

$$\mathbb{S}^{q-1} := \{\mathbf{y} \in \mathbb{R}^q : \|\mathbf{y}\| = 1\},$$

For $\mathbf{y} \in \mathbb{S}^{q-1}$, the probability density function of von Mises-Fisher distribution (vMF) is defined as

$$f_{vMF}(\mathbf{y} : \boldsymbol{\zeta}, \kappa) = C_q(\kappa) \exp(\kappa \boldsymbol{\zeta}^T \mathbf{y})$$

for a concentration parameter $\kappa \geq 0$ and mean direction $\boldsymbol{\zeta} \in \mathbb{S}^{q-1}$. Here, $C_q(\kappa)$ is a normalization constant defined as

$$C_q(\kappa) = \frac{\kappa^{q/2-1}}{(2\pi)^{q/2} I_{q/2-1}(\kappa)},$$

where $I_\nu(\cdot)$ is the modified Bessel function of the first kind at order ν .

Here, the mean direction parameter ζ is restricted on a unit sphere, so we can consider the following reparametrization to alleviate this. If we let $\boldsymbol{\theta} = \kappa \boldsymbol{\zeta}$ for $\kappa > 0$, then $\kappa = \|\boldsymbol{\theta}\|$ with $\boldsymbol{\theta} \in \mathbb{R}^q$ and we have a reparametrized vMF distribution by a mean parameter $\boldsymbol{\theta} \in \mathbb{R}^q$:

$$f_{vMF}(\mathbf{y} : \boldsymbol{\theta}) = \exp \left(\boldsymbol{\theta}^T \mathbf{y} + \log C_q(\|\boldsymbol{\theta}\|) \right). \quad (1)$$

In addition, let us express the probability density function of a random vector following the exponential family as

$$f(\mathbf{y} | \boldsymbol{\theta}, \phi) = \exp \left(\frac{\boldsymbol{\theta}^T \mathbf{y} - b(\boldsymbol{\theta})}{a(\phi)} + c(\mathbf{y}, \phi) \right),$$

where $\boldsymbol{\theta}$ and ϕ are the natural and scale parameters, $b(\cdot)$ and $c(\cdot)$ are known functions related to different exponential responses. Then, the reparametrized

$$g(\mathbb{E}[\mathbf{y}_i|\mathbf{x}_i]) = \boldsymbol{\theta}_i = \boldsymbol{\mu} + \sum_{j=1}^p x_{ij}\boldsymbol{\beta}_j = \boldsymbol{\mu} + \mathbf{B}^T \mathbf{x}_i \quad \text{and} \quad \mathbf{y}_i|\mathbf{x}_i \sim f_{vMF}(\cdot|\boldsymbol{\theta}_i),$$

- Constrained model : $\boldsymbol{\beta}_j \perp \boldsymbol{\mu}, \quad j = 1, \dots, p$
- Unconstrained model : $\boldsymbol{\beta}_j \in \mathbb{R}^q, \quad j = 1, \dots, p$

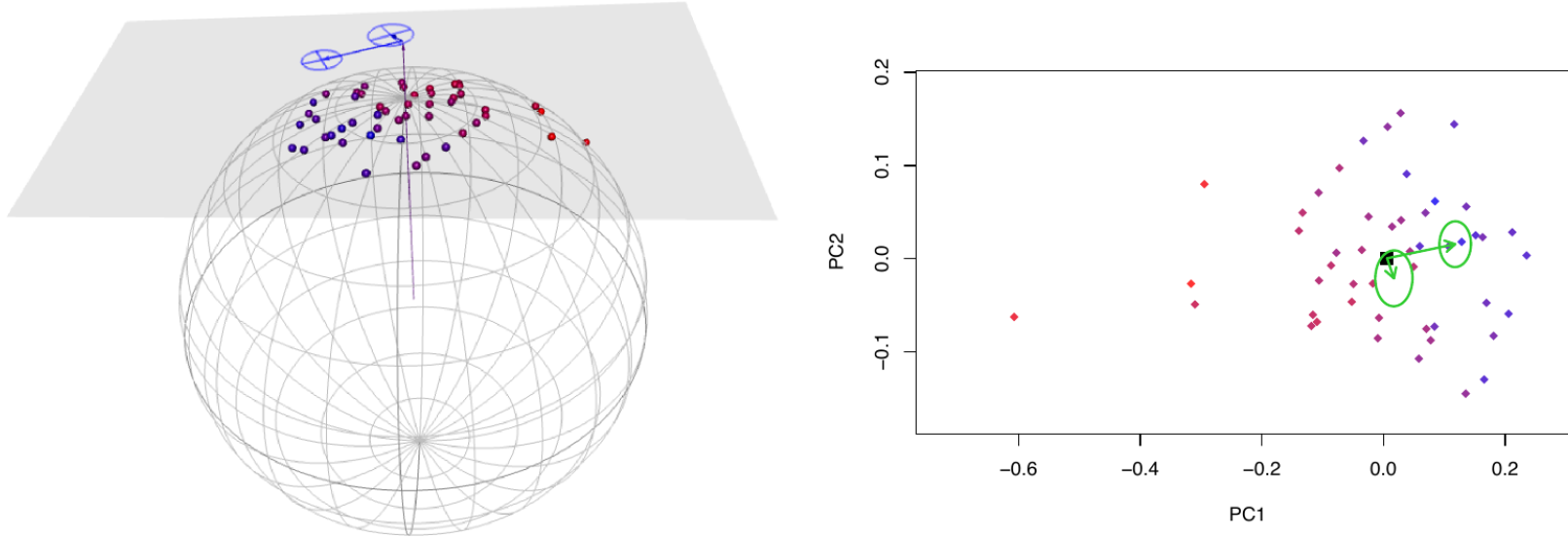


Figure 1: An example of regression model for vMF distributed spherical responses on \mathbb{S}^2 . The red points indicates the spherical responses, the blue line the estimated regression coefficients, $\hat{\boldsymbol{\beta}}_1 = (4.80, -10.22, -0.20)^T$, $\hat{\boldsymbol{\beta}}_2 = (-1.71, -1.71, 0.08)^T$ with p-values $1.19 \cdot 10^{-12}$, and 0.1542, and the purple line the estimated mean direction $\hat{\boldsymbol{\mu}} = (2.41, 0.13, 48.69)^T$. For more visibility, the coefficients were normalized to have norm = 1.25.

- Global null hypothesis

$$H_0 : \boldsymbol{\beta}^* = \mathbf{0} \tag{6}$$

- Individual null hypothesis

$$H_{0j} : \boldsymbol{\beta}_j = \mathbf{0}, \tag{7}$$

for $j = 1, \dots, p$.

- Other hypotheses for mean or dispersion difference

(1) Location difference: $H_A : \boldsymbol{\mu}^T \boldsymbol{\beta}_j = 0 \ \& \ \boldsymbol{\beta}_j \neq \mathbf{0}$

(2) Dispersion difference: $H_A : \angle(\boldsymbol{\mu}, \boldsymbol{\beta}_j) = 0 \ \& \ \boldsymbol{\beta}_j \neq \mathbf{0}$

Lemma 1. *Under the assumptions (A1)–(A3), the normed score vector*

$$\mathbf{F}_n^{-T/2} \mathbf{s}_n(\boldsymbol{\beta}_0^*) \xrightarrow{D} N(\mathbf{0}, \mathbf{I}_{pq}) \quad \text{as } n \rightarrow \infty,$$

where \mathbf{I}_{pq} is a pq -dimensional identity matrix.

Lemma 2. *Under the assumptions (A1)–(A3)*

$$\max_{\boldsymbol{\beta}^* \in N_n(\epsilon)} \|\mathbf{V}_n(\boldsymbol{\beta}^*) - \mathbf{I}_{pq}\| \xrightarrow{p} 0, \quad \forall \epsilon > 0,$$

where $\mathbf{V}_n(\boldsymbol{\beta}^) := \mathbf{F}_n^{-1/2} \mathbf{F}_n(\boldsymbol{\beta}^*) \mathbf{F}_n^{-T/2}$.*

Theorem 1. *Under the assumptions (A1)–(A3), the asymptotic normality of $\hat{\boldsymbol{\beta}}^*$ can be obtained as follows:*

$$\mathbf{F}_n^{T/2}(\hat{\boldsymbol{\beta}}^* - \boldsymbol{\beta}_0^*) \xrightarrow{d} N(\mathbf{0}, \mathbf{I}_{pq}).$$

3.2.1. Hypotheses

- Global null hypothesis

$$H_0 : \boldsymbol{\beta}^* = \boldsymbol{\beta}_0^* \quad (6)$$

- Individual null hypothesis

$$H_{0j} : \boldsymbol{\beta}_j = \boldsymbol{\beta}_{0j}, \quad (7)$$

for $j = 1, \dots, p$.

