# EM

### Lieven Clement

### statOmics, Ghent University (https://statomics.github.io)

### Contents

1	Example of model based clustering with two groups			
2	Simulate toy Data	2		
3	Parameter estimation	4		
	B.1 EM algorithm	5		
4	Example	6		
	4.1 Initialize	6		
	1.2 EM algorithm	7		

# 1 Example of model based clustering with two groups

We will use a toy example to explain the EM algorithm for model based clustering.

- Two groups k = 1, 2
- Univariate observations are observed  $x_i$  with  $i = 1, \ldots, n$
- The data of group j follows a normal distribution  $N(u_j, 1)$  with a variance  $\sigma^2 = 1$  and a mean  $\mu_j$  that depends on the group j = 1, 2.

The data follows the following mixture distribution:

$$f(x) = \pi_1 f_1(x) + (1 - \pi_1) f_2(x)$$

With

- $\pi_1$  the probability that a random sample from the population belongs to group 1
- $f_1(x)$  the density of the data in group 1, i.e.  $N(\mu_1, 1)$
- $f_2(x)$  the density of the data in group 2, i.e.  $N(\mu_2, 1)$ .
- Unknowns? Group membership  $z_i$ , the group means  $\mu_j$  and the proportion of subjects in the population of group 1  $\pi_1$  are unknown.

We can estimate the model parameters  $\boldsymbol{\theta} = [\mu_1, \mu_2, \pi_1]^T$  using maximum likelihood.

$$L(\mu_1, \mu_2, \pi_1 | \mathbf{X}) = \prod_{i=1}^{n} [\pi_1 f_1(x_i) + (1 - \pi_1) f_2(x_i)]$$

It is easier to maximise the log-likelihood:

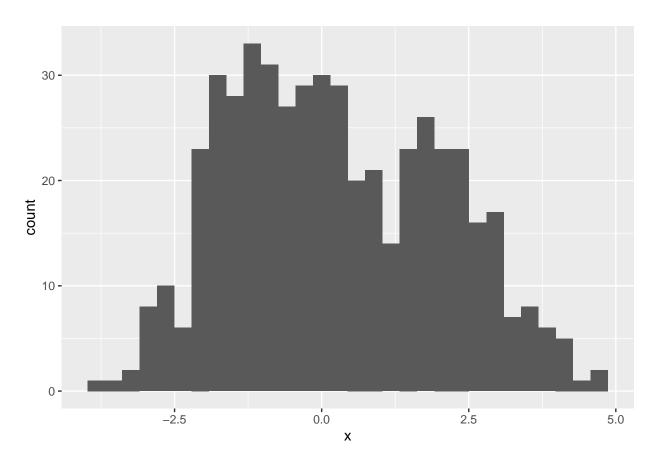
$$l(\mu_1, \mu_2, \pi_1 | \mathbf{X}) = \sum_{i=1}^n \log \left[ \pi_1 f_1(x_i) + (1 - \pi_1) f_2(x_i) \right]$$

The optimisation is difficult because we have to take the log of a sum and the likelihood does not factorise further.

## 2 Simulate toy Data

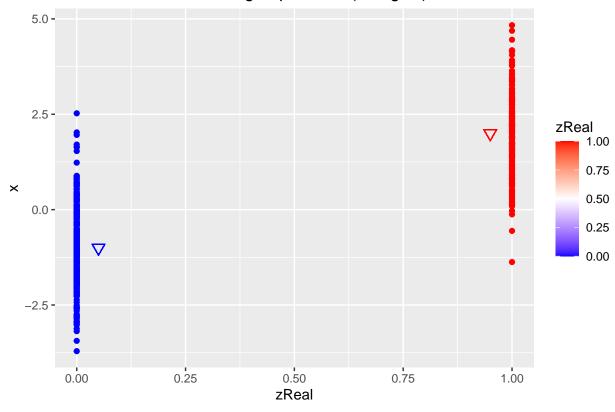
```
library(tidyverse)
mu1Real <- 2
mu2Real <- -1
pi1Real = .4
set.seed(114)
zReal <- rbinom(500, size = 1, prob = pi1Real)
x <- ifelse(zReal==1, rnorm(500, mean=mu1Real), rnorm(500, mean=mu2Real))
data.frame(x) %>%
    ggplot(aes(x = x)) +
    geom_histogram()
```

