

# Linear Models for Regression

CS534

Adapted from slides of C. Bishop

# Prediction Problems

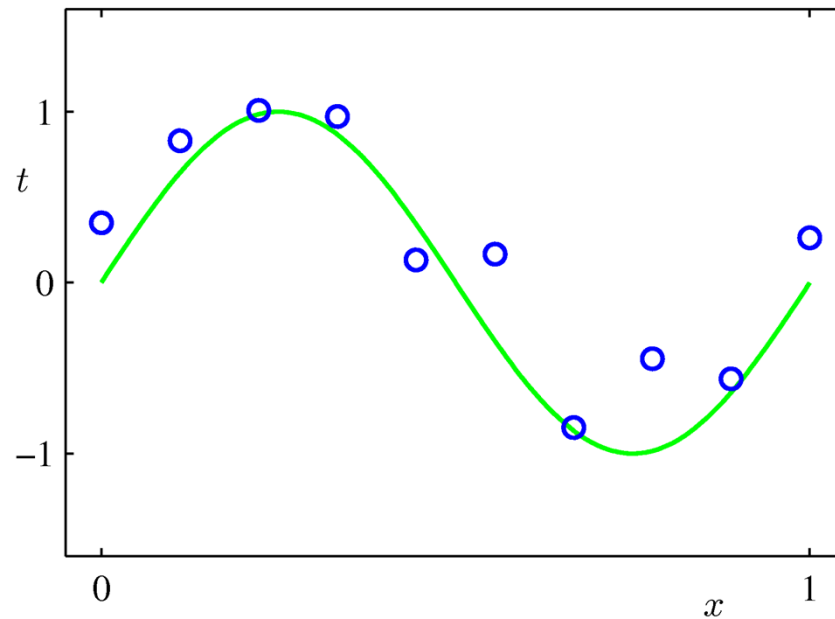
- Predict housing price based on
  - House size, lot size, Location, # of rooms ...
- Predict stock price based on
  - Price history of the past month ...
- Predict the abundance of a species based on
  - Environmental conditions

- General set up:

*Given a set of training examples  $(\mathbf{x}_i, t_i)$ ,  $i = 1, \dots, N$*

*Goal: learn a function  $\hat{y}(\mathbf{x})$  to minimize some loss function:  $L(\hat{y}, t)$*

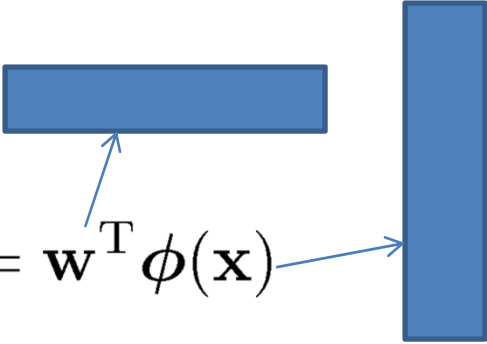
- Example: Polynomial Curve Fitting



$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + \dots + w_Mx^M = \sum_{j=0}^M w_jx^j$$

# Linear Basis Function Models (1)

- Generally

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$$


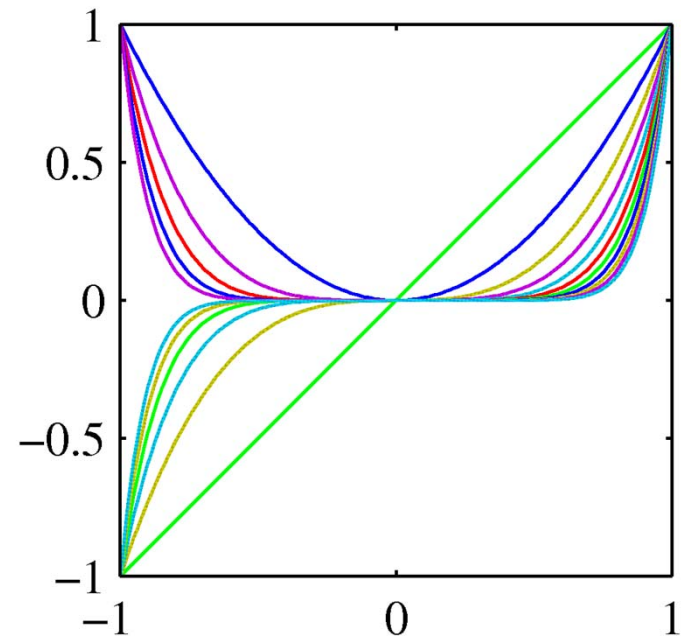
- where  $\phi_j$ 's are known as ***basis functions***.
- Typically  $\phi_0 = 1$  so that  $w_0$  acts as a ***bias***.
- In the simplest case, we use linear basis functions :  $\phi_i(\mathbf{x}) = x_i$ 
  - Multiple linear regression

# Linear Basis Function Models (2)

- Polynomial basis functions:

$$\phi_j(x) = x^j.$$

- These are global; a small change in  $x$  affects all basis functions.

