



CS489/698: Intro to ML

Lecture 04: Logistic Regression



Outline

- Announcements
- Bernoulli model
- Logistic regression
- Computation

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Announcements

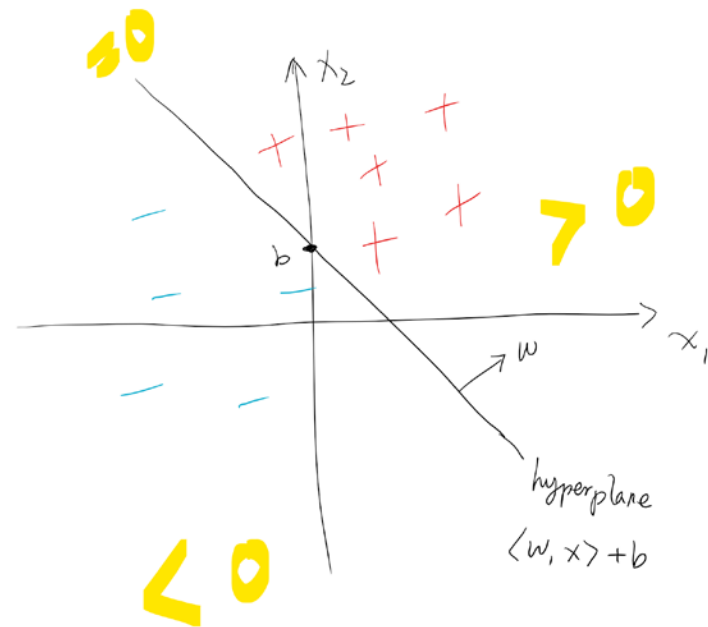
- Assignment 1 **due next Tuesday**

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Classification revisited

- $\hat{y} = \text{sign}(\mathbf{x}^T \mathbf{w} + b)$
- How confident we are about \hat{y} ?
- $|\mathbf{x}^T \mathbf{w} + b|$ seems a good indicator
 - real-valued; hard to interpret
 - ways to transform into $[0,1]$
- **Better(?)** idea: learn confidence directly



Conditional probability


- $P(Y=1 \mid X=\mathbf{x})$: conditional on seeing \mathbf{x} , what is the chance of this instance being positive, i.e., $Y=1$?
 - obviously, value in $[0,1]$
- $P(Y=0 \mid X=\mathbf{x}) = 1 - P(Y=1 \mid X=\mathbf{x})$, if two classes
 - more generally, sum to 1

Notation (Simplex). $\Delta_{k-1} := \{ \mathbf{p} \text{ in } \mathbb{R}^k : \mathbf{p} \geq 0, \sum_i p_i = 1 \}$

Reduction to a harder problem

- $P(Y=1 \mid X=\mathbf{x}) = E(1_{Y=1} \mid X=\mathbf{x})$

$$1_A = \begin{cases} 1, & A \text{ is true} \\ 0, & A \text{ is false} \end{cases}$$

- 
- Let $Z = 1_{Y=1}$, then regression function for (X, Z)
 - use linear regression for binary Z ?
 - Exploit **structure!**
 - conditional probabilities are in a simplex
 - Never reduce to **unnecessarily harder** problem

Bernoulli model

- Let $P(Y=1 \mid X=\mathbf{x}) = p(\mathbf{x}; \mathbf{w})$, parameterized by \mathbf{w}
- Conditional likelihood on $\{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$:

$$\mathbf{P}(Y_1 = y_1, \dots, Y_n = y_n \mid X_1 = \mathbf{x}_1, \dots, X_n = \mathbf{x}_n)$$

- simplifies if independence holds

$$\prod_{i=1}^n \mathbf{P}(Y_i = y_i \mid X_i = \mathbf{x}_i) = \prod_{i=1}^n p(\mathbf{x}_i; \mathbf{w})^{y_i} (1 - p(\mathbf{x}_i; \mathbf{w}))^{1-y_i}$$



- Assuming y_i is $\{0,1\}$ -valued

Naïve solution

$$\prod_{i=1}^n p(\mathbf{x}_i; \mathbf{w})^{y_i} (1 - p(\mathbf{x}_i; \mathbf{w}))^{1-y_i}$$

- Find \mathbf{w} to maximize conditional likelihood
- What is the solution if $p(\mathbf{x}; \mathbf{w})$ does not depend on \mathbf{x} ?
- What is the solution if $p(\mathbf{x}; \mathbf{w})$ does not depend on ?

