

# Modelling photosynthetic parameters via mixed effects models and the Farquhar-van Cammerer-Berry model

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## Abstract

The study of the photosynthetic process frequently involves analysis of net assimilation - intercellular CO<sub>2</sub> concentration (ACI) curves. These are used in the estimation of key parameters associated with the Farquhar-van Cammerer-Berry (FvCB) [8] model:

- $V_{c_{max}}$ : the rate of maximum Rubisco carboxylation;
- $J$ : electron transport rate;
- $R_d$ : daytime respiration; and
- $g_m$ : mesophyll conductance.

Accurate, unbiased estimation of these parameters is a non-trivial exercise with the optimal method still a matter of debate within the field [13][25][15]. The problem of selecting suitable starting values for a non-linear model is thus complicated by the lack of unanimity on what is a “good” estimate of these. The core model we are using to estimate the parameters is based on Yin et al. [26]. This method offers several steps to estimate various parameters rather than attempting estimation in a single step.

The data collected for the purpose of modelling steady-state photosynthesis is inevitably repeated measurements on individual plants. This leads to correlation between measurements, a fact that violates the assumptions in the fixed effects models used in estimating the parameters of interest ([4][2]). We introduce the use of non-linear mixed effects models in keeping with Qian et al. [15] to overcome this issue, and extend on their work by considering more modern versions of the FvCB model.

**Keywords**— ACI curves, parameter estimation, FvCB model, non-linear mixed effects models

## 1 Theory

### 1.1 FvCB model

The brilliant realisation of Farquhar et al. [8] was that the net photosynthesis rate,  $A$ , of C<sub>3</sub> plants could be accurately modeled at the leaf level as being in one of two steady-state systems. In *Rubisco activity limited net photosynthesis rate* ( $A_c$ ), the net rate of photosynthesis is limited by the properties of the enzyme ribulose 1.5 biphosphate carboxylase/oxygenase (Rubisco) (we assume saturating levels of the substrate, RuBP). In the second we assume that it is the substrate that determines the rate of assimilation; specifically its regeneration rate. Hence the second state is referred to as RuBP-regeneration limited or *electron transport limited net photosynthesis rate* ( $A_j$ )[26]. We prefer the latter name for being a more clear difference for readers unfamiliar with the biochemical constituents of the photosynthesis cycle.

### 1.1.1 Some assumptions

For parsimony's sake, we make several simplifying assumptions regarding the homogeneity of conditions within the leaf. We disregard any gradient of temperature ( $T$ ) or  $O_2$  within the leaf. According to von Caemmerer [21], these, specifically within  $C_3$  species, are unlikely to be important. Due to the thinness of leaves we also assume an equal distribution of light to all chloroplasts within the leaf; however, it has been shown that under certain conditions these assumptions and their corollary, that photosynthesis itself is homogeneous across the leaf, do not hold. Non-uniform photosynthesis can occur and this will affect the gas-exchange measurements and their interpretation [20].

We account for the *Kok effect* in our modelling. This is a physiological effect where in low light (typically  $0 - 20 \mu\text{mol m}^{-2} \text{s}^{-1}$  [19]) conditions the quantum yield of net photosynthesis is higher than at higher levels of incident light. Thus, the net assimilation recorded tends to escape description by our model. This is a well documented effect [19][18][12] and hence in some the low light data the first point (we will often consider increasing intercellular carbon dioxide ( $C_i$ ) or light intensity ( $I$ ) as pseudo-time in our descriptions) will be ignored in our models.

Several parameters require estimation, and some measurements are not feasible for the majority of experiments. For these we need either a starting value that is our initial best guess (feasible in  $C_3$  species, less so in  $C_4$ ) or a correction to some related measurement. We will show the reasoning behind our choices below.

