

PHYS 5640 - HW 4

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1 Dual Markov chain Monte Carlo

(1) Show that the probability $P(C \rightarrow C')$ is given by

$$P(C \rightarrow C') = \sum_G P(C \rightarrow G) \tilde{P}(G \rightarrow C') \quad (1)$$

We can show that Eq. 1 is true by considering the possible paths from C to C' . Since in the dual MC formulation there is no path from C to C' within the C space, then one must travel from C to C' through an intermediary state in the G space. One possible path is to travel from state C to state G to state C' . The probability of traversing this path is simply the probability of going from state C to G and then from state G to C' , expressed as:

$$P(C \rightarrow G \rightarrow C') = P(C \rightarrow G)P(G \rightarrow C').$$

This is just a single path from C to C' however. Since there is a path from C to all states in the G space, then to compute $P(C \rightarrow C')$, we need to consider all of these possible paths. Since we could have taken the single path in the example above or any other path from C to any G in the G space, and each of these paths is equiprobable, the final probability $P(C \rightarrow C')$ is given by

$$P(C \rightarrow C') = \sum_G P(C \rightarrow G) \tilde{P}(G \rightarrow C')$$

(2) Show that the strong detailed balance condition (for every pair of C and G)

$$\pi(C)P(C \rightarrow G) = \pi(G)\tilde{P}(G \rightarrow C) \quad (2)$$

implies the detailed balance of the original Markov-chain process:

$$\pi(C)P(C \rightarrow C') = \pi(C')P(C' \rightarrow C). \quad (3)$$

To prove Eq. 3, we simply start with the left hand side of Eq. 3 and use Eq. 1 and Eq. 2 to transform it to the right hand side. We use Eq. 1 on the left hand side of Eq. 3 to find

$$\pi(C)P(C \rightarrow C') = \pi(C) \sum_G P(C \rightarrow G) \tilde{P}(G \rightarrow C') = \sum_G \pi(C)P(C \rightarrow G) \tilde{P}(G \rightarrow C').$$

Invoking Eq. 2 twice we have

$$\sum_G \pi(C) P(C \rightarrow G) \tilde{P}(G \rightarrow C') = \sum_G \pi(G) \tilde{P}(G \rightarrow C) \tilde{P}(G \rightarrow C') = \sum_G \pi(C') P(C' \rightarrow G) \tilde{P}(G \rightarrow C).$$

Finally, using Eq. 1

$$\sum_G \pi(C') P(C' \rightarrow G) \tilde{P}(G \rightarrow C) = \pi(C') \sum_G P(C' \rightarrow G) \tilde{P}(G \rightarrow C) = \pi(C') P(C' \rightarrow C).$$

This then gives

$$\boxed{\pi(C) P(C \rightarrow C') = \pi(C') P(C' \rightarrow C).}$$

(3) Finally, show that the strong detailed balance condition Eq. (3) is satisfied with the following choice of transition probabilities

$$P(C \rightarrow G) = \frac{w(C, G)}{w(C)}, \quad \tilde{P}(G \rightarrow C) = \frac{w(C, G)}{w(G)} \quad (4)$$

This is easy to show by simply plugging the chosen probabilities from Eq. 4 into Eq. 2:

$$\pi(C) P(C \rightarrow G) = \pi(G) \tilde{P}(G \rightarrow C)$$

Plugging in from Eq. 4 on the left hand side we have

$$\pi(C) P(C \rightarrow G) = \pi(C) \frac{w(C, G)}{w(C)}.$$

Then, multiplying this equation by $\pi(G)/\pi(G) = 1$ and using the fact that $\pi(C) = w(C)/Z$ and $\pi(G) = w(G)/Z$ we have

$$\pi(C) \frac{\pi(G)}{\pi(G)} \frac{w(C, G)}{w(C)} = \pi(G) \frac{w(C)}{Z} \frac{Z}{w(G)} \frac{w(C, G)}{w(C)} = \pi(G) \frac{w(C, G)}{w(G)}.$$

And then finally, comparing the above to Eq. 3, we find that a choice of $\tilde{P}(G \rightarrow C) = w(C, G)/w(G)$ will make the above consistent with detailed balance (Eq. 3). So the two choices of P and \tilde{P} from Eq. 4 satisfy detailed balance.

