TWO CLASSES OF MULTI-MODEL CANDIDATE MANAGEMENT PROCEDURES FOR ATLANTIC BLUEFIN TUNA

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SUMMARY

Two classes of multi-model candidate management procedures for Atlantic bluefin tuna were developed and tested. Procedures were based on spawning biomass estimation methods scaled to five operating models selected via cluster analysis from the reference OM grid. For the empirical class, OM catchability and a constant stock mixing distribution were used to estimate area biomass from the larval indices. For model-based CMPs, five delay difference assessment models were scaled to each of the five operating models, matching stock recruit steepness and biomass for the recent historical period from 1965 - 2016. At each time step, estimates of current (empirical) or projected (model-based) biomass were generated from approved management indices and used in harvest control rules to generate area-specific TACs, and the five TACs were averaged to produce harvest advice for the East and West area. Model-based CMPs tuned to median biomass targets across the reference OM grid avoided crashing both east and west spawning stocks in over 97.5% of simulations.

RÉSUMÉ

Deux catégories de procédures de gestion multi-modèles potentielles pour le thon rouge de l'Atlantique ont été développées et testées. Les procédures étaient basées sur des méthodes d'estimation de la biomasse du stock reproducteur échelonnées à cinq modèles opérationnels sélectionnés par analyse de grappes à partir de la grille de référence du modèle opérationnel (OM). Pour la classe empirique, la capturabilité des OM et une distribution constante du mélange des stocks ont été utilisées pour estimer la biomasse de la zone à partir des indices larvaires. Pour les procédures de gestion potentielles basées sur des modèles, cinq modèles d'évaluation à différences retardées ont été échelonnés à chacun des cinq modèles opérationnels, en faisant correspondre la pente à l'origine de la relation stock-recrutement (steepness) et la biomasse pour la période historique récente de 1965 à 2016. À chaque étape, des estimations de la biomasse actuelle (empirique) ou projetée (basée sur le modèle) ont été générées à partir d'indices de gestion approuvés et utilisées dans les règles de contrôle de l'exploitation pour générer des TAC spécifiques à la zone. La moyenne des cinq TAC a été calculée pour produire un avis de capture pour la zone Est et Ouest. Les CMP basées sur des modèles et calibrées aux objectifs de biomasse médiane sur la grille de référence des OP ont évité l'effondrement des stocks reproducteurs de l'Est et de l'Ouest dans plus de 97,5% des simulations

RESUMEN

Se desarrollaron y probaron dos clases de procedimientos de ordenación candidatos con varios modelos para el atún rojo del Atlántico. Los procedimientos se basaban en los métodos de estimación de la biomasa reproductora escalada a cinco modelos operativos seleccionados de la matriz de referencia de OM mediante un análisis de conglomerados. Para la clase empírica, se utilizaron la capturabilidad de los OM y una distribución constante de la mezcla del stock para estimar la biomasa del área a partir de los índices larvarios. Para los CMP basados en modelos, cinco modelos de evaluación de diferencia retardada se escalaron a cada uno de los cinco modelos operativos, haciendo corresponder la inclinación stock recluta y la biomasa para el reciente periodo histórico desde 1965 a 2016. En cada fase temporal, las estimaciones de la biomasa actual (empírica) o proyectada (basada en el modelo) se generaron a partir de índices

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de ordenación aprobados y utilizados en normas de control de la captura para generar TAC específicos de áreas, y los cinco TAC se promediaron para formular un asesoramiento sobre captura para las zonas del este y del oeste. Los CMP basados en modelos y calibrados con los objetivos de biomasa media en la matriz de referencia del OM evitaron el colapso de los stocks reproductores del este y del oeste en más del 97,5% de las simulaciones

KEYWORDS

Atlantic bluefin tuna, Multi-model CMPs, Oms

1. Introduction

Two classes of multi-model CMPs were developed and tested in the Atlantic Bluefin tuna MSE framework. The first class is based on an empirical index-based method of estimating spawning stock biomass, and the second uses a multi-stock state-space delay difference stock assessment model. Both classes were made up of five biomass estimation methods and harvest control rules that are each scaled to one of five reference grid operating models. Catch advice from each sub-procedure is then combined to produce area-specific TACs. Model based CMP performance was tuned to three separate combinations of relative biomass targets for each spawning stock via a grid search over parameters controlling the harvest rates.

2. Methods

Our CMPs combined estimates of biomass and associated catch advice from sub-procedures tuned to a subset of the reference grid of multi-stock operating models. This multi-model formulation allowed us to capture the biological uncertainty represented by the reference grid of operating models.

We defined two classes of CMP: empirical and model-based. The empirical MPs scale larval indices by OM estimates of catchability, while the model-based MPs fit state space delay difference assessment model to management indices and pre-2015 spawning biomass.

Scaling subset of operating models

We scaled our estimation procedures on a subset of OMs which were chosen to be representative of the main OM uncertainty factors (1: steepness, 2: maturity/M, 3: stock mixing, 4: mean SSB, 5: length composition weight) and the SSB range in the OMs. To choose representative OMs, we ran a clustering algorithm on the SSB time-series to identify k cluster centers or "medoids", then checked the factor levels associated with each medoid. We considered a set of OMs as adequately representative if (i) SSB trajectories in the subset spanned approximately the same range as the SSB in the OMs, and (ii) each uncertainty factor level was represented at least once. OMs with changing recruitment steepness in the projection years were excluded from the clustering analysis, leaving a total of 64 OMs to be clustered.

We used the Time-series Anytime Density Peaks (TADPole) algorithm to cluster SSB time-series (Begum *et al.* 2016). To simultaneously cluster the SSB in each area, we first created a single SSB "time-series" for each OM by adding the west SSB time-series onto the end of the east SSB time-series. We then ran the clustering algorithm on the resulting 32 combined time-series for k=4 to k=8. The weight associated with each cluster was calculated as the proportion of individual time-series in the cluster.

Empirical estimation procedure

The empirical estimation procedure used a simple moving average of the two spawning stock larvel survey indices. For each time step t, and stock $s \in \{GOM, MED\}$ the current average larvel index is calculated as

$$\bar{I}_{s,t} = \frac{\sum_{t'=t-3}^{t} I_{s,t'}}{4},$$

where $I_{s,t}$ is the larval index for stock s at time t. We then scaled the smoothed survey indices to multiple biomass estimates, one for every operating model in the tuning subset, by using the appropriate larval survey catchability parameter, e.g

$$\widehat{B}_{s,t,j} = \frac{\overline{I}_{s,t}}{q_{s,j}},$$

where $q_{s,j}$ is the stock larval index catchability parameter from the tuning operating model $j \in \{37, 14, 53, 31, 89\}$.

Spawning stock biomass was then translated to area biomass using an assumed constant distribution of biomass for stock mixing

 $P = (p_{s,a}) = \begin{pmatrix} .898 & .102 \\ .1 & .9 \end{pmatrix},$

found by averaging the proportion of stock biomass in each area over the historical period in the tuning set of OMs. The rows of matrix P are indexed by spawning stocks $\{MED, GOM\}$, and the columns are indexed by management area $\{E, W\}$. This generates an area specific biomass

$$B_a = \sum_{S} p_{s,a} B_{s},$$

where $s \in \{MED, GOM\}, a \in \{E, W\}.$

Delay difference model

For the model-based class of procedures, the simple index scaling method of the empirical procedure is replaced by a state-space multi-stock Deriso-Schnute delay difference assessment model (Deriso 1980; Schnute 1985). The two spawning stocks are assumed to be distinct spawning stocks for the larval surveys, but are mixed using the same mixing distribution *P* as the empirical CMP for the other management indices and the observed catches.

