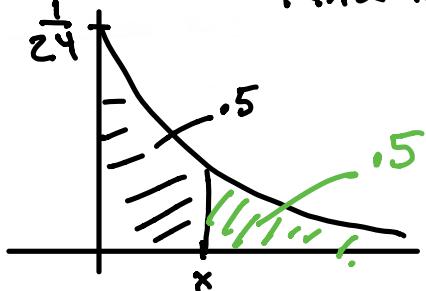


Stat 88 lec 32

Warmup: 2:00 - 2:10

3. Let X have the exponential distribution with mean 24 hours. Assume that X is measured in hours.

Find x such that $P(X \leq x) = .5$



$$f(x) = \lambda e^{-\lambda x}$$

$$E(X) = \frac{1}{\lambda}$$

$$F(x) = .5 \quad \text{find } x$$

"

$$P(X \leq x)$$

$$F(x) = .5 \Rightarrow S(x) = .5$$

$$P(X > x) = e^{-\lambda x} = e^{-\frac{1}{24}x}$$

$$e^{-\frac{1}{24}x} = .5 \quad \text{solve for } x$$

$$\log(e^{-\frac{1}{24}x}) = \log(.5)$$

$$\text{log laws} \quad -\frac{1}{24}x = \log(.5) \Rightarrow x = \frac{\log(.5)}{-\frac{1}{24}} = \frac{\log(2)}{+\frac{1}{24}}$$

$$\log\left(\frac{1}{a}\right) = -\log a$$

Called half-life
or median of
exponential distribution.

Last time

Sec 10.3 Exponential Distributions

An exponential RV,

$T \sim \text{Exp}(\lambda)$, has density $f(t) = \lambda e^{-\lambda t}$.

It models the lifetime of objects that don't age, such as a light bulb.

The cdf of T is $F(t) = 1 - e^{-\lambda t}$

$$\curvearrowleft P(T \leq t)$$

The survival function of T is $S(t) = e^{-\lambda t}$

$$\curvearrowleft P(T > t)$$

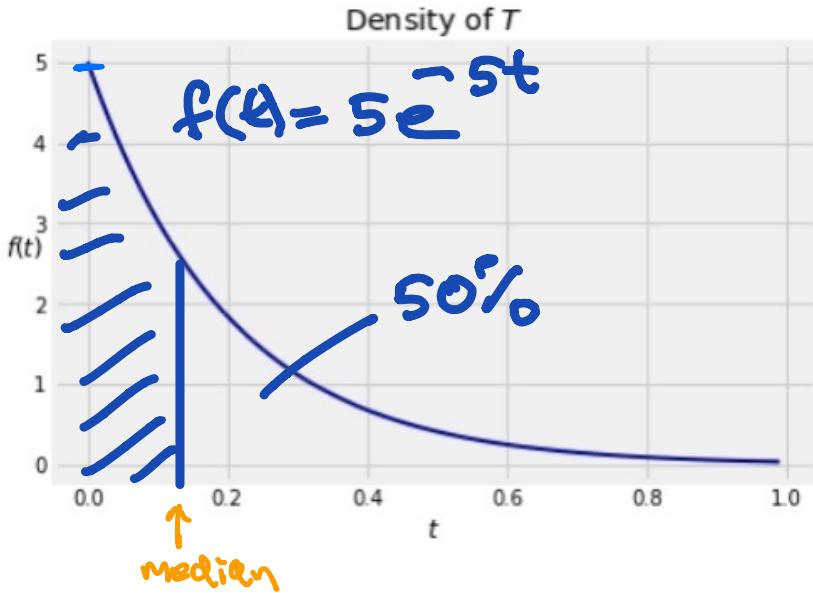
Today

- (1) Sec 10.3 half life of radioactive atoms
- (2) Sec 10.4 The normal distribution

① Sec 10.3 half life of radioactive atoms

Median

The median is the 50th percentile



Find t such that,

$$F(t) = .5 \quad \Leftrightarrow S(t) = \frac{1}{2}$$

$$\frac{1}{2} = S(t)$$

$$\Rightarrow \frac{1}{2} = e^{-\lambda t}$$

$$\Rightarrow \log\left(\frac{1}{2}\right) = -\lambda t$$

$$\Rightarrow -\log(2) = -\lambda t$$

$$\Rightarrow t = \frac{\log(2)}{\lambda}$$

half life
= median of exponential lifetime.

$$h = \frac{\log(2)}{\lambda}$$

We often model the time until a radioactive particle is emitted using the exponential distribution.

