## This is a title and this is too

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#### 2 I. INTRODUCTION

Almost 100 years of research have repeatedly shown that consonantal voicing has an effect
on preceding vowel duration: vowels followed by voiced obstruents are longer than when
followed by voiceless ones (Belasco, 1953; Chen, 1970; Durvasula and Luo, 2012; Esposito,
2002; Farnetani and Kori, 1986; Fowler, 1992; Halle and Stevens, 1967; Heffner, 1937; House
and Fairbanks, 1953; Hussein, 1994; Javkin, 1976; Klatt, 1973; Kluender et al., 1988; Laeufer,
1992; Lampp and Reklis, 2004; Lisker, 1974; Maddieson and Gandour, 1976; Peterson and
Lehiste, 1960; Raphael, 1975; Warren and Jacks, 2005). Evidence for such so called 'voicing
effect' has been found in a variety of languages, including (but not limited to) English,
German, Hindi, Russian, Italian, Arabic, and Korean (see Maddieson and Gandour, 1976,
for a more comprehensive, but still not exhaustive list). Despite of the plethora of evidence
in support of the existence of the voicing effect, still after 100 years agreement hasn't been
reached regarding the source of this effect.

Several proposal have been put forward as to where to look for the possible cause of the voicing effect (see Maddieson and Gandour, 1976, and Sóskuthy (2013) for an overview).

Most of the proposed accounts place the source of the voicing effect in properties of speech production.<sup>1</sup> One of these production accounts, which will be the focus of this study, relates the voicing effect to some constant property of speech that is held constant across contexts while the local property of voiceless vs. voiced obstruents varies, thus creating a trade-off solution within the constant property. Lindblom (1967), Slis and Cohen (1969a), Slis and Cohen (1969b), and Lehiste (1970) (among others) argue that the relevant invariant property

of speech is a constant durational interval within which segments of different duration results in different duration of other segments. Both the syllable/VC sequence (Lindblom) and the word (Lehiste, Slis) has been proposed as the fixed interval. The closure of voiced stops is shorter than that of voiceless stops. It follows that vowels followed by shorter closures (like in the case of voiced stops) are longer than vowels followed by longer closures (like in the case of voiceless stops).

However, Chen (1970) and Maddieson and Gandour (1976) criticise the compensatory temporal adjustment account on empirical grounds. Chen (1970) shows that the duration of the syllable is affected by consonant voicing (compatible with findings in Jacewicz et al., 2009), contrary to Lindblom's expectations. Maddieson and Gandour (1976) reject any compensatory account based on data from a parallel of the voicing effect, the aspiration effect, by which vowel tend to be longer when followed by aspirated stops than when followed by non-aspirated stops. They find no compensatory pattern between vowel and consonant duration: the consonant /t/, which has the shortest duration, is preceded by the shortest vowel, and vowels before /d/ and /t / have the same duration although the durations of the two consonant are different.

#### A. The present study

An exploratory study of acoustic data from Italian and Polish was conducted to investigate the relationship between vowel duration and consonant voicing in two languages that reportedly differ in the magnitude of the voicing effect. Italian has been unanimously reported as a voicing effect language (Caldognetto et al., 1979; Esposito, 2002; Farnetani and Kori, 1986). The mean difference in vowel duration when followed by voiceless vs. voiced
consonants ranges between 22 and 24 ms (with longer vowels followed by voiced consonants,
Esposito, 2002; Farnetani and Kori, 1986).<sup>2</sup> On the other hand, Polish is subject to conflicting results regarding the presence and magnitude of the effect. While Keating (1984)
reports no effect of voicing on vowel duration in data from 24 speakers, Nowak (2006) finds
that vowels followed by voiced stops are 4.5 ms longer in the 4 speakers recorded. Moreover,
Malisz and Klessa (2008) argue based on data from 40 speakers that the magnitude of the
voicing effect in Polish is highly idiosyncratic, and claim their results to be inconclusive.

They show that word duration is not affected by voicing of C2 but they don't discuss the
internal structure of the word. I will show that the Release to Release is invariant and that
this is compatible with a gestural timing in which the C2 is right-edge aligned with C1/V. I
will also offer an interpretation of Maddieson and Gandour (1976) that is compatible with

## 57 II. METHOD

## A. Participants

a compensatory temporal adjustment account.