2.1.1 Equilibria

Because there is no mixing of spawners, each stock has the simple Delay Difference equilibrium states, which can be expressed as a function of long-term fishing mortality $f = F_s$. The equations for equilibria in **Table 1** were used to initialise the model at a fished equilibrium in 1965, as well as estimate biological reference points for use in the harvest control rules.

2.1.2 Biomass time series reconstruction

Each assessment model was initialised in a fished state in 1965. To do so, we estimated an initial fishing mortality rate F(s,0), and assumed that the stock was at the fished equilibrium defined in **Table 1**. Other free parameters were unfished biomass $B_{s,0}$, catchability q_g and observation error uncertainty τ_g for biomass and abundance indices, and recruitment process error deviations $\omega_{s,t}$ for $t \in 1966, ..., 2014$ (**Table 2**). Stock recruit steepness h_s and natural mortality M_s were fixed to the estimated values from the associated operating model (Note: west h shifts from 0.6 to 0.9 in 1975 in some OMs. In these scenarios, we set west h to 0.9).

The proportion $p_{s,a}$ of stock $s \in \{MED, GOM\}$ in area $a \in \{E, W\}$ was assumed to be constant, allowing the fluctuations in stock biomass to account for the mixing dynamics. Both stocks were also assumed to be homogeneously mixed in each area, so that total stock specific catch was the sum of the area specific catches, split according to the proportion to the stock biomass in each area.

Instantaneous fishing mortality was solved via a numerical Newton-Rhapson method within the delay difference model for all time steps $t \ge 1$, which were used in the survival calculation to progress numbers and biomass to the next time step.

Recruitment process errors were modeled as a simple random walk, rather than indepent deviations from the stock recruit curve, chosen because it was more able to capture dynamic recruitment regimes. Furthermore, random walks would also make projected recruitments, which are an important component of projected biomass in a delay difference formulation, less likely to be average, and more like the estimated recruitment in the last year of the assessment period.

Because this is a model-based MP and the time-series of approved management indices are rather short, two additional indices were included. These indices were the M3 model yearly spawning stock biomass estimates for the East and West stocks. These biomass time series were assumed to be observed with catchability q=1 and an observation error CV of 5%, allowing the delay difference AMs to estimate catchability for the approved indices and scale them appropriately to the spawning stock biomass from the associated operating model. This scaling was

required for extending the fit to the approved indices in the projections, as well as making sensible estimates of biological parameters despite the complexity mismatch between the operating model and the assessment models. Finally, a log-normal prior distribution was applied to unfished biomass $B_{0,S}$ to prevent the unfished biomass from being estimated too close to the initial biomass for the assessment period, and to avoid producing optimistic estimates of initial and current biomass depletion. For each AM, the prior mean was defined as the estimate from the associated operating model with a log standard deviation of 0.05.

2.1.3 Reference Points

Reference points for each stock were estimated from the equilibrium values in (**Table 1**). The optimal fishing mortality rate F_{MSY} was found by numerically solving for the stationary point of the yield curve, and was in turn used to estimate the optimal equilibrium biomass B_{MSY} associated with that mortality rate.

Area based reference points for use in harvest control rules were defined as the stock-area biomass weighted averages of the stock-specific reference points, i.e.

$$B_{MSY,a} = \sum_{s} \frac{B_{s,a}}{B_a} B_{MSY,s},$$

$$F_{MSY,a} = \sum_{s} \frac{B_{s,a}}{B_a} F_{MSY,s},$$

where $B_{s,a} = p_{s,a}B_s$.

Harvest control rules

For each of the five sub-procedures, a ramped harvest control rule was defined with upper and lower control points and a maximum harvest rate (**Figure 1**). The control points and maximum harvest rates for these rules were either based on the associated operating model's biological reference points and a precautionary TAC cap to limit removals in cases of bias in projected biomass or reference point estimates.

Our ramped harvest control rule was based on the rules used for albacore tuna. These rules require an upper control point (UCP), a lower control point (LCP) and a maximum target harvest rate U_{max} . The general form of the rules (ignoring AM, area, and stock indices) was

$$U_{targ} = \begin{cases} 0.1 \cdot U_{max} & B \leq LCP \\ U_{max} * (0.1 + 0.9 \frac{B - LCP}{UCP - LCP}) & LCP \leq B \leq UCP \\ U_{max} & UCP \leq B \end{cases}$$

For all HCRs, we set LCP = .4UCP, and tested two options for the upper control point: $UCP = B_{MSY}$ or $UCP = .4B_0$. We tested two options for the maximum target harvest rate: (i) setting $U_{max} = F_{MSY}$, or (ii) setting U_{max} equal to some multiple of M. For the model-based MPs, we tested $U_{max} = \frac{2}{3}M$. Initial trials indicated that harvest rates under $U_{max} = \frac{2}{3}M$ for the empirical MPs were too high, so we tested a range of more conservative values for U_{max} (see 2.6). Under the empirical management procedures, control points were taken from the tuning grid of operating models, while under the model-based MPs, control points were taken as the delay difference model equilibria.

We also tested a procedure wherein TACs were capped at MSY to provide precautionary harvest rates in the presence of potentially large biomass estimation errors and biases in reference point calculations, both of which are caused by the complexity mismatch between the AMs and the OMs. Stock- and area-specific MSY estimates were based on assessment model equilibria.

Both home-stock- and area-based harvest control rules were applied for East and West area TACs. For example, in the East area, HCRs were applied based on East (MED) spawning stock biomass compared to the East stock control points and harvest rates, and East area biomass compared to mixed stock control points and harvest rates. Mixed area based control points and harvest rates were averaged over the two stocks present in the area, weighted by the proportion of stock specific biomass in that area (e.g. see area-based reference points calcs in previous section), and the lower of the two TACs was chosen. From this, the West TAC would almost always be managed according to the Gulf of Mexico spawning stock harvest control rules, and the East TAC would almost always be managed according to the mixed East Area harvest control rule, as this includes the weaker stock.

Providing catch advice by area

To provide a single TAC at each time step, the five TACs advice from the sub-procedures were averaged with the AIC based weights for the model-based CMPs, and (possibly adjusted) weights from the cluster analysis for the empirical CMPs. Averaged TACs were smoothed with respect to the previous management interval's TAC allowing a maximum increase of 20%, and a maximum decrease of 50%, compared to the previous management interval.

2.1.4 AIC weight calculation for DD models

AIC weights for each AM were calculated as

$$\omega_{\alpha} = \frac{exp(-0.5 \cdot \Delta_{\alpha})}{\sum_{\alpha'} exp(-0.5 \cdot \Delta_{\alpha})},$$

where Δ_{α} was

$$\Delta_{\alpha} = AIC_{\alpha} - \min_{\alpha'} AIC_{\alpha'}$$

and AIC values were defined as

$$AIC_{\alpha} = -\log \sum_{t=1}^{T} w_{t} \sum_{g} L(t, g)$$

where α indexed the AMs that were scaled to OM cluster medoids, L(t,g) was the log-normal probability density function value for the observation from management index g in year t,T was number of years of data being fit to at each projection time step, and w_t was an annual likelihood weighting parameter for CMP tuning. While the above AIC expression is missing the model parameter penalty, and is therefore not the originally defined AIC, all five scaled AMs were identical in parameterisation and therefore had the same number of parameters, so we may exclude the parameter penalty without affecting the AIC ranking or the AIC weights.

The AIC time-weighting schemes were either equally weighted (i.e., $w_t = 1$ for all t), using only the most recent 10 years of data (i.e., $w_t = 1$ for t = T - 9, ..., T, and $w_t = 0$ f or t < T - 9.