For testing the null hypothesis (6), we introduce the Wald statistics as:

$$W_j := \hat{\boldsymbol{\beta}}_j^T \hat{\mathbf{F}}_{n,j}(\hat{\boldsymbol{\beta}}_j) \hat{\boldsymbol{\beta}}_j, \quad j = 1, \dots, p, \quad (8)$$

where $\hat{\mathbf{F}}_{n,j}(\hat{\boldsymbol{\beta}}_j)$ is the empirical version of $\mathbf{F}_{n,j}(\boldsymbol{\beta}_{0j})$ evaluated at the maximum likelihood estimates, and $\mathbf{F}_{n,j}(\boldsymbol{\beta}_{0j})$ is the $q \times q$ dimensional submatrix consisting of the elements corresponding to the j th predictor in $\mathbf{F}_n(\boldsymbol{\beta}_0^*)$.

Theorem 2. *Under Assumptions (A1)–(A3), for $j = 1, \dots, p$, if the null hypothesis $H_{0j} : \boldsymbol{\beta}_j = \mathbf{0}$ is true,*

$$W_j \xrightarrow{d} \chi^2(q)$$

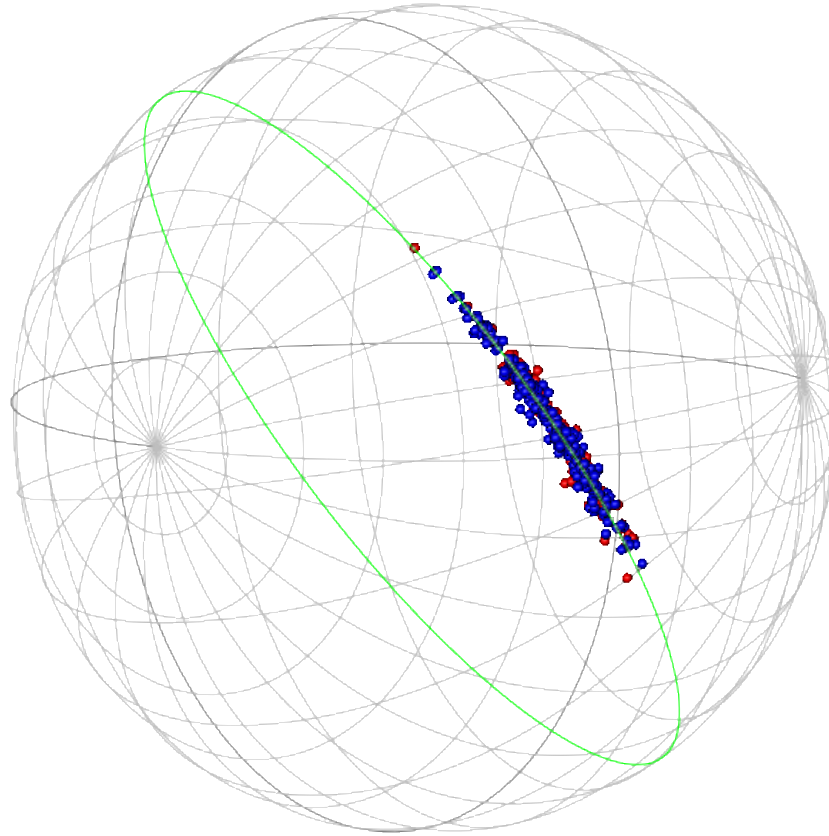
as $n \rightarrow \infty$, where $\chi^2(q)$ is the chi-squared distribution with q degrees of freedom.

Real data analysis: procedure

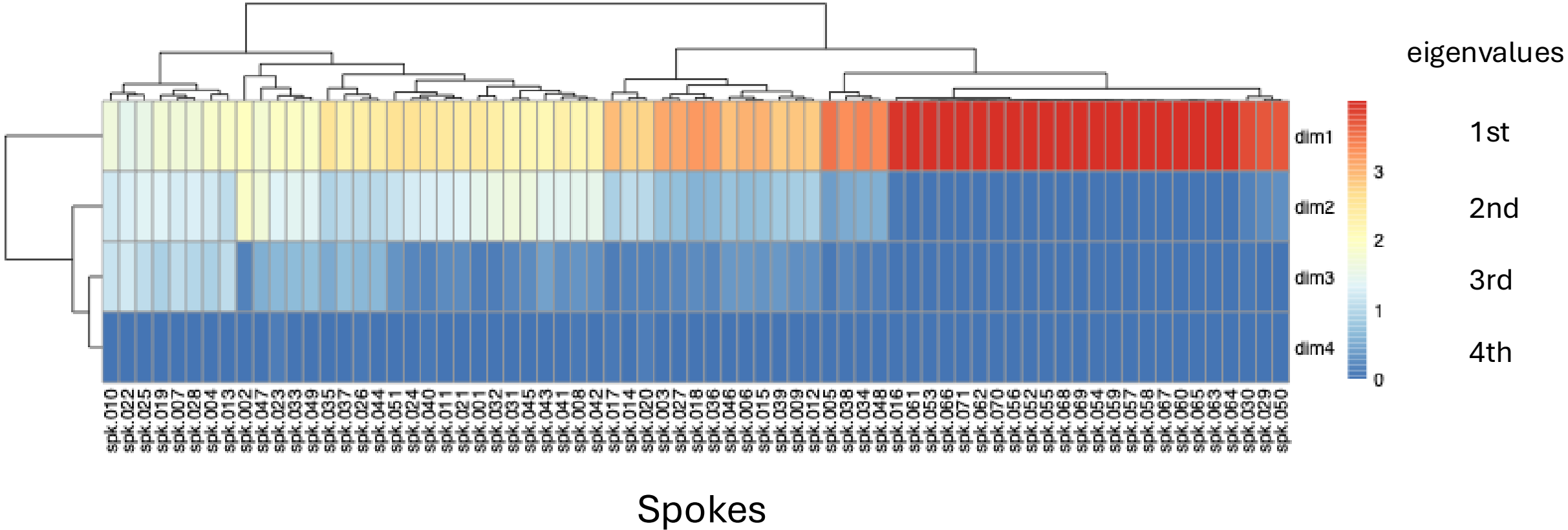
- Target responses = framesBasedOnParentsUnitQuaternion
 - Other choices: Spokes' directions, Connections' directions, ...
- Covariate of interest = Parkinson disease (the presence or absence)
- The response variables are on S^3 but actually have a rank less than 3,
>> so we use Principal Nested Sphere to get the responses projected onto S^2 .
- We fit our two models separately to the responses on S^3 and to those projected onto S^2 .
- Methods:
 - Orthogonally constrained sphereGLM
 - Unconstrained sphereGLM
- Our goal is to identify spokes associated with Parkinson's disease in terms of
 - (1) Location difference between two groups
 - (2) Dispersion difference between two groups

(though it's questionable whether this has clinical significance)
- In (2) If there is a dispersion difference, compare the sample covariance matrix for each group in terms of total var.
- If a dispersion difference is found in (2), we'll compare the sample covariance for each group in terms of total variance
- For better visibility, we coded the covariates to be -0.1 for control and +0.1 for case to increase the size of estimates.

