Lecture A2. Probability Review

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I. Probability

Probability

Definition 1 (probability)

사건 0과 1사이의 실수

Probability is a function from an event to a real number between 0 and 1.

$$\mathbb{P} : S \to [0, 1],$$

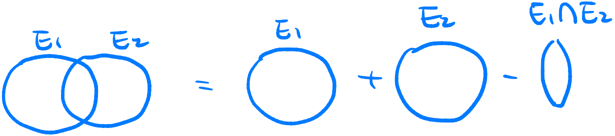
where S is the set of all possible events.

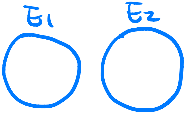
Remark 1

- For any event E, $0 \leq \mathbb{P}(E) \leq 1$. ✓
- For $E = \emptyset$, $\mathbb{P}(E) = 0$. ✓
- $\mathbb{P}(S) = 1$, where S is whole space, or a space for all possible events.

Remark 2

- $\mathbb{P}(E_1 \cup E_2) = \mathbb{P}(E_1) + \mathbb{P}(E_2) - \mathbb{P}(E_1 \cap E_2)$.
- If $E_1 \cap E_2 = \emptyset$, then $\mathbb{P}(E_1 \cup E_2) = \mathbb{P}(E_1) + \mathbb{P}(E_2)$. (mutually exclusive events)





Conditional Probabilities

Definition 2 (conditional probability)

$\mathbb{P}(E|F)$, reads as **probability of E given F**, is the probability that event E occurs given that F has occurred. In a math notation,

$$\mathbb{P}(E|F) := \frac{\mathbb{P}(E \cap F)}{\mathbb{P}(F)}$$

- Ex) (from Ross) Suppose cards numbered one through ten are placed in a hat, mixed up, and then one of the cards is drawn. If we are told that the number on the drawn card is at least five, then what is the conditional probability that it is ten?

$\geq 5 \downarrow F$ $\qquad$ $\downarrow E$

$$\checkmark\ \mathbb{P}(E|F) = \frac{\mathbb{P}(E\cap F)}{\mathbb{P}(F)} = \frac{1/10}{6/10} = \frac{1}{6}$$

$$\checkmark\ P(E) = \frac{1}{10}$$

Independent Probability

Definition 3 (independence)

If $\mathbb{P}(E) = \mathbb{P}(E|F)$, then we say the events E and F are **independent** events.

- That is, the occurence of E happening has nothing to do with the occurence of F.
- Just as Galilei once said *"And yet it rotates (even though people do not believe so)"*. People believing whether or not the earth rotates is independent event of earth's rotation.

Properties

E F

1. $\mathbb{P}(E_1 \cup E_2) = \mathbb{P}(E_1) + \mathbb{P}(E_2)$, if $E_1 \cap E_2 = \emptyset$.
2. $\mathbb{P}(E) = \mathbb{P}(E \cap F) + \mathbb{P}(E \cap F^c)$ $(\because (E \cap F) \cap (E \cap F^c) = \emptyset)$
3. If $F_i \cap F_j = \emptyset$ for all $i \neq j$ and $\mathbb{P}(F_1 \cup \cdots \cup F_n) = 1$ where $1 \leq i, j \leq n$, then $\mathbb{P}(E) = \sum_{i=1}^{n} \mathbb{P}(E \cap F_i)$

$\cup$

F_1	F_2	$\cdots$	$\cdots$	F_n

partition

Definition 4 (probabilitisic partition)

We call the series of $\{F_i, 1 \leq i \leq n\}$ as a **probabilitisic partition** if $F_i \cap F_j = \emptyset$ for all $i \neq j$ and $\mathbb{P}(F_1 \cup \cdots \cup F_n) = 1$

Remark 3

If $F_1, \ldots, F_n$ is a probabilitisic partition, it means that exactly one and only one events among $F_1, \ldots, F_n$ will occur.