2.1.5 Empirical CMP weighting schemes

By default, TACs for empirical CMPs were averaged using weights ω_{α} from the proportion of OMs in each OM biomass time-series cluster, where $\alpha = 1, ..., k$ indexes the operating model clusters.

To reduce potential overfishing caused by a weighted average TAC in the empirical procedures, we defined a trend-based TAC adjustment for the empirical CMPs. For the GOM and MED larval surveys, we calculated a proportional growth rate over the last four years of data as

$$\gamma(t) = 100 \cdot (exp(r_t) - 1)$$

where r_t was the exponential growth rate from the log-linear model

$$\log I_{s,t'} = \alpha + r_t \cdot t'$$

fit to the most recent four years (t' = t - 3, ..., t) of data for the GOM and MED larval surveys. An area's TAC was then adjusted via the weighting of TACs from all AMs based on the value of γ_t for the stock that spawns in that area (i.e., West TACs were adjusted based on GOM survey indices). Adjustments were realised either as a lower individual AM TAC when the growth rate was negative, or modifying the weights ω_{α} of each AM based on the trend. In all cases, individual TACs were ordered by size, from smallest to largest, and the range $(-1, \infty)$ of possible growth rates was broken into a number of bins equal to the number of AMs, with breaks defined by a parameter θ . Bin-specific rules were defined for modifying the TAC weights according to the growth rate γ_t (Table 3).

Simulation experiments and performance metrics

Several combinations of TAC caps, fishing mortality rates, and TAC weightings were tested for both model-based and empirical CMPs. All CMPs were evaluated over the full reference grid of 96 deterministic operating models, labeled OM_1d through OM_96d.

CMP performance was measured primarily by the ratio Br30 of spawning stock biomass to B_{MSY} in the 30^{th} projection year (one for each spawning stock), which was also the same metric used for CMP tuning (explained below). In addition to Br30, we also tracked average catch over the first 10 (C10), 20 (C20) and 30 (C30) years of the projection, biomass depletion in the 30^{th} year (D30), and the lowest biomass depletion value over the first 30 years (LD).

We chose a subset of all possible model-based (**Table 3**) and empirical (**Table 4**) CMPs to present in the results, with chosen CMPs based on representing multiple settings outlined above.

2.6.1 Tuning CMP Performance

Biomass targets were chosen for the East and West stocks (**Table 5**). Model-based CMPs that came close to Br30 = 1.0 for the MED spawning stock were tuned to meet Br30 = 1 within 0.01 via a search over the available tuning parameters. Tuning parameters were either continuous, i.e., the natural mortality multiplier λ , or discrete, such as the length of time series to use for AIC contributions and the TAC caps. Continuous tuning parameters were varied initially in a grid search, after which interpolation was used to produce proposed new parameter values until the required precision of 0.01 was reached, in a process similar to Newton-Rhapson optimisation. Discrete parameters, such as TAC caps (or a lack thereof) were tested as factors varied in combination with grid searches of continuous parameters. Due to limitations of time, not all biomass targets were reached with the required precision, and we present only the biomass targets associated with Br30 = 1.0 in the East (MED), in combination with approximations biomass targets in the West (GOM) with a precision of 0.02 in the results below.

Tuning of model-based CMPs was primarily via harvest rates controlled by the λ multiplier of natural mortality. The asymmetry in the spawning stock sizes and distribution of biomass between East and West areas was reflected in the effects of East and West area harvest rates on the GOM and MED spawning stocks, respectively. That is, modifying the East area harvest rates had considerable effect on Br30 distributions for both stocks, while modifying the West area harvest rate had a negligible effect on distribution of Br30 values for the MED stock. Therefore, tuning was a sequential process wherein the MED (East) stock median Br30 value was brought to 1.0 \pm 0.01 via a grid search over East area harvest rates. Once the East stock was tuned, then the GOM (West) stock median Br30 value was brought to within 0.02 of the three targets in **Table 6** via a grid search of West area harvest rates.

3. Results

3.1 Operating Model Cluster Analysis

Each of the cluster sets with $k \ge 5$ covered the historical OM SSB range reasonably well (**Figure 2**) and also included at least one level from each factor (**Table 1**). The differences between k = 5 and k = 6 were minimal, so in the interest of parsimony we selected k = 5 for CMP scaling. The OMs chosen under k = 5 and their cluster weights were OM 37 ($\omega_1 = 0.33$), OM 14 ($\omega_2 = 0.25$), OM 53 ($\omega_3 = 0.20$), OM 31 ($\omega_4 = 0.17$) and OM 89 ($\omega_5 = 0.05$).

The five cluster centres were well separated across the full range of biomasses, and the clusters themselves were not too spread out. The TADpole algorithm uses a non-standard loss function when optimizing clusters, and did not produce a measure of within cluster and between cluster variance, which is more interpretable for understanding cluster performance. We approximated between cluster variance by calculating pair-wise root-mean-squared-error (rMSE) between cluster medoids, and within cluster variance by calculating the average of the pairwise rMSEs for all OMs in a cluster. Average within cluster rMSE ranged within 42.91 kt – 62.34 kt, with a mean rMSE of 52.03 kt, and between cluster rMSE ranged within 111.95 kt – 556.87 kt, with a mean rMSE of 297.21 kt.

3.2 Fits of Delay Difference AMs to historical data

Fitting the delay difference assessment model to operating model biomass faithfully reproduced the historical operating model biomass series (**Figures 3** and **4**). In contrast, assessment model estimates of biological parameters and equilibria, and operating model estimates of the same quantities, were less consistently similar (**Figure 5**). While the log-normal prior on unfished biomass worked well to keep $B_{0,s}$ estimates close to the operating model values, the model equilibria were often biased. For example, in AMs 31 and 37, the DD model F_{MSY} values for the east spawning stock were around 0.30, in contrast with the operating model values, which were near 0.10. Similarly, in the same assessment models, B_{MSY} values were much higher than the corresponding operating model values. These biases have implications for the performance of MPs that rely on the AM reference point estimates to set TACs, which we describe below.

3.3 Performance in projections

3.3.1 Delay Difference Model-based CMPs

CMP performance varied for the two stocks/areas (**Figure 6**). Uncapped procedures generally produced much higher catches from the east area while procedures that set U_{max} to F_{MSY} tended to produce higher catches from the west area. In the east, the uncapped procedure that linked U_{max} to F_{MSY} produced an overharvested spawning stock (median Br30 < 1) with LD values lower than 0.05 in 6% of OMs, and produced a median Br30 value near 1.0 in the west, but with LD < 0.05 in 18% of OMs. CMPs that capped the TAC at MSY tended to be too conservative (i.e., most Br30 values were above 1) for the east stock, while most untuned model-based CMPs were too conservative for the west stock.

The CMPs that linked U_{max} to M with no cap in the east, and a MSY cap in the west (DD-MixCapFM, DD-MixCapFM-last10) had the best overall performance. Br30 values were close to 1 for the east (MED) spawning stock for nearly all OMs, while the median Br30 value for the west (GOM) spawning stock was also close to 1. Additionally, neither stock crashed for any OM under either of the MixCap CMPs (lowest depletion ranged from 0.09-0.57 with median 0.22).

