# Homework 5

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## Description

• Course: STAT638, 2022 Fall

Read Chapter 5 in the Hoff book.

Then, do the following exercises in Hoff: 5.1, 5.2, 5.5.

For 5.2, \_A denotes the mean exam score for students assigned to method A, and \_B denotes the mean exam score for students assigned to method B.

Be careful: Some of the notation differs between the lecture slides and the Hoff book.

• Deadline: Oct 12 by 12:01pm

### **Computational Environment Setup**

#### Third-party libraries

```
%matplotlib inline
import sys # system information
import matplotlib # plotting
import scipy # scientific computing
import random
import pandas as pd # data managing
from scipy.special import comb
from scipy import stats as st
from scipy.special import gamma
import numpy as np
import matplotlib.pyplot as plt
from itertools import permutations
# Matplotlib setting
plt.rcParams['text.usetex'] = True
matplotlib.rcParams['figure.dpi'] = 300
np.random.seed(20220928) # Consistent random effect
```

#### Version

```
print(sys.version)
print(matplotlib.__version__)
print(scipy.__version__)
print(np.__version__)
print(pd.__version__)

3.8.14 (default, Sep 6 2022, 23:26:50)
[Clang 13.1.6 (clang-1316.0.21.2.5)]
3.3.1
1.5.2
1.19.1
1.1.1
```

#### Problem 5.1

Studying: The files school1.dat, school2.dat and school3.dat contina data on the amount of time students from 3 high schools spent on studying or homework during an exam period. Analyze data from each of these schools separately, using the normal model with a conjugate prior distribution, in which  $\{\mu_0 = 5, \sigma_0^2 = 4, \kappa_0 = 1, \nu_0 = 2\}$  and compute or approximate following:

```
schs = [np.loadtxt("data/school{}.dat".format(i)) for i in range(1,4)]
```

(a)

Posterior means and 95% confidence intervals for the mean  $\theta$  and standard deviation  $\sigma$  from each school;

```
class NormalModel:
 def __init__(self, data, mu0, var0, k0, v0):
   self.mu0 = mu0 # prior mean
   self.var0= var0 # prior variance
   self.k0= k0 # prior observation
   self.v0 = v0
   # Data
   self.data = data
   self.mean = np.mean(data) # sample mean
   self.n = len(data) # sample counts
   self.kn = self.n + self.k0
   self.s2 = np.sum((self.data-self.mean)**2) / (self.n - 1)
   # Posterior parameters
   self.vn = self.get_vn(self.v0, self.n) # Posterior v
   self.varn = self.get_varn(self.vn, self.v0, self.var0, self.k0, self.n, self.kn, self.
   self.mun = self.get_mun(self.k0, self.mu0, self.n, self.mean, self.kn) # Posterior mea
 def get_vn(self, v0, n):
   return v0 + n
```

```
def get_varn(self, vn, v0, var0, k0, n, kn, mean, mu0):
    varn = (1/vn)*(v0*var0 + (n-1)*self.s2 + (k0*n/kn)*(mean - mu0)**2)
    return varn
  def get_mun(self, k0, mu0, n, mean, kn):
    return (k0*mu0 + n*mean)/kn
  def rv_theta_post(self):
    mu0 = self.mu0; k0 = self.k0;
    return st.norm(loc=self.mun, scale= np.sqrt(self.varn/self.kn))
  def rv_pred(self):
    mun = self.mun
    varn = self.varn
    return st.norm(loc=self.mun, scale= np.sqrt(self.varn))
setting = {"mu0": 5, "var0":4, "k0":1, "v0":2}
nms = [NormalModel(sch, **setting) for sch in schs]
rv_thps = [nm.rv_theta_post() for nm in nms]
# Display data
pd.DataFrame({"School":[i+1 for i in range(0, len(nms))], "Mean":[rv_thp.mean() for rv_thp
```

School Mean		Mean	95% confidence interval (left, right)
0	1	9.292308	(7.832422909482735, 10.752192475132652)
1	2	6.948750	(5.242990325564609, 8.654509674435392)
2	3	7.812381	(6.26475553664974, 9.360006368112167)

