

Categorical: Repeated measures

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1 Goals

1.1 Goals

1.1.1 Goals of this section

- Models for repeated measures
 - Mixed models framework
- First continuous outcomes

- Then categorical outcomes
- Two types of mixed model
 - Marginal mixed model
 - Conditional mixed model

1.1.2 Goals of this lecture

- What's the problem with repeated measures?
- Mixed model for continuous outcomes
 - Marginal model: **R** matrix
 - Conditional model: **G** matrix

2 Repeated measures

2.1 Repeated measures

2.1.1 Assumptions of GLM

- GLM (ANOVA / linear regression) assumes
 - Conditional normality of residuals
 - Constant variance of residuals
 - **Independence** of residuals

2.1.2 Assumptions of GLiM

- GLiM relaxes the assumptions of
 - Conditional normality of residuals
 - Constant variance of residuals
- But GLiM still assumes **independence** of residuals

2.1.3 Independence

- **Independent observations:** Information about one observation doesn't provide any information about other observations
 - *Lack of independence implies correlation* (but the reverse is not true)
 - * If observations are **not independent**, they will be correlated
 - * If observations are **correlated**, they are not independent
 - * If observations are **not correlated**, we don't know if they're independent or not

2.1.4 Repeated measures = non-independence

- Repeated measures from the same person are **not independent**
 - An observation from a person provides information about other observations from that person
 - Observations from the same person are more like one another than observations from different people
 - Observations from the same person are **correlated**

2.1.5 Violation of independence

- Does not impact regression coefficients
 - Impacts **standard errors**
 - Impacts **statistical significance**
- Will it be too *large* or too *small*? **It depends**
 - Hu, Goldberg, Hedeker, Flay, Pentz (1998)
 - Predictors about person: Standard errors are **too small**
 - Predictors about occasion: Standard errors are **too large**
 - Also depends on other things

2.1.6 Individual effects

- In models so far, there is **an** effect of a predictor
 - **Individual differences** in terms of *variables*
- But what if the effect of a predictor varied depending on the person?
 - **Individual differences** in terms of *effects or slopes*
- Mixed models can estimate person-specific effects

2.2 Example

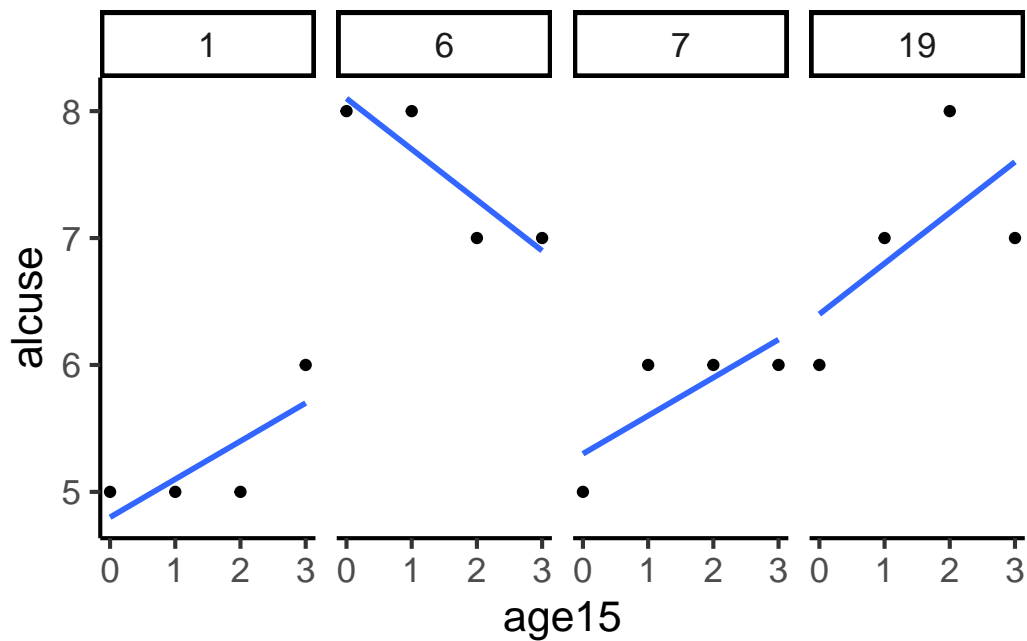
2.2.1 Example data: Substance use in adolescence