Principal Nested Sphere (PNS) : Example (spk26)

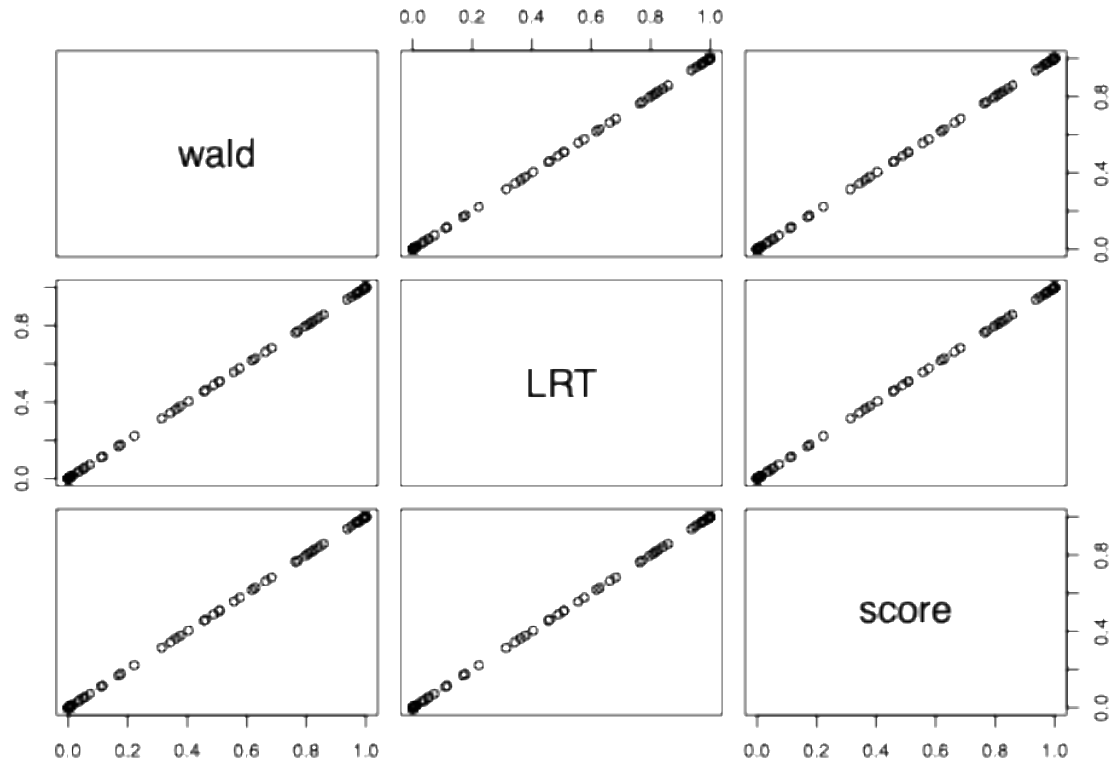


Eigenvalues for each of the 71 spokes using correlation matrix

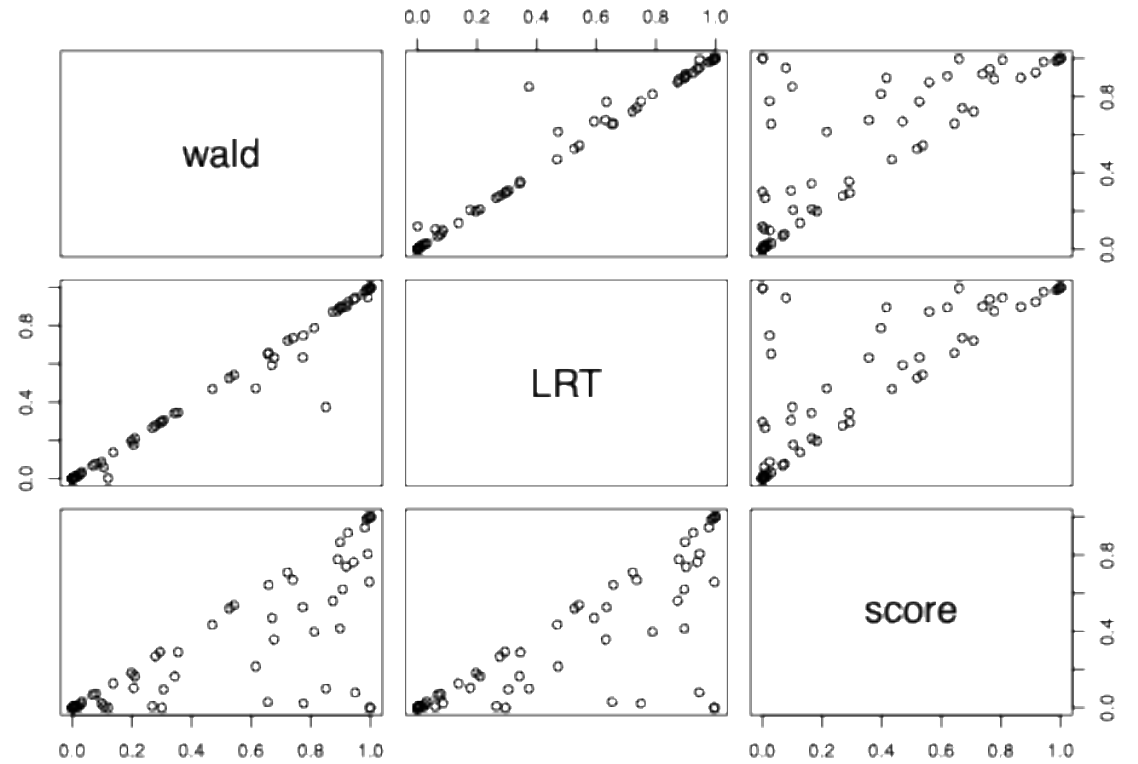


Wald vs LRT vs Score tests

constrained model



unconstrained model

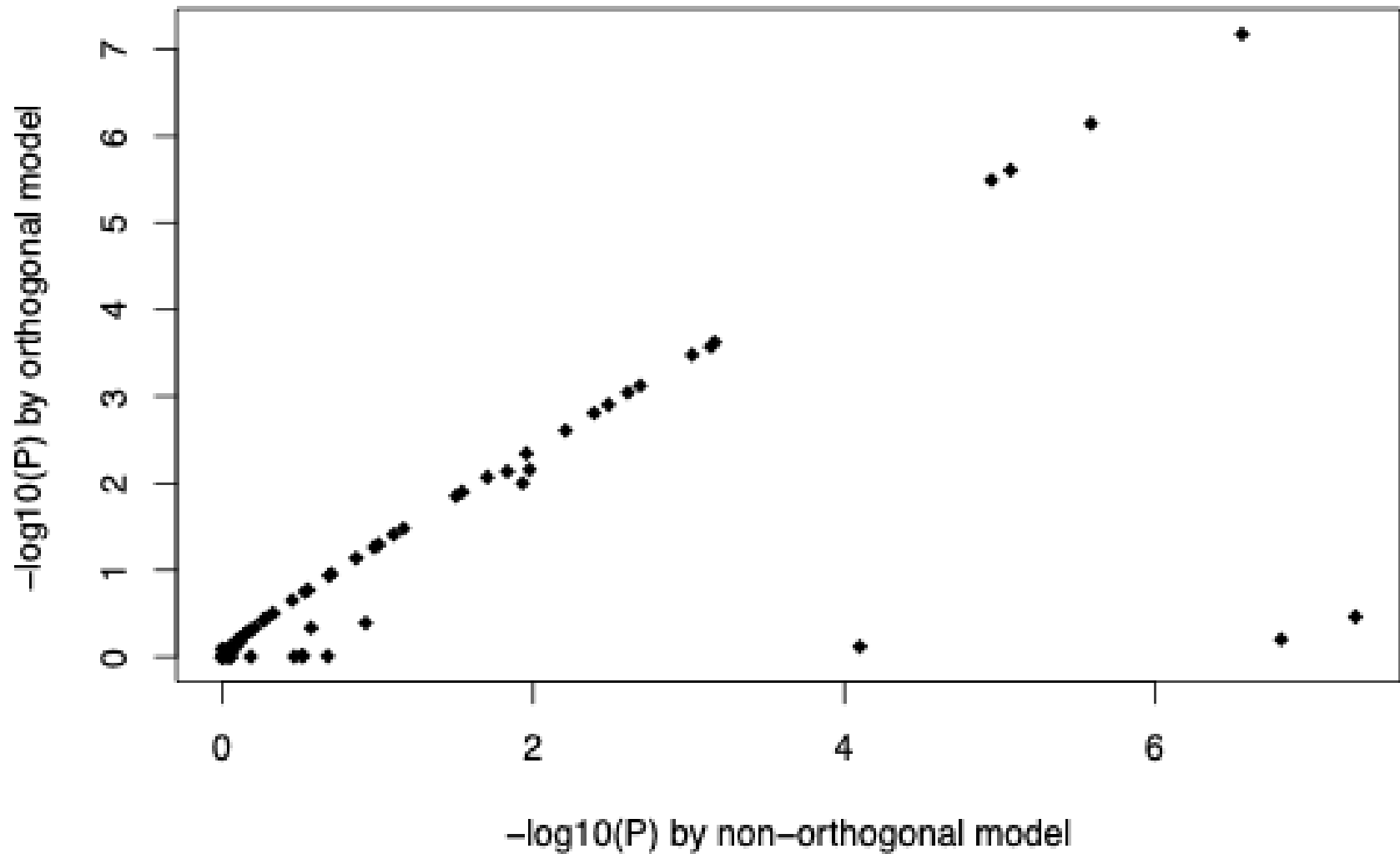


➔ LRT ??

