Optical Variability of Quasars with 20-Year Photometric Light Curves

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ABSTRACT

We study the optical gri photometric variability of a sample of 190 quasars within the SDSS Stripe 82 region that have long-term photometric coverage during $\sim 1998-2020$ with SDSS, PanSTARRS-1, the Dark Energy Survey, and dedicated follow-up monitoring with Blanco 4m/DECam. With on average ~ 200 nightly epochs per quasar per filter band, we improve the parameter constraints from a Damped Random Walk (DRW) model fit to the light curves over previous studies with 10-15 yr baselines and $\lesssim 100$ epochs. We find that the average damping timescale τ_{DRW} continues to rise with increased baseline, reaching a median value of ~ 750 days (g band) in the rest-frame of these quasars using the 20-yr light curves. Some quasars may have gradual, long-term trends in their light curves, suggesting that either the DRW fit requires very long baselines to converge, or that the underlying variability is more complex than a single DRW process for these quasars. Using a subset of quasars with better-constrained τ_{DRW} (less than 20% of the baseline), we confirm a weak wavelength dependence of $\tau_{DRW} \propto \lambda^{0.51\pm0.20}$. We further quantify optical variability of these quasars over days to decades timescales using structure function (SF) and power spectrum density (PSD) analyses. The SF and PSD measurements qualitatively confirm the measured (hundreds of days) damping timescales from the DRW fits. However, the ensemble PSD is steeper than that of a DRW on timescales less than \sim a month for these luminous quasars, and this second break point correlates with the longer DRW damping timescale.

Key words: surveys – quasars: general – quasars: supermassive black holes

1 INTRODUCTION

The optical photometric (continuum) variability of quasars encodes critical information about physical processes within the accretion disk of a rapidly accreting supermassive black hole (SMBH) that primarily emits in the rest-frame UV through optical. There has been significant progress in the past few decades in quantifying the observed optical variability of quasars with increasing sample sizes and light curve quality (e.g., Giveon et al. 1999; Hawkins 2002; Vanden Berk et al. 2004; de Vries et al. 2005; Sesar et al. 2006; Bauer et al. 2009; MacLeod et al. 2010, 2012; Sun et al. 2014; Morganson et al. 2014; Kasliwal et al. 2015; Chen & Wang 2015; Simm et al. 2016; Caplar et al. 2017; Smith et al. 2018; Sánchez-Sáez

et al. 2018; Li et al. 2018; De Cicco et al. 2019; Luo et al. 2020; Tachibana et al. 2020; Laurenti et al. 2020; Xin et al. 2020; Suberlak et al. 2021). However, the nature of optical variability of quasars is still poorly understood (e.g., Ulrich et al. 1997; Padovani et al. 2017).

Quasars are observed to vary stochastically over a broad range of timescales and wavelengths. In the rest-frame UV-optical, quasar variability amplitude increases with timescales and decreases with wavelength (e.g., Vanden Berk et al. 2004), and is observed to anticorrelate with luminosity and the Eddington ratio of the quasar (e.g., Ai et al. 2010; Rumbaugh et al. 2018). On months to years timescales, quasar optical variability typically saturates at the $\sim 10-20\%$ level. Traditionally, the characterization of quasar variability has been carried out with the structure function (SF) or power spectrum density (PSD) measurements, which quantify the variability level as a function of timescale (or frequency).

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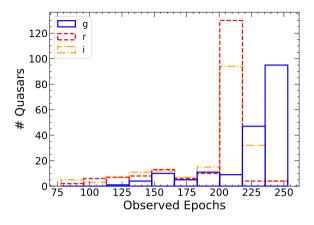


Figure 1. The number of nightly-coadded epochs observed for the 190 quasars in our sample combining SDSS, PS1 and DES (+DECam) data. The *gri* light curves have a median of [205, 209, 209] epochs for our quasars.

It has become increasingly popular in recent years to model quasar light curves in the time domain with stochastic processes (e.g., Kelly et al. 2009; Kozłowski et al. 2010; Kelly et al. 2014). This approach addresses concerns of sampling and windowing effects that come with time series analyses in the frequency domain, which are particularly relevant for ground-based quasar light curves. The Damped Random Walk (DRW) model has emerged as the simplest Gaussian random process model that can fit the optical light curves of quasars reasonably well (e.g., Kelly et al. 2009; Kozłowski et al. 2010; MacLeod et al. 2010). Deviations from the DRW model have been reported (e.g., Mushotzky et al. 2011; Zu et al. 2013; Kasliwal et al. 2015; Guo et al. 2017), although some of these claims are likely impacted by the limited duration of the light curve in the DRW fit (e.g., Kozłowski 2017). More complex Gaussian random process models, such as the continuous auto-regressive moving-average (CARMA; Kelly et al. 2014) models, can accommodate a broader range of PSD shapes, and improve the fits provided that the light curve quality is sufficiently high.

In the DRW model, the PSD is described by a f^{-2} power-law at the high-frequency end, transitioning to a white noise at the lowfrequency end. The transition frequency f_0 corresponds to the damping timescale τ_{DRW} as $f_0 = 1/(2\pi\tau_{DRW})$. The damping timescale thus describes a characteristic timescale of the optical variability. Earlier studies of quasar variability already hinted at such a characteristic variability timescale and its possible dependence on the physical properties of quasars such as the black hole mass (e.g., Collier & Peterson 2001; Kelly et al. 2009), but the exact form of the dependence is debated (e.g., MacLeod et al. 2010; Simm et al. 2016). Recently, Burke et al. (2021) measured the damping timescales using the DRW model for a sample of Active Galactic Nuclei (AGNs) with high-quality optical light curves over a large dynamic range in black hole mass. They found a strong positive correlation between τ_{DRW} and black hole mass, which extends to the stellar mass regime with optical variability measured for nova-like accreting white dwarfs (Scaringi et al. 2015). Compared with higher-order Gaussian process models, the DRW model contains a single characteristic timescale, making it easier to interpret the variability and to connect variability to the underlying physical processes (e.g., Burke et al. 2021; Sun et al. 2020).

However, as Kozłowski (2017) pointed out, in order to constrain the damping timescale τ_{DRW} when fitting the light curve with a DRW model, it is important that the duration of the light curve is

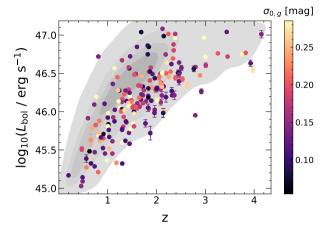


Figure 2. The distribution of 190 SDSS S82 quasars in our sample in the bolometric luminosity versus redshift plane. The individual targets are color-coded by their intrinsic rms variability in the g-band ($\sigma_{0,g}$), calculated using a maximum-likelihood approach described in Shen et al. (2019). The gray contours behind the data points represent the distribution of $L_{\rm bol}$ and z from $\sim 100,000$ SDSS DR7 quasars (Shen et al. 2011), which are on average brighter than SDSS quasars selected in the S82 region.

substantially longer than τ_{DRW} . For light curves shorter than a few times τ_{DRW} , the measured τ_{DRW} can be systematically biased low and saturated around 20–40% of the light curve duration, with elevated scatter in the measurements. Many of the DRW fits to SDSS Stripe 82 quasars in MacLeod et al. (2010) do not pass this duration test, and their reported τ_{DRW} values may be underestimated. Suberlak et al. (2021) extended the Stripe 82 light curves by another 5 years using the PanSTARRS-1 (PS1) data (Chambers et al. 2016), which alleviated this problem. But many of the updated τ_{DRW} measurements are still not short enough compared with the baseline. In addition, the number of PS1 epochs is small compared with the SDSS data, and the DRW fits are likely still dominated by the SDSS light curves.

The main purpose of this work is to study optical continuum variability of a sample of quasars with a more extended 20-yr baseline. This sample represents one of the best-quality light curve data sets to study quasar variability, with hundreds of epochs from SDSS, PS1 and the high-cadence/high-S/N monitoring from the Dark Energy Survey, as well as our dedicated follow-up photometric monitoring with DECam on the CTIO-4m Blanco telescope. We will improve the DRW measurements using these extended light curves and quantify the general optical variability properties with SF and PSD analyses.

This paper is organized as follows. In §2 we describe the sample and the photometric light curve data. In §3 we present our variability measurements, with the technical details provided in Appendix A. We discuss the implications of our results in §4 and conclude in §5. Throughout this paper we adopt a flat Λ CDM cosmology with cosmological parameters $\Omega_{M,0}=0.3$ ($\Omega_{\Lambda,0}=0.7$) and $H_0=70\,\mathrm{kms}^{-1}\mathrm{Mpc}^{-1}$. By default all timescales are in the rest-frame of the quasar unless otherwise specified.

2 DATA

To study optical quasar variability with long-term light curves, we utilize quasars identified in the SDSS Stripe 82 region (S82), a nearly 300 deg^2 stripe along the celestial equator, imaged by SDSS from \sim

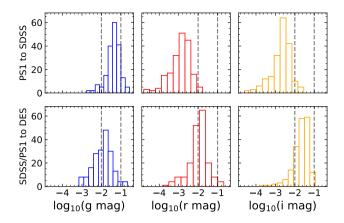


Figure 3. Histograms of the photometric offsets used for each target in each survey. The top row represents offsets from PS1 bands to SDSS bands, and the bottom row represents offsets from the combined PS1/SDSS bands to DES bands. The three columns represent the *gri* bands in corresponding order from left to right. The dashed lines represent 0.01 and 0.1 mag corrections for each band.

1998 to 2007. S82 was repeatedly observed to find supernova, being one of the most frequently observed areas in SDSS. Each target within S82 was repeatedly observed for an average of 60 epochs, albeit aperiodically and with large time gaps, as the observing window spanned 2-3 months each year. SDSS photometry has five bandpasses ((ugriz)_{SDSS}) available for each quasar, allowing for the study of variability as a function of wavelength. The SDSS light curves in S82 provide an initial 10-year baseline for quasar variability studies (e.g., MacLeod et al. 2010). To extend this baseline, we use data from PS1 (Chambers et al. 2016) spanning nearly 5 years during 2010-2014. PS1 imaged the sky in the $(grizy)_{PS1}$ bandpasses with ~ 2 epochs per year in its wide-area survey. The combined SDSS+PS1 light curves for S82 quasars have a baseline of ~ 15 yrs, and were used to study quasar variability in Suberlak et al. (2021) to improve the DRW fits. However, there were only a handful of PS1 epochs, and the DRW fits were potentially dominated by the SDSS data.