### 1.1.2 Rubisco activity limited net photosynthesis rate

The photosynthetic carbon reduction (PCR) and photorespiratory carbon oxidation (PCO) cycles are linked by an enzyme common to both, RuBP [8]. In the PCO cycle,  $0.5 \text{ mol}$  of  $\text{CO}_2$  is released, thus:

$$A = V_c - 0.5V_o - R_d \quad (1)$$

Where:

- $V_c$  is the rate of carboxylation; and
- $V_o$  is the rate of oxygenation.

We have from Farquhar [9] (assuming RuBP binding first in the reaction):

$$V_c = V_{c_{max}} \frac{C_c}{C_c + K_c (1 + O/K_o)} \cdot \frac{R}{R + K'_r} \quad (2)$$

$$V_o = V_{o_{max}} \frac{O}{O + K_o (1 + C_c/K_c)} \cdot \frac{R}{R + K'_r} \quad (3)$$

Where:

- $V_{c_{max}}$  is the maximum velocity of the carboxylase;
- $V_{o_{max}}$  is the maximum velocity of the oxygenase;
- $O$  is the partial pressure of oxygen in the chloroplast;
- $K_c$  is the Michaelis-Menten constant for  $\text{CO}_2$ ;
- $K_o$  is the Michaelis-Menten constant for  $O$ ;
- $R$  is the concentration of free RuBP; and
- $K'_r$  is the effective Michaelis-Menten constant for RuBP.

This leads to the ratio of oxygenation to carboxylation:

$$\phi = \frac{V_o}{V_c} = \frac{V_{o_{max}}}{V_{c_{max}}} \cdot \frac{O/K_o}{C_c/K_c} \quad (4)$$

We also have that the velocity of carboxylation, at saturating levels of RuBP, ( $W_c$ ) is given by [10]:

$$W_c = V_{c_{max}} \frac{C_c}{C_c + K_c(1 + O/K_o)} \quad (5)$$

Note that this arises from setting the RuBP component of (2) to 1. The compensation point for  $C_c$  at which there is no net assimilation,  $\Gamma^*$ , is given by [8]:

$$\Gamma^* = \frac{K_c O}{2K_o} \cdot \frac{V_{o_{max}}}{V_{c_{max}}} \quad (6)$$

Relating (1), (4), (5) and (6) we have:

$$A_c = V_{c_{max}} \frac{C_c - \Gamma^*}{C_c + K_c(1 + O/K_o)} - R_d \quad (7)$$

This is the Rubisco-limited rate of assimilation, assuming saturation of RuBP.

### 1.1.3 Electron transport limited net photosynthesis rate

The physiology of this model is somewhat more complex, involving analysis of the cycle for the production of NADPH [8]. A thorough understanding of this is beyond the scope of this project, but we recommend the reader to the original paper [8] and von Caemmerer [21]. We are content to state that when  $A$  is limited by RuBP regeneration:

$$A_j = J \frac{C_c - \Gamma^*}{4C_c + 8\Gamma^*} - R_d \quad (8)$$

Where  $J$  is the rate of electron transport (and hence electron transport limited net photosynthesis rate). This equation arises using assumptions that RuBP regeneration is limited because of insufficient NADPH and implicitly assumes 100% linear  $e^-$  transport. Sharkey et al. [17] recommend the use of 4 and 8 in the denominator based on the number of electrons required for  $\text{NADP}^+$  reduction; however this is not always the case; values of 4.5 and 10.5 also occur in the literature and less frequently we see 4 and 9.33 (see [17][26][24]). These different values arise from different assumptions about the complexity of linear electron ( $e^-$ ) rate. The disagreement over pairs  $\{4.5, 10.5\}$  or  $\{4, 9.33\}$  arise due to further disagreement about deeper levels of this system that are a source of some discussion in the literature ([21][26]) but are less of interest to us. To allow for this disagreement we use default values of 4 and 8 in our R model, but the functions incorporating this aspect of the model allow the user to input their own version of these coefficients.