- **id**: ID variable
- **age**: Age in years
- **alcuse**: Alcohol use
- **ciguse**: Cigarette use
- **potuse**: Marijuana use
- **Demographics**: gender, family structure, other variables

2.2.2 Tall or univariate data

id	female	twopars	peerenc	parconf	paruse	age	alcuse	ciguse	potuse
1	0	1	7	11	13	15	5	6	5
1	0	1	7	11	13	16	5	6	5
1	0	1	7	11	13	17	5	5	4
1	0	1	7	11	13	18	6	7	5
6	1	0	6	14	16	15	8	9	6
6	1	0	6	14	16	16	8	8	5

2.2.3 Plot: Individual effects



2.2.4 Two types of mixed model

- Marginal models
 - Population-averaged models or generalized estimating equations (GEE)
 - Treat the person as a *nuisance* and adjust standard errors
- Conditional models
 - Generalized linear mixed models (GLMM)
 - Explicitly model the person (and variability among people) to get person-specific effects

2.3 Repeated-measures ANOVA

2.3.1 Review: Repeated-measures ANOVA assumptions

- Covariance matrix of outcomes

$$\mathbf{S}_{YY} = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ & \sigma_2^2 & \sigma_{23} & \sigma_{24} \\ & & \sigma_3^2 & \sigma_{34} \\ & & & \sigma_4^2 \end{bmatrix}$$

- σ_1^2 = variance of outcome at time 1
- σ_{12} = covariance between outcome at time 1 and outcome at time 2

2.3.2 Review: Repeated-measures ANOVA assumptions

- **Compound symmetry** of the covariance matrix of outcomes
 - Homogeneity of variances (i.e., variances are all the same):
 - * $\sigma_1^2 = \sigma_2^2 = \sigma_3^2 = \sigma_4^2$
 - Homogeneity of covariances (i.e., covariances are all the same):
 - * $\sigma_{12} = \sigma_{13} = \sigma_{14} = \sigma_{23} = \sigma_{24} = \sigma_{34}$
- Actual assumption: **Sphericity**
 - Compound symmetry holds for **differences** between pairs of scores
 - Slightly weaker assumption

3 Linear mixed model

3.1 Linear mixed model

3.1.1 Linear mixed model

- Also known as: random coefficient model, multi-level model, nested model, hierarchical linear model, random effects model
- Many names because they were developed in parallel in different disciplines
 - Multi-level models and hierarchical linear models from education
 - Random coefficient from statistics

3.1.2 Linear mixed model

- Extension of GLM that allows for non-independence
 - Partitions variation, just like ANOVA, linear regression
 - More control over the form of non-independence
 - * Linear regression: Independence
 - * Repeated-measures ANOVA: Compound symmetry
- Two approaches
 - Correlated residuals: \mathbf{R} matrix
 - Random effects: \mathbf{G} matrix

3.1.3 Linear mixed model: Equations

The linear mixed model:

$$Y = \mathbf{X}\beta + \mathbf{Z}\gamma + \epsilon$$

- where $\gamma \sim N(0, \mathbf{G})$ and $\epsilon \sim N(0, \mathbf{R})$
- $\text{variance}(Y) = V = \mathbf{ZGZ}' + \mathbf{R}$

3.1.4 Linear mixed model: Simpler

$$Y = \mathbf{X}\beta + \mathbf{Z}\gamma + \epsilon$$

- $\mathbf{X}\beta$ are the **fixed effects**
 - Average effects
 - Think predictors (\mathbf{X}) and regression coefficients (β)
- $\mathbf{Z}\gamma$ and ϵ are *special residuals* that let us include correlations among the repeated observations
 - Specifically, among their **residuals**