To extend our baseline further, we use data from the DES survey during 2013-2019, which imaged the sky in the $(grizy)_{\rm DES}$ bandpasses. In particular, among the repeatedly observed DES Transient Survey (Deep) Fields (Hartley et al. 2022), there were two in the S82 region (SN-S1 and SN-S2, centered at J2000 coordinates 02:51:16.8+00:00:00.0 and 02:44:46.7–00:59:18.2, respectively), each with 2.7 deg² area, with > 100 epochs over six years. After DES completed its wide field survey in 2019, we continued to monitor these two S82 DES-deep fields with a dedicated long-term program (2019-2024) using the DECam imager on the CTIO-4m telescope (NOAO program 2019B-0219; PI: X. Liu) to extend the baseline further in 3 bands $((gri)_{\rm DES})$.

In this work we use the combined light curve data from SDSS, PS1, DES and DECam imaging for 190 spectroscopically confirmed quasars in SDSS that are within the two DES-deep fields in S82 (Fig. 1 and Fig. 2). These quasars are all within the SDSS DR7 quasar catalog, with derived physical properties such as bolometric luminosities and black hole masses from Shen et al. (2011). Our combined baseline is ~ 20 yrs, enabling a detailed quasar variability study over decades-long timescales. The inclusion of the DES and DECam imaging is of critical importance: it not only extends the baseline by another 6 years to improve the constraints on the damping

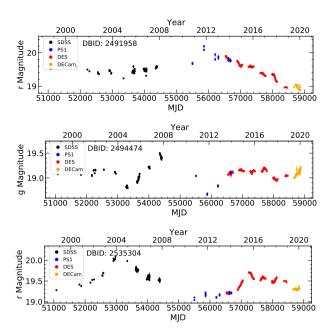


Figure 4. Example light curves from our quasar sample. These light curves are taken over a ~ 20 year baseline, across different surveys. To adjust the observed magnitudes in a common band, we apply empirical color offsets and additional small (~ 0.05 mag) offsets to merge the light curves. The data for all light curves are provided in the FITS table described in Table 1

timescale, but also provides a large number of high-S/N epochs to sample days to years timescales and to ensure the DRW fits are not dominated by the SDSS epochs.

All of these quasars have observations in the *gri* bands for all surveys, so we focus on these three bands for multi-wavelength variability. Although *z*-band data are also available across most of these surveys, the variability amplitude in this red band is lower and host contamination would be more significant, thus complicating the quasar variability measurements.

We obtain public SDSS light curve data for each of these quasars from the catalog curated in MacLeod et al. (2012), which provides light curves for nearly 9000 SDSS S82 quasars in all five *ugriz* bandpasses. We obtain public PS1 photometry for each quasar using the MAST database (https://archive.stsci.edu/), querying for all *gri* bands and excluding detections with low confidence. The proprietary DES data and our dedicated DECam imaging data are processed with the same DES pipeline (Morganson et al. 2018). We use PSF magnitudes from all these surveys for our quasars.

The filter bandpasses differ slightly between SDSS, PS1, and DES, and we apply photometric offsets to obtain merged light curves in a common bandpass for each quasar. Photometric offsets are typically constructed using colors of objects rather than magnitudes themselves, as these colors are less variable. We choose to use the mean color-based offsets described in Liu et al. (2016) to offset PS1 data into the corresponding SDSS bands, and then use the offsets described in Drlica-Wagner et al. (2018) to offset both SDSS and PS1 magnitudes into the corresponding DES bandpasses. Fig. 3 shows that most of the corrections between surveys lie under 0.1 mag for each band. PS1 *r,i* magnitudes are sufficiently similar to SDSS *r,i* magnitudes so that no correction is needed, but we opt to do so for a similar processing of all bands. All other bandpasses for each survey have small offsets, with only a handful of objects with offsets up to

0.3 mag. Therefore, the use of these mean color photometric offsets is justified for our sample.

After correcting for the zero-point offset in the same bandpass, we find that the r- and i-band light curves still display a small offset between the overlapping PS1 and DES epochs for some quasars. This additional offset is likely due to the usage of PSF magnitudes, extended host galaxy emission, seeing variations between PS1 and DES observations, as well as any residual systematics bewteen surveys. We therefore apply an additional correction (~ 0.05 mag) to manually bring the overlapping PS1 and DES epochs into agreement. We have tested w/ and w/o this minor magnitude offset between PS1 and DES and found that this detail has no effect on our variability analyses.

We show a few representative examples of the merged light curves from SDSS+PS1+DES+DECam in Fig. 4. We summarize the basic properties of our quasar sample in a FITS table along with the best-fit DRW parameters, where we compile additional properties of these quasars from the catalog in Shen et al. (2011). The columns of this FITS table are described in Table 1. We also provide all light curve data in the FITS table described in Table 1.

3 RESULTS

3.1 DRW Fits

We follow the standard practice in the literature to fit a DRW model to the quasar light curve (e.g., Kelly et al. 2009; Kozłowski et al. 2010; MacLeod et al. 2010; Suberlak et al. 2021; Burke et al. 2021). The details of the DRW modeling are provided in Appendix A2. The best-fit DRW parameters are compiled in the FITS catalog described in Table 1. An example DRW fit is shown in Fig. 5.

In Fig. 6, we show the distribution of our sample in the τ_{DRW} versus $SF_{\infty}^2 \equiv 2\sigma_{DRW}^2$ plane, where σ_{DRW} is the long-term variability amplitude in the DRW model (see Appendix A2). With the SDSS-only baselines, we reproduce the results in MacLeod et al. (2010), with a median value for $\tau_{DRW,rest}$ of ~ 540 days in the r band. Using SDSS+PS1-only baselines, however, we obtain a median value for $\tau_{DRW,rest}$ of ~ 680 days in the r band, while Suberlak et al. (2021) quoted a value of ~ 550 days. We attribute this discrepancy to the method of choosing the best-fit value from the DRW fit (discussed further in Appendix A5). By extending the baseline further with the DES+DECam data, the values of τ_{DRW} and SF_{∞} continue to rise. The median values of $\tau_{DRW,rest}$ and SF_{∞} for S82 quasars with our final baselines are ~ 750 days and 0.25 mag in g band.

Fig. 7 compares the $\tau_{DRW,obs}$ values measured with different baselines. Similar to the results shown in Fig. 6, the best-fit $\tau_{DRW,obs}$ continues to increase as the baseline increases. With longer baselines and more epochs, the constraints on τ_{DRW} are somewhat tighter, as demonstrated by the lower scatter of points with the SDSS+PS1 and SDSS+PS1+DES+DECam data than with the SDSS-only data in Fig. 7. However, the formal measurement uncertainties on τ_{DRW} are only reduced by $\sim 10\%$ on average from the SDSS-only measurements to the SDSS+PS1+DES+DECam measurements. It is possible that the formal measurement uncertainties underestimated the true uncertainties on τ_{DRW} in these studies.

Kozłowski (2017) emphasized the importance of the length of the light curve in constraining the DRW damping timescale. The best-fit τ_{DRW} could be significantly underestimated if the light curve is not long enough, as independently confirmed in other studies with simulated light curves (e.g., Suberlak et al. 2021; Burke et al. 2021). The fact that the average τ_{DRW} continues to rise as the baseline

increases indicates that even the 20-year baseline is probably not long enough to well constrain τ_{DRW} in some S82 quasars. On the other hand, the increasing τ_{DRW} as the baseline increases may be due to gradual, long-term trends in the quasar light curve (see further discussion in Appendix A2), or it is possible that these quasar light curves are more complex than a simple DRW process with only one characteristic timescale.

Nevertheless, simulations with mock light curves have shown that the systematic bias in τ_{DRW} is not significant, albeit with elevated scatter, when the measured τ_{DRW} is less than 20% of the baseline (e.g. Kozłowski 2017; Suberlak et al. 2021; Burke et al. 2021). Indeed, when we compare our best-fit DRW model to the ensemble SF and PSD measurements in §3.2 and §3.3, we find that these DRW fits and the associated damping timescales are qualitatively correct on average.

Next, we investigate the wavelength dependence of τ_{DRW} using our measurements. To reduce the impact of poorly constrained τ_{DRW} values from insufficient baselines, we only use a subset of 27 quasars with measured τ_{DRW} less than 20% of the baseline, for which we consider the constraint on the damping timescale is more reliable. Fig. 8 (left) displays the wavelength dependences of τ_{DRW} and SF_{∞} . We find a weak wavelength dependence of $\tau_{DRW} \propto \lambda^{0.51\pm0.20}$, which is slightly steeper than (but formally consistent within 2σ) the one reported in MacLeod et al. (2010) based on the much shorter SDSSonly light curves $\tau_{DRW} \propto \lambda^{0.17}$. On the other hand, we recover a weak anti-correlation between SF_∞ and wavelength, but our dynamic range in wavelength is more limited than that in MacLeod et al. (2012), given that we only use data in gri bands. These constraints on wavelength dependences are weak given the small number of quasars that pass the baseline criterion. If we use the full sample of 190 quasars instead, we find slightly different, but fully consistent results (right panel of Fig. 8).

3.2 Structure Function Analysis

The structure function measures the magnitude difference for pairs of epochs separated at different timescales, and is a simple and useful empirical tool to characterize the variability of quasars (e.g., Collier & Peterson 2001; Kozłowski 2016b). Unlike the DRW model, the SF measurements are model-independent, and provide empirical constraints on variability amplitude as a function of timescales. However, unlike the DRW and CARMA modelling, the SF approach does not rigorously deal with the flux uncertainties of each epoch, and unequal flux uncertainties for long-term pairs from different surveys may complicate the SF calculation. We therefore only use these SF measurements to provide a qualitative comparison with the more rigorous DRW and CARMA PSD fits.

For the SF analysis and the PSD analysis in §3.3, we will focus the discussion on the results using g-band data as we did not find significantly new information based on the r and i-band data. However, all the individual and ensemble SF and PSD measurements for the 3 bands are compiled in Tables 1 and 2.

We measure the SF for individual quasars in our sample as well as for the ensemble average. We have followed Kozłowski (2016b) to calculate the SF after subtracting photometric uncertainties (e.g., from flux uncertainties and additional systematics from host galaxy light and seeing variations) using close pairs separated by less than ~ 10 days in rest-frame. Fig. 9 and Fig. 10 display the ensemble SF for different subsamples, where the full sample is divided into subsamples with approximately the same number of objects in each division (either by $\tau_{\rm DRW}$ or by $L_{\rm bol}/z$).