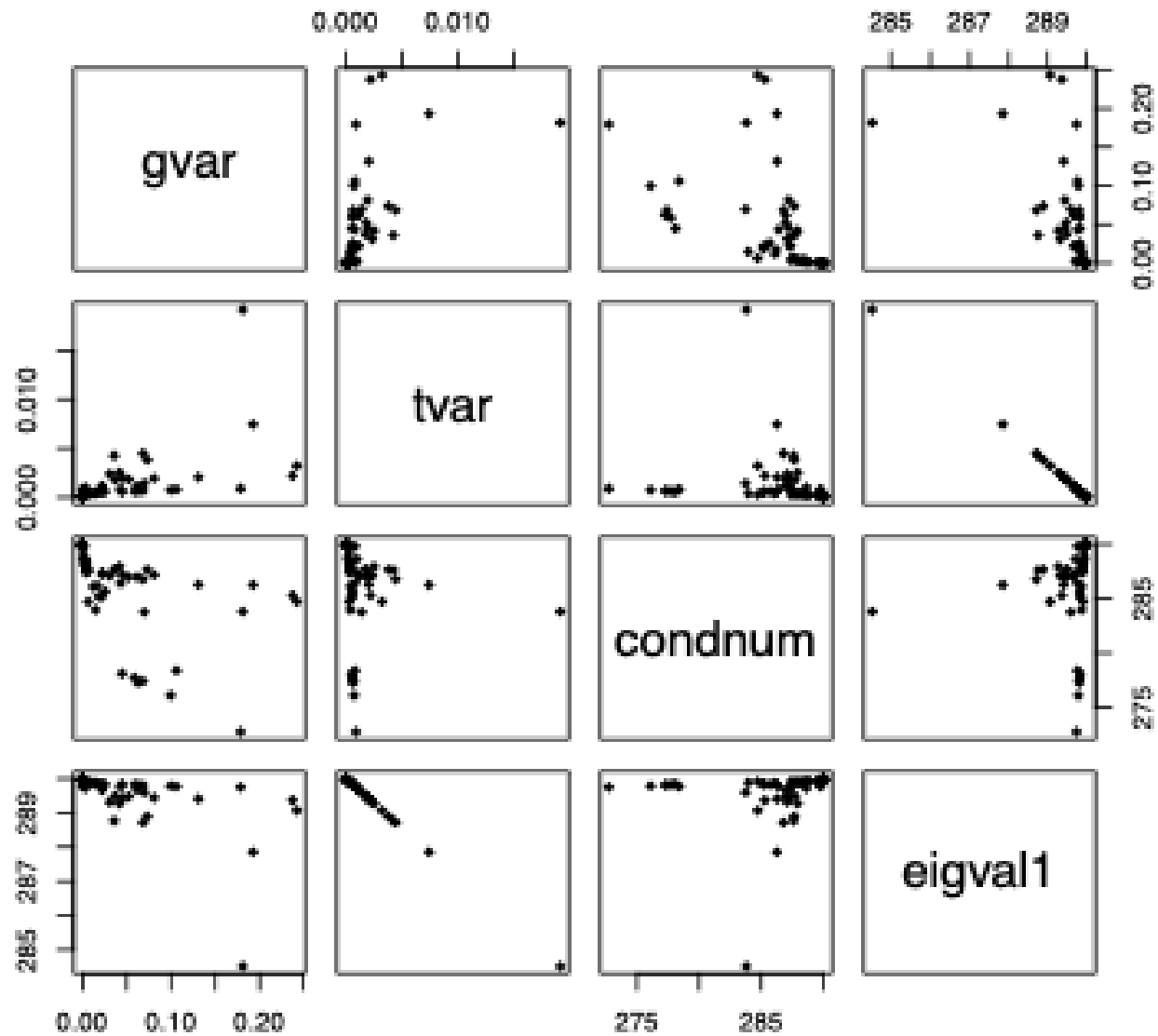
(Wald seems to be conservative in non-ortho model)

- log10 (p-values) calculated using non-orthogonal and orthogonal models

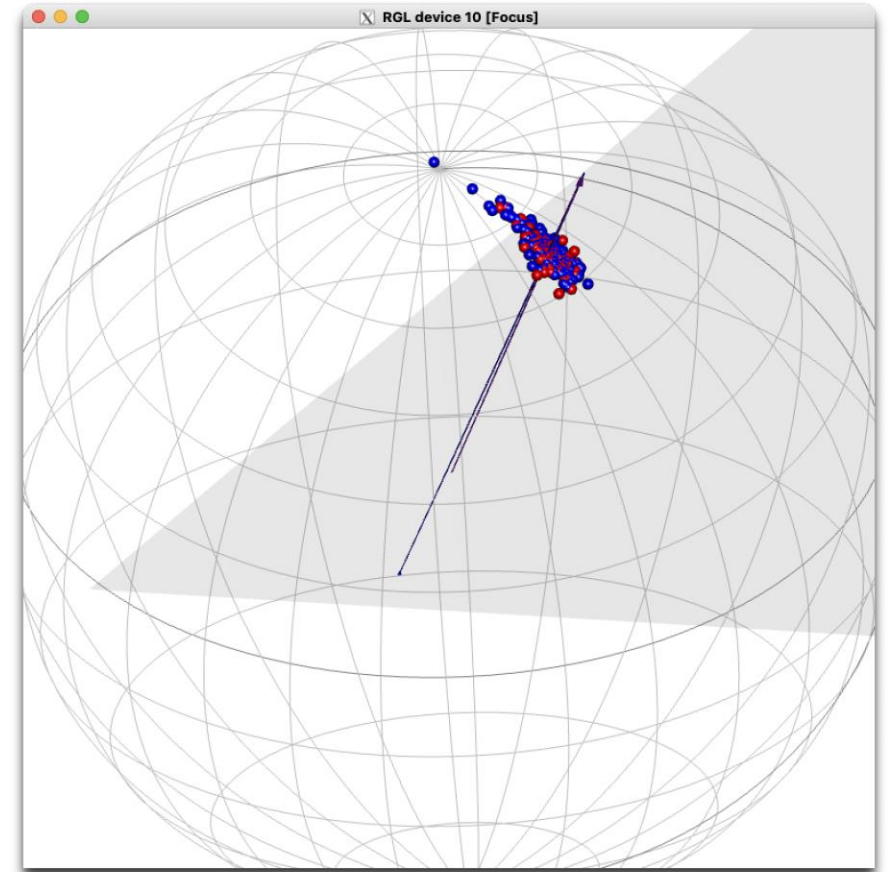
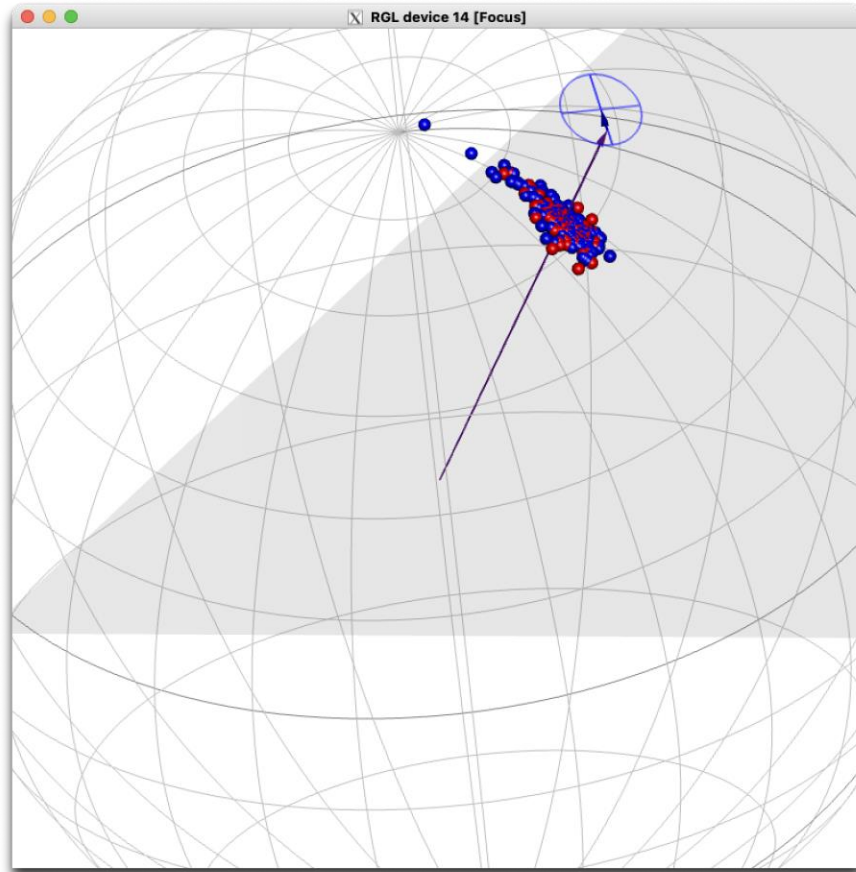
(LRT)



