

CS489/698: Intro to ML

Lecture 02: Linear Regression



Outline

Announcements

Linear Regression

Regularization

Cross-validation



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Linear Regression

Regularization



Announcements

- Assignment 1 is out.
 - Due in two weeks

TA office hour?

- Enrollment
 - CS698: permission numbers sent
 - CS489: ~10 seats available on Quest, ask CS advisors!



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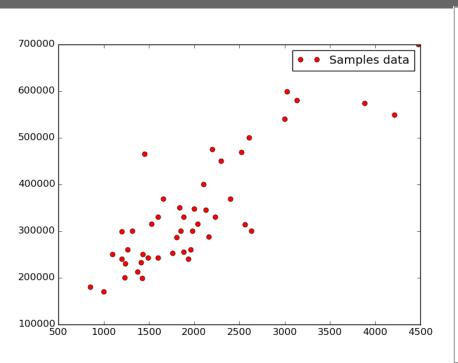
Linear Regression

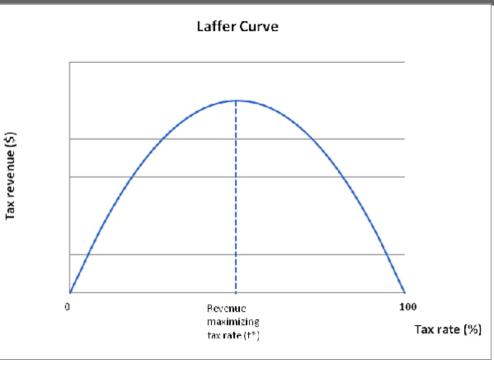
Regularization

Cross-validation



How much should I bid for?





Interpolation vs. Extrapolation

Yao-Liang Yu

Linear vs. Nonlinear



Regression

Given a pair (X, Y), find function f such that

$$f(X) \approx Y$$

- X: feature vector, d-dim real vector
- Y: response, m-dim real vector (m=1 say)

- Two problems:
 - (X,Y) is uncertain: samples from an unknown distribution

Yao-Liang Yu

How to measure the error: need a loss function



Risk minimization

Minimize the expected loss, aka risk

$$\min_{f:\mathcal{X}\to\mathcal{Y}} \mathbf{E}[L(f(X),Y)] \qquad L:\mathcal{Y}\times\mathcal{Y}\to\mathbf{R}_+$$
$$L(y,y)\equiv 0$$

- Which loss to use?
 - Not always clear; convenience dominates for now
- Least squares: $\min_f \mathbf{E} \|f(X) Y\|_2^2$



The regression function

$$\mathbf{E}\|f(X) - Y\|_{2}^{2} = \mathbf{E}\|f(X) - \mathbf{E}(Y|X)\|_{2}^{2} + \mathbf{E}\|\mathbf{E}(Y|X) - Y\|_{2}^{2}$$

Inherent noise variance

Regression function

$$f^{\star}(X) = m(X) = \mathbf{E}(Y|X)$$

Many ways to estimate m(X)

Simplest: Let's assume it is linear (affine)!



Linear regression

Assumption:
$$m(X) = \mathbf{E}(Y|X) = XA + \mathbf{b}$$

- Dream: $\min_{A,\mathbf{b}} \mathbf{E} ||XA + \mathbf{b} Y||_2^2$ distribution unknown...
- Law of Large Numbers: $\frac{1}{n} \sum_{i=1}^{n} Z_i \to \mathbf{E}(Z)$
- empirical risk empirical risk $\min_{A,\mathbf{b}} \frac{1}{n} \sum_{i=1}^{n} \|X_i A + \mathbf{b} Y_i\|_2^2$



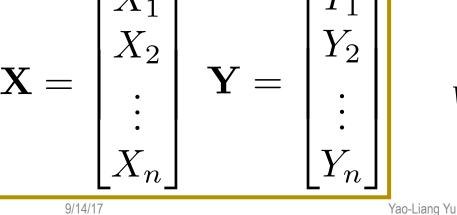


Simplification, again

$$W \leftarrow \begin{pmatrix} A \\ \mathbf{b} \end{pmatrix}$$
 $X_i \leftarrow (X_i, 1)$

$$\min_{A, \mathbf{b}} \frac{1}{n} \sum_{i=1}^{n} \|X_i A + \mathbf{b} - Y_i\|_2^2$$