The half life, h , is the median of the exponential lifetime of an atom,

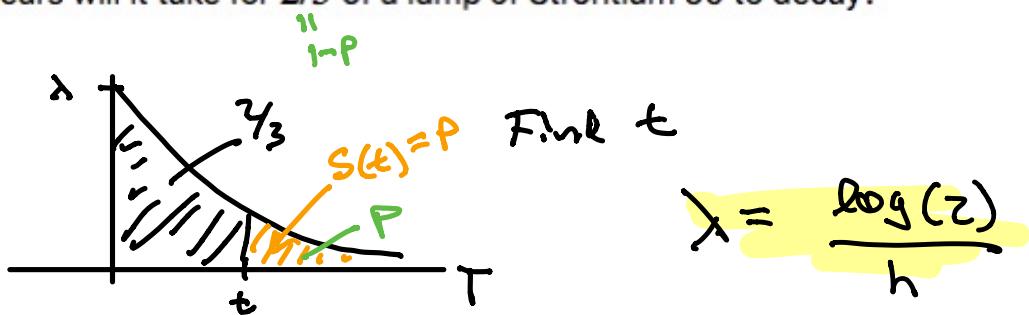
$$h = \frac{\log(2)}{\lambda}$$

defines the rate, λ , of decay.

$$\lambda = \frac{\log(2)}{h}$$

exercise 10.5.4

4. Strontium 90 has a half-life of 28.8 years. Assuming exponential decay, about how many years will it take for $2/3$ of a lump of Strontium 90 to decay?



$$\lambda = \frac{\log(2)}{h}$$

$$\lambda = \frac{\log(2)}{28.8} = .024$$

$$S(t) = e^{-\lambda t}$$

$$F(t) = \gamma_3 \Rightarrow S(t) = \gamma_3 \\ e^{-0.024t}$$

$$\Rightarrow -0.024t = \log(\frac{1}{3}) \\ t = \frac{\log(\frac{1}{3})}{-0.024}$$

For any proportion p of remaining material,

Find t such that $S(t) = p \rightarrow t = \frac{\log(p)}{-\lambda}$

radioCarbon dating

This is used to estimate the age of an object containing animal or plant material.

When an animal or plant dies, its radioCarbon ^{14}C decays exponentially with half life 5730 yrs.

If the proportion of ^{14}C to C is p , we can estimate when the animal or plant died.

$$t = \frac{-\log(p)}{\lambda} \quad \text{where } \lambda = \frac{\log(2)}{5730} \quad \begin{matrix} \text{half life} \\ \text{formula} \end{matrix}$$

$$\Rightarrow t = \frac{-\log(p)}{\log(2)/5730}$$

e.g. if the ratio of ^{14}C to C is $\frac{1}{3}$ (the proportion of remaining ^{14}C)

$$\frac{-\log(\frac{1}{3})}{\log(2)/5730}$$

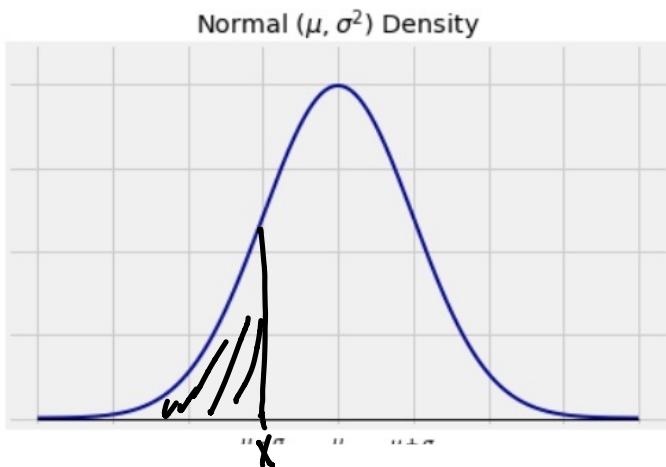
$$-np \cdot \log(0.3) / (\log(2) / 5730)$$