Seventeen subjects in total participated to this exploratory study. Eleven participants
were native speakers of Italian (5 female, 6 male), while six were native speakers of Polish (3
female, 3 male). The Italian speakers were from the North and Centre of Italy (8 speakers
from Northern Italy, 3 from Central Italy). The Polish group had 2 speakers from Poznań
and 4 speakers from Eastern Poland. For more information on the speakers, see Appendix A.

- 64 Ethical clearance was obtained for this study from the University of Manchester (REF 2016-
- 65 0099-76). The participants signed a written consent and received a monetary compensation.

## 66 B. Equipment

- The acquisition of the audio signal was achieved with the software Articulate Assistant
- <sup>68</sup> Advanced<sup>TM</sup> (AAA, v2.17.2) running on a Hawlett-Packard ProBook 6750b laptop with
- 69 Microsoft Windows 7, with a sample rate of 22050 MHz (16-bit) in a proprietary format. A
- 70 FocusRight Scarlett Solo pre-amplifier and a Movo LV4-O2 Lavalier microphone were used
- 71 for audio recording.

#### 72 C. Materials

The target stimuli were disyllabic words with  $C_1V_1C_2V_2$  structure, where  $C_1 = /p/$ ,  $V_1$  = /a, o, u/,  $C_2 = /t$ , d, k, g/, and  $V_2 = V_1$  (e.g. /pata/, /pada/, /poto/, etc.). The lexical stress of the target words was placed by speakers of both Italian and Polish on  $V_1$ , as intended. The make-up of the target words was constrained by the design of the experiment, which included ultrasound tongue imaging (UTI). Front vowels are difficult to image with UTI, since their articulation involves tongue positions which are particularly far from the ultrasonic probe, hence reducing the visibility of the tongue contour. For this reason, only central and back vowels were included. Since one of the variables of interest in the exploratory study was the closing gesture of  $C_2$ , only lingual consonants were used. A labial stop was chosen as the first consonant to reduce possible coarticulation with the following vowel (although see Vazquez-Alvarez and Hewlett 2007). The target words were

- embedded in a frame sentence, *Dico X lentamente* 'I say X slowly' in Italian (following Hajek and Stevens, 2008), and *Mówię X teraz* 'I say X now' in Polish. These sentences were chosen in order to keep the placement of stress and emphasis similar across languages, so to ensure
- 87 comparability of results.

## 88 D. Procedure

The participant was asked to read the sentences with the target words which were sequentially presented on the computer screen. The order of the sentence stimuli was randomised
for each participant. Each participant read the list of randomised sentence stimuli six times.

Due to software constraints, the order of the list was kept the same across the six repetitions
within each participant. Each speaker read a total of 72 sentences, with a grand total of
total of
total of tokens (288 per language). The reading task lasted between 15 and 20 minutes, with
optional short breaks between one repetition and the other.

#### 96 E. Data processing and measurements

The audio recordings were exported from AAA in .wav format for further processing. A forced aligned transcription was accomplished through the SPeech Phonetisation Alignment and Syllabification software (SPPAS) (Bigi, 2015). The outcome of the automatic annotation was manually corrected when necessary, according to the criteria in Table I. The releases of C1 and C2 were detected automatically by means of a Praat scripting implementation of the algorithm described in Ananthapadmanabha et al. (2014). The durations in milliseconds of the following intervals were extracted from the annotated acoustic landmarks with

TABLE I. List of measurements as extracted from acoustics.

landmark		criteria
vowel onset	(V1 onset)	appearance of higher formants in the spectrogram
		following the burst of $/p/$ (C1)
vowel offset	(V1 offset)	disappearance of the higher formants in the
		spectrogram preceding the target consonant (C2)
consonant onset	(C2 onset)	corresponds to V1 offset
closure onset	(C2 closure onset)	corresponds to V1 offset
consonant offset	(C2 offset)	appearance of higher formants of the vowel
		following C2 (V2); corresponds to V2 onset
consonant release	(C1/C2 release)	automatic detection + manual correction
		(Ananthapadmanabha et al., 2014)

Praat scripting: sentence duration, word duration, vowel duration (V1 onset to V1 offset),
consonant closure duration (V1 offset to C2 burst), and Release-to-Release duration (RR
duration, C1 release to C2 release). Syllable rate (syllables per second) was used as a proxy
to speech rate (Plug and Smith, 2018) for duration normalisation, and was calculated as
the number of syllables divided by the duration of the sentence (8 syllables in Italian, 6 in
Polish).