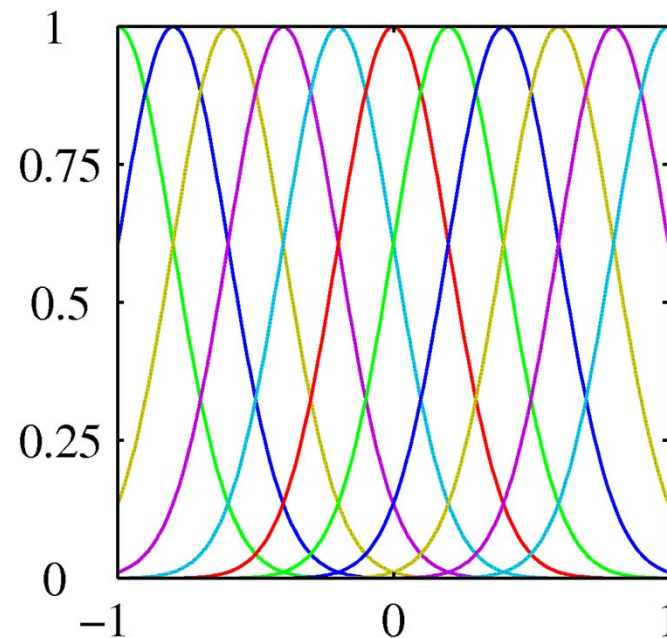


# Linear Basis Function Models (3)

- Gaussian basis functions:

$$\phi_j(x) = \exp \left\{ -\frac{(x - \mu_j)^2}{2s^2} \right\}$$

- These are local; a small change in  $x$  only affect nearby basis functions.  $\mu_j$  and  $s$  control the location and scale (width).



# Linear Basis Function Models (4)

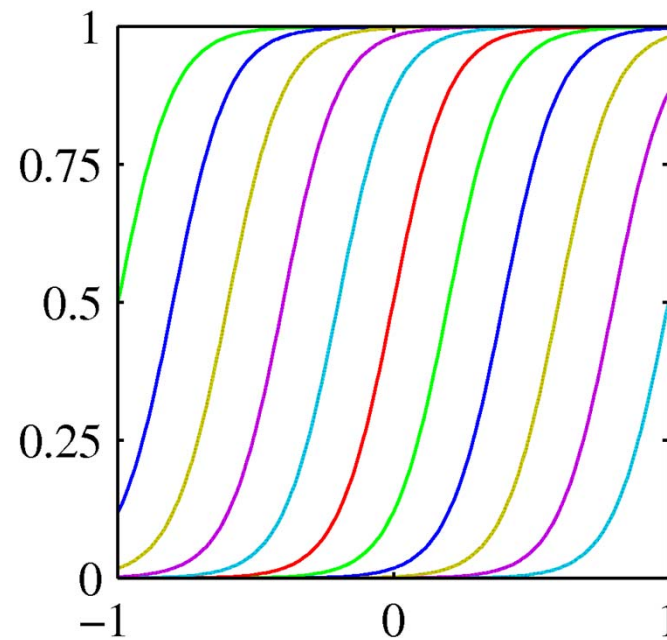
- Sigmoidal basis functions:

$$\phi_j(x) = \sigma \left( \frac{x - \mu_j}{s} \right)$$

- where

$$\sigma(a) = \frac{1}{1 + \exp(-a)}.$$

- Also these are local; a small change in  $x$  only affect nearby basis functions.  $\mu_j$  and  $s$  control location and scale (slope).



