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## I. INTRODUCTION

Almost 100 years of research have repeatedly shown that consonantal voicing has an effect on preceding vowel duration: vowels followed by voiced obstruents are longer than when followed by voiceless ones (Belasco, 1953; Chen, 1970; Durvasula and Luo, 2012; Esposito, 2002; Farnetani and Kori, 1986; Fowler, 1992; Halle and Stevens, 1967; Heffner, 1937; House and Fairbanks, 1953; Hussein, 1994; Javkin, 1976; Klatt, 1973; Kluender *et al.*, 1988; Laeufer, 1992; Lampp and Reklis, 2004; Lisker, 1974; Maddieson and Gandour, 1976; Peterson and Lehiste, 1960; Raphael, 1975; Warren and Jacks, 2005). Evidence for such so called ‘voicing effect’ has been found in a variety of languages, including (but not limited to) English, German, Hindi, Russian, Italian, Arabic, and Korean (see Maddieson and Gandour, 1976, for a more comprehensive, but still not exhaustive list). Despite of the plethora of evidence in support of the *existence* of the voicing effect, still after 100 years agreement hasn’t been reached regarding the source of this effect.

Several proposal have been put forward as to where to look for the possible cause of the voicing effect (see Sóskuthy, 2013, and Beguš (2017) for an overview). Most of the proposed accounts place the source of the voicing effect in properties of speech production.<sup>1</sup> One of these production accounts, which will be the focus of this study, relates the voicing effect to some constant property of speech that is held constant across contexts while the local property of voiceless vs. voiced obstruents varies, thus creating a trade-off solution within the constant property. Lindblom (1967), Slis and Cohen (1969b), and Lehiste (1970b) (among others) argue that the relevant invariant property of speech is a constant durational interval within which segments of different duration results in different duration of other segments. Both the syllable/VC sequence (Lindblom, 1967) and the word (Lehiste, 1970a·b; Slis

and Cohen, 1969a·b) has been proposed as the fixed interval. The closure of voiced stops is shorter than that of voiceless stops. It follows that vowels followed by shorter closures (like in the case of voiced stops) are longer than vowels followed by longer closures (like in the case of voiceless stops). However, the compensatory temporal adjustment account has been criticised in several occasions.

The proposal of the syllable or the word as the targets for compensation and encounter difficulties when confronted with empirical evidence and when scrutinised by logic. First, Lindblom's (1967) argument that the syllable is the interval within which compensation happens is not supported by the findings in Chen (1970) and Jacewicz *et al.* (2009). Chen (1970) rejects a syllable-based compensatory account on the light of the fact that the duration of the syllable is affected by consonant voicing. More recently, Jacewicz *et al.* (2009) further shows that the duration of monosyllabic words in American English does not change dependent on the voicing of the coda consonant. Second, although the results in Slis and Cohen (1969b) confirm that the word does not change in duration whether the stop following the stressed vowel is voiceless or voiced, it does not follow from this fact that compensation should necessarily target the stressed vowel. Indeed, it is possible that, for example, the following unstressed vowel could be the target of the compensation.

Maddieson and Gandour (1976) reject any compensatory account for the voicing effect based on data from Hindi on the so called 'aspiration effect', by which vowels are longer when followed by aspirated stops than when followed by non-aspirated stops. Vowels before voiceless unaspirated stops are the shortest, followed by vowels before voiced unaspirated and voiceless aspirated stops, which have similar duration between each other, followed by vowels before voiced aspirated stops, which are the longest. Maddieson and Gandour (1976) find no compensatory pattern between vowel and consonant duration: the consonant /t/, which has the shortest duration, is preceded by the short-

est vowel, and vowels before /d/ and /t<sup>h</sup>/ have the same duration although the durations of the two consonant are different.

However, an reevaluation of the way consonant duration is measured in [Maddieson and Gandour \(1976\)](#) might actually turn the situation in favour of a compensatory account. Consonant duration is in fact measured from the closure of the relevant consonant to the release of the following consonant, due to difficulties in detecting the release of the consonant of interest (e.g., in *ab sāth kaho*, the duration of /t<sup>h</sup>/ in *sāth* was calculated as the interval between the closure of /t<sup>h</sup>/ and the release of /k/). This measure includes the burst and (eventual) aspiration of the consonant. [Slis and Cohen \(1969a\)](#), however, states that the inverse correlation between vowel duration and the following consonant raises when consonant *closure* duration is taken into account, and not the entire *consonant* duration. If the correlation exists between vowel and closure duration, the inclusion of burst/aspiration duration clearly alters this relationship. Indeed, the data in [Durvasula and Luo \(2012\)](#) show that closure duration, appropriately measured, decreases from voiceless unaspirated > voiced > voiceless aspirated > voiced aspirated, which closely resembles the order of increasing vowel duration in [Maddieson and Gandour \(1976\)](#).<sup>2</sup>

More recently, [Beguš \(2017\)](#) investigated the effect of three phonation types on vowel durations in Georgian and finds that vowels are short when followed by voiceless aspirated stops, longer before ejective stops, and longest when followed by voiced stops. Crucially, stop closure duration follows a symmetrical pattern: Closure duration is short in voiced stops, longer in ejectives, and longest in voiceless aspirated stops. [Beguš \(2017\)](#) argues that these findings support a temporal compensation account, although not univocally.

To summarise, a compensatory temporal adjustment account of the voicing effect remains possible after a careful review of the critiques advanced by [Chen \(1970\)](#) and [Maddieson and Gandour \(1976\)](#), and in face of the results in [Beguš \(2017\)](#), although issues about the actual implementation of the compensation still persist. In conclusion, for the compensatory account to gain plausibility, an invariant interval within which compensation is implemented needs to be better defined, on the light of empirical data.

## A. The present study

This paper reports on results from a broader exploratory study that investigates the relationship between vowel duration and consonant voicing from an articulatory perspective. Synchronised recordings of audio, ultrasound tongue imaging, and electroglottography were carried out to enable a data-driven approach to the analysis of features related to the voicing effect in the context of disyllabic (CVCV) words in Italian and Polish. The design of the study has been constrained by the use of these articulatory techniques (see Section II). Moreover, given the exploratory nature of the study, the experimental design was not implemented to directly test the compensatory account. Here, only the results from acoustic will be discussed.

Italian and Polish reportedly differ in the magnitude of the voicing effect. Italian has been unanimously reported as a voicing effect language ([Caldognetto \*et al.\*, 1979](#); [Esposito, 2002](#); [Farnetani and Kori, 1986](#)). The mean difference in vowel duration when followed by voiceless vs. voiced consonants ranges between 22 and 24 ms (with longer vowels followed by voiced consonants, [Esposito, 2002](#); [Farnetani and Kori, 1986](#)).<sup>3</sup> On the other hand, the results regarding the presence and magnitude of the effect in Polish are mixed. While [Keating \(1984\)](#) reports no effect of voicing on

vowel duration in data from 24 speakers, [Nowak \(2006\)](#) finds that vowels followed by voiced stops are 4.5 ms longer in the 4 speakers recorded. Moreover, [Malisz and Klessa \(2008\)](#) argue based on data from 40 speakers that the magnitude of the voicing effect in Polish is highly idiosyncratic, and claim their results to be inconclusive on this matter. The difference in presence or magnitude of the voicing effect in Italian vs. Polish should enable us to find an underlying property that differs in the two languages and that might indicate a possible source for the voicing effect.

The acoustic data from the exploratory study reported here reveal that the duration of the interval between the releases of the two consonants in CVCV words (the Release to Release interval) is not affected by the voicing of the second consonant. This finding is compatible with a compensatory temporal adjustment account by which the timing of the stop closure onset within said interval determines the respective durations of the vowel and the stop closure. I further propose that the invariant duration of the Release to Release interval is congruent with current views on gestural timing ([Goldstein and Pouplier, 2014](#)) and I discuss the insights it provides in relation to our understanding of gestural organisation in speech.

## II. METHOD

### A. Participants

Seventeen subjects in total participated to this exploratory study. Eleven participants were native speakers of Italian (5 female, 6 male), while six were native speakers of Polish (3 female, 3 male). The Italian speakers were from the North and Centre of Italy (8 speakers from Northern Italy, 3 from Central Italy). The Polish group had 2 speakers from Poznań and 4 speakers from Eastern Poland.

For more information on the speakers, see Appendix B. Ethical clearance was obtained for this study from the University of Manchester (REF 2016-0099-76). The participants signed a written consent and received a monetary compensation of £10.

