

Stat 134 1ec 6

Warmup 10:00 - 10:10

You roll a fair die 600 times.

Using the normal distribution find the chance you get between 74.5 and 75.5 sixes.

Leave your answer in terms of Φ .

$$\boxed{\Phi\left(\frac{75.5 - \mu}{\sigma}\right) - \Phi\left(\frac{74.5 - \mu}{\sigma}\right)} \\ = .00101$$

$$\mu = np = \frac{600}{6} = 100$$

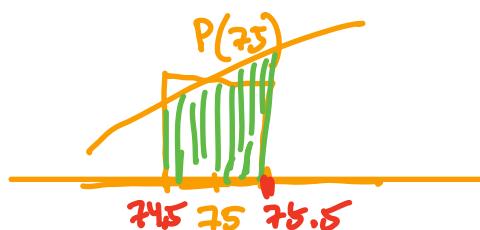
$$\sigma = \sqrt{npq} = \sqrt{100(1/6)(5/6)} = 9.1$$

$$= \sqrt{600(1/6)(5/6)} = 9.1$$

exact value that you get 75 sixes out of 600 rolls

$$\rightarrow \binom{600}{75} \left(\frac{1}{6}\right)^{75} \left(\frac{5}{6}\right)^{525} = .00087$$

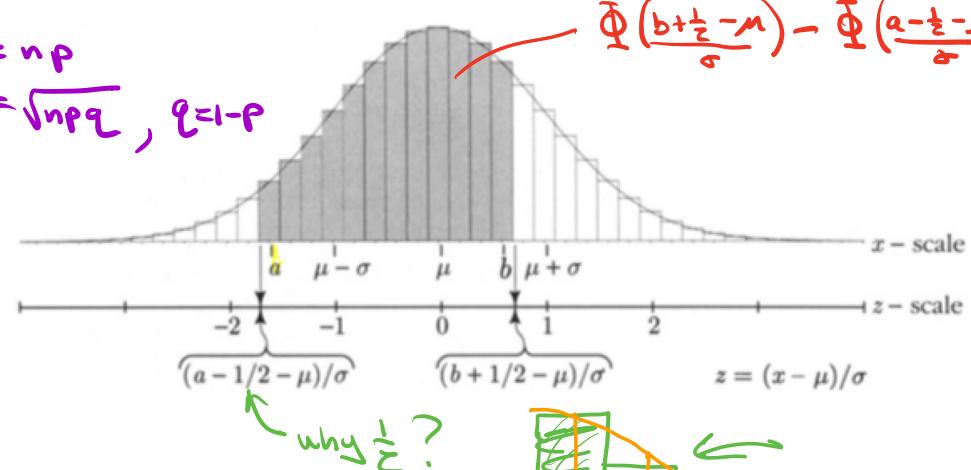
Motivation: It is difficult to calculate exact probabilities with the binomial formula. It is easier to calculate the area under the normal curve.



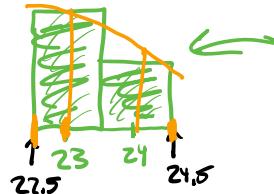
Last time Sec 2.2 Normal Approx to binomial

$$\mu = np$$

$$\sigma = \sqrt{npq}, q = 1-p$$



Continuity correction (CC)
We are approximating a discrete distribution (binomial) by a continuous one (normal)



ex

? 135

Suppose that each of 300 patients has a probability of $1/3$ of being helped by a treatment independent of its effect on the other patients. Find approximately the probability that **more than 134** patients are helped by the treatment. (Be sure to use the continuity correction. You will not receive full credit otherwise)

Hint The mean of $\text{Bin}(n, p)$ is $\mu = np$

The standard deviation of $\text{Bin}(n, p)$ is $\sigma = \sqrt{npq}$

$X = \# \text{ patients helped}$

$$P(X \geq 134.5) = 1 - \Phi(x < 134.5) = 1 - \Phi\left(\frac{134.5 - 100}{\sqrt{300(1/3)(2/3)}}\right) \approx 1 - 1 = 0$$

Tuesday

(1) Finish sec 2.2

(2) sec 2.4 Poisson approximation (skip sec 2.3)

① Sec 2.2 Normal approximation to the binomial distribution

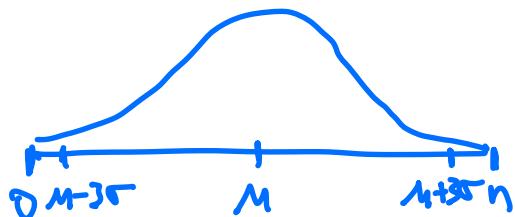
For what n, p is it ok to approximate $\text{Bin}(n, p)$ by a normal distribution $N(\mu, \sigma^2)$.

$n \geq 20$ since for fixed p , the binomial is more normal shaped as n increases.

