# EM Algorithm for Finite Mixture Regression

HW 5 of STAT 5361 Statistical Computing

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### 1 E- and M-Steps Derivations

### 1.1 E-Step Derivation

$$Q(\Psi|\Psi^{(k)}) = \sum_{Z} \left[ p(Z|\mathbf{y}, X, \Psi^{(k)}) \log p(\mathbf{y}, Z|X, \Psi) \right]$$
(1)

$$= \sum_{Z} \left[ p(Z|\mathbf{y}, X, \Psi^{(k)}) \log \prod_{i=1}^{n} p(y_i, \mathbf{z}_i | \mathbf{x}_i, \Psi) \right]$$
(2)

$$= \sum_{i=1}^{n} \sum_{Z} \left[ p(Z|\mathbf{y}, X, \Psi^{(k)}) \log p(y_i, \mathbf{z}_i | \mathbf{x}_i, \Psi) \right]$$
(3)

$$= \sum_{i=1}^{n} \sum_{\mathbf{z}_{i}} \left[ p(\mathbf{z}_{i}|\mathbf{y}, X, \Psi^{(k)}) \log p(y_{i}, \mathbf{z}_{i}|\mathbf{x}_{i}, \Psi) \right]$$

$$(4)$$

$$= \sum_{i=1}^{n} \sum_{\mathbf{z}_i} \left[ p(\mathbf{z}_i | y_i, \mathbf{x}_i, \Psi^{(k)}) \log p(y_i, \mathbf{z}_i | \mathbf{x}_i, \Psi) \right]$$
(5)

$$= \sum_{i=1}^{n} \sum_{i=1}^{m} \left[ p(\mathbf{z}_{i} = (0, \dots, 1, \dots, 0)' | y_{i}, \mathbf{x}_{i}, \Psi^{(k)}) \log p(y_{i}, \mathbf{z}_{i} = (0, \dots, 1, \dots, 0)' | \mathbf{x}_{i}, \Psi) \right]$$
(6)

$$= \sum_{i=1}^{n} \sum_{i=1}^{m} \left[ p(z_{ij} = 1 | y_i, \mathbf{x}_i, \Psi^{(k)}) \log p(y_i, \mathbf{z}_i = (0, \dots, 1, \dots, 0)' | \mathbf{x}_i, \Psi) \right]$$
 (7)

$$= \sum_{i=1}^{n} \sum_{j=1}^{m} \left[ E(z_{ij}|y_i, \mathbf{x}_i, \Psi^{(k)}) \{ \log \pi_j + \log \varphi(y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j, 0, \sigma^2) \} \right]$$
(8)

$$= \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \{ \log \pi_j + \log \varphi(y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j, 0, \sigma^2) \}$$
 (9)

(10)

where

$$Z = \begin{pmatrix} \mathbf{z}_{1}' \\ \vdots \\ \mathbf{z}_{n}' \end{pmatrix} = \begin{pmatrix} z_{11} & z_{12} & \cdots & z_{1m} \\ \vdots & \vdots & \vdots & \vdots \\ z_{n1} & z_{n2} & \cdots & z_{nm} \end{pmatrix} \quad X = \begin{pmatrix} \mathbf{x}_{1}' \\ \vdots \\ \mathbf{x}_{n}' \end{pmatrix} = \begin{pmatrix} x_{11} & x_{12} & \cdots & x_{1p} \\ \vdots & \vdots & \vdots & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{np} \end{pmatrix}$$

$$p_{ij}^{(k)} = E(z_{ij}|y_{i}, \mathbf{x}_{i}, \Psi^{(k)}) = \frac{\pi_{j}^{(k)}\varphi(y_{i} - \mathbf{x}_{i}^{T}\boldsymbol{\beta}_{j}^{(k)}, 0, \sigma^{2^{(k)}})}{\sum_{j=1}^{m} \pi_{j}^{(k)}\varphi(y_{i} - \mathbf{x}_{i}^{T}\boldsymbol{\beta}_{j}^{(k)}, 0, \sigma^{2^{(k)}})}$$

The elaboration of the above steps are

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- Step1 $\rightarrow$ Step2: Use independence among  $(y_i, \mathbf{z}_i)$
- Step3 $\rightarrow$ Step4: Marginal density of  $\mathbf{z}_i$
- Step4 $\rightarrow$ Step5: Use the fact  $\mathbf{z}_i \perp (y_1, \dots, y_{i-1}, y_{i+1}, \dots, y_n)|y_i$ , we can get rid of  $(y_1, \dots, y_{i-1}, y_{i+1}, \dots, y_n)|y_i$
- Step6-Step7: Easy to see conditional joint density is equal to condition marginal density