## (b)

The posterior probability that  $\theta_i < \theta_j < \theta_k$  for all size permutataions  $\{i, j, k\}$  of  $\{1, 2, 3\}$ ;

```
n_mc = 10000
samps_thp = [rv_thp.rvs(size=n_mc) for rv_thp in rv_thps]
perm = list(permutations([0,1,2]))
text = [""] * len(perm)
probs = [int] * len(perm)
for (a, i) in enumerate(perm):
   text[a] = "theta_{} < theta_{} < theta_{} .format(i[0]+1, i[1]+1, i[2]+1)</pre>
```

```
probs[a] = np.sum(np.logical_and(samps_thp[i[0]] < samps_thp[i[1]], samps_thp[i[1]] < samps_thp[i[1]] < samps_thp[i[1]], samps_thp[i[1]] < samps_thp[i[1]], samps_thp[i[1]] < samps_thp[i[1]], samps_thp[i[1]
```

	Permuatations	Probabilities
0	$theta_1 < theta_2 < theta_3$	0.0058
1	$theta_1 < theta_3 < theta_2$	0.0030
2	$theta\_2 < theta\_1 < theta\_3$	0.0788
3	$theta\_2 < theta\_3 < theta\_1$	0.6929
4	$theta_3 < theta_1 < theta_2$	0.0123
5	$theta\_3 < theta\_2 < theta\_1$	0.2072

(c)

The posterior probability that  $\tilde{Y}_i < \tilde{Y}_j < \tilde{Y}_k$  for all siz permutations  $\{i,j,k\}$  of  $\{1,2,3\}$ , where  $\tilde{Y}_i$  is a sample from the posterior predictive distribution of school i

```
rv_posts = [nm.rv_pred() for nm in nms]

samps_post = [rv_post.rvs(size=n_mc) for rv_post in rv_posts]
perm = list(permutations([0,1,2]))
text = [""] * len(perm)
probs = [int] * len(perm)
for (a, i) in enumerate(perm):
   text[a] = "Y_tilde_{{}} < Y_tilde_{{}} < Y_tilde_{{}}".format(i[0]+1, i[1]+1, i[2]+1)
   probs[a] = np.sum(np.logical_and(samps_post[i[0]] < samps_post[i[1]], samps_post[i[1]] <

pd.DataFrame({"Permutations": text, "Probabilities": probs})</pre>
```

	Permuatations	Probabilities
0	$Y_{tilde} = 1 < Y_{tilde} = 2 < Y_{tilde}$	0.1020
1	$Y_{tilde} = 1 < Y_{tilde} = 3 < Y_{tilde} = 2$	0.1029
2	$Y_{tilde} < Y_{tilde} < Y_{tilde} < Y_{tilde}$	0.1857
3	$Y_{tilde} < Y_{tilde} < 3 < Y_{tilde}$	0.2686
4	$Y_{tilde_3} < Y_{tilde_1} < Y_{tilde_2}$	0.1361
5	$Y\_tilde\_3 < Y\_tilde\_2 < Y\_tilde\_1$	0.2047

(d)

Compute the posterior probability that  $\theta_1$  is bigger than both  $\theta_2$  and  $\theta_3$ , and the posterior probability that  $\tilde{Y}_1$  is bigger than both  $\tilde{Y}_2$  and  $\tilde{Y}_3$ .

```
th_biggest = np.sum(np.logical_and(samps_thp[0]>samps_thp[1], samps_thp[0]>samps_thp[2]))/
post_biggest = np.sum(np.logical_and(samps_post[0]>samps_post[1], samps_post[0]>samps_post
pd.DataFrame({"Properties": ["Theta_1 is the biggest", "Tilde_1 is the biggest"], "Probabi
```

	Properties	Probabilities
0	Theta_1 is the biggest	0.9001
1	Tilde_1 is the biggest	0.4733

#### Problem 5.2

Sensitivity analysis: 32 students in a science classroom were randomly assigned to one of two study methods, A and B, so that  $n_A = n_B = 16$  students were assigned to each method. After several weeks of study, students were examined on the course material with an exam designed to give an average score of 75 with a standard deviation of 10. The scores for the two groups are summarized by  $\{\bar{y}_A = 75.2, s_A = 7.3\}$  and  $\{\bar{y}_B = 77.5, s_b = 8.1\}$ . Consider independent, conjugate normal prior distributions for each of  $\theta_A$  and  $\theta_B$ , with  $\mu_0 = 75$  and  $\sigma_0^2 = 100$  for both groups. For each  $(\kappa_0, \nu_0) \in \{(1, 1), (2, 2), (4, 4), (8, 8), (16, 16), (32, 32)\}$  (or more values), obtain  $Pr(\theta_A < \theta_B | y_A, y_B)$  via Monte Carlo Sampling. Plot this probability as a function of  $\kappa_0 = \nu_0$ . Describe how you might use this plot to convey the evidence that  $\theta_A < \theta_B$  to people of a variety of prior opinions.