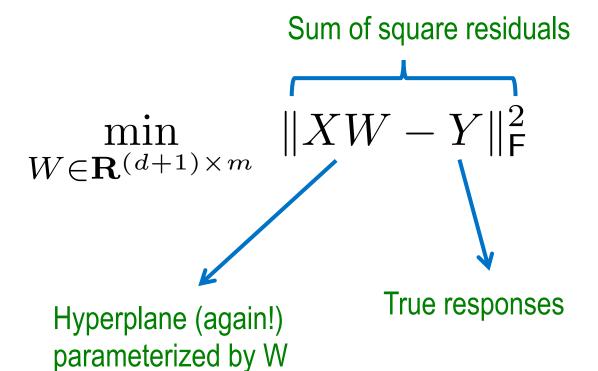
$$\min_{W} \ \frac{1}{n} \sum_{i=1}^{n} \|X_{i}W - Y_{i}\|_{2}^{2}$$

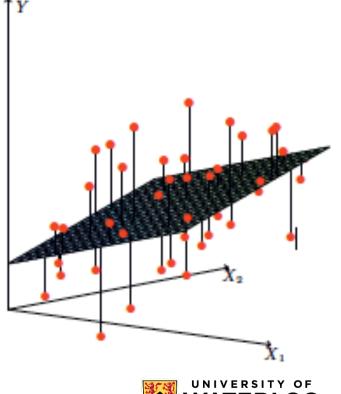






Finally







Why least squares?

$$\min_{W \in \mathbf{R}^{(d+1) \times m}} \|XW - Y\|_{\mathsf{F}}^2$$

Theorem (Sondermann'86; Friedland and Torokhti'07; Yu and Schuurmans'11)

Among all minimizers of min_W ||AWB - C||_F, W=A⁺CB⁺ is the one that has minimal F-norm.

Pseudo-inverse A⁺ is the unique matrix G such that AGA = A, GAG = G, $(AG)^T = AG$, $(GA)^T = GX$

Singular Value Decomposition

$$A=USV^T$$
 $A^+=VS^{-1}U^T$



Optimization detour

$$\min_{x} f(x)$$

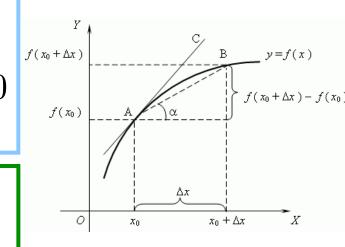
Fermat's Theorem. Necessarily Df(x) = 0

(Fréchet) Derivative at x.

$$\lim_{0 \neq \delta \to 0} \frac{|f(x+\delta) - f(x) - Df(x)\delta|}{|\delta|} = 0$$

Example.
$$f(\mathbf{x}) = \mathbf{x}^T A \mathbf{x} + \mathbf{x}^T \mathbf{b} + c$$

$$Df(\mathbf{x}) = (A + A^T) \mathbf{x} + \mathbf{b}$$





Solving least squares

$$\min_{W \in \mathbf{R}^{(d+1) \times m}} \|XW - Y\|_{\mathsf{F}}^2 = W^\top (X^\top X)W - 2W^\top XY + Y^\top Y$$

$$X^\top XW = X^\top Y$$
 Normal Equation

X^TX may not be invertible, but there is always a solution

• Even invertible, never ever compute $W = (X^TX)^{-1}XY$!