Generalized linear models (GLM)

- $y \sim \text{Bernoulli}(p)$; $p = p(\mathbf{x}; \mathbf{w})$ natural parameter
 - Logistic regression
- $y \sim \text{Normal}(\boldsymbol{\mu}, \sigma^2)$; $\boldsymbol{\mu} = \boldsymbol{\mu}(\mathbf{x}; \mathbf{w})$
 - (weighted) least-squares regression
- GLM: $y \sim \exp(\boldsymbol{\theta} \boldsymbol{\phi}(y) - A(\boldsymbol{\theta}))$
 -  sufficient statistics
 -  log-partition function

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Logit transform

- $p(\mathbf{x}; \mathbf{w}) = \mathbf{w}^\top \mathbf{x}$ $p \geq 0$ not guaranteed...

- $\log p(\mathbf{x}; \mathbf{w}) = \mathbf{w}^\top \mathbf{x}$ better!

- LHS negative, RHS real-valued...

- **Logit transform** $\log \frac{p(\mathbf{x}; \mathbf{w})}{1 - p(\mathbf{x}; \mathbf{w})} = \mathbf{w}^\top \mathbf{x}$

odds
ratio

- Or equivalently $p(\mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x})}$

Prediction with confidence

$$p(\mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x})}$$

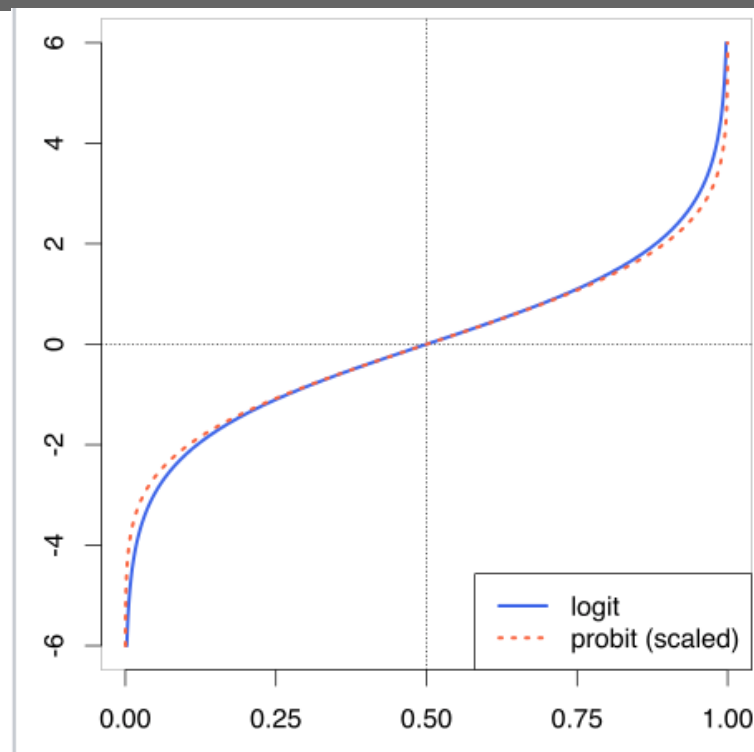
- $\hat{y} = 1$ if $p = P(Y=1 \mid X=x) > 1/2$ iff $\mathbf{w}^\top \mathbf{x} > 0$
- Decision boundary $\mathbf{w}^\top \mathbf{x} = 0$
- $\hat{y} = \text{sign}(\mathbf{w}^\top \mathbf{x})$ as before, but with confidence $p(x; w)$

Not just a classification algorithm

- Logistic regression does more than classification
 - it estimates conditional probabilities
 - under the logit transform assumption
- Having confidence in prediction is nice
 - the price is an assumption that may or may not hold
- If classification is the sole goal, then doing extra work
 - as shall see, SVM only estimates decision boundary

More than logistic regression

- $F(p)$ transforms p from $[0,1]$ to \mathbb{R}
- Then, equating $F(p)$ to a linear function $\mathbf{w}^T \mathbf{x}$
- But, there are many other choices for F !
 - precisely the **inverse of any distribution function**!