# Maximum Likelihood Estimation of $\mathbf{w}$

- **Assumption:** observations drawn from a deterministic function with added Gaussian noise:

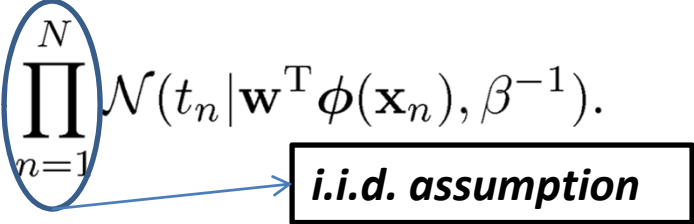
$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon \quad \text{where} \quad p(\epsilon|\beta) = \mathcal{N}(\epsilon|0, \beta^{-1})$$

- which is the same as saying,

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|y(\mathbf{x}, \mathbf{w}), \beta^{-1}) = \sqrt{\frac{\beta}{2\pi}} \exp\left\{-\frac{1}{2} \beta (t - y(\mathbf{x}, \mathbf{w}))^2\right\}$$

- Given a set of observed inputs,  $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$ , and their corresponding targets,  $\mathbf{t} = [t_1, \dots, t_N]^T$ , if we assume that the parameters take specific values  $\mathbf{w}$  and  $\beta$ , the likelihood of observing the data is:

$$p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \phi(\mathbf{x}_n), \beta^{-1}).$$

 ***i.i.d. assumption***



# Maximum Likelihood Estimation

- Taking the logarithm of the likelihood ftn, we get

$$\begin{aligned}\ln p(\mathbf{t}|\mathbf{w},\beta) &= \sum_{n=1}^N \ln N(t_n|\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1}) \\ &= \frac{N}{2} \ln \frac{\beta}{2\pi} - \frac{\beta}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n)\}^2\end{aligned}$$

- Note that

$$\operatorname{argmax}_{\mathbf{w}} \ln p(\mathbf{t}|\mathbf{w},\beta) = \operatorname{argmin}_{\mathbf{w}} E_D(\mathbf{w})$$

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2$$

- $E_D(\mathbf{w})$  is called the least -square objective.
- Maximizing likelihood = least squares

# Maximum Likelihood Estimation

- Computing the gradient and setting it to zero yields

$$\nabla_{\mathbf{w}} \ln p(\mathbf{t}|\mathbf{w}, \beta) = \beta \sum_{n=1}^N \{t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n)\} \boldsymbol{\phi}(\mathbf{x}_n)^T = \mathbf{0}.$$

- Solving for  $\mathbf{w}$ , we get

$$\mathbf{w}_{\text{ML}} = \left( \boldsymbol{\Phi}^T \boldsymbol{\Phi} \right)^{-1} \boldsymbol{\Phi}^T \mathbf{t}$$

The Moore-Penrose pseudo-inverse,  $\boldsymbol{\Phi}^\dagger$ .

- where

$$\boldsymbol{\Phi} = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \cdots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \cdots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \cdots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix}.$$


A single example

$N \times M$

A basis function

# Maximum Likelihood and Least Squares (4)

- Maximizing with respect to the bias,  $w_0$ , alone, we see that

$$\begin{aligned} w_0 &= \bar{t} - \sum_{j=1}^{M-1} w_j \bar{\phi_j} \\ &= \frac{1}{N} \sum_{n=1}^N t_n - \sum_{j=1}^{M-1} w_j \frac{1}{N} \sum_{n=1}^N \phi_j(\mathbf{x}_n). \end{aligned}$$


- We can also maximize with respect to  $\beta$ , giving

$$\frac{1}{\beta_{\text{ML}}} = \frac{1}{N} \sum_{n=1}^N \{t_n - \mathbf{w}_{\text{ML}}^T \boldsymbol{\phi}(\mathbf{x}_n)\}^2$$

# System Equation View of Linear Regression

$$\mathbf{t} = \begin{bmatrix} t_1 \\ \dots \\ t_N \end{bmatrix} \quad \boldsymbol{\phi} = \begin{bmatrix} \phi_0(\mathbf{x}_1) & \dots & \phi_{m-1}(\mathbf{x}_1) \\ \dots & \dots & \dots \\ \phi_0(\mathbf{x}_N) & \dots & \phi_{m-1}(\mathbf{x}_N) \end{bmatrix} \quad \mathbf{w} = \begin{bmatrix} w_0 \\ \dots \\ w_{m-1} \end{bmatrix}$$

$$\mathbf{t} = \boldsymbol{\phi} \mathbf{w}$$

- Over-constrained system of equations
- There exists no solution
- Maximum likelihood and Least squared solution

$$\mathbf{w}_{\text{ML}} = \left( \boldsymbol{\Phi}^T \boldsymbol{\Phi} \right)^{-1} \boldsymbol{\Phi}^T \mathbf{t}$$

