent contradicting property of bivariate normal.

bivariate normal X, Y with correlation ρ

Def'n

RVs X, Y are said to have the bivariate normal distribution with parameters $(\mu_x, \mu_y, \sigma_x^2, \sigma_y^2, \rho)$ if

$$X^* = \frac{X - \mu_x}{\sigma_x}, \quad Y^* = \frac{Y - \mu_y}{\sigma_y} \quad \text{have}$$

the std bivariate normal dist w/
correlation ρ .

Note

From Tel 38 we saw $\text{corr}(2x+3, 4y-5) = \text{corr}(x, y)$

$$\begin{aligned}\rho &= \text{corr}(X^*, Y^*) = \text{corr}\left(\frac{X - \mu_x}{\sigma_x}, \frac{Y - \mu_y}{\sigma_y}\right) \\ &= \text{corr}\left(\frac{1}{\sigma_x}X - \frac{\mu_x}{\sigma_x}, \frac{1}{\sigma_y}Y - \frac{\mu_y}{\sigma_y}\right) \\ &= \text{corr}(X, Y)\end{aligned}$$

Thm

X, Y is bivariate normal w/ corr ρ

iff $ax + by$ is normal for any $a, b \in \mathbb{R}$,

Thm

X, Y is bivariate normal w/ corr ρ

iff there exists $Z_1, Z_2 \stackrel{iid}{\sim} N(0, 1)$

such that

$$X = c_1 Z_1 + c_2 Z_2$$

$$Y = c_3 Z_1 + c_4 Z_2$$

for some constants c_1, c_2, c_3, c_4 .

Proof/

$$\text{let } Z_1 = X^*$$

$$Z_2 = \frac{-\rho}{\sqrt{1-\rho^2}} X^* + \frac{1}{\sqrt{1-\rho^2}} Y^*$$

Check $Z_1, Z_2 \stackrel{iid}{\sim} N(0, 1)$

$$X = \sigma_X Z_1 + \mu_X$$

$$Y = \sigma_Y (\rho Z_1 + \sqrt{1-\rho^2} Z_2) + \mu_Y$$

Stat 134

Friday April 27 2018

1. Let X, Y be bivariate normal. Then $2X+3Y+4$, $6X-Y-4$ is bivariate normal.

a true

b false

$$a(2x+3y+4) + b(6x-y-4)$$

is a linear combo of
 x, y hence normal.

Note

we see from this that

if x, y are bivariate normal then

$$c_1x + c_2y + c_3, c_4x + c_5y + c_6$$

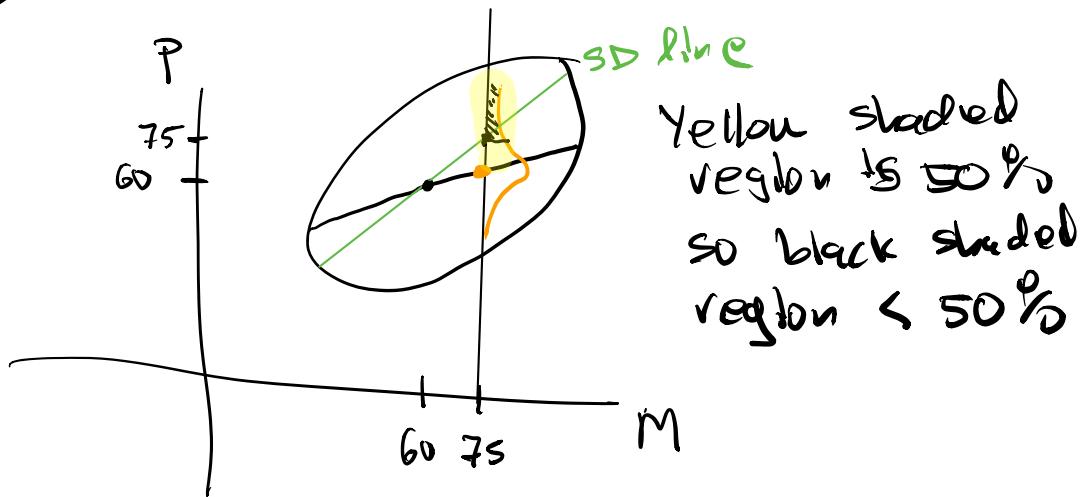
are bivariate normal, in

other words linear combinations
of bivariate normal is bivariate
normal.