## 'stat\_bin()' using 'bins = 30'. Pick better value with 'binwidth'.



```
data.frame(x,zReal) %>%
    ggplot(aes(zReal,x,color=zReal)) +
    geom_point() +
    scale_colour_gradient2(
    low = "blue",
    mid="white",
    high="red",
    midpoint = 0.5) +
    geom_point(x=.95,y=mu1Real, shape=25, col="red", size=3) +
    geom_point(x=0.05,y=mu2Real, shape=25, col="blue", size=3) +
    ggtitle("Simulated data with real group means (triangles)")
```

## Simulated data with real group means (triangles)



## 3 Parameter estimation

If we would know the cluster membership  $z_{i1}$ 

$$z_{i1} = \begin{cases} 1 & \text{if } x_i \text{ belongs to group 1} \\ 0 & \text{if } x_i \text{ belongs to group 2} \end{cases}$$

with  $z_{i1}$  follows a Bernoulli distribution

$$B(\pi_1) = \pi_1^{z_{i1}} (1 - \pi_1)^{(1 - z_{i1})}$$

Then the density of  $x_i$  given  $z_{i1}$  becomes

$$f(x_i|z_{i1}) = f_1(x_i)^{z_{i1}} f_2(x_i)^{1-z_{i1}}$$

and the joint distribution of  $z_{i1}$  and  $x_i$  then becomes

$$f(x_i, z_{i1}) = f(x_i|z_{i1})f(z_{i1})$$

$$= f_1(x_i)^{z_{i1}}f_2(x_i)^{1-z_{i1}}\pi_1^{z_{i1}}(1-\pi_1)^{(1-z_{ik})}$$
(2)

and the log likelihood of the complete data becomes

$$l(\mu_1, \mu_2, \pi_1 | \mathbf{X}, \mathbf{Z}) = \sum_{i=1}^{n} z_{i1} \log [\pi_1 f_1(x_i)] + (1 - z_{i1}) \log [(1 - \pi_1) f_2(x_i)]$$

• Note, that in the notation of the Frayley and Raftery (1998)  $z_2 = 1 - z_1$  and  $\pi_2 = 1 - \pi_1$ 

### 3.1 EM algorithm

If we would know the cluster membership, we could estimate all model parameters based on the complete likelihood.

Note, that the complete likelihood also factorises nicely!

However, the cluster membership is unknown.

The EM algorithm has been developed for missing data problems.

It is an iterative algorithm that consists of two steps:

- 1. E-step: calculate the expected log likelihood given the data and the current model parameter estimates
- 2. M-step: maximise the expected log-likelihood
- 3. Iterate between 1 and 2 until convergence.

#### 3.1.1 E-step

In iteration m+1

$$Q(\mu_{1}, \mu_{2}, \pi_{1} | \mathbf{X}, \mathbf{Z}) = E[l(\mu_{1}, \mu_{2}, \pi_{1} | \mathbf{X}, \mathbf{Z}) | \mathbf{X}, \mu_{1} = \mu_{1}^{m}, \mu_{2} = \mu_{2}^{m}, \pi_{1} = \pi_{1}^{m}]$$

$$= E\left[\sum_{i=1}^{n} z_{i1} \log [\pi_{1} f_{1}(x_{i})] + (1 - z_{i1}) \log [(1 - \pi_{1}) f_{2}(x_{i})] | \mathbf{X}, \mu_{1} = \mu_{1}^{m}, \mu_{2} = \mu_{2}^{m}, \pi_{1} = \pi_{1}^{m}\right]$$

$$= \sum_{i=1}^{n} E[z_{i1} | x_{i}, \mu_{1} = \mu_{1}^{m}, \mu_{2} = \mu_{2}^{m}, \pi_{1} = \pi_{1}^{m}] \log [\pi_{1} f_{1}(x_{i})] + \sum_{i=1}^{n} (1 - E[z_{i1} | x_{i}, \mu_{1} = \mu_{1}^{m}, \mu_{2} = \mu_{2}^{m}, \pi_{1} = \pi_{1}^{m}]$$