# Geometry of Least Squares

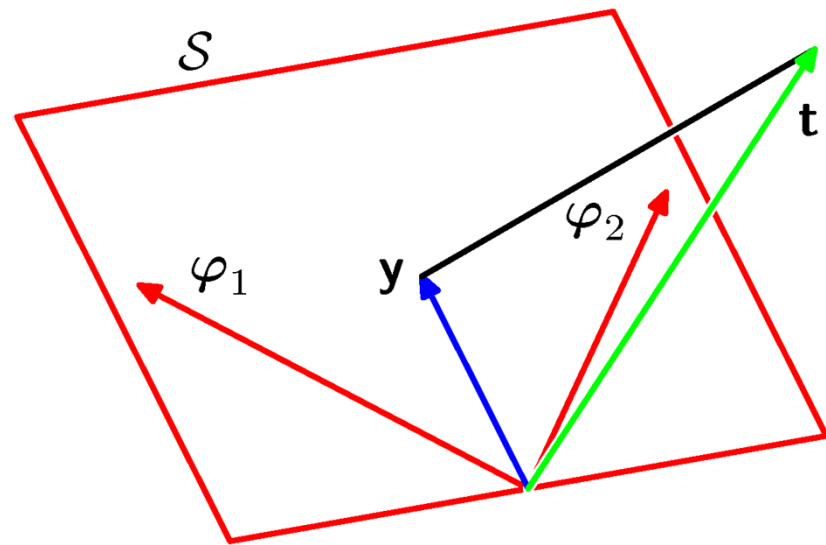
- Consider

$$\mathbf{y} = \boldsymbol{\phi} \mathbf{w}_{ML} = [\phi_0 \quad \dots \quad \phi_{m-1}] \mathbf{w}_{ML} = w_0 \phi_0 + \dots + w_{m-1} \phi_{m-1}$$

$$\mathbf{y} \in \mathcal{S} \subseteq \mathcal{T} \quad \mathbf{t} \in \mathcal{T}$$

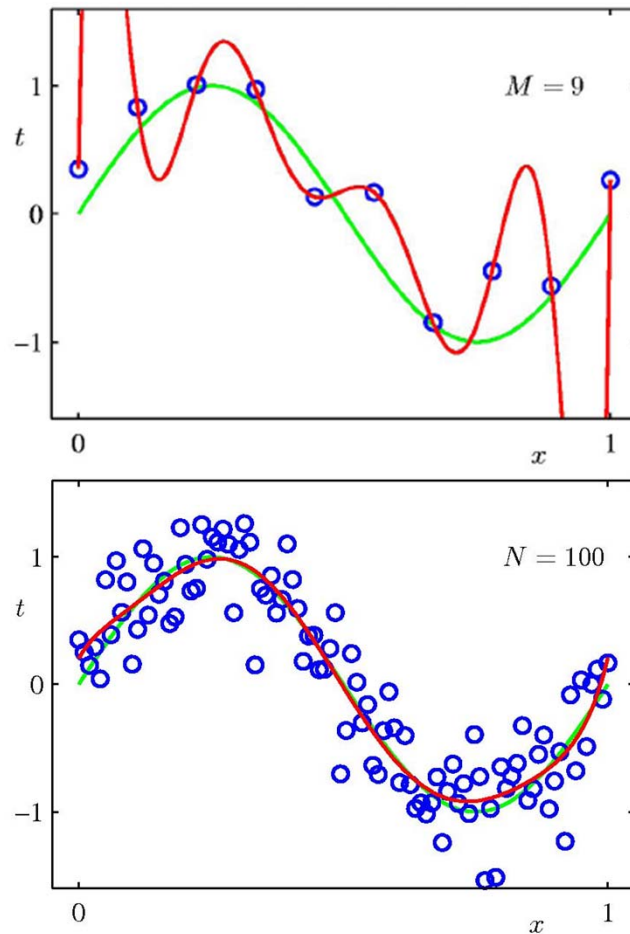
↑      ↑  
M-dimensional    N-dimensional