Fig. 9 compares the ensemble SF with the median DRW model for

Table 1. The format of the FITS table compiling the properties for our sample of 190 quasars in S82.

^a Column Name	Format	Unit	Description
DBID	int64		Database ID for each quasar; same as in MacLeod et al. (2012)
RA	float64	deg	Right ascension of the target
DEC	float64	deg	Declination of the target
Z	float64		Redshift
log_M_BH	float64	$\log_{10}(M_{\odot})$	log ₁₀ of the black hole mass
log_M_BH_ERR	float64	$\log_{10}(M_{\odot})$	Error in log ₁₀ of the black hole mass
log_LBOL	float64	$\log_{10}(\text{erg s}^{-1})$	log ₁₀ of the bolometric luminosity
log_LBOL_ERR	float64	$\log_{10}(\text{erg s}^{-1})$	Error in log ₁₀ of the bolometric luminosity
log_TAU_OBS_x	float64	log ₁₀ (days)	$\log_{10}(au_{ m DRW})$ in the observed-frame
log_TAU_OBS_x_ERR_L	float64	$log_{10}(days)$	Lower error of $\log_{10}(\tau_{DRW})$ in the observed-frame
log_TAU_OBS_x_ERR_U	float64	log ₁₀ (days)	Upper error of $\log_{10}(au_{ m DRW})$ in the observed-frame
log_TAU_REST_x	float64	$log_{10}(days)$	$\log_{10}(au_{ m DRW})$ in the rest-frame
log_TAU_REST_x_ERR_L	float64	$\log_{10}(\text{days})$	Lower error of $\log_{10}(\tau_{\mathrm{DRW}})$ in the rest-frame
log_TAU_REST_x_ERR_U	float64	$\log_{10}(\text{days})$	Upper error of $\log_{10}(au_{\mathrm{DRW}})$ in the rest-frame
log_SIGMA_x	float64	$log_{10}(mag)$	$\log_{10}(\sigma_{ m DRW})$
log_SIGMA_x_ERR_L	float64	$\log_{10}(\text{mag})$	Lower error of $\log_{10}(\sigma_{ m DRW})$
log_SIGMA_x_ERR_U	float64	$log_{10}(mag)$	Upper error of $\log_{10}(\sigma_{ m DRW})$
log_JITTER_x	float64	$\log_{10}(\text{mag})$	$\log_{10}(\sigma_{\mathrm{n}})$
log_JITTER_x_ERR_L	float64	$log_{10}(mag)$	Lower error of $\log_{10}(\sigma_n)$
log_JITTER_x_ERR_U	float64	$log_{10}(mag)$	Upper error of $\log_{10}(\sigma_n)$
SIG0_x	float64	mag	Intrinsic RMS variability
SIG0_x_ERR	float64	mag	Error in intrinsic RMS variability
LAMBDA_REST_x	float64	Å	Rest-frame wavelength the target was observed in
^b SURVEY_x	str5		Imaging survey used for the observation
^b MJD_x	float64	days	MJD of the observation
^b MAG_x	float64	mag	PSF magnitude of the observation
^b MAG_ERR_x	float64	mag	Error in the observation
^c OFFSET_x	float64	mag	Manual offset applied to the PS1 magnitudes
DT_REST_x	float64	days	Rest-frame time lags used to construct the structure function
SF_x	float64	mag	Structure function measurements
SF_x_ERR_L	float64	mag	Lower error in the structure function
SF_x_ERR_U	float64	mag	Upper error in the structure function
CARMA_P_x	int64		CARMA model p parameter
CARMA_Q_x	int64		CARMA model q parameter
REST_FREQ_x	float64	days ⁻¹	Rest-frame frequency
CARMA_PSD_x	float64	(mag) ² (days)	Median PSD constructed from the CARMA model
CARMA_PSD_x_ERR_L	float64	$(mag)^2$ (days)	Lower error in the CARMA PSD
CARMA_PSD_x_ERR_U	float64	$(mag)^2$ (days)	Upper error in the CARMA PSD
dCARMA_AR0_x	float64		0^{th} CARMA auto-regressive parameter (α_0)
dCARMA_AR1_x	float64		1 st CARMA auto-regressive parameter (α_1)
dCARMA_AR2_x	float64		2^{nd} CARMA auto-regressive parameter (α_2)
dCARMA_AR3_x	float64		$3^{\rm rd}$ CARMA auto-regressive parameter (α_3)
dCARMA_AR4_x	float64		4 th CARMA auto-regressive parameter (α_4)
^d CARMA_AR5_x	float64		5 th CARMA auto-regressive parameter (α_5)
^d CARMA_AR6_x	float64		6^{th} CARMA auto-regressive parameter (α_6)
dCARMA_AR7_x	float64		7^{th} CARMA auto-regressive parameter (α_7)
dCARMA_MA0_x	float64		0^{th} CARMA moving-average parameter (β_0)
dCARMA_MA1_x	float64		1 st CARMA moving-average parameter (β_1)
^d CARMA_MA2_x	float64		2^{nd} CARMA moving-average parameter (β_2)
^d CARMA_MA3_x	float64		3^{rd} CARMA moving-average parameter (β_3)
^d CARMA_MA4_x	float64		4 th CARMA moving-average parameter (β_4)
A	float64		5 th CARMA moving-average parameter (β_5)
^d CARMA_MA5_x ^d CARMA_MA6_x	1104104		5 CARMA moving-average parameter (p5)

^a Each column labeled with "x" is three columns, with "x" representing the value obtained from data in the g, r, or i bands.

^bFITS tables require that each entry in a column of data have the same length. However, each object has a different amount of epochs, making their data arrays unequal. To circumvent this, we have made the arrays corresponding to properties of the observations of the object (SURVEY, MJD, MAG, MAG_ERR) the same length. This length is the number of observations for the object with the maximum number of observations in the sample. For arrays with a length less than this maximum length, we fill the arrays with NaNs or empty strings until they reach this length.

^cThis manual offset is used to bring the PS1 and DES magnitudes into agreement in the overlapping region. Offsets were only applied to r-band and i-band light curves, so the "x" here corresponds to r and i only.

dAll of the entries for the CARMA parameters are given as 3-entry arrays, consisting of the 1σ errors (absolute values) and median value of the parameter. This array is formatted as [lower error, value, upper error]. If the CARMA model fit to the light curve data is not a high enough order to have a certain parameter, it will have an array filled with zeros. For example, if the CARMA p parameter is 3, all CARMA auto-regressive parameters greater than 3 will be [0, 0, 0] in the FITS table.

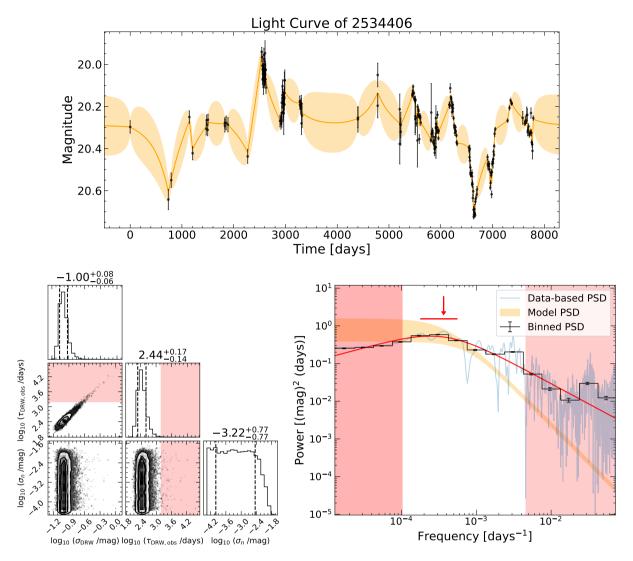


Figure 5. An example of fitting the g-band light curve with the DRW model using the fast Gaussian process solver Celerite (discussed further in A2). The top panel displays the raw light curve of the object, and the predicted light curve from the DRW model using the best-fit, maximum likelihood parameters. The orange line represents the median value of the prediction, while the shaded orange region represents the area between the 1σ uncertainty in the prediction. The plot on the lower-left displays the probability distributions of the DRW parameters fit for by Celerite, with σ_{DRW} representing the standard deviation of long-term variability, τ_{DRW} here representing the observed-frame characteristic timescale, and σ_n representing a noise term (also called jitter). The shaded regions in the probability distributions correspond to where $\tau_{DRW,obs}$ is greater than 20% of the baseline. The lower right plot shows the PSD of the light curve from both the raw data and drawing from the posterior distribution of the Celerite fit. The model PSD is shown in orange (with a band spanning the 1σ uncertainties), the Lomb-Scargle periodogram (Lomb 1976; Scargle 1982) is shown in blue, and the binned Lomb-Scargle periodogram is shown in black. The binned Lomb-Scargle periodogram was also fit to a broken power law (shown as a red line), whose break frequency (and corresponding 1σ errors) are shown with the red arrow and bar. The regions shaded red in the PSD plot correspond to regions of frequency space not sampled by the light curve (i.e. larger than the minimum cadence) as well as regions with timescales longer than 20% of the baseline (i.e. $t > t_{baseline}/5$). The difference between the Lomb-Scargle periodogram and the model PSD is caused by the difficulties of measuring the PSD accurately using the Fourier method and irregularly sampled light curves, contributions from flux uncertainties in the periodogram measurement, as well as potential deviations from a DRW model.

subsets of quasars binned by the measured τ_{DRW} . The SF does show a flattening roughly around the location of τ_{DRW} measured from the DRW fits, indicating the presence of such a damping timescale on the order of hundreds of days.

We also recover the well known dependences of variability amplitude on wavelength and luminosity of quasars using ensemble SF measurements (data required to generate these plots are provided in Table 2).

3.3 Power Spectrum Density Analysis

We measure the optical variability PSD using our sample and light curve data set. Because our light curves are irregularly sampled with large seasonal gaps, it is challenging to directly measure the PSD using the Fourier method, which suffers from aliasing and power leakage from windowing effects. Instead, we take advantage of the recent development of fitting Gaussian random process models to time series data and recovering the PSD (Kelly et al. 2014). Such an alternative approach is more robust in measuring the PSD with sparsely and irregularly sampled light curve data (e.g., Kelly et al.

