Module 7: Moving Beyond Linearity

TMA4268 Statistical learning

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Multiple linear regression

$$y_i=\beta_0+\beta_1x_{i1}+\beta_2x_{i2}+\dots\beta_kx_{ik}+\varepsilon_i,$$
 or equivalently

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon} =$$

$$\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} 1 & x_{11} & x_{12} & \dots & x_{1k} \\ 1 & x_{21} & x_{22} & \dots & x_{2k} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & x_{n2} & \dots & x_{nk} \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}.$$

Estimation

The OLS estimator for β is

$$\hat{\beta} = (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \mathbf{y}.$$

We will now change \mathbf{X} as we like, but keep $\hat{\beta}$.

Non-Linear Models

Let us focus on **one** explanatory variable X for now. We will generalize later.

$$y_i = \beta_0 + \beta_1 b_1(x_i) + \beta_2 b_2(x_i) + \dots + \beta_k b_k(x_i) + \varepsilon_i,$$

where $b_j(x_i)$ are basis functions.

Example with $y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \varepsilon_i$

We have

$$b_1(X) = X,$$

 $b_2(X) = X^2,$
 $\mathbf{x} = (6, 3, 6, 8)^T,$
 $\mathbf{y} = (3, -2, 5, 10)^T.$

This results in

$$\mathbf{X} = \begin{pmatrix} 1 & b_1(x_1) & b_2(x_1) \\ 1 & b_1(x_2) & b_2(x_2) \\ 1 & b_1(x_3) & b_2(x_3) \\ 1 & b_1(x_4) & b_2(x_4) \end{pmatrix} = \begin{pmatrix} 1 & x_1 & x_1^2 \\ 1 & x_2 & x_2^2 \\ 1 & x_3 & x_3^2 \\ 1 & x_4 & x_4^2 \end{pmatrix} = \begin{pmatrix} 1 & 6 & 36 \\ 1 & 3 & 9 \\ 1 & 6 & 36 \\ 1 & 8 & 64 \end{pmatrix}$$

and

$$\hat{\beta} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{y} = (-4.4, 0.2, 0.2)^{\mathsf{T}}.$$

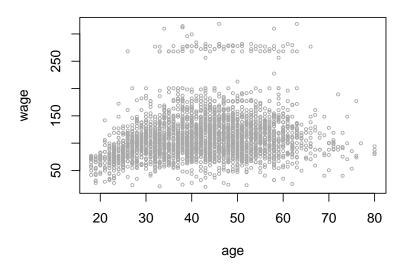
General Design Matrix

$$\mathbf{X} = \begin{pmatrix} 1 & b_1(x_1) & b_2(x_1) & \dots & b_k(x_1) \\ 1 & b_1(x_2) & b_2(x_2) & \dots & b_k(x_2) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & b_1(x_n) & b_2(x_n) & \dots & b_k(x_n) \end{pmatrix}.$$

- Rows are observations
- Columns are basis functions
- Same setup as for multiple linear regression

The Aim

Observations



Use lm(wage ~ X) and choose **X** according to method.

Polynomial Regression

The polynomial regression includes powers of X in the regression.

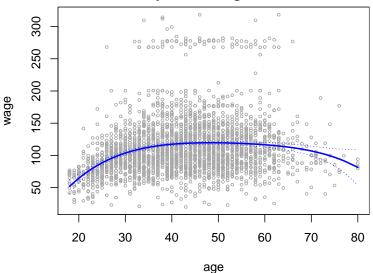
$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots \beta_k x_i^d + \varepsilon_i,$$

- ▶ In practice $d \le 4$
- ► The basis is $b_j(x_i) = x_i^j$ for $j = 1, 2 \dots, d$

$$\mathbf{X} = \begin{pmatrix} 1 & x_1 & x_1^2 & \dots & x_1^d \\ 1 & x_2 & x_2^2 & \dots & x_2^d \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_n & x_n^2 & \dots & x_n^d \end{pmatrix}.$$

```
fit = lm(wage ~ poly(age,4))
Plot(fit, main = "Polynomial Regression")
```

