

STAT 5361 Interim Report

A Simple and Effective Inequality Measure

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19 11 2020

Abstract

We'll write it later

Introduction

- 1) Examples of inequality curves and coefficients of inequality for some common income distributions and compare values of I with G.
- 2) We introduce empirical versions of these concepts and investigate their inferential properties, including robustness to outliers.
- 3) Applications to income data.

4. Application

4.1. Disposable Personal Income

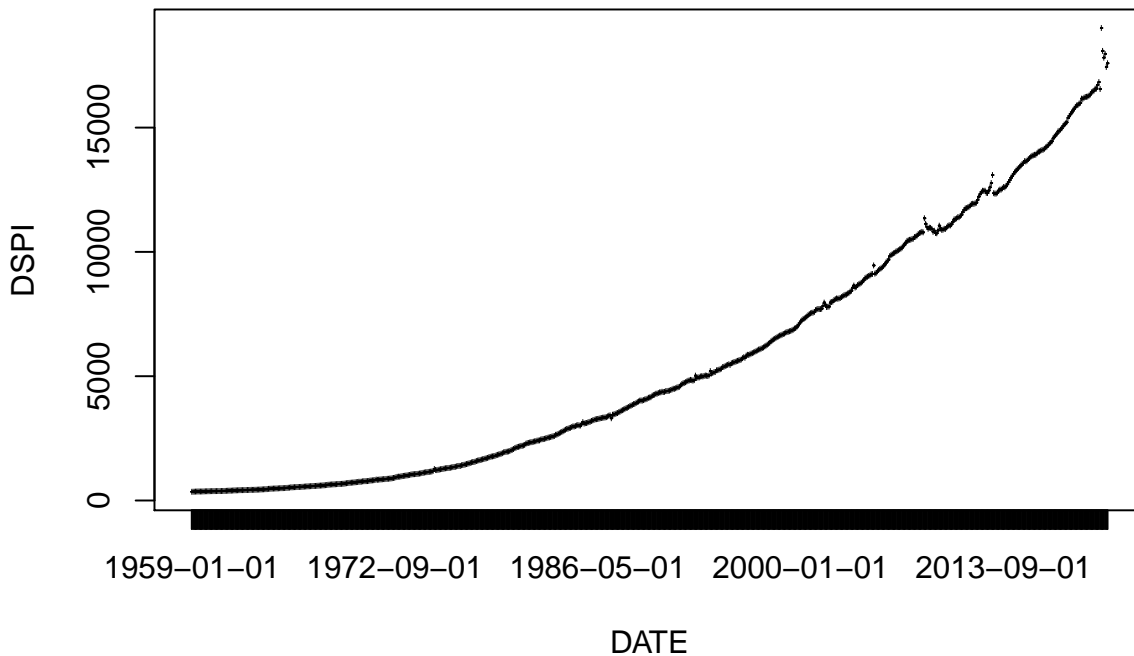
The data of Disposable Personal Income can be obtained at <https://fred.stlouisfed.org/series/DSPI>.

```
DSPI <- read.csv("C:\\Users\\Asus\\Downloads\\DSPI.csv")
plot(DSPI, type = "l", main = "Disposable Personal Income")
```