Table 2. The format of the FITS table compiling ensemble SF and PSD measurements from subsets of our full quasar sample.

^a Column Name	Format	Unit	Description
^b Subsample	str9		Description of the ensemble
DBIDs_x	float64		Database IDs of the objects included in the ensemble
DT_REST_x	float64	days	Rest-frame time lags used to construct the structure function
SF_x	float64	mag	Ensemble structure function measurements
SF_x_ERR	float64	mag	Error in structure function measurements
REST_FREQ_x	float64	days ⁻¹	Rest-frame frequency
CARMA_PSD_x	float64	(mag) ² (days)	Ensemble of the median PSDs of the optimal CARMA models for each object
CARMA_PSD_x_ERR_L	float64	(mag) ² (days)	Lower error in the ensemble PSD
CARMA_PSD_x_ERR_U	float64	$(mag)^2$ (days)	Upper error in the ensemble PSD

^a Similar to Table 1, all columns with names containing an "x" are three separate columns, where x is replaced with *gri*, corresponding to values in each of the three bands.

^bThere are four different types of ensembles described in this table in general: the total sample, the samples split by $\tau_{\text{DRW,rest}}$, and the samples split in a grid by bolometric luminosity and redshift. The total subsample is labeled "Total", the three samples split by $\tau_{\text{DRW,rest}}$ are labeled "Tau{i}" (where i = 1, 2, 3), and the five samples split by luminosity and redshift are labeled "Lz_grid{ij}" (where i, j = 1, 2, 3 represents their placement on the grid).

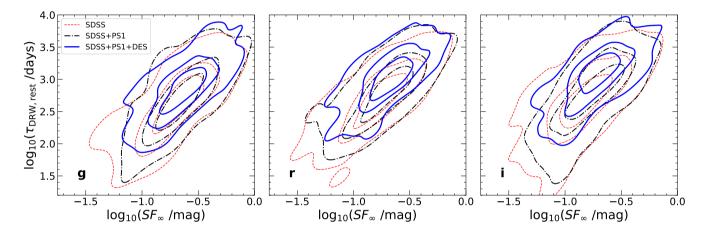


Figure 6. Contour plots showing the distribution of SF_{∞} and τ_{DRW} fitted from our quasar light curve sample. There are three contours for each band, representing data fitted from light curves using only SDSS, SDSS and PS1, and all of the data. The contours for each dataset enclose [33, 66, 100]% of the distribution respectively.

2014; Simm et al. 2016), and properly deals with uneven measurement uncertainties in the light curve.

Specifically, we use the CARMA_pack developed by Kelly et al. (2014) to find the best-fit CARMA(p,q) model to the light curve and derive the PSD, where (p,q) are the numbers of auto-regression (AR) and moving average (MA) terms, respectively. The technical details of CARMA fits are described in Appendix A3. We show an example of PSD analysis in Fig. 11, and all the individual and ensemble PSDs are provided in the FITS catalogs described in Tables 1 and 2.

We show the distributions of the best-fit values of p and q in Fig. 12. There is a tendency of clustering near $p \approx 4$ and $q \approx 1-2$, which may indicate the general similarity of quasar variability PSDs. However, we found that a forced CARMA(2,1) model fit produces very similar PSDs to the ones from the best-fit higher-order CARMA models.

Fig. 13 displays all CARMA PSDs for our sample in the rest-frame of the quasar (only showing the best-fit model), color-coded by different properties. While these individual PSDs overlap considerably given their measurement uncertainties, there are trends of the PSD amplitude and shape with luminosity and black hole mass of the quasar. In addition, the CARMA PSD tends to flatten out sooner at the low-frequency end for light curves with shorter τ_{DRW} , suggest-

ing that the DRW fits are reasonable in constraining the long-term damping timescale.

Fig. 14 shows the ensemble CARMA PSD for the full sample in the three bands. The ensemble PSDs are tightly constrained over days to decade timescales, and show a clear wavelength dependence. Fig. 15 and Fig. 16 display the ensemble PSDs for the same subsets of quasars used in our SF analysis. When divided by the best-fit τ_{DRW} , the ensemble PSD agrees with the average DRW model in the subsample reasonably well, suggesting the DRW model provides a reasonable description of the underlying PSD. However, the more flexible CARMA model reveals a sharper decline in the variability power below timescales of a few weeks than the f^{-2} power-law, consistent with earlier findings with other light curve samples (e.g., Mushotzky et al. 2011; Zu et al. 2013). In Appendix A5, we demonstrate that this PSD steepening at the highest frequencies is not due to the usage of a more flexible CARMA model or selection effects of our data, using simulated light curves. Similar to the SF analysis, the ensemble PSD shows dependences with wavelength and luminosity of the quasar, as shown in Fig. 14 and Fig. 16.

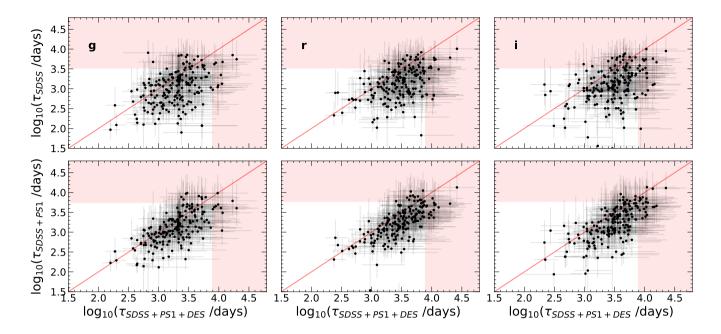


Figure 7. Comparisons between $\tau_{DRW,obs}$ measurements from DRW fits with different baselines. The upper panels compare the $\tau_{DRW,obs}$ fitted from only using SDSS data (a ~10 yr baseline) to the $\tau_{DRW,obs}$ fitted from the entire 20 yr dataset for the three bands. The lower panels compare the $\tau_{DRW,obs}$ fitted using data from SDSS and PS1 (a ~15 yr baseline) to those fitted using the full light curves. The red shaded area indicates where $\tau_{DRW,obs}$ is greater than the light curve baseline, on each respective axis. The red line running through the data shows the unity relation.

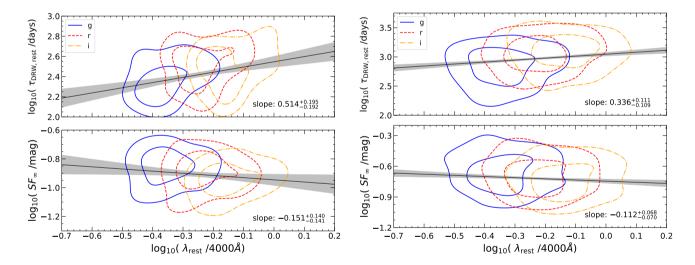


Figure 8. Wavelength dependences of both τ_{DRW} and SF_{∞} . The left panels are for a subset of 27 quasars for which all measured $\tau_{DRW,obs}$ values are less than 20% of the final baseline. The right panels are for the full sample of 190 quasars. The contours in blue, red, and orange represent the results from g, r, and i light curves, respectively, shifted to the corresponding rest-frame wavelengths of each individual quasar. The contours for each band represent 30 and 70 percent of the data. The best-fit linear regression model and 1σ uncertainties using the method described in Kelly (2007) are shown in the black line and shaded area, with the best-fit slopes marked in each panel.

4 DISCUSSION

4.1 The wavelength dependence of τ_{DRW}

We find that the DRW damping timescale only weakly depends on wavelength, consistent with earlier studies with shorter light curves. This weak wavelength dependence of the damping timescale is difficult to interpret: if τ_{DRW} tracks the local timescale of the accretion disk, e.g., the thermal timescale, then we expect a stronger wavelength dependence of this timescale because the local thermal timescale

scales with the emitting wavelength as $\tau \propto \lambda^2$ in the standard α -disk model (Shakura & Sunyaev 1973). One possibility, as suggested by Burke et al. (2021), is that the observed UV/optical variability is driven by processes in the inner (UV-emitting) part of the accretion disk, which rapidly propagates outwards at the Alfvén speed, during which the characteristic variability timescale is more or less preserved. Alternatively, the observed damping timescale may be the thermal timescale averaged over different radii, leading to a shallower wavelength dependence (e.g., Sun et al. 2020). Further development

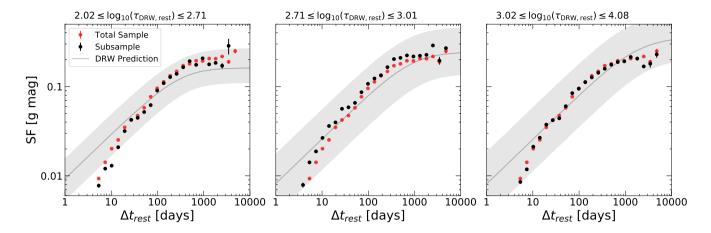


Figure 9. Ensemble structure functions for different ensembles of the 190 quasars in our sample, grouped by their fitted $\tau_{\text{DRW,rest}}$. The ensemble structure function for the full sample is overlayed in red, while the structure functions for the individual subsamples are plotted in black. The predicted structure functions using the fitted SF_{∞} and τ_{DRW} are plotted in gray. To obtain this DRW prediction, we sample 500 predicted DRW structure functions from each target in the ensemble, drawn from a Gaussian distribution using its best-fit DRW parameters and their uncertainties. We then combine the samples for all targets and use the median value in each Δt bin (shown as the gray line) as the DRW-predicted structure function, and the 16^{th} and 84^{th} in each Δt bin (colored in a gray band around the median) to construct the errors in the DRW prediction.

of these theoretical ideas, combined with dedicated global radiation MHD simulations (e.g., Jiang et al. 2019) will shed light on the nature of this long-term characteristic variability damping timescale.

4.2 Validity of the DRW prescription

Overall, we find that the DRW model, even though an empirical prescription, describes the stochastic optical quasar light curves reasonably well over rest-frame timescales from a few months to a few years. The qualitative agreement between the DRW model and SF/PSD measurements suggests that the long-term characteristic variability timescale captured by the DRW model is reliable on average. Indeed, Burke et al. (2021) tested DRW fits to non-DRW light curves with a characteristic long-term turnover timescale in the PSD and found that the best-fit $\tau_{\rm DRW}$ correctly recovers this timescale.