Recruitment appeared to be the main OM factor affecting MP performance (**Figure 7**). MPs generally performed well (i.e., Br30 close to 1) on OMs with low, constant steepness for both stocks (factor level 2). In contrast, MPs crashed most often under OMs in which steepness declines after 10 projection years (factor level 3). To further explore this result, we used a meta-modelling approach to estimate the effect of each factor level on Br30 values (**Figure 12**), fitting a generalized linear model with OM factors as the observations, and Br30 values as the response. Regression coefficients for recruitment factor level 3 were negative for all MPs, indicating that MPs performed relatively worse for OMs with shifting recruitment in the projections.

The poor performance of the CMPs that link U_{max} to F_{MSY} is caused partly by the complexity mismatch between the OM and AM, biasing the delay difference reference points, with F_{MSY} positively biased and B_{MSY} negatively biased. Because of this, although some AM estimates of current and projected biomasses were often very close to the OM biomass, the biomass relative to B_{MSY} was above the upper control point, so the positively biased maximum target harvest rate was applied, leading to overfishing.

To further explore the influence of assessment model bias on CMP performance, we calculated the relative error in biomass estimates between the OM and the top AIC-ranked assessment model at each assessment interval (**Figure 9**). Significant errors in biomass estimates exist, even for OMs which were used to tune the assessment models – only OM 31 in the west area was close to unbiased on average. Larger bias occurs because AIC does not always select the "correct" AM (i.e., when running a CMP on OM 89, the AM that was tuned to OM 37 often has the lower AIC score than the AM that was tuned to OM 89). As expected, there was a negative relationship between relative error in biomass estimates and Br30, with CMPs that under-estimate spawning biomass producing higher Br30 values, on average (**Figure 10**).

3.3.2 Empirical CMPs

All empirical CMPs were prone to over-estimate biomass for the east and, to a lesser extent, west spawning stocks. As a result, uncapped CMPs targeting $U_{max} = F_{MSY}$ (Emp-NoCap, Emp-NoCapB0) tended to over-harvest regardless of the upper control point used, producing Br30 values less than 1.0 for 97% of OMs in the east, and around 72% of OMs in the west, with both stocks crashing a large proportion of OMs (**Figure 11**). In contrast, capped CMPs targeting $U_{max} = F_{MSY}$ (Emp-msyCap, Emp-msyCapB0) produced median Br30 values very close to 1.0 in the east, although the east spawning stock crashed (LD < 0.05) in about 12% of OMs, while tending to under-harvest in the west.

To compensate for positive bias in biomass estimates, a very modest harvest rate ($\lambda = 0.3$) was required to avoid over-harvesting when there was no TAC cap (compare Emp-NoCapB0 to Emp-NoCapFMB0, Figure 11). Fishing without a cap but at the lower harvest rate produced Br30 values very close to 1.0, and while some OMs still had crashed stocks, the proportion was much lower than for CMPs with higher target harvest rates. Applying the same lower harvest rates to capped procedures increased Br30 values, as expected, but were too conservative for both east and west stocks (Emp-MsyCapFMB0).

In general, catches under the CMPs that did not react to spawner index trends tended to group according to cap in the east, but were more variable in the west. East catches were for uncapped procedures were higher in the first 20 years, but then dropped as biomass declined in years 21-30. In contrast, capped TACs were more similar across the first 30 years, with median values around 50kt (the OM cluster weighted average of MSY) for the first 20 years for all capped CMPs. In years 21-30, the capped CMPs targeting Fmsy (Emp-msyCap, Emp-msyCapB0) continued with median catches around 50kt, while the Emp-msyCapFMB0 CMP with the lower target harvest rates dropped median catches to around 35kt.

Trend-based adjustments to TAC had a minor effect on Br30 values, but a more noticeable effect on catches. When OM cluster weights were adjusted in response to increasing or decreasing trends (Emp-NoCapFMB0TrendWtd) the median Br30 value increased by around 0.02-0.04, and the range of Br30 values contracted by abount 0.1 Br30 units, indicating a small refinement. In contrast, when CMPs only adjusted TACs in response to decreases in spawn indices (Emp-NoCapFMB0TrendDwn), there were only slight differences to the trend-independent version of the CMP, with no real difference in range, but a shift upwards of around 0.1 Br30 units (**Figure 11**). For the trend-based CMPs, median and minimum catches were lower in the first 10 years term, while median and minimum catch increased slightly in years 21-30, indicating that biomass was slightly higher, given the increased precaution from the trend-based adjustments.

As with the model-based procedures, recruitment appeared to be the most important uncertainty axis (**Figures 12** and **13**). Unlike the model-based procedures, however, the OMs associated with the lowest empirical Br30 values were those with recruitment level 2.

3.3.3 Tuning Model-based CMPs

The DD-MixCapFM-last10 CMP was chosen for development tuning to the three pairs of biomass targets given in **Table 6**. The grid search of east area harvest rate multipliers λ found $U_{max} = 1.13M$ produced a median $Br30 = 1.0 \pm 0.01$ (**Figure 14**, East). Holding the east area harvest rates at 1.13M, a grid search was then conducted over west area λ values. The initial grid search of west area harvest rates did not produce west stock Br30 values within the desired precision (0.02), but the Br30 values produced by the grid search bounded the three target values. Proposed λ values were derived from the grid search results via interpolation, which produced the desired precision after one iteration (**Figure 14**, West).

Given the stronger effect of east area harvest rates on the GOM stock than west area harvest rates on the MED stock, there is considerably more risk of overfishing associated with lower Br30 target values for the west stock (**Figure 15**). Targeting Br30 = 1.0 for the west stock is likely inadvisable, as the range of Br30 values was 0.05 - 2.3 for the GOM spawning stock, in contrast to the MED stock which ranged in 0.4 - 1.95 across all three west stock targets. Similarly, the lowest depletion for the GOM stock was around 1.4% of unfished biomass for the CMP targeting Br30 = 1, as opposed to 11% for the MED stock, under the lfrDD_001 CMP. Reducing maximum harvest rates in the West area move Br30 and LD distributions away from 0, but there is still considerable risk of producing overharvested stocks. For example, the fifth percentile of LD in the East is around 14% of unfished, while the same percentiles of LD in the west are 3% (Br30 = 1), 8% (Br30 = 1.25) and 9% (Br30 = 1.5) of unfished. However, in order to achieve commensurate risk between east and west stocks, median average TACs in the west area will likely hover around 500t for the foreseeable future (**Figure 15**).

4. Discussion

In this paper, we tested empirical and model-based candidate management procedures (CMPs) for Atlantic Bluefin Tuna. These CMPs were based on multi-model inference, where TACs from 5 sub-procedures, each scaled to a specific reference grid OM, were combined as a weighted average to produce TAC advice, with weights either taken from an initial cluster analysis of historic operating model biomasses (empirical) or based on Akaike's Information Criterion (model-based).

Multi-model CMPs are a viable option for managing Atlantic Bluefin Tuna fisheries across a wide range of operating model hypotheses. Clustering OMs by biomass to choose a small representative subset for sub-procedure scaling produced CMPs with reasonable performance according to the Br30 metric. Indeed, with only a small amount of tuning, the median Br30 values across the full operating model grid were at or above Br30 = 1.0 for all capped model-based and empirical CMPs. All uncapped CMPs were, as expected, more aggressive, which crashed stocks more often.

Development tuning showed that the asymmetry in the effects of area harvest rates on each stock are likely the biggest hurdle to achieving management targets. Achieving commensurate (i.e., within 5 percentage points) fifth percentiles of lowest depletion levels between stocks requires a much lower target harvest rate in the west area than in the east area, as a large proportion of west stock fishing mortality comes from east area fishing given the order of magnitude difference in TACs, despite the low prevalence of west stock fish in the east area.