3.1.5 Two approaches

- $\mathbf{Z}\gamma$ are the **random effects**
 - γ has mean = 0 and variance given by *covariance matrix* \mathbf{G}
 - Generalized linear mixed models (GLMM) use this
- ϵ is the **error** or **residual term**
 - ϵ has mean = 0 and variance given by *covariance matrix* \mathbf{R}
 - Generalized estimating equations (GEE) and related methods use this

3.1.6 Continuous vs categorical outcomes

- **Continuous outcomes**
 - \mathbf{G} and \mathbf{R} portions of the model are *independent*
- **Categorical outcomes**
 - \mathbf{G} and \mathbf{R} portions of the model are **NOT independent**
- This week: \mathbf{G} and \mathbf{R} for the continuous model
 - Understand how they function
- Next week: Categorical outcomes

3.2 R matrix

3.2.1 R matrix

- **R** is the covariance matrix among the repeated outcomes / timepoints

$$\mathbf{R} = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ & \sigma_2^2 & \sigma_{23} & \sigma_{24} \\ & & \sigma_3^2 & \sigma_{34} \\ & & & \sigma_4^2 \end{bmatrix}$$

3.2.2 How are timepoints related?

- Linear mixed model is really flexible about what this matrix can look like
 - Unstructured
 - Compound symmetry
 - Autoregressive
 - Diagonal
 - Many others: See Kinkaid (2005)
- Repeated-measures ANOVA is a special case of LMM
 - Not flexible: Compound symmetry only

3.2.3 Unstructured R

$$\mathbf{R} = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \sigma_{13} & \sigma_{14} \\ & \sigma_2^2 & \sigma_{23} & \sigma_{24} \\ & & \sigma_3^2 & \sigma_{34} \\ & & & \sigma_4^2 \end{bmatrix}$$

- Estimate every value in the matrix
- $\frac{t(t+1)}{2}$ values: Here, 10 values

3.2.4 Compound symmetry R

$$\mathbf{R} = \begin{bmatrix} \sigma^2 + \sigma_1^2 & \sigma_1^2 & \sigma_1^2 & \sigma_1^2 \\ & \sigma^2 + \sigma_1^2 & \sigma_1^2 & \sigma_1^2 \\ & & \sigma^2 + \sigma_1^2 & \sigma_1^2 \\ & & & \sigma^2 + \sigma_1^2 \end{bmatrix}$$

- One value for all variances

- One value for all covariances
- 2 values: Here, 2 values

3.2.5 Autoregressive R

$$\mathbf{R} = \begin{bmatrix} \sigma^2 & \sigma^2\rho & \sigma^2\rho^2 & \sigma^2\rho^3 \\ & \sigma^2 & \sigma^2\rho & \sigma^2\rho^2 \\ & & \sigma^2 & \sigma^2\rho \\ & & & \sigma^2 \end{bmatrix}$$

- One value for variances
- Covariance decreases as time between points increases (ρ)
- 2 values: Here, 2 values

3.2.6 Diagonal (independence) R

$$\mathbf{R} = \begin{bmatrix} \sigma_1^2 & 0 & 0 & 0 \\ & \sigma_2^2 & 0 & 0 \\ & & \sigma_3^2 & 0 \\ & & & \sigma_4^2 \end{bmatrix}$$

- One value for variance for each time point
- t values: Here, 4 values

3.2.7 Why do we care about R?

- **R** is the residual variance matrix
 - Residual variance impacts the **standard errors of the fixed effects**
 - **R** impacts the standard error (and therefore the significance) of the fixed effects
- Variance structure you choose affects what is significant
 - Choose the variance structure that most closely reflects reality

3.2.8 Which form of R to use?

- Run models with different versions of **R** matrix
 - Compare using AIC and likelihood
 - * AIC: smaller is better
 - * Likelihood: likelihood ratio test

3.2.9 Which form of R to use?

- Unstructured: Most information, but also most parameters
 - Difficult with more than a handful of timepoints
 - First try to get an idea of what the covariance matrix looks like
- Diagonal: Independence
 - All timepoints are uncorrelated
 - Unlikely given our discussion of repeated measures
- Compound symmetry: In between
- Autoregressive: In between

