Homework 7 - STAT 5361 Statistical Computing

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Abstract

This is homework 7 for STAT 5361 - Statistical Computing.

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1 Markov Chain Monte Carlo sampling

1.1 Posterior inference of Normal mixture

For a Normal mixture of $N(\mu_1, \sigma_1^2)$ and $N(\mu_2, \sigma_2^2)$ with proportion of first Normal distribution δ , we have

$$X_i \sim \delta N(\mu_1, \sigma_1^2) + (1 - \delta) N(\mu_1, \sigma_1^2)$$

Then the likelihood function of $X_i|\mu_1,\mu_2,\sigma_1^2,\sigma_2^2,\delta$ is given as

$$L(X_i|\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \delta) = \prod_{i=1}^n [\delta \cdot \phi(x_i|\mu_1, \sigma_1^2) + (1 - \delta) \cdot \phi(x_i|\mu_1, \sigma_1^2)]$$

which can be rewritten as

$$f(x|\theta) \propto \prod_{i=1}^{n} \left[\frac{\delta}{\sigma_1} \exp\left(\frac{(x_i - \mu_1)^2}{\sigma_1^2}\right) + \frac{1 - \delta}{\sigma_2} \exp\left(\frac{(x_i - \mu_2)^2}{\sigma_2^2}\right) \right]$$

Since the prior for $1/\sigma_1^2$ and $1/\sigma_1^2$ are $\Gamma(a,b)$ with shape a=0.5 and scale b=10, then $\sigma_1^2 \sim Inv\text{-}Gamma(a,1/b)$ and $\sigma_2^2 \sim Inv\text{-}Gamma(a,1/b)$. Also, the prior for μ_1 and μ_2 are $N(0,10^2)$ and all the priors are independent. The posterior density of $\theta=(\mu_1,\mu_2,\sigma_1^2,\sigma_2^2,\delta)$ is

$$q(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \delta | x) \propto \prod_{i=1}^n \left[\frac{\delta}{\sigma_1} \exp\left(\frac{(x_i - \mu_1)^2}{\sigma_1^2}\right) + \frac{1 - \delta}{\sigma_2} \exp\left(\frac{(x_i - \mu_2)^2}{\sigma_2^2}\right) \right] \cdot \exp(\mu_1^2) \cdot \exp(\mu_2^2) \cdot \left(\sigma_1^2\right)^{-1.5} \exp\left(\frac{1}{\sigma_1^2}\right) \cdot (\sigma_2^2)^{-1.5} \exp\left(\frac{1}{\sigma_2^2}\right)$$

The full conditional distributions of $\theta = (\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \delta)$ can be shown to be log-concave, which allows adaptive rejection alogorithm.

1.2 MCMC process in R

• Firstly, generate some data from a mixtrure Normal of N(-1,0.1) and N(5,0.5) with proportion of first Normal distribution $\delta = 0.7$.

```
## generate a random sample from a mixture normal with true parameter values
delta <- 0.7 # true value to be estimated based on the data
n <- 100
set.seed(123)
u <- rbinom(n, prob = delta, size = 1)
x <- rnorm(n, ifelse(u == 1, 7, 10) , 0.5)
mydata <- data.frame(x)</pre>
```

• Secondly, calculate the posterior density.

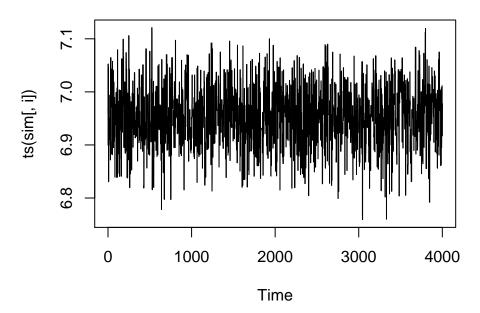
• Thirdly, An MCMC based the Gibbs sampler uses the ARMS algorithm from R package HI.