## B. Equipment

The acquisition of the audio signal was achieved with the software Articulate Assistant Advanced™ (AAA, v2.17.2) running on a Hewlett-Packard ProBook 6750b laptop with Microsoft Windows 7, with a sample rate of 22050 MHz (16-bit) in a proprietary format. A FocusRight Scarlett Solo pre-amplifier and a Movo LV4-O2 Lavalier microphone were used for audio recording.

## C. Materials

The target stimuli were disyllabic words with  $C_1V_1C_2V_2$  structure, where  $C_1 = /p/$ ,  $V_1 = /a, o, u/$ ,  $C_2 = /t, d, k, g/$ , and  $V_2 = V_1$  (e.g. /pata/, /pada/, /poto/, etc.). Most are nonce words, although inevitably some combinations lead to real words both in Italian (4 words) and Polish (2 words, see Appendix C). The lexical stress of the target words was placed by speakers of both Italian and Polish on  $V_1$ , as intended. The make-up of the target words was constrained by the design of the experiment, which included ultrasound tongue imaging (UTI). Front vowels are difficult to image with UTI, since their articulation involves tongue positions which are particularly far from the ultrasonic probe, hence reducing the visibility of the tongue contour. For this reason, only central and back vowels were included. Since one of the variables of interest in the exploratory study was the closing gesture of  $C_2$ , only lingual consonants were used. A labial stop was chosen as the first consonant to reduce possible coarticulation with the following vowel (although see Vazquez-Alvarez and Hewlett 2007). The target

words were embedded in a frame sentence, *Dico X lentamente* ‘I say X slowly’ in Italian (following Hajek and Stevens, 2008), and *Mówię X teraz* ‘I say X now’ in Polish, and presented according to the respective writing conventions. These sentences were chosen in order to keep the placement of stress and emphasis similar across languages, so to ensure comparability of results.

#### D. Procedure

The participant was asked to read the sentences with the target words which were sequentially presented on the computer screen. The order of the sentence stimuli was randomised for each participant. Each participant read the list of randomised sentence stimuli 6 times. Due to software constraints, the order of the list was kept the same across the six repetitions within each participant. Each speaker read a total of 12 sentences for 6 times (with the exceptions of IT02, who repeated the 12 sentences 5 times, and IT07, with whom words containing /u/ were not recorded due to technical difficulties relating to the ultrasound data collection).<sup>4</sup> with a grand total of 1224 tokens (792 from Italian, 432 from Polish). The reading task lasted between 15 and 20 minutes, with optional short breaks between one repetition and the other.

#### E. Data processing and measurements

The audio recordings were exported from AAA in .wav format for further processing. A forced aligned transcription was accomplished through the SPEech Phonetisation Alignment and Syllabification software (SPPAS) (Bigi, 2015). The outcome of the automatic annotation was manually corrected when necessary, according to the criteria in Table I. The releases of C1 and C2 were detected automatically by means of a Praat scripting implementation of the algorithm described in



TABLE I. List of measurements as extracted from acoustics.

landmark		criteria
vowel onset	(V1 onset)	appearance of higher formants in the spectrogram following the burst of /p/ (C1)
vowel offset	(V1 offset)	disappearance of the higher formants in the spectrogram preceding the target consonant (C2)
consonant onset	(C2 onset)	corresponds to V1 offset
closure onset	(C2 closure onset)	corresponds to V1 offset
consonant offset	(C2 offset)	appearance of higher formants of the vowel following C2 (V2); corresponds to V2 onset
consonant release	(C1/C2 release)	automatic detection + manual correction ( <a href="#">Ananthapadmanabha et al., 2014</a> )

147 [Ananthapadmanabha et al. \(2014\)](#). The durations in milliseconds of the following intervals were  
148 extracted from the annotated acoustic landmarks with Praat scripting: sentence duration, word dura-  
149 tion, vowel duration (V1 onset to V1 offset), consonant closure duration (V1 offset to C2 burst), and  
150 Release-to-Release duration (RR duration, C1 release to C2 release). Figure 1 shows an example of  
151 the segmentation of /pata/ (a) and /pada/ (b) from an Italian speaker. Syllable rate (syllables per sec-  
152 ond) was used as a proxy to speech rate ([Plug and Smith, 2018](#)) for duration normalisation, and was  
153 calculated as the number of syllables divided by the duration of the sentence (8 syllables in Italian, 6

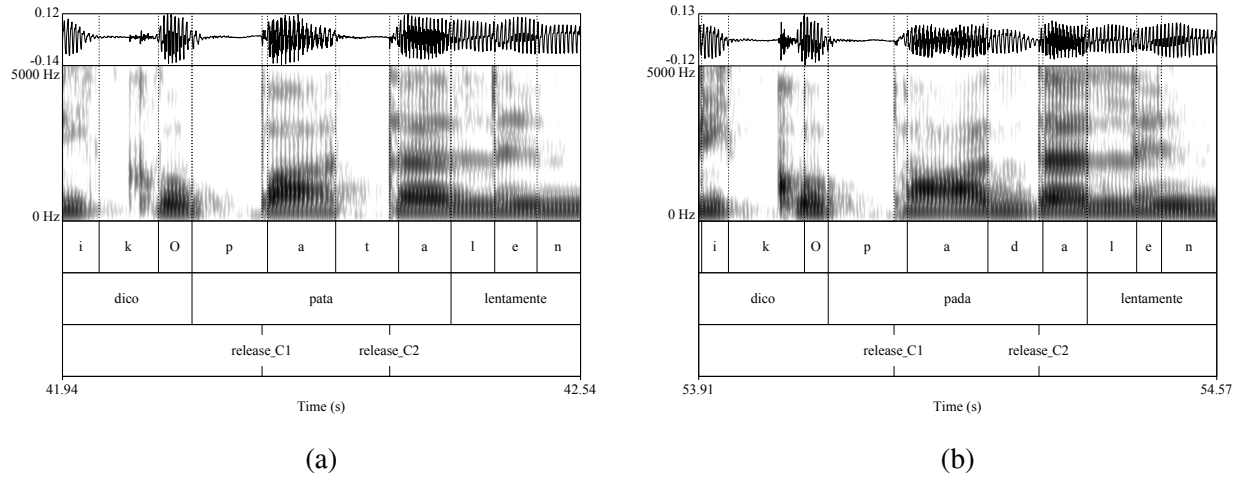


FIG. 1. Segmentation example.

154 in Polish). All further data processing and visualisation was done in R v3.5.0 (R Core Team, 2018;  
 155 Wickham, 2017).

## 156 F. Statistical analysis

157 Given the exploratory nature of the study, all statistical analyses reported here are to be considered  
 158 data-driven/hypothesis-generating rather than confirmatory/hypothesis-driven (Gelman and Loken,  
 159 2013; Kerr, 1998; Roettger, 2018). The durational measurements were analysed with linear mixed-  
 160 effects models using lme4 v1.1-17 in R (Bates *et al.*, 2015), and model estimates were extracted  
 161 with the effects package v4.0-2 (Fox, 2003). All factors were coded with treatment contrasts and  
 162 the following reference levels: voiceless (vs. voiced), /a/ (vs. /o/, /u/), coronal (vs. velar), Italian  
 163 (vs. Polish). The estimates in the results section refer to these reference levels unless interactions are  
 164 discussed. *P*-values for the individual terms were obtained with lmerTest v3.0-1, which uses the

Satterthwaite’s approximation to degrees of freedom (Kuznetsova *et al.*, 2017; Luke, 2017). A result is considered significant if the  $p$ -value is below the alpha level ( $\alpha = 0.05$ ).<sup>5</sup>

Bayes factors were used to specifically test the null hypotheses that word and RR duration are not affected by C2 voicing (i.e., the effect of C2 voicing on duration is 0). For each set of null/alternative hypotheses, a full model (with the predictor of interest) and a null model (excluding it) were fitted separately using Maximum Likelihood estimation (Bates *et al.*, 2015, p. 34). The BIC approximation was then used to obtain Bayes factors (Jarosz and Wiley, 2014; Raftery, 1995· 1999; Wagenmakers, 2007). The approximation is calculated according to the equation in 1 (Wagenmakers, 2007, p. 796).

$$BF_{01} \approx \exp(\Delta BIC_{10}/2) \quad (1)$$

where  $\Delta BIC_{10} = BIC_1 - BIC_0$ ,  $BIC_1$  is the BIC of the full model, and  $BIC_0$  is the BIC of the null model. Values of  $BF_{01} > 1$  indicate a preference of  $H_0$  over  $H_1$ . The interpretation of the Bayes factors follows the recommendations in Raftery (1995, p. 139).

