

Categorical: Contingency tables

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1 Goals

1.1 Goals

1.1.1 Goals of this lecture

- Extend 2×2 contingency tables to larger tables
 - More variable categories: 2×3 and larger
 - More variables: $2 \times 2 \times 2$ tables
- Chi-square tests for these tables

- Probing the tables
- Residuals

2 More variable categories

2.1 $I \times 2$ and $2 \times J$ tables

2.1.1 2×2 tables and beyond...

- We've only looked at 2×2 tables
 - Extend these tables in terms of rows or columns
 - * Just more rows: $I \times 2$
 - * Just more columns: $2 \times J$
 - * More rows and more columns: $I \times J$

2.1.2 Whickam2 data with a twist

| | Outcome | Smoker | Age | AgeGroup | Alive |
|----|---------|--------|-----|----------|-------|
| 1 | Alive | Yes | 23 | 18-64 | 1 |
| 2 | Alive | Yes | 18 | 18-64 | 1 |
| 3 | Dead | Yes | 71 | 65+ | 0 |
| 4 | Alive | No | 67 | 65+ | 1 |
| 5 | Alive | No | 64 | 18-64 | 1 |
| 6 | Alive | Yes | 38 | 18-64 | 1 |
| 7 | Alive | Yes | 45 | 18-64 | 1 |
| 8 | Dead | No | 76 | 65+ | 0 |
| 9 | Alive | No | 28 | 18-64 | 1 |
| 10 | Alive | No | 27 | 18-64 | 1 |

- Split the Age variable into **3 groups**
 - 18 to 40
 - 41 to 64
 - 65+

2.1.3 Convert Age into 3 categories

```
Whickham2 <- Whickham2 %>%
  mutate(AgeGroup3 = ifelse(Age %in% 18:40, 1,
    ifelse(Age %in% 41:64, 2, 3)))
head(Whickham2)
```

| | Outcome | Smoker | Age | AgeGroup | Alive | AgeGroup3 |
|---|---------|--------|-----|----------|-------|-----------|
| 1 | Alive | Yes | 23 | 18-64 | 1 | 1 |
| 2 | Alive | Yes | 18 | 18-64 | 1 | 1 |
| 3 | Dead | Yes | 71 | 65+ | 0 | 3 |
| 4 | Alive | No | 67 | 65+ | 1 | 3 |
| 5 | Alive | No | 64 | 18-64 | 1 | 2 |
| 6 | Alive | Yes | 38 | 18-64 | 1 | 1 |

2.1.4 Agegroup3 versus Alive: Observed

| | Dead | Alive | Sum |
|----------|------|-------|------|
| 18 to 40 | 19 | 521 | 540 |
| 41 to 64 | 141 | 390 | 531 |
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

2.1.5 Agegroup3 versus Alive: Expected

| | Dead | Alive | Sum |
|----------|---------|---------|------|
| 18 to 40 | 151.644 | 388.356 | 540 |
| 41 to 64 | 149.116 | 381.884 | 531 |
| 65+ | 68.240 | 174.760 | 243 |
| Sum | 369.000 | 945.000 | 1314 |

2.1.6 Agegroup3 versus Alive: Chi-square

$$\chi^2 = \sum \left(\frac{(n_{ij} - \mu_{ij})^2}{\mu_{ij}} \right) = \sum \left(\frac{(O-E)^2}{E} \right) =$$

$$\frac{(19-151.644)^2}{151.644} + \frac{(141-149.116)^2}{149.116} + \frac{(209-68.24)^2}{68.24} + \frac{(521-388.356)^2}{388.356} + \frac{(390-381.884)^2}{381.884} + \frac{(34-174.76)^2}{174.76} =$$

$$116.024 + 0.442 + 290.351 + 45.305 + 0.173 + 113.375 = 565.669$$

2.1.7 Agegroup3 versus Alive: Chi-square

- Degrees of freedom = $(I - 1) \times (J - 1) = (3 - 1) \times (2 - 1) = 2$
 - $\chi^2_{critical}(2) = 5.99$
 - $565.669 > 5.99$
 - Reject H_0 that AgeGroup3 and Alive are independent
- But then what?
 - What is different?
 - Similar to ANOVA
 - * With 3 groups (levels), which one(s) are different from each other?