3.2.10 Example: Unstructured

Generalized least squares fit by maximum likelihood

Model: alcuse ~ 1 + age15

Data: alcuse_tall

	AIC	BIC	logLik
	981.0951	1017.108	-481.5476

Correlation Structure: General

Formula: ~1 | id

Parameter estimate(s):

Correlation:

	1	2	3	
1	1			
2	0.580	1		
3	0.421	0.605	1	
4	0.324	0.366	0.340	1

Coefficients:

	Value	Std.Error	t-value	p-value
(Intercept)	5.901857	0.08824313	66.88177	0.0000
age15	0.105372	0.03342207	3.15278	0.0017

Correlation:

(Intr)

age15 -0.633

Standardized residuals:

	Min	Q1	Med	Q3	Max
	-3.418749537	-0.239414053	-0.007940977	0.858943288	2.304510243

Residual standard error: 0.9104505
 Degrees of freedom: 404 total; 402 residual

3.2.11 Example: Unstructured

- Working correlation matrix (**R**)

	[,1]	[,2]	[,3]	[,4]
[1,]	1.0000000	0.5797687	0.4208941	0.3236446
[2,]	0.5797687	1.0000000	0.6053009	0.3663275
[3,]	0.4208941	0.6053009	1.0000000	0.3404776
[4,]	0.3236446	0.3663275	0.3404776	1.0000000

3.2.12 Example: Compound symmetry

Generalized least squares fit by maximum likelihood

Model: alcuse ~ 1 + age15

Data: alcuse_tall

AIC	BIC	logLik
988.1413	1004.147	-490.0706

Correlation Structure: Compound symmetry

Formula: ~1 | id

Parameter estimate(s):

Rho
 0.4430272

Coefficients:

	Value	Std.Error	t-value	p-value
(Intercept)	5.890099	0.08302555	70.94321	0.0000
age15	0.107921	0.03036304	3.55435	0.0004

Correlation:

(Intr)
 age15 -0.549

Standardized residuals:

Min	Q1	Med	Q3	Max
-3.405625651	-0.234496379	0.002171263	0.861991319	2.313480479

Residual standard error: 0.9120029
Degrees of freedom: 404 total; 402 residual

3.2.13 Example: Compound symmetry

- Working correlation matrix (**R**)

```
      [,1]      [,2]      [,3]      [,4]  
[1,] 1.0000000 0.4430272 0.4430272 0.4430272  
[2,] 0.4430272 1.0000000 0.4430272 0.4430272  
[3,] 0.4430272 0.4430272 1.0000000 0.4430272  
[4,] 0.4430272 0.4430272 0.4430272 1.0000000
```

3.2.14 Compare

Model	AIC	-2LL	# parameters
Unstructured	981.095	963.095	10
Compound symmetry	988.141	980.141	2

- Difference in -2LL: $980.141 - 963.095 = 17.046$
- Degrees of freedom: $10 - 2 = 8$
- Critical $\chi(8)^2 = 15.507$
- Test is significant: More complex model fits better
 - Unstructured

3.2.15 Interpretation

- This is a **marginal** model
 - Treat the person as a *nuisance* and adjust standard errors
- Use **R** to account for additional correlation of repeated measures
 - But don't care about it
 - Just want to get rid of it