Bayes' rule

$\{F_i, 1 \le i \le n\}$

Theorem 1 (Bayes' rule)

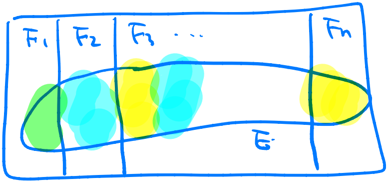
Suppose $F_1, \ldots, F_n$ is probabilitisic partition then the following holds.

$$\begin{aligned}\mathbb{P}(E) &= \mathbb{P}(E \cap F_1) + \cdots + \mathbb{P}(E \cap F_n) \\ &= \sum_{i=1}^{n} \mathbb{P}(E \cap F_i) \\ &= \sum_{i=1}^{n} \mathbb{P}(E|F_i)\mathbb{P}(F_i)\end{aligned}$$

$$\mathbb{P}(E|F_i) = \frac{\mathbb{P}(E \cap F_i)}{\mathbb{P}(F_i)}$$

$$\Downarrow$$

$$\mathbb{P}(E \cap F_i) = \mathbb{P}(E|F_i)\mathbb{P}(F_i)$$



Exercise 1

Show that $\mathbb{P}(A|B \cap C)\mathbb{P}(B|C) = \mathbb{P}(A \cap B|C)$. *(You should either derive RHS from LHS or derive LHS from RHS)*

II. Random variables

discrete vs continuous

Definition 5 (discrete random variable)

A random variable X is called a **discrete random variable** if it has a countable number of possible values.

- Example
  - flip a coin $\{H, T\}$
  - throw a dice $\{1, 2, 3 \cdots 6\}$ > finite
  - number of consecutive head in coin toss $\{1, 2, 3 \cdots \infty\}$ – infinite

Definition 6 (continuous random variable)

A random variable X is called a **continuous random variable** if it takes all values in given interval of numbers.

- Example
  - weights and heights of a person
  - temperature of a room

pmf and pdf

Definition 7 (probability mass function)

대문자 (r.v.)

A **pmf** $p(x)$ is a function that gives the *probability* that a discrete r.v. is exactly equal to some value.

prob. 소문자

$$\mathbb{P}(X = x) = p(x)$$

- Suppose you throw a dice and X be a r.v. of the outcome. What is the pmf of X?
  1. *answer in a math form*:
  2. *answer in a tabular form*:

$$p(x) = \mathbb{P}(X = x) = \begin{cases} \frac{1}{6} & \text{for } x = 1, 2, \cdots 6 \\ 0 & \text{otherwise} \end{cases}$$

x	1	2	3	4	5	6
$\mathbb{P}(X=x)$	1/6	1/6	1/6	1/6	1/6	1/6

$X \sim N(0,1) \longrightarrow f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}$

$\mathbb{P}(X=0) = 0 = \mathbb{P}(0 \le x \le 0)$

$= \int_0^0 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} dx$

$= 0$

Definition 8 (probability density function)

A **pdf** $f(x)$ is a function that gives the *relative likelihood* for this continuous r.v. to take on a given value.

$$\mathbb{P}(a \le X \le b) = \int_a^b f(x)dx$$

- Suppose a random variable takes a real values between 0 and 1 with equal likelihood, what is the pdf of X?
  1. *answer in a math form*:
  2. *draw a graph*:

Properties of pmf and pdf

- Mathematical property? *Something that always happen.*

Remark 4

The functions (pmf and pdf) are nonnegative everywhere. That is,

$$p(x) \geq 0 \text{ and } f(x) \geq 0$$

Remark 5

Its summation or integral over the entire area is equal to one. That is,

$$\sum_{x=-\infty}^{\infty} p(x) = 1 \text{ and } \int_{-\infty}^{\infty} f(x)dx = 1$$

- One can derive a cdf (cumulative distribution function) from a pmf or from a pdf. (See the upcoming definition of cdf.)