2.2 Partitioning χ^2

2.2.1 Partitioned chi-square

- Chi-square statistics can be split up (partitioned):
 - Alive by (18 to 40 vs 65+)
 - Alive by ((18 to 40 and 41 to 64) vs 65+)
 - These are two *independent* tests
- Chi-square for all *independent* tests add up to chi-square for complete table
 - Kind of
- Degrees of freedom also add up

2.2.2 Orthogonal tests

- Orthogonal partitioning of a contingency table is similar to coding orthogonal contrasts for ANOVA

| Orthogonal partitioning | | | |
|---|------|------|----|
| Alive by (18 to 40 vs 41 to 64) | +1 | -1 | 0 |
| Alive by ((18 to 40 and 41 to 64) vs 65+) | -0.5 | -0.5 | +1 |
| Not orthogonal partitioning | | | |
| Alive by (18 to 40 vs 65+) | +1 | 0 | -1 |
| Alive by (41 to 64 vs 65+) | 0 | +1 | -1 |

2.2.3 Agegroup3 versus Alive: Partition 1

- Observed:

| | Dead | Alive | Sum |
|----------|------|-------|------|
| 18 to 40 | 19 | 521 | 540 |
| 41 to 64 | 141 | 390 | 531 |
| Sum | 160 | 911 | 1071 |

- Expected:

| | Dead | Alive | Sum |
|----------|--------|--------|------|
| 18 to 40 | 80.67 | 459.33 | 540 |
| 41 to 64 | 79.33 | 451.67 | 531 |
| Sum | 160.00 | 911.00 | 1071 |

Pearson's Chi-squared test

```
data: Age3_Alive[c(1, 2), ]  
X-squared = 111.79, df = 1, p-value < 0.00000000000000022
```

2.2.4 Agegroup3 versus Alive: Partition 2

- Observed:

| | Dead | Alive | Sum |
|----------|------|-------|------|
| 18 to 64 | 160 | 911 | 1071 |
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

- Expected:

| | Dead | Alive | Sum |
|----------|--------|--------|------|
| 18 to 64 | 300.76 | 770.24 | 1071 |
| 65+ | 68.24 | 174.76 | 243 |
| Sum | 369.00 | 945.00 | 1314 |

Pearson's Chi-squared test

```
data: Age_Alive  
X-squared = 495.33, df = 1, p-value < 0.00000000000000022
```

2.2.5 Partitioned chi-square

- Overall: $\chi^2(2) = 565.67$
- Partition 1: $\chi^2(1) = 111.79$
- Partition 2: $\chi^2(1) = 495.33$
- $111.79 + 495.33 = 607.12 \approx 565.67$
 - For the χ^2 statistic, the sum will be approximate
 - * Closer for larger *samples* and larger *tables*
 - A slightly different statistic, G^2 , will always sum perfectly
 - * G^2 can also be partitioned in the same

2.2.6 G^2 statistic

- $G^2 = 2 \sum \left(n_{ij} \times \ln \left(\frac{n_{ij}}{\mu_{ij}} \right) \right)$
 - Also called “likelihood ratio test statistic”
 - Compare to χ^2 distribution with $(I - 1) \times (J - 1)$ df

2.2.7 G^2 statistic

- Overall 3×2 table

Log likelihood ratio (G-test) test of independence without correction

data: Age3_Alive

G = 584.41, X-squared df = 2, p-value < 0.000000000000000022

- Just 18 to 40 vs 41 to 64

Log likelihood ratio (G-test) test of independence without correction

data: Age3_Alive[c(1, 2),]

G = 124.02, X-squared df = 1, p-value < 0.000000000000000022

- Combined (18 to 40 and 41 to 64) vs 65+

Log likelihood ratio (G-test) test of independence without correction

data: Age_Alive

G = 460.39, X-squared df = 1, p-value < 0.000000000000000022

2.3 Residuals

2.3.1 Residuals

- Residuals exist for χ^2 just like linear regression
- **Raw residual** = observed - expected = $n_{ij} - \hat{\mu}_{ij}$
- **Standardized residual** divides by std error of raw residuals

$$- \frac{n_{ij} - \hat{\mu}_{ij}}{\sqrt{\hat{\mu}_{ij}(1 - p_{i+})(1 - p_{+j})}}$$

– where $\sqrt{\hat{\mu}_{ij}(1 - p_{i+})(1 - p_{+j})}$ is std error of raw residuals under H_0

– and $p_{i+} = n_{i+}/n$ and $p_{+j} = n_{+j}/n$

2.3.2 Observed and expected frequencies

- Observed

| | Dead | Alive |
|----------|------|-------|
| 18 to 40 | 19 | 521 |
| 41 to 64 | 141 | 390 |
| 65+ | 209 | 34 |

- Expected

| | Dead | Alive |
|----------|---------|---------|
| 18 to 40 | 151.644 | 388.356 |
| 41 to 64 | 149.116 | 381.884 |
| 65+ | 68.240 | 174.760 |

2.3.3 Residuals

- Raw residuals

| | Dead | Alive |
|----------|----------|----------|
| 18 to 40 | -132.644 | 132.644 |
| 41 to 64 | -8.116 | 8.116 |
| 65+ | 140.760 | -140.760 |