### 1.1.4 TPU limited net photosynthesis rate

Other models extend to different cases (such as  $C_4$  species) or to encompass more physiologically limiting factors (such as triose phosphate utilization ( $TPU$ ) in the synthesis of starch and sucrose).  $TPU$  is required at one third the rate of  $\text{CO}_2$  fixation. If this is not acquired, the level of free phosphate declines and the rate of photosynthesis is limited. When this limit is imposed, it can be seen that photosynthesis becomes independent of the ratio of oxygenation to carboxylation [16]. Sharkey et al. [17] model this *TPU limited net photosynthesis rate* ( $A_p$ ) using the following equation:

$$A_p = 3TPU - R_d \quad (9)$$

Where  $TPU$  is the rate of triose phosphate utilization. This extension is only relevant under high  $C_i$  and  $I$  environments and is normally modeled as an upper bound on  $A$ . We do not include this in our model, but recommend extending the model to this scenario in later versions. From von Caemmerer [21] we know that this model is too simple for the reality, please see section 2.4.3 for an extension as formulated by Harley and Sharkey [11]. As (9) is only apparent in extreme cases and we ignore it for now but recommend inclusion of this additional case in any extensions to our work.

### 1.1.5 FvCB summary

For the final model, our actual assimilation rate is calculated by combining these individual curves. The curve fitted is:

$$A = \min \{ A_c, A_j \} \quad (10)$$

Where:

- $A_c$  is the Rubisco activity limited net photosynthesis rate calculated from (7); and
- $A_j$  is the electron transport limited net photosynthesis rate calculated from (8).

### 1.1.6 Relating $C_i$ to $C_c$

One of the driving factors in net assimilation rate ( $A$ ) is the level of  $\text{CO}_2$  present at the Rubisco site of carboxylation in the chloroplasts ( $C_c$ ) [22]. Estimation of the intercellular  $\text{CO}_2$  partial pressure ( $C_i$ ) is based on gas exchange measurements and is not a matter of controversy. Less well established is estimation of  $C_c$  [25]. We can relate  $C_i$  to  $C_c$  using Flick's first law relating the diffusive flux to the concentration under steady-state assumptions. We have:

$$J = -D \frac{d\phi}{dx} \quad (11)$$

Where:

- $J$  is the diffusion flux;
- $D$  is the diffusion coefficient;
- $\phi$  is the concentration; and
- $x$  is position.

Please note that the notation used in (11) is the generic notation used normally in describing Flick's Law and will not be referring to these terms after this subsection. From this we can consider  $A$  as our flux, and the difference between  $C_c$  and  $C_i$  as our gradient in concentration:

$$A = -D(C_c - C_i) \quad (12)$$

The diffusion conductance between the substomatal cavities and the chloroplasts is  $g_m$ , the mesophyll conductance [14]. Hence we can write (12) as:

$$A = g_m(C_i - C_c) \quad (13)$$

Or equivalently:

$$C_c = C_i - \frac{A}{g_m} \quad (14)$$

Thus we can calculate  $C_c$ , the variable relevant to our models, using  $C_i$ .

### 1.1.7 Light dependence of electron transport rate

We can relate  $J_{max}$  to the incident irradiance by empirical equation:

$$\theta J^2 - J(I_2 + J_{max}) + I_2 J_{max} = 0 \quad (15)$$

Where:

- $I_2$  is the useful light absorbed by PS II;

- $J$  is the electron transport;
- $J_{max}$  is the maximum electron transport; and
- $\theta$  is an empirical curvature factor (often around 0.7 [6]).

$I_2$  is related to total incident irradiance  $I$  (sometimes referred to as the Photosynthetically Active Radiation,  $PAR$ , in the literature) by:

$$I_2 = I \times \text{abs}(1 - f)/2 \quad (16)$$

Where:

- $\text{abs}$  is the absorptance of leaves (commonly around 0.85 [21]);
- $f$  is the correction for the spectral quality of light (approximately 0.15 [7]); and
- The 2 in the denominator is due to the split of light between photosystems I and II.

