

stat 134 lec 10

Last time

Sec 3.1 Random Variables

A RV, X , is the outcome of an experiment.
 (x, y) is the joint outcome of two RVs X, Y .
The event $(X=x, Y=y)$ is the intersection of events
 $X=x$ and $Y=y$.

The probability X and Y satisfies some condition
is the sum of $P(x, y)$ over all pairs (x, y) that
satisfy that condition.

$$\text{e.g. } P(X=Y) = \sum_{(x,y): x=y} P(x,y) = \sum_{\text{all } x} P(x,x)$$

$$\text{e.g. } P(X+Y=s) = \sum_{(x,y): x+y=s} P(x,y) = \sum_{\text{all } x} P(x, s-x)$$

Independence of (X, Y, Z) means

$$P(X=x, Y=y, Z=z) = P(X=x)P(Y=y)P(Z=z) \quad \text{for all } x \in X, \\ y \in Y, z \in Z.$$

Today

- ① Student responses from Concept test last time.
- ② Sec 3.1 Sums of independent Poissons to Poisson
- ③ Sec 3.2 Expectations of a RV.

①

1. The joint distribution of X and Y is drawn below:

		$\frac{3}{8}$	$\frac{1}{2}$	$\frac{1}{8}$	$P(X)$
		$P(Y)$			
1		$\frac{1}{4}$	$\frac{1}{3}$	$\frac{1}{12}$	$\frac{3}{5}$
0		$\frac{1}{8}$	$\frac{1}{6}$	$\frac{1}{24}$	$\frac{1}{3}$
$P(X)$	$P(Y)$	0	1	2	

- a X and Y are independent
b $P(X = x|Y = 0) = P(X = x|Y = 1)$,
for all x.
c More than one of the above
d None of the above

c

X and Y are independent since each joint probability is the product of the two respective marginal probabilities. Since X and Y are independent, $P(X=x)$ should not depend on Y, making B true. Thus, C is the correct answer.

c

Since this is independent, whatever y is, it does not affect the chance of x.

d

God help me

a

The two rows $Y=0$ and $Y=1$ give different probabilities for each value of X

if we divide both row by their marginal prob we get the same answer

② Sum of independent Poisson is Poisson

informal argument:

$$\begin{aligned} X_1 &\sim \text{Bin}(1000, \frac{1}{1000}) \\ X_2 &\sim \text{Bin}(2000, \frac{1}{1000}) \end{aligned} \quad \left. \begin{array}{l} \xrightarrow{\text{indep}} \\ \approx \text{Pois}(2) \end{array} \right.$$

$$X_1 + X_2 \sim ? \quad \text{Bin}(3000, \frac{1}{1000}) \approx \text{Pois}(3)$$

$$\begin{aligned} X_1 + X_2 &= \# \text{ heads in } 1000 + 2000 = 3000 \text{ coin tosses.} \\ p &= \frac{1}{1000} \end{aligned}$$

So sum of two indep binomials with the same p is Binomial and this example suggest that sum of 2 indep Poisson is Poisson.

Let's prove this rigorously:

Recall binomial theorem

$$\begin{aligned} (a+b)^3 &= \binom{3}{0} a^3 b^0 + \binom{3}{1} a^2 b^1 + \binom{3}{2} a^1 b^2 + \binom{3}{3} a^0 b^3 \\ &= a^3 + 3a^2 b + 3ab^2 + b^3 \end{aligned}$$

$$(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^{n-k} b^k$$

Recall $X \sim \text{Pois}(\mu)$

$$P(X=k) = \frac{e^{-\mu} \mu^k}{k!}$$

Claim If $X \sim \text{Pois}(\mu)$ and $Y \sim \text{Pois}(\lambda)$ are independent then

$$S = X + Y \sim \text{Pois}(\mu + \lambda).$$

Pf/ $P(S=s) = P(X=0, Y=s) + P(X=1, Y=s-1) + \dots + P(X=s, Y=0)$

addition rule

independent of X, Y

$$= \sum_{k=0}^s P(X=k, Y=s-k)$$

$$= \sum_{k=0}^s P(X=k) P(Y=s-k)$$

$$= \sum_{k=0}^s \frac{e^{-\lambda} \lambda^k}{k!} \cdot \frac{e^{-\mu} \mu^{s-k}}{(s-k)!}$$

$$\frac{s!}{s!} = 1 \quad \xrightarrow{\text{binomial theorem}}$$

$$= e^{-(\lambda+\mu)} \frac{1}{s!} \sum_{k=0}^s \frac{s!}{k!(s-k)!} \lambda^k \mu^{s-k}$$