Instead, solve the linear system



Prediction

Once have W, can predict

$$\hat{Y} = X_{\text{test}} W$$

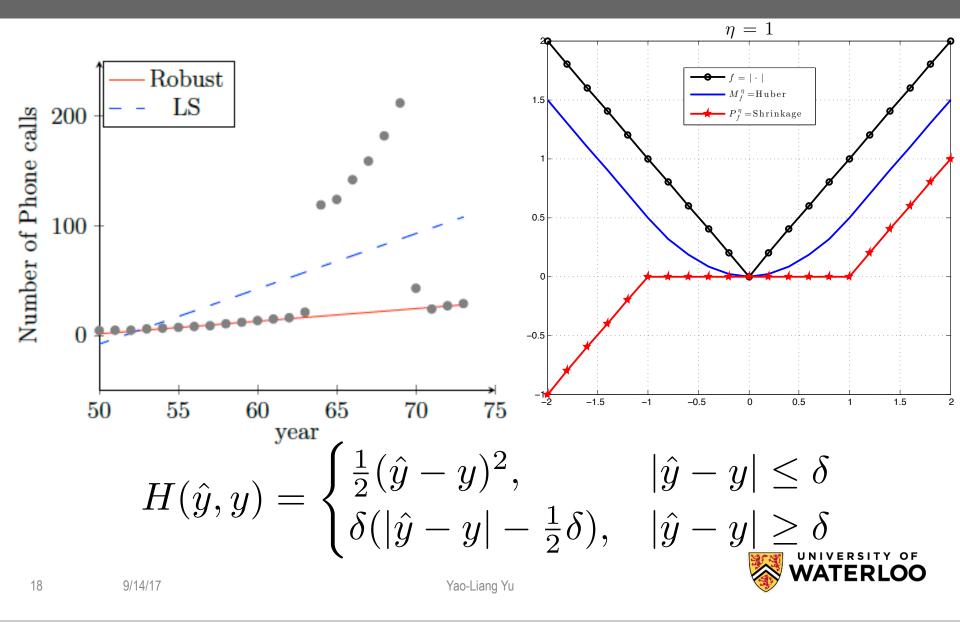
How to evaluate?

$$(Y_{\text{test}} - \hat{Y})^2$$

- Sometimes we evaluate using a different $L(Y_{\mathrm{test}}, \hat{Y})$
 - Leads to a beautiful theory of calibration



Robustness



Gauss vs. Laplace

280

S. PORTNOY AND R. KOENKER

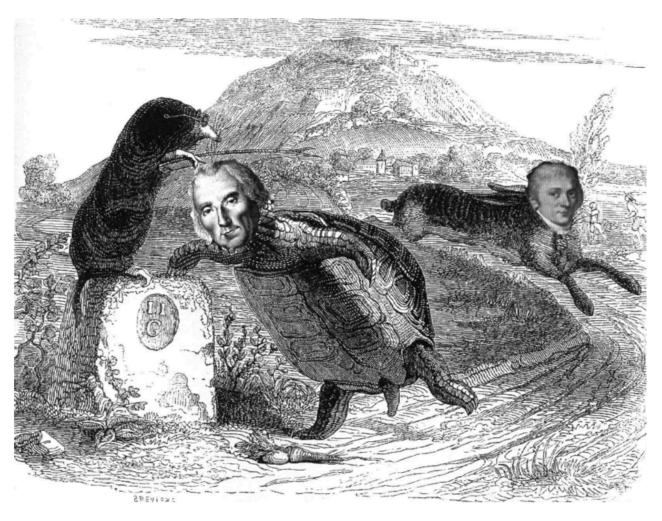


Fig. 1. The Gaussian Hare and the Laplacian Tortoise: this picture is a slightly "retouched" version of a wood engraving by J. J. Grandville from "Fables de La Fontaine" (published in Paris, 1838). The portrait of Gauss is taken from an 1803 portrait by J. C. A. Schwartz. The portrait of Laplace appears in "Cauchy: Un Mathématicien Légitimiste au XIXe Siècle," by Bruno Belhoste (Belin, Paris).