We can thus solve for  $J$  in the usual way ([15] [24]):

$$J = \frac{I_2 + J_{max} - \sqrt{(I_2 + J_{max})^2 - 4\theta \cdot I_2 \cdot J_{max}}}{2\theta} \quad (17)$$

Equations of this form are referred to as *non-rectangular hyperbola functions*.

### 1.1.8 Temperature dependency

The dependency of the rate of carboxylation and oxygenation of Rubisco is reflected in the temperature dependency of  $A$ . As a rate dependent upon temperature, we will be using Arrhenius functions, equations of the form:

$$x(T) = k \cdot \exp\left(-\frac{E_a}{R \cdot T}\right) \quad (18)$$

Where:

- $x(T)$ : the temperature dependent rate we are interested in;
- $k$ : the rate coefficient;
- $E_a$ : the activation energy, the kinetic energy of substrate required for the reaction to occur;
- $R$ : the universal gas constant ( $8.314 \text{ J K}^{-1} \text{ mol}^{-1}$ ); and
- $T$ : temperature (in Kelvin).

As most photosynthesis measurement are recorded in  $^{\circ}\text{C}$ , and specifically use a default value of  $25^{\circ}\text{C}$ , we transform the equation to this scale. In this case we have:

$$x(T') = k' \cdot \exp\left(-\frac{(25 - T')E_a}{298.15R \cdot (273.15 + T')}\right) \quad (19)$$

Where  $T'$ ,  $k'$  are the relevant form of  $T$ ,  $k$  for a scale centred on  $T = 25^{\circ}\text{C}$ . As we will not be using the Kelvin scale again, we denote  $T = T'$ ,  $k = k'$ . Many reactions in photosynthesis are reversible, with differing activation energies depending on the direction of the reaction. Thus the net activation energy may vary depending on the ratio of forward to backward reactions. This means that the Arrhenius function (19) is only semi-empirical, but it does allow easy comparison between studies.

Another frequently used method to describe a dependency on temperature is the  $Q_{10}$  *temperature coefficient*. This is a measure of the rate of change of a system as a result of raising the temperature by  $10^{\circ}\text{C}$ . It is calculated:

$$Q_{10} = \left(\frac{R(T_2)}{R(T_1)}\right)^{10^{\circ}\text{C}/(T_2 - T_1)} \quad (20)$$

Where:

- $Q_{10}$ : the factor by which the reaction rate increases when temperature is raised by 10 °C;
- $T_i$ : the temperature for the  $i^{th}$  measurement; and
- $R(T_i)$ : the rate of the reaction at temperature  $T_i$ .

Alternatively we can write this in the form:

$$R(T_2) = R(T_1)Q_{10}^{(T_2-T_1)/10^\circ\text{C}} \quad (21)$$

Again, this general form is of less interest as we have a specific, default temperature of 25 °C for which we are quite well informed. Hence, we use:

$$R(T) = R(25^\circ\text{C})Q_{10}^{(T-25)/10} \quad (22)$$

This function allows us to relate values given at the temperature  $T = 25^\circ\text{C}$  to more general temperatures.

### 1.1.9 Model parameters

Many parameters can be assigned *a priori*, leaving only the estimation of a small number of key variables. The kinetic constants of rubisco vary very little among  $C_3$  species such that one can use the same  $K_c$ ,  $K_o$  and  $\Gamma^*$  across all members of this category [21]. This means that the only parameter requiring estimation for Rubisco is the maximal Rubisco activity,  $V_{c_{max}}$ .

For RuBP-limited photosynthesis, we need to solve for  $J_{max}$  as we can relate this to  $J$  by equation (17). According to Walcroft et al. [23], the ratio  $J_{max} : V_{c_{max}}$  is expected to vary from 2 to 1.4 across the range  $[8^\circ\text{C}, 30^\circ\text{C}]$ .