Comparison between covariance matrix criteria



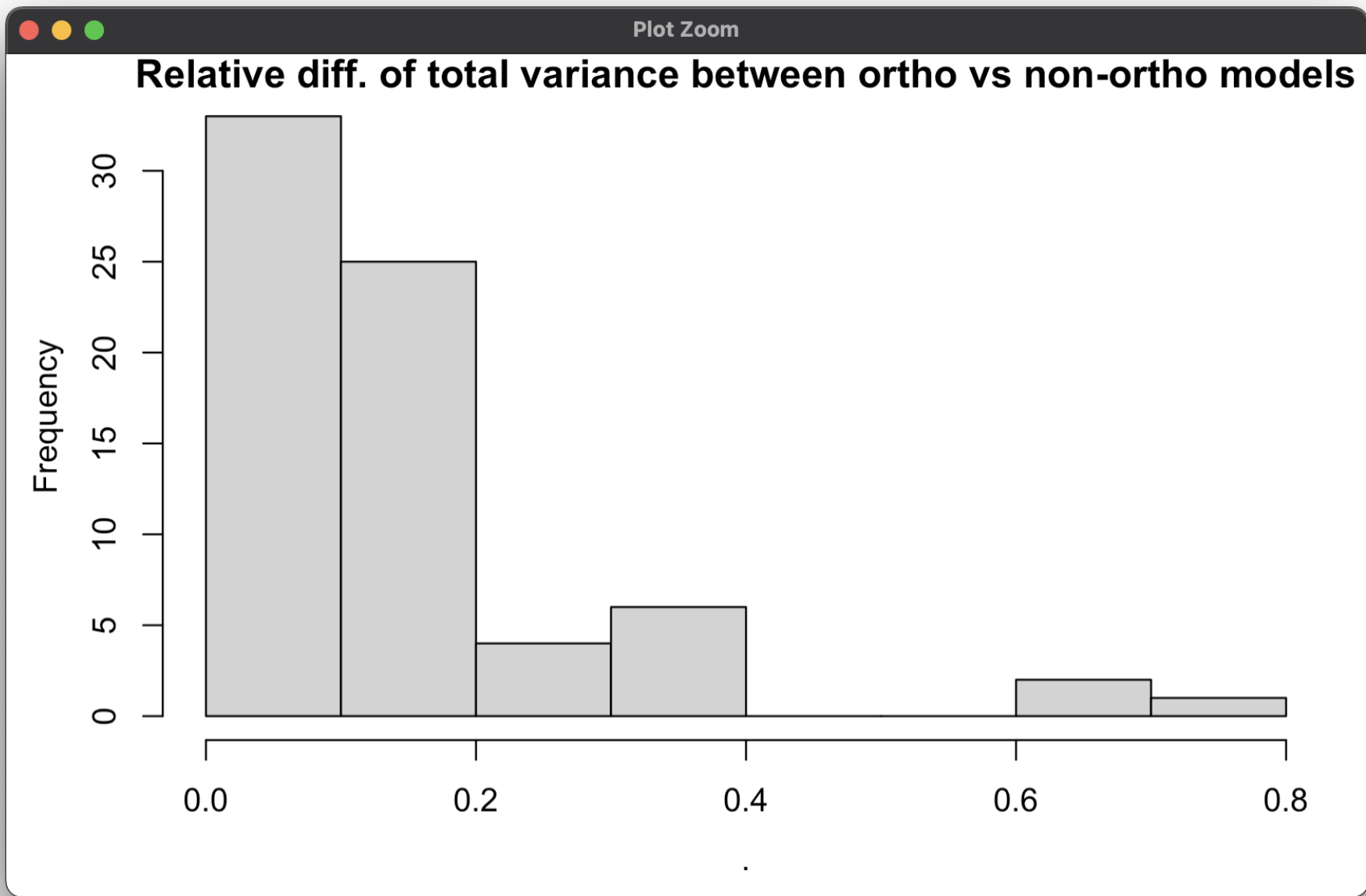
Dispersion only



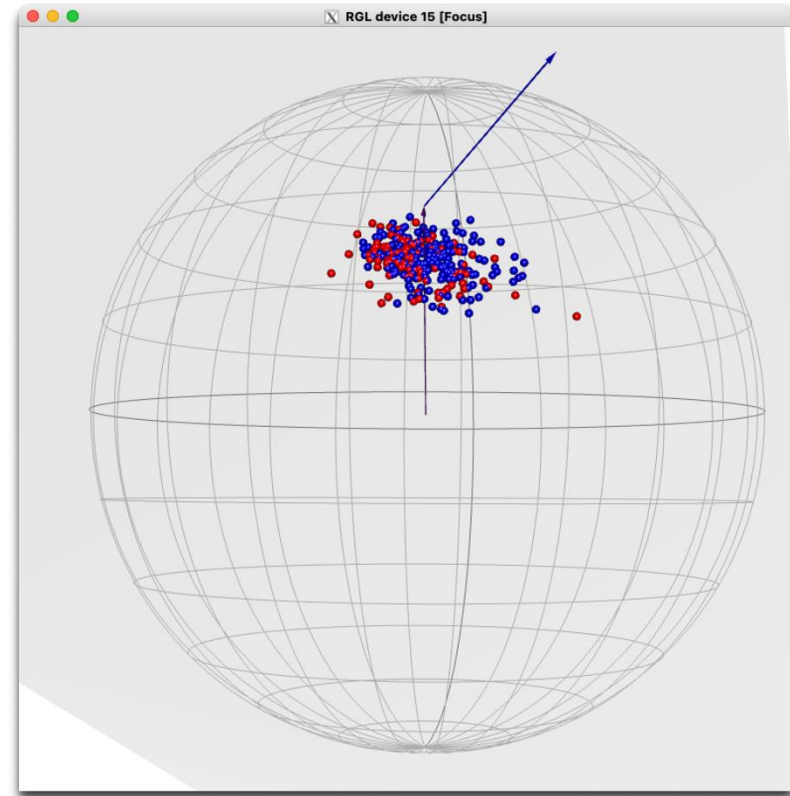
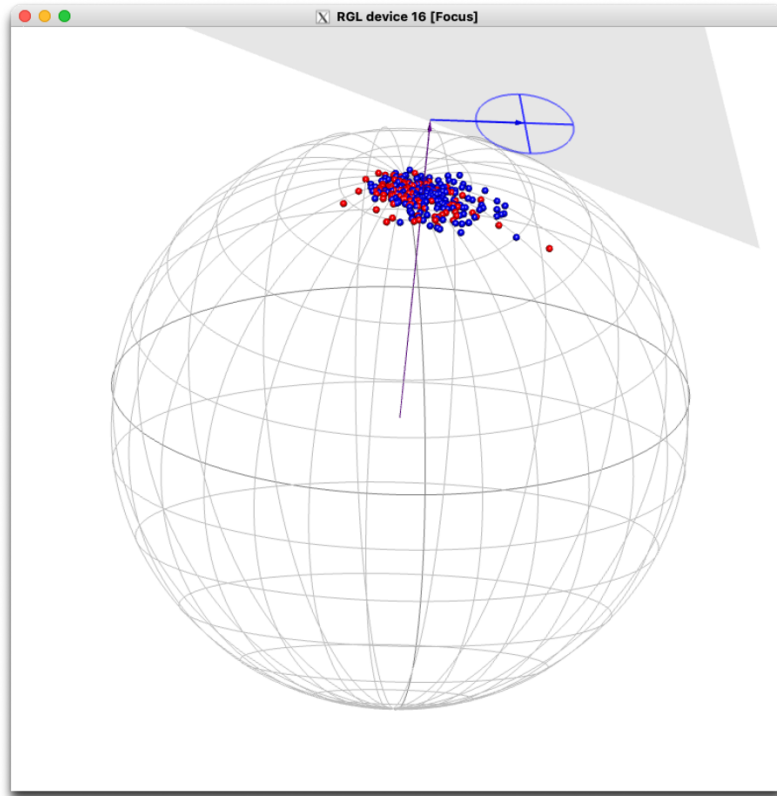
Spoke=38: Dispersion diff. but not Location diff.

```
> pvalue.array[38, "LRT"]  
      ortho  nonortho  
0.404392  0.000382
```

```
> crit.array["tvar", 38]  
Total  group1  group2  
0.00184 0.00126 0.00218
```