- Standardized residuals

| | Dead | Alive |
|----------|---------|---------|
| 18 to 40 | -16.549 | 16.549 |
| 41 to 64 | -1.015 | 1.015 |
| 65+ | 22.256 | -22.256 |

2.3.4 Standardized residuals

- Under H_0 , variables are **independent**
 - Observed cell frequencies = expected cell frequencies
 - * Residuals tell you how much each cell deviates from this
 - * Large standardized residual = cell shows lack of fit from H_0
- Standardized \approx normal distribution
 - Expect about 5% of residuals to be greater than ± 2
 - * Look at standardized residual greater than ± 2
 - * In small tables, this is way off

2.3.5 Standardized residuals

- Observed

| | Dead | Alive |
|----------|------|-------|
| 18 to 40 | 19 | 521 |
| 41 to 64 | 141 | 390 |
| 65+ | 209 | 34 |

- Standardized residuals

| | Dead | Alive |
|----------|---------|---------|
| 18 to 40 | -16.549 | 16.549 |
| 41 to 64 | -1.015 | 1.015 |
| 65+ | 22.256 | -22.256 |

3 More variables

3.1 Conditional and marginal effects

3.1.1 Adding a 3rd variable

- **Control for** a potentially confounding third variable, like **smoking**

- Relationship between **AgeGroup** (X) and **Alive** (Y)
 - * What if **smokers** have one relationship between X and Y
 - * But **non-smokers** have a different relationship between X and Y ?
- From last time: **Smoker** and **Alive** had an unexpected pattern
 - * Smokers were **less likely to die** than non-smokers?
 - * What if smokers are younger than non-smokers and that's what's really going on?

3.1.2 Adding a third variable: $Z = \text{Smoker}$

| AgeGroup | Alive | Smoker | Freq |
|----------|-------|--------|------|
| 18 to 64 | Dead | No | 65 |
| 65+ | Dead | No | 165 |
| 18 to 64 | Alive | No | 474 |
| 65+ | Alive | No | 28 |
| 18 to 64 | Dead | Yes | 95 |
| 65+ | Dead | Yes | 44 |
| 18 to 64 | Alive | Yes | 437 |
| 65+ | Alive | Yes | 6 |

3.1.3 3 variables = 3-way = 3D

- 2 variables = 2-way or 2D table
 - 3 variables = 3-way or 3D table
- Two ways to look at a 3D table
 - **Partial table** (a.k.a. conditional table)
 - **Marginal table**

3.1.4 Partial tables

- Slice 3D table into more 2D tables
 - 2-way table of X vs Y for each level of Z
- *Conditional* on levels of Z
 - Remove effect of Z by holding it constant at specific levels
- *Conditional associations*
 - e.g., conditional χ^2

, , Smoker = No

| | Alive | | |
|----------|-------|-------|-----|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 65 | 474 | 539 |
| 65+ | 165 | 28 | 193 |
| Sum | 230 | 502 | 732 |

, , Smoker = Yes

| | Alive | | |
|----------|-------|-------|-----|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 95 | 437 | 532 |
| 65+ | 44 | 6 | 50 |
| Sum | 139 | 443 | 582 |

, , Smoker = Sum

| | Alive | | |
|----------|-------|-------|------|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 160 | 911 | 1071 |
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

3.1.5 Marginal table

- 2D table ignoring Z
 - 2-way table of X vs Y
- Collapse across levels of Z
 - Add up across
 - No information about Z
- *Marginal associations*

| | Alive | | |
|----------|-------|-------|------|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 160 | 911 | 1071 |
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

3.1.6 Alive vs AgeGroup: Conditional on Smoker

- Frequencies

```
, , Smoker = No
```

| | Alive | | |
|----------|-------|-------|-----|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 65 | 474 | 539 |
| 65+ | 165 | 28 | 193 |
| Sum | 230 | 502 | 732 |

```
, , Smoker = Yes
```

| | Alive | | |
|----------|-------|-------|-----|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 95 | 437 | 532 |
| 65+ | 44 | 6 | 50 |
| Sum | 139 | 443 | 582 |

```
, , Smoker = Sum
```

| | Alive | | |
|----------|-------|-------|------|
| AgeGroup | Dead | Alive | Sum |
| 18 to 64 | 160 | 911 | 1071 |
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

- Proportions

```
# A tibble: 4 x 6
# Groups:   AgeGroup, Smoker [4]
  AgeGroup Smoker Alive count prop totaln
  <fct>    <fct>   <int> <int> <dbl>   <int>
1 18-64    No         1    474 0.879    539
2 18-64    Yes         1    437 0.821    532
3 65+      No         1     28 0.145    193
4 65+      Yes         1      6 0.12     50
```