3.3 G matrix

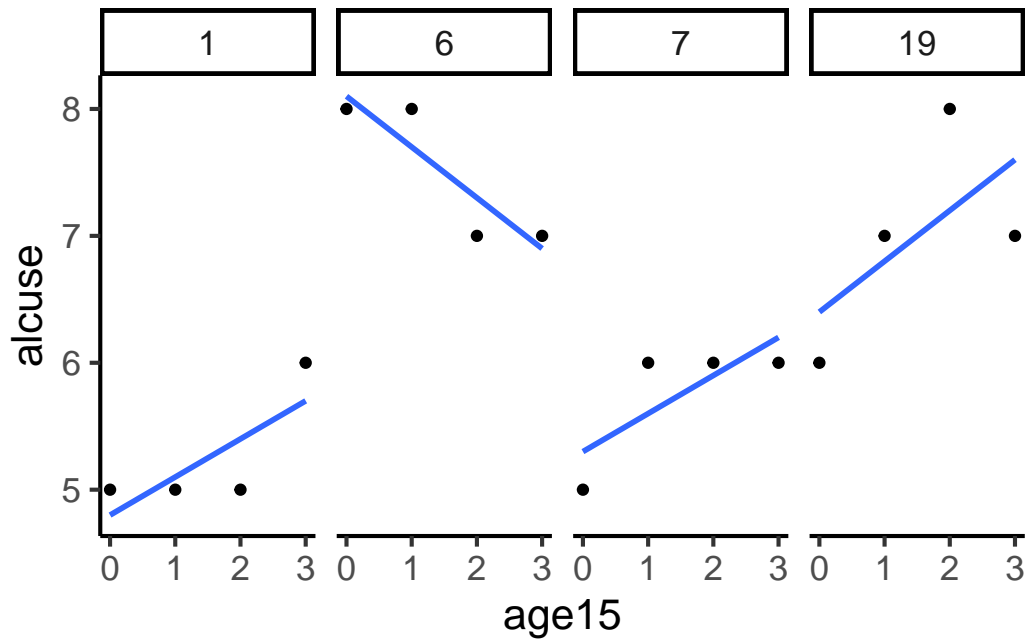
3.3.1 Individual effects

- Before:
 - There is **an** effect of a predictor
 - * **Individual differences** in terms of *variables*
- Now:
 - Effect of a predictor varies depending on the person
 - * **Individual differences** in terms of *effects or slopes*

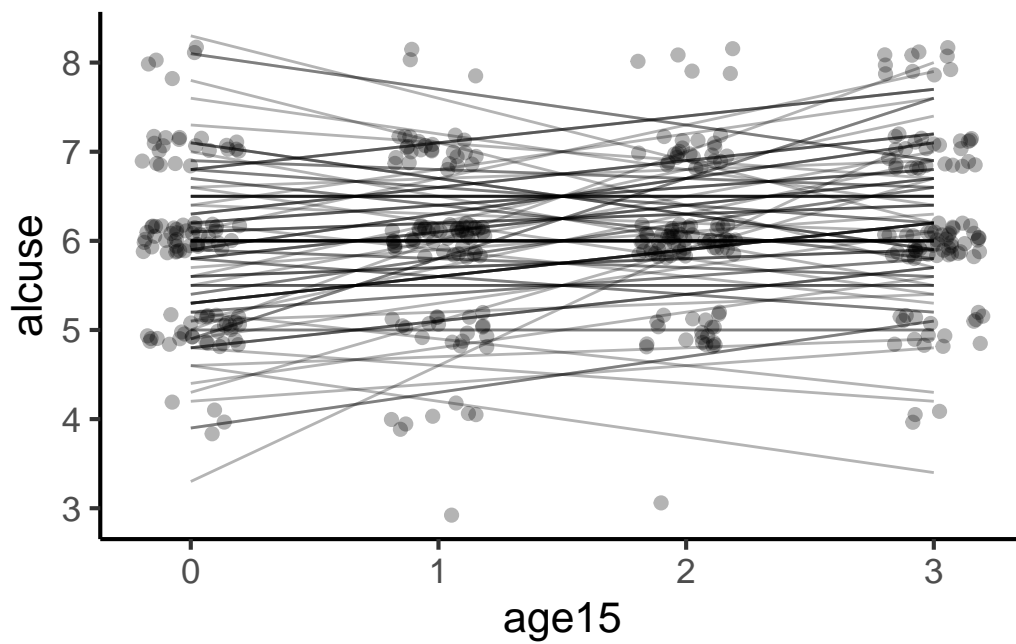
3.3.2 Conceptually

- Perform a regression on each person's data
 - Predictor: **time**
 - Outcome: Outcome
- Separate regression for **each person** in the study
 - Each person has an intercept
 - Each person has a slope