pmf → cdf
pdf → cdf

Expectation

pmf, pdf → $\mathbb{E}X$

Remark 6

For a discrete random variable X with pmf $p(x)$,

- $\mathbb{E}X = \sum_{x=-\infty}^{\infty} xp(x)$
- $\mathbb{E}[g(X)] = \sum g(x)p(x)$
- Ex) $\mathbb{E}X^2 = \sum x^2 p(x)$
- Ex) $\mathbb{E}(X^2 - 2X) = \sum (x^2 - 2x)p(x)$

$Ex^2 = \sum \boxed{x^2} p(x)$

Remark 7

For a continuous random variable X with pdf $f(x)$,

- $\mathbb{E}X = \int_{-\infty}^{\infty} xf(x)dx$
- $\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx$
- Ex) $\mathbb{E}X^2 = \int_{-\infty}^{\infty} x^2 f(x)dx$

cdf

Definition 9 (cumulative distribution function)

For a random variable X, the cdf(cumulative distribution function) $F(x)$ is a function of probability that the random variable X is found at a value less than or equal to x.

$$F(x) = \mathbb{P}(X \leq x)$$

- If discrete,

pmf에 정의.

$$F(x) := \mathbb{P}(X \leq x) = \sum_{y=-\infty}^{y=x} \mathbb{P}(X = y) = \sum_{y=-\infty}^{x} p(y)$$

- If continuous,

$$F(x) := \mathbb{P}(X \leq x) = \int_{-\infty}^{x} f(y)dy$$

$$= \mathbb{P}(-\infty \leq X \leq x)$$

III. Uniform distribution

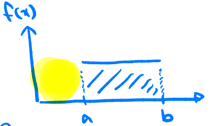
(일양분포)

Definition 10 (uniform distribution)

$\sim$: \sim

A continuous random variable X is said to follow uniform distribution with parameter a and b, and write $X \sim U(a, b)$ (*reads "X follows a uniform distribution with parameter a and b."*) if

$$\text{pdf } f(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \le x \le b \\ 0 & \text{otherwise} \end{cases}$$



$\sim U(a,b)$

$F(x) = \mathbb{P}(X \le x) = \int_{-\infty}^{x} f(u)\,du = *$

i) $x \le a$

$* = \int_{-\infty}^{x} f(u)\,du = \int_{-\infty}^{x} 0 \cdot du = 0$

ii) $a \le x \le b$

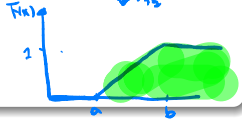
$* = \int_{-\infty}^{a} f(x)\,dx + \int_{a}^{x} f(x)\,dx = \int_{-\infty}^{a} 0\,dx + \int_{a}^{x} \frac{1}{b-a}\,dx = \frac{x-a}{b-a}$

$\int_{-\infty}^{\infty} f(x) = 1$. $\sum_{x=-\infty}^{\infty} p(x) = 1$

$$\text{cdf } F(x) = \begin{cases} 0 & \text{if } x \le a \\ \frac{x-a}{b-a} & \text{if } a \le x \le b \\ 1 & \text{if } x > b \end{cases}$$

$\parallel$

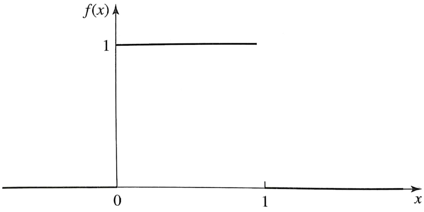
$\mathbb{P}(X \le x)$



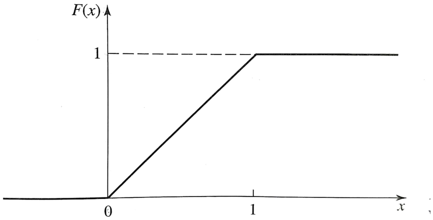
- (*The value $\frac{1}{b-a}$ is chosen as a constant so that it will make integrations over $-\infty$ to ∞ to be equal to 1.*)