- $\mathbf{t}$  is a n-d vector
- $\mathcal{S}$  is the space spanned by  $\phi'_i$ s and  $\mathbf{t} \notin \mathcal{S}$
- $\mathbf{w}_{ML}$  minimizes the distance between  $\mathbf{t}$  and  $\mathcal{S}$  by finding the projection of  $\mathbf{t}$  onto  $\mathcal{S}$



$$\mathbf{y} = \boldsymbol{\phi} (\boldsymbol{\phi}^T \boldsymbol{\phi})^{-1} \boldsymbol{\phi}^T \mathbf{t}$$

# Over-fitting issue



- What can we do to curb overfitting
  - Use less complex model
  - Use more training examples
  - Regularization

# Regularized Least Squares (1)

- Consider the error function:

$$E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$$

Data term + Regularization term (penalize complex models)

- With the sum-of-squares error function and a **quadratic regularizer**, we get

$$\frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2 + \frac{\lambda}{2} \mathbf{w}^T \mathbf{w}$$

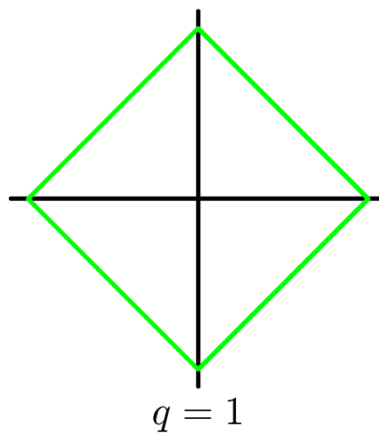
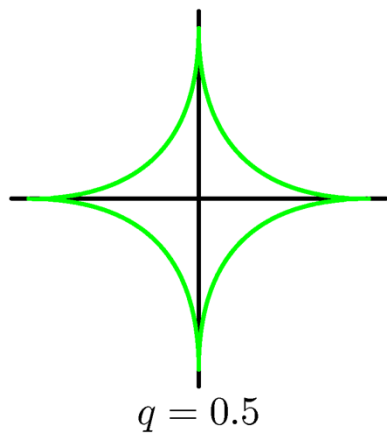
- which is minimized by

$$\mathbf{w} = \left( \lambda \mathbf{I} + \Phi^T \Phi \right)^{-1} \Phi^T \mathbf{t}.$$

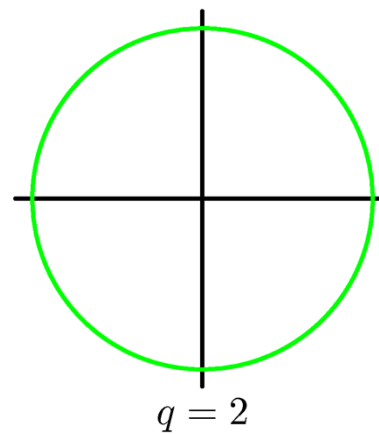
# Regularized Least Squares (2)

- With a more general regularizer, we have

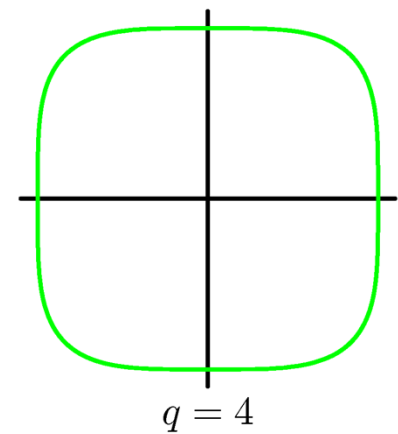
$$\frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2 + \frac{\lambda}{2} \sum_{j=1}^M |w_j|^q$$