2. A test score in Math and Physics is bivariate normal, $\rho > 0$. The average is 60 on both tests and the SDs are the same. Of students scoring 75 on the Math test:

- a** about half scored over 75 on Physics
- b** more than half scored over 75 on Physics

- c** less than half scored over 75 on Physics

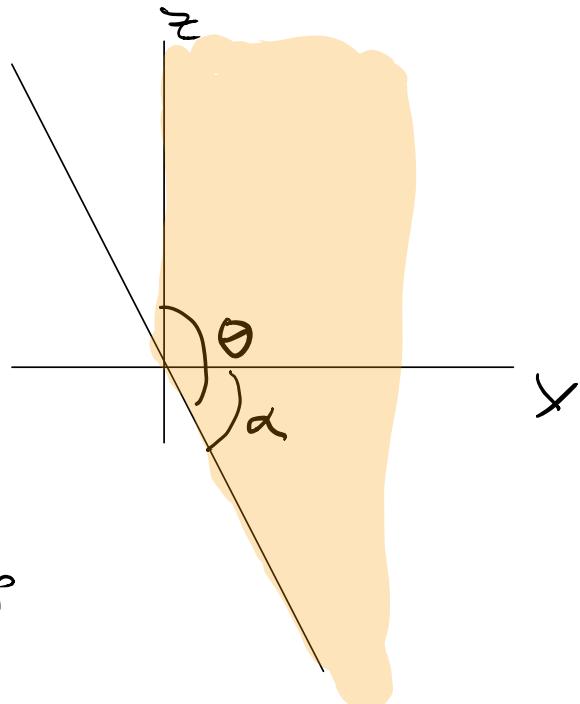


$\Leftrightarrow X, Y$ sto bivariate normal, $\rho > 0$

$$\begin{aligned} P(X > 0, Y > 0) &= P(X > 0, \rho X + \sqrt{1-\rho^2} Z > 0) \\ &= P(X > 0, Z > \frac{-\rho}{\sqrt{1-\rho^2}} X) \end{aligned}$$

$$= \frac{|\Theta|}{360}$$

$$|\Theta| = 90^\circ + |\alpha|$$



$$\tan \alpha = \frac{-\rho}{\sqrt{1-\rho^2}}$$

$$\alpha = \arctan \left(\frac{-\rho}{\sqrt{1-\rho^2}} \right)$$

$$= \frac{90^\circ + \arctan \left(\frac{-\rho}{\sqrt{1-\rho^2}} \right)}{360}$$

ex ^{Another example} $M = \text{midterm score}$

$F = \text{final score}$

M and F have bivariate $(66, 73, 10^2, 8^2, 0.6)$

let overall score $S = .3M + .7F$

What is joint distribution of F and S ?

Need 5 param: $M_F = 73$

$$\sigma_F^2 = 8^2$$

$$M_S = 0.3(66) + (0.7)(73) = 70.9$$

$$\sigma_S^2 - ? \text{ — derived below}$$

$$\text{corr}(F, S) - ? \text{ — derived below}$$

$$\begin{aligned} \text{Cov}(M, F) &= \text{corr}(M, F) \sigma_M \sigma_F \\ &= .6(10)(8) = \boxed{48} \end{aligned}$$

$$\begin{aligned} \text{Cov}(F, S) &= \text{Cov}(F, S) = \text{Cov}(F, .3M + .7F) \\ &= .3 \text{Cov}(F, M) + .7 \text{Cov}(F, F) \\ &= .3(48) + .7(8^2) \\ &= \boxed{59.2} \end{aligned}$$

$$\begin{aligned}
 \text{Var}(S) &= \text{Var}(0.3M + 0.7F) \\
 &= 0.3^2 \text{Var}(M) + (0.7)^2 \text{Var} F + 2\text{Cov}(0.3M, 0.7F) \\
 &= 0.3^2(10^2) + 0.7^2(8^2) + 2(0.3)(0.7)(48) \\
 &= 60.2
 \end{aligned}$$

$$\text{Corr}(F, S) = \frac{\text{Cov}(F, S)}{\text{SD}(F)\text{SD}(S)} = \frac{59.2}{\sqrt{60.2}} = 0.95$$