## F. Statistical analysis

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Given the exploratory nature of the study, all statistical analyses reported here are to be 111 considered data-driven or hypothesis-generating rather than hypothesis-driven (Kerr, 1998). The durational measurements were analysed with linear mixed-effects models using lme4 113 v1.1-17 in R v3.5.0 (Bates et al., 2015; Team, 2018). All factors were coded as treatment 114 contrasts. P-values for the individual terms were obtained with lmerTest v3.0-1, which uses the Satterthwaite's approximation to degrees of freedom (Kuznetsova et al., 2017). P-values 116 below the alpha level 0.05 were considered significant. The estimates of the relevant effects 117 are then calculated by refitting the models including only the significant terms (step-down 118 approach, Diggle et al., 2002; Zuur et al., 2009, pp. 121–122). Bayes factors were used 119 to specifically test the null hypotheses that word and RR duration are not affected by C2 120 voicing (i.e., the effect of C2 voicing on duration is 0). For each set of null/alternative 121 hypotheses, a full model (with the predictor of interest) and a null model (excluding it) 122 were fitted separately using Maximum Likelihood estimation (Bates et al., 2015, p. 34). The 123 BIC approximation was then used to obtain Bayes factors (Jarosz and Wiley, 2014; Raftery, 124 1995, 1999; Wagenmakers, 2007). The approximation is calculated according to the equation 125 in 1 (Wagenmakers, 2007, p. 796).

$$BF_{01} \approx exp(\Delta BIC_{10}/2)$$
 (1)

where  $\Delta BIC_{10} = BIC_1 - BIC_0$ ,  $BIC_1$  is the BIC of the full model, and  $BIC_0$  is the BIC of the null model. Values of  $BF_{01} > 1$  indicate a preference of  $H_0$  over  $H_1$ . The interpretation of the Bayes factors follows the recommendations in Raftery (1995, p. 139).

## 130 III. RESULTS

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Only the most relevant terms will be presented. For the others see tables and appendixes.

## A. Vowel duration

A linear mixed-effects model was fitted with the following terms: vowel duration as the 133 outcome variable; fixed effects for C2 voicing (voiceless, voiced), vowel (a, o, u), language 134 (Italian, Polish), and speech rate (as syllables per second); by-speaker and by-word random intercept with by-speaker random slopes for C2 voicing. All possible interactions between 136 C2 voicing, vowel, and language were included. The following terms are significant: C2 137 voicing, vowel, language, and speech rate. No interaction was significant. The vowel /a/ (when followed by voiceless stops) has a duration of 202.5 ms (se = 8.5 ms). The vowel 139 /o/ is 9.5 ms shorter (se = 3 ms), while the vowel /u/ is 30.5 ms shorter (se = 3). Vowels 140 are 11 ms longer (se = 3) when followed by a voiced stop. Polish has on average shorter vowels than Italian ( $\hat{\beta} = -28 \text{ ms}$ , se = 8 ms), although the effect of voicing is estimated to be 142 the same in both languages (the interaction of language and C2 voicing is not significant). 143 Speech rate has a negative effect on vowel duration, such that faster rates correlate with shorter vowel durations ( $\hat{\beta} = -15 \text{ ms}$ , se = 1 ms).

#### B. Consonant closure duration

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The same maximally specified model as with vowel duration has been fitted to consonant closure durations as the outcome variable. Only C2 voicing and speech rate were significant. Stop closure is 13 ms shorter (se = 3 ms) if the stop is voiced. Finally, faster speech rates correlate with shorter closure durations ( $\hat{\beta}$  = -8 ms, se = 1 ms).

