Exercises: Week 4

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2021-02-25

Problem 1 (Coding Exercise)

This exercise asks you to implement and assess the performance of the bootstrap for the linear regression model. Suppose you have the linear regression model:

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

where,

- $x_i \sim U[0, 2]$
- $\epsilon_i | x_i \sim U[-1, 1]$
- $\beta_0 = \beta_1 = 1$

We ask you to answer the following questions:

1. Write a code that generates i.i.d. samples of sizes n=10,50,200 from that distribution, computes (1) the least squares estimator for β , (2) the t-ratio for the least squares coefficient β_1 , $t_n = \frac{\hat{\beta}_{1,LS} - 1}{s.\hat{e}.(\hat{\beta}_{1,LS})}$, and (3) the least square residuals $\hat{\epsilon}_i = y_i - \hat{\beta}_{0,LS} - \hat{\beta}_{1,LS}x_i$.

```
# Easier with true beta_1 = 0
sim sample \leftarrow function(n, betas = c(1, 0)){
  d \leftarrow tibble(x = runif(n, 0, 2),
                eps = runif(n, -1, 1),
                y = betas[1] + betas[2] * x + eps)
  d
}
ols <- function(d){</pre>
  n \leftarrow nrow(d)
  y_  <  d\$y - mean(d\$y)
  x_ < - d$x - mean(d$x)
  b1 <- sum(x_ * y_) / sum(x_ * x_)
  b0 \leftarrow mean(d\$y) - (b1 * mean(d\$x))
  e \leftarrow d\$y - b0 - b1 * d\$x
  t_b1 \leftarrow b1 / (sum(e^2) / (n - 2))
  pars <- c(n, b0, b1, t_b1) %% set_names(c("n", "b0", "b1", "t_b1")) %>% round(2)
  pars
}
```

```
d10 <- sim_sample(10)</pre>
d10 %>% ols
       n
            b0
                b1 t_b1
## 10.00 0.91 -0.09 -0.15
d50 \leftarrow sim_sample(50)
d50 %>% ols
                  b1 t_b1
       n
            b0
## 50.00 0.77 0.23 0.52
d200 <- sim_sample(200)
d200 %>% ols
##
              b0
                      b1
                           t_b1
            0.92
## 200.00
                    0.03
                           0.09
```

- 2. Write a code for drawing n times at random from the discrete uniform distribution over the estimated residuals $\hat{\epsilon}_1,...,\hat{\epsilon}_n$ (i.e. with replacement).
- 3. Use your code from parts (a) and (b) to implement the residual bootstrap assuming that ϵ_i and x_i are independent to estimate the 95th percentiles of the respective distributions of $\hat{\beta}_{1,LS}$ and t_n

```
bs_par <- function(data) {</pre>
  d <- data
  boots <- bootstraps(d, times = nrow(d), apparent = T)</pre>
  lm_bs <- function(splits) {</pre>
    lm(formula(y ~ x), analysis(splits))
  boot_models <-
    boots %>%
    mutate(model = map(splits, lm_bs),
           coef_info = map(model, tidy))
  boot_coefs <-
    boot_models %>%
    unnest(coef_info) %>%
    filter(term != "(Intercept)")
  boot_coefs
  }
rbind(bs_par(d10) %>% select(estimate, statistic) %>% summarise(across(everything(), quantile, 0.95)),
      bs_par(d10) %>% select(estimate, statistic) %>% summarise(across(everything(), quantile, 0.95)),
      bs_par(d10) %>% select(estimate, statistic) %>% summarise(across(everything(), quantile, 0.95))
      ) %>%
  rename("b1 95%" = estimate, "t 95%" = statistic) %>%
  round(2) %>%
  mutate(sim = c(10, 50, 200)) \%\%
  kableExtra::kbl(booktabs = T)
```

| b1 95% | t 95% | $_{ m sim}$ |
|--------|-------|-------------|
| 0.06 | 0.14 | 10 |
| 0.49 | 2.05 | 50 |
| 0.23 | 1.02 | 200 |

4. Repeat part (a) for sample size n=10, 50, 200 with 200 replications, where you keep the initial draws of $x_1, ..., x_n$ from part (a) and only generate new residuals from their conditional distribution. Compute $\hat{\beta}_{1,LS}$ and the statistic t_n using 200 independent samples of size n. Use your results to compute a simulated estimate for the 95th percentiles of the respective sampling distributions for $\hat{\beta}_{1,LS}$ and t_n .

```
bs_par <- function(data, reps = 200) {</pre>
  d <- data
  boots <- bootstraps(d, times = reps, apparent = T)</pre>
  lm bs <- function(splits) {</pre>
    lm(formula(y ~ x), analysis(splits))
  }
  boot_models <-
    boots %>%
    mutate(model = map(splits, lm_bs),
           coef_info = map(model, tidy))
  boot_coefs <-
    boot_models %>%
    unnest(coef_info) %>%
    filter(term != "(Intercept)")
  boot_coefs
rbind(bs_par(d10) %>% select(estimate, statistic) %>%
                                                        summarise(across(everything(), quantile, 0.
      bs_par(d10) %>% select(estimate, statistic) %>% summarise(across(everything(), quantile, 0.95)),
      bs_par(d10) %>% select(estimate, statistic) %>% summarise(across(everything(), quantile, 0.95))
      ) %>%
  rename("b1 95%" = estimate, "t 95%" = statistic) %>%
  round(2) %>%
  mutate(sim = c(10, 50, 200)) \%
  kableExtra::kbl(booktabs = T)
```

| b1 95% | t 95% | $_{ m sim}$ |
|--------|-------|-------------|
| 0.45 | 1.97 | 10 |
| 0.58 | 2.19 | 50 |
| 0.60 | 2.02 | 200 |

5. Compare your results from (c) and (d). What do you conclude about the performance of the bootstrap? How does it compare to the 95th percentile of the asymptotic distribution of t_n ?

I'm not sure here: it seems like that the low number of bootstraps makes anything possible. The asymptotic distribution of t_n follows a standard normal distribution: $\Phi^{-1}(0.95) = 1.645$.

6. See if you can construct a subsampled confidence interval for n = 200 with $a_n = 25$. How does it compare to the bootstrapped CI?