Note, that the expected log likelihood simplifies to replacing the unknown class memberships in the complete likelihood by their expected values.

$$E[z_{i1}|\mathbf{X}, \mu_1 = \mu_1^m, \mu_2 = \mu_2^m, \pi_1 = \pi_1^m] = 1 \times f(z_{i1} = 1|x_i, \mu_1 = \mu_1^m, \mu_2 = \mu_2^m, \pi_1 = \pi_1^m) + 0 \times f(z_{i1} = 0|x_i, \mu_1 = \mu_1^m, \mu_2 = f(z_{i1} = 1|x_i, \mu_1 = \mu_1^m, \mu_2 = \mu_2^m, \pi_1 = \pi_1^m)$$

$$= \frac{f(z = 1, x_i)}{f(x_i)}$$

We will refer to the expected class membership as  $\hat{z}_{i1}^{m+1}$ 

 $= \frac{\pi_1^m f_1(x_i)}{\pi_1^m f_1(x_i) + (1 - \pi_1^m) f_2(x_i)}$ 

#### 3.1.2 M-step

Maximise the expected log-likelihood to obtain the unknown model parameters:

$$Q(\mu_1, \mu_2, \pi_1 | \mathbf{X}, \mathbf{Z}) = \sum_{i=1}^{n} \hat{z}_{i1}^{m} \log \pi_1 + \sum_{i=1}^{n} \hat{z}_{i1}^{m} \log f_1(x_i) + \sum_{i=1}^{n} (1 - \hat{z}_{i1}) \log(1 - \pi_1) + \sum_{i=1}^{n} (1 - \hat{z}_{i1}) \log f_2(x_i)$$

So we observe that the expected log likelihood implies an estimation orthogonality between the normal distributions and the Bernoulli distribution.

So the parameter estimates that maximise the expected log-likelihood become:

$$\begin{split} \hat{\pi}_1^{m+1} &= \frac{\sum_{i=1}^n \hat{z}_{i1}^{m+1}}{n} \\ \hat{\mu}_1^{m+1} &= \frac{\sum_{i=1}^n \hat{z}_{i1}^{m+1} x_i}{\sum_{i=1}^n \hat{z}_{i1}^{m+1}} \\ \hat{\mu}_2^{m+1} &= \frac{\sum_{i=1}^n (1 - \hat{z}_{i1}^{m+1}) x_i}{\sum_{i=1}^n (1 - \hat{z}_{i1}^{m+1})} \end{split}$$

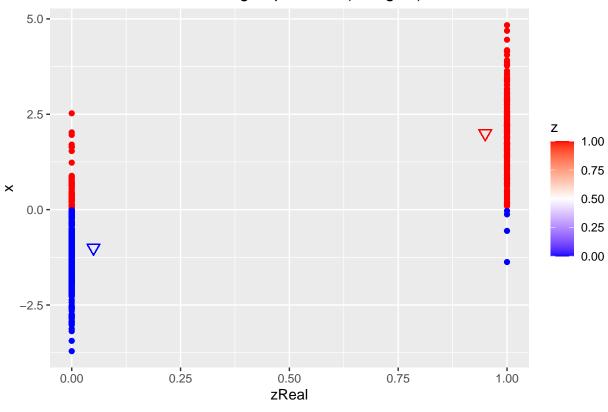
## 4 Example

### 4.1 Initialize

```
z <- as.double(x > 0)
n <- length(z)

p <- data.frame(x,zReal,z) %>%
    ggplot(aes(zReal,x,color=z)) +
    geom_point() +
    scale_colour_gradient2(
    low = "blue",
        mid="white",
        high="red",
        midpoint = 0.5) +
    geom_point(x=.95,y=mu1Real, shape=25, col="red", size=3) +
    geom_point(x=0.05,y=mu2Real, shape=25, col="blue", size=3) +
    ggtitle("Simulated data with real group means (triangles)")
```