$U(0,1)$

- pdf



- cdf



Exercise 2

$X \sim U(10, 20)$, *then what is* $F(10)$ *and* $F(15)$?

$$F(x) = \begin{cases} 0 & \text{if } x < 10 \\ \frac{x-10}{10} & \text{if } 10 \leq x < 20 \\ 1 & \text{if } x \geq 20 \end{cases}$$

$F(10) = \frac{10-10}{10} = 0$

$F(15) = \frac{15-10}{10} = 0.5.$

IV. Exponential distribution

Definition 11

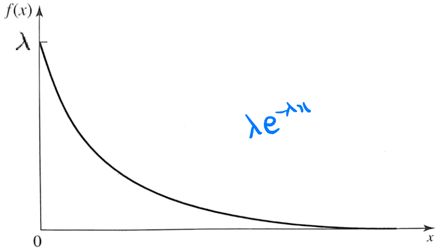
A nonnegative continuous random variable X is said to follow exponential distribution with parameter λ and write $X \sim exp(\lambda)$, if

$$\text{pdf } f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

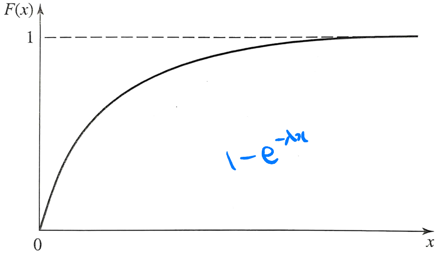
$$\text{cdf } F(x) = \begin{cases} 1 - e^{-\lambda x}, & \text{if } x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$exp(\lambda)$

- pdf



- cdf



- TFAE (The followings are all equivalent - *If any one of them is true, the rest of them are all true. If any one of them is false, then the rest of them are all false.*)
  1. $X \sim exp(\lambda)$
  2. pdf $f(x) = \lambda e^{-\lambda x}$, if $x \geq 0$; $f(x) = 0$, otherwise
  3. cdf $F(x) = 1 - e^{-\lambda x}$, if $x \geq 0$; $F(x) = 0$, otherwise

Exercise 3

Prove that pdf $\longrightarrow$ cdf

$F(x) = \mathbb{P}(X \leq x)$

i) $x < 0$, $F(x) = \mathbb{P}(X \leq x) = \int_{-\infty}^{x} f(x)\, dx = \int_{-\infty}^{x} 0\, dx = 0$

ii) $x \geq 0$, $F(x) = \mathbb{P}(X \leq x) = \int_{-\infty}^{x} f(x)\, dx = \int_{-\infty}^{0} f(x)\, dx + \int_{0}^{x} f(x)\, dx$

$= \int_{-\infty}^{0} 0\, dx + \int_{0}^{x} \lambda e^{-\lambda x}\, dx$

$= \dots$

$= 1 - e^{-\lambda x}$.

More statistics (통계량) for a r.v.

- Variance: $Var(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2$
- Standard Deviation: $sd(X) = \sqrt{Var(X)}$
- Coefficient of Variation (cv): $cv(X) = \frac{sd(X)}{\mathbb{E}X}$ >
- (*A cv measures 'relative' variation and often more meaningful.*)
- Examples
  1. cv of $N(10, \sqrt{10}^2)$ vs $N(10000, 10^2)$.
  2. For an exponential r.v., $X \sim exp(\lambda)$, then $cv(X) =$ (*next slide*)
  3. For a normal r.v. $X \sim N(\mu, \sigma^2)$, $cv(X) =$
  4. For deterministic variable X, ~~c_x~~ $X =$

$cv(X) = \frac{sd(X)}{\mathbb{E}X} = \frac{0}{\mathbb{E}X} = 0$

Properties of exponential distribution

- Statistics for a r.v. $X \sim exp(\lambda)$,
  1. $EX = 1/\lambda$
  2. $Var(X) = 1/\lambda^2$ $sd(X) = \sqrt{Var(X)} = 1/\lambda$
  3. $cv(X) = 1$ *(Exponential r.v. has c.v. equal to 1, always.)*