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Disposable Personal Income



```
QOR.ln <- function(u){
  # QOR function for the log-normal
  q.n.0 <- 1/dnorm(qnorm(u))
  q.n.1 <- qnorm(u)*q.n.0^2
  q.n.2 <- (1 + 2*qnorm(u)^2)*q.n.0^3
  1/(q.n.0^2 + 3*q.n.1 + q.n.2/q.n.0)
}

QOR.gld <- function(u, lambda = NULL){
  # QOR function for the GLD FKML parameterization
  if(is.null(lambda)) lambda <- fit.fkml(x)$lambda
  l3 <- lambda[3]
  l4 <- lambda[4]
  (u^(l3 - 1) + (1 - u)^(l4 - 1))/(u^(l3 - 3)*(l3 - 2)*(l3 - 1) +
    (1 - u)^(l4 - 3)*(l4 - 2)*(l4 - 1))
}

hatI <- function(x, J = 100, conf.level = 0.95, bw.correct = TRUE, QOR.FUN = QOR.ln, ...){
  n <- length(x)

  us <- ((1:J) - 0.5)/J
  Rs <- (xu2 <- quantile(x, us/2))/(x1u2 <- quantile(x, 1 - us/2))
}
```

```

I <- sum(1 - Rs)/J

if(!is.null(conf.level)){
  v <- c(us/2, 1 - us/2)
  qor <- QOR.FUN(v, ...)
  bw <- 15^(1/5)*abs(qor)^(2/5)/n^(1/5)
  if (bw.correct) bw[v <= bw] <- v[v <= bw]

  kernepach <- function(u) 3/4*(1 - u^2)*(abs(u) <= 1)
  m1 <- matrix(v, nrow = 2*J, ncol = n, byrow = FALSE)
  m2 <- matrix(1:n, nrow = 2*J, ncol = n, byrow = TRUE)

  consts <- kernepach((m1 - (m2 - 1)/n)*(1/bw))*(1/bw) -
    kernepach((m1 - m2/n)*(1/bw))*(1/bw)

  x.sorted <- sort(x)
  q.hat <- c(consts%%x.sorted)
  q.hat.1 <- q.hat[1:(length(q.hat)/2)]
  q.hat.2 <- q.hat[-(1:(length(q.hat)/2))]

  rc <- matrix(Rs, ncol = J, nrow = J, byrow = FALSE)

  covm <- ((1/x1u2)%%t(1/x1u2))*(((us/2)%%t(1 - us/2))*(q.hat.1%%t(q.hat.1) + Rs%%t(Rs)*
    ((us/2)%%t(us/2))*((q.hat.1%%t(q.hat.2))*t(rc) + (q.ha

  sigma.p2 <- (us/2)*(1 - us/2)*q.hat.1^2
  sigma.q2 <- (1 - us/2)*(us/2)*q.hat.2^2
  sigma.pq <- (us/2)^2*q.hat.1*q.hat.2
  a0 <- sigma.p2/x1u2^2
  a1 <- -2*sigma.pq/x1u2^2
  a2 <- sigma.q2/x1u2^2
  Vs <- (a0 + a1*Rs + a2*Rs^2)/n

  V <- (sum(Vs) + 2*sum(covm[row(covm) < col(covm)]))/J^2
  SE <- sqrt(V)
  conf.int <- I + c(-1, 1)*qnorm(1 - (1 - conf.level)/2)*sqrt(V)
} else{
  V <- NULL
  SE <- NULL
  conf.int <- NULL
}

list(I = I, SE = SE, conf.int = conf.int)
}

hatG <- function(x, conf.level = 0.95){

```

```

indices <- 1:(n <- length(x))
ordered.x <- sort(x)
sx <- sum(ordered.x*(indices - 1/2))
mu.hat <- mean(x)
Gv <- 2/mu.hat/n^2*sx - 1

Z.hat <- -(Gv + 1)*ordered.x + (2*indices - 1)/n*ordered.x - 2/n*cumsum(ordered.x)
Z.bar <- mean(Z.hat)

V <- 1/n^2/mu.hat^2*sum((Z.hat - Z.bar)^2)
conf.int <- Gv + c(-1, 1)*qnorm(1 - (1 - conf.level)/2)*sqrt(V)

list(G = Gv, SE = sqrt(V), conf.int = conf.int)
}

```

```

DSPI <- DSPI$DSPI
I <- hatI(DSPI)
I

```

```

## $I
## [1] 0.7560689
##
## $SE
## [1] 0.01064179
##
## $conf.int
## [1] 0.7352114 0.7769264