3.1.7 Alive vs Smoker: Conditional on AgeGroup

- Frequencies

```
, , AgeGroup = 18-64
```

| | | Alive | | |
|--------|------|-------|------|--|
| Smoker | Dead | Alive | Sum | |
| No | 65 | 474 | 539 | |
| Yes | 95 | 437 | 532 | |
| Sum | 160 | 911 | 1071 | |

```
, , AgeGroup = 65+
```

| | | Alive | | |
|--------|------|-------|-----|--|
| Smoker | Dead | Alive | Sum | |
| No | 165 | 28 | 193 | |
| Yes | 44 | 6 | 50 | |
| Sum | 209 | 34 | 243 | |

```
, , AgeGroup = Sum
```

| | | Alive | | |
|--------|------|-------|------|--|
| Smoker | Dead | Alive | Sum | |
| No | 230 | 502 | 732 | |
| Yes | 139 | 443 | 582 | |
| Sum | 369 | 945 | 1314 | |

- Proportions

```
# A tibble: 4 x 5
# Groups:   Smoker, AgeGroup [4]
  Smoker AgeGroup Alive count prop
<fct>   <fct>     <int> <int> <dbl>
1 No     18-64         1   474 0.879
2 No     65+         1    28 0.145
3 Yes    18-64         1   437 0.821
4 Yes    65+         1     6 0.12
```

3.1.8 Alive vs AgeGroup: Ignoring (marginal on) Smoker

- Frequencies

| | | Alive | | |
|----------|------|-------|------|--|
| AgeGroup | Dead | Alive | Sum | |
| 18 to 64 | 160 | 911 | 1071 | |

| | | | |
|-----|-----|-----|------|
| 65+ | 209 | 34 | 243 |
| Sum | 369 | 945 | 1314 |

- Proportions

```
# A tibble: 2 x 4
# Groups:   AgeGroup [2]
  AgeGroup Alive count prop
  <fct>     <int> <int> <dbl>
1 18-64      1   911 0.851
2 65+        1    34 0.140
```

3.1.9 Alive vs Smoker: Ignoring (marginal on) AgeGroup

- Frequencies

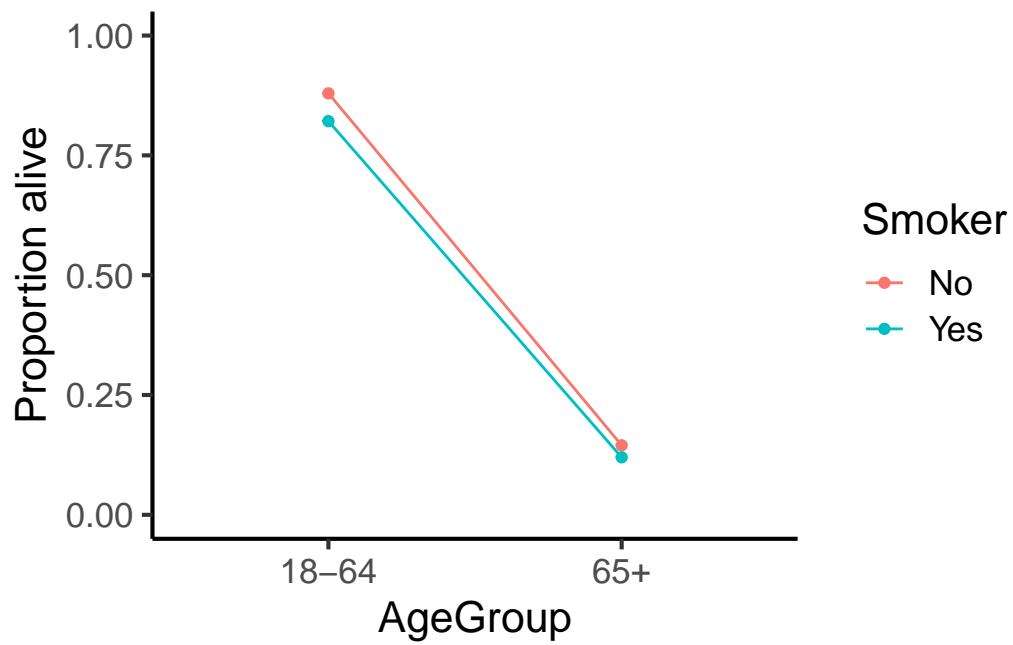
| | | | |
|--------|-------|-------|------|
| | Alive | | |
| Smoker | Dead | Alive | Sum |
| No | 230 | 502 | 732 |
| Yes | 139 | 443 | 582 |
| Sum | 369 | 945 | 1314 |