$$= e^{-(\lambda+\mu)} \frac{1}{s!} (\mu+\lambda)^s$$

$$\Rightarrow S \sim \text{Pois}(\mu + \lambda).$$

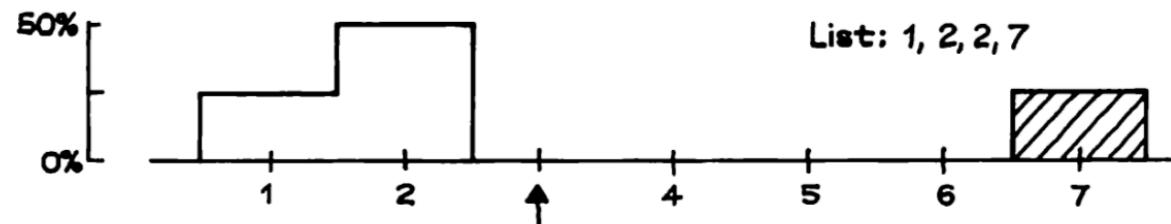
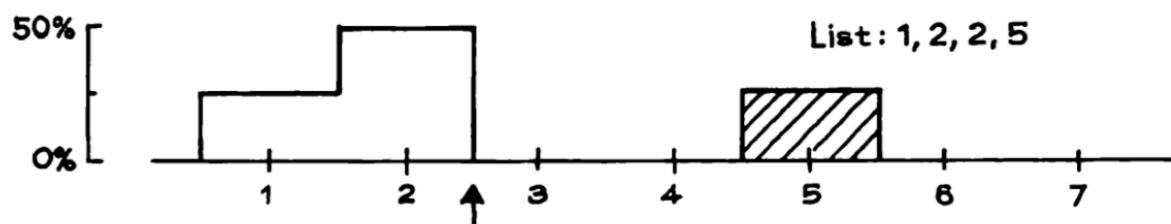
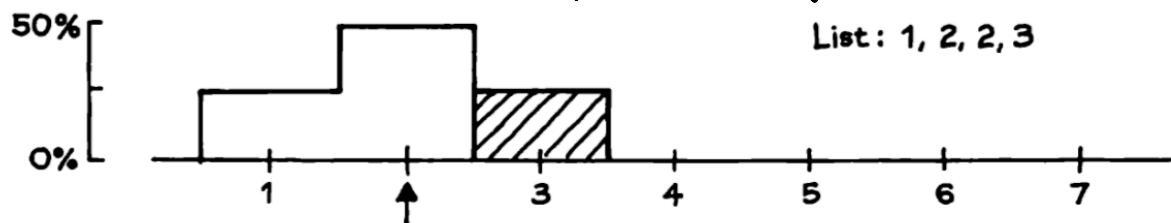
□

(3)

Sec 3.2 Expectation

$$E(X) = \sum_{x \in X} x \cdot P(X=x)$$

$$E(X) = \frac{1}{4} \cdot 1 + \frac{1}{4} \cdot 2 + \frac{1}{4} \cdot 2 + \frac{1}{4} \cdot 3 = 2$$



Properties of Expectation - Pitman

$$\textcircled{1} \quad E(c) = c$$

$$\textcircled{2} \quad E(X+Y) = E(X)+E(Y) \quad (X, Y \text{ don't need to be independent})$$

$$\textcircled{3} \quad E(aX+b) = aE(X)+b$$

Indicators

An indicator is a RV that has only 2 values 1 (w/prob p) and 0 (w/prob 1-p).

$$I = \begin{cases} 1 & \text{prob } p \\ 0 & \text{prob } 1-p \end{cases} \quad \text{Same as a Bernoulli p trial.}$$

$$E(I) = 1 \cdot p + 0 \cdot (1-p) = p$$

RV that are counts can often be written as a sum of indicators.

$$\text{ex: } X \sim \text{Bin}(n, p)$$

↙ # successes in n Bernoulli p trials,

ex: $X = \# \text{ heads in } n \text{ flips at } p \text{ coin}$

$$X = I_1 + I_2 + \dots + I_n$$

where $I_j = \begin{cases} 1 & \text{if } j^{\text{th}} \text{ trial success} \\ 0 & \text{else} \end{cases}$

$$E(X) = E(I_1) + \dots + E(I_n) \quad \boxed{np}$$

indicators are independent since trials are indep.

$$\text{Def } X \sim \text{Hyper}(N, G, n)$$

Ex $X = \# \text{ aces in a poker hand from a deck of cards}$

$$N = 52$$

$$G = 4$$

$$n = 5$$

a) What are the range of values of X ?