The extracted measurements were filtered before statistical analysis. Measures of vowel duration, closure duration, word duration, and RR duration that are 3 standard deviations lower or higher than the respective means were excluded from the final dataset (which generally corresponds to a data loss of around 2.5%). This operation yields a total of 920 tokens of vowel and closure durations, 1176 tokens of word duration, and 848 tokens of RR duration.

### III. RESULTS

The following sections report the results of the study in relation to the durations of vowels, consonant closure, word, and the Release to Release interval. When discussing the output of statistical

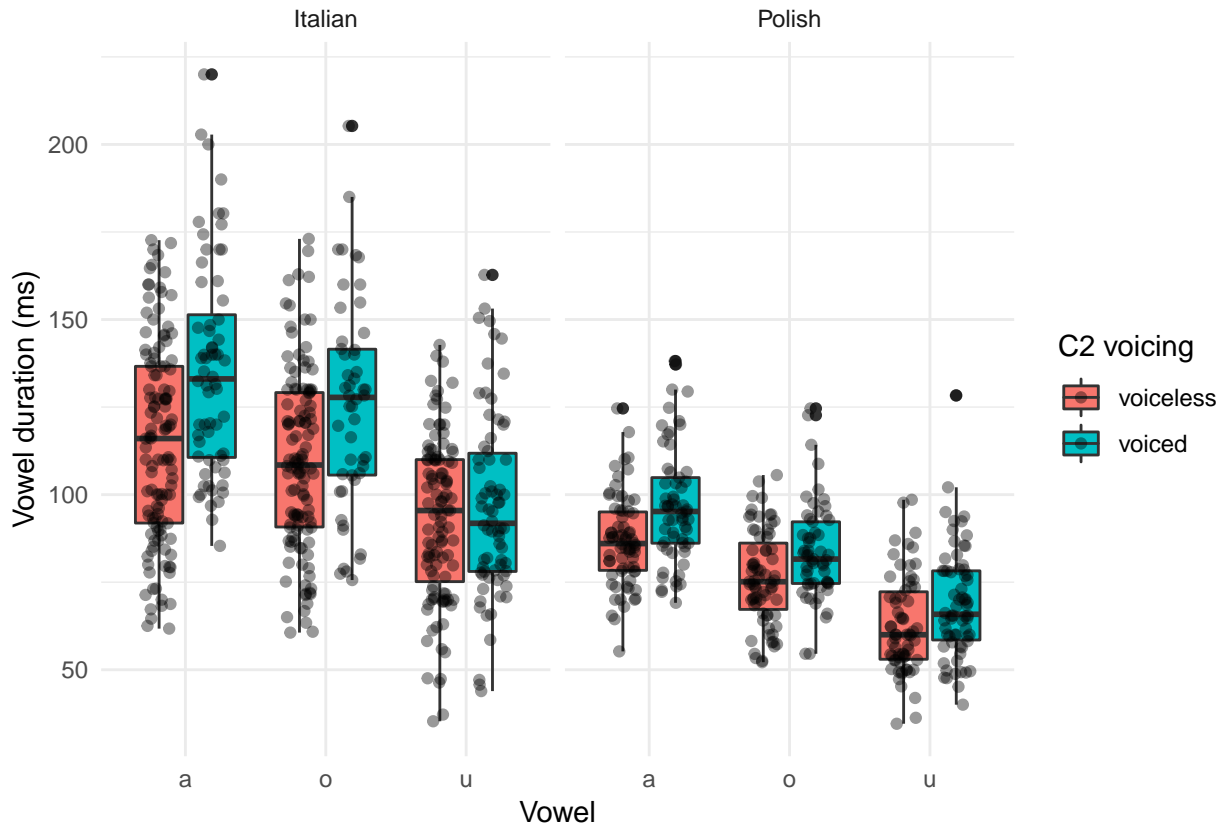


FIG. 2. Vowel duration in Italian and Polish.

modelling, only the relevant predictors and interactions will be presented. To avoid the cluttering generated by model parameters and alleviate the reader, the full output of statistical models and respective  $p$ -values are included in Appendix A.

### A. Vowel duration

Figure 2 shows boxplots and the raw data of vowel duration in Italian (on the left) and Polish (on the right) for the three vowels /a, o, u/. Vowel tend to be longer when followed by a voiced stop both in Italian and Polish. The effect appears to be greater in Italian than in Polish, especially for the vowels /a/ and /o/. There is no clear effect of C2 voicing in /u/ in Italian, but the effect is discernible

in Polish /u/. In Italian, vowels have a mean duration of 106 ms (sd = 27) before voiceless stops, and a mean duration of 118 ms (sd = 33) before voiced stops. Polish vowels are on average 75 ms long (sd = 16) when followed by a voiceless stop, and 83 ms long (sd = 19) if a voiced stop follows. The difference in vowel duration based on the raw means is 12 ms in Italian and 8 ms in Polish.

A linear mixed-effects model with vowel duration as the outcome variable was fitted with the following predictors: fixed effects for C2 voicing (voiceless, voiced), C2 place of articulation (coronal, velar), vowel (a, o, u), language (Italian, Polish), and speech rate (as syllables per second); by-speaker and by-word random intercept with by-speaker random slopes for C2 voicing. All possible interactions between C2 voicing, vowel, and language were included. The following terms are significant according to *t*-tests with Satterthwaite's approximation to degrees of freedom: C2 voicing, vowel, language, and speech rate. Only the interaction between C2 voicing and vowel is significant. Vowels are 19 ms longer (se = 4.4) when followed by a voiced stop (C2 voicing). The effect of C2 voicing is smaller with /u/ (around 5 ms,  $\hat{\beta} = -14.4$  ms, se = 6). Polish has on average shorter vowels than Italian ( $\hat{\beta} = -28$  ms, se = 8), and the effect of voicing is estimated to be about 11 ms (although recall that the interaction between language and C2 voicing is deemed not significant). Speech rate has unsurprisingly a negative effect on vowel duration, such that faster rates correlate with shorter vowel durations ( $\hat{\beta} = -15$  ms, se = 1).

## B. Consonant closure duration

Figure 3 illustrates stop closure durations with boxplots and individual raw data points. A pattern opposite to that with vowel duration can be noticed: Closure duration is shorter for voiced than for voiceless stops. The closure of voiceless stops in Italian is 77 ms long (sd = 20), while the voiced

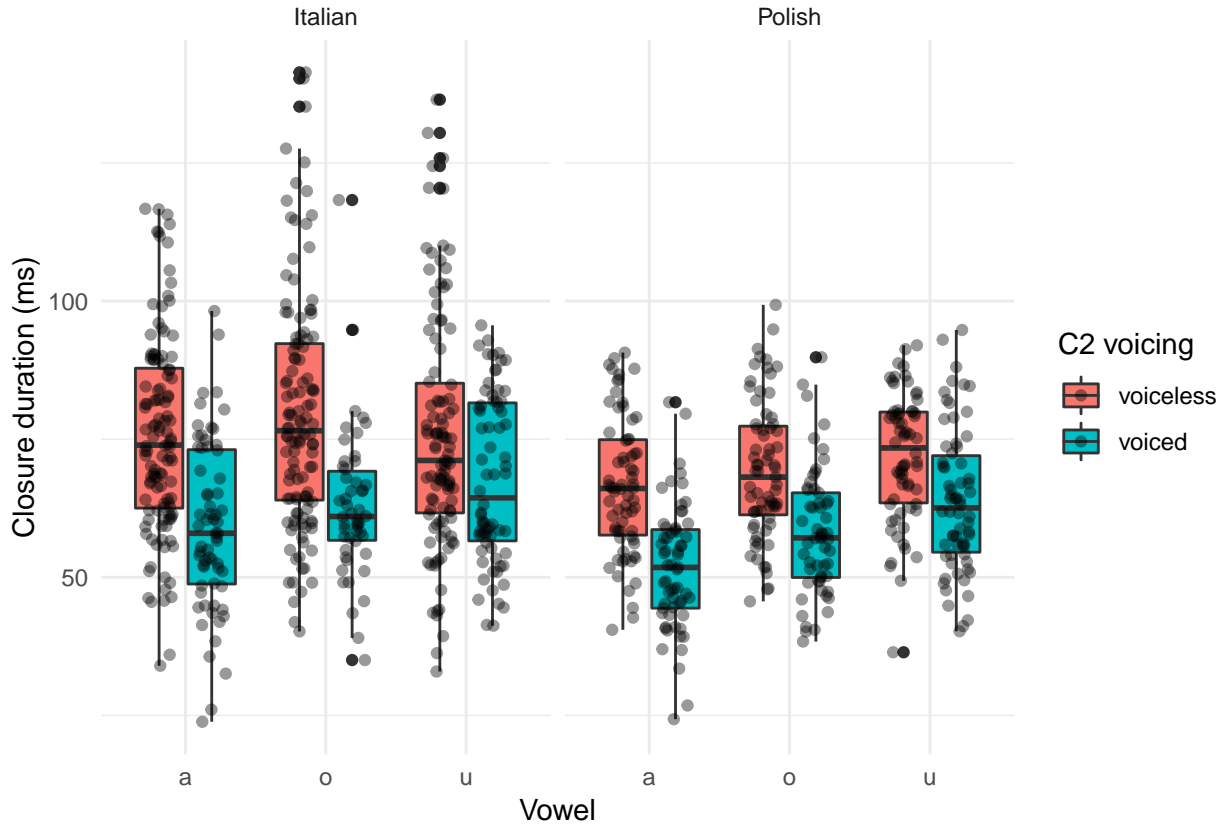


FIG. 3. Stop closure duration in Italian and Polish.