Multi-task learning

$$X^{\top}XW = X^{\top}Y$$

Everything we've shown still holds if Y is m-dim

But, can solve each column of Y independently

$$X^{\top}XW_{:j} = X^{\top}Y_{:j}$$

Things are more interesting if we had regularization



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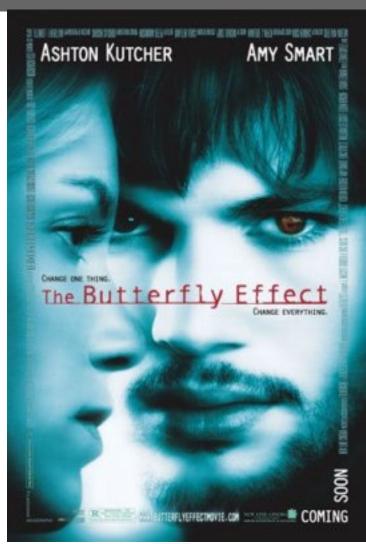


III-posedness

• Let $x_1=0$, $x_2=\epsilon$, $y_1=1$, $y_2=-1$

• w=X-1y=

 Slight perturbation leads to chaotic behaviour





Tiknohov regularization (Hoerl and Kennard'70)

$$\min_{\substack{W \in \mathbf{R}^{(d+1)\times m} \\ \text{Ridge} \\ \text{regression}}} \|XW - Y\|_{\mathsf{F}}^2 + \lambda \|W\|_{\mathsf{F}}^2$$
 Reg. constant (hyperparameter)

 With positive lambda, slight perturbation in input leads to proportional (wrt 1/lambda) perturbation in output



Data augmentation

$$\min_{W \in \mathbf{R}^{(d+1) \times m}} \|XW - Y\|_{\mathsf{F}}^2 + \lambda \|W\|_{\mathsf{F}}^2$$



$$\min_{W \in \mathbf{R}^{(d+1) \times m}} \|\tilde{X}W - \tilde{Y}\|_{\mathsf{F}}^2$$

$$ilde{X} = egin{bmatrix} X \ \sqrt{\lambda}I \end{bmatrix} & ilde{Y} = egin{bmatrix} Y \ \mathbf{0} \end{bmatrix}$$



Sparsity

Ridge regression weight is always dense

- Computationally heavy
- Interpretationally cumbersome
- Lasso (Tibshirani'96)

$$\min_{\|W\|_1 \le C} \|XW - Y\|_{\mathsf{F}}^2$$

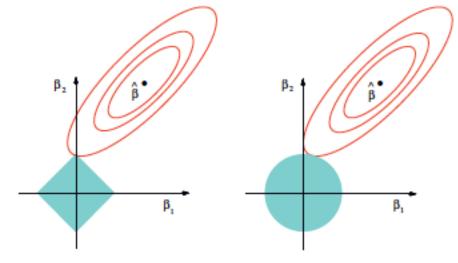
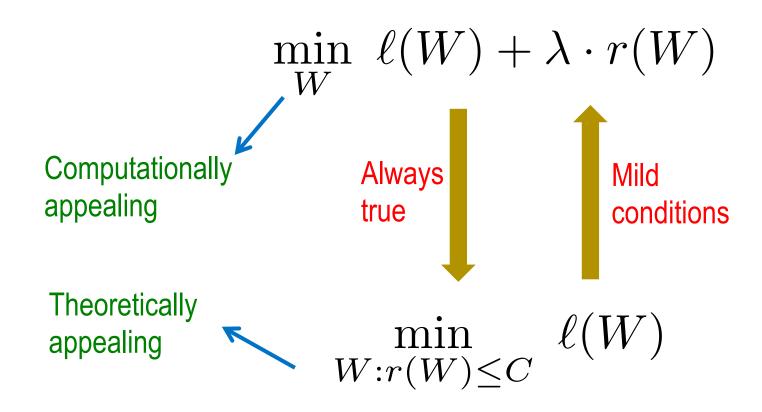


FIGURE 3.11. Estimation picture for the lasso (left) and ridge regression (right). Shown are contours of the error and constraint functions. The solid blue areas are the constraint regions $|\beta_1| + |\beta_2| \le t$ and $\beta_1^2 + \beta_2^2 \le t^2$, respectively, while the red ellipses are the contours of the least squares error function.