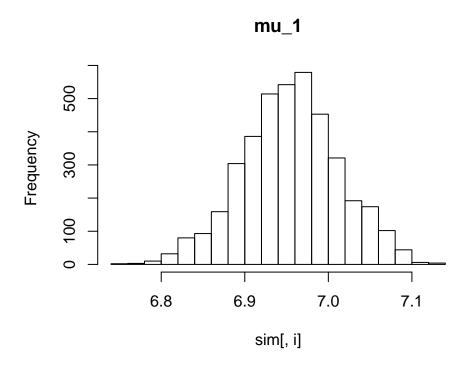
```
mymcmc <- function(niter, thetaInit, data, nburn) {</pre>
  p <- length(thetaInit)</pre>
  thetaCurrent <- thetaInit</pre>
  ## define a function for full conditional sampling
  logFC <- function(th, idx) {</pre>
    theta <- thetaCurrent
    theta[idx] <- th</pre>
    logpost(theta, data)
  out <- matrix(thetaInit, niter, p, byrow = TRUE)
  ## Gibbs sampling
  for (i in 2:niter) {
    for (j in 1:p) {
      ## general-purpose arms algorithm
      # Indicator function
      indF <- function(x, idx) {</pre>
         if (idx==1 \mid idx==2) \{(x<30)*(x>-30)\}
        else if (idx==3 \mid idx==4) \{(x<10)*(x>0)\}
         else \{(x<1)*(x>=0)\}
      }
      out[i, j] <- thetaCurrent[j] <-</pre>
        HI::arms(thetaCurrent[j], logFC, indF, 1, idx = j)
    }
  }
  out[-(1:nburn), ]
```

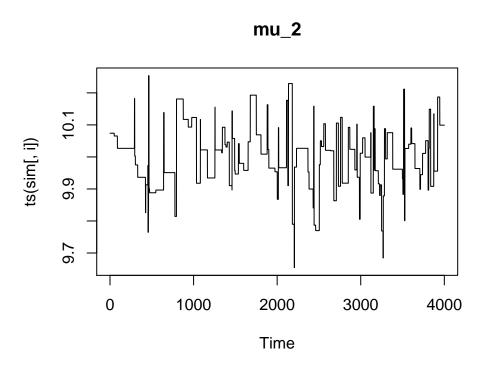
• Finally, run a simulation, drop the first 1000 observations and draw corresponding histogram for all parameters.

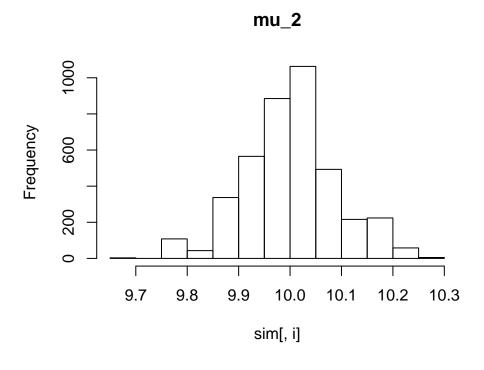
```
## simulation
niter <- 5000; nburn <- 1000
thetaInit <- c(0,0,0.2,0.2,0.5)
sim <- mymcmc(niter, thetaInit, mydata, nburn)
name_para <- c("mu_1","mu_2","sigma2_1","sigma2_2","delta")
for (i in 1:5) {
    plot(ts(sim[,i]),main=(paste(name_para[i])))
    hist(sim[,i],main=(paste(name_para[i])))
}</pre>
```

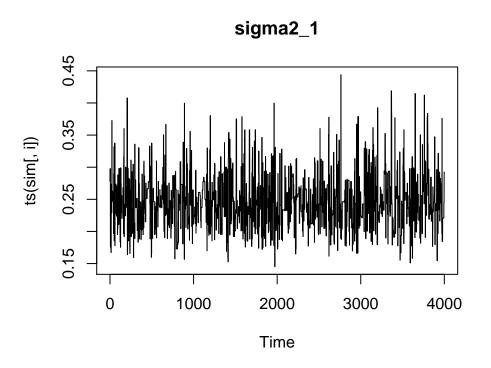
mu_1



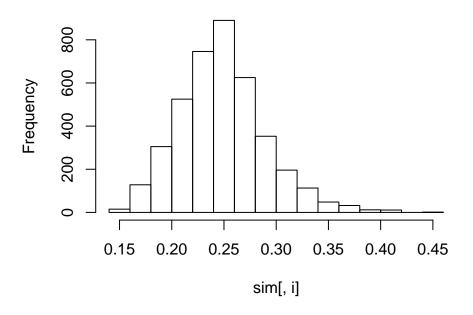








sigma2_1



sigma2_2