$$0, 1, 2, 3, 4$$

b) Write X as a sum of indicator \downarrow

$$X = I_1 + I_2 + I_3 + I_4 + I_5$$

c) How is I_2 defined?

$$I_2 = \begin{cases} 1 & \text{if 2nd card is an ace} \\ 0 & \text{else} \end{cases}$$

d) Find $E(I_2)$

$$E(I_2) = P(\text{2nd card is an ace}) = \boxed{\frac{4}{52}}$$

e) Find $E(X)$

$$E(X) = 5 \cdot E(I_1) = \boxed{5 \cdot \left(\frac{4}{52}\right)}$$

Note One student in class suggested
we define $I_2 = \begin{cases} 1 & \text{if get 2 aces} \\ 0 & \text{else} \end{cases}$

$$\text{so } X = I_1 + 2 \cdot I_2 + 3 \cdot I_3 + 4 \cdot I_4$$

This is also correct but more complicated
than my solution.

We have

$$E(I_1) = \frac{\binom{4}{1}\binom{48}{4}}{\binom{52}{5}}$$

$$E(I_2) = \frac{\binom{4}{2}\binom{48}{3}}{\binom{52}{5}}$$

$$E(I_3) = \frac{\binom{4}{3}\binom{48}{2}}{\binom{52}{5}}$$

$$E(I_4) = \frac{\binom{4}{4}\binom{48}{1}}{\binom{52}{5}}$$

$$\text{so } E(X) = \frac{1}{\binom{52}{5}} \left[\binom{4}{1}\binom{48}{4} + 2 \cdot \binom{4}{2}\binom{48}{3} + 3 \cdot \binom{4}{3}\binom{48}{2} + 4 \cdot \binom{4}{4}\binom{48}{1} \right]$$

wow! $\rightarrow 5 \cdot \left(\frac{4}{52}\right)$ \leftarrow I checked this
using R

Ex Suppose a fair die is rolled 10 times.

Let $X =$ Number of different faces
that appear in 10 rolls.

Ex If roll 2, 3, 4, 2, 3, 5, 2, 3, 3, 2 then $X = 4$

a) What are the range of values of X ?

1, 2, 3, 4, 5, 6

b) Write X as a sum of indicator

$$X = I_1 + I_2 + I_3 + I_4 + I_5 + I_6$$

c) How is I_2 defined?

$$I_2 = \begin{cases} 1 & \text{if 2 appears at least once,} \\ 0 & \text{else} \end{cases}$$

d) Find $E(I_2)$

$$= 1 - P(2 \text{ never appears}) = 1 - \left(\frac{5}{6}\right)^{10}$$

e) Find $E(X)$

$$6 \cdot \left(1 - \left(\frac{5}{6}\right)^{10}\right)$$

Stat 134

Chapter 3 Wednesday February 13 2019

1. A forgetful valet is attempting to return n cars to their n rightful owners. For each driver, the valet remembers the car correctly 5% of the time; otherwise the valet retrieves a car at random (possibly the correct car). What is the probability that the second driver retrieves his own car.

- a .05
b $.05 + .95/n$
c $.05 + .95/(n - 1)$
d none of the above

we have,
 $= P(\text{Correct, remembers}) + P(\text{Correct, forgets})$
 $= P(\text{Correct} | \text{remember}) \cdot P(\text{remember}) + P(\text{correct} | \text{forget}) \cdot P(\text{forget})$

$\frac{1}{1}$ $\frac{.05}{.05}$ $\frac{1}{n}$ $\frac{.95}{.95}$

*we aren't conditioning on the outcome of returning anyone else's car.
Everyone has the same chance of getting their car back!*

2. A forgetful valet is attempting to return n cars to their n rightful owners. For each driver, the valet remembers the car correctly 5% of the time; otherwise the valet retrieves a car at random (possibly the correct car). Let N be the number of drivers who retrieve their own car. $E(N)$ is:

- a**. $05n + .95$
- b**. $.05n$
- c**. $.05n + 1$
- d** none of the above

$N = \# \text{ drivers who retrieve their own car}$
 $\hookrightarrow 0, 1, 2, \dots, n.$

$$N = I_1 + I_2 + \dots + I_n, \quad I_j = \begin{cases} 1 & \text{if } j^{\text{th}} \text{ driver retrieves their own car,} \\ 0 & \text{else.} \end{cases}$$

$$E(N) = \mathbb{E} \left(\sum_{j=1}^n I_j \right) = \sum_{j=1}^n \mathbb{E}(I_j) = \sum_{j=1}^n \left(0.05 + \frac{0.95}{n} \right) = \boxed{0.05n + \frac{0.95}{n}}$$