stops have a mean closure duration of 63 ms (sd = 15). In Polish, the closure duration is 69 ms (sd =  
 12) in voiceless stops and 58 ms (sd = 13) in voiced stops. The difference in closure duration based  
 on the raw means is 14 ms in Italian and 11 ms in Polish. The same model specification as with vowel  
 duration has been fitted with consonant closure durations as the outcome variable. C2 voicing, C2  
 place, and speech rate are significant. Stop closure is 16.5 ms shorter (se = 3) if the stop is voiced and  
 3.5 ms longer (se = 1.5) if velar. Finally, faster speech rates correlate with shorter closure durations  
 ( $\hat{\beta} = -8.5$  ms, se = 1 ms).

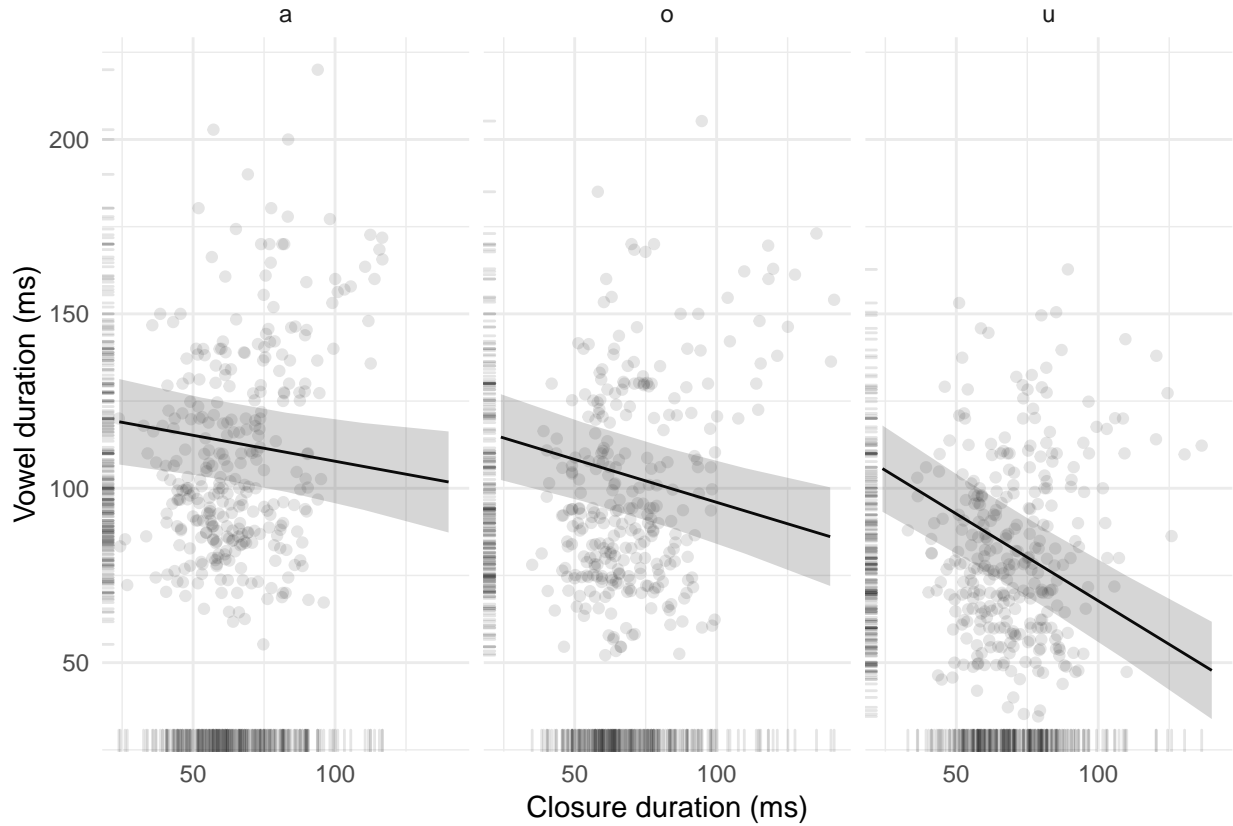


FIG. 4. Linear regression of closure and vowel duration per vowel.

### C. Vowel and closure duration

A model addressing the relationship between vowel and stop closure duration was fitted with the following terms and interactions: vowel duration as the outcome variable; as fixed effects, closure duration, vowel, speech rate; an interaction between closure duration and vowel; by-speaker and by-word random intercepts, and by-speaker random slopes for C2 voicing. Closure duration has a significant effect on vowel duration ( $\hat{\beta} = -0.15$  ms,  $se = 0.06$  ms). The effect with /u/ is greater than with /a/ and /o/ ( $\hat{\beta} = -0.35$  ms,  $se = 0.06$  ms). In general, closure duration is inversely correlated with vowel duration. However such correlation is quite weak. A 1 ms increase in closure duration corresponds to a 0.2–0.5 ms decrease in vowel duration. Figure 4 shows for each of /a, o, u/ the

individual data points and the regression lines with confidence intervals extracted from the linear model.

#### **D. Word duration**

Words with a voiceless stop are on average 397 ms long ( $sd = 81$ ) in Italian and 356 ms long ( $sd = 39$ ) in Polish. Words with a voiced stop have a mean duration of 396 ms ( $sd = 72$ ) in Italian and 362 ms ( $sd = 39$ ) in Polish. The following full and null models were fitted to test for the effect of C2 voicing on word duration. The full model has the following fixed effects: C2 voicing, C2 place, vowel, speech rate, and language. The model also includes by-speaker and by-word random intercepts, and a by-speaker random slope for C2 voicing. The null model excludes the fixed effect of C2 voicing. The Bayes factor of the null model against the full model is 24. Thus, the null model (in which the effect of C2 voicing is 0) is 24 times more likely under the observed data than the full model. This indicates that there is strong evidence for word duration not being affected by C2 voicing.

#### **E. Release to Release interval duration**

In Figure 5, boxplots show the durations of the Release to Release interval in words with a voiceless vs. a voiced C2 stop, in Italian (left side) and Polish (right side). It can be seen, also from the single data points, that the distributions and main statistics of the durations in the two conditions do not differ much within both languages. In Italian, the mean duration of the Release to Release interval is 210 ms ( $sd = 44$ ) if C2 is voiceless, and 209 ms ( $sd = 41$ ) if C2 is voiced. In Polish, the means are respectively 173 ( $sd = 22$ ) and 172 ( $sd = 21$ ) ms. The models specifications for the Release to Release duration are the same as for word duration. The Bayes factor of the null model against the