However, the length of the light curve will affect the constraints on τ_{DRW} in a DRW fit (Kozłowski 2017). With our 20-yr baseline, we find that the median $\tau_{DRW,rest}$ for S82 quasars is ~ 750 days in the g band, longer than the median g-band $\tau_{DRW,rest}$ of 450 and 470 days if we use the shorter SDSS-only or SDSS+PS1 light curves. For comparison, MacLeod et al. (2010) and Suberlak et al. (2021) report a median r-band $\tau_{DRW,rest}$ of 570 days using SDSS-only light curves. For r-band and using the SDSS-only light curves, we measure a median $\tau_{DRW,rest}$ of \sim 540 days, consistent with these earlier studies. Details in the adopted "best-fit" DRW parameters and the MCMC convergence criterion in the fitting do not seem to impact our results much (see Appendix A4). It is unclear if this is because some quasars have much longer intrinsic τ_{DRW} than what can be realistically constrained by our current light curves, or because the light curve cannot be described by a single DRW process. For example, if the quasar light curve contains a long-term gradual trend in addition to a DRW process, the best-fit τ_{DRW} will increase as the baseline increases (Appendix A2). Continued monitoring of these quasars in our dedicated DECam program will address this question with even more extended light curves.

On timescales shorter than \sim a month, however, the slope of the PSD is noticeably steeper than -2. In Appendix A5, we use simulated

DRW light curves matched to the observed cadences and S/N of our sample to test if a more flexible CARMA fit would lead to an artificially steeper high-frequency-end slope. We find that the resulting CARMA PSD has a high-frequency-end slope of -2, which confirms that the steeper PSD slope observed in our sample is real. While the exact asymptotic slope of the PSD is likely impacted by the CARMA model restrictions, the locations of the slope transitions are largely determined by the data. There is evidence (e.g., Fig. 15) that this short timescale cutoff of power is positively correlated with the long-term damping timescale. To further illustrate this point, we fit a doubly-broken power-law model to the three ensemble PSDs divided by the measured $\tau_{\text{DRW,rest}}$ in Fig. 15: $P \propto 1/[1 + (f/f_0)^2 +$ $(f/f_1)^4$]. This PSD model fits the three ensemble PSDs reasonably well over years to days timescales, as shown in Fig. 17. The two break timescales, $\tau_0 = 1/(2\pi f_0)$ and $\tau_1 = 1/(2\pi f_1)$, indeed vary in concordance in the three ensembles.

While our sample is small and the dynamic range in black hole mass or quasar luminosity is limited, there is also some tentative evidence that this high-frequency-end break occurs at shorter timescales for lower-luminosity (and less massive) quasars (Fig. 16). This point is further illustrated in Fig. 18, where we compare the ensemble PSDs for subsamples divided by black hole mass. If we assume both break timescales scale with black hole mass as $M_{\rm BH}^{0.5}$ (Burke et al. 2021), we expect much shorter high-frequency break timescales in low-redshift Seyferts ($M_{\rm BH} \sim 10^7~M_{\odot}$) than in SDSS quasars ($M_{\rm BH} \sim 10^9~M_{\odot}$). This may explain the much shorter (a few days) cutoff timescales found for low-redshift, low-luminosity AGNs that are two orders of magnitude less massive than SDSS quasars (e.g., Mushotzky et al. 2011).

The physical origin of the suppression of variability power on timescales shorter than ~ 1 month is unclear. It could be due to the intrinsic shape of the variability PSD, e.g., resulting from the break in the driving variability PSD and/or damping processes in the accretion disk (e.g., Sun et al. 2020). An alternative explanation, as pointed out by, e.g., Tachibana et al. (2020), is due to an averaging effect. Even if the flux of the accretion disk varies coherently, emission from different parts of the disk or from more spatially-extended regions

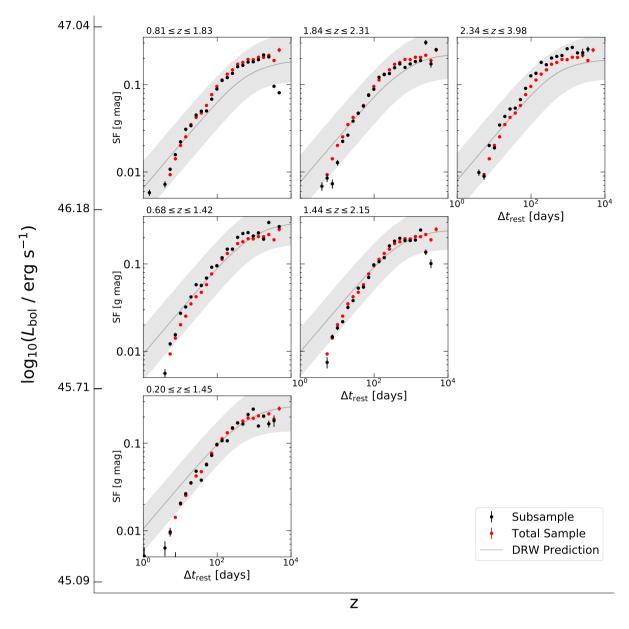


Figure 10. Ensemble g-band structure functions for different subsets of the full sample, grouped by their bolometric luminosity and redshift. Each subsample contains \sim 30 objects. The redshift ranges are listed above each subsample, and the L_{bol} ranges are shown on the leftmost axis, being [45.09, 45.71], [45.71, 46.18] and, [46.18, 47.04]. Each subsample in a given row has the same range of L_{bol} . We have subtracted an "SF floor" seen in time lags below \sim 10 days, to remove contamination from PSF variations and host-galaxy flux (discussed further in Appendix A1). The ensemble SF for the full sample and the DRW-prediction for each subsample are also shown for reference. We constructed the ensemble DRW-predicted structure functions in the same manner as those presented in Fig. 9.

(e.g., an extended diffuse continuum emission region or the broadline region) will arrive at different times. The observed variable flux is then the convolution of the intrinsic variability pattern with the transfer function describing the time delays from different locations. Tachibana et al. (2020) showed that, with a likely transfer function form (a semi-circle with a characteristic timescale of a month), the short-time variability power will be reduced due to averaging, producing a PSD slope close to -4 beyond this characteristic frequency. In general, such transfer functions would reduce the high-frequency power, leading to a steeper high-frequency end slope in the observed PSD. In both the intrinsic PSD scenario and the "smearing" scenario, it is possible that the characteristic timescale of this second high-frequency-end break, which reflects some characteristic size of

the emission region, depends on the physical properties of the quasar, such as the black hole mass (Sun et al. 2020; Tachibana et al. 2020), in a similar way as the long-term damping timescale τ_{DRW} .

5 CONCLUSIONS

Given the simplicity of the DRW model and its reasonable success to fit quasar light curves, it has become increasingly popular to use the DRW prescription to describe stochastic quasar variability. However, the validity of the DRW prescription has yet to be tested with high-quality light curve data that are well sampled, have sufficient baselines and high S/N flux measurements.

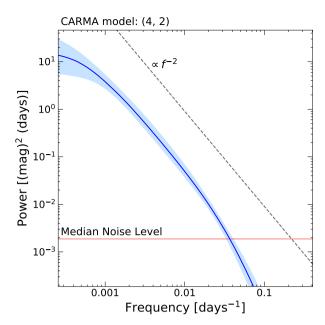


Figure 11. An example of CARMA model-fit PSDs for our quasar sample. The CARMA-predicted PSD (discussed further in A3) is shown in blue, where the median from the posterior is the solid line and the shaded region encloses the 1σ uncertainty range. The median noise level derived from the raw light curve data $(2\times \text{median}(\Delta t)\times \text{median}(\sigma_y^2))$ is shown as the red horizontal line. The gray dashed line indicates a $\propto f^{-2}$ PSD. The CARMA-predicted PSDs for individual targets are compiled in the FITS catalog described in Table 1.

In this work we have measured the optical continuum variability of a sample of 190 quasars from the SDSS Stripe 82 region. Our quasar sample has been photometrically monitored in the SDSS, PS1, DES surveys, as well as our continued monitoring with DECam. The light curves of our sample span a baseline of ~ 20 years with ~ 200 epochs in each of the gri bands. We fit these light curves with the DRW model, and measured the structure function and power spectrum density using the CARMA models. The main findings from our work are the following:

- The best-fit DRW parameters (τ_{DRW} and SF_{∞}) continue to rise with our light curve data, compared with earlier studies with shorter (e.g., 10-yr and 15-yr) baselines from SDSS-only (MacLeod et al. 2010) and SDSS+PS1 (Suberlak et al. 2021). The average rest-frame $\tau_{DRW} \sim 750$ days in g band for S82 quasars with our 20-yr light curves.
- While the τ_{DRW} measurements for many S82 quasars are still not well constrained with the 20-yr light curves, we believe that the bias from insufficient baselines is reduced compared with earlier studies based on shorter baselines, if the underlying variability process is indeed a single DRW. However, we caution that realistic quasar light curves may be more complicated than a single DRW process, e.g., multiple variability processes with different characteristic timescales and/or non-stationary variability processes could be at work. In such cases, the results from a single DRW fit will depend on the baseline. More extended light curves are required to test this possibility.
- Using a subset of 27 quasars for which we have relatively better-constrained $\tau_{\rm DRW}$ in the g, r and i bands, we confirm a weak wavelength dependence of $\tau_{\rm DRW} \propto \lambda^{0.51\pm0.20} \, (\tau_{\rm DRW} \propto \lambda^{0.34\pm0.10}$ for the full sample). This wavelength dependence is slightly stronger than previous results $\tau_{\rm DRW} \propto \lambda^{0.17}$ based on 10-yr light curves (MacLeod et al. 2010), although these results are formally consistent within 2σ .

• We also measured the optical SF and PSD of our quasar sample. The baseline and sampling of our light curves enabled reliable constraints of the ensemble PSD over days to decades timescales. Comparisons between the ensemble SF and PSD with predictions from the best-fit DRW models suggest that the DRW prescription provides a reasonably good description of the variability properties of quasars over months to years timescales. But the average PSD slope on timescales shorter than \sim a month is noticeably steeper than the DRW model, consistent with earlier findings (e.g., Mushotzky et al. 2011; Zu et al. 2013). There is tentative evidence that this high-frequency cutoff timescale correlates with the low-frequency damping timescale $\tau_{\rm DRW}$, hence both timescales may have similar dependences on physical properties of the quasar (e.g., Burke et al. 2021).