Outcomes for $\text{Bin}(n, p)$ are $0, 1, 2, \dots, n$

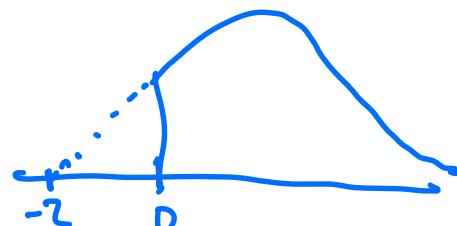
all of our data lie between $\mu \pm 3\sigma$

so replace $\mu - 3\sigma > 0$ and $\mu + 3\sigma < n$



ex Can we approx. $\text{Bin}(20, \frac{1}{10})$ by the normal?

$$\begin{aligned} n &= 20 & \checkmark \\ M - 3\sigma &= 2 - 4 = -2 & \times \\ M &= 2 & \text{green oval} \\ \sigma &= \sqrt{20 \left(\frac{1}{10}\right)\left(\frac{9}{10}\right)} & \text{green oval} \end{aligned}$$



check $M + 3\sigma = 6 < 20 \checkmark$

- . (3 pts) Southwest overbooks its flights, knowing some people will not show up. Suppose for a plane that seats 350 people, 360 tickets are sold. Based on previous data, the airline claims that each passenger has a 90% chance of showing up. Approximately, what is the chance that at least one empty seat remains? (There are no assigned seats.)

$X = \# \text{ people who show up}$

$$P(X \leq 349.5)$$

$$= \Phi\left(\frac{349.5 - \mu}{\sigma}\right)$$

$$= \Phi(4.47) = \boxed{1}$$

$$\mu = 360(.9) = 324$$

$$\sigma = \sqrt{360(.9)(.1)} = 5.7$$

$$\mu + 3\sigma < n \quad \checkmark$$

$$\mu - 3\sigma > 0 \quad \checkmark$$

$$n > 20 \quad \checkmark$$

(2) Sec 2.4 (skip 2.3) Poisson approx to Binomial

The normal approximation has almost 100% of data $\pm 3\sigma$ from the mean M . For this reason we approximated the binomial w/ the normal only when $M \leq 3\sigma$ is between 0 and n .

For cases when p is small (or p is close to 1)

and n is large, we approximate

$\text{Bin}(n, p)$ by $\text{Pois}(u=np)$

Picture

$$p = \frac{1}{6}$$

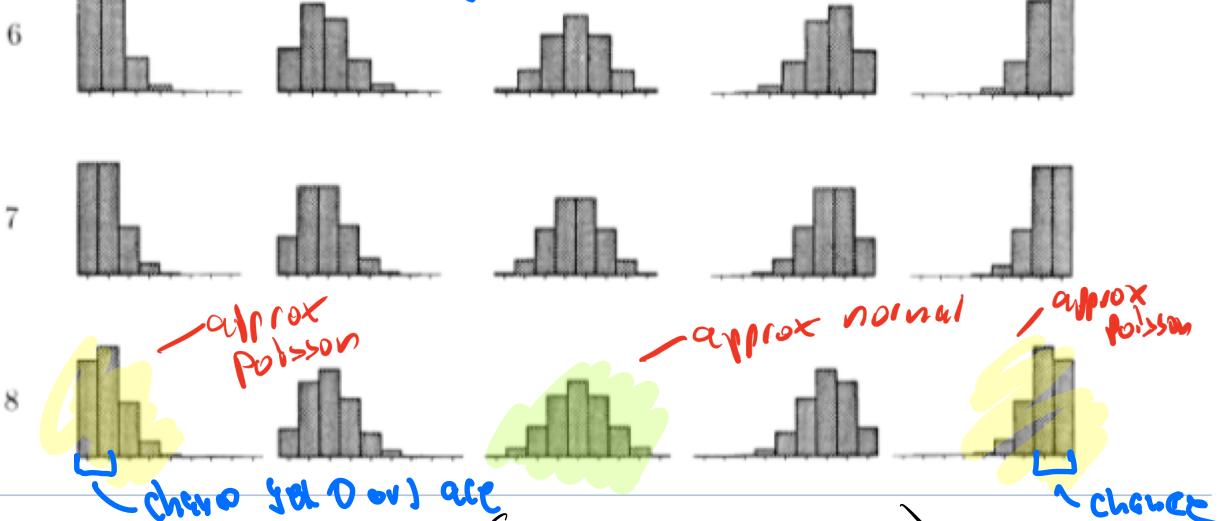
$$p = \frac{1}{4}$$

$$p = \frac{1}{2}$$

$$p = \frac{3}{4}$$

$$p = \frac{7}{8}$$

Suppose $p = \text{chance you get ace in a fair 8-sided die} = \frac{1}{8}$
 $n = \# \text{times you roll die}$



Defn $\text{Poisson}(u)$ (written $\text{Poi}(u)$)

intuitively many outcomes.

$$P(k) = \frac{e^{-u} u^k}{k!} \text{ for } k=0, 1, 2, \dots$$

You can just define the $\text{Poisson}(u)$ distribution this way or think of it as a limit of the Binomial formula when n is large and p is small and $np \rightarrow u$.