Lasso



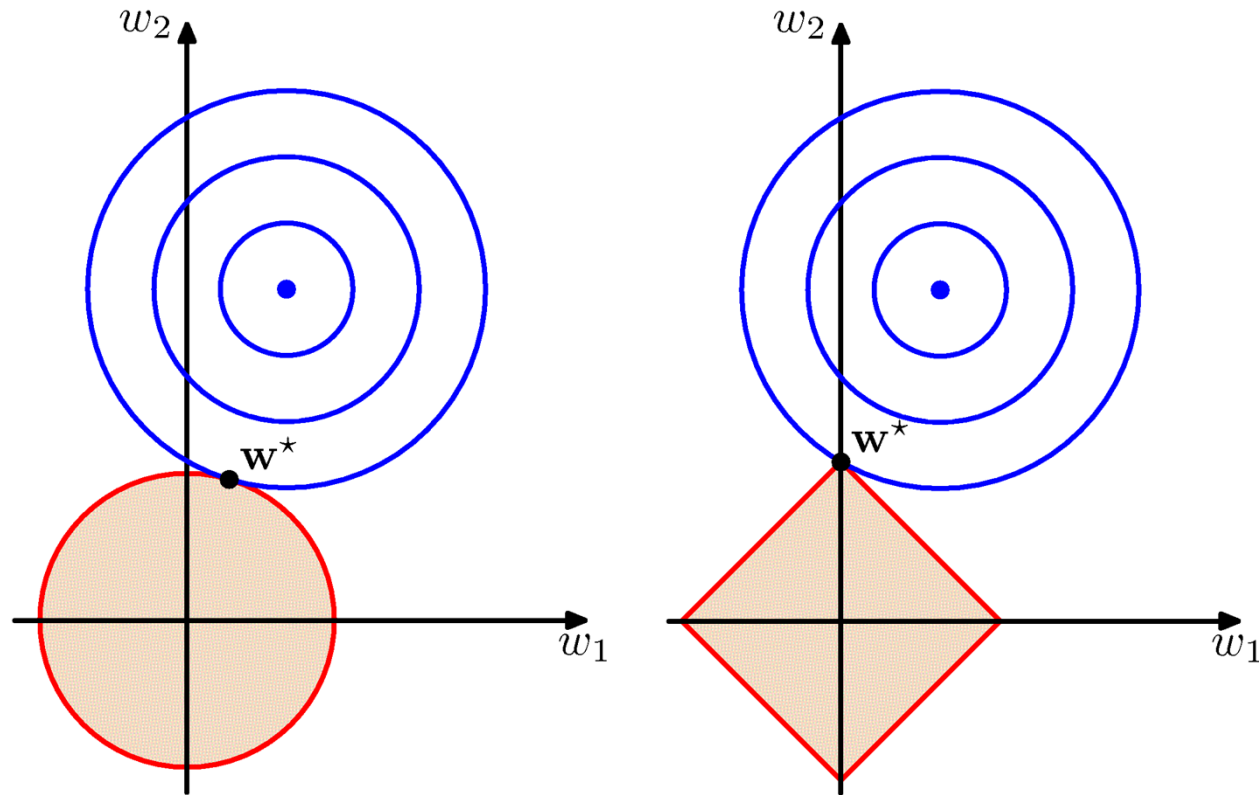
Quadratic





# Regularized Least Squares (3)

- Lasso tends to generate sparser solutions (majority of the weights shrink to zero) than a quadratic regularizer.



# The Bias-Variance Decomposition (1)

- Consider the *expected squared loss*,

$$\mathbb{E}[L] = \int \{y(\mathbf{x}) - h(\mathbf{x})\}^2 p(\mathbf{x}) d\mathbf{x} + \underbrace{\iint \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) d\mathbf{x} dt}_{\text{noise}}$$

- where

$$h(\mathbf{x}) = \mathbb{E}[t|\mathbf{x}] = \int tp(t|\mathbf{x}) dt.$$

- The second term of  $\mathbb{E}[L]$  corresponds to the noise inherent in the random variable  $t$ .
- What about the first term?

# The Bias-Variance Decomposition (2)

- Suppose we were given multiple data sets, each of size  $N$ . Any particular data set,  $\mathcal{D}$ , will give a particular function  $y(\mathbf{x}; \mathcal{D})$ . We then have

$$\begin{aligned} & \{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2 \\ &= \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] + \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 \\ &= \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2 + \{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 \\ &\quad + 2\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}\{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}. \end{aligned}$$

# The Bias-Variance Decomposition (3)

- Taking the expectation over  $\mathcal{D}$  yields

$$\begin{aligned} & \mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2] \\ &= \underbrace{\{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2}_{(\text{bias})^2} + \underbrace{\mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2]}_{\text{variance}}. \end{aligned}$$