• Increase  $(\kappa_0, \nu_0)$  can descrease the probability of B bigger that than A. But, the probability of B larger than A can hardly lower than 0.5.

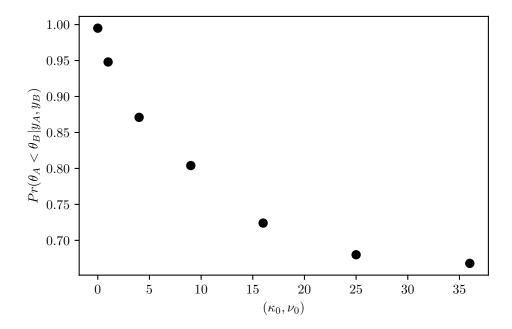
```
class NormalModel:
```

```
def __init__(self, sample_size,sample_mean, sample_var, mu0, var0, k0, v0):
    # Prior
    self.mu0 = mu0 # prior mean
    self.var0= var0 # prior variance
    self.k0= k0 # prior observation
    self.v0 = v0
    # Data
    self.mean = sample mean # sample mean
```

```
self.n = sample_size # sample counts
    self.kn = self.n + self.k0
    self.s2 = sample_var
    # Posterior parameters
    self.vn = self.get_vn(self.v0, self.n) # Posterior v
    self.varn = self.get_varn(self.vn, self.v0, self.var0, self.k0, self.n, self.kn, self.
    self.mun = self.get_mun(self.k0, self.mu0, self.n, self.mean, self.kn) # Posterior mea
  def get_vn(self, v0, n):
    return v0 + n
  def get_varn(self, vn, v0, var0, k0, n, kn, mean, mu0):
    varn = (1/vn)*(v0*var0 + (n-1)*self.s2 + (k0*n/kn)*(mean - mu0)**2)
    return varn
  def get_mun(self, k0, mu0, n, mean, kn):
    return (k0*mu0 + n*mean)/kn
  def rv_theta_post(self):
    mu0 = self.mu0; k0 = self.k0;
    return st.norm(loc=self.mun, scale= np.sqrt(self.varn/self.kn))
  def rv_pred(self):
    mun = self.mun
    varn = self.varn
    return st.norm(loc=self.mun, scale= np.sqrt(self.varn))
def expAB(k0, v0, size=1000):
  setting = {"sample_size":16, "mu0": 75, "var0": 100, "k0":k0, "v0":v0}
  rvtA = NormalModel(sample_mean=75.2,\
    sample_var=7.3, **setting).rv_theta_post()
  rvtB = NormalModel(sample_mean=77.5,\
    sample_var=8.1, **setting).rv_theta_post()
  spA, spB = rvtA.rvs(size=size), rvtB.rvs(size=size)
  return np.sum(spA < spB)/size
pars = [i**2 for i in range(0, 7)]
vals = [float]*len(pars)
```

```
for (i, p) in enumerate(pars):
    vals[i] = expAB(p,p)

# Plotting
figV, axV = plt.subplots()
axV.plot(pars,vals, 'o', color="k")
axV.set_xlabel("$(\\kappa_0, \\nu_0)$")
axV.set_ylabel("$Pr(\\theta_A < \\theta_B | y_A, y_B)$");</pre>
```



### Problem 5.5

Unit information prior: Obtain a unit information prior for the normal model as follows:

(a)

Reparameterize the normal model as  $p(y|\theta,\psi)$ , where  $\psi=\frac{1}{\sigma^2}$ . Write out the log likelihood  $l(\theta,\psi|y)=\sum \log p(y_i|\theta,\psi)$  in terms of  $\theta$  and  $\psi$ .