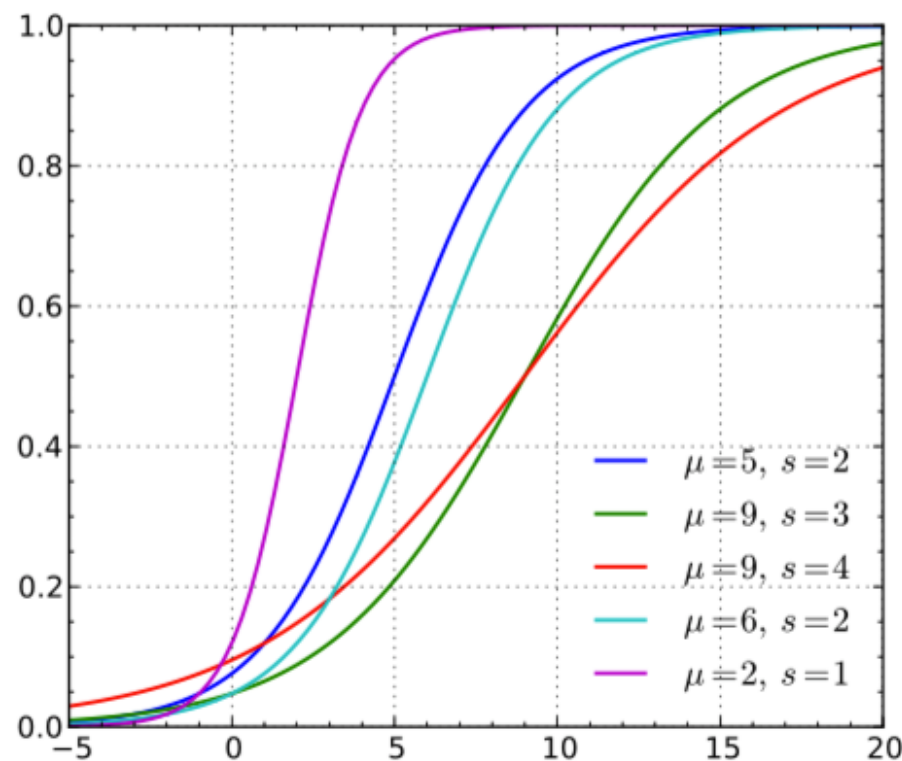
Comparison of the **logit function** with a scaled probit (i.e. the inverse **CDF** of the **normal distribution**), comparing $\text{logit}(x)$ vs. $\Phi^{-1}(x) / \sqrt{\frac{\pi}{8}}$, which makes the slopes the same at the origin.

Logistic distribution

- Cumulative Distribution Function

$$F(x; \mu, s) = \frac{1}{1 + \exp\left(-\frac{x-\mu}{s}\right)}$$

- Mean μ , variance $s^2\pi^2/3$



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Maximum likelihood

$$\prod_{i=1}^n p(\mathbf{x}_i; \mathbf{w})^{y_i} (1 - p(\mathbf{x}_i; \mathbf{w}))^{1-y_i}$$
$$p(\mathbf{x}; \mathbf{w}) = \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x})}$$

- Minimize negative log-likelihood

$$\sum_i \log(e^{(1-y_i)\mathbf{w}^\top \mathbf{x}_i} + e^{-y_i\mathbf{w}^\top \mathbf{x}_i}) \equiv \sum_i \log(1 + e^{-\tilde{y}_i\mathbf{w}^\top \mathbf{x}_i})$$

Newton's algorithm

$$\mathbf{w}_{t+1} \leftarrow \mathbf{w}_t - \eta_t [\nabla^2 f(\mathbf{w}_t)]^{-1} \cdot \nabla f(\mathbf{w}_t)$$

$$\nabla f(\mathbf{w}_t) = X^\top (\mathbf{p} - \mathbf{y})$$

$$\nabla^2 f(\mathbf{w}_t) = \sum_i p_i (1 - p_i) \mathbf{x}_i \mathbf{x}_i^\top$$

PSD

$$p_i = \frac{1}{1 + e^{-\mathbf{w}_t^\top \mathbf{x}_i}}$$

Uncertain predictions get bigger weight

- $\eta = 1$: iterative weighted least-squares

A word about implementation

- Numerically computing exponential can be tricky
 - easily underflows or overflows
- The usual trick
 - estimate the range of the exponents
 - shift the mean of the exponents to 0

Robustness

$$\ell(t) = \log(1 + e^t) \quad L(\hat{y}, y) = \ell(-\hat{y}y) \quad \hat{y} = \mathbf{w}^\top \mathbf{x}$$

- Bounded derivative

$$\ell'(t) = \frac{e^t}{1 + e^t} = \frac{1}{1 + e^{-t}}$$

- Variational exponential

Larger exp loss gets smaller weights

$$\log(1 + e^t) = \min_{0 \leq \eta \leq 1} \boxed{\eta e^t} - \log(\eta) + \eta - 1$$

More than 2 classes

- Softmax

$$\mathbf{P}(Y = c | \mathbf{x}, W) = \frac{\exp(\mathbf{w}_c^\top \mathbf{x})}{\sum_{q=1}^k \exp(\mathbf{w}_q^\top \mathbf{x})}$$

- Again, nonnegative and sum to 1

- Negative log-likelihood (y is one-hot)

$$-\log \prod_{i=1}^n \prod_{c=1}^k p_{ic}^{y_{ic}} = -\sum_{i=1}^n \sum_{c=1}^k y_{ic} \log p_{ic}$$

Questions?