$$9952.812854572363$$

(2) Sec 10.4 The normal distribution

The random variable X has the *normal* (μ, σ^2) distribution if the density of X is given by

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}, \quad -\infty < x < \infty$$



To find
 $P(X < x) =$
 $P\left(\frac{X-\mu}{\sigma} < \frac{x-\mu}{\sigma}\right)$
 " "
 $Z \sim N(0,1)$

Sums of independent normal variables

Fact: the sum of two independent normal RVs is a normal RV.

$$\begin{aligned} X &\sim N(\mu_x, \sigma_x^2) \\ Y &\sim N(\mu_y, \sigma_y^2) \end{aligned} \quad \left. \begin{array}{l} \\ \end{array} \right\} \text{Indep}$$

$$X+Y \sim N(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2)$$

This result extends to linear combinations of independent normals,
 e.g. what distribution is

$$X - 2Y + 3 ?$$

$$N(\mu_x - 2\mu_y + 3, \sigma_x^2 + 4\sigma_y^2)$$

$$\begin{aligned} E(X - 2Y + 3) &= E(X) - 2E(Y) + 3 \\ &= \mu_x - 2\mu_y + 3 \end{aligned}$$

$$\begin{aligned} \text{Var}(X - 2Y + 3) &= \text{Var}(X) + 4\text{Var}(Y) \\ &= \sigma_x^2 + 4\sigma_y^2 \end{aligned}$$

Exercise 10.5.5

5. The weights of five randomly sampled people are i.i.d. normally distributed random variables with mean 150 pounds and SD 20 pounds. Let W be the total weight of the five people, measured in pounds. If possible, find w such that $P(W > w) = 0.05$. If this is not possible, explain why not.

$$X_1, X_2, X_3, X_4, X_5 \stackrel{iid}{\sim} N(150, 20^2)$$

$$W = X_1 + X_2 + X_3 + X_4 + X_5 \sim N(5(150), 5(20^2))$$

$$P(W > w) = .05$$

$$P(W < w) = .95$$

" " same as \leq

$$P(Z < \frac{w - 5(150)}{\sqrt{5}(20)})$$

$$\Rightarrow \frac{w - 5(150)}{\sqrt{5}(20)} = \Phi^{-1}(0.95) = 1.64$$

$$w = 5(150) + 1.64(\sqrt{5}(20)) = \boxed{823.34}$$

stats.norm.pdf (.95)

Confidence interval for the difference between means

- X_1, X_2, \dots, X_n are i.i.d. with mean μ_X and SD σ_X
- Y_1, Y_2, \dots, Y_m are i.i.d. with mean μ_Y and SD σ_Y

} indep

$\bar{X} - \bar{Y}$ is an unbiased estimator for $\mu_X - \mu_Y$

$$E(\bar{X} - \bar{Y}) = E(\bar{X}) - E(\bar{Y}) = \mu_X - \mu_Y \quad \checkmark$$

$$\text{Var}(\bar{X} - \bar{Y}) = \text{Var}(\bar{X}) + \text{Var}(\bar{Y}) \quad \text{since } \bar{X} \text{ and } \bar{Y} \text{ is indep.}$$

$$= \frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}$$

So a 95% CI for $\mu_x - \mu_y$ is

$$\bar{X} - \bar{Y} \pm 2 \sqrt{\frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}}$$

ex

As an example, suppose you have drawn samples of people independently from two cities, and suppose you have collected the following data.