```

```

G <- hatG(DSPI)
G

```

```

## $G
## [1] 0.487252
##
## $SE
## [1] 0.009500346
##
## $conf.int
## [1] 0.4686317 0.5058723

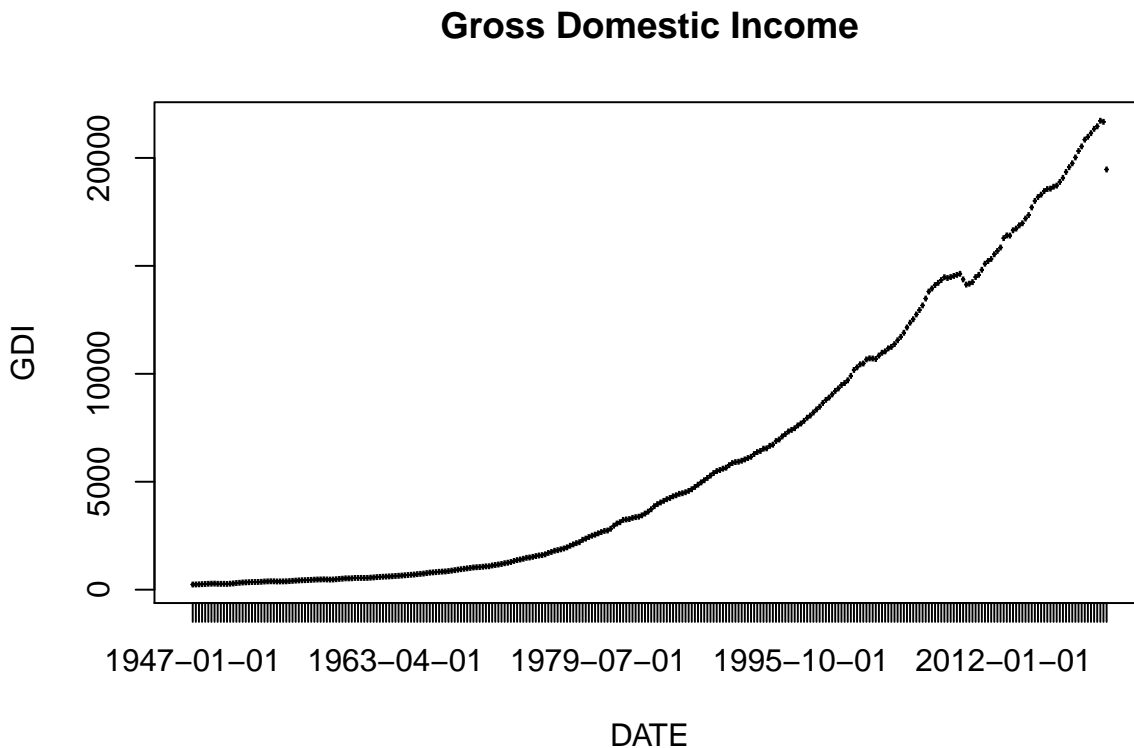
```

Firstly, I and G were calculated for the Disposable Personal Income Data. It appears that $\hat{I} = 0.7561$, which is an estimate of I , with a standard error of 0.0106. Additionally, 95% confidence interval of \hat{I} was calculated to be (0.7352, 0.7769). While $\hat{G} = 0.4872$ is significantly lower, with a standard error of 0.0095 and respective 95% confidence interval of (0.4686, 0.5059).

4.2. Gross Domestic Income

The data on Gross Domestic Income can be found here: <https://fred.stlouisfed.org/series/GDI>.

```
GDI <- read.csv("C:\\Users\\Asus\\Downloads\\GDI.csv")
plot(GDI, type = "l", main = "Gross Domestic Income")
```



```
GDI <- GDI$GDI
I_gdi <- hatI(GDI)
I_gdi
```

```
## $I
## [1] 0.8204182
##
## $SE
## [1] 0.01624378
##
## $conf.int
## [1] 0.7885809 0.8522554
```

```
G_gdi <- hatG(GDI)
G_gdi
```

```
## $G
## [1] 0.549877
##
## $SE
## [1] 0.01613956
```