## Simulated data with real group means (triangles)

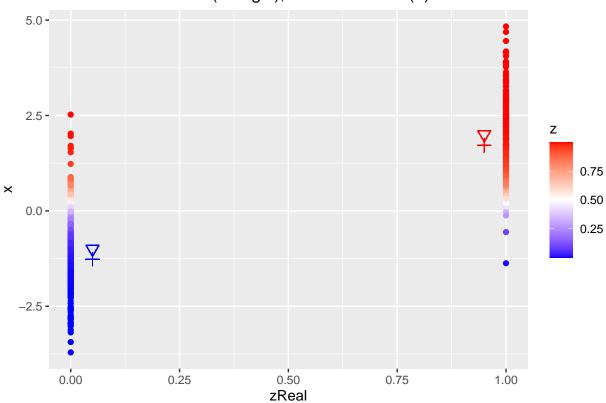


### 4.2 EM algorithm

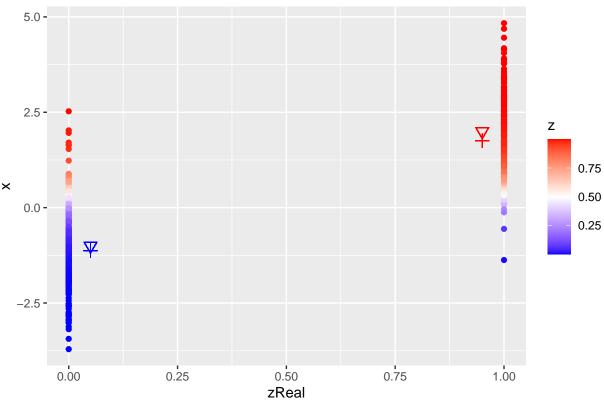
```
for (k in 1:10)
  ## M-step
  pi1 \leftarrow sum(z)/n
  mu1 \leftarrow sum(z * x)/sum(z)
  mu2 <- sum((1-z) * x)/sum(1-z)
  ## E-step
  d1 <- dnorm(x,mean=mu1)</pre>
  d2 <- dnorm(x,mean=mu2)</pre>
  d \leftarrow pi1 * d1 + (1-pi1)*d2
  z \leftarrow pi1*d1/d
  p <- data.frame(x,zReal,z) %>%
    ggplot(aes(zReal,x,color=z)) +
    geom_point() +
    scale_colour_gradient2(
      low = "blue",
      mid="white",
      high="red",
      midpoint = 0.5) +
      geom_point(x=.95,y=mu1Real, shape=25, col="red", size=3) +
      geom_point(x=0.05,y=mu2Real, shape=25, col="blue", size=3) +
```

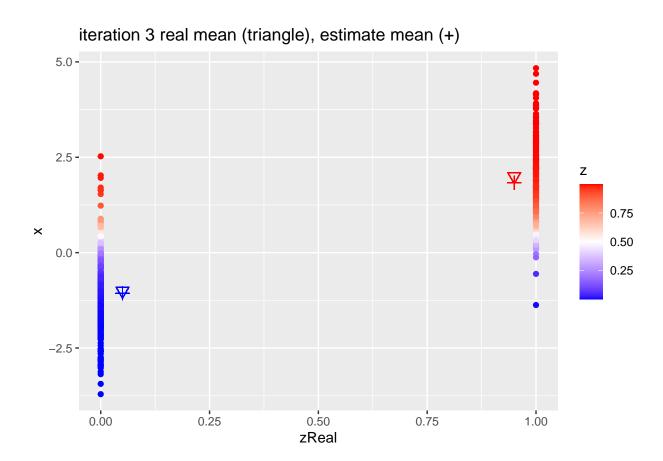
```
geom_point(x=.95,y=mu1, shape=3, col="red", size=3) +
    geom_point(x=.05,y=mu2, shape=3, col="blue", size=3) +
    ggtitle(paste0("iteration ",k, " real mean (triangle), estimate mean (+)"))
print(p)
}
```