Trend-based adjustments to empirical CMPs showed some promise in reducing Br30 variance for the East and West stocks. Current implementations of the trend-based adjustments reduce west area TACs when the west stock index is reduced, or east area TACs in response to east stock indices. However, there may be potential to refinine west stock Br30 distributions via east area TAC adjustments in response to west stock trends given the asymmetry in area TAC effects on each stock's biomass. If so, achieving lower west stock Br30 targets may be less risky than we have found here.

There was no evidence that scaling model-based sub-procedures by fitting to historical OM biomass bestowed any kind of omniscience to the model-based CMPs, nor did it cause the CMPs to "overfit" to the OM grid. The AMs only fit to biomass from the scaling set of OMs in the historical period, as a method of correctly scaling the catchability parameters for the set of management indices, and any updated operating model information was not fed to the AMs during the projection period. Further, the performance of model-based CMPs on the set of OMs outside of the selected five scaling OMs constitutes a form of hold-out analysis similar to a cross validation, and while CMP performance was acceptable on the full grid, it was still quite variable over combinations of OM grid factors that were not included in the sub-procedure scaling. Further, with the exception of the GOM spawning stock under OM 31, the weighted estimates of spawning stock biomass combined from each sub-procedure at each time step were on average biased on the scaling set of OMs, with mean relative errors around 50% in absolute value, showing that omniscience was not bestowed even for a "self-test" of the CMPs on the scaling OM subset.

There may be some limitations associated with the method used to select the scaling set of OMs. Here, each OM was represented by a concatenated East/West historical biomass in the clustering algorithm, and the number of clusters was chosen to include each OM factor level at least once. The concatenation may have caused problems, where the higher variance in the East required a larger number of clusters than the more concentrated West biomasses. Further, clustering by biomass is only one dimension in which the operating models vary, and another choice may yield a different set. However, based on preliminary results not presented here where we selected particular OMs from the grid to ensure some of the more extreme scenarios were considered, it appears that conservation outcomes (as measured by the Br30 metric) are not very sensitive to scaling set choice, as long as the scaling set of operating models is in some sense representative (either of biomass history or over the OM grid factor

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Table 1. Unfished and fished equilibrium quantities for the delay-difference population dynamics.

Description	Equation
Survivorship	$S^{(f)} = e^{-M-f}$
Average Weight	$\overline{w}^{(f)} = \frac{S^{(f)}\alpha + w^{(k)}(1 - S^{(f)})}{1 - \rho S^{(f)}}$
Unfished Numbers	$N = B_0 / \overline{w}^{(f=0)}$
Unfished Recruitment	$R_0 = (1 - S^{(f=0)})N_0$
Stock-Recruit	$a = \frac{4hR_0}{B_0(1-h)}, b = \frac{5h-1}{B_0(1-h)}$
Biomass	$B^{(f)}$
	$=\frac{S^{(f)}(\alpha+\rho\overline{w}^{(f)})+\overline{w}^{(f)}(aw^{(k)}-1)}{b(\overline{w}^{(f)}-\rho S^{(f)}\overline{w}^{(f)}-\alpha S^{(f)})}$
Recruitment	$R^{(f)} = \frac{aB^{(f)}}{1 + bB^{(f)}}$
Yield	$Y^{(f)} = \frac{f}{M+f} (1 - e^{-M-f}) B^{(f)}$

Table 2. Process and observation model components of the delay difference stock assessment model used in BC Sablefish management procedures. Initialisation values for biomass, numbers, and recruitment are equilibrium unfished values from Table 2.

No.	Equation
A2.1	$\theta_s = \{\log F_{s,0}, \log B_{s,0}, q_g, \tau_g, \overrightarrow{\omega}_{s,t}\}.$
A2.2	$B_{s,1} = B_s^{(f=F_{s,0})}$
A2.3	$N_{s,1} = N_s^{(f=F_{s,0})}$
A2.4	$R_{s,1} = R_s^{(f = F_{s,0})}$
A2.5	$B_{s,a,t} = p_{s,a} \cdot B_{s,t}$
A2.6	$B_{a,t} = \sum_{s} B_{s,a,t}$
A2.7	$C_{s,t} = \sum_{a}^{s} \frac{B_{s,a,t}}{B_{a,t}} C_{a,t}$
A2.8	$(R_{s,1}e^{\omega_{s,t-k}-\sigma^2/2} \qquad t \le k$
	$R_{s,t} = \begin{cases} R_{s,1} e^{\omega_{s,t-k} - \sigma^2/2} & t \le k \\ \frac{aB_{s,t-k}}{1 + bB_{s,t-k}} e^{\omega_{s,t-k} - \sigma^2/2} & t > k \end{cases}$
A2.9	$Z_{s,t-1} = M_s + F_{s,t-1}$
A2.10	$N_{s,t} = e^{-Z_{s,t-1}} N_{s,t-1} + R_{s,t}$
A2.11	$B_{s,t} = e^{-Z_{s,t-1}}(\alpha N_{s,t-1} + \rho B_{s,t-1}) + w_s^{(k)} R_{s,t}$
A2.12	$\hat{I}_{g,t} = q_g \cdot B_{s,t}$
A2.13	$\hat{I}_{g,t} = q_g \cdot B_{a,t}$
A2.14	$\log I_{g,t} \sim N(\log \hat{I}_{g,t}, \tau_g)$
A2.15	$\omega_{s,t} \sim N(0,\sigma)$

Table 3. TAC weighting adjustments under different trend-based TAC adjustments for 5 AMs. The TrendWt rule uses multipliers for original weights, $(w)_{\alpha}$, when ordered by TAC size from lowest to highest, with * defined as component-wise multiplication, while the TrendDwn rule gives the final weights for each bin.

Growth rate bin	TrendWt	TrendDwn	
$(-1,-2\theta]$	$(w)_{\alpha} *(4,2,1,.5,.25)$	(1,0,0,0,0)	
$(-2\theta, -\theta]$	$(w)_{\alpha} * (2,4,2,1,.5)$	(.5,.5,0,0,0)	
$(-\theta, \theta]$	$(w)_{\alpha} * (1,2,4,2,1)$	$(w)_{\alpha}$	
$(\theta, 2\theta]$	$(w)_{\alpha} * (.5,1,2,4,2)$	$(w)_{\alpha}$	
<u>(2θ,∞</u>]	$(w)_{\alpha} * (.25, .5, 1, 2, 4)$	$(w)_{\alpha}$	

Table 4. List of model-based (delay-difference) CMPs tested for Atlantic Bluefin Tuna.

Name	TAC Cap (East; West)	U_{max} (East; West)	AIC Weights
DD-NoCap	None; None	$F_{MSY}; F_{MSY}$	All equal
DD-MsyCap	MSY; MSY	F_{MSY} ; F_{MSY}	All equal
DD-NoCapFM	None; None	$M; \frac{2}{3}M$	All equal
DD-MsyCapFM	MSY; MSY	$M; \frac{2}{3}M$	All equal
DD-MixCapFM	None; MSY	$M; \frac{2}{3}M$	All equal
DD-MixCapFM-last10	None; MSY	$M; \frac{2}{3}M$	Last 10 years

Table 5. List of empirical CMPs tested for Atlantic Bluefin Tuna.