We continue to monitor our quasar sample during 2020-2024 as part of our ongoing effort to photometrically monitor deep extragalactic fields with ample multi-wavelength and time-domain data. With another ~ 5 year extension of the baseline and seamlessly merging with light curves from the Vera C. Rubin Observatory Legacy Survey of Space and Time (Ivezić et al. 2019), this quasar sample will become a prime sample to study quasar optical continuum variability. Such studies will further test the applicability of the DRW model and the stationarity of the stochastic variability process, as well as provide insights on the physical origin of quasar variability.

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This research made use of matplotlib, a Python library for publication quality graphics (Hunter 2007). This research made use of SciPy (Virtanen et al. 2020). This research made use of Astropy, a community-developed core Python package for Astronomy (Astropy Collaboration et al. 2018). This research made use of NumPy (Harris et al. 2020). This research made use of pandas (McKinney 2011).

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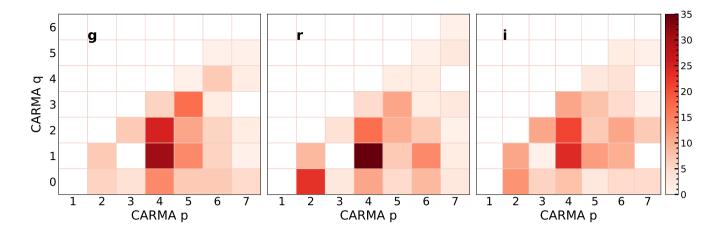


Figure 12. The distribution of best-fit CARMA (p, q) parameters when fitting our quasar light curves to a generalized CARMA model (requiring q < p for stationary processes). Darker colors indicate higher incidence. There is a tendency of clustering of quasars around $(p, q) \approx (4, 2)$.

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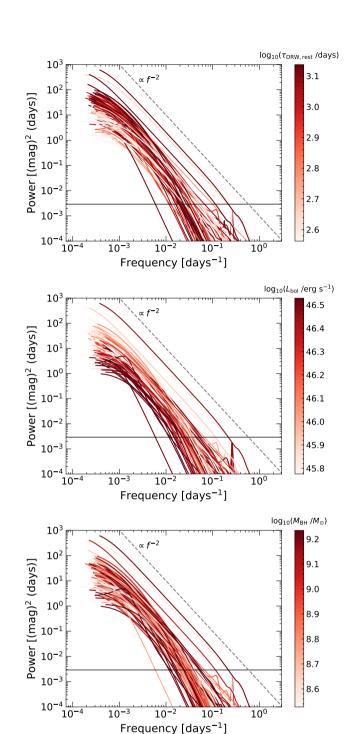


Figure 13. Rest-frame CARMA model PSDs (*g*-band) of all quasars in our parent sample, color-coded by different attributes of the target: $\log_{10}(\tau_{\text{DRW,rest}})$, $\log_{10}(L_{\text{bol}})$, and $\log_{10}(M_{\text{BH}})$. In each panel, the black horizontal line represents the median noise level of the individual light curves, and the dashed gray line represents a $\propto f^{-2}$ PSD.

DATA AVAILABILITY

We provide all light curve data and time series measurements in two online FITS tables. The format of these FITS tables is described in Tables 1 and 2.

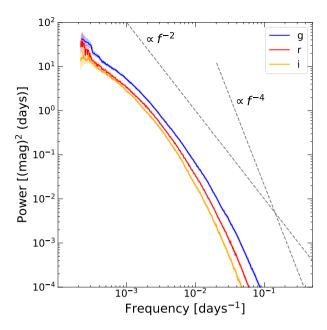


Figure 14. Rest-frame ensemble PSDs for the full sample in gri bands. Each quasar light curve was fit using CARMA_pack, a code designed to fit timeseries data to CARMA models using the method described in Kelly et al. (2014), with optimized (p,q) parameters for the CARMA model. For each PSD, the darker line shows the median value from the ensemble, and the light shaded area (nearly negligible at $f>10^{-3}\,\mathrm{days^{-1}}$) indicates the nominal uncertainty of the ensemble PSD.

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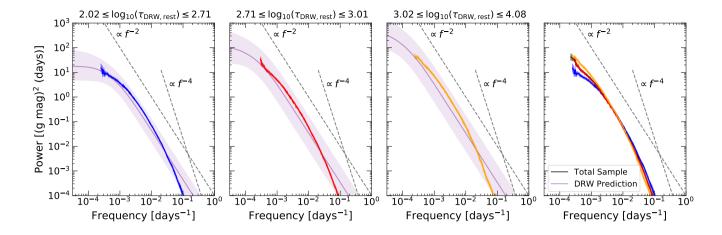


Figure 15. Ensemble CARMA PSDs for subsamples divided by their best-fit τ_{DRW} from §3.1. The first three panels show these ensemble PSDs corresponding to each subsample, whose τ_{DRW} ranges are shown above each panel. The DRW-predicted ensemble PSDs are shown in the purple-shaded area for each ensemble. The ensemble DRW-predicted PSDs are constructed in the same manner as the ensemble DRW-predicted structure functions, shown in Fig. 9. The rightmost panel shows the PSDs of all three ensembles superimposed on the ensemble PSD for the full sample (shown in black). The two gray dashed lines in each panel correspond to f^{-2} and f^{-4} power laws.

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APPENDIX A: LIGHT CURVE ANALYSIS

A1 Structure Function

One of the more traditional ways of modeling the variability of quasars is through structure function analysis. This method describes the change in magnitude as a function of time $\log \Delta t$ between two observations. Since the SF calculation is model-independent, it provides an empirical view of quasar variability with no underlying assumptions. The most basic way to define the structure function is the root-mean-square magnitude difference for a given grid of time lags.

However, without accounting for the flux measurement uncertainties, structure function measurements at small Δt will level off to a certain "SF floor". Therefore, using the method described by Kozłowski (2017), we subtract the measurement errors of both observations in the pair in quadrature:

$$SF(\Delta t) = \sqrt{\frac{1}{N_{\Delta t}} \sum_{i < j} \left((y_i - y_j)^2 - \sigma_i^2 - \sigma_j^2 \right)}, \tag{A1}$$

where σ_i and σ_j are the measurement errors in observations y_i and y_j respectively.

The structure function is related to the auto-correlation function (ACF) of the light curve. Assuming that the variability of the source is stationary, we can use the covariance of two signals to compute the structure function:

$$SF(\Delta t) = \sqrt{2\sigma_s^2(1 - ACF(\Delta t))},$$
 (A2)

where σ_s is the variability amplitude intrinsic to the source. Taking the limit as $\Delta t \to \infty$, we obtain:

$$SF(\Delta t) = SF_{\infty} \sqrt{1 - ACF(\Delta t)}$$
, (A3)

where $SF_{\infty}^2 \equiv 2\sigma_s^2$ is the value of the structure function as $\Delta t \to \infty$. Assuming the variability is stationary (meaning the mean value of the light curve does not change), the difference between signals at large time lags will approach a constant value proportional to the

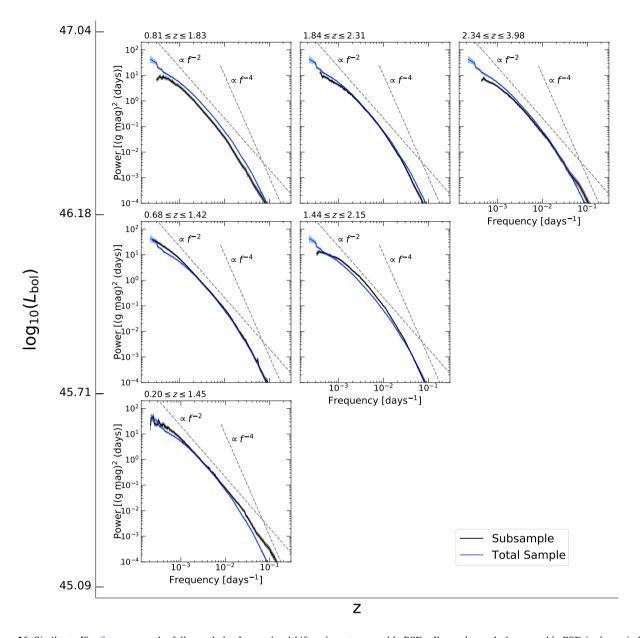


Figure 16. Similar to Fig. 9, we group the full sample by $L_{\rm bol}$ and redshift and create ensemble PSDs. For each panel, the ensemble PSD is shown in blue, the ensemble PSD for the whole sample is shown in black, and the two gray dashed lines indicate a f^{-2} PSD and a f^{-4} PSD. The range of redshifts of each subsample is labeled above each panel, and the range of bolometric luminosities is indicated on the axis on the left. Subsamples in the same row have the same luminosity range (this is not true for the same column with slightly different redshift ranges). The artificial turnover of power at the lowest frequencies is caused by the limited number of objects (~ 10) with the proper temporal coverage.

intrinsic variability amplitude (white noise). The structure function will also flatten to white noise at very short time lags, where the change in magnitude is on the order of the measurement uncertainty.

We utilize Eqn. A1 to make all of our structure function measurements, where time lags are shifted to the rest-frame of the quasar. We also make ensemble structure function measurements for various subsamples of our 190 quasar dataset. To derive ensemble structure functions from individual objects, we bin the structure functions of each individual object into the same Δt grid. We then take the median of each bin to be the ensemble measurement for that time lag, and use the uncertainty on the median (the standard deviation of the samples in each bin, divided by \sqrt{N} the number of samples in the bin) to represent the uncertainty in that measurement. We create these

ensemble structure functions for the total sample, three subsamples grouped by their fitted $\tau_{\rm DRW}$, and six subsamples grouped by their bolometric luminosity and redshift.

For *g*-band measurements, through visual inspection, we observed that the structure function began to rise near time lags of 10 days. However, when measuring these ensemble structure functions, we noticed that they started to flatten at time lags less than days to weeks in the quasar rest frame). This proved to be more prevalent for structure functions in the r and i bands, where the structure function would be almost constant until $\Delta t_{\rm rest} \sim 5$ -10 days and then jump. We attribute this flattening to PSF seeing variations on short timescales, measurement uncertainties, as well as host galaxy contamination. To minimize the effect of this flattening, we perform linear regression

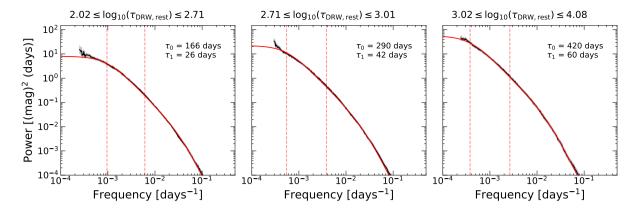


Figure 17. The ensemble PSD in the three subsamples divided by the best-fit $\tau_{\text{DRW,rest}}$ (the black line), and a doubly-broken power-law fit (red line) with two break frequencies f_0 and f_1 . The corresponding break timescales $\tau = 1/(2\pi f)$ are marked in each panel. Both break timescales vary in concordance.