Regularization vs. Constraint





26

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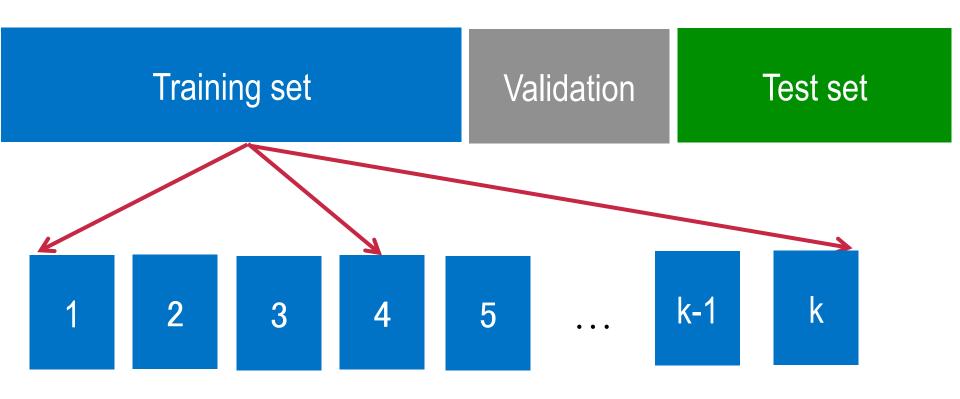
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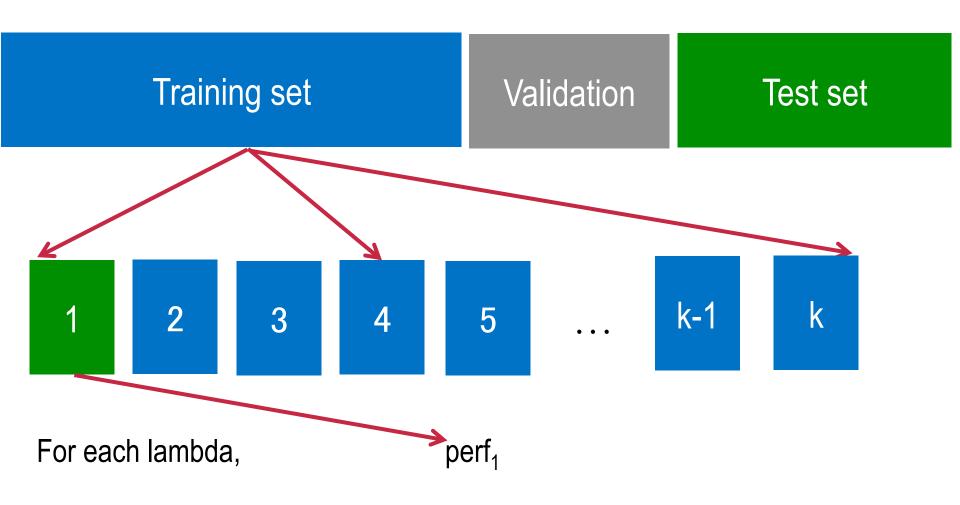
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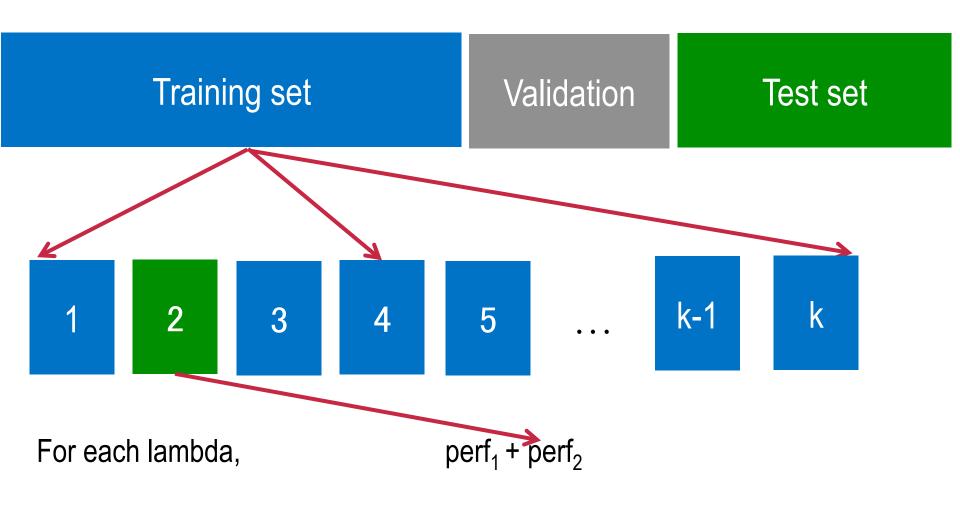