- The incomes of the 400 sampled people in City X have an average of 70,000 dollars and an SD of 40,000 dollars.
- The incomes of the 600 sampled people in City Y have an average of 80,000 dollars and an SD of 50,000 dollars.

Find a 95% CI for $\mu_x - \mu_y$

$$70,000 - 80,000 \pm 2 \sqrt{\frac{40,000^2}{400} + \frac{50,000^2}{600}}$$

$$= -10,000 \pm 2(2858) = (-15,716, -4824)$$

Can you reject the null $\mu_x - \mu_y = 0$ (for alternative $\mu_x - \mu_y < 0$ at level 5%) — yes since 0 not in 95% CI.

Test for the equality of the means (A/B test)

We wish to determine if two independent populations have the same mean (i.e. $\mu_y - \mu_x = 0$)

pt

- The incomes of the 400 sampled people in City X have an average of 70,000 dollars and an SD of 40,000 dollars.
- The incomes of the 600 sampled people in City Y have an average of 80,000 dollars and an SD of 50,000 dollars.

H_0 : the mean income in City X is the same as the mean income in City Y ($\mu_X = \mu_Y$)

H_1 : $\mu_Y > \mu_X$

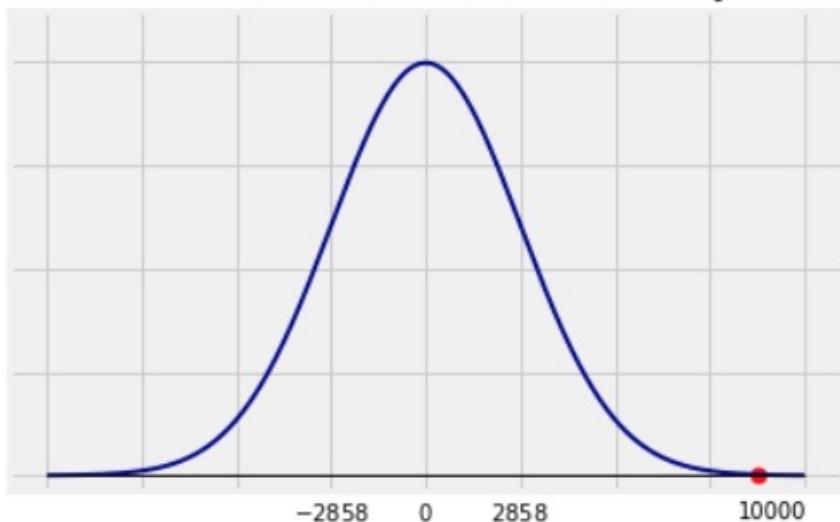
Our T.S. is $\bar{Y} - \bar{X} = 10,000$
we reject the null if $\bar{Y} - \bar{X}$ is large.

under H_0 :

$$\bar{Y} - \bar{X} \sim N(0, \frac{\sigma_x^2}{400} + \frac{\sigma_y^2}{600})$$

$\mu_Y - \mu_X = 0$ 2858^2

Distribution of Test Statistic Under H_0



$$P\text{-val} = 1 - \Phi\left(\frac{10,000 - 0}{2858}\right) = 1 - \Phi(3.5) = .0002 < .05$$

so reject null.

Exercise 10.9.7

7. A simple random sample of 200 students is taken at University A. Independently, a simple random sample of 300 students is taken at University B.

- The number of football games watched by students in Sample A has an average of 1.5 and an SD of 2. The number of football games watched by students in Sample B has an average of 4 and an SD of 1.5.

a) Construct an approximate 95% confidence interval for the difference between the average number of football games watched by students in the two universities.

$$(\bar{B} - \bar{A}) \pm 2 \sqrt{\frac{\sigma_B^2}{300} + \frac{\sigma_A^2}{200}}$$

$$(4 - 1.5) \pm 2 \sqrt{\frac{1.5^2}{300} + \frac{2^2}{200}} = (2.17, 2.83)$$

So we can reject the null hypothesis $\mu_B - \mu_A = 0$

since 0 not in our CI.