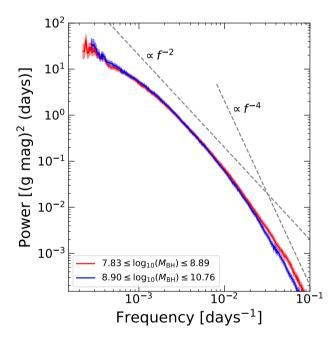


Figure 18. Ensemble PSDs for subsamples divided by the median virial black hole mass in the parent sample. While the dynamic range in black hole mass is limited in our sample, there is some evidence that the short-term break timescale is longer for the high-mass subsample. The mass dependence for the long-term (DRW) break timescale is less obvious, which would require a larger dynamic range in black hole mass (e.g., Burke et al. 2021).

(in log-space) on this floor using the method of Kelly (2007) and subtract the best-fit line from the full ensemble structure function. This floor stopped at different time lags for each band, [5,20,40] days for *gri* measurements respectively, which we use to set the linear regression range.

For each ensemble structure function, excluding for the total sample, we also overlay the DRW model prediction for comparison. The DRW-predicted structure functions are also ensembles, using the predicted $\tau_{\rm DRW}$ and $\sigma_{\rm DRW}$ and Eqn. A6. These ensemble DRW-predicted structure functions are obtained in a similar manner to the ensemble structure functions themselves: we create structure functions for each target in the ensemble using the best-fit DRW parameters, then we bin the structure functions onto a common Δt grid. We

then take the median of each bin to get an ensemble DRW-predicted structure function.

A2 The Damped Random Walk Model

The Damped Random Walk model, also known as the Ornstein-Uhlenbleck process, is a statistical model used to describe the stochastic variability from the accretion disk emission of quasars. This Gaussian process is the simplest model of a family of Gaussian processes known as continuous auto-regressive moving-average (CARMA) models. General CARMA models, discussed in §A3, specify that the output of the model is linear in the current and past terms in the time-series. This is seen in the DRW model (a CARMA(1,0) model), as it has a term that pushes large deviations from the mean of the time-series back towards the mean. It is useful to model light curves with the DRW model as it has parameters that can potentially connect to physical parameters of the quasar, and it can be modelled directly in the time domain instead of the frequency domain. Quasar variability studies in the frequency domain are subject to windowing effects, as large gaps in the data can lead to power leakage and aliases. Using a DRW model (or any CARMA model) can mediate these adverse effects.

All Gaussian processes require a covariance matrix (also known as a kernel), governing the relationship between two points in a time series. In the case of a DRW process, the covariance matrix is

$$k(t_{nm}) = \sigma_{\text{DRW}}^2 \exp(-t_{nm}/\tau_{\text{DRW}}), \qquad (A4)$$

where $t_{nm} = |t_n - t_m|$ and t_n , t_m are times within the time series. σ is the long-term standard deviation of variability, and τ defines a characteristic timescale where the PSD of this process breaks. We can relate this model to the structure function and the PSD in the following way:

$$SF^2(\Delta t) = 2\sigma_{\text{DRW}}^2 \left(1 - e^{-|\Delta t|/\tau_{\text{DRW}}}\right),$$
 (A5)

$$P(f) = \frac{4\sigma_{\text{DRW}}^2 \tau_{\text{DRW}}}{1 + (2\pi f \tau_{\text{DRW}})^2},$$
(A6)

where P(f) is the PSD as a function of frequency f. By comparing Eqn. A5 and Eqn. A1, we have $SF_{\infty}^2 = 2\sigma_{\rm DRW}^2$ and $ACF(\Delta t) = \exp(-|\Delta t|/\tau_{\rm DRW})$. This PSD follows white noise at low frequencies ($\propto f^0$), and transitions to a f^{-2} PSD at higher frequencies below the characteristic timescale $\tau_{\rm DRW}$.

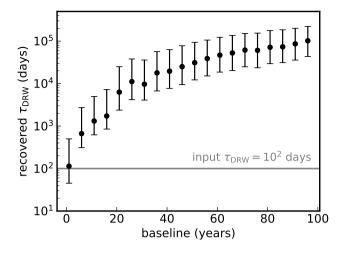


Figure A1. Recovered damping timescales for simulated non-stationary DRW light curves with input $\tau_{DRW} = 100$ days plus a linear trend of 0.0365 mag yr⁻¹ at varying baselines.

In this study, we model our quasar light curves using the fast Gaussian process solver Celerite (Foreman-Mackey et al. 2017). which uses Gaussian process regression to fit the time series to a specified kernel. Given a number of terms in the kernel, and a method to maximize, Celerite can fit a time series to derive the best-fit parameters to said kernel. In our case, we utilize a DRW kernel (specified in Eqn. A4), as well as a term to characterize the effect of a white noise floor from unknown systematic flux errors (σ_n) , also called jitter. We use uniform priors on all parameters within the input Celerite kernel (in log-space), and allow Celerite to minimize the log-likelihood in parameter-space to obtain a set of parameters to fit the light curve. We then use the MCMC sampler emcee (Foreman-Mackey et al. 2013) implemented in Python to draw from the joint posterior probability distribution output from Celerite. The final parameters compiled in the FITS table described in Table 1 are the median samples from these MCMC samples. The upper and lower errors for these parameters are obtained from the 16th and 84th percentiles of the samples. One example Celerite fit is shown in Fig. 5.

However, there are potentially additional features in the light curve that can skew the results of the DRW fit. Here, we investigate the effects of a long-term trend in the light curve on the recovery of $\tau_{\text{DRW,obs}}$ using Celerite and simulated data. We input simulated DRW light curves with input $\tau_{\text{DRW,obs}} = 100$ days, but add a long-term linear trend (non-stationarity) to the light curve, in this case of 1×10^{-4} mag per day. We generate mock light curves using this hybrid model with different baselines, and use Celerite to extract a $\tau_{\text{DRW,obs}}$ from the simulated light curve. The results (shown in Fig. A1), show that as the baseline of the non-stationary light curve increases, the extracted $\tau_{\text{DRW,obs}}$ increases as well. In this test, the input $\tau_{\text{DRW,obs}}$ is 100 days, and is reasonably recovered for short baselines (less than ~ 10 years). However, as the baseline increases, the linear long-term trend starts to skew the recovery of $\tau_{\text{DRW,obs}}$ towards longer and longer damping timescales.

A3 PSD Analysis with CARMA models

While many studies have shown that the DRW model can describe quasar light curve variability to a reasonable degree, we understand that it is not the only model available. It has been shown that stochasic processes generated from non-DRW models can be modeled with DRW (Kozłowski 2016a), albeit with biased DRW parameters. Therefore, to get a true sense of the PSD of quasar light curves and the stochastic processes occurring within their accretion disks, we utilize the more general CARMA model to obtain PSD measurements. Whereas DRW-modeled PSDs are restricted to having a white noise at low frequencies and a f^{-2} PSD at higher frequencies (with a characteristic break timescale in between them), CARMA-predicted PSDs are not restricted to such a shape.

The PSD of a CARMA model is described in the following manner:

$$P(f) = \sigma^2 \frac{|\sum_{j=0}^{q} \beta_j (2\pi i f)^j|^2}{|\sum_{k=0}^{p} \alpha_k (2\pi i f)^k|^2},$$
 (A7)

where σ^2 is the variance of the modeled white noise process, α_j are the auto-regressive parameters of the model, and β_k are the moving-average parameters of the model. The order of the CARMA model is defined by the p and q parameters, which define the number of auto-regressive and moving-average components respectively. The requirement that CARMA processes are stationary also requires that q < p. By convention, we set $\beta_0 = 1$ and $\alpha_p = 1$. When setting p = 1 and q = 0, we recover the DRW PSD, as well as the covariance matrix, where $\tau_{\text{DRW}} = 1/\alpha_0$ and $\sigma_{\text{DRW}} = \sigma \sqrt{\tau_{\text{DRW}}}/2$.

While we have used Celerite to fit the quasar light curves to the DRW model, we now opt to utilize the widely used CARMA_pack code (Kelly et al. 2014) to fit our light curves to a generalized CARMA(p,q) model. While the covariance matrix for the DRW model is somewhat simple, it becomes increasingly complex as the order of the CARMA model is increased, and therefore increasingly more complex to implement into Celerite. The generality of the kernel terms available in Celerite allows the implementation of a large variety in the kernels that can be used, but formulating the CARMA PSD in terms of Celerite's kernel terms is highly involved. CARMA_pack also includes the functionality of choosing an optimal (p,q) of the model used to fit the time series.

We perform the CARMA modeling using time series in the restframe of each quasar. To model our light cures to a generalized CARMA model with CARMA_pack, we obtain the optimal (p,q) of the model. CARMA_pack does this by finding the maximum likelihood estimate of the CARMA models produced from a user-input grid of (p,q) values. We choose to search a parameter space where 1 and all <math>q < p. After using 100 different optimizers initialized to random values, the code picks the (p,q) combination which minimizes the corrected Akaike Information Criterion (AICc, Akaike 1973) provided by Hurvich & Tsai (1989). After choosing the optimal CARMA model for a given object, we use CARMA_pack to derive the maximum likelihood posterior distribution for all of the CARMA parameters. We then use CARMA_pack's MCMC implementation to sample the CARMA parameters, given the order of the model. After testing the effect of the number of iterations of the MCMC on the convergence of fitted parameters (discussed in Appendix A4), we found the results are well convergent for 60,000 iterations and 30,000 burnin samples. After running the MCMC sampler, CARMA_pack will then output samples for all of the CARMA parameters using the posterior distribution of the object's fitted CARMA model. We can then use CARMA_pack to sample the PSD of the light curve given the fitted CARMA model, where we opt to use 10,000 samples. Sim-

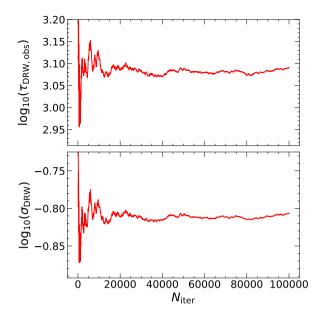


Figure A2. Convergence of fitted DRW parameters for an example quasar light curve, using a CARMA(1,0) model in CARMA_pack, as a function of the number of samples generated by the MCMC sampler. In this study, we opt to use $N_{\text{samp}} = 60,000$, well within the range where both of these parameters cease to vary significantly.

ilar to our structure function analysis, we use the median value of the CARMA parameters and PSD as the best-fit value, and the 16th and 84th percentiles of the samples to obtain the uncertainties in the values.