Name	U_{max}	UCP	TAC Cap (East; West)	Trend Rule
Emp-MsyCap	F_{MSY}	B_{MSY}	MSY; MSY	None
Emp-NoCap	F_{MSY}	B_{MSY}	None; None	None
Emp-MsyCapB0	F_{MSY}	$0.4B_{0}$	MSY; MSY	None
Emp-NoCapB0	F_{MSY}	$0.4B_{0}$	None; None	None
Emp-MsyCapFMB0	M/λ	$0.4B_{0}$	None; None	None
Emp-NoCapFMB0	M/λ	$0.4B_{0}$	None; None	None
Emp-NoCapFMB0_trWtd	M/λ	$0.4B_{0}$	None; None	TrendWtd; $\theta = 0.1$
Emp-NoCapFMB0_trDwn	M/λ	$0.4B_{0}$	None; None	TrendDown; $\theta = 0.1$

Table 6. Target Br30 values used in model-based CMP tuning, and the label of the model-based CMP tuned to those targets.

East Br30	West Br30	Label
1.0	1.0	lfrDD_001
1.0	1.25	lfrDD_002
1.0	1.5	lfrDD_002

Table 7. Median performance metrics over the reference grid of 96 deterministic OMs for the three CMPs tuned to biomass targets described in Table 6. The λ column shows the final M multiplier used to set maximum harvest rates arrived at for each CMP, while the remaining columns show median average catch over the first (C10), second (C20) and third (C30) ten year intervals of the projection (kt), biomass depletion in year 30 (D30), lowest depletion over the projection (LD), and biomass relative to spawning stock Bmsy in year 30 (Br30).

East							
CMP	λ	C10	C20	C30	D30	LD	Br30
ZeroC		0	0	0	0.85	0.46	3.41
1frDD_001	1.13	57.58	61.61	29.77	0.26	0.26	1.00
1frDD_002	1.13	57.58	62.10	29.84	0.26	0.26	1.01
1frDD_003	1.13	57.58	62.24	29.90	0.27	0.27	1.02
West							
CMP	λ	C10	C20	C30	D30	LD	Br30
ZeroC		0	0	0	0.75	0.28	2.78
1frDD_001	1.09	1.78	1.60	1.49	0.30	0.20	0.99
1frDD_002	0.52	0.96	0.90	0.94	0.38	0.22	1.24
1frDD 003	0.23	0.60	0.43	0.45	0.46	0.23	1.48

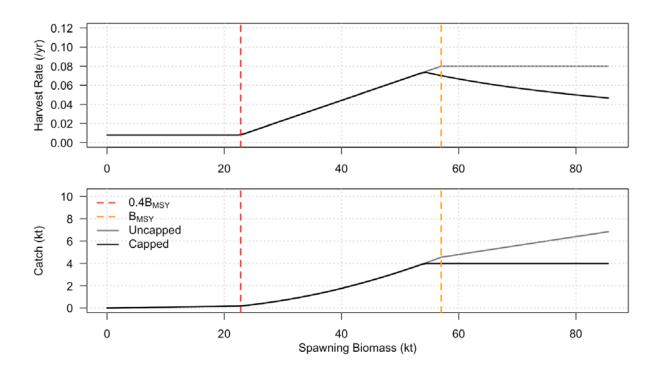


Figure 1. An example harvest control rule with maximum harvest rate $U_{\text{max}} = 0.08$, an upper control point of $B_{\text{MSY}} = 57$, and a TAC cap at 4 kt.

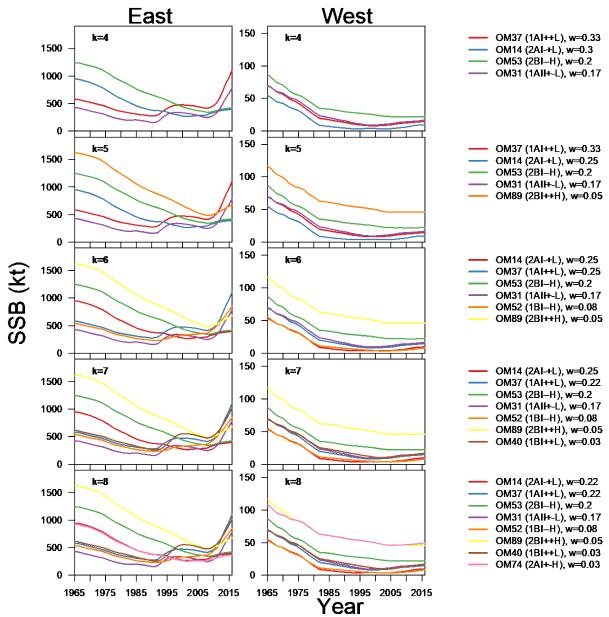


Figure 2. Spawning stock biomasses for the reduced set of 32 operating models for both the Mediterranean stock (East) and Gulf of Mexico stock (West). Rows represent different clustering analyses, each with a different number of clusters (k). Coloured lines represent cluster centers, each of which corresponds to a different OM from the reference grid.

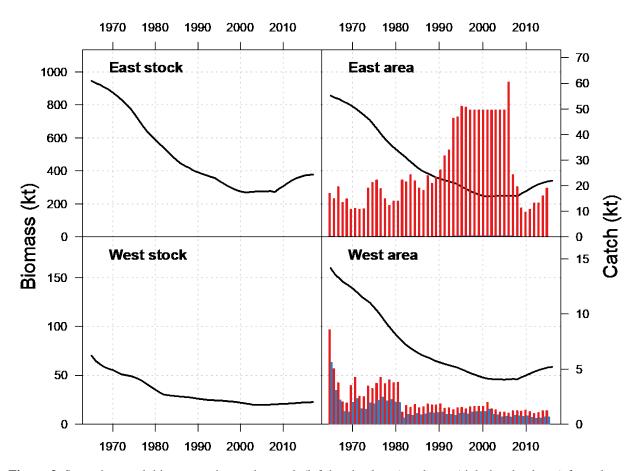


Figure 3. Spawning stock biomass estimates by stock (left hand column) and area (right hand column) from the delay difference stock assessment model fit to OM 50. Total area catch is shown split into East stock (red bars) and West stock (blue bars). Note that the catch scale is exaggerated with respect to the biomass, so that West stock catch is visible in the East area.

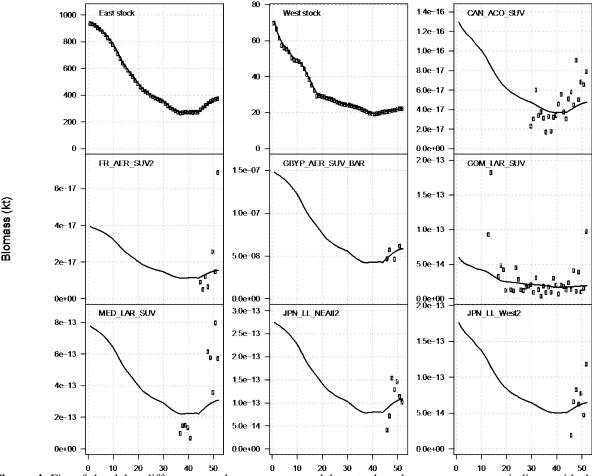


Figure 4. Fits of the delay difference stock assessment model to stock and area management indices, with the associated catchability estimates. Data are shown as circles, while the lines indicate the model biomass scaled by catchability. East Stock and West Stock indices are the spawning stock biomass estimates from operating model 50.