Polynomial Regression



Step Functions

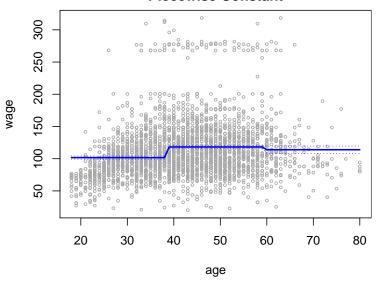
- Divide age into bins
- Model wage as a constant in each bin
- ightharpoonup The basis functions indicate which bin x_i belongs to
- ightharpoonup Cutpoints c_1, c_2, \ldots, c_K

$$b_j(x_i) = I(c_j \leq x_i < c_{j+1})$$

$$\mathbf{X} = \begin{pmatrix} 1 & I(x_1 < c_1) & I(c_1 \le x_1 < c_2) & \dots & I(c_K \le x_1) \\ 1 & I(x_2 < c_1) & I(c_1 \le x_2 < c_2) & \dots & I(c_K \le x_2) \\ \vdots & \vdots & & \vdots & \ddots & \vdots \\ 1 & I(x_n < c_1) & I(c_1 \le x_n < c_2) & \dots & I(c_K \le x_n) \end{pmatrix}.$$

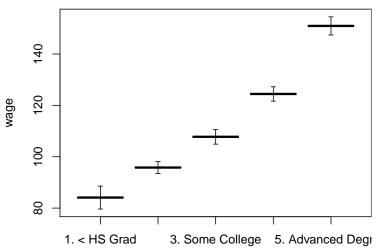
```
fit = lm(wage ~ cut(age,3))
Plot(fit, main = "Piecewise Constant")
```

Piecewise Constant



```
fit = lm(wage ~ education)
Plot(fit, main = "Piecewise Constant")
```

Piecewise Constant

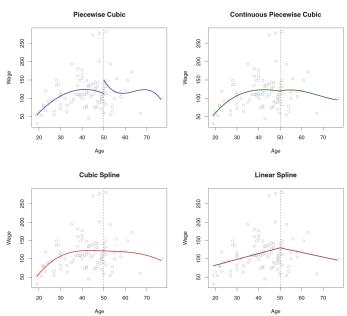


Regression Splines

A degree-d spline is a piecewise degree-d polynomial, with continuity in derivatives up to degree d-1 at each knot.

- Combination of polynomials and step functions
- \blacktriangleright Knots c_1, c_2, \ldots, c_K
- ▶ Continous derivatives up to order d-1 at each knot.

Regression Splines



Regression Splines

An order-d spline with K knots has the basis

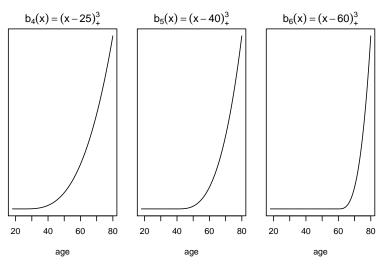
$$b_j(x_i) = x_i^j$$
 , $j = 1, ..., d$
 $b_{d+k}(x_i) = (x_i - c_k)_+^d$, $k = 1, ..., K$,

where

$$(x-c_j)_+^d = \begin{cases} (x-c_j)^d & , x > c_j \\ 0 & , \text{otherwise.} \end{cases}$$

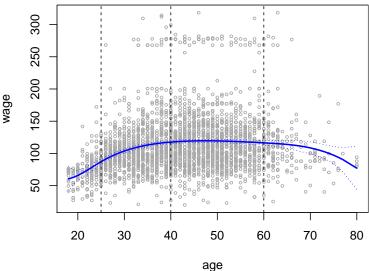
Cubic Splines

- ightharpoonup A spline with d=3 is cubic
- ► The basis is $X, X^2, X^3, (X c_1)^3_+, (X c_2)^3_+, \dots, (X c_K)^3_+$