Dispersion and Location



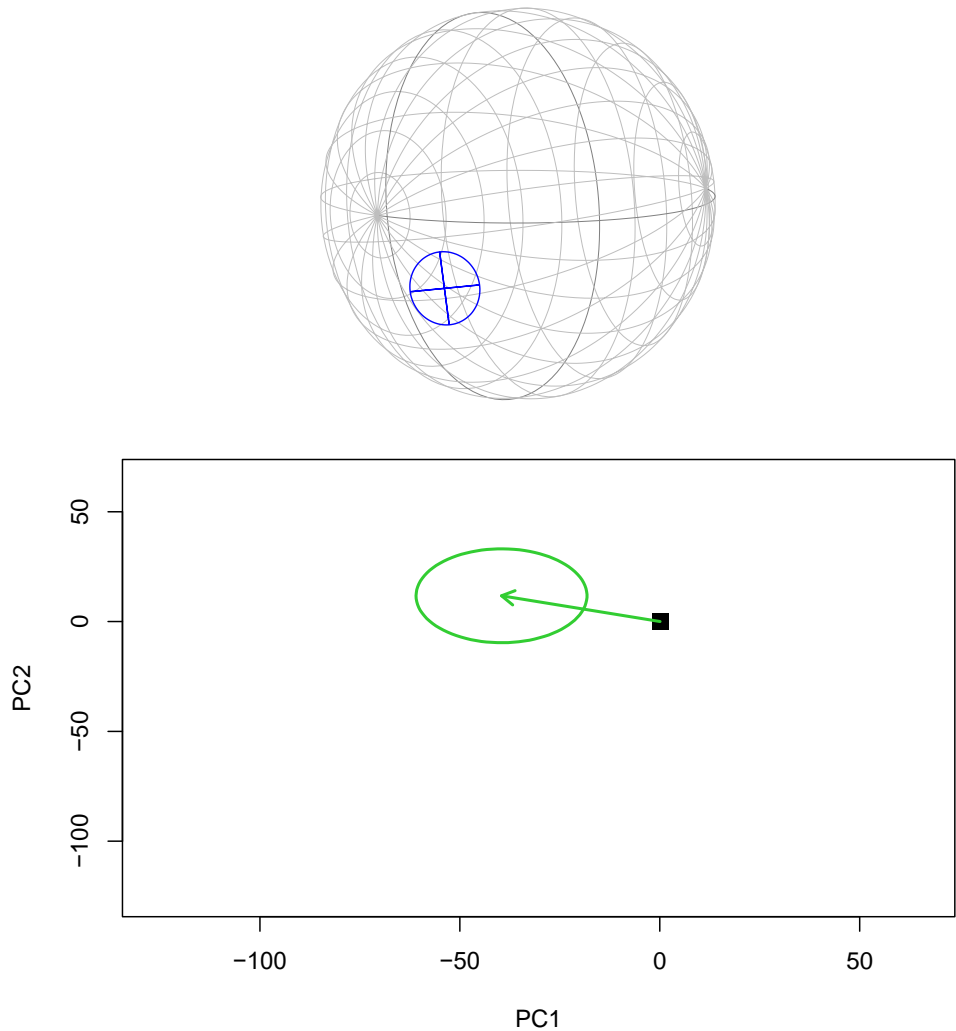
Spoke=9; Location \& Dispersion difference

```
> pvalue.array[idx.MeanDiff.VarDiff, , "LRT"]
```

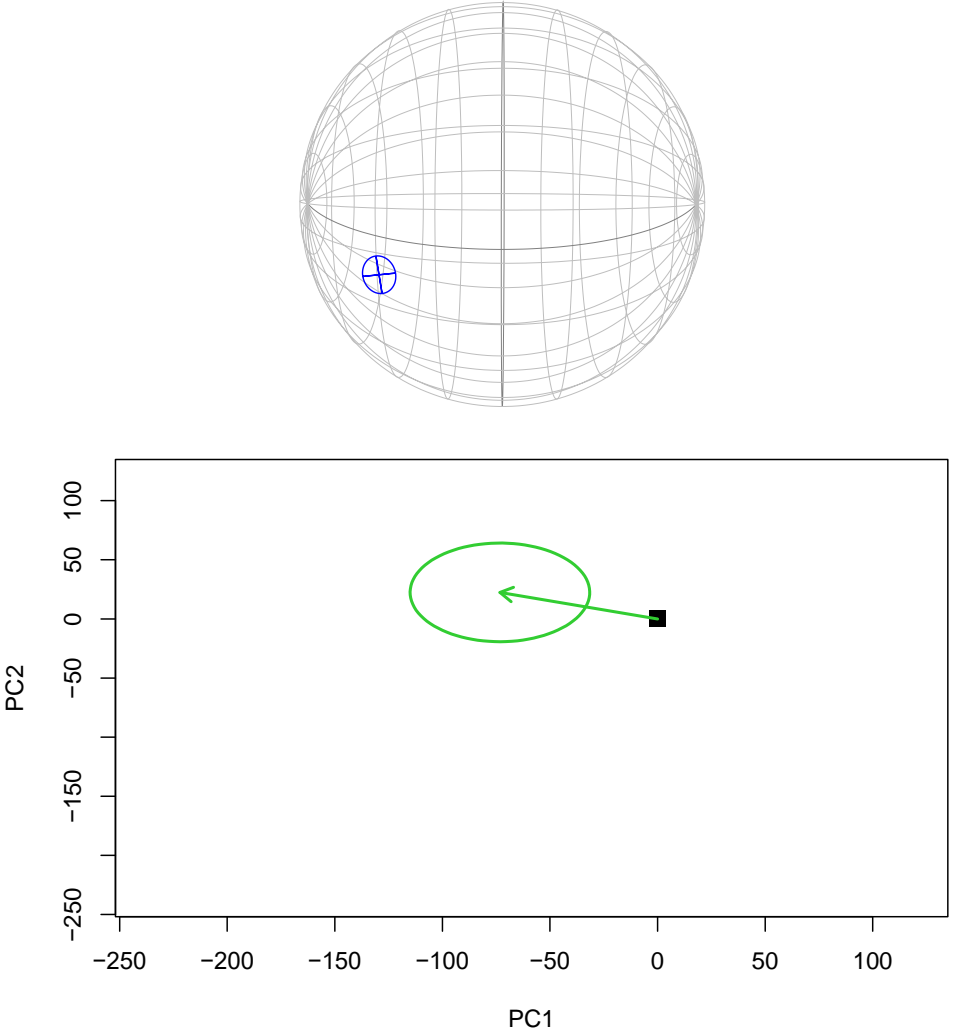
	ortho	nonortho
spk9	7.186e-07	2.574e-06
spk12	2.463e-06	8.491e-06
spk14	3.222e-06	7.429e-06
spk26	6.668e-08	2.512e-07

The two significant spokes were identified by both methods based on the adjusted p-values.