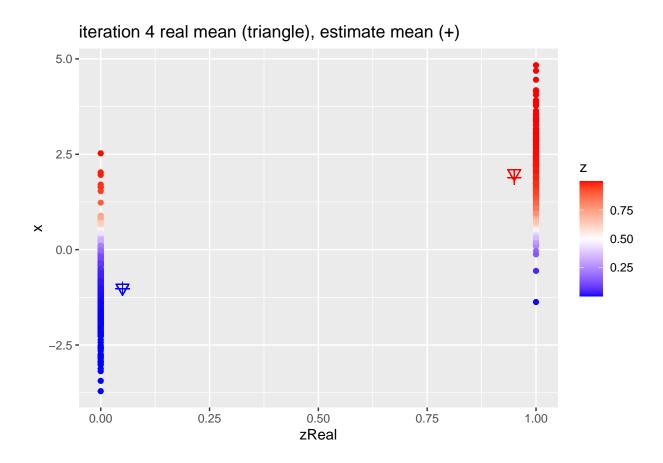
# iteration 1 real mean (triangle), estimate mean (+)

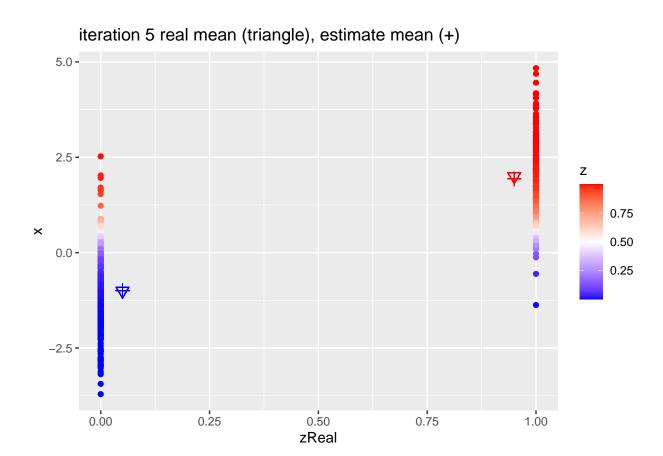


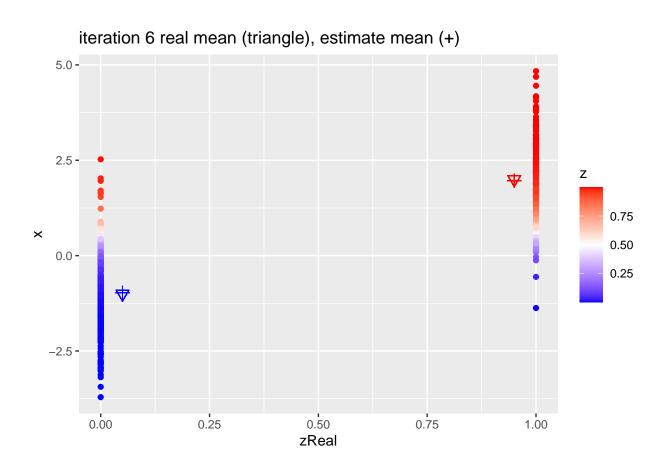


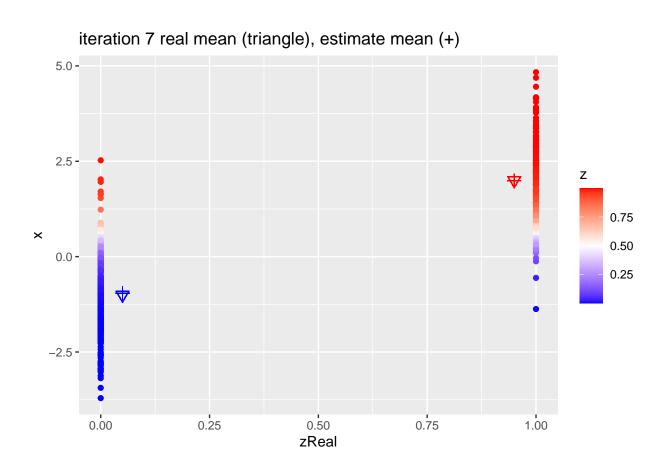