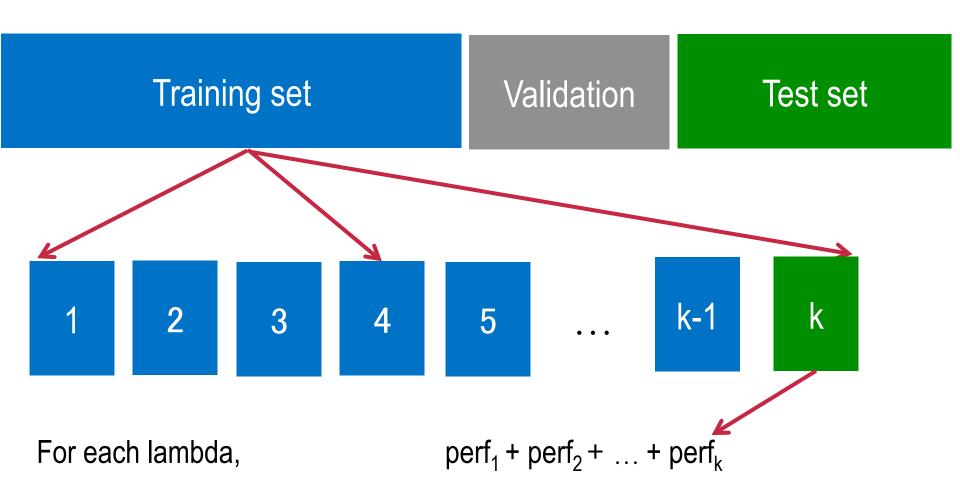




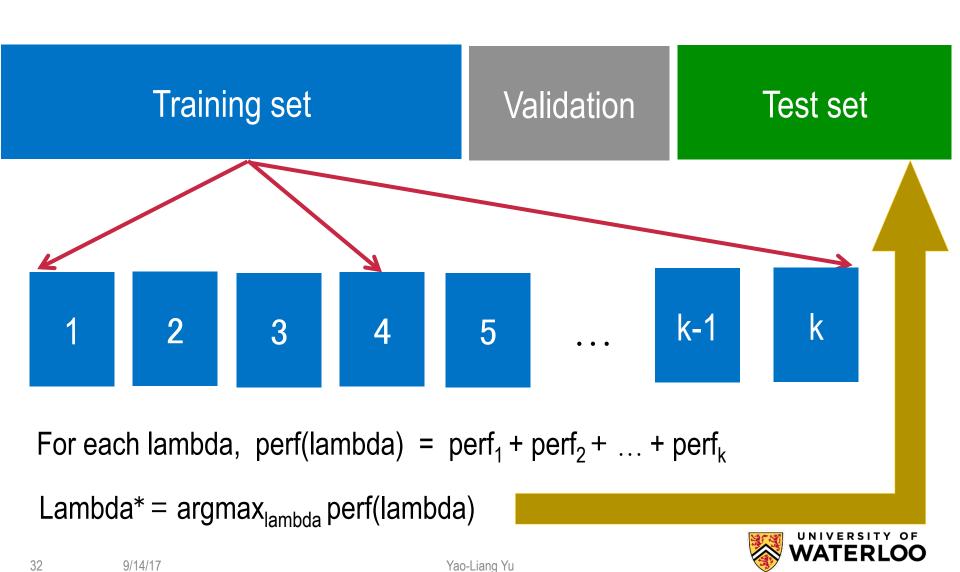












Questions?