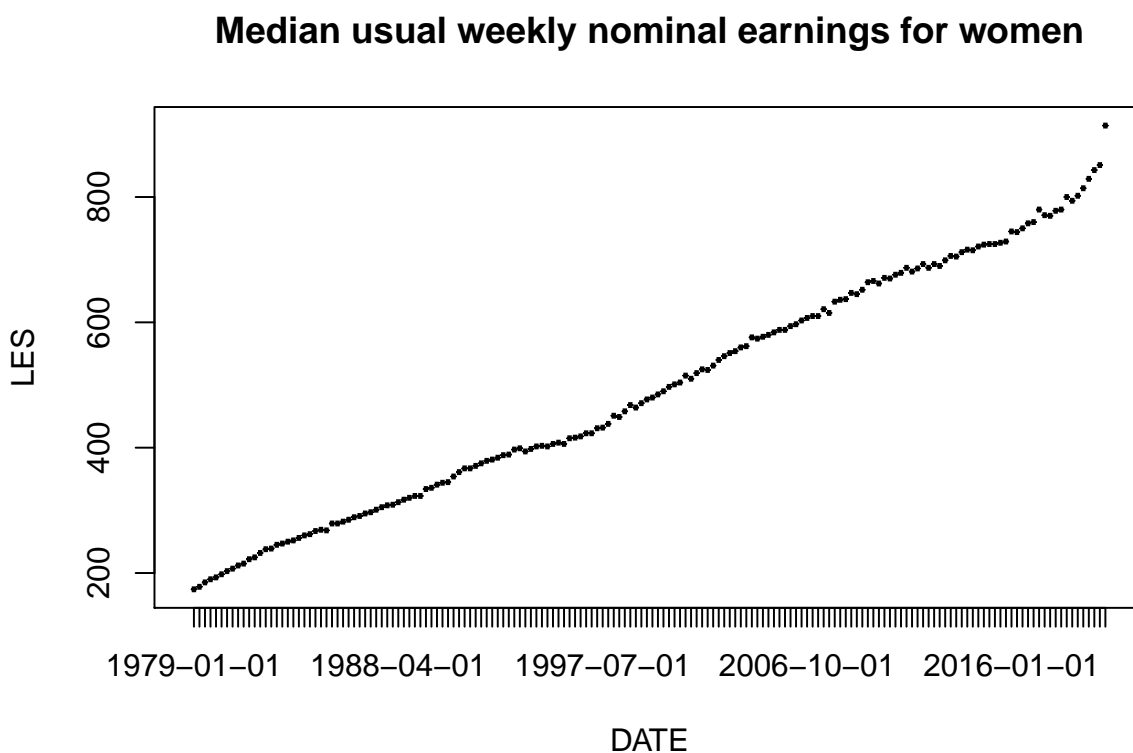
```
##
## $conf.int
## [1] 0.5182440 0.5815099
```

As a result, ratio of symmetric quantiles of the GDI is $\hat{I} = 0.8204$ with a standard error 0.0162 and confidence interval (0.7886, 0.8523). At the same time $\hat{G} = 0.5499$ with a standard error of 0.0161 and CI (0.5182, 0.5815).

4.3. Earnings data of women

The data on Earnings of women in the US between 1979 and 2020 can be obtained from: <https://fred.stlouisfed.org/series/LES1252882700Q>.

```
LES <- read.csv("C:\\Users\\Asus\\Downloads\\LES.csv")
plot(LES, type = "l", main = "Median usual weekly nominal earnings for women")
```



```
LES <- LES$LES
I_les <- hatI(LES)
I_les
```

```
## $I
## [1] 0.4665993
##
## $SE
## [1] 0.01831842
```

```
##  
## $conf.int  
## [1] 0.4306959 0.5025028
```

```
G_les <- hatG(LES)  
G_les
```

```
## $G  
## [1] 0.2199384  
##  
## $SE  
## [1] 0.009415175  
##  
## $conf.int  
## [1] 0.2014850 0.2383918
```

Results reveal that $\hat{I} = 0.4666$ with a standard error 0.0183 and CI (0.4307, 0.5025). While Gini coefficient is $\hat{G} = 0.2199 \pm 0.0094$ with CI (0.2015, 0.2384).