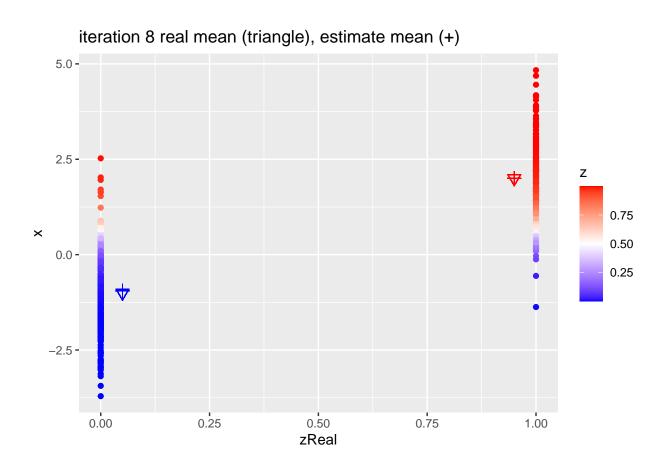


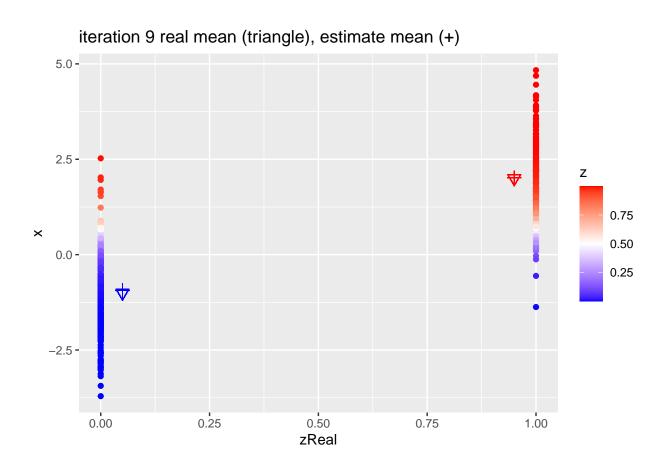




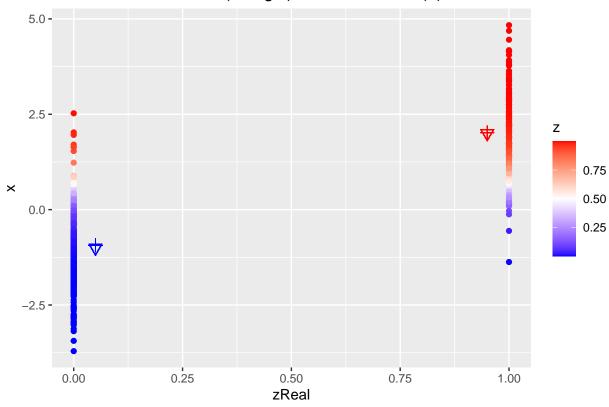












## 4.2.1 Estimates

parameter	population	initial	estimate
mu1	2.0	1.715	2.020
mu2	-1.0	-1.270	-0.935
pi1	0.4	0.512	0.404