Prob 140 Fall 2018 Final Exam Reference	inal Exar	n Reference Sheet			A	A. Adhikari
name and parameters	values	mass fn. or pdf	cdf or survival fn.	expectation	variance	MGF M(t)
Uniform	$m \le k \le n$	1/(n-m+1)		(m+n)/2	$((n-m+1)^2-1)/12$	
Bernoulli (p)	0, 1	$p_1=p$ , $p_0=q$		р	bd	$q + pe^t$
Binomial $(n, p)$	$0 \le k \le n$	$\binom{n}{k} p^k q^{n-k}$		du	bdu	$(q+pe^t)^n$
Poisson $(\mu)$	<i>k</i> ≥ 0	$e^{-\mu}\mu_k/k$ i		η	щ	$\exp(\mu(e^t-1))$
Geometric ( <i>p</i> )	$k \geq 1$	$q^{k-1}p$	$P(X > k) = q^k$	$1/\rho$	$q/p^2$	
"Negative binomial" $(r, p)$	$k \geq r$	$inom{(k-1)}{(r-1)} p^{r-1} q^{k-r} p$		r/p	$rq/p^2$	
Geometric ( <i>p</i> )	$k \geq 0$	$d_{k}^{b}$	$P(X>k)=q^{k+1}$	d/b	$q/p^2$	
Negative binomial $(r, p)$	$k \geq 0$	$inom{(^{k+r-1})}{r-1}igphi^{r-1}q^k p$		rq/p	$rq/p^2$	
Hypergeometric $(N, G, n)$	$0 \le g \le n$	$\binom{G}{g}\binom{B}{b}/\binom{N}{n}$		$n\frac{G}{N}$	$n \frac{G}{N} \cdot \frac{B}{N} \cdot \frac{N-n}{N-1}$	
Uniform	$x \in (a, b)$	1/(b-a)	F(x) = (x-a)/(b-a)	(a+b)/2	$(b-a)^2/12$	
Beta (r, s)	$x \in (0,1)$	$\frac{\Gamma(r+s)}{\Gamma(r)\Gamma(s)} x^{r-1} (1-x)^{s-1}$	ts if r, s integers	r/(r+s)	$rs/((r+s)^2(r+s+1))$	
Exponential $(\lambda) = Gamma \; (1,\lambda)$	0 < ×	$\lambda e^{-\lambda x}$	$F(x) = 1 - e^{-\lambda x}$	$1/\lambda$	$1/\lambda^2$	
Gamma $(r,\lambda)$	0 < ×	$\frac{\lambda^r}{\Gamma(r)} \chi^{r-1} e^{-\lambda x}$		r/\	$r/\lambda^2$	$(\lambda/(\lambda-t))^r$ for $t<\lambda$
Chi-square $(n)$	0 ≤ x	same as gamma $(n/2, 1/2)$		n	2 <i>n</i>	
Normal (0, 1)	$x \in \mathbb{R}$	$\phi(x)=rac{1}{\sqrt{2\pi}}\mathrm{e}^{-rac{1}{2}x^2}$	cdf: $\phi(x)$	0	1	$\exp(t^2/2)$
Normal $(\mu,\sigma^2)$	$x \in \mathbb{R}$	$\frac{1}{\sigma}\phi((x-\mu)/\sigma)$	cdf: $\Phi((x-\mu)/\sigma)$	$\mu$	$\sigma^2$	
Rayleigh	0 < x	$xe^{-\frac{1}{2}x^2}$	$F(x) = 1 - e^{-\frac{1}{2}x^2}$	$\sqrt{\pi/2}$	$(4-\pi)/2$	
Cauchy	$x \in \mathbb{R}$	$1/\pi(1+x^2)$	$F(x)=rac{1}{2}+rac{1}{\pi}arctan(x)$			

• If  $X_1, X_2, \ldots, X_n$  are i.i.d. with variance  $\sigma^2$ , then  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$  is an unbiased estimator of  $\sigma^2$  but  $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$  is not.

• For r>0, the integral  $\Gamma(r)=\int_0^\infty x^{r-1}e^{-x}dx$  satisfies  $\Gamma(r+1)=r\Gamma(r)$ . So  $\Gamma(r)=(r-1)!$  if r is an integer. Also,  $\Gamma(1/2)=\sqrt{\pi}$ .

• If  $Z_1$  and  $Z_2$  are i.i.d. standard normal then  $\sqrt{Z_1^2+Z_2^2}$  is Rayleigh.

• If Z is standard normal then  $E(|Z|) = \sqrt{2/\pi}$ 

• The kth order statistic  $U_{(k)}$  is kth smallest of  $U_1, U_2, \ldots, U_n$  i.i.d. uniform (0,1), so  $U_{(1)}$  is min and  $U_{(n)}$  is max. Density of  $U_{(k)}$  is beta (k, n-k+1).

• If  $S_n$  is the number of heads in n tosses of a coin whose probability of heads was chosen according to the beta (r,s) distribution, then the distribution of  $S_n$  is beta-binomial (r,s,n) with  $P(S_n=k)=\binom{n}{k}C(r,s)/C(r+k,s+n-k)$  where  $C(r,s)=\digamma(r+s)/(\digamma(r)\digamma(s))$  is the constant in the beta (r,s) density.

ullet If **X** has mean vector  $\mu$  and covariance matrix  $\Sigma$  then  ${\sf AX}+{\sf b}$  has mean vector  ${\sf A}\mu+{\sf b}$  and covariance matrix  ${\sf A}\Sigma{\sf A}^T$ .

• If X has the multivariate normal distribution with mean vector  $\mu$  and covariance matrix  $\Sigma$ , then X has density  $f(\mathbf{x}) = \frac{1}{(\sqrt{2\pi})^n \sqrt{\det(\Sigma)}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mu)^T \Sigma^{-1}(\mathbf{x} - \mu)\right)$ 

• The least squares linear predictor of Y based on the  $p \times 1$  vector  $\mathbf{X}$  is  $\hat{Y} = \mathbf{b}^T(\mathbf{X} - \mu_{\mathbf{x}}) + \mu_Y$  where  $\mathbf{b} = \Sigma_{\mathbf{x}}^{-1} \Sigma_{\mathbf{x},Y}$ . Here the ith element of the  $p \times 1$  vector  $\Sigma_{\mathbf{x},Y}$  is  $Cov(X_i,Y_i)$ . In the case p=1 this is the equation of the regression line, with slope  $Cov(X_i,Y_i)/Var(X) = rSD(Y_i)/SD(X_i)$  and intercept  $E(Y_i)$  slope $E(X_i)$ .

• If  $W = Y - \hat{Y}$  is the error in the least squares linear prediction, then E(W) = 0 and  $Var(W) = Var(Y) - \Sigma_{Y,X}\Sigma_X^{-1}\Sigma_{X,Y}$ . In the case p = 1,  $Var(W) = (1 - r^2)Var(Y)$ .

• If Y and X are multivariate normal then the formulas in the above two bullet points are the conditional expectation and conditional variance of Y given X.

• If Y and X are standard bivariate normal with correlation r, then  $Y = rX + \sqrt{1 - r^2}Z$  for some standard normal Z independent of X.

ullet Under the multiple regression model  ${f Y}={f X}eta+\epsilon$ , the least squares estimate of eta is  $\hat{eta}=({f X}^{\sf T}{f X})^{-1}{f X}^{\sf T}{f Y}$ .