- Proportions

```
# A tibble: 2 x 5
# Groups:   Smoker [2]
  Smoker Alive count prop totaln
  <fct>   <int> <int> <dbl> <int>
1 No      1   502 0.686   732
2 Yes     1   443 0.761   582
```

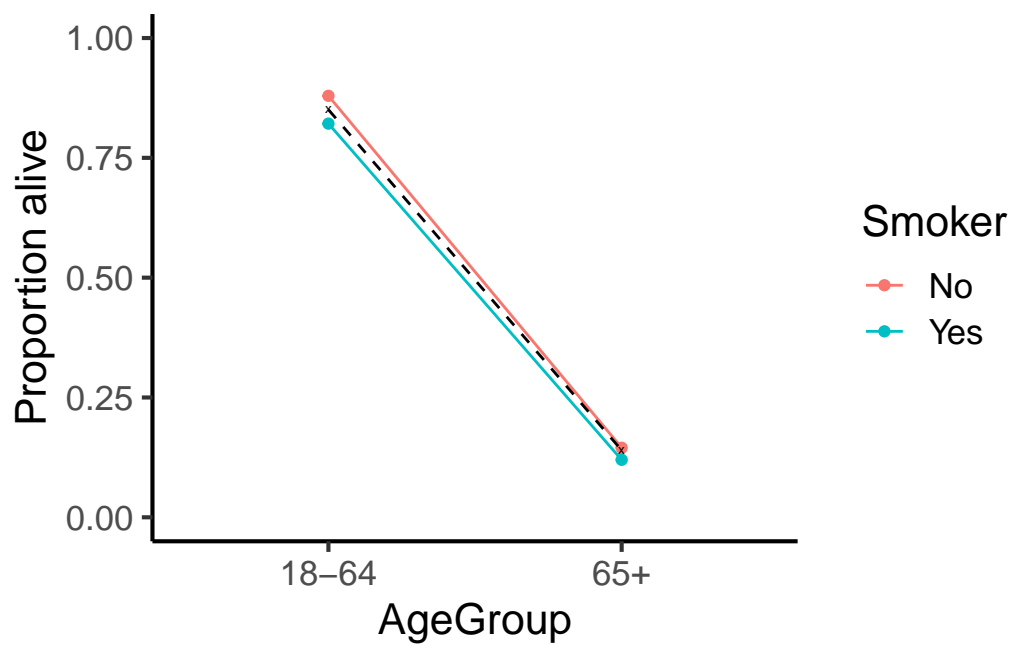
3.1.10 Plot: Proportion Alive by AgeGroup, by Smoker

- Conditional effect of AgeGroup, conditional on Smoker...



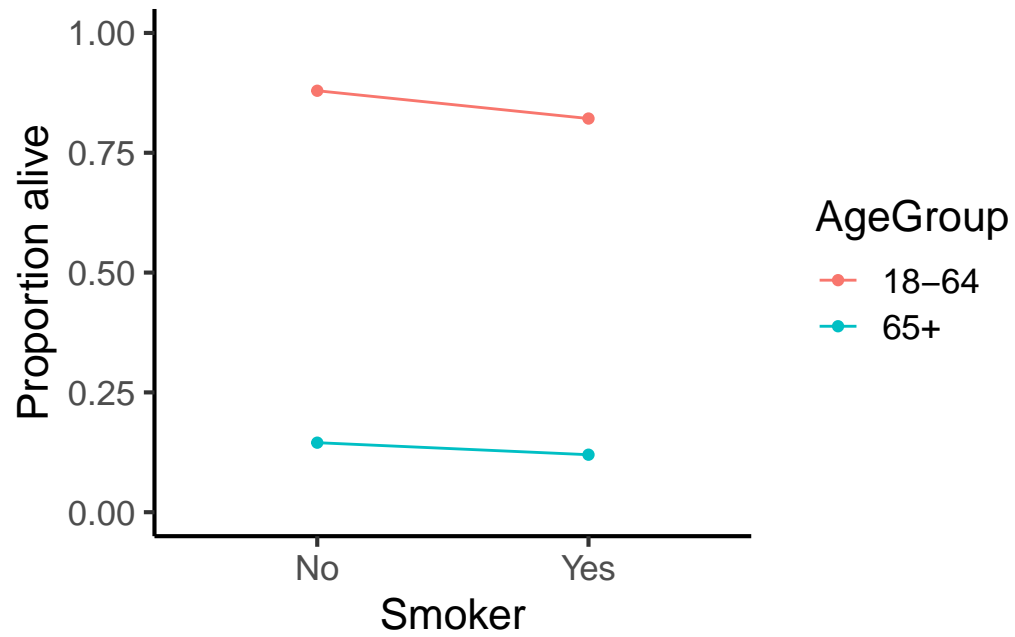
3.1.11 Plot: Proportion Alive by AgeGroup, by Smoker