3.3.3 Time vs outcome for the first four participants



3.3.4 Time vs outcome for all participants



3.3.5 Assumptions

- Figures are a good way to think about the model
 - But still violate assumptions of linear regression
- Observations for each line are **not independent**
 - Multiple observations from the same person
- Good news!
 - Lack of independence only impacts **standard errors**
 - Estimates of the **intercepts** and **slopes** are good

3.3.6 That's a lot of lines!

- Intercept and slope for every single person in the sample
 - Can't report all of those
 - Summarize the intercepts and slopes in some way
- Variances and covariances
 - Variance of intercepts
 - Variance of slopes
 - Covariances between intercepts and slopes

3.3.7 G

- **G** is the variance-covariance matrix of intercepts and slopes

$$\mathbf{G} = \begin{bmatrix} \sigma_{int}^2 & \sigma_{int-slope} \\ & \sigma_{slope}^2 \end{bmatrix}$$

- σ_{int}^2 is the variance of the intercepts
- σ_{slope}^2 is the variance of the slopes
- $\sigma_{int-slope}$ is the covariance between intercepts and slopes

3.3.8 Why do we care about G?

$$Y = \mathbf{X}\beta + \mathbf{Z}\gamma + \epsilon$$

- $\mathbf{Z}\gamma$ is where the random effects (**G**) are
 - It accounts for some variation in scores
- Remaining variation ends up in ϵ , the residual variance
 - Residual variance impacts the **standard errors of the fixed effects**
 - **G** impacts s.e.s (and therefore significance) of fixed effects
- Random effects affect what is significant

3.4 Example: Random intercept and slope

Linear mixed model fit by maximum likelihood . t-tests use Satterthwaite's method [lmerModLmerTest]

Formula: alcuse ~ 1 + age15 + (1 + age15 | id)

Data: alcuse_tall

AIC	BIC	logLik	deviance	df.resid
985.7	1009.7	-486.8	973.7	398

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.71801	-0.60424	-0.07987	0.54069	2.97288

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
id	(Intercept)	0.55174	0.7428	
	age15	0.03706	0.1925	-0.59
	Residual	0.40148	0.6336	

Number of obs: 404, groups: id, 101

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	5.89010	0.09080	100.99824	64.866	< 0.0000000000000002 ***
age15	0.10792	0.03409	100.99947	3.166	0.00204 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:


```

      (Intr)
age15 -0.653

```

3.4.1 Example: Random intercept and slope

- Random effects (and residual variance)

```

Groups   Name             Std.Dev. Corr
id       (Intercept) 0.74279
          age15      0.19252 -0.586
Residual                      0.63363

```

3.4.2 Interpretation

- This is a **conditional** model
 - Individual variability in intercept and slope are of interest
- Explicitly model this using **G**
 - Person-specific intercepts and slopes

3.5 Comparing both sets of models

3.5.1 Compare fixed effects

Model	Intercept	Slope
R unstructured	5.902	0.105
G with random intercept and slope	5.89	0.108

- **R** unstructured: Correlated residuals
- **G** with random intercept and slope: Random effects

3.5.2 Continuous models vs categorical models

- Continuous outcome (here)
 - Interpretations are different
 - But numerical results are basically identical
 - * Assuming some things

- Categorical outcome
 - Interpretations are different
 - Numerical results are **very** different
 - * Nonlinearity for categorical outcomes

4 Summary

4.1 Summary

4.1.1 Summary of this week

- Repeated measures violate the assumption of independence
- Mixed model (/ multilevel model / hierarchical linear model)
 - Marginal and conditional models
 - * **R** and **G** matrix approaches
 - * Equivalent (numerically) for continuous outcomes

4.1.2 Next weeks

- Extend mixed models to categorical outcomes
 - Marginal: **R** matrix, population averaged, GEE, cluster robust
 - Conditional: **G** matrix, generalized linear mixed models (GLMM)
- Different interpretations, different numbers