     $= \frac{sd(X)}{EX} = \frac{1/\lambda}{1/\lambda} = 1$
- Theorems
  1. Memoryless property
  2. Suppose $X_1 \sim exp(\lambda_1)$, $X_2 \sim exp(\lambda_2)$, and independent, then $\mathbb{P}(X_1 < X_2) = \frac{\lambda_1}{\lambda_1+\lambda_2}$
  3. If $X_1 \sim exp(\lambda_1)$, $X_2 \sim exp(\lambda_2)$, and independent, then $min(X_1, X_2) \sim exp(\lambda_1 + \lambda_2)$

Exercise 4

Show that $EX = 1/\lambda$

$$\mathbb{E}X = \int_{-\infty}^{\infty} f(x)\,dx$$

$$= \int_{-\infty}^{0} f(x)\,dx + \int_{0}^{\infty} f(x)\,dx$$

$$= \int_{-\infty}^{0} 0\,dx + \int_{0}^{\infty} \lambda e^{-\lambda x}\,dx$$

$$= \ldots$$

$$= \ldots$$

$$= 1/\lambda.$$

Exercise 5

Show that $Var(X) = 1/\lambda^2$. *(Hint: need to do* EX^2 *first)*

$Var(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2$

$\mathbb{E}X^2 = \int_{-\infty}^{\infty} x^2 f(x)\, dx = \int_{-\infty}^{0} x^2 \underbrace{f(x)}_{0}\, dx + \int_{0}^{\infty} x^2 \underbrace{f(x)}_{\lambda e^{-\lambda x}}\, dx$

$= \cdots$

$= \cdots$

$= \cdots$

$(\mathbb{E}X)^2 = \frac{1}{\lambda^2}$

Common usage of exponential distribution

- Since exponential r.v. is continuous and nonnegative, it is frequently used to describe time.
- In such cases, the parameter λ can be understood as "rate" as an inverse value of the time.
- For example, $X \sim exp(\overset{0.2}{\lambda})$ and $\mathbb{E}X = 1/\lambda = 5$ yrs imply $\lambda = 1/5$ per year.
- For example, consider the series of events that occur in every time interval that follows exponential distribution with mean 3 minutes. (In post office, you expect one customer arrival in 3 minutes on average and the time follows exponential distribution). We say that it follows exponential arrival times with rate 1/3 per minute.

3분에 1명 = 1분에 1/3 명

Memoryless property

① $\mathbb{P}(X>30 \mid X>20)$
② $\mathbb{P}(X>80 \mid X>70)$

$\mathbb{P}(X>30 \mid X>10) \neq \mathbb{P}(X>20)$

1) 전구. ✓
2) 거북이
3) 버스 ✓
4) 전화 ✓
5) 지나가는 [illegible]

Definition 12

시간

A r.v. X is *memoryless*, if $\mathbb{P}(\underbrace{X > s + t}_{0.5+2} | \underbrace{X > t}_{2}) = \mathbb{P}(\underbrace{X > s}_{0.5})$, for $s, t \geq 0$.

Theorem 2

✓ *Exponential random variable is memoryless.*

$\mathbb{E}X = 5\text{min.}$

1/5 per minutes.

- For example, we shall assume that bus arrival time follows $exp(1/5)$. In other words, it follows an exponential distribution and its expected arrival time is 5 minutes. You are waiting for bus, and have been waiting for 3 minutes, what is the probability that bus will not come in 5 minutes from now.

$$\mathbb{P}(X>5+3 \mid X>3) = \mathbb{P}(X>5)$$

memoryless property

Exercise 6

Prove the previous theorem.