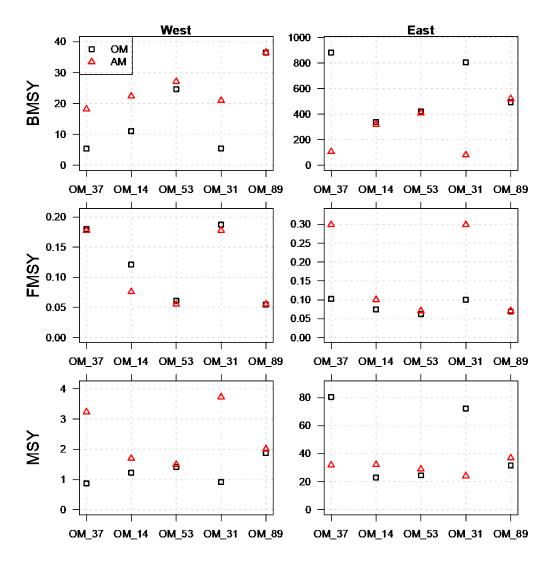


Figure 5. Comparison of B_{MSY} (top row), F_{MSY} (middle row) and MSY (bottom row) values between the operating models (black squares) and delay-difference assessment models (red triangles).

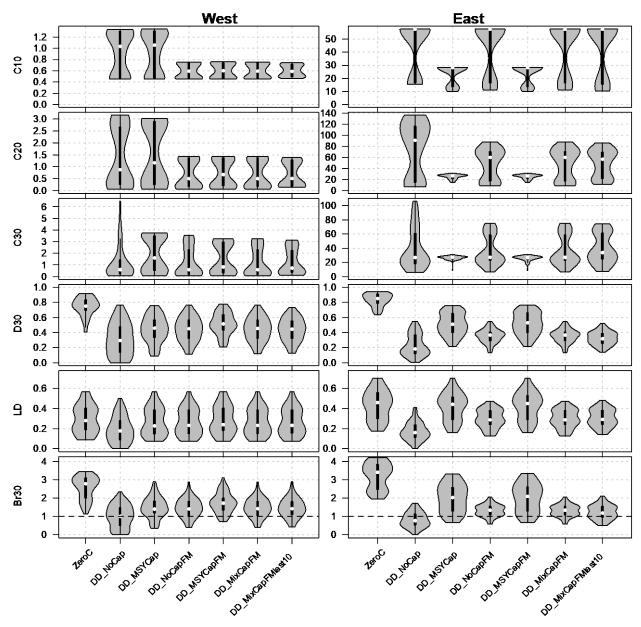


Figure 6. Violin plots of MSE performance metrics for model-based CMPs. The thin line, thick line and white circle within each "violin" represents a boxplot of values across OMs, while either side of boxplot shows a rotated kernel density plot of the distribution of values.

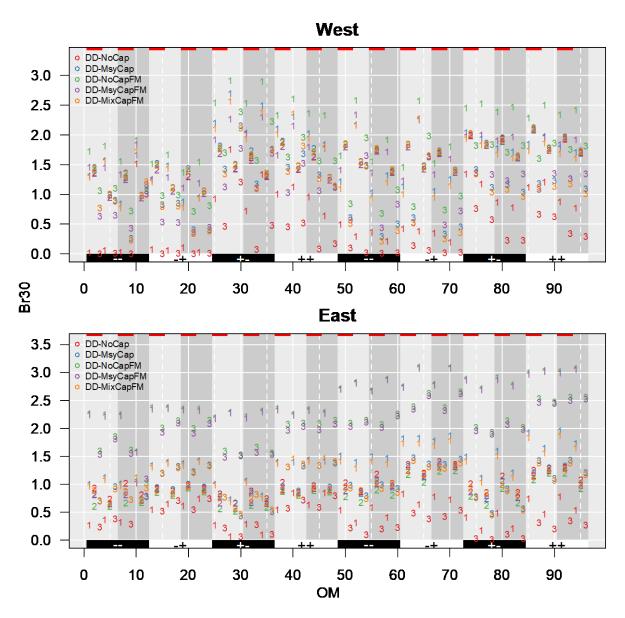


Figure 7. Estimated Br30 values for model-based MPs. Numerical values indicate the steepness factor for each OM (1=high, 2=low, 3=shifting). Younger spawning/higher M scenarios are indicated by the red line at the top of each plot. Migration factor levels are indicated by the background shading (light grey indicates low stock migration, dark grey indicates high). SSB factor levels are indicated by the black and white symbols on the bottom of each plot (-- indicates low West/low East; -+ indicates low/high; +- indicates high/low; ++ indicates high/high). The first 48 OMs have low length composition weighting while the latter 48 have high length composition weighting.

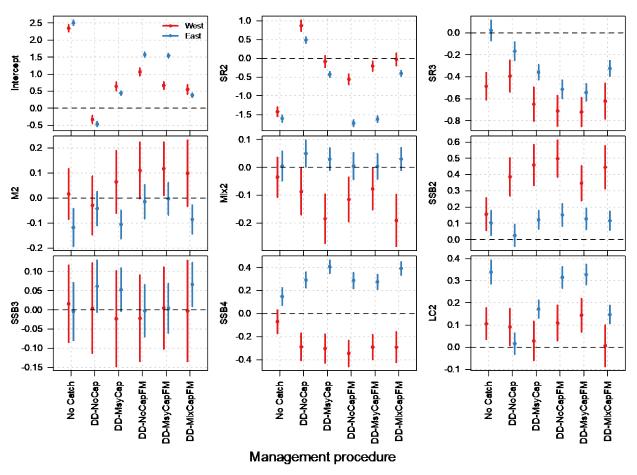


Figure 8. Coefficients from regression of Br30 on OM factor levels for model-based CMPs.

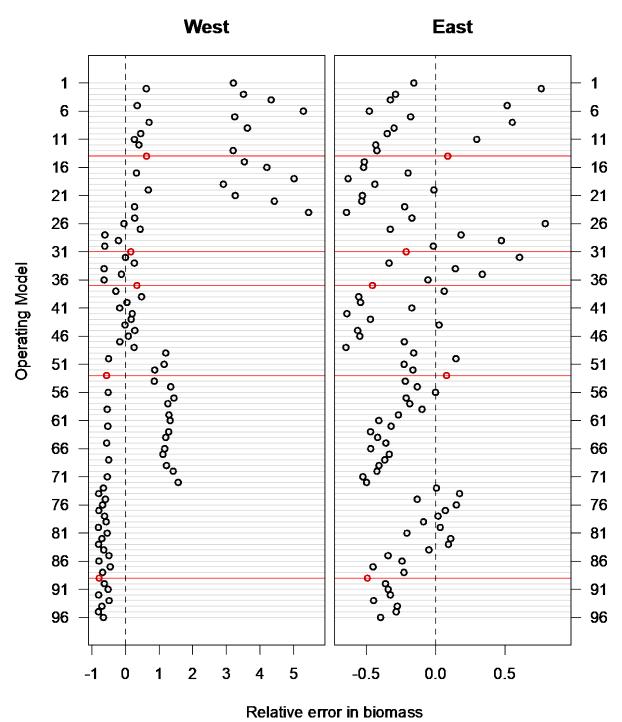


Figure 9. Average bias is biomass estimates from the top AIC-ranked model in each assessment interval for the MP_Mixed procedure. OMs used to condition the MP are coloured red.

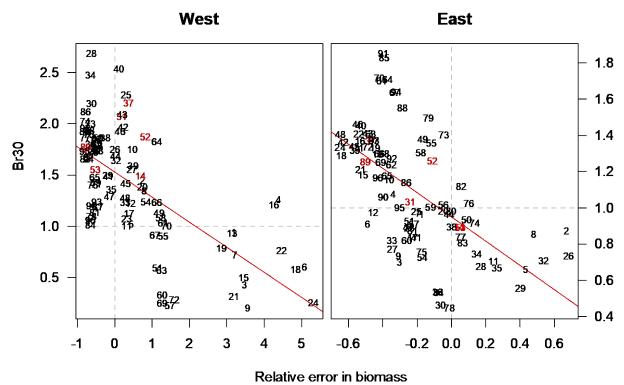


Figure 10. Relationship between Br30 and the average bias is biomass estimates from the top AIC-ranked model in each assessment interval for the MP_Mixed procedure for each of the 96 OMs. OMs used to condition the MP are coloured red.