Evans and von Caemmerer [5] recommend setting  $g_m = 0.0045V_{c_{max}}$ . Other values can be seen, with an assumption of  $g_m = \infty$  used frequently, but this is controversial and can lead to biased estimates of  $V_{c_{max}}$  and  $J_{max}$  [25]. Specific plants in specific conditions can see diverging values, but we feel this is a sufficient initial value.

This means that if we have an empirically driven initial value for  $V_{c_{max}}$  we can initialise all our parameters to some default temperature, however these relationships vary with temperature, the relationship of which we state below.

We use table 1 for our initial values of photosynthetic parameters in our model. We can combine these with (22) to calculate initial values for our model across a range of temperatures. Many of the values in table 1 can be found in von Caemmerer [21].

The assumption that  $g_m = \infty$ , equivalent to  $C_c = C_i$  is now considered redundant but would severely change these values. We include a sample of this impact to enable comparison with older papers and research.

## 1.2 Mixed-effects models

### 1.2.1 Linear mixed-effects models

Traditional parametric models incorporate only *fixed effects*. That is, they have a set of parameters describing with the entire population. For example, consider a system of  $n \in \mathbb{N}$  observations of dependent variable,  $Y = (y_1, \dots, y_n)$ , and the  $n \times p$  matrix of associated independent variable measurements,  $X = \{x_{i,j}\}$  for  $i \in [1, n]$ ,  $j \in [1, p]$  for some  $p \in \mathbb{N}$ , and a  $p$ -vector of weights,  $\beta$ , represented:

$$Y = X\beta + \epsilon \quad (23)$$

Here  $X$  is assumed to contain a column of 1's in the first position (hence the +1 in the description of its dimensionality) and  $\epsilon$  is the  $n$ -vector of associated errors. We assume that  $\epsilon_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$  for

Parameter	unit	Value	E (kJ mol <sup>-1</sup> )	Q <sub>10</sub> (T = 25 °C)
K <sub>c</sub>	μbar	260 or 404 <sup>a</sup>	59.36 <sup>b</sup>	2.24
K <sub>o</sub>	mbar	179 or 248 <sup>a</sup>	35.94	1.63
Γ*	μbar	38.6 or 37 <sup>a</sup>	23.4	1.37
V <sub>cmax</sub>	μmol m <sup>-2</sup> s <sup>-1</sup>	80	58.52	2.21
V <sub>o,max</sub>	μmol m <sup>-2</sup> s <sup>-1</sup>	0.25 × V <sub>cmax</sub>	58.52	2.21
R <sub>d</sub>	μmol m <sup>-2</sup> s <sup>-1</sup>	0.01 – 0.02 × V <sub>cmax</sub>	66.4	2.46
J <sub>max</sub>	μmol m <sup>-2</sup> s <sup>-1</sup>	1.4 – 2.0 × V <sub>cmax</sub>	37	1.65
g <sub>m</sub>	μmol m <sup>-2</sup> s <sup>-1</sup>	0.0045V <sub>cmax</sub> <sup>c</sup> or ∞		
θ		0.7		

<sup>a</sup> The first value is appropriate when an internal diffusion conductance is included; the second value should be used if the internal conductance is not included (i.e. g<sub>m</sub> = ∞) and C<sub>c</sub> is assumed to equal C<sub>i</sub>.