- ...plus **marginal effect**, ignoring Smoker



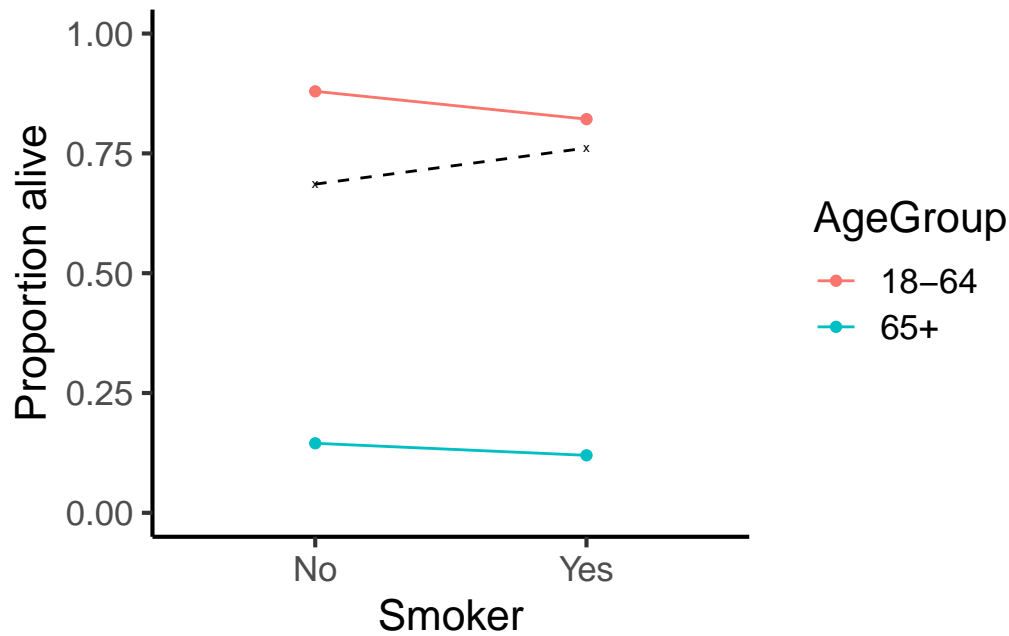
3.1.12 Plot: Proportion Alive by Smoker, by AgeGroup

- Conditional effect of Smoker, conditional on AgeGroup...



3.1.13 Plot: Proportion Alive by Smoker, by AgeGroup

- ...plus marginal effect, ignoring AgeGroup



3.2 Simpson's paradox

3.2.1 Conditional vs marginal effects

- **Conditional effects** take Z into account
 - Smokers **less likely to be alive** than non-smokers among *young people*
 - Smokers **less likely to be alive** than non-smokers among *old people*
- **Marginal effects** ignore Z
 - Smokers **more likely to be alive** than non-smokers

3.2.2 Simpson's paradox

- Opposite direction *marginal* vs *conditional* effects
- Not unique to contingency tables
 - Can happen any time you *ignore a confounder*
 - Change in direction of effect after adding covariate
- Related to
 - Lord's paradox: ANCOVA versus difference scores
 - Suppression effects: Change relationship by adding covariate
 - Ecological fallacy: Opposite direction effects at higher vs lower level

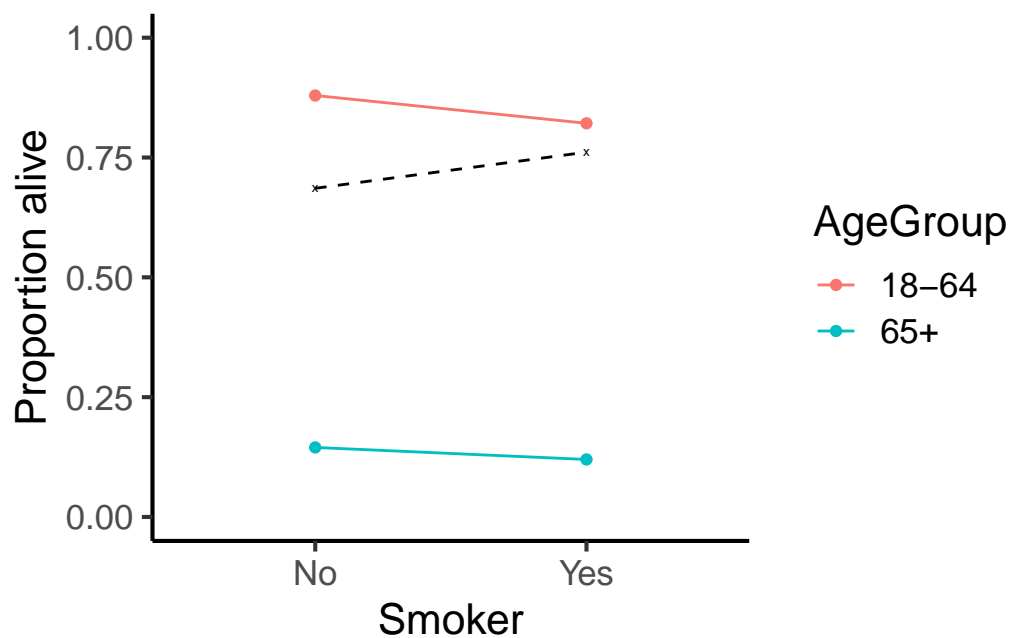
3.2.3 Why does it happen?