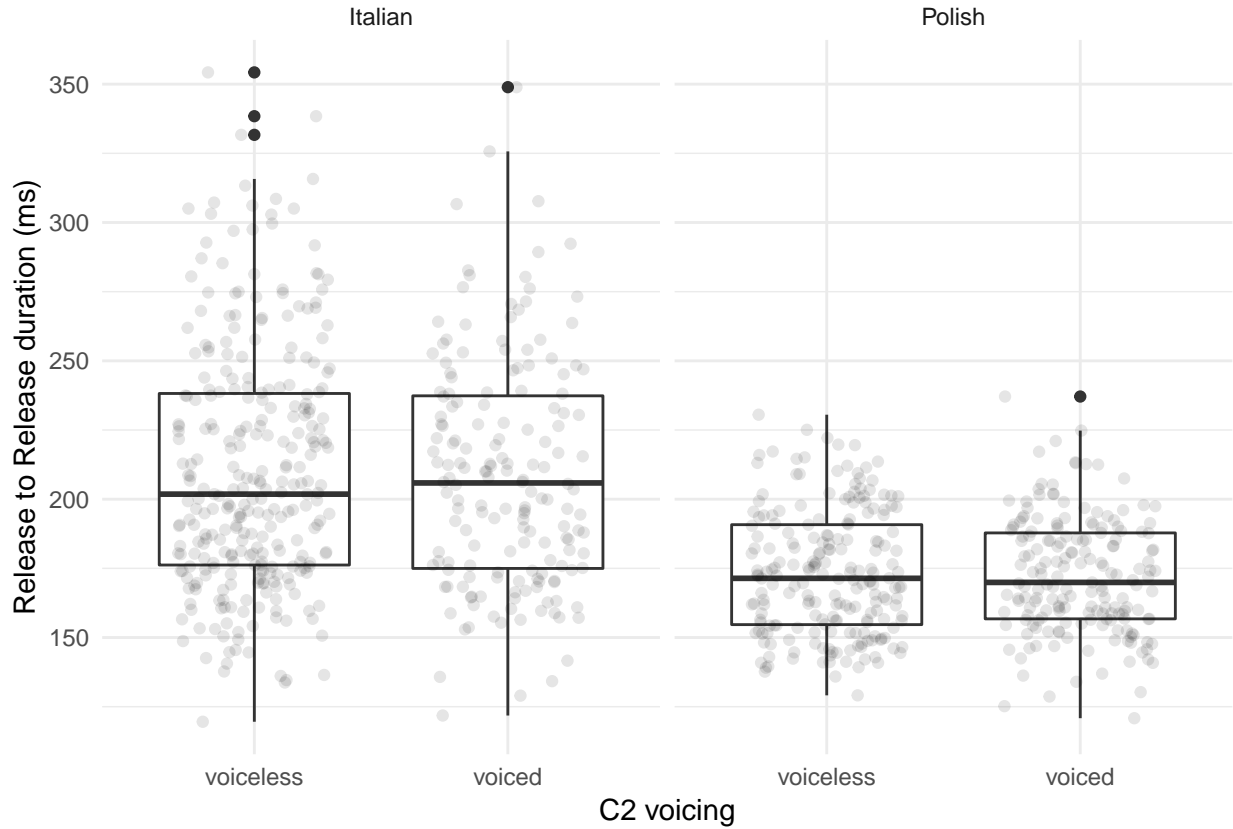


FIG. 5. Release to Release interval duration.

full model is 23, which means that the null model (without C2 voicing) is 23 times more likely than the full model. The data suggests there is positive evidence that duration of the RR interval is not affected by C2 voicing.

#### IV. DISCUSSION

The data and statistical analyses of this exploratory study suggest that the duration of interval between the releases of two consecutive consonants in  $C\acute{V}CV$  words (the Release to Release interval) is insensitive to the phonological voicing of the second consonant (C2) in Italian and Polish. In accordance with a compensatory temporal adjustment account (Lehiste, 1970b; Slis and Cohen, 1969b),

the difference in vowel duration before voiceless vs. voiced stops can be seen as the outcome of differences in stop closure duration. More specifically, the timing of the closure onset of C2 within the invariant Release to Release interval determines the duration of the preceding vowel. An earlier closure onset (like in the case of voiceless stops), relative to the onset of the preceding vowel, causes the vowel to be shorter. On the other hand, a later closure onset (like with voiced stops) produces a longer vowel. Hence, the voicing effect would be a by-product of gestural timing, rather than a consequence of characteristics of voicing *per se*.

A limitation of the compensatory temporal adjustment account, in the variant of [Slis and Cohen \(1969b\)](#) and [Lehiste \(1970b\)](#), is the lack of a precise identification of the internal structure of the word as the unit within which the compensation happens. It is not clear, for example, why compensation should target the preceding stressed vowel, rather than the following unstressed vowel or any other segment in the word. The invariance of Release to Release interval allows us to formulate a better defined internal organisation of the word. Since the Release to Release interval contains (in a simplified view) just the vowel (broadly defined as a vocoid gesture) and the consonant closure, it follows by logic that differences in closure duration must affect the preceding vowel.

The invariance of the Release to Release interval can be seen as the explanans of the voicing effect, and it gives the compensatory account a more logical consistency by narrowing the scope of the temporal adjustment action. On the other hand, the Release to Release invariance becomes in turn an explanandum and it thus requires to be motivated. In the following section, I offer a gestural organisation account that allows the invariance of such interval to follow from the relative timing of the articulatory gestures in a CVC sequence.

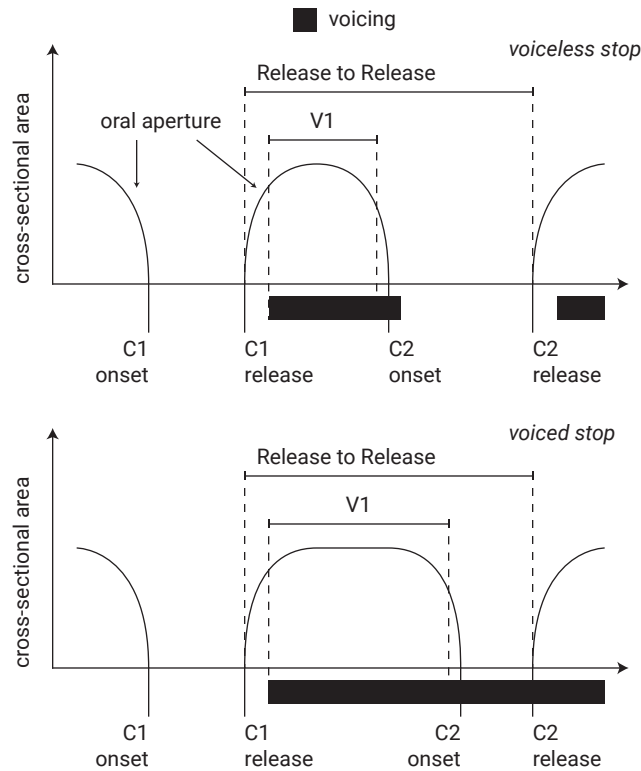


FIG. 6. A schematic representation of the voicing effect as a compensatory temporal adjustment phenomenon.

The schematic show the gestural unfolding of a  $CVC$  sequence when  $C2$  is voiceless (top panel), or voiced (bottom panel). Oral cavity aperture (on the  $y$ -axis, as the inverse of oral constriction) through time (on the  $x$ -axis) is represented with a changing black line that represents the movement trajectory of an articulator. Lower values represent a more constricted oral tract (a contoid configuration), while higher values indicate a more open oral tract (a vocoid configuration). The black bars below the time axis represent voicing (vocal fold vibration). Various landmarks and intervals are indicated in the schematic.

## A. Gestural alignment

According to the coupled oscillator model of syllabic structure (Browman and Goldstein, 1988; 2000; Goldstein *et al.*, 2006; Goldstein and Pouplier, 2014), articulatory gestures can be timed ac-

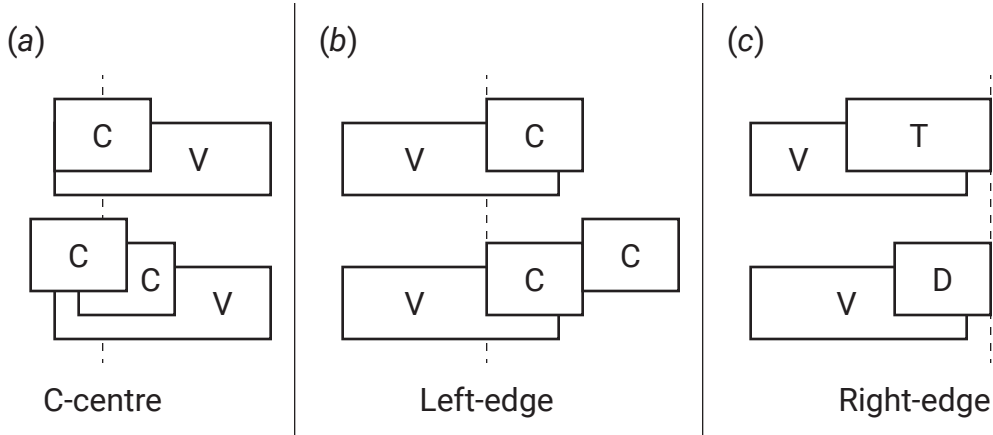


FIG. 7. Gestural organisation patterns for onsets (a), codas (b), heterosyllabic onsets (c). See Section IV A for details. Based on [Marin and Pouplier \(2010\)](#).