# The Bias-Variance Decomposition (4)

- Thus we can write

$$\text{expected loss} = (\text{bias})^2 + \text{variance} + \text{noise}$$

- where

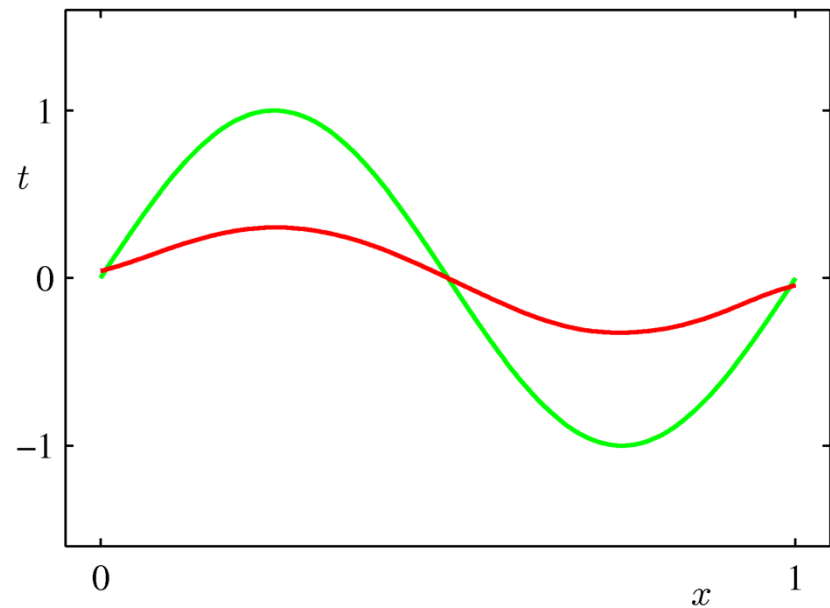
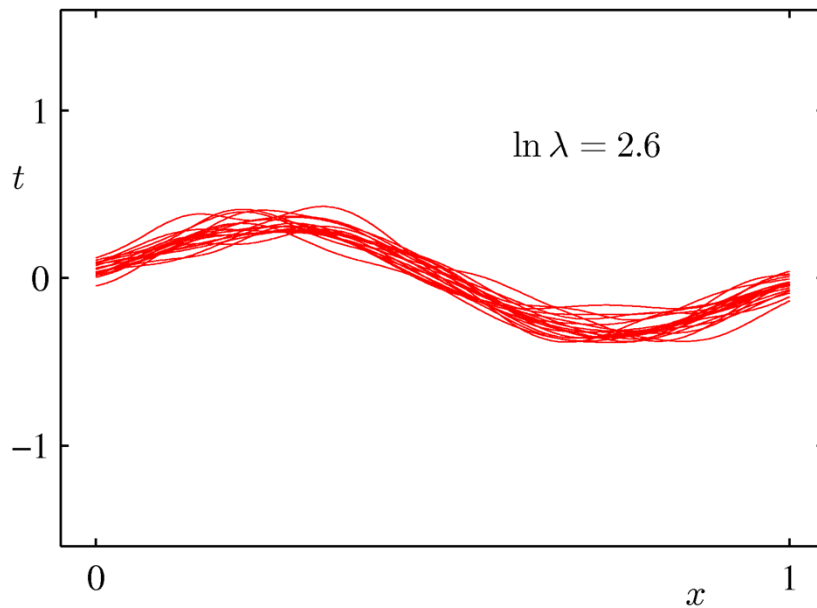
$$(\text{bias})^2 = \int \{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{variance} = \int \mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2] p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{noise} = \iint \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt$$