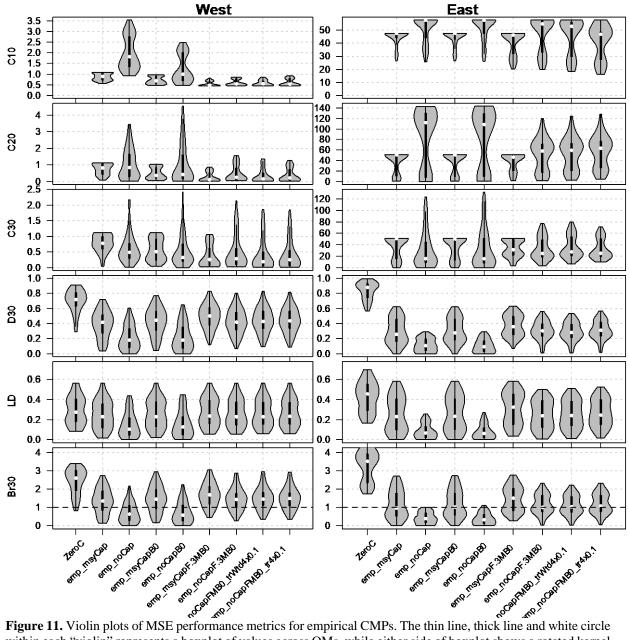


Figure 11. Violin plots of MSE performance metrics for empirical CMPs. The thin line, thick line and white circle within each "violin" represents a boxplot of values across OMs, while either side of boxplot shows a rotated kernel density plot of the distribution of values.

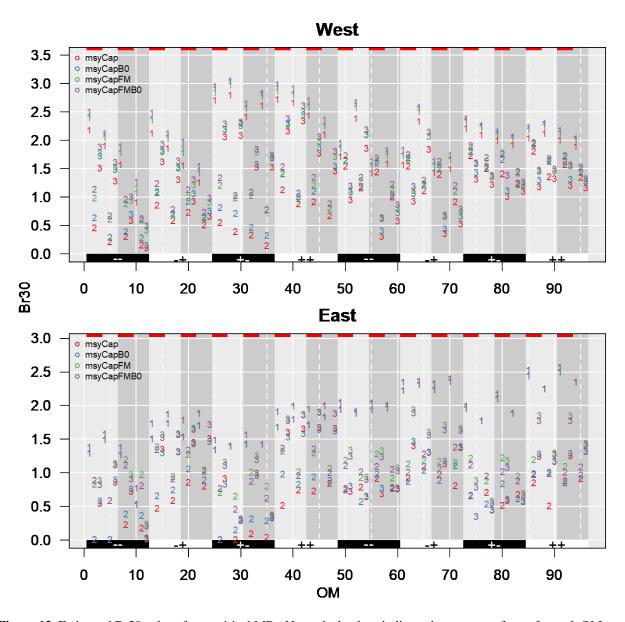


Figure 12. Estimated Br30 values for empirical MPs. Numerical values indicate the steepness factor for each OM (1=high, 2=low, 3=shifting). Younger spawning/higher M scenarios are indicated by the red line at the top of each plot. Migration factor levels are indicated by the background shading (light grey indicates low stock migration, dark grey indicates high). SSB factor levels are indicated by the black and white symbols on the bottom of each plot (-- indicates low West/low East; -+ indicates low/high; +- indicates high/low; ++ indicates high/high). The first 48 OMs have low length composition weighting while the latter 48 have high length composition weighting.

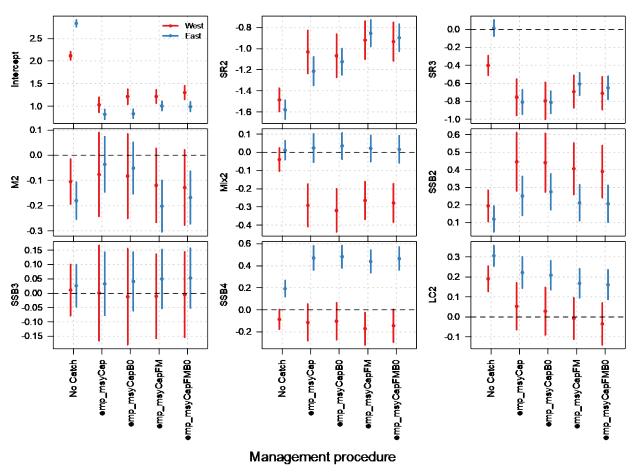


Figure 13. Coefficients from regression of Br30 on OM factor levels for empirical CMPs.

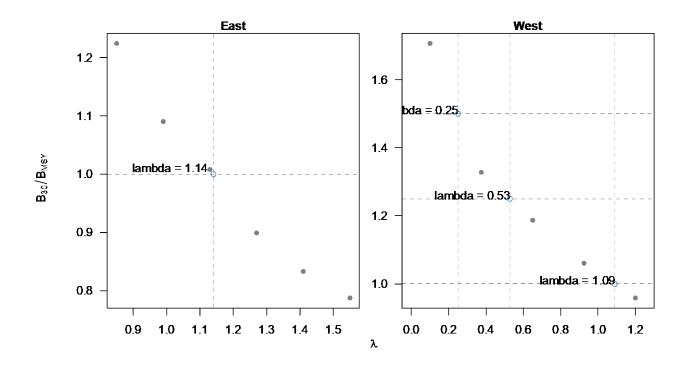


Figure 14. Responses of MED and GOM spawning stock Br30 to the grid search over λ values for the east and west areas, respectively. Points show grid search treatments (x-axis) and responses (y-axis), and dashed vertical and horizontal lines show proposed target λ values (labeled) found by interpolating the two nearest points to the target Br30 value.

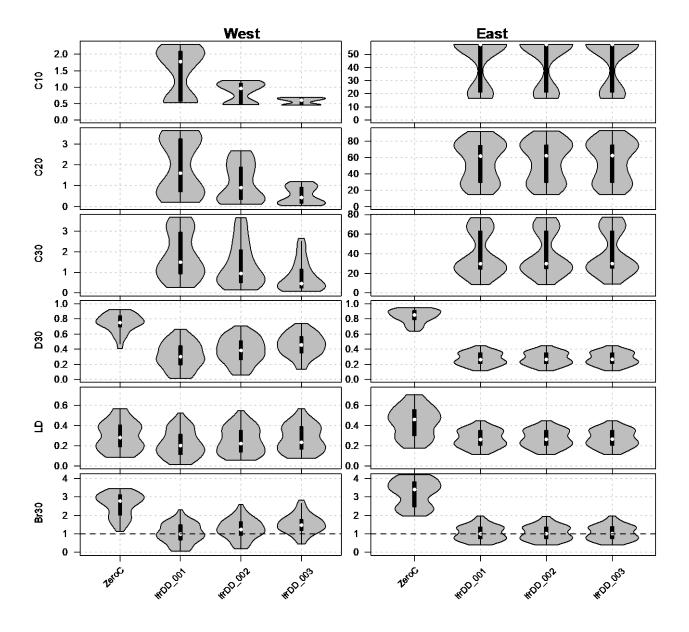


Figure 15. Violin plots of MSE performance metrics for model-base CMPs tuned to meet targets in Table 6. The thin line, thick line and white circle within each "violin" represents a boxplot of values across OMs, while either side of boxplot shows a rotated kernel density plot of the distribution of values.