281 cording to two coupling modes: in-phase (synchronous) mode, by which two gestures start in syn-  
 282 chrony, or anti-phase (sequential) mode, in which one gesture starts when the preceding one has  
 283 reached its target. [Marin and Pouplier \(2010\)](#) showed that onset consonants in American English are  
 284 in-phase with the vowel nucleus and anti-phase with each other. Such phasing pattern establishes a  
 285 stable relationship between the centre of the consonant or consonant cluster and the following vowel.  
 286 Independent of the number of onset consonants, the midpoint of the onset, the so-called ‘C-centre’,  
 287 is maintained at a fixed distance from the vowel, such that increasing number of consonants in the  
 288 onset does not change the C-centre/vowel distance (Figure 7(a)). On the other hand, coda consonants  
 289 are timed anti-phase with the preceding vowel and between themselves. Stability in codas is seen in  
 290 the lag between the vowel and the left-most edge of the coda, which is not affected by the number  
 291 of coda consonants (Figure 7(b)). Other studies found further evidence for the synchronous and se-  
 292 quential coupling modes (see extensive review in [Marin and Pouplier \(2010\)](#) and [Marin and Pouplier](#)

(2014)), although the use of one mode over the other depends on the language and the consonants under study.

Consonants can thus be said to follow either a C-centre organisation pattern or a left-edge organisation pattern. In both cases, of course, the pattern is relative to the tautosyllabic vowel (the following vowel for onsets, the preceding vowel for codas). To the best of my knowledge, no study has reported the timing of onset consonants relative to the *preceding* (heterosyllabic) vowel. The results from this acoustic study on Italian and Polish are compatible with a right-edge organisation pattern for onset consonants and preceding stressed vowels Figure 7(c). The release of C2 (which is the onset of the second syllable in CVCV words)—which can be thought as the acoustic parallel of the articulatory right edge of C2—is invariantly timed relative to V1 (which is the nucleus of the first syllable).

A consequence of a right-edge organisation pattern of C2 relative to V1 in CVCV words is that differences in C2 closure duration do not affect the lag between V1 and the release of C2, as shown by the results of this study. The invariance of the lag between the release of C1 and that of C2 then can be seen to follow from the invariance in timing between, on the one hand, C1 (which is always /p/ in this study) and V1, and, on the other, between V1 and the right edge of C2.

A right-edge organisation account is compatible with findings from electromyographic, x-ray microbeam, and ultrasonic data by, respectively, Raphael (1975), de Jong (1991), and Celata *et al.* (2018). Celata *et al.* (2018) show that vowels before tautosyllabic clusters have the same duration as before heterosyllabic clusters. However, vowels followed by geminates are shorter than when followed by singletons, although from a syllabic structure point of view geminates correspond to heterosyllabic clusters and singletons to tautosyllabic clusters (i.e., V-final syllables followed by singletons and tautosyllabic clusters are open, while those followed by geminates and heterosyllabic clusters are

closed). [Celata \*et al.\* \(2018\)](#) argue that these results corroborate a rhythmic account in which the relevant unit is the rhythmic syllable, i.e. the VC(C) sequence (independent of the traditional syllabic structure), which is kept constant. Such view reflects a gestural timing view in which the timing of the right edge of the consonant is held constant relative to the vowel.

[de Jong \(1991\)](#) reports that the closing gesture of voiceless stops (following stressed vowels) is faster than that of voiced stops, and that also it is timed earlier with respect to the opening gesture of the stressed vowel. According to [de Jong \(1991\)](#), the differences in vowel duration are driven by the timing of the consonantal closing gesture relative to the vocalic opening gesture (also see [Hertrich and Ackermann 1997](#)). Moreover, the data in [de Jong \(1991\)](#) show that the final portion of the vocalic opening gesture is prolonged before voiced stops. This finding corresponds to what [Raphael \(1975\)](#) reported based on electromyographic data. The electromyographic signal corresponding to the vocalic gesture reaches its plateaux at the same time in the voiceless and voiced context, but the plateaux is held for longer in the case of vowels followed by voiced stops, indicating that muscular activation is kept for longer.

These studies taken together, plus the results from this study, bring evidence to the view that two factors contribute to the difference in vowel duration observed before consonants varying in their voicing specification. These two factors are: (1) the right-edge alignment of coda consonants following stressed vowels relative to the latter, and (2) the differential timing of the closing gesture onset for voiceless vs. voiced stops. These two factors together can be synthesised into a compensatory temporal adjustment account, in which the fixed interval is generated by factor (1) and the temporal adjustment is brought about by factor (2).

## B. Limitations and future work

The generalisations reported in this paper strictly apply to disyllabic words with a stressed vowel in the first syllable. It is possible that the organisation pattern found in this context does not occur in sequences including an unstressed vowel. For example, it is known that the difference in closure duration between voiceless and voiced stops is not stable when the stops precede a stressed vowel, although the vowels preceding the pre-stress stops have different durations (Davis and Van Summers, 1989). According to the gestural interpretation given here, the absence in differences of closure duration should correspond to no difference in vowel duration. Data from different contexts and different languages is thus needed to assess the generality of the claims put forward in this paper.

The constraints on experimental material enforced by the use of ultrasound tongue imaging have been previously mentioned in Section II C. Given these constraints, temporal information from other vowels (like front vowels) and places of articulation is a desideratum. Section IV A discusses the interpretation of the Release to Release invariance in CVCV words as a consequence of the timing of C2 rather than of a holistic CVC motor plan in which the RR interval is held constant. Although beyond the scope of this paper, disambiguating between these two interpretations on articulatory grounds is fundamental for a general understanding of a theory of gestural organisation.

The compensatory temporal adjustment account presented here extends to other durational effects discussed in the literature. In particular, the account bears predictions on the direction of the durational difference led by phonation types different from voicing, like aspiration and ejection. For example, the mix of results with regard to the effect of aspiration (Durvasula and Luo, 2012) suggests that the conditions for a temporal adjustment might differ across the contexts and languages studied.

357 In light of the results in [Beguš \(2017\)](#), future studies will have to investigate the durational invariance  
358 of speech intervals in relation to a variety of phonation contrasts.

## 359 **V. CONCLUSION**

## 360 **ACKNOWLEDGMENTS**

361 Thanks to...



362 **APPENDIX A: OUTPUT OF STATISTICAL MODELS**363 **1. Vowel duration**

364

term	estimate	std.error	df	statistic	p.value	conf.low	conf.high
(Intercept)	202.5289	8.6169	134.7948	23.5036	0.0000	185.6400	219.4178
c2_phonationvoiced	18.9669	4.3898	12.7785	4.3207	0.0009	10.3631	27.5707
vowelo	-6.1457	3.9512	8.6900	-1.5554	0.1555	-13.8899	1.5985
vowelu	-26.3039	3.9772	8.9199	-6.6136	0.0001	-34.0991	-18.5087
languagePolish	-24.2194	8.1708	21.7230	-2.9642	0.0072	-40.2338	-8.2050
c2_placevelar	-8.1827	1.6984	10.5938	-4.8178	0.0006	-11.5116	-4.8539
syl_rate	-15.2920	1.2679	775.7483	-12.0608	0.0000	-17.7771	-12.8070
c2_phonationvoiced:vowelo	-2.0453	5.8662	10.5314	-0.3487	0.7342	-13.5428	9.4522
c2_phonationvoiced:vowelu	-14.4536	5.8040	10.0977	-2.4903	0.0318	-25.8292	-3.0780
c2_phonationvoiced:languagePolish	-7.9928	6.4252	14.2528	-1.2440	0.2336	-20.5860	4.6005
vowelo:languagePolish	-3.6121	5.7389	9.6704	-0.6294	0.5437	-14.8601	7.6360
vowelu:languagePolish	1.6149	5.7695	9.8777	0.2799	0.7853	-9.6931	12.9230
c2_phonationvoiced:vowelo:languagePolish	-2.9987	8.3627	10.8862	-0.3586	0.7268	-19.3894	13.3920
c2_phonationvoiced:vowelu:languagePolish	7.9601	8.3077	10.6040	0.9582	0.3593	-8.3227	24.2428