2 Swendsen-Wang algorithm

(1) Show that the partition function of the Ising model can be expressed as

$$Z = \sum_C \sum_G w(C, G) = e^{N_b \beta J} \sum_{\{\sigma_i\}} \sum_{\{n_{ij}\}} \prod_{\langle ij \rangle} w_{ij}(\sigma_i, \sigma_j; n_{ij}) \quad (5)$$

We begin by expressing Z as

$$Z = \text{Tr} (e^{-\beta \mathcal{H}}) = \sum_{\{\sigma_i\}} e^{-\beta \mathcal{H}}$$

where

$$\mathcal{H} = -J \sum_{\langle ij \rangle} \sigma_i \sigma_j.$$

We can reexpress $\exp(-\beta\mathcal{H})$ as

$$e^{\beta J \sum_{\langle ij \rangle} \sigma_i \sigma_j} = \prod_{\langle ij \rangle} e^{\beta J \sigma_i \sigma_j}.$$

We then consider cases on σ_i and σ_j , arriving at the following values for $\exp(\beta J \sigma_i \sigma_j)$:

$$e^{\beta J \sigma_i \sigma_j} = \begin{cases} e^{\beta J}, & \sigma_i = \sigma_j \\ e^{-\beta J}, & \sigma_i \neq \sigma_j \end{cases} = e^{\beta J} \delta_{\sigma_i, \sigma_j} + e^{-\beta J} (1 - \delta_{\sigma_i, \sigma_j}).$$

Rearranging terms we have

$$e^{\beta J \sigma_i \sigma_j} = e^{\beta J} [(1 - e^{-2\beta J}) \delta_{\sigma_i, \sigma_j} + e^{-\beta J}] = e^{\beta J} [(1 - p) + p \delta_{\sigma_i, \sigma_j}]$$

where $p = 1 - \exp(-2\beta J)$ and $1 - p = \exp(-2\beta J)$. Introducing the variable n_{ij} where $n_{ij} = 0$ indicates the lack of a bond between spins and $n_{ij} = 1$ indicates the presence of a bond between spins at points i and j we have

$$e^{\beta J \sigma_i \sigma_j} = \sum_{\{n_{ij}\}} e^{\beta J} [(1 - p) \delta_{n_{ij}, 0} + p \delta_{\sigma_i, \sigma_j} \delta_{n_{ij}, 1}].$$

Finally, this allows us to express the partition function as

$$Z = \sum_{\{\sigma_i\}} \prod_{\langle ij \rangle} \sum_{\{n_{ij}\}} e^{\beta J} [(1 - p) \delta_{n_{ij}, 0} + p \delta_{\sigma_i, \sigma_j} \delta_{n_{ij}, 1}].$$

We can rearrange the sum over bond variables and the product over nearest neighbors and also pull out the constant factor of $\exp(\beta J)$:

$$\prod_{\langle ij \rangle} e^{\beta J} = e^{\sum_{\langle ij \rangle} \beta J} = e^{N_b \beta J}$$

where $N_b = 2N$ is the number of bonds. This gives finally

$$Z = e^{N_b \beta J} \sum_{\{\sigma_i\}} \sum_{\{n_{ij}\}} \prod_{\langle ij \rangle} w_{ij}(\sigma_i, \sigma_j; n_{ij})$$

where

$$w_{ij}(\sigma_i, \sigma_j; n_{ij}) = (1 - p) \delta_{n_{ij}, 0} + p \delta_{\sigma_i, \sigma_j} \delta_{n_{ij}, 1} = \begin{cases} 1 - p, & n_{ij} = 0 \\ p, & n_{ij} = 1, \sigma_i = \sigma_j \\ 0, & n_{ij} = 1, \sigma_i \neq \sigma_j \end{cases}.$$

(2) Show that the local transition probability of connecting the bonds is

$$P(\sigma_i, \sigma_j \rightarrow n_{ij} = 1) = \begin{cases} p, & \sigma_i = \sigma_j \\ 0, & \sigma_i \neq \sigma_j \end{cases} \quad (6)$$

To show the above, we simply use Eq. 4 with $C = \sigma_i, \sigma_j$ and $G = n_{ij}$:

$$P(\sigma_i, \sigma_j \rightarrow n_{ij} = 1) = \frac{w_{ij}(\sigma_i, \sigma_j; 1)}{\sum_{\{n_{ij}\}} w_{ij}(\sigma_i, \sigma_j; n_{ij})}.$$