Proven in appendix at end of lecture notes,

Then

$$P_n(k) = \binom{n}{k} p^k (1-p)^{n-k} \xrightarrow{n \rightarrow \infty} \frac{e^{-\lambda} \lambda^k}{k!} \quad \text{as } n \rightarrow \infty \text{ and } p \rightarrow 0 \quad \text{with } np \rightarrow \lambda$$

Ex Bet 500 times, independently, on a bet with λ_{1000} large Small

Approximate the chance of winning at least once.

Dont use CC since Poisson is discrete.

$$P(\text{win} \geq 1 \text{ bet})$$

Defⁿ Poisson(λ)

$$P(k \geq 1) = 1 - P(0)$$

$$P(k) = \frac{e^{-\lambda} \lambda^k}{k!} \text{ for } k=0, 1, 2, \dots$$

$$= 1 - e^{-\lambda} \left(\frac{\lambda}{\lambda}\right)^0 = 1 - e^{-\lambda} \quad \lambda = 500 \left(\frac{1}{1000}\right) = \frac{1}{2}$$

$$=.3943$$

exactly (binomial)

$$1 - P(0) = 1 - \left(\binom{500}{0} \left(\frac{1}{1000}\right)^0 \left(\frac{999}{1000}\right)^{500}\right)$$

$$= 1 - \left(\frac{999}{1000}\right)^{500} = .3936$$

Calculating $P(k)$ using the Poisson formula versus the Binomial formula is a little easier. The main point I want to make is that $\text{Pois}(\lambda)$ is related to $\text{Bin}(n, p)$,

What about those binomials with p close to 1?

p = chance of success

q = chance of failure

If $p \approx 1$ then $q \approx 1-p \approx 0$

$\text{Bin}(n, q) \approx \text{Pois}(\lambda = nq)$ for large n , small q .

$\underline{p \text{ large}}$

$\underline{\text{large}}$

$\approx 97.8\%$ of approx 30 million poor families in the US. have a fridge. If you randomly sample 100 of these families roughly what is the chance 98 or more have a fridge?

Defn Poisson (μ)

$$P(k) = \frac{e^{-\mu} \mu^k}{k!} \text{ for } k=0, 1, 2, \dots$$

Think about this for next time.

Appendix

Then let $P_n(k) = \binom{n}{k} p^k (1-p)^{n-k}$ (binomial formula)

$$P_n(k) = \binom{n}{k} p^k (1-p)^{n-k} \xrightarrow{\substack{n \rightarrow \infty \\ p \rightarrow 0 \\ np = \mu}} \frac{e^{-\mu} \mu^k}{k!} \text{ as } n \rightarrow \infty \text{ and } p \rightarrow 0 \text{ with } np = \mu ?$$

Pf/ The claim follows if we show these 2 facts:

$$\textcircled{1} \quad P_n(0) \approx e^{-\mu}$$

$$\textcircled{2} \quad P_n(k) = P_n(k-1) \frac{\mu}{k}$$

$$\text{so } P_n(1) = e^{-\mu} \frac{\mu}{1}$$

$$P_n(2) = P_n(1) \frac{\mu}{2} = e^{-\mu} \frac{\mu}{1} \cdot \frac{\mu}{2} = e^{-\mu} \frac{\mu^2}{2!}$$

etc

Proof of fact $\textcircled{1}$: $P_n(0) \approx e^{-\mu}$

Remember from Calculus $\log(1+x) \approx x$ for x small

let $P_n(k) = \binom{n}{k} p^k (1-p)^{n-k}$ binomial formula

$$P_n(0) = (1-p)^n \quad \begin{matrix} \text{PSmall} \\ np = \mu \end{matrix}$$

$$\Rightarrow \log P_n(0) = n \log(1-p) \approx n(-p) = -\mu$$

$$\Rightarrow P_n(0) = e^{-\mu}$$

$$P_n(k) = P_n(k-1) \frac{m}{k}$$

Proof of fact(2):

Remember from sec 2.1 p85, $\frac{P_n(k)}{P_n(k-1)} = \left[\frac{n-k+1}{k} \right] \frac{p}{q}$

$$\begin{aligned} \Rightarrow P_n(k) &= P_n(k-1) \left[\frac{n-(k-1)}{k} \right] \frac{p}{q} \\ &= P_n(k-1) \left[\frac{np - (k-1)p}{k} \right] \frac{1}{q} \approx P_n(k-1) \frac{m}{k} \quad \square \end{aligned}$$