In a similar manner to Simm et al. (2016), we define a median noise level, $2 \times \text{median}(\Delta t) \times \text{median}(\sigma_y^2)$, for each PSD to define where the PSD is credible. In this expression, Δt is a list of time lags in a given time-series, and σ_y is the measurement uncertainty in the light curve fluxes.

A4 Sampling Methods

One significant step in generating DRW and CARMA parameters for each light curve is the generation of samples from the posterior probability distribution through the use of MCMC sampling. In Celerite, this is done using the popular python-based MCMC sampler emcee, while the sampling in CARMA_pack is done through a custom, C++ MCMC sampler. One important parameter of sampling is the number of burn-in samples and actual samples to use for a given dataset. The burn-in samples for an MCMC sampler help to initialize the sampler to the data and allow it to converge properly. The number of actual samples for an MCMC sampler affects how well the posterior probability distribution for a parameter is sampled. For Celerite DRW fits, we opt to use 500 burn-in samples and 2,000 actual samples, which we found to be the optimal values through trial and error. For the CARMA_pack fits, we adjust the number of burn-in samples relative to the total samples as well as the number of total samples to see where the results from the sample would converge and have low fluctuations. In Fig. A2 we show the evolution of the two DRW parameters over iterations of the sampler when fitting one of our quasar light curves to a CARMA(1,0) model in CARMA_pack. We can see that there are large fluctuations in the sampled value in the early iterations, but the value converges to a set value after $\sim 50,000$ iterations of the sampler. In Fig. A3 we show the evolution

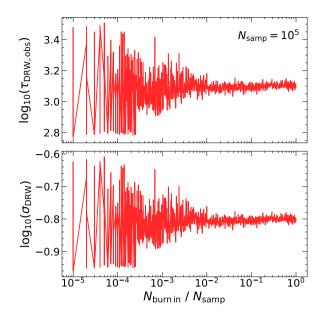


Figure A3. DRW parameters recovered from a DRW fit to an example quasar light curve, using a CARMA(1,0) model in CARMA_pack, as the number of burn-in samples for the MCMC sampler increases. We choose $N_{\rm samp}=10^5$, as Fig. A2 showed that for this number of iterations, both parameters have already converged. We opt to use the default number of burn-in samples that CARMA_pack chooses $(0.5N_{\rm samp})$ for which the output parameters have already converged.

of the sampled parameter values for these DRW parameters, for the same quasar, as the number of burn-in samples increases for a fixed total number of samples of 100,000 (which we have seen has a converged parameter value). This shows that the uncertainty of the value produced with a relatively low number of burn-in samples is high, but decreases to a nearly constant value at $\sim 2,000$ burn-in samples (.02 $N_{\rm burn\,in}/N_{\rm samp}$). The default value for the number of burn-in samples is half the total number of samples, which adequately allows for the initialization of the sampler. Therefore, we opt to use 60,000 samples and 30,000 burn-in samples for each CARMA_pack fit.

A5 Fitting simulated DRW light curves

Here we test if simulated DRW light curves with the same sampling and S/N as our real data would produce a steep high-frequency slope in the CARMA PSD. First, we generate mock g-band DRW light curves for all quasars in our sample using the best-fit $\tau_{\rm DRW,obs}$ and $\sigma_{\rm DRW}$ in §3.1. These mock light curves are sampled at the same times and with the same S/N as the real light curves in our sample.

Next, we use CARMA_pack to fit these mock DRW light curves with a generalized CARMA model following the same procedures described in Appendix A3. We then extract a PSD from each mock light curve from the best-fit CARMA model, and construct an ensemble PSD. The results are shown in Fig. A4, where we compare the PSDs from the expected DRW model and recovered by CARMA_pack. We find that CARMA_pack successfully recovers a DRW PSD for these simulated light curves, as expected. This test confirms that the steep high-frequency-end PSD slope seen in real data is not due to effects of light curve cadence and S/N or the use of a more flexible CARMA model to fit the light curves.

We also use these simulated DRW light curves to investigate different choices of the best-fit parameters in Celerite or CARMA_pack.

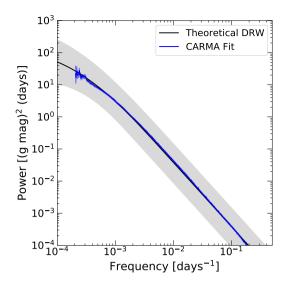


Figure A4. The ensemble PSD from mock DRW light curves generated using the best-fit DRW parameters from the real quasar light curves in §3.1. The expected DRW ensemble (median) PSD is shown in the black solid line with 1σ uncertainties highlighted in gray. The blue line shows the ensemble (median) PSD from generalized CARMA fits with CARMA_pack. The generalized CARMA fits correctly recover the DRW PSD, with no evidence for slope steepening at the highest frequencies sampled here.

In this work, we opt to use the median of the posterior distribution of samples as the fiducial best-fit parameters for all DRW and general CARMA model fits. Other works may use different choices for their best-fit parameters (such as the maximum-a-posteriori (MAP) (MacLeod et al. 2010) or the expectation value of the marginalized posterior (Suberlak et al. 2021)). Here, we discuss the differences in these choices of the best-fit parameters.

When modeling our quasar light curves with Celerite or CARMA_pack, we are given a number of samples for each parameter, output by a certain MCMC algorithm. The posterior probability distribution is simply the normalized distribution of the output parameters themselves. Using the median of the posterior is less susceptible to large fluctuations in the probability due to insufficient sampling of the distribution. The MAP, however, can prove to be unreliable, as it can be easily influenced by these fluctuations. The marginalized posterior utilizes the joint-posterior distribution of multiple parameters, giving a more robust look into the relationships between parameters, and taking that into account to choose the best possible value. The expectation value of this distribution (as opposed to the MAP or median) can aid if the posterior distribution has multiple peaks.

We compare different choices, including: median posterior, MAP, and expectation value. Both the MAP and the expectation value of a parameter's distribution are obtained by using the marginalized distribution of each parameter. This is done through the use of the likelihoods output from the Celerite fitting, for each quasar. One of the functions implemented in Celerite allows one to obtain a likelihood for a given set of parameters and data, given the model fit to a certain set of data. Therefore, for each sample from a given quasar light curve fit, we can construct a grid in parameter space, performing this likelihood calculation for an arbitrary number of points to obtain an n-dimensional posterior distribution, where n is the number of parameters. In this case, Celerite fits for both DRW parameters and a noise term, making this posterior three-dimensional. We can then marginalize over this distribution for each of the parameters, and obtain a best-fit parameter for each light curve.

We compare these different choices of best-fit parameters in Fig. A5. We obtained these values from fitting our simulated DRW light curves with a DRW model with Celerite, as well as CARMA_pack (in the latter case, a DRW or CARMA(1,0) model is enforced). We find that all these choices perform similarly for both DRW parameters with a similar amount of scatter, when compared with the input DRW parameters used to construct the simulated light curves. Overall, $\sigma_{\rm DRW}$ is better recovered than $\tau_{\rm DRW,obs}$. For long input $\tau_{\rm DRW}$, the recovered $\tau_{\rm DRW}$ is generally biased low due to an insufficient baseline of the light curve (e.g., Kozłowski 2017). Interestingly, the median posterior from the Celerite fit produces the least overall bias in $\tau_{\rm DRW}$ for our sample, justifying our choice of this particular definition of best-fit DRW parameters in this work.

Fig. A6 shows the comparison of the three different choices of the best-fit DRW parameters, using Celerite for the same simulated DRW light curves described above. While there are correlations among these different choices, there are also systematic offsets among them. For this study, we have chosen the median posterior as our fiducial best-fit parameters, given its performance in recovering the input DRW parameters as demonstrated in Fig. A5.

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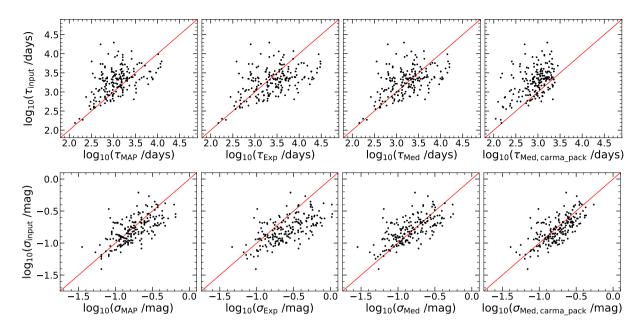


Figure A5. Comparison of the recovered and input DRW parameters ($\tau_{DRW,obs}$, σ_{DRW}) from our test using simulated DRW light curves. The best-fit value for the recovered parameter was obtained via three different methods: (1) the MAP from Celerite fitting samples, (2) the expectation value of the marginalized posterior using Celerite fitting samples, (3) the median value of the Celerite fitting samples, and (4) the median value of the CARMA_pack DRW fit samples. For each panel, the unity relation is shown as a red line.

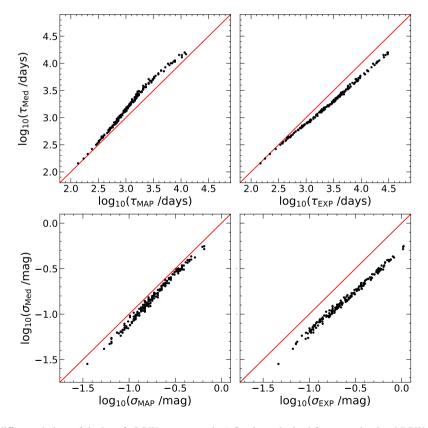


Figure A6. Comparison of different choices of the best-fit DRW parameters in Celerite, obtained from our simulated DRW light curves. These $\tau_{\text{DRW},\text{obs}}$ and σ_{DRW} values are determined using the same sets of posterior samples. The method used to obtain these MAP values and expectation values uses the marginalized posterior distribution of the samples. The red line in each panel indicates the unity relation.

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