#### C. Vowel and closure duration

The full model was specified with the following terms and interactions: vowel duration 152 as the outcome variable; as fixed effects, closure duration, vowel, speech rate; an interaction 153 between closure duration and vowel; by-speaker and by-word random intercepts, and by-154 speaker random slopes for C2 voicing. Closure duration has a significant effect on vowel 155 duration when the vowel is  $/a/(\hat{\beta} = -0.2 \text{ ms}, \text{ se} = 0.06 \text{ ms})$ . The effect with the vowel /o/156 does not significantly differ from the one with /a/, while with /u/ the effect is -0.5 ms ( $\hat{\beta}$  = 157 -0.3 ms, se = 0.06 ms). In general, then, closure duration is inversely correlated with vowel 158 duration after controlling for speech rate. However the correlation is quite weak. A 1 ms 159 increase in closure duration corresponds to a 0.2–0.5 ms decrease in vowel duration. 160

## D. Word duration

The following full and null models were fitted to test whether word duration is affected by
C2 voicing. In the full model, I entered as fixed effects: C2 voicing, C2 place, vowel, speech
rate, and language. The model also included by-speaker and by-word random intercepts,

plus a by-speaker random slope for C2 voicing. The null model was the same as the full model, with the exclusion of the fixed effect of C2 voicing. The Bayes factor of the null model against the full model is 21.5. Thus, the null model (in which the effect of C2 voicing is 0) is 21.5 times more likely under the observed data than the full model. This indicates that there is strong evidence for word duration not being affected by C2 voicing.

### E. Release to Release interval (RR) duration

The models specifications for the RR duration were the same as for word duration. The
Bayes factor of the null model against the full model for RR duration is 19, which means
that the null model (without C2 voicing) is 19 times more likely than the full model. The
data suggests there is positive evidence that duration of the RR interval is not affected by
C2 voicing.

#### 176 IV. DISCUSSION

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A major drawback of the analysis in Maddieson and Gandour (1976) is that the consonant to the release nant duration in fact was measured from the closure of the relevant consonant to the release of the following consonant, due to difficulties in detecting the release of the consonant of interest (e.g., in  $ab \ s\bar{a}th \ kaho$ , the duration of /t / in  $s\bar{a}th$  was calculated as the interval between the closure of /t / and the release of /k/). This measure includes the burst and (eventual) aspiration of the consonant. Slis and Cohen (1969a), however, states that the inverse correlation between vowel duration and the following consonant raises when consonant closure duration is taken into account, and not entire consonant duration. If the correlation

exists between vowel and closure duration, the inclusion of burst/aspiration duration clearly alters this relationship.

## 7 ACKNOWLEDGMENTS

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#### 189 APPENDIX A: SOCIO-LINGUISTIC INFORMATION OF PARTICIPANTS

- <sup>1</sup>Two accounts that point to perceptual features are Javkin (1976) and Kluender *et al.* (1988). To the

  best of my knowledge, Javkin (1976)'s proposal remains empirically untested, while see Fowler (1992) for

  arguments against Kluender *et al.* (1988).

  <sup>2</sup>These estimates should be taken as a gross approximation. There are several issues: number of speakers,
- different contexts, statistical modelling.

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Participant ID	Age	Sex	Native language	Other languages	City of birth	$S_{J}$
it01	29	Male	Italian	English, Spanish	Verbania	V
it02	26	Male	Italian	Friulian, English, Ladin-Venetan	Udine	T
it03	28	Female	Italian	English, German	Verbania	V
it04	54	Female	Italian	Calabrese	Verbania	V
it05	28	Female	Italian	Engligh	Verbania	V
it07	29	Male	Italian	English	Tradate	С
it09	35	Female	Italian	English	Vignola (MO), Italy	V
it11	24	Male	Italian	english	Monza	Μ
it13	20	Female	Italian	English, French, Arabic, Farsi	Ancona	С
it14	32	Male	Italian	English, Spanish	Frosinone	F
pl02	32	Female	Polish	English, Norwegian, French, German, Dutch	Koło	Ρ
pl03	26	Male	Polish	Russian, English, French, German	Nowa Sol	Ρ
pl04	34	Female	Polish	Spanish, English, French	Warsaw	W
pl05	42	Male	Polish	English, French	Przasnysz	W
pl06	33	Male	Polish	English	Zgierz	$\mathbf{Z}_{i}$
pl07	32	Female	Polish	English, Russian	Bielsk Podlaski	В

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