<sup>b</sup> From Farquhar et al. [8]

<sup>c</sup> From Evans and von Caemmerer [5]

Table 1: Photosynthetic parameters and their activation energy for T = 25 °C

$i = (1, \dots, n)$ . If we consider the intuitive case of *biological* and *technical replicates*, where we have  $n$  biological replicates and for the  $i^{\text{th}}$  sample  $m_i$  associated technical replicates. As each of the  $m_i$  measurements is on the same sample we expect there to be a non-negligible correlation between the  $m_i$  technical replicates for each  $i$ . The model in (23) does not allow for the within group effects due to the correlation between technical replicates. Consider the simplest possible model for the fixed effects model including only the intercept:

$$y_{ij} = \beta + \epsilon_{ij}, \quad i = 1, \dots, n, \quad j = 1, \dots, m_i \quad (24)$$

Some of this model's limitations become apparent if one considers the case of unbalanced data. Continuing the previous example of biological and technical replicates, consider the case that  $m_i \neq m_j$  for any  $i, j \in (1, \dots, n)$ . In this case the model is skewed by the within sample data rather than by the true observations, the biological replicates. A possible solution is the inclusion of a individual intercept for each group of technical replicates, accommodating the within-sample variability, or *random effects*:

$$y_{ij} = \beta_i + \epsilon_{ij}, \quad i = 1, \dots, n, \quad j = 1, \dots, m_i \quad (25)$$

While this better describes the observed data, it has certain inherent flaws; most notably the number of parameters scales linearly with the number of observations and the model only describes the measurements included in the sample. Consider as a solution a combination of these models, containing both a sample mean,  $\beta$ , and a random variable for each group representing the deviation from the population mean,  $b_i$ , i.e. a *mixed effects* model:

$$y_{ij} = \beta + b_i + \epsilon_{ij}, \quad i = 1, \dots, n, \quad j = 1, \dots, m_i \quad (26)$$

The linear mixed effects model described in (26) contains information at both a population level (in the fixed effects) but also at the individual level (in the random effects).

For now we assume  $b_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma_b^2)$  for  $i = 1, \dots, n$ . This means that the variance of the observations is divided into two parts,  $\sigma_b^2$  for the biological variability and  $\sigma^2$  for the technical variability:

$$b_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma_b^2), \quad \epsilon_{ij} \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2) \quad (27)$$

The assumption of normality can be modified if deemed inappropriate and it is possible to generalise the model to allow for heteroscedasticity.

The  $b_i$  are called *random effects* as they are associated with the experimental unit and selected at random from the population of interest (at least in theory, obviously there are limitations on this particularly in the area of medicine) They represent that the effect of choosing the sample  $i$  is to shift the mean expression of  $Y$  from  $\beta$  to  $\beta + b_i$  - i.e. they effect a deviation from an overall mean. Technical replicates share the same random effect  $b_i$  and are correlated. The covariance between technical replicates on the same experimental unit is  $\sigma_b^2$ ; this corresponds to a correlation of  $\frac{\sigma_b^2}{(\sigma_b^2 + \sigma^2)}$ .

### 1.2.2 Non-linear mixed effects models

Non-linear mixed effects models, also known as *non-linear hierarchical models*, are an extension to the more traditional linear mixed effects models. They are used in scenarios where all of the following features are present [3]:

1. Repeated observations of a continuous variable on each of several *experimental units* (in our case these are individuals, thus individual is considered equivalent to empirical unit in the following section) over time or another condition (e.g. measurements at given heights on a tree) (the *condition variable*);
2. We expect the relationship between the response variable and the condition variable to vary across individuals; and
3. Availability of a scientifically relevant model characterising the behaviour of the individual response in terms of meaningful parameters that vary across individuals and dictate variation in patterns of condition-response (for us this will be the Farquhar-van Cammerer-Berry model).

The final point from the list in 1.2.2 is where the non-linear aspect is introduced. This is often a mechanistic function describing a physical or chemical system (for example in toxicokinetics physiologically-based pharmacokinetics models are used or HIV dynamics in “precision medicine”); that is the model is described by meaningful, interpretable parameters rather than being an empirical best fit. It is expected that the mechanistic model will better describe data beyond the range of the measurements used here (whereas an empirical fit might describe the data over the range captured in the measurement data but then misbehave woefully beyond these boundaries).