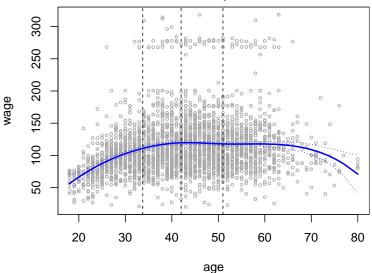
```
fit = lm(wage ~ bs(age, knots = c(25,40,60)))
Plot(fit, main = "Cubic spline")
```





```
fit = lm(wage ~ bs(age, df = 6))
Plot(fit, main = "Cubic spline")
```





Natural Cubic Splines

- Cubic spline that is linear at the ends
- ► The idea is to reduce variance
- ▶ Straight line outside $c_0 = 18$ and $c_{K+1} = 80$
- We call these points boundary knots

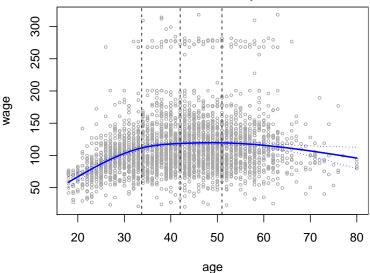
The basis is

$$b_1(x_i) = x_i, \quad b_{k+2}(x_i) = d_k(x_i) - d_K(x_i), \ k = 0, \dots, K - 1,$$

$$d_k(x_i) = \frac{(x_i - c_k)_+^3 - (x_i - c_{K+1})_+^3}{c_{K+1} - c_k}.$$

```
fit = lm(wage ~ ns(age, df = 4))
Plot(fit, main = "Natural Cubic Spline")
```

Natural Cubic Spline



Recap

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}.$$

- Non-linear methods, but linear regression.
- ▶ Each method defined by a basis, $\mathbf{X}_{ij} = b_j(x_i)$.
- ► And simply $\hat{\beta} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{y}$
- ▶ We will now move from $\mathbf{X}\beta$ to f(X)

$$\mathbf{y} = \mathbf{f}(X) + \varepsilon.$$

Smoothing Splines

- Different idea than regression splines
- Minimize the prediction error
- Bias-variance approach

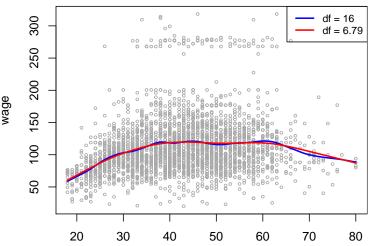
A smoothing spline is the function g that minimizes

$$\sum_{i=1}^n (y_i - g(x_i))^2 + \lambda \int g''(t)^2 dt.$$

- ▶ What happens as $\lambda \to \infty$?
- ▶ What happens as $\lambda \to 0$?

```
fit = smooth.spline(age, wage, df = 16)
Plot(fit, main = "Smoothing Splines")
fit = smooth.spline(age, wage, cv = T)
Plot(fit, legend = 16)
```

Smoothing Splines



The Smoother Matrix

The fitted values are

$$\hat{\mathbf{y}} = \mathbf{S}\mathbf{y}$$
.

The effective degrees of freedom is

$$df_{\lambda}=\mathrm{tr}(\mathbf{S}).$$

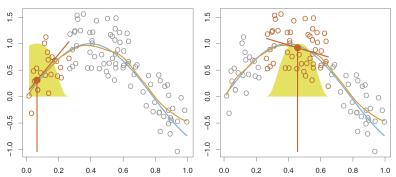
The leave-one-out cross-validation error is

$$RSS_{cv}(\lambda) = \sum_{i=1}^{n} \left(\frac{y_i - \hat{y}_i}{1 - \mathbf{S}_{ii}} \right)^2.$$

Local Regression

- ightharpoonup Smoothed k-nearest neighbor algorithm
- ightharpoonup Run for each x_0
- ▶ Draw a line $\beta_0 + \beta_1 x$ through neighborhood
- ► Close observations are weighted more heavily
- ► The fitted value is $\hat{\beta}_0 + \hat{\beta}_1 x_0$