$$p(y|\theta,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}(\frac{y-\theta}{\sigma})^2}$$

$$p(y|\theta,\psi) = \frac{\sqrt{\psi}}{\sqrt{2\pi}} e^{-\frac{1}{2}(y-\theta)^2\psi}$$

$$\log p(y_i|\theta,\psi) = \log \left(\sqrt{\frac{\psi}{2\pi}}\right) - \frac{1}{2}(y_i - \theta)^2 \psi$$

$$l(\theta, \psi | y) = \sum_{i=1}^{n} \log(y_i | \theta, \psi) = \log\left(\frac{\psi}{2\pi}\right)^{n/2} - \frac{1}{2}\psi \sum_{i=1}^{n} (y_i - \theta)^2 \tag{1}$$

$$= \frac{n}{2} \log \left(\frac{\psi}{2\pi}\right) - \frac{\psi}{2} \sum_{i=1}^{n} (y_i - \theta)^2 \tag{2}$$

$$= \frac{n}{2} \log \left( \frac{\psi}{2\pi} \right) - \frac{\psi}{2} \sum_{i=1}^{n} (y_i^2 - 2\theta y_i + \theta^2)$$
 (3)

$$=\frac{n}{2}\log\left(\frac{\psi}{2\pi}\right)-\frac{\psi}{2}(\sum y_i^2-2\theta\sum y_i+n\theta^2) \tag{4}$$

$$= \frac{-\psi}{2} \sum_{i=1}^{n} y_i^2 + \psi \theta \sum_{i=1}^{n} y_i - \frac{n\psi \theta^2}{2} + \frac{n}{2} \log \left(\frac{\psi}{2\pi}\right) \tag{5}$$

(b)

Find a probability density  $P_U(\theta, \psi)$  so that  $\log p_U(\theta, \psi) = \frac{l(\theta, \psi|y)}{n} + c$ , where c is a constant that does not depend on  $\theta$  or  $\psi$ .

Hint: Write  $\sum (y_i-\theta)^2$  as  $\sum (y_i-\bar{y}+\bar{y}-\theta)^2=\sum (y_i-\bar{y})^2+n(\theta-\bar{y})^2$ , and recall that  $\log p_U(\theta,\psi)=\log p_U(\theta|\psi)+\log p_U(\psi)$ .

$$l(\theta, \psi|y) = \frac{n}{2} \log\left(\frac{\psi}{2\pi}\right) - \frac{\psi}{2} \sum_{i=1}^{n} (y_i - \theta)^2$$
 (6)

$$= \frac{n}{2} \log \left( \frac{\psi}{2\pi} \right) - \frac{\psi}{2} \left[ \sum (y_i - \bar{y})^2 + n(\theta - \bar{y})^2 \right] \tag{7}$$

$$= \frac{n}{2} \log \left(\frac{\psi}{2\pi}\right) - \frac{\psi}{2} \left[ (n-1)s^2 + n(\theta - \bar{y})^2 \right]$$
 (8)

$$= \frac{n}{2}\log\left(\frac{\psi}{2\pi}\right) - \frac{\psi}{2}\left((n-1)s^2 + n\theta^2 - 2n\theta\bar{y} + n\bar{y}^2\right) \tag{9}$$

$$= \frac{n}{2} \log \left( \frac{\psi}{2\pi} \right) - \frac{\psi}{2} (n-1)s^2 - \frac{\psi n\theta^2}{2} + n\psi \theta \bar{y} - \frac{\psi n\bar{y}^2}{2}$$
 (10)

$$= \frac{n}{2} \log \left( \frac{\psi}{2\pi} \right) - \frac{\psi}{2} (n-1)s^2 - \frac{\psi n \bar{y}^2}{2} + n\psi \theta \bar{y} - \frac{\psi n \theta^2}{2}$$
 (11)

$$\frac{l(\theta, \psi|y)}{n} + c = \underbrace{\frac{1}{2}\log\left(\frac{\psi}{2\pi}\right) - \frac{\psi}{2n}(n-1)s^2 - \frac{\psi\bar{y}^2}{2}}_{=h(\psi|y)} + \underbrace{\psi\theta\bar{y} - \frac{\psi\theta^2}{2}}_{=g(\theta|\psi,y)} + c$$
(12)

$$p_U(\theta, \psi) = \exp\left(\frac{l(\theta, \psi|y)}{n} + c\right) \tag{13}$$

$$\propto \exp\left(\frac{l(\theta,\psi|y)}{n}\right)$$
 (14)