- Relationship between the confounder and other variables besides Y
 - In this case, the relationship between **AgeGroup** and **Smoker**

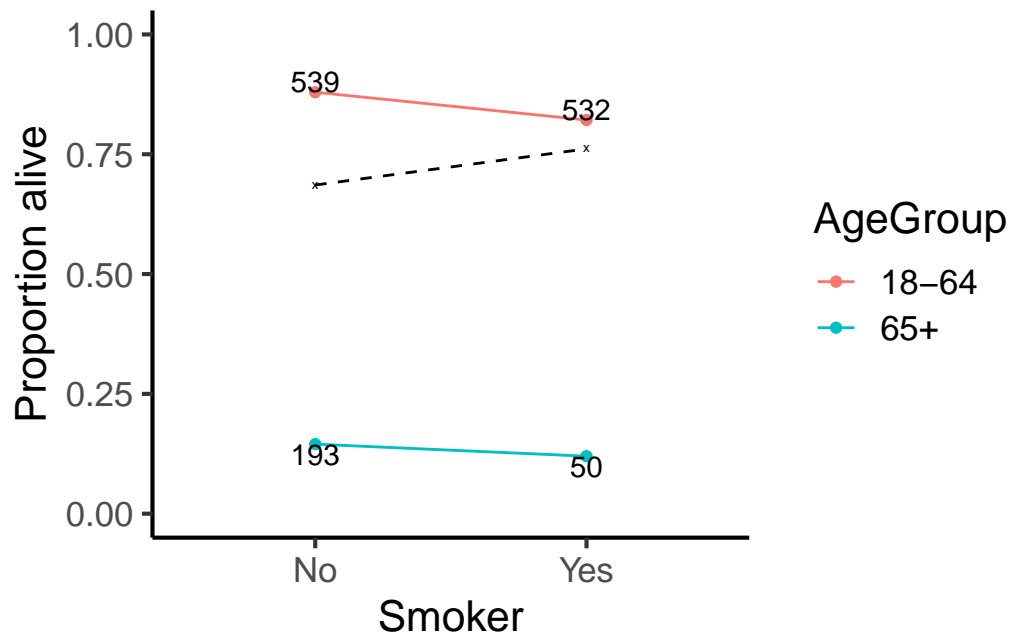
| AgeGroup | Smoker | | Sum |
|----------|--------|-----|------|
| | No | Yes | |
| 18 to 64 | 539 | 532 | 1071 |
| 65+ | 193 | 50 | 243 |
| Sum | 732 | 582 | 1314 |

- Younger people are $\frac{532/539}{50/193} = \frac{0.987}{0.259} = 3.81$ times more likely to be smokers than older people
 - $1071/243 = 4.41$ times more young people than older people
 - A lot of young people, who are more likely to smoke and less likely to die

3.2.4 Plot: Simpson's paradox



3.2.5 Plot: Total n for each proportion



3.2.6 Simpson's paradox

- Not really a paradox
 - Just a different way of looking at the information
- When you have 3 variables
 - Focus is XY relationship
 - But also pay attention to the ZY and XZ relationships

3.3 Conditional and marginal odds ratios

3.3.1 Conditional and marginal effects

- Partial tables
 - Conditional associations odds ratios, chi-square
 - Association between X and Y , at a given value of Z
- Marginal tables
 - Marginal associations, odds ratios, chi-square
 - Association between X and Y , ignoring Z

3.3.2 Partial tables

, , AgeGroup = 18-64

| Alive | | | |
|--------|------|-------|------|
| Smoker | Dead | Alive | Sum |
| No | 65 | 474 | 539 |
| Yes | 95 | 437 | 532 |
| Sum | 160 | 911 | 1071 |

, , AgeGroup = 65+

| Alive | | | |
|--------|------|-------|-----|
| Smoker | Dead | Alive | Sum |
| No | 165 | 28 | 193 |
| Yes | 44 | 6 | 50 |
| Sum | 209 | 34 | 243 |