The analysis tends to have the goal of understanding one or more of the following:

1. The “typical” behaviour of the phenomena (i.e. mean or median values) represented by the model parameters;
2. The variation of these parameters, and hence the phenomena, between individuals; and
3. If some of the variation is inherently associated with individual characteristics.

Individual level prediction can also be of interest (e.g. in medical treatment with highly individual reaction), but is less relevant to this project. From the individual level we are interested in investigating the level of variation between individuals and questioning if this is sufficiently small to allow the “all-purpose” models generally used in photosynthesis describing the parameters and systems of interest.

Consider an experiment involving repeated measurements of some response variable,  $Y$ , across a condition variable  $T$  for  $n$  individuals. Each individual’s characteristics are recorded in  $A$  (this is assumed to be time-independent measurements, like an individual’s height over the course of a drugs trial) and possible additional initial conditions in  $U$  (for example the dose of a drug the individual was receiving at time  $t = 0$ ). Let  $y_{i,j}$  denote the  $j$ th measurement of the response under



condition  $t_{i,j}$ ,  $j = 1, \dots, n_i$ , for individual  $i$ ,  $i = 1, \dots, n$  with additional initial conditions  $u_i$ . Thus:

$$\begin{aligned} Y &= (y_1, \dots, y_n) \\ y_i &= (y_{i,1}, \dots, y_{i,n_i}) \\ T &= (t_1, \dots, t_n) \\ t_i &= (t_{i,1}, \dots, t_{i,n_i}) \\ A &= (a_1, \dots, a_n) \\ U &= (u_1, \dots, u_n) \end{aligned} \tag{28}$$

We denote  $x_{i,j} = (t_{i,j}, u_i)$ . Often  $T$  is time and  $U = \emptyset$ , but it might be the case that each  $t_{i,j}$  is a  $p_1$ -vector of measurements and  $u_i$  is a  $p_2$ -vector such that  $p_1 + p_2 = p$ , returning to the notation of (26). The assumption that the triplets  $(y_i, u_i, a_i)$  are independent across  $i$  is often included to reflect the belief that individuals are unrelated (this will hold for us, but might require more thought in other situations). For some function  $f$  regulating the within-individual behaviour defined by a vector of parameters  $\beta_i$  unique to individual  $i$ , we have:

$$y_{i,j} = f(x_{i,j}, \beta_i) + \epsilon_{i,j}, \quad j = 1, \dots, n_i \tag{29}$$

For us this will be the Farquhar-van Cammerer-Berry model of steady-state photosynthesis. We assume that  $\mathbb{E}(\epsilon_{i,j} | u_i, \beta_i) = 0$  and  $\epsilon_{i,j} \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$  for all  $i, j$  (the assumption of normality can be relaxed, but this extension exceeds the reach of this project). This model in (29) is called the *individual level model*. To model the population parameters we consider  $d$ , a  $p$ -dimensional function depending on an  $r$ -vector of fixed parameters, or *fixed effects*,  $\beta$ , and a  $k$ -vector of *random effects*,  $b_i$ , associated with individual  $i$ :

$$\beta_i = d(a_i, \beta, b_i), \quad i = 1, \dots, n \tag{30}$$

Here, the *population model* in (30) describes how  $\beta_i$  varies among individuals due to both individual attributes  $a_i$  and biological variation in  $b_i$ . We assume that the  $b_i$  are independent of the  $a_i$ , i.e.:

$$\begin{aligned} \mathbb{E}(b_i | a_i) &= \mathbb{E}(b_i) = 0 \\ \mathbb{V}ar(b_i | a_i) &= \mathbb{V}ar(b_i) = D \end{aligned} \tag{31}$$

Here,  $D$  is an unstructured covariance matrix and is common to all  $i$ . It characterises the degree of unexplained variation in the elements of  $\beta_i$  and associations among them; the ubiquitous assumption is  $b_i \sim \mathcal{N}(0, D)$ . However, if this set of assumptions regarding the conditional distribution of  $b_i$  on  $a_i$  is found to be insufficient, then  $b_i \sim \mathcal{N}(0, D(a_i))$  is frequently used.