**2. Closure duration**

term	estimate	std.error	df	statistic	p.value	conf.low	conf.high
(Intercept)	119.7338	7.2100	128.2742	16.6065	0.0000	105.6023	133.8652
c2_phonationvoiced	-16.5825	4.3129	17.8144	-3.8449	0.0012	-25.0356	-8.1294
vowelo	3.6830	3.4951	9.0918	1.0538	0.3192	-3.1672	10.5333
vowelu	-1.9898	3.5174	9.3243	-0.5657	0.5849	-8.8837	4.9041
languagePolish	-6.9400	6.8688	22.0443	-1.0104	0.3233	-20.4027	6.5226
c2_placevelar	3.4024	1.4976	10.9532	2.2719	0.0443	0.4672	6.3376
syl_rate	-8.4278	1.0550	557.6472	-7.9887	0.0000	-10.4954	-6.3601
c2_phonationvoiced:vowelo	1.1040	5.1738	10.8916	0.2134	0.8350	-9.0364	11.2445
c2_phonationvoiced:vowelu	9.9882	5.1257	10.4981	1.9486	0.0786	-0.0581	20.0344
c2_phonationvoiced:languagePolish	1.6759	6.5019	20.0145	0.2578	0.7992	-11.0675	14.4194
vowelo:languagePolish	-0.2681	5.0672	10.0440	-0.0529	0.9588	-10.1997	9.6635
vowelu:languagePolish	7.1432	5.0932	10.2505	1.4025	0.1903	-2.8393	17.1256
c2_phonationvoiced:vowelo:languagePolish	1.5022	7.3707	11.2269	0.2038	0.8422	-12.9441	15.9485
c2_phonationvoiced:vowelu:languagePolish	-3.2088	7.3279	10.9696	-0.4379	0.6700	-17.5711	11.1536

3. Vowel and closure duration

term	estimate	std.error	df	statistic	p.value	conf.low	conf.high
(Intercept)	219.3142	10.4477	123.5512	20.9917	0.0000	198.8371	239.7913
closure_duration	-0.1487	0.0632	50.3807	-2.3532	0.0226	-0.2726	-0.0249
vowelo	-2.0462	5.4702	81.5530	-0.3741	0.7093	-12.7675	8.6751
vowelu	-5.0236	5.5582	86.7938	-0.9038	0.3686	-15.9176	5.8703
syl_rate	-17.5364	1.2855	896.1529	-13.6415	0.0000	-20.0559	-15.0168
closure_duration:vowelo	-0.0973	0.0615	876.5971	-1.5835	0.1137	-0.2178	0.0231
closure_duration:vowelu	-0.3500	0.0619	895.3921	-5.6582	0.0000	-0.4712	-0.2288

APPENDIX B: SOCIO-LINGUISTIC INFORMATION OF PARTICIPANTS

APPENDIX C: TARGET WORDS

<sup>1</sup>Two accounts that point to perceptual features are [Javkin \(1976\)](#) and [Kluender \*et al.\* \(1988\)](#). To the best of my knowledge, [Javkin \(1976\)](#)'s proposal remains empirically untested, while see [Fowler \(1992\)](#) for arguments against [Kluender \*et al.\* \(1988\)](#).

<sup>2</sup>[Durvasula and Luo \(2012\)](#) does not find a negative correlation between vowel duration and consonant closure duration, but rather a (small) positive effect: vowel duration increases with closure duration when including an voicing and aspi-

TABLE II. Participants' sociolinguistic information.

ID	Age	Sex	Native L	Other Ls	City of birth	Spent most time in	> 6 mo
it01	29	Male	Italian	English, Spanish	Verbania	Verbania	Yes
it02	26	Male	Italian	Friulian, English, Ladin-Venetan	Udine	Tricesimo	Yes
it03	28	Female	Italian	English, German	Verbania	Verbania	No
it04	54	Female	Italian	Calabrese	Verbania	Verbania	No
it05	28	Female	Italian	English	Verbania	Verbania	No
it09	35	Female	Italian	English	Vignola	Vignola	Yes
it11	24	Male	Italian	English	Monza	Monza	Yes
it13	20	Female	Italian	English, French, Arabic, Farsi	Ancona	Chiaravalle	Yes
it14	32	Male	Italian	English, Spanish	Frosinone	Frosinone	Yes
pl02	32	Female	Polish	English, Norwegian, French, German, Dutch	Koło	Poznań	Yes
pl03	26	Male	Polish	Russian, English, French, German	Nowa Sol	Poznań	Yes
pl04	34	Female	Polish	Spanish, English, French	Warsaw	Warsaw	No
pl05	42	Male	Polish	English, French	Przasnysz	Warsaw	No
pl06	33	Male	Polish	English	Zgierz	Zgierz	Yes
pl07	32	Female	Polish	English, Russian	Bielsk Podlaski	Bielsk Podlaski	Yes

TABLE III. Target words.

Italian			Polish		
pata	poto*	putu	pata	poto	putu
pada	podo	pudu	pada*	podo	pudu
paca*	poco*	pucu	paka*	poko	puku
paga*	pogo	pugu	paga	pogo	pugu

376 ration as fixed effects. However, it is likely that this result is a consequence of not controlling for speech rate, so it will  
377 not be discussed here.

378 <sup>3</sup>These estimates should be taken as a gross approximation. There are several issues: number of speakers, different  
379 contexts, statistical modelling.

380 <sup>4</sup>IT01 and IT02 (the first two participants of this study) read also sentences with words starting with /b/, which were later  
381 excluded from the experimental design. The data from /b/-initial words are not included in the analysis reported in this  
382 paper.

383 <sup>5</sup>[Luke \(2017\)](#) argues that the common approach of using likelihood ratio tests for statistical inference with mixed models  
384 leads to inflated Type I error rates. [Luke \(2017, 1501\)](#) also warns that ‘results should be interpreted with caution,  
385 regardless of the method adopted for obtaining *p*-values’.

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387 Ananthapadmanabha, T. V., Prathosh, A. P., and Ramakrishnan, A. G. (2014). “Detection of the  
388 closure-burst transitions of stops and affricates in continuous speech using the plosion index,” The  
389 Journal of the Acoustical Society of America **135**(1), 460–471.

- 390 Bates, D., Mächler, M., Bolker, B., and Walker, S. (2015). “Fitting linear mixed-effects models using  
391 lme4,” *Journal of Statistical Software* **67**(1), 1–48.
- 392 Beguš, G. (2017). “Effects of ejective stops on preceding vowel duration,” *The Journal of the Acous-  
393 tical Society of America* **142**(4), 2168–2184, doi: [10.1121/1.5007728](https://doi.org/10.1121/1.5007728).
- 394 Belasco, S. (1953). “The influence of force of articulation of consonants on vowel duration,” *The  
395 Journal of the Acoustical Society of America* **25**(5), 1015–1016.
- 396 Bigi, B. (2015). “SPPAS - Multi-lingual approaches to the automatic annotation of speech,” *The  
397 Phonetician* **111–112**, 54–69.
- 398 Browman, C. P., and Goldstein, L. (1988). “Some notes on syllable structure in articulatory phonol-  
399 ogy,” *Phonetica* **45**(2-4), 140–155.
- 400 Browman, C. P., and Goldstein, L. (2000). “Competing constraints on intergestural coordination and  
401 self-organization of phonological structures,” *Bulletin de la communication parlée* (5), 25–34.
- 402 Caldognetto, E. M., Ferrero, F., Vaggies, K., and Bagno, M. (1979). “Indici acustici e indici percettivi  
403 nel riconoscimento dei suoni linguistici (con applicazione alle consonanti occlusive dell’italiano),”  
404 *Acta Phoniatica Latina* **2**, 219–246.
- 405 Celata, C., Meluzzi, C., and Bertini, C. (2018). “Stressed vowel durational variations and articulatory  
406 cohesiveness: Italian data” Poster presented at LabPhon 16, Lisbon.
- 407 Chen, M. (1970). “Vowel length variation as a function of the voicing of the consonant environment,”  
408 *Phonetica* **22**(3), 129–159.
- 409 Davis, S., and Van Summers, W. (1989). “Vowel length and closure duration in word-medial VC  
410 sequences,” *The Journal of the Acoustical Society of America* **17**, 339–353.