#### 1.2 M-Step Derivation

Since we have

$$\begin{split} Q(\Psi|\Psi^{(k)}) &= \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \{\log \pi_{j} + \log \varphi(y_{i} - \mathbf{x}_{i}^{T}\boldsymbol{\beta}_{j}, 0, \sigma^{2})\} \\ &= \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \log \pi_{j} - \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \log \sqrt{2\pi}\sigma - \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \frac{(y_{i} - \mathbf{x}_{i}^{T}\boldsymbol{\beta}_{j})^{2}}{2\sigma^{2}} \\ &= \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \log \frac{\pi_{j}}{\sqrt{2\pi}} - \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \log \sigma^{2} - \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} \frac{(y_{i} - \mathbf{x}_{i}^{T}\boldsymbol{\beta}_{j})^{2}}{\sigma^{2}} \\ &= I_{1} - \frac{1}{2} I_{2} - \frac{1}{2} I_{3} \end{split}$$

From the above, we can see only  $I_3$  contains  $\beta_j$  and

$$I_3 = \sum_{i=1}^n \sum_{j=1}^m p_{ij}^{(k)} \frac{(y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j)^2}{\sigma^2} = \sum_{j=1}^m \sum_{i=1}^n p_{ij}^{(k)} \frac{(y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j)^2}{\sigma^2}$$

To minimize  $I_3$ , we only need to fix j and optimize with regard to  $\beta_j$ . We can directly use the formula from generazied least square method and obtain

$$\boldsymbol{\beta}_{j}^{(k+1)} = (X'V^{-1}X)^{-1}X'V^{-1}\mathbf{y} = \left(\sum_{i=1}^{n} \mathbf{x}_{i}\mathbf{x}_{i}^{T}p_{ij}^{(k)}\right)^{-1} \left(\sum_{i=1}^{n} \mathbf{x}_{i}p_{ij}^{(k)}y_{i}\right) \quad j = 1, \dots, m$$

where

$$V^{-1} = diag(p_{1j}^{(k)}, \cdots, p_{nj}^{(k)})$$

Only  $I_2$  and  $I_3$  contains  $\sigma^2$ , since

$$I_2 + I_3 = \sum_{i=1}^n \sum_{j=1}^m p_{ij}^{(k)} \log \sigma^2 + \sum_{i=1}^n \sum_{j=1}^m p_{ij}^{(k)} \frac{(y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j)^2}{\sigma^2}$$

We minimize it with regard to  $\sigma^2$  given  $\boldsymbol{\beta}_j = \boldsymbol{\beta}_j^{(k+1)}$  and obtain

$$\sigma^{2^{(k+1)}} = \frac{\sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} (y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j^{(k+1)})^2}{\sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)}} = \frac{\sum_{i=1}^{n} \sum_{j=1}^{m} p_{ij}^{(k)} (y_i - \mathbf{x}_i^T \boldsymbol{\beta}_j^{(k+1)})^2}{n}$$

Only  $I_1$  contains  $\pi_i$  and

$$I_1 = \sum_{i=1}^n \sum_{j=1}^m p_{ij}^{(k)} \log \frac{\pi_j}{\sqrt{2\pi}} = -\frac{1}{2} \log(2\pi) \sum_{i=1}^n \sum_{j=1}^m p_{ij}^{(k)} + \sum_{j=1}^m \left(\sum_{i=1}^n p_{ij}^{(k)}\right) \log \pi_j$$

In order to maximize  $I_1$  under constraint  $\pi_1 + \cdots + \pi_m = 1$ . We use Lagrange multiplier method

$$L(\pi_1, \dots, \pi_m) = \sum_{j=1}^m \left(\sum_{i=1}^n p_{ij}^{(k)}\right) \log \pi_j - \lambda \left(\sum_{j=1}^m \pi_j - 1\right)$$

with  $\lambda$  a Lagrange multiplier. We can obtain

$$\pi_j^{(k+1)} = \frac{\sum_{i=1}^n p_{ij}^{(k)}}{\sum_{j=1}^m \sum_{i=1}^n p_{ij}^{(k)}} = \frac{\sum_{i=1}^n p_{ij}^{(k)}}{n} \quad j = 1, \dots, m$$

# 2 A Function to Implement EM Algorithm