, , AgeGroup = Sum

| Alive | | | |
|--------|------|-------|------|
| Smoker | Dead | Alive | Sum |
| No | 230 | 502 | 732 |
| Yes | 139 | 443 | 582 |
| Sum | 369 | 945 | 1314 |

3.3.3 Conditional odds ratios: Young people

- $\hat{\theta} = \frac{474/65}{437/95} = \frac{7.292}{4.6} = 1.585$
- Odds of *non-smoker* being alive = 7.292
 - A non-smoker is 7.292 times more likely to be alive than dead
- Odds of *smoker* being alive = 4.6
 - A smoker is 4.6 times more likely to be alive than dead
- Odds ratio = 1.585: Odds of a *non-smoker* being alive is 1.585 times the odds of an *smoker* being alive
 - Non-smokers are more likely to be alive than smokers

3.3.4 Conditional odds ratios: Older people

- $\hat{\theta} = \frac{28/165}{6/44} = \frac{0.170}{0.136} = 1.248$
- Odds of *non-smoker* being alive = 0.170
 - A non-smoker is 0.170 times more likely to be alive than dead
- Odds of *smoker* being alive = 0.136
 - A smoker is 0.136 times more likely to be alive than dead
- Odds ratio = 1.248: Odds of a *non-smoker* being alive is 1.248 times the odds of an *smoker* being alive
 - Non-smokers are more likely to be alive than smokers

3.3.5 Marginal table

| | Alive | | |
|--------|-------|-------|------|
| Smoker | Dead | Alive | Sum |
| No | 230 | 502 | 732 |
| Yes | 139 | 443 | 582 |
| Sum | 369 | 945 | 1314 |

3.3.6 Marginal odds ratio

- $\hat{\theta} = \frac{502/230}{443/139} = \frac{2.183}{3.187} = 0.685$
- Odds of *non-smoker* being alive = 2.183
 - A non-smoker is 2.183 times more likely to be alive than dead
- Odds of *smoker* being alive = 3.187
 - A smoker is 3.187 times more likely to be alive than dead
- Odds ratio = 0.685: Odds of a *non-smoker* being alive is 0.685 times the odds of an *smoker* being alive
 - Non-smokers are **less likely to be alive** than smokers

3.3.7 Conditional and marginal effects

- Neither tell the whole story usually
 - Look at both
- We looked at odds ratios
 - Also consider conditional and marginal difference in proportion and relative risk, if applicable

3.4 Conditional and marginal independence

3.4.1 Conditional vs marginal independence

- Variables can be
 - Conditionally independent
 - Marginally independent
 - Both
 - Neither
- Depends on XY relationship, as well as ZY and XZ relationships

3.4.2 Conditional independence...

- Clinic 1

| | Success | Failure | |
|-------------|---------|---------|----|
| Treatment A | 18 | 12 | 30 |
| Treatment B | 12 | 8 | 20 |
| | 30 | 20 | 50 |

- Odds of success is 1.5 times odds of failure for both treatments
- $OR = \frac{18/12}{12/8} = \frac{1.5}{1.5} = 1$

- Clinic 2

| | Success | Failure | |
|-------------|---------|---------|----|
| Treatment A | 2 | 8 | 10 |
| Treatment B | 8 | 32 | 40 |
| | 10 | 40 | 50 |

- Odds of failure is 4 times odds of success for both treatments
- $OR = \frac{2/8}{8/32} = \frac{.25}{.25} = 1$

3.4.3 ... but not marginal independence

- Combine both clinics into a single large sample

| | Success | Failure | |
|-------------|---------|---------|-----|
| Treatment A | 20 | 20 | 40 |
| Treatment B | 20 | 40 | 60 |
| | 40 | 60 | 100 |

- Odds of success for treatment A is twice odds of success for treatment B
- $OR = \frac{20/20}{20/40} = \frac{1}{0.5} = 2$

3.4.4 What to do?

- Look at both **partial** and **marginal** tables
- Unless XY relationship is same at every level of Z , **partial and marginal effects will tell different stories**
- Very similar to covariates in ANOVA or regression
 - Main effect for XY relationship doesn't tell you the complete story if
 - * Covariate is related to the outcome
 - * Covariate and predictor are related
 - * Interaction of covariate and predictor

4 Summary

4.1 Summary

4.1.1 Summary of this week

- Extend to **more levels** and/or **more variables**
 - Additional complexity with
 - * Partitioning effects
 - * Partial vs marginal effects

4.1.2 Summary of this section

- Contingency tables
 - 2×2 and larger
- Measures of relationship
 - Difference in proportion
 - Relative risk
 - Odds ratio
 - Chi-square tests
- Extend these to larger tables