9th spoke

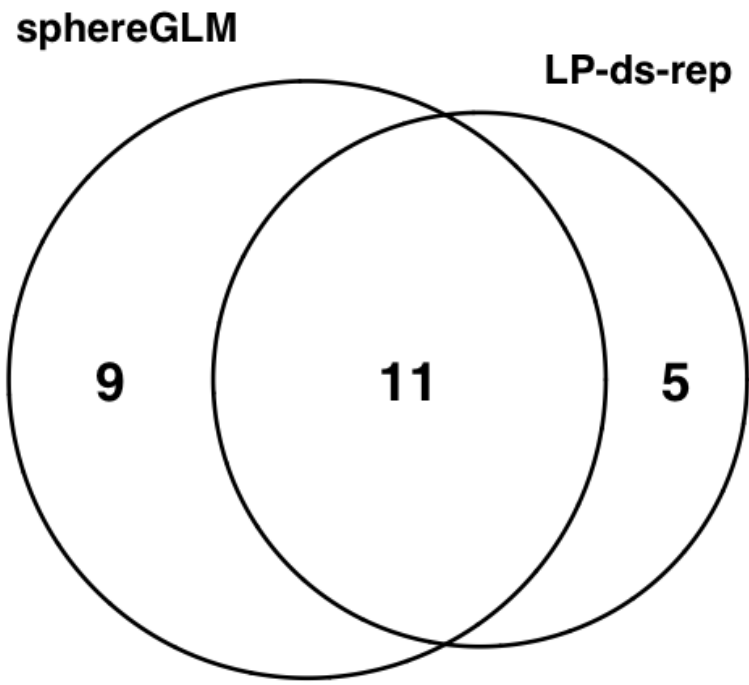


12th spoke

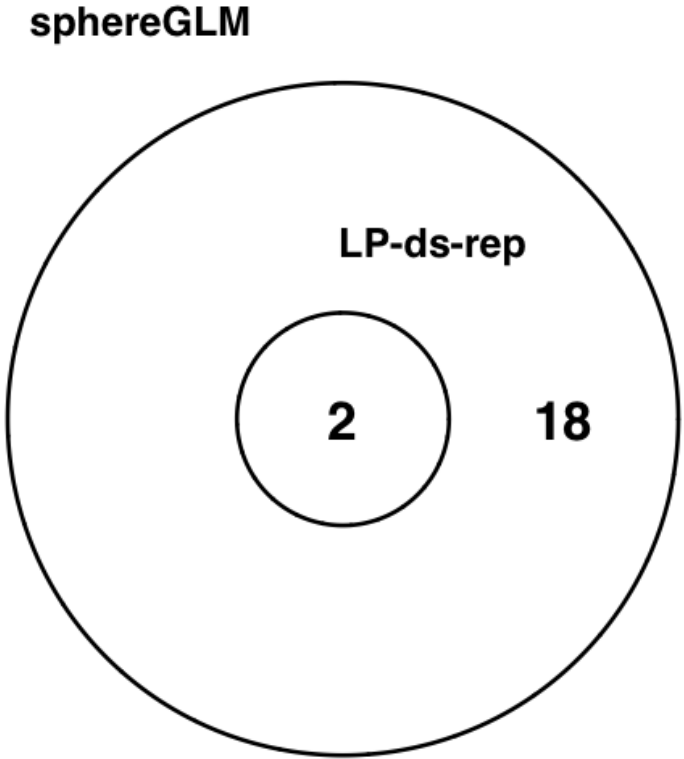


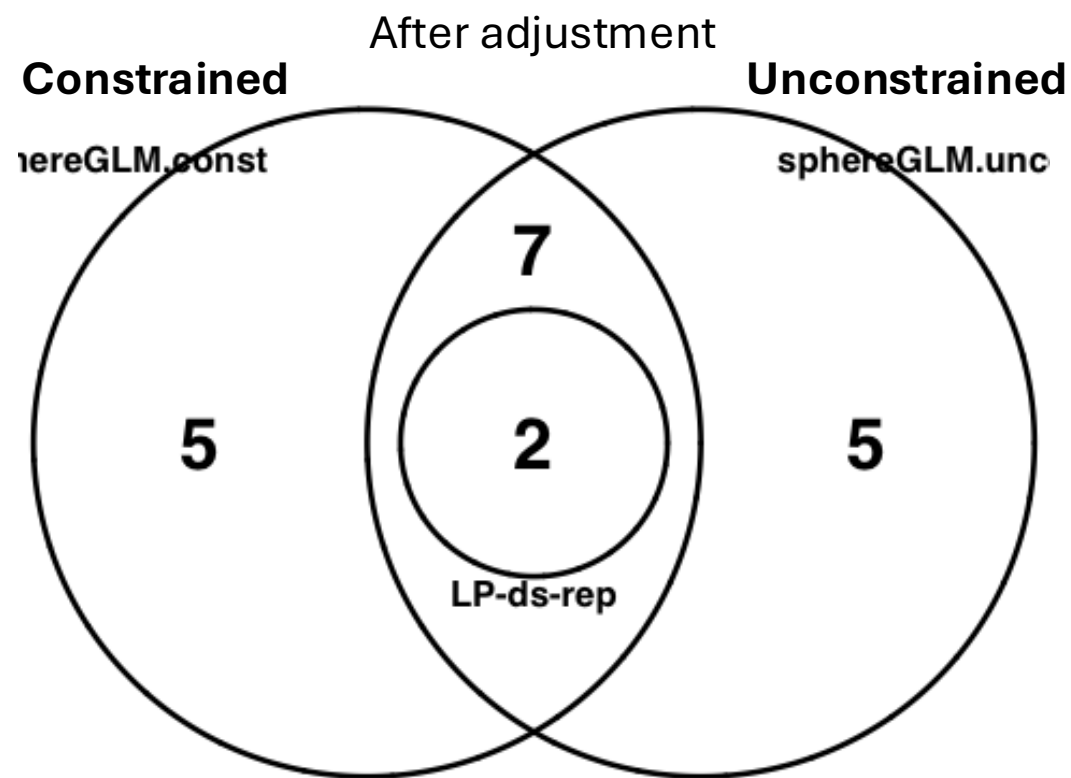
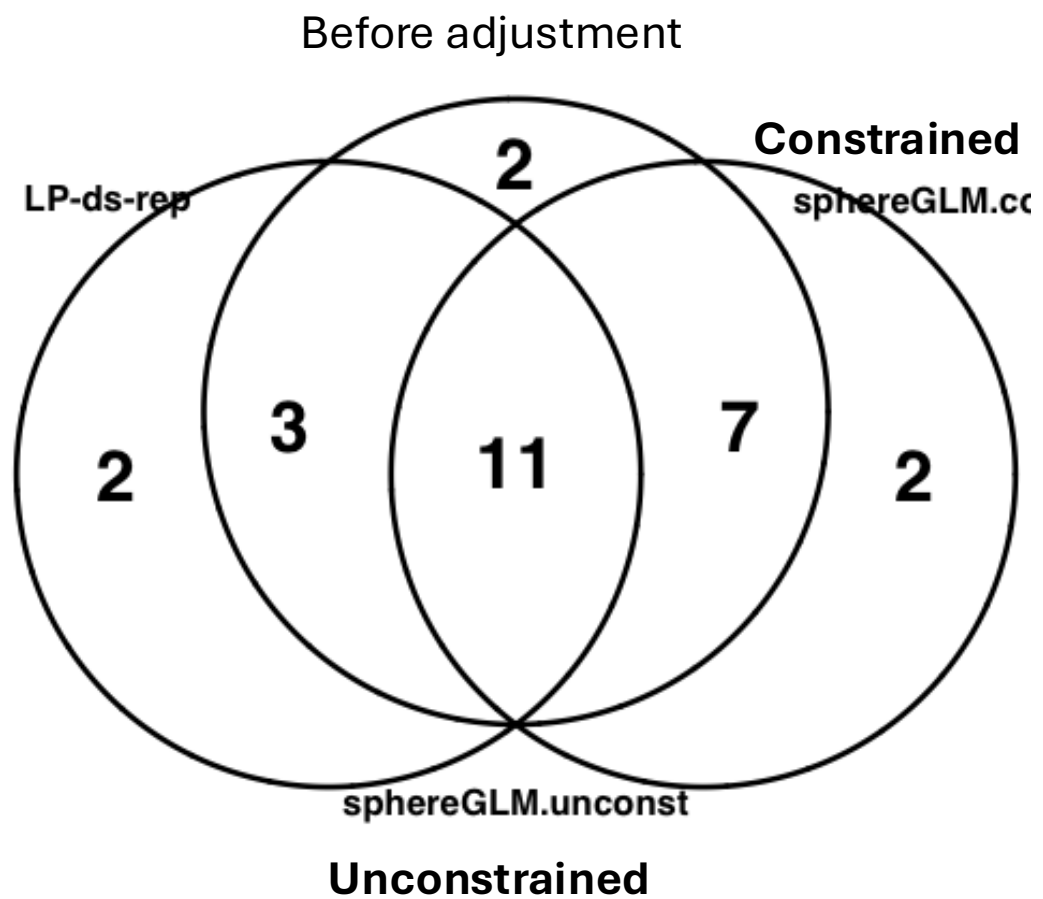
vs LP-ds-rep