In (30),  $\beta_i$  is considered to have an associated random effect, reflecting the belief that each component varies non-negligibly in the population even after systematic relationships with subject characteristics are accounted for. It may happen that “unexplained” variance in a component of  $\beta_i$  may be very small in magnitude relative to that in the remaining elements. In this situation it is common to drop the negligible quantity entirely. This lacks biological sense as each parameter is part of the “scientifically relevant model” and thus is unlikely to have no associated unexplained variation. Hence, one must recognise that this omission of an element of  $\beta_i$  is adopted to achieve numerical stability in fitting rather than to reflect belief in perfect biological consistency across individuals and analyses in the literature to determine whether elements of  $\beta_i$  are fixed or random effects should be interpreted so.

### 1.3 Our model

Much of the preceding concepts and theory make for very pleasant reading, but we consider it useful to explicitly state our model and its assumptions in one place.

First, recall the FvCB model (10):

$$A = \min \{A_c, A_j\} \quad (32)$$

$$= FvCB(C_c, \Gamma_*, I, K_c, K_o, O_{i,j}, J_{max}, V_{c_{max}}, \alpha, \theta, R_{d_i}) \quad (33)$$

Where  $A_c$  and  $A_j$  are two curves describing separate biochemical limits on the rate of photosynthesis:

$$A_c = V_{c_{max}} \frac{C_c - \Gamma_*}{C_c + K_c(1 + O/K_o)} - R_d \quad (34)$$

$$A_j = J \frac{C_c - \Gamma_*}{4C_c + 8\Gamma_*} - R_d \quad (35)$$

Furthermore:

$$J = \frac{\alpha \cdot I + J_{max} - \sqrt{(\alpha \cdot I + J_{max})^2 - 4\theta \cdot \alpha \cdot I \cdot J_{max}}}{2\theta} \quad (36)$$

We plug this into (29) for each of our experimental effects (in Rachel Schipper's data, SIDE and TREATMENT) in line with Bates and Pinheiro [1]. Thus, we have:

- Our response variable,  $Y$ , is net photosynthesis rate,  $A$ ;
- Our individual level function,  $f(\cdot)$ , is  $FvCB(\cdot)$ ;
- Our measured condition variables,  $X$ , are  $C_c, \Gamma_*, I, K_c, K_o$ , and  $O$ ; and
- Our fixed effects,  $\beta$ , are  $J_{max}, V_{c_{max}}, \alpha, \theta$  and  $R_d$ .

Notice that  $U = \emptyset$ . Therefore our individual-level model is:

$$\begin{aligned} A_{i,j} &= FvCB(C_{c_{i,j}}, \Gamma_{*_{i,j}}, I_{i,j}, K_c, K_o, O_{i,j}, J_{max_i}, V_{c_{max_i}}, \alpha_i, \theta_i, R_{d_i}) + \epsilon_{i,j} \\ \epsilon_{i,j} &\stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2) \\ \mathbb{E}(\epsilon_{i,j} | J_{max_i}, V_{c_{max_i}}, \alpha_i, \theta_i, R_{d_i}) &= 0 \end{aligned} \quad (37)$$

For our specific version of (30), we cannot know in advance what kind of model we will use. This will be based on the ability of different models to converge. We expect, given the paucity of data that we will have a diagonal covariance structure with 0's in the non-diagonal entries imposed on our mixed effects model (as we expect that there is not enough data for more flexible formats to converge). This corresponds to a statement of independence between random effects which may not be the ideal assumption, and with this in mind we will attempt other formats. However we do know that our photosynthetic parameters of interest ( $J_{max}, V_{c_{max}}, \alpha, \theta$  and  $R_d$ ) will all be represented in this aspect of the model to consider each individual plant with unique photosynthetic parameters.

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