Let $X \sim \exp(\lambda)$, then its CDF is $F(x) = 1 - e^{-\lambda x}$ for $x \geq 0$

Claim, $\mathbb{P}(X > s+t \mid X > t) = \mathbb{P}(X > s)$

$$(\text{LHS}) = \frac{\mathbb{P}(X > s+t, X > t)}{\mathbb{P}(X > t)} = \frac{\mathbb{P}(X > s+t)}{\mathbb{P}(X > t)} = \cdots = \mathbb{P}(X > s)$$

Theorem 3

Suppose $X_1 \sim exp(\lambda_1), X_2 \sim exp(\lambda_2)$ *and they are independent, then* $\mathbb{P}(X_1 < X_2) = \frac{\lambda_1}{\lambda_1+\lambda_2}$

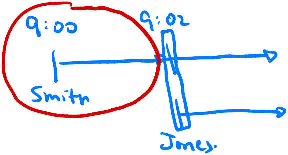
- Smith and Jones came to post office together and they are served by two clerks, A and B, respectively. Server A has service time following $exp(1/3)$ and server B has service time following $exp(1/5)$. What is the chance that Smith will be done first?

$X_1 \sim exp(1/3)$
$X_2 \sim exp(1/5)$

$\mathbb{P}(X_1 < X_2) = \frac{\lambda_1}{\lambda_1+\lambda_2} = \frac{1/3}{1/3+1/5} = \frac{5}{8}$

X_1 Jones X_2

- Suppose that Smith came to post office earlier than Jones by 2 minutes, but Smith was still being served at the moment that Jones started to being served. Would this assumption change your previous answer? Why or why not?



NO. because of exp. dist's memoryless property!

V. Poisson distribution

$X \sim Poi(\lambda)$

Definition 13

nonnegative

A discrete random variable X is said to follow Poisson distribution with parameter λ, and write $X \sim poi(\lambda)$, if its pmf is

$$\mathbb{P}(X = k) = \frac{\lambda^k e^{-\lambda}}{k!} \text{ for } k = 0, 1, 2, \ldots$$

$0! = 1$
$1! = 1$
$2! = 2$
$3! = 3 \times 2 \times 1 = 6$

- What is cdf of $poi(\lambda)$?

$$\mathbb{P}(X \le k) = \sum_{i=0}^{k} \mathbb{P}(X = i) = \sum_{i=0}^{k} \frac{\lambda^i e^{-\lambda}}{i!}, \text{ for } k = 0, 1, \ldots$$

Remark 8

$$\mathbb{E}X = Var(X) = \lambda$$

Exercise 7

For $X \sim poi(\lambda)$, prove that $\mathbb{E}X = \lambda$.

- cf) $e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \cdots = \sum_{n=0}^{\infty} \frac{x^n}{n!}$ ✓
- pf) We have $\mathbb{P}(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$ for $k = 0, 1, 2, \ldots$, and

$$\mathbb{E}X = \sum_{x=-\infty}^{\infty} xp(x) \text{ (this is common for all discrete r.v.)}$$

$$= \sum_{x=-\infty}^{-1} xp(x) + \sum_{x=0}^{\infty} xp(x)$$

$$= 0 + \sum_{x=0}^{\infty} x \cdot \frac{\lambda^x e^{-\lambda}}{x!} = \sum_{x=1}^{\infty} x \cdot \frac{\lambda^x e^{-\lambda}}{x!}$$

$$= 1 \cdot \frac{\lambda e^{-\lambda}}{1!} + \sum_{x=2}^{\infty} \frac{\lambda^x e^{-\lambda}}{(x-1)!} = \cdots = \lambda$$

Your work.