Before adjustment



After adjustment





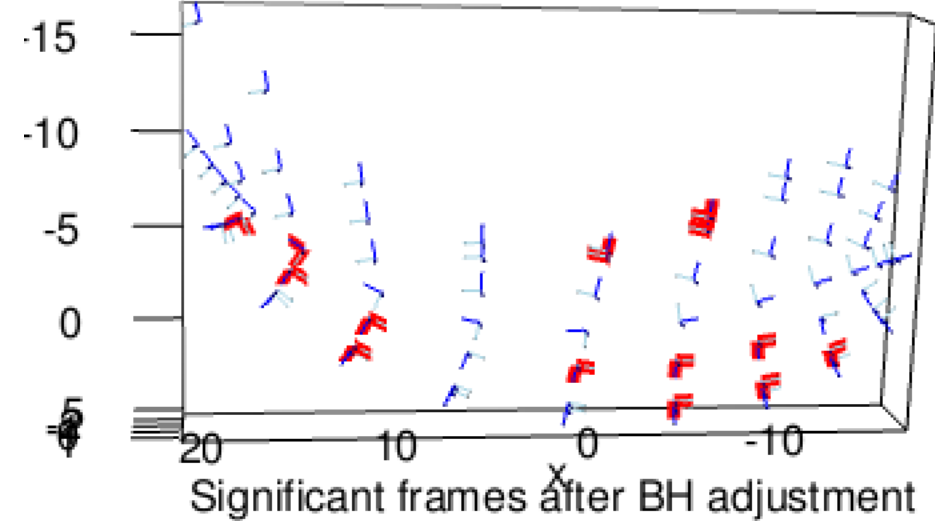
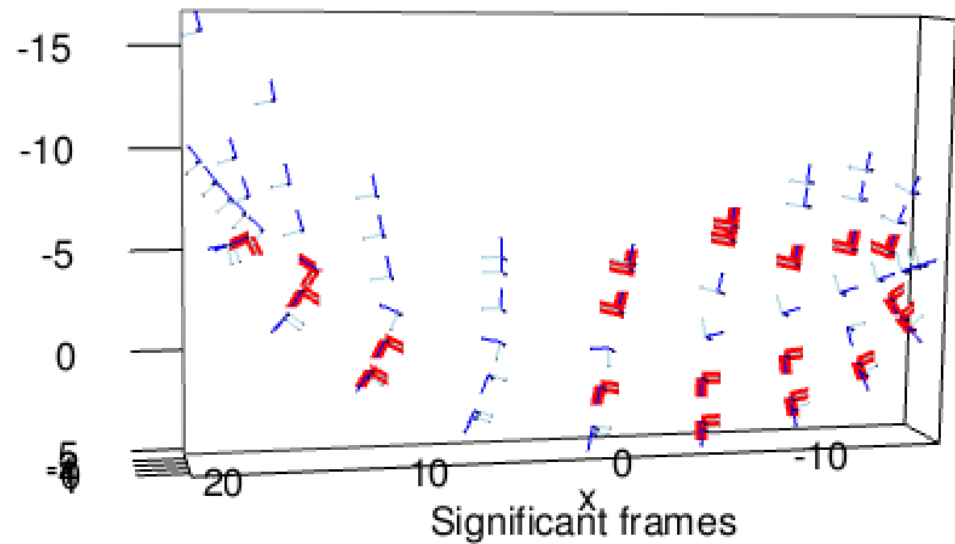
vs LP-ds-rep

Frames' directions

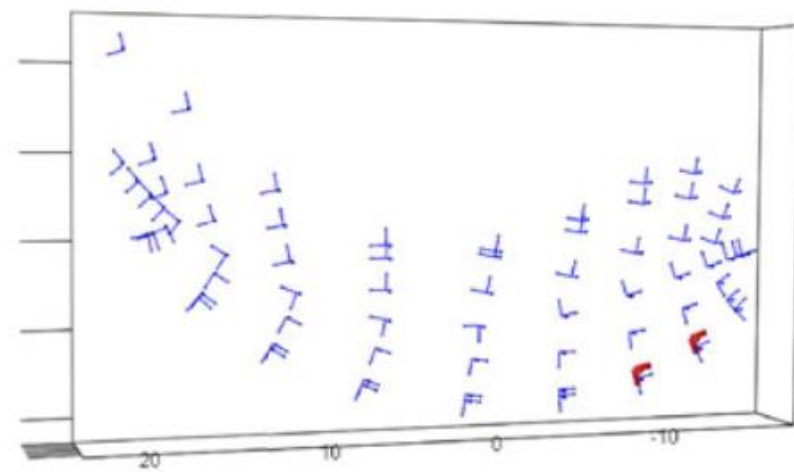
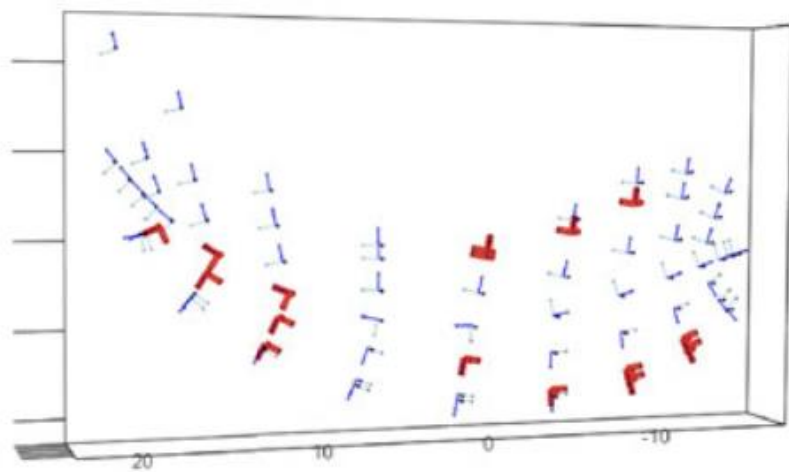
Before adjustment

After adjustment

Constrained
sphereGLM



LP-ds-rep



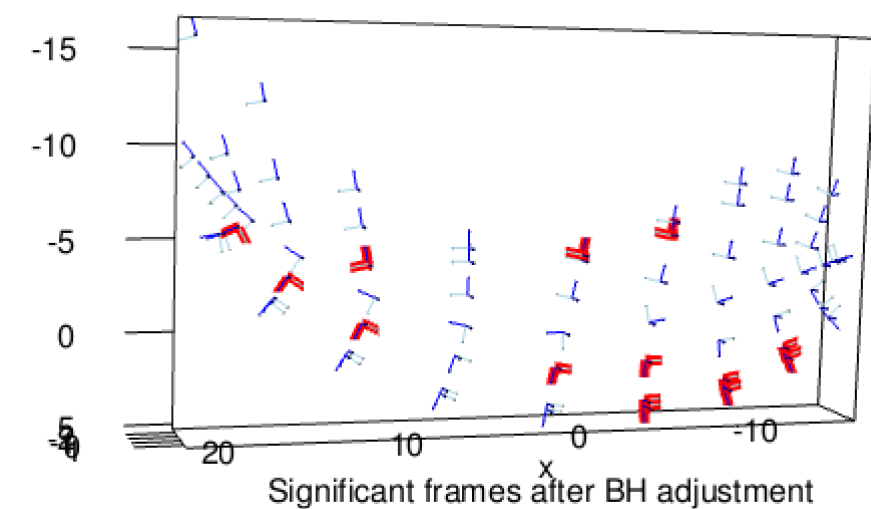
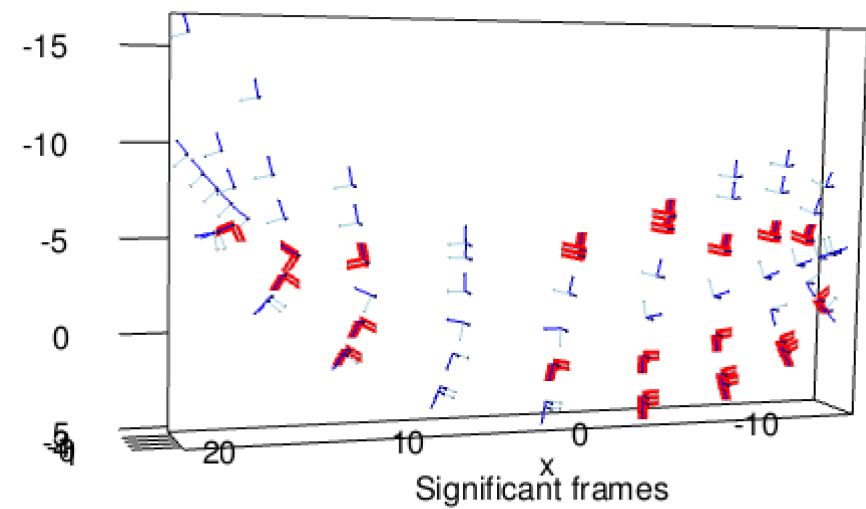
vs LP-ds-rep

Frames' directions

Before adjustment

After adjustment

Unconstrained
sphereGLM



LP-ds-rep