Now, we consider cases on the values of σ_i and σ_j . Since the value of w_{ij} for $n_{ij} = 1$ depends only on whether σ_i and σ_j are the same or different, then we'll consider those two cases. If $\sigma_i = \sigma_j$ we have

$$P(\sigma_i, \sigma_j \rightarrow n_{ij} = 1) = \frac{p}{1 - p + p} = p.$$

If $\sigma_i \neq \sigma_j$ we have

$$P(\sigma_i, \sigma_j \rightarrow n_{ij} = 1) = \frac{0}{1 - p + p} = 0.$$

This then tells us that

$$P(\sigma_i, \sigma_j \rightarrow n_{ij} = 1) = \begin{cases} p, & \sigma_i = \sigma_j \\ 0, & \sigma_i \neq \sigma_j \end{cases}$$

(3) Show that the transition probability of assigning spins for a given bond variable is

$$\tilde{P}(n_{ij} \rightarrow \sigma_i, \sigma_j) = \begin{cases} \frac{1}{4}, & n_{ij} = 0 \\ \frac{1}{2}\delta_{\sigma_i, \sigma_j}, & n_{ij} = 1 \end{cases}. \quad (7)$$

Similarly to above, we use Eq. 4 again giving

$$\tilde{P}(n_{ij} \rightarrow \sigma_i, \sigma_j) = \frac{w_{ij}(\sigma_i, \sigma_j; n_{ij})}{\sum_{\{\sigma_i\}} w_{ij}(\sigma_i, \sigma_j; n_{ij})}.$$

We then consider cases on values of n_{ij} . First, for $n_{ij} = 0$ we have

$$\tilde{P}(0 \rightarrow \sigma_i, \sigma_j) = \frac{w_{ij}(\sigma_i, \sigma_j; 0)}{\sum_{\{\sigma_i\}} w_{ij}(\sigma_i, \sigma_j; 0)}.$$

Since $w_{ij} = 1 - p$ when $n_{ij} = 0$ regardless of the values of σ_i and σ_j we have

$$\tilde{P}(0 \rightarrow \sigma_i, \sigma_j) = \frac{1 - p}{4(1 - p)} = \frac{1}{4}.$$

Now, for $n_{ij} = 1$ we have

$$\tilde{P}(0 \rightarrow \sigma_i, \sigma_j) = \frac{w_{ij}(\sigma_i, \sigma_j; 1)}{\sum_{\{\sigma_i\}} w_{ij}(\sigma_i, \sigma_j; 1)}.$$

Since w_{ij} does depend on the values of σ_i and σ_j when $n_{ij} = 1$, we will consider cases on $\sigma_i = \sigma_j$ and $\sigma_i \neq \sigma_j$. First, if $\sigma_i = \sigma_j$ we have

$$\tilde{P}(0 \rightarrow \sigma_i, \sigma_j) = \frac{p}{p + p + 0 +} = \frac{1}{2}.$$

If $\sigma_i \neq \sigma_j$, then

$$\tilde{P}(0 \rightarrow \sigma_i, \sigma_j) = \frac{0}{p + p + 0 +} = 0.$$

This gives finally

$$\tilde{P}(n_{ij} \rightarrow \sigma_i, \sigma_j) = \begin{cases} \frac{1}{4}, & n_{ij} = 0 \\ \frac{1}{2}\delta_{\sigma_i, \sigma_j}, & n_{ij} = 1 \end{cases}.$$

The results of 2.2 and 2.3 make sense in the context of the Swendsen-Wang algorithm. Eq. 6 tells us that if two neighboring spins are aligned, then they should be bonded with probability p . If they are not, then they should never be bonded. This allows us to create a lattice of bonded similarly aligned spins in the G space. Eq. 7 tells us that if a bond is not present between two spins, then we can assign the values of those spins arbitrarily; if the spins are not already in the cluster, then the two spins can take on values $\uparrow\uparrow$, $\downarrow\downarrow$, $\uparrow\downarrow$, or $\downarrow\uparrow$ with equal probability. If there is a bond present, then there is a restriction on the values of the spins; for a bond to be present, the spins must be aligned. But the specific orientation of the spins does not matter, and Eq. 7 tells us that we can choose either $\uparrow\uparrow$ or $\downarrow\downarrow$ each with the same probability for the values of σ_i and σ_j . This interpretation of Eq. 6 and Eq. 7 tells us how to proceed in our implementation of the Swendsen-Wang algorithm when proposing a configuration in G from a configuration in C or when proposing a configuration in C from a configuration in G .