Local Regression



Local Regression

Finding the $\hat{\beta}_0$ and $\hat{\beta}_1$ that minimize

$$\sum_{i=1}^n K_{i0}(y_i - \beta_0 - \beta_1 x_i)^2,$$

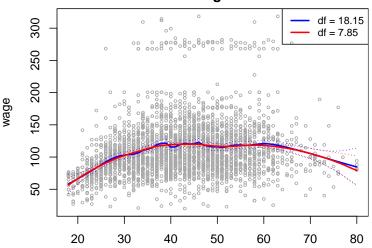
where

$$\mathcal{K}_{i0} = \left(1 - \left|\frac{x_0 - x_i}{x_0 - x_\kappa}\right|^3\right)_+^3.$$

Local Regression

```
fit = loess(wage ~ age, span = .2)
Plot(fit, main = "Local Regression")
Plot(loess(wage ~ age, span=.5),legend=fit$trace.hat)
```

Local Regression



Additive Models

Combines the models we have discussed so far. For example

$$y_i = f_1(x_{i1}) + f_2(x_{i2}) + \varepsilon_i$$

= $f(x_i) + \varepsilon_i$.

If each component is on the form $\mathbf{X}\beta$, so is f.

Component 1

- ightharpoonup Cubic spline with $X_1 = age$
- Knots at 40 and 60

The design matrix when excluding the intercept is

$$\mathbf{X}_{1} = \begin{pmatrix} x_{11} & x_{11}^{2} & x_{11}^{3} & (x_{11} - 40)_{+}^{3} & (x_{11} - 60)_{+}^{3} \\ x_{21} & x_{21}^{2} & x_{21}^{3} & (x_{21} - 40)_{+}^{3} & (x_{21} - 60)_{+}^{3} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ x_{n1} & x_{n1}^{2} & x_{n1}^{3} & (x_{n1} - 40)_{+}^{3} & (x_{n1} - 60)_{+}^{3} \end{pmatrix}.$$

Component 2

- Natural spline with X₂ = year
- ▶ Knot at $c_1 = 2006$
- ▶ Boundary knots at $c_0 = 2003$ and $c_2 = 2009$

The design matrix when excluding the intercept is

$$\mathbf{X}_{2} = \begin{pmatrix} x_{12} & \left[\frac{1}{6} (x_{12} - 2003)^{3} - \frac{1}{3} (x_{12} - 2006)^{3}_{+} \right] \\ x_{22} & \left[\frac{1}{6} (x_{22} - 2003)^{3} - \frac{1}{3} (x_{22} - 2006)^{3}_{+} \right] \\ \vdots & \vdots \\ x_{n2} & \left[\frac{1}{6} (x_{n2} - 2003)^{3} - \frac{1}{3} (x_{n2} - 2006)^{3}_{+} \right] \end{pmatrix}.$$

Component 3

- Factor $X_3 =$ education
- ► Levels < HS Grad, HS Grad (HSG) , Some College (SC) , College Grad (CG) and Advanced Degree (AD)
- Dummy variable coding

The design matrix when excluding the intercept is

$$\mathbf{X}_{3} = \begin{pmatrix} I(x_{13} = \mathrm{HSG}) & I(x_{13} = \mathrm{SC}) & I(x_{13} = \mathrm{CG}) & I(x_{13} = \mathrm{AD}) \\ I(x_{23} = \mathrm{HSG}) & I(x_{23} = \mathrm{SC}) & I(x_{23} = \mathrm{CG}) & I(x_{23} = \mathrm{AD}) \\ \vdots & \vdots & \vdots & \vdots \\ I(x_{n3} = \mathrm{HSG}) & I(x_{n3} = \mathrm{SC}) & I(x_{n3} = \mathrm{CG}) & I(x_{n3} = \mathrm{AD}) \end{pmatrix}.$$