Theorem 4 (Exponential Dist Section에 들어가야 함)

If $X_1 \sim exp(\lambda_1)$, $X_2 \sim exp(\lambda_2)$, *and independent, then*
$min(X_1, X_2) \sim exp(\lambda_1 + \lambda_2)$

Exercise 8

Prove the above theorem.

$$\mathbb{P}(\min(X_1, X_2) \le x) = 1 - \mathbb{P}(\min(X_1, X_2) > x)$$
$$= 1 - \mathbb{P}(X_1 > x, X_2 > x)$$
$$= 1 - \mathbb{P}(X_1 > x)\,\mathbb{P}(X_2 > x)$$

cf) $\mathbb{P}(X_1 \le x) = 1 - e^{-\lambda_1 x}$

$\mathbb{P}(X_2 \le x) = 1 - e^{-\lambda_2 x}$

VI. Some exercises

Maximum and minimum

Definition 14

$$x \wedge y = min(x, y) = \begin{cases} x & \text{, if } x \leq y \\ y & \text{, otherwise} \end{cases}$$

$$x \vee y = max(x, y) = \begin{cases} x & \text{, if } x \geq y \\ y & \text{, otherwise} \end{cases}$$

$$x^+ = max(x, 0) \text{ (positive part)}$$

"positive part of x"

- Ex) $x \wedge 25 = min(x, 25)$
- Ex) $(25 - x)^+ = max(25 - x, 0)$

Exercise 9

For $X \sim U(20, 40)$*, evaluate* $\underline{\mathbb{E}[X \wedge 25]}$ *and* $\underline{\mathbb{E}[(25 - X)^+]}$.

$$\mathbb{E}[(25-X)^+] = \int_{-\infty}^{\infty} (25-x)^+ f(x)\, dx$$

$$= \int_{-\infty}^{20} (25-x)^+ \cancel{f(x)}^{\,0}\, dx + \int_{20}^{40} (25-x)^+ \underbrace{f(x)}_{\frac{1}{40-20}}\, dx + \int_{40}^{\infty} (25-x)^+ \cancel{f(x)}^{\,0}\, dx$$

"Divide & conquer"

$$= \int_{20}^{40} \max(25-x, 0) \cdot \frac{1}{20}\, dx$$

$$= \int_{20}^{25} \max(25-x, 0) \frac{1}{20}\, dx + \int_{25}^{40} \underline{\max(25-x, 0)} \frac{1}{20}\, dx$$

$$= \int_{20}^{25} (25-x) \frac{1}{20}\, dx + \int_{25}^{40} 0 \cdot \frac{1}{20}\, dx$$

$$= \ldots$$

Exercise 10

$p(x) = \mathbb{P}(X = x) = \frac{\lambda^x e^{-\lambda}}{x!}$

For $X \sim Poi(8)$,

1. $\mathbb{P}(X = 0) =$ $\frac{8^0 e^{-8}}{0!} = e^{-8}$
2. $\mathbb{P}(2 \leq X \leq 4) =$
3. $\mathbb{P}(X > 2) =$

Exercise 11

For $X \sim exp(7)$, *evaluate* $\mathbb{E}[max(X,7)]$.

$$\mathbb{E}[\max(X,7)] = \int_{-\infty}^{\infty} \max(x,7) f(x)\, dx$$

$$= \int_{-\infty}^{0} \max(x,7) \cancelto{0}{f(x)}\, dx + \int_{0}^{\infty} \underbrace{\max(x,7)} f(x)\, dx$$

$$= \int_{0}^{7} \max(x,7)\, 7e^{-7x}\, dx + \int_{7}^{\infty} \max(x,7)\, 7e^{-7x}\, dx$$

$$= \int_{0}^{7} 7 \cdot 7e^{-7x}\, dx + \int_{7}^{\infty} x \cdot 7e^{-7x}\, dx$$

Exercise 12

For $X \sim exp(8)$, *find* x^* *such that* $F(x^*) = \underline{0.6}$.

CDF $F(x) = \begin{cases} 1 - e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$

$$F(x^*) = 1 - e^{-\lambda x^*} = 0.6$$

$$e^{-\lambda x^*} = 0.4$$

$$-\lambda x^* = \log 0.4$$

$$x^* = -\frac{1}{8} \log 0.4.$$

Exercise 13

For $X \sim U(10, 20)$*, find* x^* *such that* $F(x^*) = 0.7$.

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"Man can learn nothing unless he proceeds from the known to the unknown. - Claude Bernard"
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