- 411 de Jong, K. (1991). “An articulatory study of consonant-induced vowel duration changes in english,”  
412 *Phonetica* **48**(1), 1–17.
- 413 Durvasula, K., and Luo, Q. (2012). “Voicing, aspiration, and vowel duration in Hindi,” *Proceedings*  
414 *of Meetings on Acoustics* **18**, 1–10.
- 415 Esposito, A. (2002). “On vowel height and consonantal voicing effects: Data from Italian,” *Phonetica*  
416 **59**(4), 197–231.
- 417 Farnetani, E., and Kori, S. (1986). “Effects of syllable and word structure on segmental durations in  
418 spoken Italian,” *Speech communication* **5**(1), 17–34.
- 419 Fowler, C. A. (1992). “Vowel duration and closure duration in voiced and unvoiced stops: There are  
420 no contrast effects here,” *Journal of Phonetics* **20**(1), 143–165.
- 421 Fox, J. (2003). “Effect displays in R for generalised linear models,” *ournal of Statistical Software*  
422 **8**(15), 1–27, doi: [10.18637/jss.v008.i15](https://doi.org/10.18637/jss.v008.i15).
- 423 Gelman, A., and Loken, E. (2013). “The garden of forking paths: Why multiple comparisons can be  
424 a problem, even when there is no “fishing expedition” or “p-hacking”and the research hypothesis  
425 was posited ahead of time,” *Department of Statistics, Columbia University* .
- 426 Goldstein, L., Byrd, D., and Saltzman, E. (2006). “The role of vocal tract gestural action units in  
427 understanding the evolution of phonology,” in *Action to Language via the Mirror Neuron System*,  
428 edited by M. A. Arbib (Cambridge: Cambridge University Press), pp. 215–249.
- 429 Goldstein, L., and Pouplier, M. (2014). “The temporal organization of speech,” in *The Oxford hand-*  
430 *book of language production*, edited by V. Ferreira, M. Goldrick, and M. Miozzo (Oxford: Oxford  
431 University Press).

- 432 Hajek, J., and Stevens, M. (2008). "Vowel duration, compression and lengthening in stressed syllables  
433 in central and southern varieties of standard italian," ISCA.
- 434 Halle, M., and Stevens, K. (1967). "Mechanism of glottal vibration for vowels and consonants," The  
435 Journal of the Acoustical Society of America **41**(6), 1613–1613.
- 436 Heffner, R.-M. (1937). "Notes on the length of vowels," American Speech **12**, 128–134.
- 437 Hertrich, I., and Ackermann, H. (1997). "Articulatory control of phonological vowel length contrasts:  
438 Kinematic analysis of labial gestures," The Journal of the Acoustical Society of America **102**(1),  
439 523–536.
- 440 House, A. S., and Fairbanks, G. (1953). "The influence of consonant environment upon the secondary  
441 acoustical characteristics of vowels," The Journal of the Acoustical Society of America **25**(1), 105–  
442 113.
- 443 Hussein, L. (1994). "Voicing-dependent vowel duration in Standard Arabic and its acquisition by  
444 adult american students," Ph.D. thesis, The Ohio State University.
- 445 Jacewicz, E., Fox, R. A., and Lyle, S. (2009). "Variation in stop consonant voicing in two regional  
446 varieties of American English," Journal of the International Phonetic Association **39**(3), 313–334,  
447 doi: [10.1017/S0025100309990156](https://doi.org/10.1017/S0025100309990156).
- 448 Jarosz, A. F., and Wiley, J. (2014). "What are the odds? a practical guide to computing and reporting  
449 Bayes factors," The Journal of Problem Solving **7**(1), 2–9, doi: [10.7771/1932-6246.1167](https://doi.org/10.7771/1932-6246.1167).
- 450 Javkin, H. R. (1976). "The perceptual basis of vowel duration differences associated with the  
451 voiced/voiceless distinction," Report of the Phonology Laboratory, UC Berkeley **1**, 78–92.
- 452 Keating, P. A. (1984). "Universal phonetics and the organization of grammars," UCLA Working  
453 Papers in Phonetics **59**.



- Kerr, N. L. (1998). "HARKing: Hypothesizing after the results are known," *Personality and Social Psychology Review* **2**(3), 196–217.
- Klatt, D. H. (1973). "Interaction between two factors that influence vowel duration," *The Journal of the Acoustical Society of America* **54**(4), 1102–1104.
- Kluender, K. R., Diehl, R. L., and Wright, B. A. (1988). "Vowel-length differences before voiced and voiceless consonants: An auditory explanation.," *Journal of Phonetics* **16**, 153–169.
- Kuznetsova, A., Bruun Brockhoff, P., and Haubo Bojesen Christensen, R. (2017). "lmerTest package: Tests in linear mixed effects models," *Journal of Statistical Software* **82**(13), doi: [10.18637/jss.v082.i13](https://doi.org/10.18637/jss.v082.i13).
- Laeuffer, C. (1992). "Patterns of voicing-conditioned vowel duration in French and English," *Journal of Phonetics* **20**(4), 411–440.
- Lampp, C., and Reklis, H. (2004). "Effects of coda voicing and aspiration on Hindi vowels," *The Journal of the Acoustical Society of America* **115**(5), 2540–2540.
- Lehiste, I. (1970a). "Temporal organization of higher-level linguistic units," *The Journal of the Acoustical Society of America* **48**(1A), 111–111.
- Lehiste, I. (1970b). "Temporal organization of spoken language," in *Working Papers in Linguistics*, Vol. 4, pp. 96–114.
- Lindblom, B. (1967). "Vowel duration and a model of lip mandible coordination," *Speech Transmission Laboratory Quarterly Progress Status Report* **4**, 1–29.
- Lisker, L. (1974). "On "explaining" vowel duration variation," in *Proceedings of the Linguistic Society of America*, pp. 225–232.

- 475 Luke, S. G. (2017). "Evaluating significance in linear mixed-effects models in R," Behavior Research  
476 Methods **49**(4), 1494–1502, doi: [10.3758/s13428-016-0809-y](https://doi.org/10.3758/s13428-016-0809-y).
- 477 Maddieson, I., and Gandour, J. (1976). "Vowel length before aspirated consonants," in *UCLA Work-*  
478 *ing papers in Phonetics*, Vol. 31, pp. 46–52.
- 479 Malisz, Z., and Klessa, K. (2008). "A preliminary study of temporal adaptation in Polish vc groups,"  
480 in *Proceedings of Speech Prosody*, pp. 383–386.
- 481 Marin, S., and Pouplier, M. (2010). "Temporal organization of complex onsets and codas in American  
482 English: Testing the predictions of a gestural coupling model," Motor Control **14**(3), 380–407.
- 483 Marin, S., and Pouplier, M. (2014). "Articulatory synergies in the temporal organization of liquid  
484 clusters in Romanian," Journal of Phonetics **42**, 24–36.
- 485 Nowak, P. (2006). "Vowel reduction in Polish," Ph.D. thesis, University of California, Berkeley.
- 486 Peterson, G. E., and Lehiste, I. (1960). "Duration of syllable nuclei in english," The Journal of the  
487 Acoustical Society of America **32**(6), 693–703.
- 488 Plug, L., and Smith, R. (2018). "Segments, syllables and speech tempo perception" Talk presented  
489 at the 2018 Colloquium of the British Association of Academic Phoneticians (BAAP 2018).
- 490 R Core Team (2018). "R: A language and environment for statistical computing" R Foundation for  
491 Statistical Computing, Vienna, Austria, <https://www.R-project.org>.
- 492 Raftery, A. E. (1995). "Bayesian model selection in social research," Sociological methodology 111–  
493 163.
- 494 Raftery, A. E. (1999). "Bayes factors and BIC: Comment on "A critique of the Bayesian information  
495 criterion for model selection"," Sociological Methods & Research **27**(3), 411–427.

- 496 Raphael, L. J. (1975). "The physiological control of durational differences between vowels preceding  
497 voiced and voiceless consonants in English," *Journal of Phonetics* **3**(1), 25–33.
- 498 Roettger, T. B. (2018). "Researcher degrees of freedom in phonetic sciences" Pre-print available at  
499 PsyArXiv, doi: [10.31234/osf.io/fp4jr](https://doi.org/10.31234/osf.io/fp4jr).
- 500 Slis, I. H., and Cohen, A. (1969a). "On the complex regulating the voiced-voiceless distinction I,"  
501 *Language and speech* **12**(2), 80–102.
- 502 Slis, I. H., and Cohen, A. (1969b). "On the complex regulating the voiced-voiceless distinction II,"  
503 *Language and speech* **12**(3), 137–155.
- 504 Sóskuthy, M. (2013). "Phonetic biases and systemic effects in the actuation of sound change," Ph.D.  
505 thesis, University of Edinburgh.
- 506 Vazquez-Alvarez, Y., and Hewlett, N. (2007). "The 'trough effect': an ultrasound study," *Phonetica*  
507 **64**(2-3), 105–121.
- 508 Wagenmakers, E.-J. (2007). "A practical solution to the pervasive problems of p values," *Psycho-*  
509 *nomic bulletin & review* **14**(5), 779–804.
- 510 Warren, W., and Jacks, A. (2005). "Lip and jaw closing gesture durations in syllable final voiced and  
511 voiceless stops," *The Journal of the Acoustical Society of America* **117**(4), 2618–2618.
- 512 Wickham, H. (2017). "tidyverse: Easily install and load the 'tidyverse'" R package version 1.2.1.,  
513 <https://CRAN.R-project.org/package=tidyverse>.