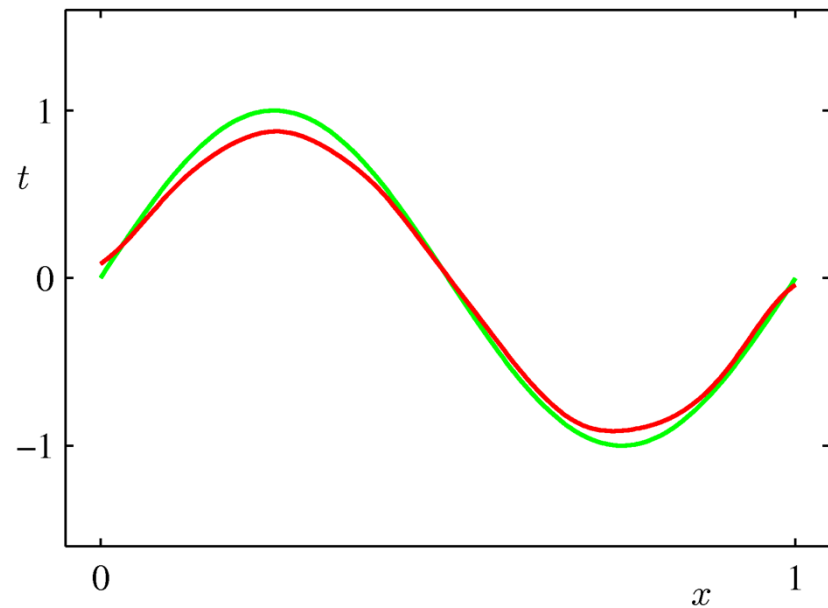
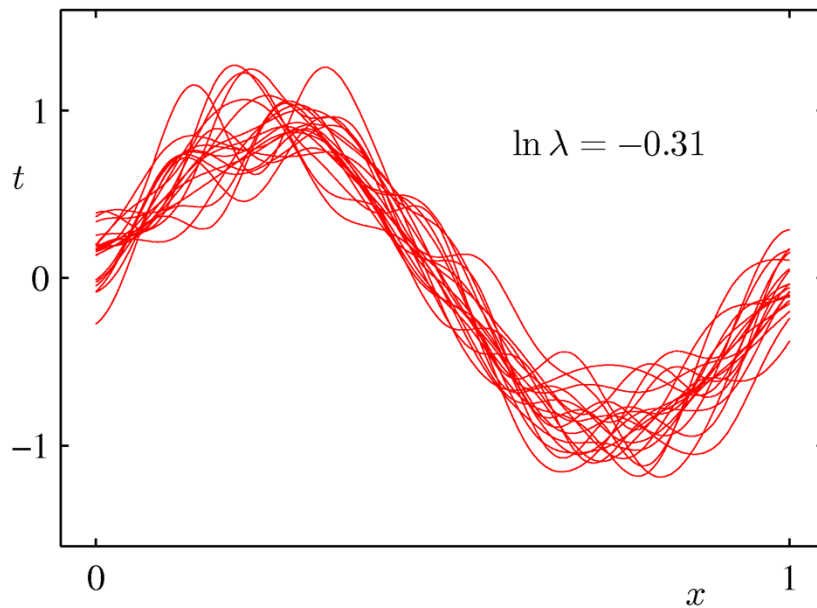
# The Bias-Variance Decomposition (5)

- Example: 25 data sets from the sinusoidal, varying the degree of regularization,  $\lambda$ .



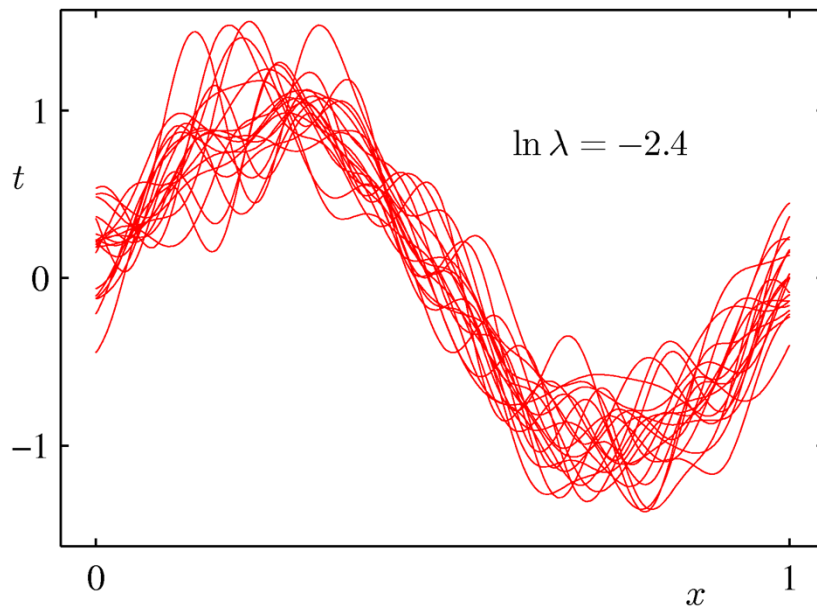
# The Bias-Variance Decomposition (6)

- Example: 25 data sets from the sinusoidal, varying the degree of regularization,  $\lambda$ .



# The Bias-Variance Decomposition (7)

- Example: 25 data sets from the sinusoidal, varying the degree of regularization,  $\lambda$ .





# The Bias-Variance Trade-off

- From these plots, we note that an over-regularized model (large  $\lambda$ ) will have a high bias, while an under-regularized model (small  $\lambda$ ) will have a high variance.