Additive Model

Combine the components to

$$\mathbf{y}_i = f_1(x_{i1}) + f_2(x_{i2}) + f_3(x_{i3}) + \varepsilon_i.$$

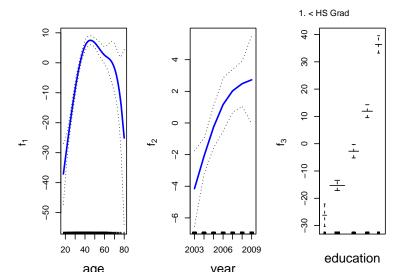
Since each component is linear, we can write

$$\mathbf{y}=\mathbf{X}\boldsymbol{\beta}+\boldsymbol{\varepsilon},$$

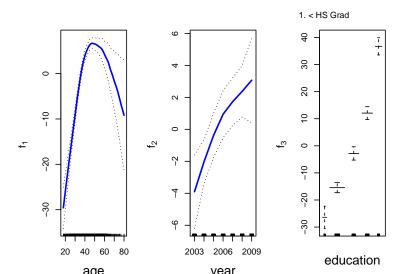
where

$$\mathbf{X} = \begin{pmatrix} \mathbf{1} & \mathbf{X}_1 & \mathbf{X}_2 & \mathbf{X}_3 \end{pmatrix}.$$

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Qualitative Responses

- ► Logistic regression
- Y = 0 or Y = 1
- $p(X) = \Pr(Y = 1|X)$

The generalized logistic regression model is

$$\log\left(\frac{p(X)}{1-p(X)}\right)=f(X).$$

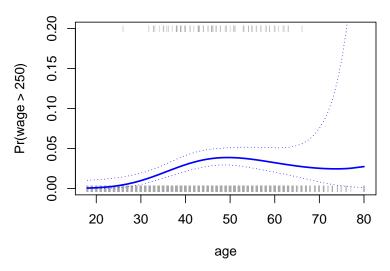
Choose f from the methods we have learned.

Polynomial Logistic Regression

With degree 4 we have

$$\log\left(\frac{p(X_1)}{1-p(X_1)}\right) = \beta_0 + \beta_1 X_1 + \beta_2 X_1^2 + \beta_3 X_1^3 + \beta_4 X_1^4.$$

Polynomial

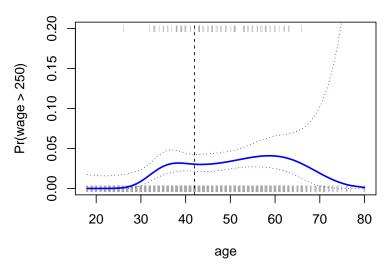


Cubic Spline Logistic Regression

- ► A cubic spline in age
- ► Knot at 42

$$\log\left(\frac{p(X_1)}{1-p(X_1)}\right) = \beta_0 + \beta_1 X_1 + \beta_2 X_1^2 + \beta_3 X_1^3 + \beta_4 (X_1 - 42)_+^3.$$

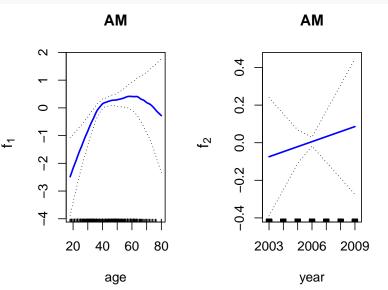
Cubic spline



GAM

- $ightharpoonup f_1$ is a local regression in age
- ▶ f₂ is a simple linear regression in year

$$\log\left(\frac{p(X_1,X_2)}{1-p(X_1,X_2)}\right) = \beta_0 + f_1(X_1) + f_2(X_2).$$



References

- Friedman, Jerome, Trevor Hastie, and Robert Tibshirani. 2001. *The Elements of Statistical Learning*. Vol. 1. Springer series in statistics New York.
- James, Gareth, Daniela Witten, Trevor Hastie, and RobertTibshirani. 2013. An Introduction to Statistical Learning. Vol. 112. Springer.
- Reinsch, Christian H. 1967. "Smoothing by Spline Functions." *Numerische Mathematik* 10 (3): 177–83.
- Rodriguez, German. 2001. "Smoothing and Non-Parametric Regression."
- Tibshirani, Ryan, and L Wasserman. 2013. "Nonparametric Regression." Statistical Machine Learning, Spring.