$$= \exp(h(\psi|y)) \exp(g(\theta|\psi, y)) \tag{15}$$

$$\exp(h(\psi|y)) = (\frac{\psi}{2\pi})^{1/2} \exp\left(-(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})\psi\right)$$
 (16)

$$\propto \psi^{\frac{1}{2}} \exp\left(-\frac{\psi}{(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})^{-1}}\right)$$
 (17)

$$\sim Gamma(\psi, \frac{3}{2}, (\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})^{-1})$$
 (18)

$$\exp(g(\theta|\psi,y)) = \exp\left(\psi\theta\bar{y} - \frac{\psi\theta^2}{2}\right) \tag{19}$$

$$= \exp\left(-\frac{\psi}{2}\theta^2 + \psi\bar{y}\theta - \frac{\psi\bar{y}^2}{2} + \frac{\psi\bar{y}^2}{2}\right) \tag{20}$$

$$= \exp\left(-\frac{\psi}{2}(\theta^2 - 2\bar{y}\theta + \bar{y}^2) + \frac{\psi\bar{y}^2}{2}\right) \tag{21}$$

$$= \exp\left(-\frac{\psi}{2}(\theta - \bar{y})^2 + \frac{\psi\bar{y}^2}{2}\right) \tag{22}$$

$$\propto \exp\left(-\frac{\psi}{2}(\theta-\bar{y})^2\right)$$
 (23)

$$= \exp\left(-\frac{1}{2}\left(\frac{\theta - \bar{y}}{\psi^{-1/2}}\right)^2\right) \tag{24}$$

$$\propto Normal(\theta, \bar{y}, \psi^{-1/2}) \tag{25}$$

Thus<sup>1</sup>,

<sup>&</sup>lt;sup>1</sup>Use the PDF formula on Wiki

$$P_{U}(\theta, \psi) = P_{U}(\theta|\psi)P_{U}(\psi) \propto Gamma(\psi, \frac{3}{2}, (\frac{(n-1)s^{2}}{2n} + \frac{\bar{y}^{2}}{2})^{-1})$$
 (26)

$$\times Normal(\theta, \bar{y}, \psi^{-1/2}) \tag{27}$$

(28)

(c)

Find a probability density  $p_U(\theta, \psi|y)$  that is proportional to  $p_U(\theta, \psi)$  ×  $p(y_1, \dots, y_n | \theta, \psi)$ . It may be convenient to write this joint density as  $p_U(\theta|\psi,y) \times p_U(\psi|y)$ . Can this joint density be considered a posterior density?

$$P_U(\theta, \psi|y) = P_U(\psi|y) \times p_U(\theta|\psi, y) \tag{29}$$

$$= Gamma(\psi, \frac{3}{2}, (\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})^{-1})$$
(30)

$$\times Normal(\theta, \bar{y}, \psi^{-1/2}) \tag{31}$$

$$\propto \psi^{1/2} \exp\left(-\left(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2}\right)\psi\right) \times \exp\left(-\frac{1}{2}\left(\frac{\theta - \bar{y}}{\psi^{-1/2}}\right)^2\right)$$
(32)

$$\propto \psi^{1/2} \exp\left(-(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})\psi\right) \times \exp\left(-\frac{\psi}{2} \left(\theta - \bar{y}\right)^2\right) \tag{33}$$

$$\propto \psi^{1/2} \exp\left(-(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2})\psi\right) \times \exp\left(-\frac{\psi}{2}\left(\theta^2 - 2\theta\bar{y} + \bar{y}^2\right)\right) \tag{34}$$

$$\psi^{1/2} \exp\left(-\left(\frac{(n-1)s^2}{2n} + \frac{\bar{y}^2}{2} + \frac{\bar{y}^2}{2}\right)\psi\right) \times \exp\left(-\frac{\psi}{2}(\theta^2 - 2\theta\bar{y})\right) \tag{35}$$

- $p_U(\theta|\psi,y) \propto Normal(\theta,\bar{y},\psi^{-1/2})$   $p_U(\psi|\theta,y) \sim Gamma(\frac{3}{2},(\frac{(n-1)s^2}{2n}+\frac{\bar{y}^2}{2})^{-1})$  Yes, it can be considered as a posterior density because the unit information prior is the product of two proper distribution.

#### References